# 2 Transformation of pairs of r.v.s

Given 2 random variables and their joint function, and given a function of them, find its distribution.  $f_{Y_1,Y_2}(y_1,y_2) = \sum_{i=1}^k f_{X_1,X_2}(h_1(y_1,y_2),h_2(y_1,y_2))|J_i|$ 

$$J_{1,2} = \begin{vmatrix} \frac{\partial h_1}{\partial y_1} & \frac{\partial h_2}{\partial y_2} \\ \frac{\partial h_1}{\partial y_1} & \frac{\partial h_2}{\partial y_2} \\ \frac{\partial h_2}{\partial y_1} & \frac{\partial h_2}{\partial y_2} \end{vmatrix} = \frac{\partial h_1}{\partial y_1} \frac{\partial h_2}{\partial y_2} - \frac{\partial h_2}{\partial y_1} \frac{\partial h_1}{\partial y_2}$$

 $U \sim Geom(p=\frac{1}{2}), u=1,2..$  the number of trials needed to get the first head.

 $V \sim NBin(p=\frac{1}{2},r=2), v=2,3..$  the number of trials needed to get two heads in repeated tosses of a fair coin.

the distribution of (U,V) is  $\{(u,v): u=1,2,...; v=u+1,u+2,...\}$  is not a cross-product set. U and V are not independent.

### 3 Covariance and Correlation

Given a joint pdf, find the covariance, correlation, conditional expectation, conditional variance.

## Conditional expectations and variances

$$f(x_{2}|x_{1}) = \frac{f(x_{1},x_{2})}{f_{1}(x_{1})}; p(x_{2}|x_{1}) = \frac{p(x_{1},x_{2})}{p_{1}(x_{1})}$$

$$4.2.1 \ f(y|x) = P(Y = y|X = x) = \frac{f(x,y)}{f_{X}(x)}$$

$$4.2.3 \ E(g(x_{2})|x_{1}) = \sum_{all \ x_{2}} g(x_{2})p(x_{2}|x_{1}), E(g(x_{2})|x_{1}) = \int_{all \ x_{2}} g(x_{2})f(x_{2}|x_{1})dx_{2}$$

$$4.2.4 \ Var(y|x) = E(Y^{2}|x) - (E(Y|x))^{2}, \ V[X|Y] = E[(X - E[X|Y])^{2}|Y]$$

$$4.2.7 \ f(x,y) = g(x)h(y) \iff \text{indep}$$

$$4.2.10 \ E(g(X)h(Y)) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x)h(y)f(x,y)dxdy = \int_{-\infty}^{\infty} h(y)f_{Y}(y) \left[\int_{-\infty}^{\infty} g(x)f_{X}(x)dx\right]dy = E[g(x)]E[h(y)] = (Eg(X))(Eh(Y))$$

$$F_{XY}(x,y) = P(X \le x, Y \le y)$$

# Properties of $\bar{X}$ and $S^2$

$$\begin{split} \bar{X}_{n+1} &= \frac{\sum_{i=1}^{n+1} X_i}{n+1} = \frac{\sum_{i=1}^{n} X_i + X_{n+1}}{n+1} = \frac{n\bar{X}_n + X_{n+1}}{n+1} \\ nS_{n+1}^2 &= (n-1)S_n^2 + \left(\frac{n}{n+1}\right)(X_{n+1} - \bar{X}_n)^2 \\ \text{when } X_1, .., X_n \text{ be iid } n(\mu, \sigma^2), \ \frac{(n-1)S^2}{\sigma^2} \sim \chi_{n-1}^2, Var\chi_{n-1}^2 = 2(n-1) \\ Var[\frac{(n-1)S^2}{\sigma^2}] &= \frac{(n-1)^2}{\sigma^4} Var[S^2] = 2(n-1) \implies Var[S^2] = \frac{2\sigma^4}{n-1} \end{split}$$

# 1 Connection between $N, \chi^2, t, F$

5.3

 $X_1...X_n \sim iidN(\mu,\sigma^2)$ . Given some function of these, find the distribution.

W Normal 
$$|\bar{X} \sim N(\mu, \frac{\sigma^2}{n}) - \bar{X}, S^2 \text{ indep} - \frac{(n-1)S^2}{\sigma^2} \sim \chi_{n-1}^2$$
  
W/o Normal  $|E[\bar{X}] = \mu$   $|V[\bar{X}] = \frac{\sigma^2}{n}$   $|E[S^2] = \sigma^2$ 

 $\frac{\text{W/o Normal}}{\text{S.2.8 Distribution of the mean: }} \frac{E[\bar{X}] = \mu}{V[\bar{X}] = \frac{\sigma^2}{n})} \frac{\sigma^2}{E[S^2] = \sigma^2}$ 5.2.8 Distribution of the mean:  $\bar{X} \sim gamma(n\alpha, \beta/n)$  5.2.9 Convolution formula Z = X + Y $f_Z(z) = \int_{-\infty}^{\infty} f_X(w) f_Y(z-w) dw$ 

5.3.2 
$$Z \sim n(0,1), Z^2 \sim \chi_1^2, X_1...X_n$$
 indep  $X_i \sim \chi_{p_i}^2, X_1 = ... = X_n \sim \chi_{p_1+...p_n}^2$ 

 $\chi_p^2 \sim gamma(p/2,2)$ 

5.3.5 
$$X_1..X_n, Y_1..Y_m$$
 indep.  $X_i \sim n(\mu_X, \sigma_X^2)Y_j \sim n(\mu_Y, \sigma_Y^2), \frac{S_X^2}{\sigma_X^2} \sim \chi^2 F: \frac{S_X^2/\sigma_X^2}{S_Y^2/\sigma_Y^2}$ 

#### 4 Order statistics

Find the distribution of X(k) or find the joint distribution of  $X_{(i)}, X_{(k)}$ 

$$5.4.4 f_K(x) = \frac{n!}{(j-1)!(n-j)!} K[F_X(x)]^{k-1} [1 - F_X(x)]^{n-k} f(x)$$

1-31p4-8

5.4.5 uniform order pdf

1-31p2-3

5.4.6

$$f_{X_{(i)},X_{(j)}}(u,v) = \frac{n!}{(i-1)!(j-1-i)!(n-j)!} f_X(u) f_X(v) [F_X(u)]^{i-1} [F_X(v) - F_X(u)]^{j-1-i} [1 - F_X(v)]^{n-j}$$

$$f_{X_{(1)},...,X_{(n)}}(x_1,...,x_n) = \begin{cases} n! f_X(x_1) \cdot ... \cdot f_X(x_n) & -\infty < x_1 < .. < x_n < \infty \\ 0 & \text{otherwise} \end{cases}$$

joint pmf

$$\begin{array}{l} \text{Conditional/Baye's:} \ P(A|B) = \frac{P(A \cap B)}{P(B)} = \frac{P(B|A)P(A)}{P(B)}; \\ \text{disjoint} \\ P(B) = P(B|A_1)P(A_1) + ..P(B|A_k)P(A_k) = \sum_{i=1}^k P(B|A_i)P(A_i) \\ P(A \cap B) = P(A)P(B) & P(X \leq x, Y \leq y) = P(X \leq x)P(Y \leq y) \\ \textbf{Independence:} \ E(XY) = E(X)E(Y) & F_{X,Y}(x,y) = F_X(x)F_Y(y) \\ V(X \pm Y) = VX + VY & f_{X,Y}(s,t) = f_X(s)f_Y(t) \end{array}$$

# pdf/pmf

CDF: 
$$F(x) = P(X \le x) = \sum_{y:y \le x} p(y) = \int_{-\infty}^{x} f(y) dy$$

probabilities: a < b:

PMF: 
$$p(x) = P(X = x) = P(\forall w \in \mathcal{W} : X(w) = x)$$

$$P(a \le X \le b) = F(b) - F(a^{-}); P(a < X \le b) = F(b) - F(a); P(a \le X \le a) = p(a);$$

 $P(a < X < b) = F(b^{-}) - F(a)$ ; (where  $a^{-}$  is the largest possible X value strictly less than a); Taking a = b yields P(X = a) = F(a) - F(a - 1) as desired.

PDF: 
$$P(\forall w \in \mathcal{W} : a \leq X(w) \leq b) = \int_a^b f(x) dx$$

$$P(a \le X \le b) = P(a < X \le b) = P(a < X \le b) = P(a < X < b); P(X > a) = 1 - F(a); P(a \le X \le b) = F(b) - F(a)$$

#### CDF

Condition: 
$$f(x) \ge 0 \forall x \text{ pmf } \sum_{x} f_X(x) = 1, pdf \int_{-\infty}^{\infty} f(x)dx = 1$$

$$c \sin x \ge 0, 0 < x < \pi/2 \qquad \left| \begin{array}{c} \int_0^{\frac{\pi}{2}} c \sin x dx = 1 \\ c e^{-|x|} \ge 0, -\infty < x < \infty \end{array} \right| \begin{array}{c} c = 1 \\ c \int_{-\infty}^{\infty} e^x dx + c \int_0^{\infty} e^{-x} dx = 1 \end{array} \right| \begin{array}{c} c = 1 \\ c = \frac{1}{2} \end{array}$$

## mean/variance

$$\begin{split} E[g(x)] &= \mu = \sum_{x \in D} h(x) p(x) = \int_{-\infty}^{\infty} g(x) f(x) dx \\ E[(X - \mu)^n] &= \mu_n = \sum (x - \mu)^n p(x) = \int (x - \mu)^n f(x) dx \\ E(aX + b) &= aE(X) + b \\ V(aX + b) &= a^2 \sigma^2 \\ \sigma_{ax + b} &= |a| \cdot \sigma_x \end{split} \qquad \begin{array}{c} E(X) = \int_{-\infty}^{\infty} F_X(t) dt + \int_0^{\infty} F_X(t) dt \\ V(X \pm Y) &= VX + VY \pm 2Cov(X, Y) \\ Cov(x, y) &= E(XY) - E(X)E(Y) \end{split}$$

$f_X(x)$	EX	$EX^2$	V(X)
$\frac{1}{n}$	$\sum_{1}^{n} x \frac{1}{n}$	$\sum_{1}^{n} x^{2} \frac{1}{n}$	
	$\frac{1}{n}\frac{n(n+1)}{2}$	$\frac{1}{n} \frac{n(n+1)(2n+1)}{6}$	$\frac{n^2-1}{12}$
$ax^{a-1}$	$\int_0^1 x a x^{a-1} dx$	$\int_{0}^{1} x^{2} a x^{a-1} dx$	$\frac{\frac{1}{a}}{a+2} - \left(\frac{a}{a+1}\right)^2$
	$\int_{0}^{1} x a x^{a-1} dx$ $\frac{a}{a+1} x^{a+1} \Big _{0}^{1}$	$\int_{0}^{1} x^{2} a x^{a-1} dx$ $\frac{a}{a+2} x^{a+2} \Big _{0}^{1}$	$\frac{a+2}{(a+2)(a+1)^2}$
$\frac{3}{2}(x-1)^2$	$\int_0^2 x \frac{3}{2} (x-1)^2 dx$	$\int_0^2 x^2 \frac{3}{2} (x-1)^2 dx$	$\frac{8}{5} - 1^2$
	$\int_0^2 x \frac{3}{2} (x-1)^2 dx$ $\frac{3}{2} \left(\frac{x^4}{4} \Big _0^2\right)$	$\frac{3}{2}(\frac{x^5}{5} _0^2)$	$\frac{3}{5}$

### transform

$$\begin{array}{c|c} g(x) \uparrow & F_Y(y) = F_X(g^{-1}(y)) \\ g(x) \downarrow & F_Y(y) = 1 - F_X(g^{-1}(y)) \end{array}$$

not monotone 
$$: X \leq 0$$
 is  $\emptyset : P(X \leq -\sqrt{y}) = 0$ 

not monotone : 
$$X \le 0$$
 is  $\emptyset$  :  $P(X \le -\sqrt{y}) = 0$   
 $F_Y(y) = P(Y \le y) = P(X^2 \le y) = P(-\sqrt{y} \le X \le \sqrt{y}) = P(X \le \sqrt{y}) - P(X \le -\sqrt{y}) = F_X(\sqrt{y}) = Y(y) =$ 

$$\begin{aligned} \mathbf{monotone} : & f_Y(y) = f_X(g^{-1}(y)) |\frac{d(g^{-1}(y))}{dy}| & 2 & e^X & |f_X(x) \\ & 2 & e^X & \ln Y & x^n(1-x)^m, 0 < x < 1 \\ & 3 & -2lnX & e^{-Y/2} & 1, 0 < x < 1 \\ & 4 & X^2 & \sqrt{Y} & 1, 0 < x < 1 \\ & 5 & d \tan X & \arctan \frac{Y}{d} & \frac{1}{b-a} = \frac{2}{\pi} \end{aligned}$$

$$\begin{array}{l} 1 \ (e^{-y})^n (1-e^{-y})^m |\frac{d(e^{-y})}{dy}| = e^{-(n+1)y} (1-e^{-y})^m, \ 0 < y < \infty \\ 2 \ \frac{1}{\sigma^2} \ln y e^{-(\ln y/\sigma)^2/2} |\frac{d(\ln y)}{dy}| = \frac{\ln(y)}{\sigma^2 y} e^{-(\ln(y)/\sigma)^2/2}, 1 < y < \infty \end{array}$$

$$2 \frac{1}{\sigma^2} \ln y e^{-(\ln y/\sigma)^2/2} \left| \frac{d(\ln y)}{dy} \right| = \frac{\ln(y)}{\sigma^2 y} e^{-(\ln(y)/\sigma)^2/2}, 1 < y < \infty$$

$$31\left|\frac{de^{-Y/2}}{du}\right| = \frac{1}{2}e^{-Y/2}, 0 < y < \infty$$

$$4 \ 1 \left| \frac{d\sqrt{y}}{dy} \right| = \frac{1}{2\sqrt{y}}, 0 < y < 1$$

$$5 \frac{2}{\pi} \left| \frac{d(\arctan \frac{y}{d})}{dy} \right| = \frac{2}{\pi} \frac{d(\frac{y}{d})}{\left[ (\frac{y}{d})^2 + 1 \right]} = \frac{2}{\pi d\left[ (\frac{y}{d})^2 + 1 \right]} 0 < y < \infty$$

## find mean/variance by mgf

$$\begin{split} E(e^{tx}) &= \int e^{tx} f(x) dx = \sum e^{tx} f(x); \ M_{aX+b}(t) = e^{tb} M_{aX}(t) \\ M_X(t) &= E(e^{tx}) & M_X(0) = 1 \\ M_X'(t) &= E(xe^{tx}) & M_X'(0) = E(X) \\ M_X''(t) &= E(x^2 e^{tx}) & M_X''(0) = E(X^2) \\ M_X^n(t) &= E(x^n e^{tx}) & M_X^n(0) = E(X^n) \end{split}$$

### Series

$$\int k dx = kx$$

$$\int e^{u} du = e^{u}$$

$$\int \ln aa^{x} = a^{x}$$

$$\int \int \ln a dx = c(b-a)$$

$$\int \ln a dx = c(b-a)$$

$$\begin{split} &=uv|_a^b - \int_a^b v du \\ &= -xe^{-x} + \int xe^{-x} dx = -xe^{-x} - e^{-x} + c \\ &= xlnx|_3^5 - \int_3^5 dx = (xlnx - x)|_3^5 = 5ln5 - 3ln3 - 2 \\ &\int_a^b f(x) dx = F(b) - F(a) = - \int_b^a f(x) dx = \int_a^c f(x) dx + \int_c^b f(x) dx; \\ &\frac{d}{dx} \int_{v(x)}^{u(x)} f(t) dt = u'(x) f[u(x)] - v'(x) f[v(x)] \end{split}$$

## Derivatives

Derivatives 
$$(cf)' = cf'(x) \qquad (fg)' = f'g + fg' \qquad (f \pm g)' = f'(x) \pm g'(x)$$
 
$$\frac{dx}{dx} = 1 \qquad (\frac{f}{g})' = \frac{(f'g - fg')}{g^2} \qquad (f(g(x)))' = f'(g(x))g'(x)$$
 
$$\frac{de^x}{dx} = e^x \qquad \frac{d(x^n)}{dx} = nx^{n-1} \qquad \frac{d\sin^n x}{dx} = \frac{1}{1+x^2}$$
 
$$\frac{d\cos_x x}{dx} = -\sin x \qquad \frac{d\tan x}{dx} = \sec^2 x \qquad \frac{d\sin(x-1)}{dx} = \ln x$$

Distribution	$\mathbf{CDF}$	P(X=x),f(x)	$\mu$	$EX^2$	Var	MGF	M'(t)	M"(t)	$M^n(t)$
Bern(p)		$p^x q^{1-x}, x \in \{1, 0\}$	p	p	pq	$pe^t + q$			
$\operatorname{Bin}(n,p)$	$I_{1-p}(n-x,x+1)$	$\binom{n}{x} p^x q^{n-x}; x \in \{0, 1n\}$	np	$\mu(\mu+q)$	$q\mu$	$(pe^t + q)^n$			
Geom(p)	$1 - q^x$	$pq^{x-1}, x \in 1, 2, \dots$	$\frac{1}{p}$	$\frac{p+2q}{p^2}$ $\frac{q^2+q}{p^2}$	$\frac{q}{p^2}$	$\frac{pe^t}{1-qe^t}, t < -\ln q$	4	_	
	$1 - q^{x+1}$	$pq^x, x \in 0, 1, \dots$	$\frac{q}{p}$	$\frac{q^2+q}{p^2}$	$\frac{q}{p^2}$	$\frac{p}{1-qe^t}, qe^t < 1$	$\frac{pqe^t}{(1-qe^t)^2}$	$\frac{2pqe^t}{(1-qe^t)^3} - M'(t)$	
$\operatorname{NBin}(r,p)$		$\binom{x-1}{r-1} p^r q^{x-r}, x \in r, r+1$	$\frac{r}{p}$		$\frac{rq}{p^2}$ $\frac{rq}{p^2}$	$(\frac{pe^t}{1-qe^t})^r$			
		$\binom{x+r-1}{r-1} p^r q^x, x \in 0, 1$	$\frac{rq}{p}$		$\frac{rq}{p^2}$	$(\frac{\vec{p}}{1 - qe^t})^r, qe^t < 1$			
$\mathrm{HGeom}(N,m,k)$		$\frac{\binom{m}{x}\binom{N-m}{k-x}}{\binom{N}{k}}$	$\frac{km}{N}$		$\mu \tfrac{(N-m)(N-k)}{N(N-1)}$				
$\mathrm{HGeom}(w,b,k)$		$\frac{\binom{w}{x}\binom{\acute{b}}{k-x}}{\binom{w+b}{k}}$	$\frac{kw}{w+b}$		$\mu_{\frac{b(w+b-k)}{(w+b)(w+b-1)}}$				
$Pois(\mu)$	$e^{-\mu} \sum_{i=0}^{x} \frac{\mu^i}{i!}$	$\frac{\mu^x}{x!}e^{-\mu}, x \in 0, 1$	$\mu$	$\mu^2 + \mu$	$\mu$	$e^{\mu(e^t-1)}$	$\mu e^t M(t)$	$\mu e^t (1 + \mu e^t) M(t)$	
Unif(n)		$\frac{1}{n}, x \in 1, 2n$	$\frac{n+1}{2}$	$\frac{(n+1)(2n+1)}{6}$	$\frac{(n^2-1)}{12}$	$\frac{\sum_{i=1}^{n} e^{ti}}{n}$			
$\mathrm{Unif}(a,b)$	$\frac{x-a}{b-a}$	$\frac{1}{b-a}, x \in (a,b)$	$\frac{a+b}{2}$		$\frac{(b-a)^2}{12}$	$\frac{e^{tb} - e^{ta}}{t(b-a)}$			
$\mathcal{N}(\mu,\sigma^2)$		$\frac{1}{\sigma\sqrt{2\pi}}e^{-\frac{(x-\mu)^2}{2\sigma^2}}$	$\mu$	$\mu^2 + \sigma^2$	$\sigma^2$	$e^{\mu t + \frac{\sigma^2 t^2}{2}}$	$(\mu+\sigma^2t)M(t)$	$[(\mu+\sigma^2t)^2+\sigma^2]M(t)$	
$\mathcal{N}(0,1)$		$\frac{1}{e^{-\frac{x^{2}}{2}}}$	0	1	1	$e^{\frac{t^2}{2}}$			
$\mathcal{LN}(\mu, \sigma^2)$		$\frac{1}{\frac{1}{x\sigma\sqrt{2\pi}}e^{\frac{-(\ln x - \mu)^2}{2\sigma^2}}}$	$e^{\mu + \frac{\sigma^2}{2}}$	$e^{2\mu+2\sigma^2}$	$\theta^2(e^{\sigma^2}-1)$	×			
$Cauchy(\theta, \sigma^2)$		$\pi \sigma \frac{1+(\frac{x-\theta}{2})^2}{1+(\frac{x-\theta}{2})^2}$	×	×	×				
$D$ Expo $(\mu, \sigma^2)$		$\frac{1}{2\pi\sigma}e^{-\left \frac{x-\mu}{\sigma}\right }$	$\mu$	$\mu^2 + 2\sigma^2$	$2\sigma^2$	$\frac{e^{\mu t}}{1 - \sigma^2 t^2}$			
$\text{Expo}(\lambda)$	$1 - e^{-\lambda x}$	$\lambda e^{-\lambda x}, x \in (0, \infty)$	$\frac{1}{\lambda}$		$\frac{1}{\lambda^2}$	$\frac{\lambda}{\lambda - t}, t < \lambda$			
$\operatorname{Expo}(\beta)$		$\frac{1}{\beta}e^{-\frac{x}{\beta}}$	β		$\beta^2$	$\frac{1}{1-\beta t}$	$\beta(1-\beta t)^{-2}$	$2\beta^2(1-\beta t)^{-3}$	
$\operatorname{Gamma}(a,\lambda)$		$\frac{\lambda^a}{\Gamma(a)} x^{a-1} e^{-\lambda x}, x \in (0, \infty)$	$\frac{a}{\lambda}$	$\frac{a}{\lambda^2}$	$\left(\frac{\lambda}{\lambda - t}\right)^a, t < \lambda$				
$\operatorname{Gamma}(\alpha,\beta)$		$\frac{1}{\Gamma(a)\beta^{\alpha}}x^{a-1}e^{-x/\beta}$	lphaeta	$\alpha \beta^2$	$\left(\frac{1}{1-\beta t}\right)^a, t < \frac{1}{\beta}$				
Beta(a,b)		$\frac{\Gamma(a+b)}{\Gamma(a)\Gamma(b)}x^{a-1}(1-x)^{b-1}$	$\frac{a}{a+b}$		$\frac{\mu(1-\mu)}{(a+b+1)}$				$\frac{\Gamma(\alpha+n)\Gamma(\alpha+\beta)}{\Gamma(\alpha+\beta+n)\Gamma(\alpha)}$
$B(\alpha, \beta) =$		$\Gamma(\alpha)\Gamma(\beta)$ $\sigma \in (0, 1)$		$\frac{a(a+1)}{(a+b)(a+b+1)}$	$\frac{ab}{(a+b)^2(a+b+1)}$				
$\chi_p^2$		$\frac{1}{2^{p/2}\Gamma(p/2)} x^{p/2-1} e^{-x/2}$	p	$2p + p^2$	2p	$(1-2t)^{-p/2}, t < \frac{1}{2}$			
$t_n$		$\frac{\Gamma(\frac{n+1}{2})}{\sqrt{n\pi}\Gamma(\frac{n}{2})} \left(1 + \frac{x^2}{n}\right)^{-\frac{n+1}{2}}$	0, n > 1		$\frac{n}{n-2}, n > 2$	×			