

4.1 4.2 Joint, Marginal, Conditional

$f_X(x) = \sum_{y \in \mathbf{R}} f_{X,Y}(x,y) = \int_{-\infty}^{\infty} f(x,y)dy$  4.1.6/10  $f(x|y) = P(X = x|Y = y) = \frac{f(x,y)}{f_Y(y)}$  4.2.1  
Indep  $f(x,y) = f_X(x)f_Y(y)$  4.2.5  $= g(x)h(y)$  4.2.7  $X, Y$  indep r.v.  $g(X), h(Y)$  indep 4.3.5  
 $F(x,y) = P(X \leq x, Y \leq y) = \int_{-\infty}^y \int_{-\infty}^x f(s,t)dsdt$  — Indep  $F_{X,Y}(x,y) = F_X(x)F_Y(y)$   
 $M_Z(t) = M_X(t)M_Y(t)$  4.2.12 —  $M_Z(t) = (M_X(t))^n$  4.6.7 —  $M_{\bar{X}}(t) = [M_X(\frac{t}{n})]^n$  5.2.7  
 $M_W(t) = M_X(t)M_Y(t) = e^{\mu_1(e^t-1)}e^{\mu_2(e^t-1)} = e^{(\mu_1+\mu_2)(e^t-1)}$  Use 4.2.12 proof 4.3.2  
 $U \sim \text{Geom}(\frac{1}{2}), u = 1, 2, ..$   $V \sim \text{NBin}(2, \frac{1}{2}), v = 2, 3, ..$   $\{(u,v) : u = 1, 2, ..; v = u + 1, u + 2, ..\}$  not indep

Expectations

$E[X] = \int_{-\infty}^{\infty} xf(x)dx = E[E(X|Y)]$  4.4.3  $E[g(x)] = \sum_{x \in D} h(x)p(x) = \int_{-\infty}^{\infty} g(x)f(x)dx$   
 $E[(X - \mu)^n] = \mu_n = \sum (x - \mu)^n p(x) = \int (x - \mu)^n f(x)dx$  —  $E(aX + b) = aE(X) + b$   
 $E[g(\bar{X})] = \sum \cdots \sum_{all \bar{x}} g(\bar{x})p(\bar{x}) = \int \cdots \int_{\mathbf{R}^n} g(\bar{x})f(\bar{x})d\bar{x}$   
 $E[g(X, Y)] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x,y)f(x,y)dxdy$  4.1.10  
 $E[g(X)|y] = \sum_x g(x)f(x|y) = \int_{-\infty}^{\infty} g(x)f(x|y)dx$  r.v.(y) 4.2.3  $E[g(X)h(Y)|y] = h(Y)E[g(X)|y]$   
 $E[XY] = E[X]E[Y] = E[g(x)]E[h(y)]$   $P(X \leq x, Y \leq y) = P(X \leq x)P(Y \leq y)$  indep 4.2.10  
 $E[g_1(X_1) \cdots g_n(X_n)] = E[g_1(X_1)] \cdots E[g_n(X_n)]$   $X_1, .., X_n$  indep 4.6.6  $E[g(X)|y] = E[g(X)]$  indep

Variances

$V[X] = \sigma_x^2 = E(X^2) - [E(X)]^2 = E[V(X|Y)] + V[E(X|Y)]$  4.4.7 —  $V[aX + b] = a^2\sigma^2$   
 $V[X|Y] = E[(X - E[X|Y])^2|Y] = \sum_x [x - E[X|Y]]^2 f(x|y) = E[X^2|y] - (E[X|y])^2$  4.2.4 —  
 $\sigma_{ax+b} = |a|\sigma_x$   
 $V[aX \pm bY] = a^2VX + b^2VY \pm 2abCov(X, Y)$  —  $V[X \pm Y] = VX + VY$  Indep 4.5.6  
 $E[\sum_{i=1}^n g(X_i)] = nE(g(X_1))$  —  $V[\sum_{i=1}^n g(X_i)] = nVar(g(X_1))$  r.s. 5.25

4.5 Covariance and Correlation

$Cov(X, Y) = E[(X - \mu_X)(Y - \mu_Y)] = E(XY) - E(X)E(Y) = E[XY] - \mu_X\mu_Y = \sigma_{XY}$  4.5.1/3  
 $Cov(aX, bY) = abCov(X, Y)$   $Cov(X, Y + Z) = Cov(X, Y) + Cov(X, Z)$   $Cov(X, c) = 0$   
 $\rho_{XY} = \frac{\sigma_{XY}}{\sigma_X\sigma_Y}$   $Corr(X, Y) = \frac{Cov(X,Y)}{\sqrt{Var(X)Var(Y)}}$  4.5.2

4.3 Transform

$f_{U,V}(u,v) = f_{X,Y}(h_1(u,v), h_2(u,v))|J|$  4.3.2  $= \sum_{i=1}^k f_{X,Y}(h_{1i}(u,v), h_{2i}(u,v))|J_i|$  4.3.5  
 $\begin{vmatrix} u = g_1(x,y) & x = h_1(u,v) \\ v = g_2(x,y) & y = h_2(u,v) \end{vmatrix} J = \begin{vmatrix} \frac{\partial x}{\partial u} & \frac{\partial y}{\partial v} \\ \frac{\partial x}{\partial u} & \frac{\partial y}{\partial v} \end{vmatrix} = \frac{\partial x}{\partial u} \frac{\partial y}{\partial v} - \frac{\partial y}{\partial u} \frac{\partial x}{\partial v}$   
 $f_Z(z) = \int_{-\infty}^{\infty} f_X(w)f_Y(z-w)dw$   $Z = X + Y$  5.2.9 Convolution

Distribution

$X \sim n(\mu, \sigma^2), Y \sim n(\gamma, \tau^2)$  —  $X + Y \sim n(\mu + \gamma, \sigma^2 + \tau^2)$  indep 4.2.14  
 $X, Y \sim n(0, 1)$  —  $X + Y, X - Y \sim n(0, 2), X/Y \sim \text{cauchy}(0, 1)$  indep 4.3.4/6  
 $X \sim \text{Poisson}(\theta), Y \sim \text{Poisson}(\lambda)$  —  $X + Y \sim \text{Poisson}(\theta + \lambda)$  indep 4.3.2  
 $X \sim \text{Beta}(\alpha, \beta), Y \sim \text{Beta}(\alpha + \beta, \gamma)$  —  $XY \sim n(\alpha, \beta + \gamma)$  indep 4.3.3  
 $X|Y \sim \text{Bin}(Y, p), Y|\Lambda \sim \text{Pois}(\Lambda), \Lambda \sim \text{Expo}(\beta)$  or  $X|Y \sim \text{Bin}(Y, p), Y \sim \text{NBin}(1, \frac{1}{1+\beta})$  4.4.5  
 $Y|\Lambda \sim \text{Pois}(\Lambda), \Lambda \sim \text{Gamma}(\alpha, \beta)$  then  $Y \sim \text{NBin}(\alpha, \frac{1}{1+\beta\beta})$  Pois-Gamma  
 $X|P \sim \text{Bin}(n, P), P \sim \text{Gamma}(\alpha, \beta)$   $EX = E[E(X|P)] = E[nP] = n\frac{\alpha}{\alpha+\beta}$  Beta-Bin 4.4.6  
 $V[X] = V[E(X|P)] + E[V(X|P)] = \frac{n^2\alpha\beta}{(\alpha+\beta)^2(\alpha+\beta+1)} + \frac{n\alpha\beta}{(\alpha+\beta)(\alpha+\beta+1)} = \frac{n\alpha\beta(\alpha+\beta+n)}{(\alpha+\beta)^2(\alpha+\beta+1)}$  Beta-Bin 4.4.8  
bivariate normal  $f_X(x) \sim n(\mu_X, \sigma_X^2)$  —  $f_{X|Y}(x|y) \sim n(\mu_X + \rho\frac{\sigma_X}{\sigma_Y}(x - \mu_Y), \sigma_X^2(1 - \rho^2))$   
4.5.10  $\rho_{XY} = \rho$   $f_Y(y) \sim n(\mu_Y, \sigma_Y^2)$  —  $aX + bY \sim n(a\mu_X + b\mu_Y, a^2\mu_X^2 + b^2\mu_Y^2 + 2ab\rho\sigma_X\sigma_Y)$   
 $X_1, ..X_n \sim \text{Gamma}(\alpha, \beta), X_1 + ..X_n \sim \text{Gamma}(\alpha_1 + ..\alpha_n, \beta)$   $\bar{X} \sim \text{Gamma}(n\alpha, \beta/n)$  indep 4.6.8  
 $X \sim \text{Cauchy}(0, \sigma), Y \sim \text{Cauchy}(0, \tau)$  —  $X + Y \sim \text{Cauchy}(0, \sigma + \tau)$  5.2.10  
 $X_1, .., X_n \sim \text{Cauchy}(0, \sigma)$  —  $\bar{X} \sim \text{Cauchy}(0, \sigma), \sum_1^n X \sim \text{Cauchy}(0, n\sigma)$

5.3 Sampling from N

$\bar{X} = \frac{X_1 + .. + X_n}{n} = \frac{1}{n} \sum_{i=1}^n X_i$  5.2.2  $\bar{X}_{n+1} = \frac{\sum_{i=1}^{n+1} X_i}{n+1} = \frac{\sum_{i=1}^n X_i + X_{n+1}}{n+1} = \frac{n\bar{X}_n + X_{n+1}}{n+1}$  5.3.1  
 $S^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i^2 - n\bar{X}^2)$  5.2.3  $nS_{n+1}^2 = (n-1)S_n^2 + (\frac{n}{n+1})(X_{n+1} - \bar{X}_n)^2$   
5.2.6  $X_1, .., X_n$  iid |  $E[\bar{X}] = \mu$  |  $E[S^2] = \sigma^2$  |  $V[\bar{X}] = \frac{\sigma^2}{n}$  | W/o Normal  
5.3.1  $\sim n(\mu, \sigma^2)$  |  $\bar{X}, S^2$  indep |  $\bar{X} \sim N(\mu, \frac{\sigma^2}{n})$  |  $\frac{(n-1)S^2}{\sigma^2} \sim \chi_{n-1}^2$  | W Normal

$X \sim n(\mu, \sigma^2), \frac{x-\mu}{\sigma}, \frac{\bar{x}-\mu}{\sigma/\sqrt{n}} \sim n(0, 1), \frac{\bar{X}-\mu}{S/\sqrt{n}} = \frac{\bar{x}-\mu}{\sigma/\sqrt{n}} = \frac{U}{\sqrt{S^2/\sigma^2}} = \frac{U}{\sqrt{\frac{\chi_{n-1}^2}{n-1}}} \sim t_{n-1}$  5.3.4 —  $t_1 = \text{Cauchy}(0, 1)$

$x_i \sim n(0, 1), \sum_{i=1}^n x_i^2 \sim \chi_n^2$  —  $x_i \sim n(0, \sigma^2), \sum_{i=1}^n x_i^2 \sim \sigma^2\chi_n^2, \sum_{i=1}^n (x_i - \bar{x})^2 \sim \sigma^2\chi_{n-1}^2$   
 $\chi_2^2 \Leftrightarrow \text{Expo}(2)$   $\chi_p^2 \sim \text{Gamma}(\frac{p}{2}, 2)$  4.6.8  
 $X_1, ..X_n \sim \chi_{p_i}^2, X_1 + ..X_n \sim \chi_{p_1+..p_n}^2$  5.3.2  $U \sim \chi_m^2, V \sim \chi_n^2, U + V \sim \chi_{m+n}^2$   
 $X_i \sim n(\mu_X, \sigma_X^2)Y_j \sim n(\mu_Y, \sigma_Y^2), X_1..X_n, Y_1..Y_m$  indep,  $\frac{S_X^2}{\sigma_X^2} \sim \chi^2 - \frac{S_X^2/\sigma_X^2}{S_Y^2/\sigma_Y^2} \sim F$  5.3.5  
 $X \sim F_{p,q}, \frac{1}{X} \sim F_{q,p}$  —  $X \sim T_q, X^2 \sim F_{1,q}$  —  $X \sim F_{p,q}, \frac{\frac{p}{1}X}{\frac{q}{1}X} \sim \text{Beta}(\frac{p}{2}, \frac{q}{2})$

$V\chi_{n-1}^2 = V[\frac{(n-1)S^2}{\sigma^2}] = \frac{(n-1)^2}{\sigma^4}Var[S^2] = 2(n-1) \implies Var[S^2] = \frac{2\sigma^4}{n-1}$

5.4 Order statistics

5.4.4  $f_K(x) = K\binom{n}{k}[F_X(x)]^{k-1}[1 - F_X(x)]^{n-k}f(x)$  or  $\frac{n!}{(k-1)!(n-k)!}$  1-29p9  
5.4.6  $f_{X_{(1)}, ..., X_{(n)}}(x_1, ..., x_n) = \begin{cases} n!f_X(x_1) \cdots f_X(x_n) & -\infty < x_1 < .. < x_n < \infty \\ 0 & \text{otherwise} \end{cases}$

$f_{X_{(i)}, X_{(j)}}(u, v) = \frac{n!}{(i-1)!(j-1-i)!(n-j)!}f_X(u)f_X(v)[F_X(u)]^{i-1}[F_X(v) - F_X(u)]^{j-1-i}[1 - F_X(v)]^{n-j}$   
**1 from N to T to Chi to F**

Given some function of these, find the distribution.

2 Transformation of pairs of r.v.s

Given 2 r.v.s, f(x,y), and a function of them, find its distribution. 1-10p1  
 $f(x,y) = \frac{1}{4}e^{-\frac{x+y}{2}}, 0 < x < \infty, 0 < y < \infty, u = \frac{X-Y}{2}$   
1.  $\begin{vmatrix} U = \frac{x-y}{2} & X = 2u + v \\ V = y & Y = v \end{vmatrix}$  2.  $J = \begin{vmatrix} \frac{\partial x}{\partial u} & \frac{\partial x}{\partial v} \\ \frac{\partial y}{\partial u} & \frac{\partial y}{\partial v} \end{vmatrix} = \begin{vmatrix} 2 & 1 \\ 0 & 1 \end{vmatrix} = 2$   
3.  $g(u,v) = f(x,y)|J| = \frac{1}{4}e^{-\frac{x+y}{2}}2 = \frac{1}{2}e^{-\frac{2u+v+v}{2}} = \frac{1}{2}e^{-(u+v)}$   
4.  $0 < x < \infty, 0 < y < \infty \implies 0 < 2u + v < \infty, 0 < v < \infty \implies v > -2u$   
5. Double Exponential(Laplace)  
 $g_U(u) = \begin{vmatrix} \int_{-2u}^{\infty} \frac{1}{2}e^{-(u+v)}dv = \frac{1}{2}e^{-u} \int_{-2u}^{\infty} e^{-v}dv = \frac{1}{2}e^{-u} [-e^{-v}]_{-2u}^{\infty} = \frac{1}{2}e^{-u} [0 + e^{2u}] \\ \frac{1}{2}e^{|u|} \int_0^{\infty} \frac{1}{2}e^{-(u+v)}dv = \frac{1}{2}e^{-u} \int_0^{\infty} e^{-v}dv = \frac{1}{2}e^{-u} [-e^{-v}]_0^{\infty} = \frac{1}{2}e^{-u} [0 + 1] \end{vmatrix} \begin{matrix} u < 0 \\ u \geq 0 \end{matrix}$

3

Given  $f(x,y)$ , find the  
 $f_X(x) = -E[X] = -V[X] =$   
 $f_Y(y) = -E[Y] = -V[Y] =$   
 $f(x,y) = -E[XY] =$   
 $Cov(x,y) = E[XY] - EXEY =$   
 $f(X|Y) = -E[X|Y] = -V[X|Y] =$   
 $\rho = Cov(x,y)/\sqrt{V[X]V[Y]} =$   
4.5.7  $E[Y|X] = a + bx, E[Y] = E[E[Y|X]] = E[a + bX] = a + bE[X]$  by 4.4.3,  
 $E[XE[Y|X]] = E[X(a + bX)] = aE[X] + bE[X^2], E[XE[Y|X]] = \int_{-\infty}^{\infty} xE[Y|x]f_X(x)dx$  by 2.2.1,  
 $= \int_{-\infty}^{\infty} x \left[ \int_{-\infty}^{\infty} yf_Y(y|x)dy \right] f_X(x)dx$  by 4.2.3,  $= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} xyf(x,y)dxdy = E[XY]$  by 4.1.10,  
 $\sigma_{XY} = E[XY] - \mu_X\mu_Y = a\mu_X + bE[X^2] - \mu_X\mu_Y = a\mu_X + b(\sigma_X^2 + \mu_X^2) - \mu_X(a + b\mu_X) = b\sigma_X^2$   
 $\rho_{XY} = \frac{\sigma_{XY}}{\sigma_X\sigma_Y} = \frac{b\sigma_X^2}{\sigma_X\sigma_Y} = b\frac{\sigma_X}{\sigma_Y}$   
**4 Order statistics**  
Find the distribution of  $X(k)$  or  $X_{(j)}, X_{(k)}$   
joint pmf 1-31p2-3  
5.4.5 uniform order pdf 1-31p4-8  
**pdf/pmf**  
CDF:  $F(x) = P(X \leq x) = \sum_{y: y \leq x} p(y) = \int_{-\infty}^x f(y)dy$

**probabilities:**  $a \leq b$ :  
PMF:  $p(x) = P(X = x) = P(\forall w \in \mathcal{W} : X(w) = x)$   
 $P(a \leq X \leq b) = F(b) - F(a^-); P(a < X \leq b) = F(b) - F(a); P(a \leq X < a) = p(a);$   
 $P(a < X < b) = F(b^-) - F(a);$  (where  $a^-$  is the largest possible X value strictly less than a); Taking  
 $a = b$  yields  $P(X = a) = F(a) - F(a^-)$  as desired.  
PDF:  $P(\forall w \in \mathcal{W} : a \leq X(w) \leq b) = \int_a^b f(x)dx$   
 $P(a \leq X \leq b) = P(a < X \leq b) = P(a < X < b); P(X > a) = 1 - F(a); P(a \leq X \leq b) = F(b) - F(a)$

CDF

Condition: $f(x) \geq 0 \forall x$  pmf  $\sum_x f_X(x) = 1, pdf \int_{-\infty}^{\infty} f(x)dx = 1$

mgf

$E(e^{tx}) = \int e^{tx} f(x)dx = \sum e^{tx} f(x); M_{aX+b}(t) = e^{tb} M_{aX}(t)$

$M_X(t) = E(e^{tx})$	$M_X(0) = 1$
$M'_X(t) = E(xe^{tx})$	$M'_X(0) = E(X)$
$M''_X(t) = E(x^2e^{tx})$	$M''_X(0) = E(X^2)$
$M^n_X(t) = E(x^ne^{tx})$	$M^n_X(0) = E(X^n)$

transform

$g(x) \uparrow$	$F_Y(y) = F_X(g^{-1}(y))$
$g(x) \downarrow$	$F_Y(y) = 1 - F_X(g^{-1}(y))$

**not monotone**  $\because X \leq 0$  is  $\emptyset \therefore P(X \leq -\sqrt{y}) = 0$   
 $F_Y(y) = P(Y \leq y) = P(X^2 \leq y) = P(-\sqrt{y} \leq X \leq \sqrt{y}) = P(X \leq \sqrt{y}) - P(X \leq -\sqrt{y}) = F_X(\sqrt{y}) = \sqrt{y}, 0 < \sqrt{y} < 1$

**monotone:** $f_Y(y) = f_X(g^{-1}(y))| \frac{d(g^{-1}(y))}{dy} |$

Series

Finite	Binomial
$\sum_{k=1}^n k = \frac{n(n+1)}{2}$	$\sum_{k=0}^n \binom{n}{k} = 2^n$
$\sum_{k=1}^n (2k-1) = n^2$	$\sum_{k=0}^n \binom{r+k}{k} = \binom{r+n+1}{n}$
$\sum_{k=1}^n k^2 = \frac{n(n+1)(2n+1)}{6}$	$\sum_{k=0}^n \binom{k}{m} = \binom{n+1}{m+1}$
$\sum_{k=1}^n k^3 = (\frac{n(n+1)}{2})^2$	$\sum_{k=0}^r \binom{m}{k} \binom{n}{r-k} = \binom{m+n}{r}$
$\sum_{k=0}^n c^k = \frac{c^{n+1}-1}{c-1}$	$\sum_{k=0}^n \binom{n}{k} a^{n-k} b^k = (a+b)^n$
Infinite $ p  < 1$	$\sum_{k=0}^{\infty} kp^{k-1} = \frac{d}{dp} (\sum_{k=0}^{\infty} p^k) = \frac{1}{(1-p)^2}$
$\sum_{k=0}^{\infty} p^k = \frac{1}{1-p}$	$\sum_{k=0}^{\infty} \binom{r+k-1}{k} x^k = (1-x)^{-r} r \in \mathbb{N}^+$
$\sum_{k=1}^{\infty} p^k = \frac{p}{1-p}$	$\sum_{k=0}^{\infty} \binom{\alpha}{k} p^k = (1+p)^{\alpha} \quad  p  < 1, \alpha \in \mathbb{C}$
$\Gamma(n) = (n-1)!$	$\sum_{n=0}^{\infty} \frac{x^n}{n!} = e^x$
$\Gamma(a+1) = a\Gamma(a)$	$\sum_{n=0}^{\infty} ar^n = \frac{a}{1-r},  r  < 1$
$\Gamma(1/2) = \sqrt{\pi}$	$\sum_{k=0}^{n-1} ar^n = \frac{a(1-r^n)}{1-r}$
$\Gamma(0) = \Gamma(-1) = \infty$	$\int_0^1 x^{a-1} (1-x)^{b-1} dx = \frac{\Gamma(a)\Gamma(b)}{\Gamma(a+b)}$
$\Gamma(-1/2) = -2\Gamma(1/2)$	$\int_0^{\infty} x^{t-1} e^{-x} dx = \Gamma(t)$

Integrals+c

$\int k dx = kx$	$\int_a^b c dx = c(b-a)$	$ \int_a^b f(x) dx  \leq \int_a^b  f(x)  dx$
$\int e^u du = e^u$	$\int \ln u du = u \ln(u) - u$	$\int \frac{1}{a^2+u^2} du = \frac{1}{a} \tan^{-1}(\frac{u}{a})$
$\int x^n = \frac{x^{n+1}}{n+1}$	$\int \frac{1}{ax+b} = \frac{1}{a} \ln ax+b $	$\int \frac{1}{\sqrt{a^2-u^2}} du = \frac{1}{\sin(\frac{u}{a})}$
$\int \ln aa^x = a^x$		

**Substitution**

$u = g(x)$	$du = g'(x)dx$	$\int_a^b f(g(x))g'(x)dx = \int_{g(a)}^{g(b)} f(u)du$
$x^3$	$du = 3x^2 dx$	$\int_1^2 5x^2 \cos(x^3)dx = \int_1^{8\frac{5}{3}} \cos(u)du$

**Integreation by parts**

$u$	$x$	$\frac{du}{dx}$	$\frac{dv}{dx}$	$v$	$\int_a^b u dv$
$x$	$\ln x$	$\frac{1}{x}$	$\frac{1}{x}$	$e^{-x}$	$\int_a^b x e^{-x} dx$
				$x$	$\int_3^5 x \ln x dx$

$= uv|_a^b - \int_a^b v du$   
 $= -xe^{-x} + \int xe^{-x} dx = -xe^{-x} - e^{-x} + c$   
 $= x \ln x|_3^5 - \int_3^5 dx = (x \ln x - x)|_3^5 = 5 \ln 5 - 3 \ln 3 - 2$   
 $\int_a^b f(x)dx = F(b) - F(a) = - \int_b^a f(x)dx = \int_a^c f(x)dx + \int_c^b f(x)dx;$   
 $\frac{d}{dx} \int_{v(x)}^{u(x)} f(t)dt = u'(x)f[u(x)] - v'(x)f[v(x)]$

Derivatives

$(cf)' = cf'(x)$	$(fg)' = f'g + fg'$	$(f \pm g)' = f'(x) \pm g'(x)$
$\frac{dx}{dx} = 1$	$(\frac{f}{g})' = \frac{(f'g-fg')}{g^2}$	$(f(g(x)))' = f'(g(x))g'(x)$
$\frac{de^x}{dx} = e^x$	$\frac{d(x^n)}{dx} = nx^{n-1}$	$\frac{d \ln(x)}{dx} = \frac{1}{x}, x > 0$
$\frac{da^x}{dx} = a^x \ln(a)$	$\frac{d \tan^{-1} x}{dx} = \frac{1}{1+x^2}$	$\frac{d \log_a(x)}{dx} = \frac{1}{x \ln a}, x > 0$
$\frac{d \cos x}{dx} = -\sin x$	$\frac{d \tan x}{dx} = \sec^2 x$	$\frac{dx(\ln x - 1)}{dx} = \ln x$
	Replace	No Replace

Fixed # trials ( $n$ )	Binomial (Bern if $n = 1$ )	HGeom
Draw until $r$ success	NBin (Geom if $r = 1$ )	NHGeom

$U \sim Geom(\frac{1}{2}), u = 1, 2..$  # of trials needed to get the first head.  
 $V \sim NBin(2, \frac{1}{2}r), v = 2, 3..$  # of trials needed to get two heads in repeated tosses of a fair coin.

Distribution	CDF	P(X=x),f(x)	$\mu$	$EX^2$	Var	MGF	M'(t)	M''(t)	$M^n(t)$
Bern( $p$ )		$p^x q^{1-x}, x \in \{1, 0\}$	$p$	$p$	$pq$	$pe^t + q$			
Bin( $n, p$ )	$I_{1-p}(n-x, x+1)$	$\binom{n}{x} p^x q^{n-x}; x \in \{0, 1..n\}$	$np$	$\mu(\mu + q)$	$\mu q$	$(pe^t + q)^n$			
Geom( $p$ )	$1 - q^x$	$pq^{x-1}, x \in 1, 2, ..$	$\frac{1}{p}$	$\frac{p+2q}{p^2}$	$\frac{q}{p^2}$	$\frac{pe^t}{1-qe^t}, t < -\ln q$			
	$1 - q^{x+1}$	$pq^x, x \in 0, 1, ..$	$\frac{q}{p}$	$\frac{q^2+q}{p^2}$	$\frac{q}{p^2}$	$\frac{p}{1-qe^t}, qe^t < 1$	$\frac{pqe^t}{(1-qe^t)^2}$	$\frac{2pqe^t}{(1-qe^t)^3} - M'(t)$	
NBin( $r, p$ )		$\binom{x-1}{r-1} p^r q^{x-r}, x \in r, r+1..$	$\frac{r}{p}$		$\frac{rq}{p^2}$	$(\frac{pe^t}{1-qe^t})^r$			
		$\binom{x+r-1}{r-1} p^r q^x, x \in 0, 1..$	$\frac{rq}{p}$		$\frac{rq}{p^2}$	$(\frac{p}{1-qe^t})^r, qe^t < 1$			
HGeom( $N, m, k$ )		$\frac{\binom{m}{x} \binom{N-m}{k-x}}{\binom{N}{k}}$	$\frac{km}{N}$		$\mu \frac{(N-m)(N-k)}{N(N-1)}$				
HGeom( $w, b, k$ )		$\frac{\binom{w}{x} \binom{b}{k-x}}{\binom{w+b}{k}}$	$\frac{k w}{w+b}$		$\mu \frac{b(w+b-k)}{(w+b)(w+b-1)}$				
Pois( $\mu$ )	$e^{-\mu} \sum_{i=0}^x \frac{\mu^i}{i!}$	$\frac{\mu^x}{x!} e^{-\mu}, x \in 0, 1..$	$\mu$	$\mu^2 + \mu$	$\mu$	$e^{\mu(e^t-1)}$	$\mu e^t M(t)$	$\mu e^t (1 + \mu e^t) M(t)$	
Unif( $n$ )		$\frac{1}{n}, x \in 1, 2..n$	$\frac{n+1}{2}$	$\frac{(n+1)(2n+1)}{6}$	$\frac{(n^2-1)}{12}$	$\sum_{i=1}^n \frac{e^{ti}}{n}$			
Unif( $a, b$ )	$\frac{x-a}{b-a}$	$\frac{1}{b-a}, x \in (a, b)$	$\frac{a+b}{2}$		$\frac{(b-a)^2}{12}$	$\frac{e^{tb}-e^{ta}}{t(b-a)}$			
$\mathcal{N}(\mu, \sigma^2)$		$\frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$	$\mu$	$\mu^2 + \sigma^2$	$\sigma^2$	$e^{\mu t + \frac{\sigma^2 t^2}{2}}$	$(\mu + \sigma^2 t) M(t)$	$[(\mu + \sigma^2 t)^2 + \sigma^2] M(t)$	
$\mathcal{N}(0, 1)$		$\frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}}$	0	1	1	$e^{-\frac{t^2}{2}}$			
$\mathcal{LN}(\mu, \sigma^2)$		$\frac{1}{x\sigma\sqrt{2\pi}} e^{\frac{-(\ln x - \mu)^2}{2\sigma^2}}$	$e^{\mu + \frac{\sigma^2}{2}}$	$e^{2\mu + 2\sigma^2}$	$\theta^2 (e^{\sigma^2} - 1)$	$\times$			
Cauchy( $\theta, \sigma^2$ )		$\frac{1}{\pi\sigma} \frac{1}{1+(\frac{x-\theta}{\sigma})^2}$	$\times$	$\times$	$\times$				
DExpo( $\mu, \sigma^2$ )		$\frac{1}{2\pi\sigma} e^{- \frac{x-\mu}{\sigma} }$	$\mu$	$\mu^2 + 2\sigma^2$	$2\sigma^2$	$\frac{e^{\mu t}}{1-\sigma^2 t^2}$			
Expo( $\lambda$ )	$1 - e^{-\lambda x}$	$\lambda e^{-\lambda x}, x \in (0, \infty)$	$\frac{1}{\lambda}$		$\frac{1}{\lambda^2}$	$\frac{\lambda}{\lambda-t}, t < \lambda$			
Expo( $\beta$ )		$\frac{1}{\beta} e^{-\frac{x}{\beta}}$	$\beta$		$\beta^2$	$\frac{1}{1-\beta t}$	$\beta(1-\beta t)^{-2}$	$2\beta^2(1-\beta t)^{-3}$	
Gamma( $a, \lambda$ )		$\frac{\lambda^a}{\Gamma(a)} x^{a-1} e^{-\lambda x}, x \in (0, \infty)$	$\frac{a}{\lambda}$	$\frac{a}{\lambda^2}$	$\left(\frac{\lambda}{\lambda-t}\right)^a, t < \lambda$				
Gamma( $\alpha, \beta$ )		$\frac{1}{\Gamma(a)\beta^\alpha} x^{a-1} e^{-x/\beta}$	$\alpha\beta$	$\alpha\beta^2$	$\left(\frac{1}{1-\beta t}\right)^a, t < \frac{1}{\beta}$				
Beta( $a, b$ )		$\frac{\Gamma(a+b)}{\Gamma(a)\Gamma(b)} x^{a-1} (1-x)^{b-1}$	$\frac{a}{a+b}$		$\frac{\mu(1-\mu)}{(a+b+1)}$				$\frac{\Gamma(\alpha+n)\Gamma(\alpha+\beta)}{\Gamma(\alpha+\beta+n)\Gamma(\alpha)}$
$B(\alpha, \beta) =$		$\frac{\Gamma(\alpha)\Gamma(\beta)}{\Gamma(\alpha+\beta)}, x \in (0, 1)$		$\frac{a(a+1)}{(a+b)(a+b+1)}$	$\frac{ab}{(a+b)^2(a+b+1)}$				
$\chi_p^2$		$\frac{x^{\frac{p}{2}-1}}{\Gamma(\frac{p}{2}) 2^{\frac{p}{2}}} e^{-\frac{x}{2}}$	$p$	$2p + p^2$	$2p$	$(1-2t)^{-p/2}, t < \frac{1}{2}$			
$t_p$		$\frac{\Gamma(\frac{p+1}{2})}{\Gamma(\frac{p}{2})\sqrt{p\pi}} \left(1 + \frac{x^2}{p}\right)^{-\frac{p+1}{2}}$	$0, p > 1$		$\frac{p}{p-2}, p > 2$	$\times$			
$F$	$x > 0$	$\frac{\Gamma(\frac{p+q}{2})}{\Gamma(\frac{p}{2})\Gamma(\frac{q}{2})} \left(\frac{p}{q}\right)^{\frac{p}{2}} \frac{x^{\frac{p}{2}-1}}{(1+\frac{p}{q}x)^{\frac{p+q}{2}}}$	$\frac{q}{q-2}$	$q > 2$	$2\left(\frac{q}{q-2}\right)^2 \frac{p+q-2}{p(q-4)}$	$q > 4$			
Arcsine	$\frac{1}{\pi \arcsin \sqrt{x}}$	$\frac{1}{\pi \sqrt{x(1-x)}}, x \in [0, 1]$	$\frac{1}{2}$		$\frac{1}{8}$				Beta( $\frac{1}{2}, \frac{1}{2}$ )
Dirichlet	$B(a) = \frac{\prod_{i=1}^k \Gamma(a_i)}{\Gamma(\sum_{i=1}^k a_i)}$	$\frac{1}{B(a)} \prod_{i=1}^k x_i^{a_i-1}, x \in (0, 1)$	$\sum_k \frac{a_i}{a_k}$	$\sum_{i=1}^k x_i = 1$	$\frac{a_i(a_0-a_i)}{a_0^2(a_0+1)}$	$Cov(X_i, X_j) =$	$\frac{-a_i a_j}{a_0^2(a_0+1)}$	$a_0 = \sum_{i=1}^k a_i$	