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## Assignment 2

Sep 27, 2023

3. Suppose we have a data set with five predictors,  $X_1 = \text{GPA}$ ,  $X_2 = \text{IQ}$ ,  $X_3 = \text{Level}$  (1 for College and 0 for High School),  $X_4 = \text{Interaction between GPA and IQ}$ , and  $X_5 = \text{Interaction between GPA and Level}$ . The response is starting salary after graduation (in thousands of dollars). Suppose we use least squares to fit the model, and get  $\hat{\beta}_0 = 50, \hat{\beta}_1 = 20, \hat{\beta}_2 = 0.07, \hat{\beta}_3 = 35, \hat{\beta}_4 = 0.01, \hat{\beta}_5 = -10$ .

**(a) Which answer is correct, and why?**

$$\text{Salary} = b_0 + b_1x_1 + b_2x_2 + b_3x_3 + b_4x_4 + b_5x_5 = 50 + 20x_1 + 0.07x_2 + 35x_3 + 0.01x_4 - 10x_5$$

for fixed IQ and GPA at  $x_1$  and  $x_2$ : - Salary (high school) =  $50 + 20x_1 + 0.07x_2 + 35*(0) + 0.01(x_1.x_2) - 10(x_1.0) = 50 + 20x_1 + 0.07x_2 + 0.01(x_1.x_2)$

$$\text{Salary (college)} = 50 + 20x_1 + 0.07x_2 + 35*(1) + 0.01(x_1.x_2) - 10(x_1.1) = 50 + 20x_1 + 0.07x_2 + 35 + 0.01(x_1.x_2) - 10(x_1) = \text{Salary (high school)} + 35 - 10(x_1)$$

From here:

$$\text{Salary (college)} - \text{Salary (high school)} = 35 - 10x_1$$

Assuming the salary difference to be more than equal to zero, we get:

$$35 - 10x_1 \geq 0 \rightarrow x_1 \leq 3.5$$

Assuming the salary difference to be less than equal to zero, we get:

$$35 - 10x_1 \leq 0 \rightarrow x_1 \geq 3.5$$

Hence, for a fixed value of IQ and GPA, high school graduates earn more, on average, than college graduates provided that the GPA is more than equal to 3.5.

**The correct answer is “iii”.**

**(b) Predict the salary of a college graduate with IQ of 110 and a GPA of 4.0.**

$$\text{Ans: Salary} = 50 + 20(4) + 0.07(110) + 35 + 0.01(110 \times 4) - 10(4) = 137.1$$

Hence, the predicted salary would be \$137,100.

**(c) True or false: Since the coefficient for the GPA/IQ interaction term is very small, there is very little evidence of an interaction effect. Justify your answer.**

Ans: This statement is false because the magnitude of coefficient is not an indicator of statistical significance.

10. This question should be answered using the **Carseats** data set.

→ Check python file as well for detailed answer to these questions.

(a) Fit a multiple regression model to predict **Sales** using **Price**, **Urban**, and **US**.

```
In [5]: Carseats = load_data("Carseats")
Carseats.columns

Out[5]: Index(['Sales', 'CompPrice', 'Income', 'Advertising', 'Population', 'Price',
              'ShelveLoc', 'Age', 'Education', 'Urban', 'US'],
              dtype='object')
```

```
In [15]: import patsy
f = 'Sales ~ Price + Urban + US'
y, X = patsy.dmatrices(f, Carseats, return_type='dataframe')

model = sm.OLS(y, X).fit()
print(model.summary())
```

OLS Regression Results

Dep. Variable:	Sales	R-squared:	0.239
Model:	OLS	Adj. R-squared:	0.234
Method:	Least Squares	F-statistic:	41.52
Date:	Wed, 27 Sep 2023	Prob (F-statistic):	2.39e-23
Time:	18:04:40	Log-Likelihood:	-927.66
No. Observations:	400	AIC:	1863.
Df Residuals:	396	BIC:	1879.
Df Model:	3		
Covariance Type:	nonrobust		

	coef	std err	t	P> t	[0.025	0.975]
Intercept	13.0435	0.651	20.036	0.000	11.764	14.323
Urban[T.Yes]	-0.0219	0.272	-0.081	0.936	-0.556	0.512
US[T.Yes]	1.2006	0.259	4.635	0.000	0.691	1.710
Price	-0.0545	0.005	-10.389	0.000	-0.065	-0.044

Omnibus: 0.676 Durbin-Watson: 1.912  
 Prob(Omnibus): 0.713 Jarque-Bera (JB): 0.758  
 Skew: 0.093 Prob(JB): 0.684  
 Kurtosis: 2.897 Cond. No. 628.

Notes:  
 [1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

(b) Provide an interpretation of each coefficient in the model. Be careful—some of the variables in the model are qualitative!

- On average the unit sales in urban location are 21.9 units less than in rural location all other predictors remaining fixed.
- A store located in the US sells, on average, 1200 more car seats than a store situated abroad.
- When the price rises by \$1000, and all other variables remain unchanged, the sales figures decrease by 54.5 units. In simpler terms, a \$1000 increase in price leads to a reduction in car seat sales by 54.5 units.

(c) Write out the model in equation form, being careful to handle the qualitative variables properly.

$$\text{Sales} = 13.0435 + (-0.0545) \times \text{Price} + (-0.0219) \times \text{Urban} + (1.2006) \times \text{US} + \varepsilon$$

with  $\text{Urban}=1$  if the store is in an urban location and 0 if not, and  $\text{US}=1$  if the store is in the US and 0 if not.

(d) For which of the predictors can you reject the null hypothesis  $H_0 : \beta_j = 0$ ?

We can reject the null hypothesis for “Price” and “US” variables.

(e) On the basis of your response to the previous question, fit a smaller model that only uses the predictors for which there is evidence of association with the outcome.

```
In [16]: f = 'Sales ~ Price + US'
y, X = patsy.dmatrices(f, Carseats, return_type='dataframe')

model2 = sm.OLS(y, X).fit()
print(model2.summary())
```

OLS Regression Results						
Dep. Variable:	Sales	R-squared:	0.239			
Model:	OLS	Adj. R-squared:	0.235			
Method:	Least Squares	F-statistic:	62.43			
Date:	Wed, 27 Sep 2023	Prob (F-statistic):	2.66e-24			
Time:	18:05:15	Log-Likelihood:	-927.66			
No. Observations:	400	AIC:	1861.			
Df Residuals:	397	BIC:	1873.			
Df Model:	2					
Covariance Type:	nonrobust					
	coef	std err	t	P> t	[0.025	0.975]
Intercept	13.0308	0.631	20.652	0.000	11.790	14.271
US[T.Yes]	1.1996	0.258	4.641	0.000	0.692	1.708
Price	-0.0545	0.005	-10.416	0.000	-0.065	-0.044
Omnibus:	0.666	Durbin-Watson:	1.912			
Prob(Omnibus):	0.717	Jarque-Bera (JB):	0.749			
Skew:	0.092	Prob(JB):	0.688			
Kurtosis:	2.895	Cond. No.	607.			

Notes:  
[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

(f) How well do the models in (a) and (e) fit the data?

R - squared and Adjusted R squared values are same for both the models. But the smaller model has a higher f -squared value that means smaller model (e) is a better fit compared to (a).

(g) Using the model from (e), obtain 95% confidence intervals for the coefficient(s).

```
In [11]: confidence_intervals = results2.conf_int(alpha=0.05)
```

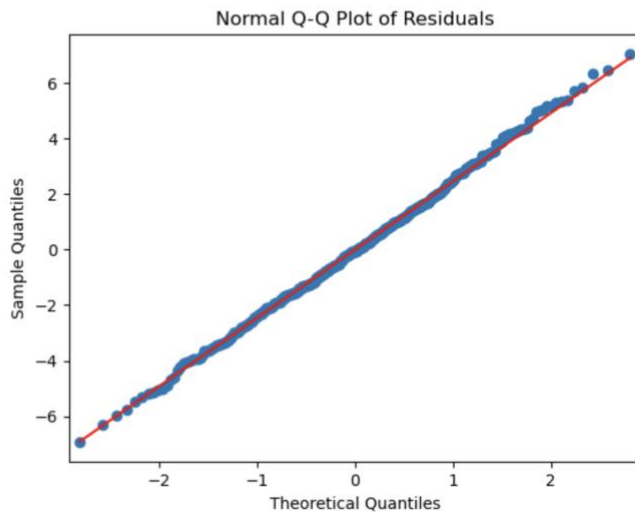
```
In [12]: print(confidence_intervals)
```

	0	1
intercept	11.79032	14.271265
Price	-0.06476	-0.044195
US[Yes]	0.69152	1.707766

(h) Is there evidence of outliers or high leverage observations in the model from (e)?

```
In [67]: # 10 (h)
residuals = results2.resid

sm.qqplot(residuals, line='s')
plt.title('Normal Q-Q Plot of Residuals')
plt.show()
```



Answer: Considering that most of the residuals align closely with the diagonal line, it indicates that they exhibit an approximate normal distribution. While there are a few outliers represented by minor deviations from the diagonal line, these outliers do not appear to pose significant issues.

#### 14. This problem focuses on the *collinearity* problem.

(a) Perform the following commands in **Python**:

```
rng = np.random.default_rng(10)
x1 = rng.uniform(0, 1, size=100)
x2 = 0.5 * x1 + rng.normal(size=100) / 10
y = 2 + 2 * x1 + 0.3 * x2 + rng.normal(size=100)
```

The last line corresponds to creating a linear model in which **y** is a function of **x1** and **x2**. Write out the form of the linear model. What are the regression coefficients?

*The form of the linear model is:*

$$y = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \epsilon$$

$\epsilon \sim N(0,1)$  random variable. The regression coefficients are respectively 2, 2 and 0.3.

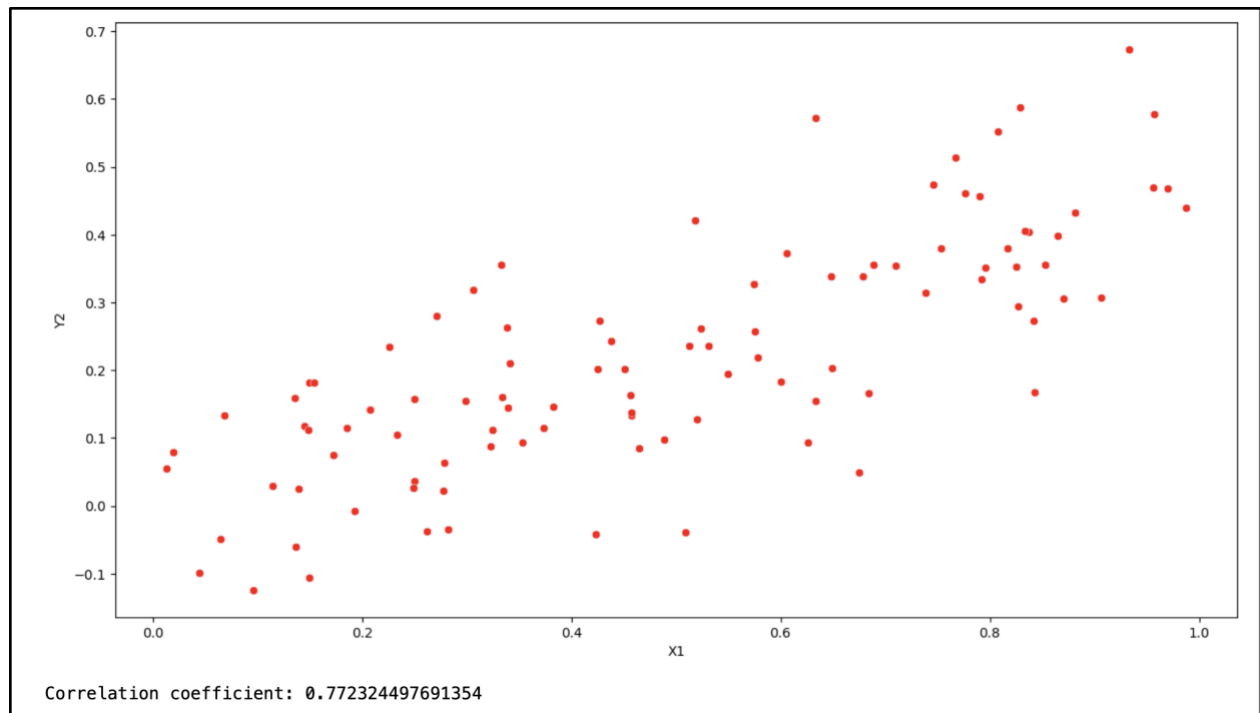
(b) What is the correlation between **x1** and **x2**? Create a scatterplot displaying the relationship between the variables.

```
In [28]: fig = plt.figure(figsize=(15,8))
ax = fig.add_subplot(111)
ax = sns.scatterplot(x= x1, y= x2, color='r')

ax.set_xlabel("x1")
ax.set_ylabel("y2")

plt.show()

print("Correlation coefficient: " + str(np.corrcoef(x1, x2)[0][1]))
```



The correlation between  $x_1$  and  $x_2$  is 0.77.

(c) Using this data, fit a least squares regression to predict  $y$  using  $x_1$  and  $x_2$ . Describe the results obtained. What are  $\hat{\beta}_0$ ,  $\hat{\beta}_1$ , and  $\hat{\beta}_2$ ? How do these relate to the true  $\beta_0$ ,  $\beta_1$ , and  $\beta_2$ ? Can you reject the null hypothesis  $H_0 : \beta_1 = 0$ ? How about the null hypothesis  $H_0 : \beta_2 = 0$ ?

```
In [31]: X = np.stack((x1, x2), axis=-1)
X = sm.add_constant(X, prepend=True)
```

```
model = sm.OLS(y, X)
result = model.fit()
print(result.summary())
```

```

=====
                        OLS Regression Results
=====
Dep. Variable:          y      R-squared:                0.291
Model:                  OLS    Adj. R-squared:           0.276
Method:                 Least Squares    F-statistic:       19.89
Date:                   Wed, 27 Sep 2023    Prob (F-statistic):  5.76e-08
Time:                   18:41:26    Log-Likelihood:     -130.62
No. Observations:      100    AIC:                267.2
Df Residuals:          97    BIC:                275.1
Df Model:               2
Covariance Type:       nonrobust
=====

```

	coef	std err	t	P> t	[0.025	0.975]
const	1.9579	0.190	10.319	0.000	1.581	2.334
x1	1.6154	0.527	3.065	0.003	0.569	2.661
x2	0.9428	0.831	1.134	0.259	-0.707	2.592

```

=====
Omnibus:                 0.051    Durbin-Watson:           1.964
Prob(Omnibus):           0.975    Jarque-Bera (JB):         0.041
Skew:                   -0.036    Prob(JB):                 0.979
Kurtosis:                2.931    Cond. No.:                11.9
=====

```

Notes:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

The coefficients  $\hat{\beta}_0$ ,  $\hat{\beta}_1$ , and  $\hat{\beta}_2$  are respectively 1.9579, 1.6154 and 0.9428. As the p-value is less than 0.05, we may reject null hypothesis for  $\beta_0$  and  $\beta_1$ , however we may not reject  $H_0$  for  $\beta_2$  as the p-value is higher than 0.05.

(d) Now fit a least squares regression to predict **y** using only **x1**. Comment on your results. Can you reject the null hypothesis  $H_0 : \beta_1 = 0$ ?

```
In [38]: X = sm.add_constant(x1, prepend=True)
```

```
model2 = sm.OLS(y, X)
result2 = model2.fit()
print(result2.summary())
```

```

=====
                        OLS Regression Results
=====
Dep. Variable:          y      R-squared:                0.281
Model:                  OLS    Adj. R-squared:           0.274
Method:                 Least Squares    F-statistic:       38.39
Date:                   Wed, 27 Sep 2023    Prob (F-statistic):  1.37e-08
Time:                   19:04:23    Log-Likelihood:     -131.28
No. Observations:      100    AIC:                266.6
Df Residuals:          98    BIC:                271.8
Df Model:               1
Covariance Type:       nonrobust
=====

```

	coef	std err	t	P> t	[0.025	0.975]
const	1.9371	0.189	10.242	0.000	1.562	2.312
x1	2.0771	0.335	6.196	0.000	1.412	2.742

```

=====
Omnibus:                 0.204    Durbin-Watson:           1.931
Prob(Omnibus):           0.903    Jarque-Bera (JB):         0.042
Skew:                   -0.046    Prob(JB):                 0.979
Kurtosis:                3.038    Cond. No.:                4.65
=====

```

Notes:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

The coefficient for “x1” in this model is different from the one with “x1” and “x2” as predictors. In this case “x1” is highly significant as its p-value is very low, so we may reject  $H_0 : \beta_1 = 0$ .

**(e) Now fit a least squares regression to predict  $y$  using only  $x_2$ . Comment on your results. Can you reject the null hypothesis  $H_0 : \beta_1 = 0$ ?**

```
In [43]: X = sm.add_constant(x2, prepend=True)
```

```
model3 = sm.OLS(y, X)
result3 = model3.fit()
print(result3.summary())
```

```

=====
                        OLS Regression Results
=====
Dep. Variable:          y      R-squared:          0.222
Model:                  OLS    Adj. R-squared:       0.214
Method:                 Least Squares    F-statistic:      27.99
Date:                   Wed, 27 Sep 2023    Prob (F-statistic): 7.43e-07
Time:                   19:08:43    Log-Likelihood:    -135.24
No. Observations:       100    AIC:                274.5
Df Residuals:           98    BIC:                279.7
Df Model:                1
Covariance Type:        nonrobust
=====

```

	coef	std err	t	P> t	[0.025	0.975]
const	2.3239	0.154	15.124	0.000	2.019	2.629
x1	2.9103	0.550	5.291	0.000	1.819	4.002

```

=====
Omnibus:                 0.191    Durbin-Watson:       1.943
Prob(Omnibus):           0.909    Jarque-Bera (JB):     0.373
Skew:                    -0.034    Prob(JB):             0.830
Kurtosis:                2.709    Cond. No.             6.11
=====

```

Notes:  
[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

The coefficient for “x2” in this model is different from the one with “x1” and “x2” as predictors. The value for  $\hat{\beta}_1$  is 2.9103. In this case “x2” is highly significant as its p-value is very low, so we may again reject  $H_0 : \beta_1 = 0$ .

**(f) Do the results obtained in (c)–(e) contradict each other? Explain your answer.**

The variations in significance levels do not inherently conflict with one another; instead, they emphasize the significance of factoring in the broader context and additional predictors when interpreting the significance of individual predictors within a multiple regression framework.

**(g) Suppose we obtain one additional observation, which was unfortunately mismeasured. We use the function `np.concatenate()` to add this additional observation to each of  $x_1$ ,  $x_2$  and  $y$ .**

```
x1 = np.concatenate([x1, [0.1]])
x2 = np.concatenate([x2, [0.8]])
y = np.concatenate([y, [6]])
```

Re-fit the linear models from (c) to (e) using this new data. What effect does this new observation have on the each of the models? In each model, is this observation an outlier? A high-leverage point? Both? Explain your answers.

→ Check Python file for this answer.

**5.8 Twitter users and news, Part I.** A poll conducted in 2013 found that 52% of U.S. adult Twitter users get at least some news on Twitter<sup>12</sup>. The standard error for this estimate was 2.4%, and a normal distribution may be used to model the sample proportion. Construct a 99% confidence interval for the fraction of U.S. adult Twitter users who get some news on Twitter, and interpret the confidence interval in context.

52% of adult Twitter users obtain some of their news from the platform, with a standard error estimate of 2.4%. Given this data, a normal distribution can be applied. The critical value for a 99% confidence interval is approximately 2.575829.

Calculating the confidence interval:

Lower limit:  $0.52 - (2.575829)(0.024) = 0.4582$

Upper limit:  $0.52 + (2.575829)(0.024) = 0.5818$

This results in a confidence interval of 0.4582 to 0.5818. The interpretation of the confidence interval is as follows: "We are 99% confident that the fraction of U.S. adult twitter users who get some news on twitter is between 45.82% and 58.18%. The margin of error in this context is 6.18%. This margin of error pertains to the percentage of users who receive news on Twitter and not the margin of error for the difference between the percentage obtaining news on Twitter and those who do not.

**5.16 Identify hypotheses, Part II.** Write the null and alternative hypotheses in words and using symbols for each of the following situations.

- (a) Since 2008, chain restaurants in California have been required to display calorie counts of each menu item. Prior to menus displaying calorie counts, the average calorie intake of diners at a restaurant was 1100 calories. After calorie counts started to be displayed on menus, a nutritionist collected data on the number of calories consumed at this restaurant from a random sample of diners. Do these data provide convincing evidence of a difference in the average calorie intake of a diners at this restaurant?
- (b) The state of Wisconsin would like to understand the fraction of its adult residents that consumed alcohol in the last year, specifically if the rate is different from the national rate of 70%. To help them answer this question, they conduct a random sample of 852 residents and ask them about their alcohol consumption.

a)

Null Hypothesis ( $H_0$ ): There is no change in the average calorie intake for diners.

Alternative Hypothesis ( $H_A$ ): There is a difference in calorie intake for diners.

$H_0: \mu = 1100$

$H_A: \mu \neq 1100$



b)

Null Hypothesis ( $H_0$ ): The fraction of Wisconsin adults who consume alcohol is equal to national average of 0.7.

Alternative Hypothesis ( $H_A$ ): The fraction of Wisconsin adults who consume alcohol is different from national average of 0.7.

Let, 'p' be the population of Wisconsin's adult who consumed alcohol past year.

$H_0: p = 0.70$

$H_A: p \neq 0.70$