APPLICATIONS

· Probability estimation

Let A be an event (set of outcomes of a random phenomenon) of which we want to estimate the probability P(A).

Examples

- · A = { the outcome of flipping a coin is head}
- A = { a person who tested positive to COVID-19 is asymptomatic}
- $A = \{ X(t) = x \}$, where X(t) is the state of a stochastic timed automaton at time t.

Consider N independent observations of the random phenomenon, and let NA be the number of times that A was observed. Then, an estimate of P(A) can be computed as:

$$P(A) = \frac{N_A}{N}$$

The properties of this estimator can be studied with the Law of Large Numbers. Let we be the i-th outcome of the random phenomenon, and define the random variable:

$$1_{A}(\omega_{i}) \stackrel{\triangle}{=} \begin{cases} 1 & \text{if event A is observed in } \omega_{i} \\ 0 & \text{otherwise} \end{cases}$$

Indicator function of event A

1 is a random variable such that:

•
$$E[I_A] = 0 \cdot [1 - P(A)] + 1 \cdot P(A) = P(A)$$

•
$$Var(1_A) = [0 - P(A)]^2 [1 - P(A)] + [1 - P(A)]^2 P(A)$$

$$= P(A)[1 - P(A)][P(A) + 1 - P(A)]$$

$$= P(A)[1 - P(A)]$$

- => Since $E[1_A] = P(A)$, an estimate of $E[1_A]$ is an estimate of P(A).
- => E[1A] can be estimated using the Law of Large Numbers, and with the theoretical guarantees thereof.

In this respect, notice that:

$$\frac{\sum_{i=1}^{N} \mathcal{1}_{A}(\omega_{i})}{N} = \frac{N_{A}}{N}$$

Arithmetic mean of the observations $1_A(\omega_i)$

Hence, according to the Law of Large Numbers, $\frac{N_A}{N}$ is an unbiased, consistent estimate of P(A).

For the choice of N (whenever this is possible), consider the Following.

Assume that a desired accuracy $\Delta > 0$ of the estimate is given. The problem is to choose N such that

$$|P(A) - P(A)| \leq \Delta$$

For the Central Limit Theorem, we know that

$$\hat{P(A)} \sim \mathcal{N}(M, \frac{\sigma^2}{N})$$
,

where

•
$$\sigma^2 = Var(1_A) = P(A)[1-P(A)]$$

for N sufficiently large.

Hence,

$$P\left(|\hat{P(A)} - P(A)| \le \frac{3\sigma}{\sqrt{N}}\right) \simeq 0.9973$$

If we accept that 3 times out of 1000 (on average)
the estimate differs from the true value more than

 $\frac{3\sigma}{\sqrt{N}}$, we can set

From the tables of the Normal distribution: if $X \sim N(\mu, \sigma^2)$, $P(|X-\mu| \leq \sigma) \simeq 0.6827$ $P(|X-\mu| \leq 2\sigma) \simeq 0.3545$ $P(|X-\mu| \leq 3\sigma) \simeq 0.9973$

$$\Delta = \frac{3\sigma}{\sqrt{N}} \implies N = \frac{9\sigma^2}{\Delta^2} \quad (*)$$

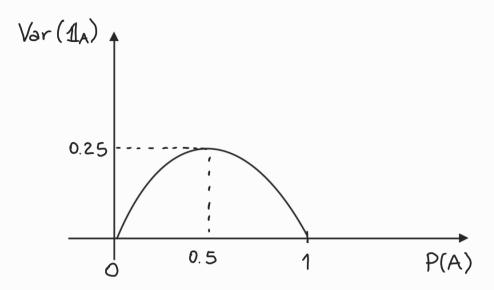
The problem with this formula is that

$$\sigma^2 = P(A)[1-P(A)] \Rightarrow It$$
 depends on the quantity we want to estimate

However, if we plot this function versus P(A), we observe that

$$P(A)[1-P(A)] \le 0.25 \quad \forall 0 \le P(A) \le 1$$

The maximum is achieved for P(A)=0.5.



This makes it possible to replace of in (*) with its upper bound 0.25, thus obtaining the Following Formula For the choice of N:

$$N = \frac{2.25}{\Delta^2}$$

We used that 952 < 9.0.25 = 2.25.

EXAMPLE

We have an unfair coin.

The probability to get head is P=0.7.

Assume that we do not know p, and we want to estimate it from the results of flipping the coin repeatedly.

Let $\Delta = 0.01$ be the required accuracy for estimating p.

Using the approximated formula for N, we have

$$N = \frac{2.25}{(0.04)^2} = 22500$$

With this choice of N, we expect that, on average, no more than 3 times out of 1000 the estimate \hat{p} should differ from the true value p more than Δ .

We compute M=1000 estimates $\hat{\rho}$ of p. Each estimate is computed using N=22500 observations of the random experiment. One observation consists of flipping the coin, and recording the result (head or tail).

The figure shows the histogram of the Mestimates. All the estimates except one are within the interval $[p-\Delta, p+\Delta] = [0.69, 0.71]$.

