4.1 Joint and Marginal 4.2 conditional and independent

$$\begin{array}{c|ccccc} f_X(x) & F(x,y) & E[g(\vec{X})] \\ \sum_{y \in \mathbf{R}} f_{X,Y}(x,y) & P(X \leq x, Y \leq y) & \sum_{y \in \mathbf{R}} f(x,y) & \sum_{y \in \mathbf{R}} f(x,y) & \sum_{z \in \mathbf{R}} g(\vec{x}) & \sum_{z \in \mathbf{R}} g(x) & \sum_{z \in \mathbf{$$

4.2.1
$$f(x|y) = P(X = x|Y = y) = \frac{f(x,y)}{f_Y(y)}$$
 4.2.5 $f(x,y) = f_X(x)f_Y(y)$

$$4.1.10 \ Eg(X,Y) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x,y) f(x,y) dx dy \ 4.2.10 \ E(g(X)h(Y)) = E[g(x)]E[h(y)]$$

4.6.6 $X_1, ..., X_n$ indep, $E(g_1(X_1) \cdots g_n(X_n)) = (E(g_1(X_1)) \cdots (E(g_n(X_n)))$

Independence

Use 4.2.12 proof 4.3.2 $M_W(t) = M_X(t) M_Y(t) = e^{\mu_1(e^t - 1)} e^{\mu_2(e^t - 1)} = e^{(\mu_1 + \mu_2)(e^t - 1)}$

 $U \sim Geom(p = \frac{1}{2}), u = 1, 2... V \sim NBin(p = \frac{1}{2}, r = 2), v = 2, 3...$ (U,V) is

 $\{(u,v): u=1,2,...; v=u+1,u+2,..\}$ not independent.

4.2.14 — indep $X \sim n(\mu, \sigma^2), Y \sim n(\gamma, \tau^2)$ — $X + Y \sim n(\mu + \gamma, \sigma^2 + \tau^2)$

4.3.2 — indep
$$X \sim Poisson(\theta), Y \sim Poisson(\lambda)$$
 — $X + Y \sim Poisson(\theta + \lambda)$

4.3.3 — indep
$$X \sim Beta(\alpha, \beta), Y \sim Beta(\alpha + \beta, \gamma)$$
 — $XY \sim n(\alpha, \beta + \gamma)$

4.3.4/6—indep
$$X, Y \sim n(0,1)$$
 — $X + Y, X - Y \sim n(0,2), X/Y \sim cauchy(0,1)$

$$\begin{array}{l} 4.3.2 \; f_{Y_1,Y_2}(y_1,y_2) = \sum_{i=1}^k f_{X_1,X_2}(h_1(y_1,y_2),h_2(y_1,y_2))|J| \\ 4.3.5 \; f_{U,V}(u,v) = \sum_{i=1}^k f_{X,Y}(h_{1i}(u,v),\;h_{2i}(u,v))|J_i| \end{array}$$

4.3.5
$$f_{UV}(u,v) = \sum_{i=1}^{k} f_{VV}(h_{1i}(u,v), h_{2i}(u,v))|J_{i}|$$

$$J_{1,2} = \begin{vmatrix} \frac{\partial h_1}{\partial y_1} & \frac{\partial h_2}{\partial y_2} \\ \frac{\partial h_1}{\partial y_1} & \frac{\partial h_2}{\partial y_2} \end{vmatrix} = \frac{\partial h_1}{\partial y_1} \frac{\partial h_2}{\partial y_2} - \frac{\partial h_2}{\partial y_1} \frac{\partial h_1}{\partial y_2}$$

4.4 Mixture Distributions

$$4.2.3 E(g(x_2)|x_1) = \sum_{all \ x_2} g(x_2)p(x_2|x_1), E(g(x_2)|x_1) = \int_{all \ x_2} g(x_2)f(x_2|x_1)dx_2$$

Expectations

$$E(aX + b) = aE(X) + b - E(X) = \int_{-\infty}^{0} F_X(t)dt + \int_{0}^{\infty} F_X(t)dt$$

$$E[g(x)] = \mu = \sum_{x \in D} h(x)p(x) = \int_{-\infty}^{\infty} g(x)f(x)dx$$

$$E[(X - \mu)^n] = \mu_n = \sum (x - \mu)^n p(x) = \int (x - \mu)^n f(x) dx$$

$$V[X] = \sigma_x^2 = E(X^2) - [E(X)]^2 - V[X|Y] = E[(X - E[X|Y])^2|Y] - V[aX + b] = a^2\sigma^2$$

$$\sigma_{ax+b} = |a|\sigma_x$$

 $4.2.4 V[y|x] = E[Y^2|x] - (E[Y|x])^2$

4.4.3 EX = E[E(X|Y)]

4.4.7 V[X] = E[V(X|Y)] + V[E(X|Y)]

$$4.4.5~X|Y\sim Bin(Y,p), Y|\Lambda\sim Pois(\Lambda), \Lambda\sim Expo(\beta)~{\rm or}~X|Y\sim Bin(Y,p), Y\sim NBin(p=\frac{1}{1+\beta},r=1)$$

Pois-Gamma: $Y | \Lambda \sim Pois(\Lambda), \Lambda \sim Gamma(\alpha, \beta)$ then $Y \sim NBin(\alpha, \frac{1}{1+p\beta})$

4.4.6 Beta-Bin $X|P \sim Bin(n,P), P \sim Gamma(\alpha,\beta)$ $EX = E[E(X|P)] = E[nP] = n\frac{\alpha}{\alpha+\beta}$

4.4.8 Beta-Bin

$$V[X] = V[E(X|P)] + E[V(X|P)] = \frac{n^2 \alpha \beta}{(\alpha+\beta)^2(\alpha+\beta+1)} + \frac{n\alpha \beta}{(\alpha+\beta)(\alpha+\beta+1)} = n\frac{\alpha \beta(\alpha+\beta+n)}{(\alpha+\beta)^2(\alpha+\beta+1)}$$

4.5 Covariance and Correlation

$$4.5.1/3 \ Cov(X,Y) = E[(X-\mu_X)(Y-\mu_Y)] = E(XY) - E(X)E(Y) = E[XY] - \mu_X\mu_Y = \sigma_{XY}$$

$$4.5.2 \ \rho_{XY} = \frac{\sigma_{XY}}{\sigma_X\sigma_Y} \ Corr(X,Y) = \frac{Cov(X,Y)}{\sqrt{Var(X)Var(Y)}}$$

 $Cov(aX, bY) = abCov(X, Y) \ Cov(X, Y + Z) = Cov(X, Y) + Cov(X, Z) \ Cov(X, c) = 0$

 $4.5.6 Var(aX \pm bY) = a^{2}VarX + b^{2}VarY \pm 2abCov(X,Y)$

 $4.5.7 \ E[Y|X] = a + bx, \ E[Y] = E[E[Y|X]] = E[a + bX] = a + bE[X]$ by 4.4.3,

 $E[XE[Y|X]] = E[X(a+bX)] = aE[X] + bE[X^2], E[XE[Y|X]] = \int_{-\infty}^{\infty} xE[Y|x]f_X(x)dx$ by 2.2.1,

 $= \int_{-\infty}^{\infty} x \left[\int_{-\infty}^{\infty} y f_Y(y|x) dy \right] f_X(x) dx \text{ by } 4.2.3, = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} x y f(x,y) dy dx = E[XY] \text{ by } 4.1.10,$

 $\sigma_{XY} = E[XY] - \mu_X \mu_Y = a\mu_X + bE[X^2] - \mu_X \mu_Y = a\mu_X + b(\sigma_X^2 + \mu_X^2) - \mu_X(a + b\mu_X) = b\sigma_X^2$

 $\rho_{XY} = \frac{\sigma_{XY}}{\sigma_{X}\sigma_{Y}} = \frac{b\sigma_{X}^{2}}{\sigma_{X}\sigma_{Y}} = b\frac{\sigma_{X}}{\sigma_{Y}}$

4.5.10 bivarialte normal pdf with $\mu_X, \mu_Y, \sigma_X^2, \sigma_Y^2, \rho$

 $f_X(x) \sim n(\mu_X, \sigma_X^2) f_Y(y) \sim n(\mu_Y, \sigma_Y^2)$

 $f_{Y|X}(y|x) \sim n\left(\mu_Y + \rho \frac{\sigma_Y}{\sigma_X}(x - \mu_X), \sigma_Y^2(1 - \rho^2)\right)$

 $aX + bY \sim n(a\mu_X + b\mu_Y, a^2\mu_X^2 + b^2\mu_Y^2 + 2ab\rho\sigma_X\sigma_Y)$ 4.6.8 — indep $X_1, ..., X_n \sim gamma(\alpha, \beta)$ — $X_1 + ... + X_n \sim Gamma(\alpha_1 + ... + \alpha_n, \beta)$

5.2 Sum of r.s.

$$5.2.2 \ \bar{X} = \frac{X_1 + \dots + X_n}{n} = \frac{1}{n} \sum_{i=1}^n X_i - 5.2.3 \ S^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i^2 - n\bar{X}^2)$$

$$5.2.5 \ E\left(\sum_{i=1}^{n} g(X_i)\right) = nE(g(X_1)) \ Var\left(\sum_{i=1}^{n} g(X_i)\right) = nVar(g(X_1)) - 5.2.6 \ E\bar{X} = \mu,$$

$$Var\bar{X} = \frac{\sigma^2}{n}, ES^2 = \sigma^2$$

5.2.7
$$M_{\overline{X}}(t) = [M_X(\frac{t}{n})]^n$$
 — 5.2.9 Convolution formula $Z = X + Y$ $f_Z(z) = \int_{-\infty}^{\infty} f_X(w) f_Y(z-w) dw$

$$5.2.8 - X_1, ..., X_n \sim N(\mu, \sigma^2) - \bar{X} \sim N(\mu, \sigma^2/n)$$

indep —
$$X_1, ..., X_n \sim gamma(\alpha, \beta)$$
 — $\bar{X} \sim gamma(n\alpha, \beta/n)$

5.2.10—
$$X \sim cauchy(0, \sigma), Y \sim cauchy(0, \tau)$$
 — $X + Y \sim cauchy(0, \sigma + \tau)$

indep — $X_1,...,X_n \sim cauchy(0,\sigma)$ — $\bar{X} \sim cauchy(0,\sigma), \sum_{1}^{n} X \sim cauchy(0,n\sigma)$

5.3 Sampling from N

Properties of \bar{X} and S^2 5.3.1 when $X_1,...,X_n$ be iid $n(\mu,\sigma^2)$

$$\begin{array}{c|cccc} \text{W Normal} & \bar{X} \sim N(\mu, \frac{\sigma^2}{n}) & \bar{X}, S^2 \text{ indep} & \frac{(n-1)S^2}{\sigma^2} \sim \chi^2_{n-1} \\ \hline \text{W/o Normal} & E[\bar{X}] = \mu & V[\bar{X}] = \frac{\sigma^2}{n}) & E[S^2] = \sigma^2 \\ \end{array}$$

$$f_{\chi^2}(x) = \frac{x^{\frac{p}{2}-1}}{\Gamma \frac{p}{2} 2^{\frac{p}{2}}} e^{-\frac{x}{2}}, x > 0 - \chi_p^2 \sim gamma(\frac{p}{2}, 2) \text{ see } 4.6.8$$

$$5.3.2a - Z \sim n(0,1) - Z^2 \sim \chi_1^2$$

$$5.3.2b - X_1, ..., X_n \sim \chi_{n_1}^2 - X_1^2 + ... + X_n \sim \chi_{n_1 + n_2}^2$$

$$5.3.2a - Z \sim n(0,1) - Z^{2} \sim \chi_{1}^{2}$$

$$5.3.2b - X_{1}, ..., X_{n} \sim \chi_{p_{i}}^{2} - X_{1} + ... + X_{n} \sim \chi_{p_{1}+...p_{n}}^{2}$$

$$5.3.1 \text{ proof } \bar{X}_{n+1} = \frac{\sum_{i=1}^{n+1} X_{i}}{n+1} = \frac{\sum_{i=1}^{n} X_{i} + X_{n+1}}{n+1} = \frac{n\bar{X}_{n} + X_{n+1}}{n+1}$$

$$nS_{n+1}^{2} = (n-1)S_{n}^{2} + (\frac{n}{n+1})(X_{n+1} - \bar{X}_{n})^{2}$$

$$nS_{n+1}^2 = (n-1)S_n^2 + (\frac{n}{n+1})(X_{n+1} - \bar{X}_n)$$

$$Var\chi_{n-1}^2 = Var[\frac{(n-1)S^2}{\sigma^2}] = \frac{(n-1)^2}{\sigma^4} Var[S^2] = 2(n-1) \implies Var[S^2] = \frac{2\sigma^4}{n-1}$$

Connection between N, χ^2, t, F

$$5.3.4 X_1, ... X_n \sim n(\mu, \sigma^2) - \frac{\bar{X} - \mu}{\sigma/\sqrt{n}} \sim t_{n-1}$$

$$f_T(x) = \frac{\Gamma(\frac{p+1}{2})}{\Gamma(\frac{p}{2})\sqrt{p\pi}} \left(1 + \frac{x^2}{p}\right)^{-\frac{p+1}{2}}, -\infty < x < \infty - t_1 = Cauchy(0, 1)$$

$$5.3.5 \ X_1..X_n, Y_1..Y_m \ \text{indep.} \ X_i \sim n(\mu_X, \sigma_X^2) Y_j \sim n(\mu_Y, \sigma_Y^2), \frac{S_X^2}{\sigma_X^2} \sim \chi^2 - \frac{S_X^2/\sigma_X^2}{S_Y^2/\sigma_Y^2} \sim F$$

$$5.3.6 f_F(x) = \frac{\Gamma(\frac{p+q}{2})}{\Gamma(\frac{p}{2})\Gamma(\frac{q}{2})} {\binom{p}{q}}^{\frac{p}{2}} \frac{x^{\frac{p}{2}-1}}{(1+\frac{p}{q}x)^{\frac{p+q}{2}}}, x > 0$$

a.
$$X \sim F_(p,q) \ \frac{1}{X} \sim F_(q,p)$$
 — b. $X \sim T_(q) \ X^2 \sim F_(1,q)$ — c. $X \sim F_(p,q) \ \frac{\frac{p}{q} \ X}{1+\frac{p}{q} \ X} \sim Beta(\frac{p}{2},\frac{q}{2})$

5.4 Order statistics

5.4.4
$$f_K(x) = \frac{n!}{(i-1)!(n-j)!} K[F_X(x)]^{k-1} [1 - F_X(x)]^{n-k} f(x)$$
 1-29p9

$$5.4.6 \ f_{X_{(1)},...,X_{(n)}}(x_1,...,x_n) = \begin{cases} n! f_X(x_1) \cdot ... \cdot f_X(x_n) & -\infty < x_1 < .. < x_n < \infty \\ 0 & \text{otherwise} \end{cases}$$

$$f_{X_{(i)},X_{(j)}}(u,v) = \frac{n!}{(i-1)!(j-1-i)!(n-j)!} f_X(u) f_X(v) [F_X(u)]^{i-1} [F_X(v) - F_X(u)]^{j-1-i} [1 - F_X(v)]^{n-j}$$

1 from N to T to Chi to F

Given some function of these, find the distribution.

2 Transformation of pairs of r.v.s

Given 2 random variables and their joint function, and given a function of them, find its distribution. 1-10p1

$$\begin{array}{c|c} f(x,y) & 0 < x < \infty, 0 < y < \infty \\ \frac{1}{4}e^{-\frac{x+y}{2}} & u = \frac{X-Y}{2} \\ 1. \ V = Y \rightarrow X = 2u+v, Y = v \end{array}$$

1.
$$V = Y \rightarrow X = 2u + v, Y = v$$

2.
$$J = \begin{vmatrix} \frac{\partial x}{\partial u} & \frac{\partial x}{\partial v} \\ \frac{\partial y}{\partial u} & \frac{\partial y}{\partial v} \end{vmatrix} = \begin{vmatrix} 2 & 1 \\ 0 & 1 \end{vmatrix} = 2$$

3.
$$g(u,v) = f(x,y)|J| = \frac{1}{7}e^{-\frac{x+y}{2}} = \frac{1}{7}e^{-\frac{2u+v+v}{2}} = \frac{1}{7}e^{-(u+v)}$$

3.
$$g(u,v) = f(x,y)|J| = \frac{1}{4}e^{-\frac{x+y}{2}}2 = \frac{1}{2}e^{-\frac{2u+v+v}{2}} = \frac{1}{2}e^{-(u+v)}$$

4. $0 < x < \infty, 0 < y < \infty \implies 0 < 2u + v < \infty, 0 < v < \infty \implies v > -2u$

$$\begin{array}{ll} g_U(u) = \left| \begin{array}{l} \int_{-2u}^{\infty} \frac{1}{2} e^{-(u+v)} dv = \frac{1}{2} e^{-u} \int_{-2u}^{\infty} e^{-v} dv = \frac{1}{2} e^{-u} \left[-e^{-v} \right]_{-2u}^{\infty} = \frac{1}{2} e^{-u} \left[0 + e^{2u} \right] \end{array} \right| \quad u < \\ = \frac{1}{2} e^{|u|} \left| \begin{array}{l} \int_{0}^{\infty} \frac{1}{2} e^{-(u+v)} dv = \frac{1}{2} e^{-u} \int_{0}^{\infty} e^{-v} dv = \frac{1}{2} e^{-u} \left[-e^{-v} \right]_{0}^{\infty} = \frac{1}{2} e^{-u} \left[0 + 1 \right] \end{array} \right| \quad u < \\ u > 0 \end{array}$$

Distribution: Double Exponential(Laplace)

3

Given
$$f(x, y)$$
, find the $f_X(x) = -E[X] = -V[X] = f_Y(y) = -E[Y] = -V[Y] = f(x, y) = -E[XY] = Cov(x, y) = E[XY] - EXEY = f(X|Y) = -E[X|Y] = -V[X|Y] = \rho = Cov(x, y) / \sqrt{V[X]V[Y]} = 0$

4 Order statistics

* Find the distribution of X(k) or find the joint distribution of $X_{(i)}, X_{(k)}$ joint pmf 1-31p2-3

5.4.5 uniform order pdf 1-31p4-8

pdf/pmf

CDF:
$$F(x) = P(X \le x) = \sum_{y:y \le x} p(y) = \int_{-\infty}^{x} f(y) dy$$
 probabilities: $a \le b$:

PMF: $p(x) = P(X = x) = P(\forall w \in \mathcal{W} : X(w) = x)$

$$P(a \le X \le b) = F(b) - F(a^-); \ P(a < X \le b) = F(b) - F(a); \ P(a \le X \le a) = p(a);$$

$$P(a < X < b) = F(b^-) - F(a); \ \text{(where } a^- \text{ is the largest possible X value strictly less than } a); \ \text{Taking } a = b \ \text{yields } P(X = a) = F(a) - F(a - 1) \ \text{as desired.}$$

$$PDF: \ P(\forall w \in \mathcal{W} : a \le X(w) \le b) = \int_a^b f(x) dx$$

$$P(a \le X \le b) = P(a < X \le b) = P(a < X \le b) = P(a < X \le b) = F(b) - F(a)$$

CDF

Condition: $f(x) \ge 0 \forall x \text{ pmf } \sum_{x} f_X(x) = 1, pdf \int_{-\infty}^{\infty} f(x) dx = 1$

mgf

$$\begin{split} E(e^{tx}) &= \int e^{tx} f(x) dx = \sum e^{tx} f(x); \ M_{aX+b}(t) = e^{tb} M_{aX}(t) \\ M_X(t) &= E(e^{tx}) & M_X(0) = 1 \\ M_X'(t) &= E(xe^{tx}) & M_X'(0) = E(X) \\ M_X''(t) &= E(x^2 e^{tx}) & M_X''(0) = E(X^2) \\ M_X^n(t) &= E(x^n e^{tx}) & M_X^n(0) = E(X^n) \end{split}$$

transform

$$\begin{array}{l} g(x) \uparrow & | F_Y(y) = F_X(g^{-1}(y)) \\ g(x) \downarrow & | F_Y(y) = 1 - F_X(g^{-1}(y)) \\ \text{not monotone} \because X \leq 0 \ is \ \emptyset \therefore P(X \leq -\sqrt{y}) = 0 \\ F_Y(y) = P(Y \leq y) = P(X^2 \leq y) = P(-\sqrt{y} \leq X \leq \sqrt{y}) = P(X \leq \sqrt{y}) - P(X \leq -\sqrt{y}) = F_X(\sqrt{y}) = \sqrt{y}, \ 0 < \sqrt{y} < 1 \\ \text{monotone:} f_Y(y) = f_X(g^{-1}(y)) |\frac{d(g^{-1}(y))}{dy}| \end{array}$$

Series

Integrals+c

Integreation by parts
$$\begin{bmatrix} u & du & dv & v \\ x & dx & e^{-x} & -e^{-x} \end{bmatrix} \int_a^b u dv \int_a^b x e^{-x} dx$$

$$\begin{aligned} &=uv|_a^b - \int_a^b v du \\ &= -xe^{-x} + \int xe^{-x} dx = -xe^{-x} - e^{-x} + c \\ &= xlnx|_3^5 - \int_3^5 dx = (xlnx - x)|_3^5 = 5ln5 - 3ln3 - 2 \\ &\int_a^b f(x) dx = F(b) - F(a) = - \int_b^a f(x) dx = \int_a^c f(x) dx + \int_c^b f(x) dx; \\ &\frac{d}{dx} \int_{v(x)}^{u(x)} f(t) dt = u'(x) f[u(x)] - v'(x) f[v(x)] \end{aligned}$$

Derivatives

Derivatives
$$(cf)' = cf'(x) \qquad (fg)' = f'g + fg' \qquad (f \pm g)' = f'(x) \pm g'(x) \\ \frac{dx}{dx} = 1 \qquad (\frac{f}{g})' = \frac{(f'g - fg')}{g^2} \qquad (f(g(x)))' = f'(g(x))g'(x) \\ \frac{de^x}{dx} = e^x \qquad \frac{d(x^n)}{dx} = nx^{n-1} \qquad \frac{d \ln(x)}{dx} = \frac{1}{x}, x > 0 \\ \frac{d \cos x}{dx} = -\sin x \qquad \frac{d \tan x}{dx} = \sec^2 x \qquad \frac{d \ln(x - 1)}{dx} = \ln x$$

Distribution	CDF	P(X=x),f(x)	μ	EX^2	Var	MGF	M'(t)	M" (t)	$M^n(t)$
Bern(p)		$p^x q^{1-x}, x \in \{1, 0\}$	p	p	pq	$pe^t + q$			
$\operatorname{Bin}(n,p)$	$I_{1-p}(n-x,x+1)$	$\binom{n}{x} p^x q^{n-x}; x \in \{0, 1n\}$	np	$\mu(\mu+q)$	$q\mu$	$(pe^t + q)^n$			
Geom(p)	$1 - q^x$	$pq^{x-1}, x \in 1, 2, \dots$	$\frac{1}{p}$	$\frac{p+2q}{p^2}$	$\frac{q}{p^2}$	$\frac{pe^t}{1 - qe^t}, t < -\ln q$			
	$1 - q^{x+1}$	$pq^x, x \in 0, 1, \dots$	$\frac{q}{p}$	$\frac{p+2q}{p^2}$ $\frac{q^2+q}{p^2}$	$\frac{q}{p^2}$	$\frac{p}{1-qe^t}, qe^t < 1$	$\frac{pqe^t}{(1-qe^t)^2}$	$\frac{2pqe^t}{(1-qe^t)^3} - M'(t)$	
$\overline{\mathrm{NBin}(r,p)}$		$\binom{x-1}{r-1} p^r q^{x-r}, x \in r, r+1$	$\frac{r}{p}$	•		$\left(\frac{pe^t}{1-qe^t}\right)^r$			
, ,		$\binom{r-1}{r-1} p^r q^x, x \in 0, 1$	$\frac{rq}{p}$		$\frac{rq}{p^2}$ $\frac{rq}{p^2}$	$\left(\frac{p}{1-qe^t}\right)^r, qe^t < 1$			
$\overline{\text{HGeom}(N,m,k)}$		$\frac{\binom{m}{x}\binom{N-m}{k-x}}{\binom{N}{k}}$	$\frac{km}{N}$		$\mu \frac{(N-m)(N-k)}{N(N-1)}$	- 4-			
$\mathrm{HGeom}(w,b,k)$		$\frac{\binom{w}{x}\binom{k}{b}}{\binom{w+b}{k}}$	$\frac{kw}{w+b}$		$\mu \frac{b(w+b-k)}{(w+b)(w+b-1)}$				
$Pois(\mu)$	$e^{-\mu} \sum_{i=0}^{x} \frac{\mu^i}{i!}$	$\frac{\mu^x}{x!}e^{-\mu}, x \in 0, 1$	μ	$\mu^2 + \mu$	μ	$e^{\mu(e^t-1)}$	$\mu e^t M(t)$	$\mu e^t (1 + \mu e^t) M(t)$	
$\operatorname{Unif}(n)$		$\frac{1}{n}, x \in 1, 2n$	$\frac{n+1}{2}$	$\frac{(n+1)(2n+1)}{6}$	$\frac{(n^2-1)}{12}$	$\frac{\sum_{i=1}^{n} e^{ti}}{n}$			
$\overline{\mathrm{Unif}(a,b)}$	$\frac{x-a}{b-a}$	$\frac{1}{b-a}, x \in (a,b)$	$\frac{a+b}{2}$		$\frac{(b-a)^2}{12}$	$\frac{e^{tb} - e^{ta}}{t(b-a)}$			
$\mathcal{N}(\mu, \sigma^2)$		$\frac{1}{\sigma\sqrt{2\pi}}e^{-\frac{(x-\mu)^2}{2\sigma^2}}$	μ	$\mu^2 + \sigma^2$	σ^2	$e^{\mu t + \frac{\sigma^2 t^2}{2}}$	$(\mu + \sigma^2 t)M(t)$	$[(\mu + \sigma^2 t)^2 + \sigma^2]M(t)$	
$\mathcal{N}(0,1)$		$\frac{1}{2}e^{-\frac{x}{2}}$	0	1	1	$e^{\frac{t^2}{2}}$			
$\mathcal{LN}(\mu, \sigma^2)$		$\frac{1}{x\sigma\sqrt{2\pi}}e^{\frac{-(\ln x-\mu)^2}{2\sigma^2}}$	$e^{\mu + \frac{\sigma^2}{2}}$	$e^{2\mu+2\sigma^2}$	$\theta^2(e^{\sigma^2}-1)$	×			
$Cauchy(\theta, \sigma^2)$		$\frac{\pi\sigma}{1+(\frac{x-\theta}{2})^2}$	×	×	×				
D Expo (μ, σ^2)		$\frac{1}{2\pi\sigma}e^{-\left \frac{\omega-\mu}{\sigma}\right }$	μ	$\mu^2 + 2\sigma^2$	$2\sigma^2$	$\frac{e^{\mu t}}{1-\sigma^2 t^2}$			
$-$ Expo (λ)	$1 - e^{-\lambda x}$	$\lambda e^{-\lambda x}, x \in (0, \infty)$	$\frac{1}{\lambda}$		$\frac{1}{\lambda^2}$	$\frac{\frac{e^{t}}{1-\sigma^{2}t^{2}}}{\frac{\lambda}{\lambda-t}, t < \lambda}$			
$\text{Expo}(\beta)$		$\frac{1}{\beta}e^{-\frac{x}{\beta}}$	β		$\hat{\beta}^2$	$\frac{1}{1-\beta t}$	$\beta(1-\beta t)^{-2}$	$2\beta^2(1-\beta t)^{-3}$	
$Gamma(a, \lambda)$		$\frac{\lambda^a}{\Gamma(a)} x^{a-1} e^{-\lambda x}, x \in (0, \infty)$	$\frac{a}{\lambda}$	$\frac{a}{\lambda^2}$	$\left(\frac{\lambda}{\lambda - t}\right)^a, t < \lambda$				
$\operatorname{Gamma}(\alpha,\beta)$		$\frac{1}{\Gamma(a)\beta\alpha}x^{a-1}e^{-x/\beta}$	lphaeta	$lphaeta^2$	$\left(\frac{1}{1-\beta t}\right)^a, t < \frac{1}{\beta}$				
Beta (a,b)		$\frac{\Gamma(a+b)}{\Gamma(a)\Gamma(b)}x^{a-1}(1-x)^{b-1}$	$\frac{a}{a+b}$		$\frac{\mu(1-\mu)}{(a+b+1)}$				$\frac{\Gamma(\alpha+n)\Gamma(\alpha+\beta)}{\Gamma(\alpha+\beta+n)\Gamma(\alpha)}$
$B(\alpha, \beta) =$		$\frac{\Gamma(a)b}{\Gamma(a)\Gamma(b)}x^{a-1}(1-x)^{b-1}$ $\frac{\Gamma(a)\Gamma(b)}{\Gamma(a+\beta)}, x \in (0,1)$	$a_{\mp b}$	$\frac{a(a+1)}{(a+b)(a+b+1)}$	$\frac{ab}{(a+b)^2(a+b+1)}$				$\Gamma(\alpha+\beta+n)\Gamma(\alpha)$
χ_p^2		$\frac{1}{2^{p/2}\Gamma(n/2)}x^{p/2-1}e^{-x/2}$	p	$\frac{(a+b)(a+b+1)}{2p+p^2}$	2p	$(1-2t)^{-p/2}, t < \frac{1}{2}$			
$\overline{t_n}$		$\frac{\Gamma(\frac{2}{2})}{\sqrt{n\pi}\Gamma(\frac{n}{2})}(1+\frac{x^2}{n})^{-\frac{N+2}{2}}$	0, n > 1		$\frac{n}{n-2}, n > 2$	×			
F	x > 0	$\frac{\Gamma(\frac{p+q}{2})}{\Gamma(\frac{p}{2})\Gamma(\frac{q}{2})} \left(\frac{p}{q}\right)^{\frac{p}{2}} \frac{x^{\frac{p}{2}-1}}{\left(1+\frac{p}{q}x\right)^{\frac{p+q}{2}}}$	$\frac{q}{q-2}$	q > 2	$2(\frac{q}{q-2})^2 \frac{p+q-2}{p(q-4)}$	q > 4			
Arcsine	$\frac{1}{\pi \arcsin \sqrt{x}}$	$\frac{1}{\pi\sqrt{x(1-x)}}, x \in [0,1]$	$\frac{1}{2}$		$\frac{1}{8}$				$\operatorname{Beta}(\frac{1}{2},\frac{1}{2})$
Dirichlet	$B(a) = \frac{\prod_{i=1}^{k} \Gamma(a_i)}{\Gamma(\sum_{i=1}^{k} a_i)}$	$\frac{1}{B(a)} \prod_{i=1}^{k} x_i^{a_i - 1}, x \in (0, 1)$	$\frac{a_i}{\sum_k a_k}$	$\sum_{i=1}^{k} x_i = 1$	$\frac{a_i(a_0\!-\!a_i)}{a_0^2(a_0\!+\!1)}$	$Cov(X_i, X_j) =$	$^{\frac{-a_{i}a_{j}}{a_{0}^{2}(a_{0}+1)}}$	$a_0 = \sum_{i=1}^k a_i$	

 $U \sim Geom(p=\frac{1}{2}), u=1,2..$ the number of trials needed to get the first head. $V \sim NBin(p=\frac{1}{2},r=2), v=2,3..$ the number of trials needed to get two heads in repeated tosses of a fair coin.