# EL-GY 6063: Information Theory Lecture 11

May 14, 2019

#### 1 Rate-Distortion

Reference: Chapter 10, Elements of Information Theory.

**Definition 1.** A distortion function is a mapping  $d: \mathcal{X} \times \mathcal{X} \to \mathbb{R}^0$ .

Example. Assume that the source X is binary. The distortion function  $d_H(x, \hat{x}) = \mathbb{1}(x \neq \hat{x}), x, \hat{x} \in \{0, 1\}$  is called the binary Hamming distortion function.

Example. Assume that X and Y are binary random variables with joint distribution  $P_{X,Y}$ . Find  $\mathbb{E}(d_H(X,Y))$ .

$$\mathbb{E}(d_{H}(X,Y)) = \sum_{x,y} d(x,y) P_{X,Y}(x,y) = \sum_{x,y} \mathbb{1}(x \neq y) P_{X,Y}(x,y) = \sum_{x,y: x \neq y} P_{X,Y}(x,y) = P(X \neq Y).$$

Example. Assume that the source X is m-ary (i.e.  $\mathcal{X} = \{1, 2, \dots, m\}$ ). The distortion function  $d_H(x, \hat{x}) = \mathbb{1}(x \neq \hat{x}), x, \hat{x} \in \{1, 2, \dots, m\}$  is called the m-ary Hamming distortion function.

Example. Assume that the source X is a real-valued source (i.e.  $\mathcal{X} = \mathbb{R}$ ). Then, the squared error distortion is defined as  $d_2(x,\hat{x}) = (x-\hat{x})^2, x, \hat{x} \in \mathbb{R}$ .

So far, we have only defined distortion between single variables X and  $\hat{X}$ . Next, we define distortion between vectors of variables.

**Definition 2.** For a given distortion function  $d: \mathcal{X} \times \mathcal{X} \to \mathbb{R}^+$ , the distortion between the pair of vectors  $x^n$  and  $\hat{x}^n$  is defined as:

$$d(x^n, \hat{x}^n) = \frac{1}{n} \sum_{i=1}^n d(x_i, \hat{x}_i).$$

In the source coding problem, the input alphabet  $\mathcal{X}$ , the source distribution X and the distortion function d are given.

**Definition 3.** A source coding problem is characterized by the triple  $(\mathcal{X}, P_X, d)$ , where  $\mathcal{X}$  is called the source alphabet,  $P_X$  is called the source distribution and d is called the distortion crieteria.

Given the source coding problem, we are required to construct the coding strategy.

**Definition 4.** For the source coding problem characterized by  $(\mathcal{X}, P_X, d)$ , and natural numbers  $n, k \in \mathbb{N}$ , and (n,k)-coding strategy is characterized by a pair of functions  $e: \mathcal{X}^n \to \{0,1\}^k$  and  $f: \{0,1\}^k \to \mathcal{X}^n$  are called the encoding and decoding functions, respectively. The natural number n is called the blocklength of the coding strategy.

We often write  $\hat{X}^n$  instead of  $f(e(X^n))$  to denote the reconstruction of  $X^n$  at the decoder. The goal is to design coding strategies which have small average distortion.

**Definition 5.** For the source coding problem characterized by  $(\mathcal{X}, P_X, d)$ , and natural numbers  $n, k \in \mathbb{N}$ , and the coding strategy (e, f), the average distortion is defined as

$$\mathbb{E}(d(X^{n}, \hat{X}^{n})) = \sum_{x^{n}} P_{X^{n}}(x^{n}) d(x^{n}, f(e(x^{n}))).$$

The other objective is to use the minimum number of bits to store the source.

**Definition 6.** For the source coding problem characterized by  $(\mathcal{X}, P_X, d)$ , and natural numbers  $n, k \in \mathbb{N}$ , and (n, k)-coding strategy, the rate of the coding strategy is defined as  $r = \frac{k}{n}$ .

**Example 1.** Let n = 2, k = 1 and  $\mathcal{X} = \{0, 1\}$  and  $d_H$  is the Hamming directorion. Let e(00) = e(01) = 0 and e(11) = e(10) = 1. Also, assume that f(0) = 00 and f(1) = 11. Then,  $r = \frac{k}{n} = \frac{1}{2}$  and

$$D = \mathbb{E}(d(X^2, \hat{X}^2)) = \sum_{x^2 \in \{0,1\}^2} P_{X^2}(x^2) d(x^2, f(e(x^2))) =$$

$$P_{X^2}(0,0)d_H(00,00) + P_{X^2}(0,1)d_H(01,00) + P_{X^2}(1,0)d_H(10,11) + P_{X^2}(1,1)d_H(11,11) = \frac{1}{2}(P_{X^2}(01) + P_{X^2}(10))$$

**Definition 7.** For the source coding problem characterized by  $(\mathcal{X}, P_X, d)$ , the rate distortion pair (R, D) is said to be achievable if there exists a coding strategy (e, f) with rate at most R and average distortion at most D.

For a given distortion D, we are interested in finding the minimum rate R which can be used to store the source. This is formalized before:

**Definition 8.** For the source coding problem characterized by  $(\mathcal{X}, P_X, d)$ , the rate-distortion function is defined below:

$$R(D) = min_{(e,f): \frac{1}{n} \sum_{i=1}^{n} d(X_i, \hat{X}_i) \le D} r, \forall D \ge 0.$$

The following theorem states Shannon's source coding theorem:

**Theorem 1.** For the source coding problem characterized by  $(\mathcal{X}, P_X, d)$ , the rate-distortion function is characterized as follows:

$$R(D) = \min_{P_{\hat{X}|X}: \mathbb{E}(d(X,\hat{X})) \le D} I(X; \hat{X}).$$

## 2 Proof of Achievability

Let the distribution  $P_{\hat{X}|X}$  be the distribution that optimizes the rate-distortion formula. Also, fix  $n \in \mathbb{N}$  and  $\epsilon > 0$ . Let  $k = \lceil nR \rceil$ .

**Codebook Generation:** Construct a codebook containing  $2^k$  codewords each of length n where each codeword is chosen independently based on the distribution  $\prod_{i=1}^n P_{\hat{X}}(x_i)$ .

The matrix  $\mathcal{C}$  is generated by choosing each of its entries independently of other entries and based on the distribution  $P_{\hat{X}}$ .

**Encoding:** In order to store the sequence  $X^n$ , the encoder finds a codeword  $\underline{C}_i$  which is jointly typical with  $X^n$  with respect to the distribution  $P_{X,\hat{X}} = P_X P_{\hat{X}|X}$ . If such a codeword exists (not necessarily uniquely), then the index of the codeword is sent using k bits (note that there are a total of  $2^k$  codewords) otherwise an error is declared.

**Decoding:** The decoder receives the index i and declared the codeword  $\underline{C}_i$  as the reconstruction  $\hat{X}^n$ . The strategy fails if there is no codeword in the codeword which is jointly typical with  $X^n$ . We have:

$$\begin{split} &P(\mathcal{E}) = P(\nexists i: (X^n,\underline{C}_i) \in \mathcal{T}_{\epsilon}^n(X,\hat{X})) \\ &= \prod_{i=1}^{2^k} P((X^n,\underline{C}_i) \notin \mathcal{T}_{\epsilon}^n(X,\hat{X})) \\ &= (1 - P((X^n,\underline{C}_1) \in \mathcal{T}_{\epsilon}^n(X,\hat{X})))^{2^k}. \end{split}$$

Note that

$$\begin{split} &P((X^n,\underline{C}_1) \in \mathcal{T}^n_{\epsilon}(X,\hat{X})) = \sum_{(x^n,\hat{x}^n) \in \mathcal{T}^n_{\epsilon}(X,\hat{X})} P(X^n = x^n,\underline{C}_1 = \hat{x}^n) \\ &\stackrel{(a)}{=} \sum_{(x^n,\hat{x}^n) \in \mathcal{T}^n_{\epsilon}(X,\hat{X})} P(X^n = x^n) P(\underline{C}_1 = \hat{x}^n) \\ &\geq 2^{n(H(X,\hat{X}) - \epsilon)} 2^{-n(H(X) + \epsilon)} 2^{-n(H(\hat{X}) + \epsilon)} \\ &= 2^{-n(I(X;\hat{X}) + 3\epsilon)}. \end{split}$$

where (a) holds since  $X^n$  is produced by the source independently of  $\underline{C}_1$ . So,

$$P(\mathcal{E}) \le (1 - 2^{-n(I(X;\hat{X}) + 3\epsilon)})^{2^k}$$

Next, we use the following well-known inequality:  $(1-y)^n \leq 2^{-ny}, \forall y \in [0,1], n \in \mathbb{N}$ . We have:

$$P(\mathcal{E}) \le 2^{-2^{-n(I(X;\hat{X})+3\epsilon)} \cdot 2^k}.$$

The exponent can be further simplified as follows:

$$2^{-n(I(X;\hat{X})+3\epsilon)} \cdot 2^k = 2^{n(-I(X;\hat{X})-3\epsilon+\frac{k}{n})}$$

Note that  $\frac{k}{n} \geq R = I(X; \hat{X})$ . So,  $2^{-n(I(X; \hat{X}) + 3\epsilon + \frac{k}{n})}$  goes to infinity as  $n \to \infty$ . Define  $a_n = 2^{-n(I(X; \hat{X}) + 3\epsilon + \frac{k}{n})}$ . Then, we have shown that  $P(\mathcal{E}) \leq 2^{-a_n} \to 0$  as  $n \to \infty$ .

Next, we investigate the rate and distortion criteria. The rate of the coding strategy is  $\frac{k}{n} \approx R$  for large n. As for the distortion, the strategy guarantees that  $X^n$  and  $\hat{X}^n$  are jointly typical with respect to  $P_{X,\hat{X}}$ .

$$\frac{1}{n} \sum_{i=1}^{n} d(X_i, \hat{X}_i) \approx \mathbb{E}_{P_{X, \hat{X}}}(d(X, \hat{X})) \le D.$$

Where the last inequality follows from the assumption that  $P_{\hat{X}|X}$  is the optimizing distribution in the rate-distortion formula.

#### 3 Proof of converse

We first prove that the rate-distortion function is convex.

**Lemma 1.** For the source coding problem  $(X, P_X, d)$ , the rate-distortion function defined as  $R(D) = \min_{P_{\hat{X}|X}: \mathbb{E}(d(X,\hat{X}) \leq D} I(X;\hat{X})$  is convex. Alternatively, for any positive  $D_1$  and  $D_2$  and  $\lambda \in [0,1]$ , we have  $R(\bar{D}) \leq \lambda R(D_1) + (1-\lambda)R(D_2)$ , where  $\bar{D} = \lambda D_1 + (1-\lambda)D_2$ .

proof. It is enough to show that there exists a distribution  $P_{X,\hat{X}}$  for which  $\mathbb{E}_{P_{X,\hat{X}}}(d(X,\hat{X}) \leq \bar{D})$  and  $I_{P_{X,\hat{X}}}(X;\hat{X}) \leq \lambda R(D_1) + (1-\lambda)R(D_2)$ .

Assume that  $P_{\hat{X}|X}^i$ ,  $i \in \{1,2\}$  is the optimizing distribution for  $R(D_i)$ . Then, define  $P_{X,\hat{X}}^i = P_X P_{\hat{X}|X}^i$ . Also, define  $P_{X,\hat{X}} = \lambda P_{X,\hat{X}}^1 + (1-\lambda)P_{\hat{X}|X}^2$ . We have:

$$\begin{split} &\mathbb{E}_{P_{X,\hat{X}}}(d(X,\hat{X})) \\ &= \sum_{x,\hat{x}} P_{X,\hat{X}}(x,\hat{x}) d(x,\hat{x}) \\ &= \sum_{x,\hat{x}} (\lambda P^1_{X,\hat{X}}(x,\hat{x}) + (1-\lambda) P^2_{\hat{X}|X}(x,\hat{x})) d(x,\hat{x}) \\ &= \lambda \mathbb{E}_{P^1_{X,\hat{X}}}(d(X,\hat{X})) + (1-\lambda) \mathbb{E}_{P^2_{X,\hat{X}}}(d(X,\hat{X})) \\ &\leq \lambda D_1 + (1-\lambda) D_2 = D. \end{split}$$

Also, we know that mutual information  $I(X; \hat{X})$  is convex in  $P_{\hat{X}|X}$ . So,

$$I_{P_{X,\hat{X}}}(X;\hat{X}) \leq \lambda I_{P_{X,\hat{X}}^{1}}(X;\hat{X}) + (1-\lambda)I_{P_{X,\hat{X}}^{2}}(X;\hat{X}) = \lambda R(D_{1}) + (1-\lambda)R(D_{2}).$$

This completes the proof of the lemma.

We proceed to prove the converse. Let  $U^k = e(X^n)$  be the stored bits. Then, clearly  $H(U^k) \le \sum_{i=1}^k H(U_i) \le k$ . So,

$$k \geq H(e(X^{n})) \geq H(f(e(X^{n}))) = H(\hat{X}^{n})$$

$$= I(X^{n}; \hat{X}^{n}) + H(\hat{X}^{n}|X^{n})$$

$$\stackrel{(a)}{=} I(X^{n}; \hat{X}^{n})$$

$$= H(X^{n}) - H(X^{n}|\hat{X}^{n})$$

$$\stackrel{(b)}{=} \sum_{i=1}^{n} H(X_{i}) - H(X^{n}|\hat{X}^{n})$$

$$\stackrel{(c)}{=} \sum_{i=1}^{n} H(X_{i}) - H(X_{i}|\hat{X}_{i})$$

$$= \sum_{i=1}^{n} I(X_{i}, \hat{X}_{i}).$$

where (a) follows from the fact that  $\hat{X}^n = f(e(X^n))$  is a function of  $X^n$ , and (b) follows from the fact that the source is generated i.i.d, (c) follows from the fact that conditioning reduces entropy and the chain rule of entropy (exercise). Let  $d_i = \mathbb{E}(d(X_i, \hat{X}_i))$ , then, from the definition of the rate-distortion function, we have that  $I(X_i; \hat{X}_i) \geq R(d_i)$ . So,

$$k \ge \sum_{i=1}^{n} R(d_i) = n \sum_{i=1}^{n} \frac{1}{n} R(d_i) \ge n R(\frac{1}{n} \sum_{i=1}^{n} d_i),$$

where we have used the convexity of the rate-distortion function. On the other hand,

$$\frac{1}{n} \sum_{i=1}^{n} d_i = \frac{1}{n} \sum_{i=1}^{n} \mathbb{E}(d(X_i, \hat{X}_i))$$
$$= \mathbb{E}(\frac{1}{n} \sum_{i=1}^{n} d(X_i, \hat{X}_i)) \le D,$$

where in the last line we have used the fact that by assumption  $(\frac{1}{n}\sum_{i=1}^n d(X_i, \hat{X}_i) \leq D$ .

On the other hand it is straightforward to show that the rate-distortion function is decreasing in D (exercise). So,

$$k \ge nR(\frac{1}{n}\sum_{i=1}^{n}d_i) \ge nR(D) \Rightarrow \frac{k}{n} \ge R(D).$$

This completes the proof of the converse.

### 4 Calculating the rate-distortion function

**Example 2.** Let X be a Be(p) random variable where  $p \in [0, \frac{1}{2}]$ . Let  $d_H$  be the binary Hamming distortion. Then, for the source coding problem  $(\mathcal{X}, P_X, d_H)$ , the rate-distortion function is given by:

$$R(D) = \begin{cases} H(p, 1-p) - H(D, 1-D) & \text{if } D$$

proof. Assume that D < p. Note that  $\mathbb{E}(d_H(X, \hat{X})) = P(X \neq \hat{X}) = P(X \oplus \hat{X} = 1)$ . So,

$$I(X; \hat{X}) = H(X) - H(X|\hat{X}) = H(X) - H(X \oplus \hat{X}|\hat{X})$$
  
 
$$\geq H(X) - H(X \oplus \hat{X}) \geq H(p, 1 - p) - H(D, 1 - D).$$

So, we have shown that  $R(D) \geq H(p, 1-p) - H(D, 1-D)$ . On the other hand, let  $\hat{X} = X \oplus N_D$ , where  $N_D$  is a Be(D) random variable which is independent of  $\hat{X}$  (as an exercise, prove that such a variable always exists, in other words, prove that under these assumptions  $\hat{X}$  has a valid distribution and find the distribution  $P_{\hat{X}}$ ). Then,  $\mathbb{E}(d_H(X,\hat{X})) = P(X \oplus \hat{X} = 1) = P(N_D = 1) = D$ , and

$$I(X; \hat{X}) = H(X) - H(X|\hat{X}) = H(p, 1-p) - H(N_D|\hat{X}) = H(p, 1-p) - H(N_D) = H(p, 1-p) - H(D, 1-D).$$

This show that  $R(D) \le H(p, 1-p) - H(D, 1-D)$  which completes the proof for the case when D < p. The proof for the case when  $D \ge p$  is left as an exercise.

#### 5 Gaussian Sources

**Definition 9.** A Gaussian source coding problem is characterized by the triple  $(\mathcal{X}, f_X, d)$ , where  $\mathcal{X} = \mathbb{R}$ ,  $f_X$  is a Gaussian distribution and  $d_2$  is the squared distribution.

**Definition 10.** For the Gaussian source coding problem characterized by given the natural numbers  $n, k \in \mathbb{N}$ , and (n, k)-coding strategy is characterized by a pair of functions  $e : \mathbb{R}^n \to \{0, 1\}^k$  and  $f : \{0, 1\}^k \to \mathbb{R}^n$  are called the encoding and decoding functions, respectively.

**Theorem 2.** For the Gaussian source coding problem characterized by  $(\mathbb{R}, f_X, d_2)$ , the rate-distortion function is defined below:

$$R(D) = \min_{(e,f): \frac{1}{n} \sum_{i=1}^{n} d(X_i, \hat{X}_i) \leq D} r, \forall D \geq 0.$$

**Example 3.** Let X be a Gaussian random variable with variance  $\sigma^2$ . Then, for the Gaussian source coding problem  $(\mathbb{R}, f_X, d_2)$ , the rate-distortion function is given by:

$$R(D) = \begin{cases} \frac{1}{2} \log \frac{\sigma^2}{D} & \text{if } D < \sigma^2 \\ 0 & \text{Otherwise.} \end{cases}$$

proof. Assume that  $D < \sigma^2$ . First, we show that  $R(D) \ge \frac{1}{2} \log \frac{\sigma^2}{D}$ .

$$I(X; \hat{X}) = h_d(X) - h_d(X|\hat{X}) = h_d(X) - h_d(X - \hat{X}|\hat{X})$$
$$\geq h_d(X) - h_d(X - \hat{X}) = \frac{1}{2} \log 2\pi e \sigma^2 - h_d(X - \hat{X}).$$

On the other hand, we know that  $h_d(X - \hat{X}) \leq \frac{1}{2} \log 2\pi e \sigma_{X - \hat{X}}^2 = \frac{1}{2} \log 2\pi e D$ . So,  $I(X; \hat{X}) \geq \frac{1}{2} \log \frac{\sigma^2}{D}$ .

So, we have shown that  $R(D) \geq \frac{1}{2}\log \frac{\sigma^2}{D}$ . Next, let  $\hat{X}$  be a Gaussian random variable with variance  $\sigma^2 - D$  and  $N_D$  another Gaussian variable with variance D which is independent of  $\hat{X}$ . Let  $X = \hat{X} + N_D$  (check that this is valid). Then,  $\mathbb{E}(d_2(X,\hat{X})) = Var(X - \hat{X}) = Var(N_D) = D$ . Also,

$$I(X; \hat{X}) = h_d(X) - h_d(X|\hat{X}) = h_d(X) - h_d(N_D) = \frac{1}{2} \log \frac{\sigma^2}{D}.$$

 $R(D) \geq H(p,1-p) - H(D,1-D)$ . On the other hand, let  $\hat{X} = X \oplus N_D$ , where  $N_D$  is a Be(D) random variable which is independent of  $\hat{X}$  (as an exercise, prove that such a variable always exists, in other words, prove that under these assumptions  $\hat{X}$  has a valid distribution and find the distribution  $P_{\hat{X}}$ ). Then,  $\mathbb{E}(d_H(X,\hat{X})) = P(X \oplus \hat{X} = 1) = P(N_D = 1) = D$ , and

$$I(X; \hat{X}) = H(X) - H(X|\hat{X}) = H(p, 1-p) - H(N_D|\hat{X}) = H(p, 1-p) - H(N_D) = H(p, 1-p) - H(D, 1-D).$$