

1. (Poisson Regression) The independent random variables $Y_i, i = 1, 2, \dots, n$ represent the outcomes of a Poisson experiment where the mean μ_i is proportional to the value of x_i . That is, $Y_i \sim \text{Poisson}(\mu_i)$ and $\mu_i = \gamma x_i$. Assume that the x_i , values are known constants.

a) Find the MLE of γ

$$L(\gamma) = \prod_{i=1}^n \left(\frac{\mu_i^{y_i}}{y_i!} e^{-\mu_i} \right) = \prod_{i=1}^n \frac{(\gamma x_i)^{y_i} e^{-\gamma x_i}}{y_i!} = \frac{\gamma^{\sum_{i=1}^n y_i} \prod_{i=1}^n x_i^{y_i}}{\prod_{i=1}^n y_i!} e^{-\gamma \sum_{i=1}^n x_i}, \quad y_i \in 0, 1, 2, \dots$$

$$l(\gamma) = \ln \gamma \sum_{i=1}^n y_i + \sum_{i=1}^n x_i^{y_i} - \sum_{i=1}^n \ln y_i! - \gamma \sum_{i=1}^n x_i$$

$$l'(\gamma) = \frac{\sum_{i=1}^n y_i}{\gamma} - \sum_{i=1}^n x_i \stackrel{\text{set}}{=} 0$$

$$\hat{\gamma}_{MLE} = \frac{\sum_{i=1}^n y_i}{\sum_{i=1}^n x_i}$$

b) Find the mean and variance of $\hat{\gamma}_{MLE}$

For x_i are known constants. $Y_i \sim \text{Poisson}(\mu_i)$, $E[y_i] = \text{Var}[y_i] = \mu_i = \gamma x_i$,

$$E[\hat{\gamma}_{MLE}] = \frac{E[\sum_{i=1}^n y_i]}{\sum_{i=1}^n x_i} = \frac{\sum_{i=1}^n E[y_i]}{\sum_{i=1}^n x_i} = \frac{\sum_{i=1}^n \gamma x_i}{\sum_{i=1}^n x_i} = \gamma$$

For Y_i are independent random variables, $\text{Cov}(y_i, y_j) = 0, i \neq j$, $\text{Var}[\sum_{i=1}^n y_i] = \sum_{i=1}^n \text{Var}[y_i]$

$$\text{Var}[\hat{\gamma}_{MLE}] = \frac{\text{Var}[\sum_{i=1}^n y_i]}{(\sum_{i=1}^n x_i)^2} = \frac{\sum_{i=1}^n \text{Var}[y_i]}{(\sum_{i=1}^n x_i)^2} = \frac{\sum_{i=1}^n \gamma x_i}{(\sum_{i=1}^n x_i)^2} = \frac{\gamma}{\sum_{i=1}^n x_i}$$

2. Consider the regression model $Y_i = \beta_0 + \beta_1 x_i + \varepsilon_i, i = 1, \dots, n$. Find the maximum likelihood estimates of the parameters if:

a) $\varepsilon_i \sim N(0, \sigma^2 x_i^2)$, independent for $i = 1, \dots, n$.

For $E[\varepsilon_i] = 0$, $\text{Var}[\varepsilon_i] = \sigma^2 x_i^2$, x_i and ε_i are independent,

$$E[y_i] = E[\beta_0 + \beta_1 x_i] + E[\varepsilon_i] = \beta_0 + \beta_1 x_i$$

$$\text{Var}[y_i] = \text{Var}[\beta_0 + \beta_1 x_i] + \text{Var}[\varepsilon_i] = \sigma^2 x_i^2$$

From Slutsky's theorem? C.L.T.?

$$Y_i \sim N(\beta_0 + \beta_1 x_i, \sigma^2 x_i^2)$$

$$f_Y(y_i) = \frac{1}{x_i \sqrt{2\pi\sigma^2}} e^{\frac{-1}{2\sigma^2 x_i^2} (y_i - \beta_0 - \beta_1 x_i)^2}$$

$$L(\sigma) = \prod_{i=1}^n (x_i \sqrt{2\pi\sigma^2})^{-1} e^{\sum_{i=1}^n \frac{-1}{2\sigma^2 x_i^2} (y_i - \beta_0 - \beta_1 x_i)^2} = (2\pi\sigma^2)^{-\frac{n}{2}} \left(\prod_{i=1}^n x_i \right)^{-1} e^{\sum_{i=1}^n \frac{-1}{2\sigma^2 x_i^2} (y_i - \beta_0 - \beta_1 x_i)^2}$$

$$l(\sigma) = -n \ln \sigma^2 - \frac{n}{2} \ln(2\pi) - \ln\left(\prod_{i=1}^n x_i\right) - \sum_{i=1}^n \frac{1}{\sigma^2 x_i^2} (y_i - \beta_0 - \beta_1 x_i)^2$$

$$l'(\sigma) = -\frac{n}{\sigma^2} + \frac{1}{\sigma^4} \sum_{i=1}^n \frac{1}{x_i^2} (y_i - \beta_0 - \beta_1 x_i)^2 \stackrel{\text{set}}{=} 0$$

$$\hat{\sigma}_{MLE} = \frac{1}{n} \sum_{i=1}^n \frac{1}{x_i^2} (y_i - \beta_0 - \beta_1 x_i)^2$$

b) $\varepsilon_i \sim i.i.d. f(\varepsilon; \lambda) = \frac{\lambda}{2} e^{-\lambda|x|}.$

$$\varepsilon_i = y_i - \beta_0 - \beta_1 x_i +$$

$$f_Y(y_i) = \frac{\lambda}{2} e^{-\lambda|y_i - \beta_0 - \beta_1 x_i|}$$

$$L(\lambda) = \prod_{i=1}^n \left(\frac{\lambda}{2} e^{-\lambda|y_i - \beta_0 - \beta_1 x_i|} \right) = \lambda^n 2^{-n} e^{-\lambda \sum_{i=1}^n |y_i - \beta_0 - \beta_1 x_i|}$$

$$l(\lambda) = n \ln \lambda - n \ln 2 - \lambda \sum_{i=1}^n |y_i - \beta_0 - \beta_1 x_i|$$

$$l'(\lambda) = \frac{n}{\lambda} - \sum_{i=1}^n |y_i - \beta_0 - \beta_1 x_i| \stackrel{\text{set}}{=} 0$$

$$\hat{\lambda}_{MLE} = \frac{n}{\sum_{i=1}^n |y_i - \beta_0 - \beta_1 x_i|}$$

3. Finde the finite breakdown point and the infinite breakdown point for

a) the Mean Absolute Deviation, or $\frac{1}{n} \sum_{i=1}^n |X_i - \bar{X}_i|.$

The finite breakdown point is the smallest proportion m/n of the sample values such that $|\hat{\theta}^* - \hat{\theta}|$ can be made arbitrarily large by corrupting m data values and computing $\hat{\theta}^*$, where n is the sample size, $\hat{\theta}$ is the estimator. The limit as $n \rightarrow \infty$ is called the breakdown point.

Replace X_i with X_i^*

$$|\hat{\theta}^* - \hat{\theta}| = \left| \frac{1}{n} \sum_{i=1}^n |X_n^* - \bar{X}_i| - \frac{1}{n} \sum_{i=1}^n |X_n - \bar{X}_i| \right| = \frac{1}{n} |X_n^* - X_n| = \frac{1}{n}$$

The breakdown point = $\frac{1}{n}$

$$\lim_{n \rightarrow \infty} |\hat{\theta}^* - \hat{\theta}| = \lim_{n \rightarrow \infty} \frac{1}{n} = 0$$

b) the Median Absolute Deviation, or $\text{Median}\{(X_1 - \bar{X}_i), \dots, (X_n - \bar{X}_i)\}.$

When n is even,

$$|\hat{\theta}^* - \hat{\theta}| = |(X_{\frac{n}{2}}^* - \bar{X}_i) - (X_{\frac{n}{2}} - \bar{X}_i)| = |X_{\frac{n}{2}}^* - X_{\frac{n}{2}}|$$

$$\text{The breakdown point} = \frac{n/2}{n} = \frac{1}{2}$$

$$\lim_{n \rightarrow \infty} |\hat{\theta}^* - \hat{\theta}| = \lim_{n \rightarrow \infty} \left(\frac{1}{2}\right) = \frac{1}{2}$$

When n is odd,

$$|\hat{\theta}^* - \hat{\theta}| = |(X_{\frac{n+1}{2}}^* - \bar{X}_i) - (X_{\frac{n+1}{2}} - \bar{X}_i)| = |X_{\frac{n+1}{2}}^* - X_{\frac{n+1}{2}}|$$

$$\text{The breakdown point} = \frac{(n+1)/2}{n} = \frac{1}{2} + \frac{1}{2n}$$

$$\lim_{n \rightarrow \infty} |\hat{\theta}^* - \hat{\theta}| = \lim_{n \rightarrow \infty} \left(\frac{1}{2} + \frac{1}{2n}\right) = \frac{1}{2}$$

4. Assume that X_1, X_2, \dots, X_n are i.i.d. $Uniform(a, b)$. Find the asymptotic relative efficiency of the sample median to the sample mean.

$$\text{For } X \sim Uniform(a, b), E[X] = \frac{a+b}{2}, Var[X] = \frac{(b-a)^2}{12}, \bar{X} = \frac{1}{n} \sum_{i=1}^n x_i,$$

$$E[\bar{X}] = E\left[\frac{1}{n} \sum_{i=1}^n x_i\right] = \frac{1}{n} \sum_{i=1}^n E[x_i] = \frac{1}{n} \sum_{i=1}^n \frac{a+b}{2} = \frac{a+b}{2}$$

For X_i are independent,

$$Var[\bar{X}] = Var\left[\frac{1}{n} \sum_{i=1}^n x_i\right] = \frac{1}{n^2} \sum_{i=1}^n Var[x_i] = \frac{1}{n^2} \sum_{i=1}^n \frac{(b-a)^2}{12} = \frac{(b-a)^2}{12n}$$

From Slutsky's theorem? C.L.T.?

$$\bar{X} \sim N\left(\frac{a+b}{2}, \frac{(b-a)^2}{12n}\right)$$

For large n , the sample median $m_n \approx N(M, \frac{1}{4nf^2(M)})$, where M is the population median, $f(x)$ is the p.d.f. of X

$$E[m_n] = M$$

$$Var[m_n] = \frac{1}{4nf^2(M)} = \frac{(b-a)^2}{4n}$$

The asymptotic relative efficiency of m_n to \bar{X}

$$= \frac{Var[\bar{X}]}{Var[m_n]} = \frac{\frac{(b-a)^2}{12n}}{\frac{(b-a)^2}{4n}} = \frac{1}{3}$$

Therefore, the sample mean is asymptotic more efficiency than sample median.