

Answers (mock exam)

1.

$$Y_n \sim \text{Poisson}\left(\frac{1}{\sqrt{n}}\right)$$
$$Y_n \xrightarrow{L_2} ?$$

First we need to decide where it converges. Let's try $E[Y_n]$

$$E[Y_n] = \frac{1}{\sqrt{n}}$$
$$\lim_{n \rightarrow \infty} E[Y_n] = \lim_{n \rightarrow \infty} \frac{1}{\sqrt{n}} = 0$$

Does it converge to a δ_0 ? Let's apply the definition

$$\lim_{n \rightarrow \infty} E[(Y_n - 0)^2] = \lim_{n \rightarrow \infty} E[Y_n^2]$$

The second moment of a Poisson is

$$\lambda = E[Y^2] - E[Y]^2$$
$$\lambda = E[Y^2] - \lambda^2 \Rightarrow E[Y^2] = \lambda + \lambda^2 \Rightarrow E[Y_n^2] = \frac{1}{\sqrt{n}} + \frac{1}{n}$$
$$\lim_{n \rightarrow \infty} \frac{1}{\sqrt{n}} + \frac{1}{n} = 0 \Rightarrow Y_n \xrightarrow{L_2} \delta_0$$

2. $X_1 \sim N(0, 2)$ and $X_2 \sim N(-2, 1)$ and covariance -1. Compute $COV(X_1 + X_2, X_1 - X_2)$.

$$\begin{aligned} COV(X_1 + X_2, X_1 - X_2) &= \\ &= COV(X_1, X_1) - \cancel{COV(X_1, X_2)} + \cancel{COV(X_2, X_1)} - COV(X_2, X_2) = \\ &= VAR(X_1) - VAR(X_2) = 2 - 1 = 1 \end{aligned}$$

3. $X \sim Ber(\frac{1}{2})$

$$Y = \frac{1}{2} + \frac{X}{2}$$

$$M_X(t) = 1 - p + pe^t$$

$$M_{aX+b}(t) = e^{bt}M_x(at) \Rightarrow M_Y(t) = e^{\frac{t}{2}}M_X(\frac{t}{2}) = e^{\frac{t}{2}}(1 - p + pe^{\frac{t}{2}})$$

Now if

$$p = \frac{1}{2} \Rightarrow M_Y(t) = e^{\frac{t}{2}}\left(\frac{1}{2} + \frac{e^{\frac{t}{2}}}{2}\right) = \frac{1}{2}\left(e^{\frac{t}{2}} + e^t\right)$$

4. $f_X(x) = \frac{x}{2}$, $X \in [0, 2]$

$$F_Y(y) = ?, Y = 6X - 3$$

$$\begin{aligned} F_Y(y) &= \left| \frac{\partial}{\partial y} g^{-1}(y) \right| f_X(g^{-1}(y)) \\ g(X) &= 6X - 3 \quad g^{-1}(Y) = \frac{Y + 3}{6} \\ F_Y(y) &= \frac{1}{6} \left(\frac{Y + 3}{6 * 2} \right) = \frac{1}{6} \left(\frac{Y + 3}{12} \right) \end{aligned}$$

5. $X \sim U(\theta - 2, \theta + 1)$

$$T(X) = \bar{X} + \frac{1}{2}.$$

A sufficient condition is that

$$\lim_{n \rightarrow \infty} MSE(T_n(X)) = 0$$

We need

$$\begin{aligned} E[T_n(X)] &= E\left[\bar{X} + \frac{1}{2}\right] = E[\bar{X}] + \frac{1}{2} = \\ &= \frac{\theta - 2 + \theta + 1}{2} + \frac{1}{2} = \frac{2\theta}{2} - \frac{1}{2} + \frac{1}{2} = \theta \end{aligned}$$

$$V\left[\bar{X} + \frac{1}{2}\right] = \frac{V(X)}{n} = \frac{1}{12n}(\theta + 1 - \theta + 2)^2 = \frac{9}{12} \frac{1}{n} = \frac{3}{4n}$$

from which $B(T_n(X)) = 0$ and $\lim_{n \rightarrow \infty} V(T_n(X)) = 0$

6. $f(X, \theta) = \theta x^{\theta-1}$, for $0 < X < 1$, ($X_1 = 0.2, X_2 = 0.5, X_3 = 0.7, X_4 = 0.8$).

We need $E(X)$

$$\begin{aligned} E(X) &= \int_0^1 x \theta x^{\theta-1} dx = \int_0^1 x^{\frac{1}{x}} \theta x^{\theta} dx = \theta \int_0^1 x^{\theta} dx = \\ &= \theta \left[\frac{x^{\theta+1}}{\theta+1} \right]_0^1 = \frac{\theta}{\theta+1} \end{aligned}$$

$$\bar{X} = \frac{0.2 + 0.5 + 0.7 + 0.8}{4} = 0.55$$

$$\Rightarrow \frac{\theta}{\theta+1} = 0.55$$

$$\theta = 0.55\theta + 0.55$$

$$0.45\theta = 0.55$$

$$\hat{\theta} = 1.22$$

$$7. Y_n \sim \text{Exp}\left(\frac{n}{2}\right)$$

Find n such that $\Pr\{Y_n > 0.25\} \leq 0.80$.

According to the Markov Inequality

$$P(X \geq \lambda E[X]) \leq \frac{1}{\lambda}$$

Now we assume $\frac{1}{\lambda} = 0.8 \Rightarrow \lambda = 1.25$

$E[X] = \frac{2}{n}$ We can write the Markov inequality as

$$P(X \geq 1.25 \frac{2}{n}) \leq 0.8$$

$$P(X \geq \frac{2.5}{n}) \leq 0.8$$

$$\frac{2.5}{n} = 0.25$$

$$n = 10$$

$$8. X = (X_1, X_2)^T, \quad f(x_1, x_2) = kx_2, \quad 0 < x_1 < x_2 < 1$$

In order to compute k

$$\int_0^1 \int_0^{x_2} kx_2 dx_1 dx_2 = 1$$

$$\begin{aligned}\int_0^1 \int_0^{x_2} kx_2 dx_1 dx_2 &= k \int_0^1 x_2 \int_0^{x_2} 1 dx_1 dx_2 = k \int_0^1 x_2 [x_1]_0^{x_2} dx_2 = \\ &= k \int_0^1 x_2^2 dx_2 = k \left[\frac{x_2^3}{3} \right]_0^1 = \frac{k}{3}\end{aligned}$$

$$\frac{k}{3} = 1, \quad k = 3$$

9.

$$\xrightarrow{L_4} \xrightarrow{\text{a.s.}}$$