

$y_i = \beta_0 + \beta_1 x_i + \varepsilon_i, i_{1:n} \text{SLR}; \mathbf{Y}_{n \times 1} = \mathbf{X}_{n \times (k+1)} \boldsymbol{\beta}_{(k+1) \times 1} + \boldsymbol{\varepsilon}_{n \times 1} \text{MLR}; p=k+1$

Assumptions: constant variance, zero mean, independent; $\varepsilon_{ijkl} \stackrel{iid}{\sim} N(0, \sigma^2 I)$
 $Y'Y = \sum y_i^2 = S_{yy} + n\bar{y}^2$ $(Y - \bar{y})'(Y - \bar{y}) = \sum (y_i - \bar{y})^2 = \sum y_i^2 - n\bar{y}^2 = (n-1)S_y^2 = S_{yy}$
 $X'Y = [\sum y_i, \sum x_{1i}y_i, \dots, \sum x_{ki}y_i]$ $(X - \bar{x})'(Y - \bar{y}) = \sum (x_i - \bar{x})(y_i - \bar{y}) = \sum xy - n\bar{x}\bar{y} = \bar{S}_{xy}$
 $X'X = \begin{bmatrix} n & \sum x_{ki} \\ \sum x_{ki} & \sum x_{ki}^2 \end{bmatrix}_{p \times p}$ $(X - \bar{x})(X - \bar{x})' : [\sum (x_i - \bar{x})^2 = \sum x_i^2 - n\bar{x}^2 = (n-1)S_x^2 = S_{xx}]$

$(X'X)^{-1} = \frac{1}{nS_{xx}} \begin{bmatrix} \sum x_{ki}^2 & -\sum x_{ki} \\ -\sum x_{ki} & n \end{bmatrix}; (X'X)^{-1}X'y = \begin{bmatrix} \bar{y} - \bar{x} \frac{S_{xy}}{S_{xx}} \\ \frac{S_{xy}}{S_{xx}} \end{bmatrix} = \begin{bmatrix} \bar{y} - \hat{\beta}_1 \bar{x} \\ \hat{\beta}_1 \end{bmatrix} = \begin{bmatrix} \hat{\beta}_0 \\ \hat{\beta}_1 \end{bmatrix}$
 $\hat{V}[\hat{\beta}] = \sigma^2 (X'X)^{-1} = E[(\hat{\beta} - E\hat{\beta})(\hat{\beta} - E\hat{\beta})'] = \begin{bmatrix} \hat{V}[\hat{\beta}_1] & \text{Cov}(\hat{\beta}_0, \hat{\beta}_1) \\ \dots & \hat{V}[\hat{\beta}_1] \end{bmatrix} = \begin{bmatrix} \frac{\sigma^2 \sum x_i^2}{nS_{xx}} & -\frac{\sigma^2 \bar{x}}{S_{xx}} \\ -\frac{\sigma^2 \bar{x}}{S_{xx}} & \frac{\sigma^2}{S_{xx}} \end{bmatrix}$
 $\hat{V}[\hat{\beta}_0] = \hat{\sigma}^2 (\frac{1}{n} + \frac{\bar{x}^2}{S_{xx}}); \hat{V}[\hat{\beta}_1] = \frac{\hat{\sigma}^2}{S_{xx}}; \text{Cov}(\hat{\beta}_0, \hat{\beta}_1) = -\frac{\hat{\sigma}^2 \bar{x}}{S_{xx}}; \text{Cov}(\bar{y}, \hat{\beta}_1) = 0$

$\text{Cov}[\bar{a}'\hat{\beta}] = \sigma^2 a' (X'X)^{-1} a$; $\text{se}(\bar{a}'\hat{\beta} | \{\hat{\beta}_i\} | \{\hat{\beta}_i - \hat{\beta}_i\}) = \hat{\sigma} \sqrt{a' (X'X)^{-1} a} \{C_{ii} | \{C_{ii} + C_{jj} - 2C_{ij}\}\}$
 $\hat{Y} = \hat{\beta}_0 + \hat{\beta}_1 x_{1i} + \dots + \hat{\beta}_k x_{ki} = \hat{X}\hat{\beta} = (X'X)^{-1}X'Y = HY; \hat{y}_0 = \hat{\beta}_0 + \hat{\beta}_1 x_0; \hat{y}_0 = X_0\hat{\beta}$
 $Y'Y = \hat{\beta}_0 n\bar{y} + \hat{\beta}_1 \sum x_{1i}y_i + \dots + \hat{\beta}_k \sum x_{ki}y_i = \hat{\beta}'X'X\hat{\beta} = \hat{\beta}'X'Y = Y'X(X'X)^{-1}X'Y = Y'HY$

kdf	SS				E[SS]; $\hat{\sigma}^2 = \text{MSE}$
R1[k]	$\sum (\hat{y} - \bar{y})^2$	$Y'Y - n\bar{y}^2$	$\hat{\beta}'X'Y - n\bar{y}^2$		
E n-2(p)	$\sum (y_i - \hat{y})^2$	$Y'Y - Y'Y\hat{Y} - Y'\hat{Y} - \hat{\beta}'X'Y$	$Y(Y - X\hat{\beta})'(Y - X\hat{\beta})$	$\hat{\beta}_1 S_{xy} = \beta_1^2 S_{xx}$	$k\sigma^2 + \beta_{1:k}'X'X\beta_{1:k}$
T n-1	$\sum (y_i - \bar{y})^2$	$Y'Y - n\bar{y}^2$	$\sum y_i^2 - n\bar{y}^2$	$(Y - \bar{y})'(Y - \bar{y})$	$S_{yy} - \hat{\beta}_1 S_{xy} \quad (n-p)\sigma^2$
				$(n-1)S_y^2$	$(n-1)\sigma^2 + \hat{\beta}_1 S_{xx}\hat{\beta}_1$

Hypothesis $\text{Cov}[a(A)Y] = a(A)\text{Cov}[Y]a(A)'$

$H_0 : \hat{\beta}_i = c\{\hat{\beta}_i\} | \{\bar{a}'\hat{\beta}\}; t_0 = \frac{\hat{\beta}_i - c\{\hat{\beta}_i - \hat{\beta}_i\} | \{\bar{a}'\hat{\beta}\}}{\text{se}(\hat{\beta}_i | \{\hat{\beta}_i - \hat{\beta}_i\} | \{\bar{a}'\hat{\beta}\})} < t_{\frac{\alpha}{2}, n-2(p)}; t \text{ test } \hat{\beta}_i = \text{partial F test on } a_i$
 $H_0: \hat{\beta}_{1:k} = 0; \frac{SSR}{\sigma^2} \sim \chi^2_{(k)}; \frac{SSE}{\sigma^2} = \frac{Y'(I-H)Y}{\sigma^2} \sim \chi^2_{n-2(p)}; \frac{MSR}{MSE} \sim F_{\alpha, k, n-p}$
 $H_0 : T\hat{\beta}_{ij} = c; \beta_1 = 2\beta_3, \beta_2 = \beta_3, \beta_5 = 0; df F_{Red} - df F_{Full} = n - (p-r) - (n-p) = r - j - i + 1$

$T = \begin{bmatrix} 0 & 0 & 0 & -2 \\ 0 & 0 & 0 & -1 \\ 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 \end{bmatrix}_{3 \times 7} \begin{bmatrix} \hat{\beta}_0 \\ \hat{\beta}_1 \\ \hat{\beta}_2 \\ \hat{\beta}_3 \end{bmatrix}_{7 \times 1} = \begin{bmatrix} 0 \\ 0 \\ 0 \\ 0 \end{bmatrix}_{3 \times 1} \text{rank}(T) = 3 \text{ independent};$
 $F_{\alpha, r, df F_{Full}} = \frac{(SSE_{Red} - SSE_{Full})/r}{SSE_{Full}/df F_{Full}} = \frac{(SSR_{Full} - SSR_{Red})}{r \cdot MSE_{Full}} = \frac{(T\hat{\beta} - c)'T'(X'X)^{-1}T'(T\hat{\beta} - c)}{r \cdot MSE_{Full}}$

$CI\hat{\beta}_i | \{\hat{\beta}_i - \hat{\beta}_i\} \pm \Delta_{n-2(p)} \text{se}(\hat{\beta}_i | \{\hat{\beta}_i - \hat{\beta}_i\}); \Delta_{Bonf} = t_{\frac{\alpha}{2m}} m: \# \text{ of hypo}; \Delta_{Shef} = \sqrt{2\{p\}} F_{\alpha, p, n-p}$
 $\sigma^2: (\frac{SSE}{\lambda_{2, dfE}} / \frac{(n-2)\hat{\sigma}^2}{\lambda_{1, n/2, dfE}}) \text{ Joint: } H_0: \hat{\beta} = \beta, \frac{(\hat{\beta} - \beta)'X'X(\hat{\beta} - \beta)}{\sigma^2(X'X)^{-1}} \sim \chi^2_p; P(\frac{(\hat{\beta} - \beta)'X'X(\hat{\beta} - \beta)}{p \cdot MSE} < F_{\alpha, p, n-p}) = 1 - \alpha$

Elliptically reg: $n(\hat{\beta}_0 - \beta_0)^2 + 2\sum x(\hat{\beta}_0 - \beta_0)(\hat{\beta}_1 - \beta_1) + \sum x^2(\hat{\beta}_1 - \beta_1)^2 \leq \text{cpMSE}$
 $\hat{y}_0 \pm t_{\frac{\alpha}{2}, n-2} \hat{\sigma} \sqrt{\frac{1}{k} + (\frac{1}{n} + \frac{(x_0 - \bar{x})^2}{S_{xx}})} \{x'_0 (X'X)^{-1} x_0\} \# \text{ new valu(PI)} k=1; \infty(CI); \hat{\beta}_0 \text{ as } x_0 = 0$

$R^2 = \frac{SSR}{SST} = \frac{S_{xy}^2}{S_{xx}S_{yy}}; r = \frac{S_{xy}}{(n-1)S_xS_y} = \frac{\text{Cov}(x,y)}{\sqrt{V(X)Y}} = \frac{\text{Cov}(\hat{\beta}_1, \hat{\beta}_0)}{\text{se}\hat{\beta}_1 \text{se}\hat{\beta}_0}; R^2_{adj} = 1 - \frac{SSE/(n-p)}{SST/(n-1)}; R^2_{pre} = 1 - \frac{PRESS}{SST}$
coefficient of determination is the proportion of variation explained by regressor x
Standardized $W = b_0 + b_1 Z = \frac{y_i - \bar{y}}{s_y} = r \frac{x_i - \bar{x}}{s_x}; b_0 = 0; b_1 = \frac{s_{xy}}{s_{xx}} = r$

Multicollinearity $XZ=0; VIF = \frac{1}{1-R^2};$ for the Reg of X_i on all of the other predictors .

Calibration given a $y^*, \hat{x}^* = \frac{y^* - \hat{\beta}_0}{\hat{\beta}_1} V[\hat{x}^*] = \frac{\hat{\sigma}^2}{\hat{\beta}_1^2} [1 + \frac{1}{n} + \frac{(\hat{x}^* - \bar{x})^2}{S_{xx}}]$

least-squares estimators $\frac{\partial \bar{a}\bar{x}}{\partial \bar{x}} = \bar{a}; \frac{\partial \bar{x}'\bar{A}\bar{x}}{\partial \bar{x}} = 2\bar{A}\bar{x}$

$SSE = \sum^n (y_i - \beta_0 - \beta_1 x_i - \beta_2 x_i^2)^2 = e'e = (Y - \hat{Y})'(Y - \hat{Y}) = (Y - X\hat{\beta})'(Y - X\hat{\beta});$
 $\frac{\partial SSE}{\partial \beta_{0,1,2}} = 2\sum^n (y_i - \hat{\beta}_0 - \hat{\beta}_1 x_i - \hat{\beta}_2 x_i^2)(-1, -x_i, -x_i^2) = 0 = -2X'Y + 2X'X\hat{\beta}$
 $\beta_0 = \beta_1 h(y_i) = \beta_1 f(1 + x_i) + \beta_2 g(x_i^2); \beta_0 = 0, h(y_i) = \beta_1 f(x_i) + \beta_2 g(x_i^2),$
 $\hat{\beta}_1 = \frac{\sum g(x^2)^2 \sum f(x)h(y) - \sum g(x^2)f(x)\sum g(x^2)h(y)}{\sum g(x^2)^2 \sum f(x)^2 - [\sum g(x^2)f(x)]^2}, \hat{\beta}_2 = \frac{\sum f(x)^2 \sum g(x^2)h(y) - \sum g(x^2)f(x)\sum f(x)h(y)}{\sum g(x^2)^2 \sum f(x)^2 - [\sum g(x^2)f(x)]^2}$
 $d\beta_0 = \beta_2, \beta_1 = c, h(y_i - cx_i) = \beta_0 g(1 + dx_i^2), \hat{\beta}_0 = \frac{\sum g(1+dx^2)h(y-cx)}{\sum g(1+dx^2)^2}$

Generalized least-squares solution $V = PAP'; V^{-1} = L'L; L = \Lambda^{-\frac{1}{2}}P'$

$V[e] = \sigma^2 V; \text{Cov}[\gamma] = L\text{Cov}[e]L' = \Lambda^{-\frac{1}{2}}P'\sigma^2 V(\Lambda^{-\frac{1}{2}}P')^{-1} = \sigma^2 I; W(LY) = Z(LX)\beta + [Le]\gamma$
 $\hat{\beta} = (Z'Z)^{-1}Z'W = (X'V^{-1}X)^{-1}X'V^{-1}Y; \text{Cov}(\hat{\beta}) = \sigma^2 (X'V^{-1}X)^{-1}$
 $H^2 = X(X'X)^{-1}X'X(X'X)^{-1}X' = X(X'X)^{-1}X' = H^2 = H, \text{idempotent matrix, symmetric}$
Residual: $e = (I-H)Y; \text{Cov}(e) = \sigma^2(I-H); \text{Var}(e_i) = \sigma^2(1-h_{ii}); \text{Cov}(e_i, e_j) = -\sigma^2 h_{ij}$
Standardized $d_i = \frac{e_i}{\sqrt{MSE}}; PRESS(e_i) = y_i - \hat{y}_i = \frac{e_i}{1-h_{ii}}; \text{Studentized } r_i = \frac{e_i}{\sqrt{MSE(1-h_{ii})}}$

$h_{ii} = h'_i h_i = \sum h_{ii}^2 = h_{ii}^2 + \sum_{j \neq i} h_{ij}^2; h_{ii} - h_{ii}^2 = h_{ii}(1 - h_{ii}) = \sum_{j \neq i} h_{ij}^2 \geq 0; 0 \leq h_{ij} \leq 1$
Leverage of the i^{th} data point h_{ii} : how much y_i contributes to \hat{y}_i ; $\sum^n h_{ii} = p$
 $\hat{y}_i = (i^{th} \text{ row of } H)y = \sum_{j=1}^n h_{ij}y_j = h_{ii}y_i + \sum_{j \neq i} h_{ij}y_j$ examine the point with $> \frac{2p}{n}$

Omit the i^{th} point: Cook's $D_i = \frac{(\hat{\beta}_{(i)} - \hat{\beta})'X'X(\hat{\beta}_{(i)} - \hat{\beta})}{p \cdot MSE} \cdot \frac{e_i^2 h_{ii}}{(1-h_{ii})^2 p \cdot MSE} \cdot \frac{r_i^2 h_{ii}}{(1-h_{ii})^p}$ examine if > 1

DIFFITS $_i = \frac{e_i \sqrt{h_{ii}}}{(1-h_{ii}) \sqrt{MSE(i)}}$; examie if $> 2\sqrt{p/n}$ $X = [X_1 | X_2]; \hat{\beta} = \frac{(X_1'X_1)^{-1}X_1'Y}{(X_1'X_2)^{-1}X_2'Y}$

Lack of fit y_{ij} denote the j^{th} observation on the response at $x_{i,lm}$ m distinct x values
 $\sum_i \sum_j n_{ij} (y_{ij} - \hat{y}_i)^2 (SSE) = \sum_i \sum_j n_{ij}^2 (y_{ij} - \bar{y}_i)^2 (SSPE) + \sum_m n_i (\bar{y}_i - \hat{y}_i)^2 (SSLoF)$

n-p(dfE)=n-m(dfPE)+m-p(dfLoF); $F = \frac{SSLoF/dfLoF}{MSPE/dfPE}; H_0$: no LOF, model is appropriate
Indicator $y_i = \beta_0 + \beta_1 x_i + w_i(\beta_2 + \beta_3 x_i) + \varepsilon_i, w_i = m+1, n=1, 0, \text{o.w.}$
 $y_i = \beta_0 + \beta_1 x_i + w_{1i}(\gamma_0 + \gamma_1 x_i) + w_{2i}(\delta_0 + \delta_1 x_i) + \varepsilon_i, w_{1i} = m+1, k=1, w_{2i} = k+1, n=1, 0, \text{o.w.}$
 $y_i = \beta_0 + \beta_1 x_{1i} + \beta_2 x_{2i} + \beta_3 x_{3i} + \varepsilon_i, x_{1i} = \frac{k}{k+1}, k=1, x_{2i} = 1, 4, 7, \dots, -1, x_{3i} = 3, 6, 9, \dots, 1, 0, \text{o.w.}$

$Y_{1:n} = \begin{bmatrix} 1 & x_1 & 0 & 0 & 0 & 0 \\ 1 & x_m & 0 & 0 & 0 & 0 \\ 1 & x_{m+1} & 1 & x_{m+1} & 0 & 0 \\ 1 & x_k & 1 & x_k & 0 & 0 \\ 1 & x_{k+1} & 0 & 0 & 1 & x_{k+1} \\ 1 & x_n & 0 & 0 & 1 & x_n \end{bmatrix} \begin{bmatrix} \hat{\beta}_0 \\ \hat{\beta}_1 \\ \gamma_0 \\ \gamma_1 \\ \delta_0 \\ \delta_1 \end{bmatrix} + \epsilon'_{1:n}; \begin{bmatrix} 1 & 0 & 0 & -1 & x_{31} \\ 1 & 0 & 0 & 0 & \vdots \\ 1 & 1 & 1 & -1 & \vdots \\ 1 & 1 & 1 & 0 & \vdots \\ 1 & 1 & 1 & 1 & x_{3n} \end{bmatrix} \beta'_{1:4}$

μ overall mean; $y_{ijkl}(\varepsilon_{ijkl})$ is response(random error) (for the k^{th} replicate EU) when i^{th} (Latin(Greek) letter) treatment(level of Factor X) is applied (at j^{th} block) (in the k^{th} row and k^{th} column);
 τ_i is fixed(random) (main) effect of i^{th} (Latin(Greek) letter) treatment (block; level of Factor X; row; column);
 $(\tau\beta)_{ij}$ is interaction effect of i^{th} level of Factor A and j^{th} level of Factor B;

RCDDf		SS; $y_{ij} = \mu + \tau_i + \varepsilon_{ij}; i_{1:n}, j_{1:n}; \sum_i \tau_i = 0$		EMS		y_{11} 100.0	y_{1n} 100.0	y_{21} 010.0	y_{2n} 010.0	y_{a1} 00.01	y_{an} 00.01
Tre	a-1	$n \sum_i (\bar{y}_i - \bar{y}_{..})^2; \frac{1}{b} \sum_i y_{i.}^2 - \frac{1}{N} y_{..}^2$		$\sigma^2 + \frac{a-1}{n} \sigma_{\tau}^2$						μ_1	μ_a
E	a(n-1)	$\sum_i \sum_j^n (y_{ij} - \bar{y}_i - \bar{y}_{.j})^2; \sum_i \sum_j^n y_{ij}^2 - \frac{1}{n} \sum_i y_{i.}^2 - \frac{1}{n} \sum_j y_{.j}^2; (n-1) \sum_i S_i^2$		σ^2							
T	an-1	$\sum_i \sum_j^n (y_{ij} - \bar{y}_{..})^2; \sum_i \sum_j^n y_{ij}^2 - \frac{1}{N} y_{..}^2$		$N \times a; a \times 1$							

$\hat{\mu} = \bar{y}_{..}; \bar{y}_i - \bar{y}_{..} = \hat{\mu}_i; y_{ij} \sim N(\mu_i, \sigma_{\tau}^2); \bar{y}_i \sim (\mu_i, \frac{\sigma_{\tau}^2}{n}); y_{ij} - \bar{y}_{..} \sim (0, \frac{n-1}{n} \sigma^2); \bar{y}_i - \bar{y}_{.j} \sim (\tau_i - \tau_j, \frac{2\sigma^2}{n})$
 $X'X = nI_{a \times a}; (X'X)^{-1} = \frac{1}{n} I_{a \times a}; X'Y = [y_{1.}, \dots, y_{a.}]; \beta = [\bar{y}_{1.}, \dots, \bar{y}_{a.}]$

$CI\mu_i: \bar{y}_i \pm t_{\frac{\alpha}{2}} \sqrt{\frac{MSE}{n}}; \sqrt{\frac{MSE}{n_i}} (\text{unb})$

paired $\mu_i - \mu_j: \bar{y}_i - \bar{y}_j \pm t_{\frac{\alpha}{2}, dfE} \sqrt{\frac{2MSE}{n}}; \sqrt{\frac{MSE}{n_i} (\frac{1}{n_i} + \frac{1}{n_j})} (\text{unb});$ Fisher huge Type I
Multiple Comparison: Type I experimentwise error rate (Prob of making at least one I error when performing a set of comparisons)

Contrast $\sum_i c_i = 0; C = \sum_i c_i \bar{y}_i; E[C] = \Gamma = \sum_i c_i \mu_i; V[C] = \sigma^2 \sum_i \frac{c_i^2}{n_i \{n_i\}}$

$SS_C = \frac{n(\text{of mean})C^2}{\sum_i c_i^2} \sim \chi^2_1; \frac{C-0}{\sqrt{\frac{MSE}{n} \sum_i c_i^2}} \sim t_{N-a}; \frac{C^2}{\frac{MSE}{n} \sum_i c_i^2} \sim F_{1, N-a}$

Compare 'a' means, a-1 contrast. Orthogonal $\bar{c} \perp \bar{d}; \sum_i c_i d_i = 0$
 $\mu_1 - \mu_2 = 0; \mu_1 + \mu_2 - 2\mu_3 = 0; \mu_1 + \mu_2 + \mu_3 - 3\mu_4 = 0 \bar{c} = [1, -1, 0, 0]; \bar{d} = [1, 1, -2, 0]; \bar{e} = [1, 1, 1, -3]$
 $\sum_1^{-1} SS_C = SS_{Trt}; \sum_1^{-1} df_C = df_{Trt} = a-1; \text{Type I error rate: } 1 - (1-\alpha)^{a-1};$
Simultaneous Scheffe CI: $C \pm \sqrt{(a-1)F_{\alpha} V[C]};$ control α_e , higher Type II;

Bonferroni: $t_{\frac{\alpha}{2m}} \sqrt{\frac{2MSE}{n}}, m = \binom{a}{2} \# \text{ of comparisons; control } \alpha_e;$

Tukey: $q_{\alpha}(a, df_E) \sqrt{\frac{MSE}{n}}, a = \# \text{ of compared means, } \sqrt{\frac{MSE}{2} (\frac{1}{n_i} + \frac{1}{n_j})} (\text{unb})$

Dunnett: $d_{\alpha}(a-1, df_E) \sqrt{\frac{2MSE}{n}}$ One control group vs. each other group
Hypothesis H_0 : main: $\tau_i | \{\beta_j\} | \{\text{inter: } (\tau\beta)_{ij}\} = 0 \forall i, j; H_1$ at least one $\neq 0$.

RCBDf		SS; $y_{ij} = \mu + \tau_i + \beta_j + \varepsilon_{ij}; i_{1:n}, j_{1:b}; \sum_j \beta_j = 0$		EMS		EMS _r	
Tre	a-1	$b \sum_i (\bar{y}_i - \bar{y}_{..})^2; \frac{1}{b} \sum_i y_{i.}^2 - \frac{1}{N} y_{..}^2$		$\sigma^2 + \frac{b \sum_i \tau_i^2}{a-1}$		$\sigma^2 + b\sigma_{\tau}^2$	
B	b-1	$a \sum_j (\bar{y}_{.j} - \bar{y}_{..})^2; \frac{1}{a} \sum_j y_{.j}^2 - \frac{1}{N} y_{..}^2$		$\sigma^2 + \frac{a \sum_j \beta_j^2}{b-1}$		$\sigma^2 + a\sigma_{\beta}^2$	
E	(a-1)(b-1)	$\sum_i \sum_j^n (y_{ij} - \bar{y}_i - \bar{y}_{.j} + \bar{y}_{..})^2$		σ^2			
T	ab-1	$\sum_i \sum_j^n (y_{ij} - \bar{y}_{..})^2; \sum_i \sum_j^n y_{ij}^2 - y_{..}^2 / N$				$\hat{\sigma}_{\tau}^2 = \frac{MS_{Trt} - MSE}{n}$	

Missing: blk 1st ANOVA; Exact: full reg model & no Trt $F_0 = \frac{(SSE_{red} - SSE_{full})/r}{MSE_{full}};$

Approx: $y_{..}' = y_{..} - x_{..} y_{i.}' = y_{i.} - x_{i.} y_{.j}' = y_{.j} - x_{.j} \hat{x}; \hat{x} = \frac{ay_{i.} + by_{.j} - y_{..}}{(a-1)(b-1)}; df_E - 1; \text{LaSq: } \frac{p(y_{i.}' + y_{.j}' + y_{..}' - 2y_{..}')}{(p-2) \frac{n_i n_j (n-1)}{n}} \dots$
Unbalanced: $n_{ij} = 1$ $i_{1:a}^{th}$ Trt appears in $j_{1:b}^{th}$ Blk; $= 0$ o.w.; proportional: $n_{ij} = \frac{n_i n_j}{n_{..}}$

Yates: $V[y_{ijk}]: \hat{\sigma}^2 = \frac{\sum \sum \sum (y_{ijk} - \bar{y}_{ij.})^2}{n_{..} - ab};$ cell mean $\bar{V}(\bar{y}_{ij.}) = \frac{\sigma^2}{ab} \sum_j \sum_i \frac{1}{n_{ij}};$ ma'ly MSE, dfE: $n_{..} - ab$

Relative Efficiency

$(df_E(rcbd) + 1)(df_E(crd) + 3)MSE_{crd} \cdot (df_E(La) + 1)(df_E(rcbd) + 3)MSE_{rcbd} \cdot (df_E(La) + 1)(df_E(crd) + 3)MSE_{crd}$
 $(df_E(rcbd) + 3)(df_E(La) + 1)MSE_{rcbd} \cdot (df_E(La) + 3)(df_E(rcbd) + 1)MSE_{La} \cdot (df_E(La) + 3)(df_E(crd) + 1)MSE_{La}$
eliminate two nuisance sources of variability, blocking in two directions. the rows and columns actually represent two restrictions on randomization.

LaSqdf		SS $y_{ijk} = \mu + \tau_i + \alpha_j + \beta_k + \varepsilon_{ijk}, i, j, k: 1, p$		y_{111} 11101010	y_{123} 11001-1-1	y_{132} 1101-1011	y_{212} 10110011	y_{221} 10101110	y_{233} 101-1-1-1-1	y_{313} 1-1-110101	y_{321} 1-1-101011	y_{221} 1-1-1-1-110
Tre	p-1	$p \sum_i (\bar{y}_{i...} - \bar{y}_{....})^2; \frac{\sum_i y_{i...}^2}{p} - \frac{y_{....}^2}{p^2}$										
RC	p-1	$p \sum_{j\{k\}} (\bar{y}_{j.. \{..k\}} - \bar{y}_{....})^2; \frac{1}{p} \sum_j y_{j.. \{..k\}}^2 - \frac{1}{N} y_{..}^2$		$\frac{1}{12} B$	$\frac{1}{24} C$	$\frac{1}{24} B$	$\frac{1}{24} C$	$\frac{1}{24} B$	$\frac{1}{24} C$	$\frac{1}{24} B$	$\frac{1}{24} C$	$\frac{1}{24} B$
E	(p-1)(p-2)	$\sum_i \sum_j \sum_k (y_{ijk} - \bar{y}_{i..} - \bar{y}_{.j.} - \bar{y}_{..k} + 2\bar{y}_{....})^2$		$\frac{1}{36} C$	$\frac{1}{36} A$	$\frac{1}{36} C$	$\frac{1}{36} A$	$\frac{1}{36} C$	$\frac{1}{36} A$	$\frac{1}{36} C$	$\frac{1}{36} A$	$\frac{1}{36} C$
T	$p^2 - 1$	$\sum_p \sum_p \sum_p (y_{ijk} - \bar{y}_{....})^2; 9 \times 7; 7 \times 1$										

rep123	df	SS												
Tre	p-1	$n p \sum_i^p (\bar{y}_{i...} - \bar{y}_{....})^2; \frac{\sum_i^p y_{i...}^2}{np} - \frac{y_{....}^2}{np^2}$												
RC	p-1	$n p \sum_{j\{k\}}^p (\bar{y}_{j.. \{..k\}} - \bar{y}_{....})^2; \frac{1}{np} \sum_p y_{j.. \{..k\}}^2 - \frac{1}{N} y_{..}^2$												
Rep-RC	n(p-1)	$p \sum_l^p \sum_{j\{k\}}^p (\bar{y}_{j..l \{..kl\}} - \bar{y}_{....})^2; \frac{1}{p} \sum_p \sum_p y_{j..l \{..kl\}}^2 - \frac{1}{p^2} y_{..}^2$												
Rep	n-1	$p^2 \sum_l^n (\bar{y}_{l...} - \bar{y}_{....})^2; \frac{1}{p^2} \sum_p y_{l...}^2 - \frac{1}{N} y_{..}^2$												
E		(p-1)(np+n-3);(p-1)(np-2);(p-1)(np-n-1)												
T	np ² -1	Rep each point; change R C; change R&C												
Graeco-Latin			$y_{ijkl} = \mu + \tau_i + \gamma_j + \alpha_k + \beta_l + \varepsilon_{ijkl}, i, j, k, l: 1, p$											
Gr;Lap	p-1	$p \sum_{i\{j\}}^p (\bar{y}_{i... \{..j\}} - \bar{y}_{....})^2$	Aα Bβ	Cγ Dδ	a1 b2 c3	d4 e5								
R;C	p-1	$p \sum_{k\{l\}}^p (\bar{y}_{..k \{..l\}} - \bar{y}_{....})^2$	Bδ Aγ	Dβ Cα	b3 c4 d5	e1 a2								
E	(p-1)(p-3)	SS _T - SS	Cβ Dα	Aδ Bγ	c5 d1	e2 a3 b4								
T	p ² - 1	$\sum \sum \sum \sum (y_{ijkl} - \bar{y}_{....})^2$	Dγ Cδ	Bα Aβ	d2 e3 a4	b5 c1	e4 a5 b1	c2 d3						

[illegible]

$$\begin{aligned} \beta_0 &= \hat{\beta}_1 = \frac{\sum x_i^2 \sum (1+x_i) y_i - \sum x_i^2 (1+x_i) \sum x_i^2 y_i}{\sum x_i^2 \sum (1+x_i)^2 - [\sum x_i^2 (1+x_i)]^2}; \hat{\beta}_2 = \frac{\sum (1+x_i)^2 \sum x_i^2 y_i - \sum x_i^2 (1+x_i) \sum (1+x_i) y_i}{\sum x_i^2 \sum (1+x_i)^2 - [\sum x_i^2 (1+x_i)]^2} \\ y_i &= \beta_1 (1+x_i) + \beta_2 x_i^2 + \varepsilon_i \\ \beta_1 &= 0; \hat{\beta}_2 = 2\hat{\beta}_0 = \frac{2 \sum_{i=1}^n (1+2x_i^2) y_i}{\sum_{i=1}^n (1+2x_i^2)^2}; y_i = \beta_0 (1+2x_i^2) + \varepsilon_i \\ \beta_1 &= 1; \hat{\beta}_0 = \hat{\beta}_2 = \frac{\sum_{i=1}^n (1+x_i^2) (y_i - x_i)}{\sum_{i=1}^n (1+x_i^2)^2}; y_i - x_i = \beta_2 (1+x_i^2) + \varepsilon_i \\ \hat{\beta}_1 &= 1; \hat{\beta}_0 = 2\hat{\beta}_2 = \frac{2 \sum_{i=1}^n (2+x_i^2) (y_i - x_i)}{\sum_{i=1}^n (2+x_i^2)^2} \\ \beta_0 &= 1; \hat{\beta}_1 = \hat{\beta}_2 = \frac{\sum_{i=1}^n (x_i + x_i^2) (y_i - 1)}{\sum_{i=1}^n (x_i + x_i^2)^2}; y_i - 1 = \beta_2 (x_i + x_i^2) + \varepsilon_i \\ \beta_1 &= 0; \hat{\beta}_0 = 2\hat{\beta}_2 = \frac{2 \sum_{i=1}^n (2+x_i^2) y_i}{\sum_{i=1}^n (2+x_i^2)^2}; y_i = \beta_2 (2+x_i^2) + \varepsilon_i \\ \beta_0 &= 0; \hat{\beta}_1 = \frac{\sum_{i=1}^n x_i^2 \sum_{i=1}^n x_i y_i - \sum_{i=1}^n x_i^2 \sum_{i=1}^n x_i^2 y_i}{\sum_{i=1}^n x_i^4 \sum_{i=1}^n x_i^2 - (\sum_{i=1}^n x_i^3)^2}; \hat{\beta}_2 = \frac{\sum_{i=1}^n x_i^2 \sum_{i=1}^n x_i^2 y_i - \sum_{i=1}^n x_i^3 \sum_{i=1}^n x_i y_i}{\sum_{i=1}^n x_i^4 \sum_{i=1}^n x_i^2 - (\sum_{i=1}^n x_i^3)^2} \\ y_i &= \beta_1 x_i + \beta_2 x_i^2 + \varepsilon_i \\ y_i - 3 &= \beta_1 x_i + \varepsilon_i; \hat{\beta} = \frac{\sum (y+3)x_i}{\sum x_i^2}; E[\hat{\beta}] = \beta, V[\hat{\beta}] = \frac{\sigma^2}{\sum x_i^2}; \text{sel}[\hat{\beta}] = \frac{\sigma^2}{\sum x_i^2} \\ \text{SSE} &= \sum (y+3 - \beta x)^2 = \sum (y+3)^2 - 2\beta \sum (y+3)x + \beta^2 \sum x^2 = \sum (y+3)^2 - \beta^2 \sum x^2 \\ E[\hat{\beta}] &= V[\sum (y+3)] + E[\sum (y+3)]^2 - \sum x^2 (E[\beta]^2 + V[\beta]) = n\sigma^2 + \beta \sum x^2 \sum x^2 (\beta^2 + \frac{\sigma^2}{\sum x_i^2}) = (n-1)\sigma^2 \end{aligned}$$

$(\theta)_{ijkl} = \mu + \tau_i + \beta_{j(i)} + \gamma_k + \tau\gamma_{ik} + \beta\gamma_{jk(i)} + \varepsilon_{(ijk)l}$ where $\varepsilon_{(ijk)l} \stackrel{iid}{\sim} N(0, \sigma^2)$; $\beta_{j(i)} \stackrel{iid}{\sim} N(0, \sigma_\beta^2)$; $\beta\gamma_{jk(i)} \stackrel{iid}{\sim} N(0, \frac{(a-1)}{a} \sigma_\beta^2 \sigma_\gamma^2)$ $\sum_{i=1}^a \tau_i = 0$; $\sum_{k=1}^c \gamma_k = 0$; $\sum_{i=1}^a \tau\gamma_{ik} = \sum_{k=1}^c \tau\gamma_{ik} = 0$ An exact F test is available for each factor. The SS and df for AB go to B(A) and for ABC to BC(A).													
Factor	F_{15}^R	F_{15}^R	F_{15}^R	F_{15}^R	F_{15}^R	EMS	F_{den}	Source	SS	df	MS	F	Pr(>F)
τ_i	0	5	2	3		$15 \sum_i \tau_i^2 + 6\sigma_\tau^2 + \sigma^2$	B(A)	A	17	2	8.5	$\frac{8.5}{4.75} = 1.79$.209
$\beta_{j(i)}$	1	1	2	3		$6\sigma_\beta^2 + \sigma^2$	E	B(A)	57	124	75	$\frac{4.75}{2.5} = 2.07$.033
γ_k	3	5	0	3		$45 \sum_k \gamma_k^2 + 3\sigma_\gamma^2 + \sigma^2$	B(A)C	C	4	1	4	$\frac{4}{2} = 2$.183
$\tau\gamma_{ik}$	0	5	0	3		$7.5 \sum_i \sum_k \tau\gamma_{ik}^2 + 3\sigma_\tau^2 + \sigma^2$	B(A)C	AC	5	2	2.5	$\frac{2.5}{2.5} = 1.25$.321
$\beta\gamma_{jk(i)}$	1	1	0	3		$3\sigma_\beta^2 + \sigma^2$	E	BC(A)	24	12	2	$\frac{2}{2.5} = 0.87$.581
$\varepsilon_{(ijk)l}$	1	1	1	1		σ^2	Total	Err	138	60	2.3		
								Total	324	589			

Factor	F_{15}^R	F_{15}^R	F_{15}^R	F_{15}^R	F_{15}^R	EMS	F_{den}	Source	SS	df	MS	F	Pr(>F)
τ_i	1	3	4	2		$24\sigma_\tau^2 + \sigma^2$	MSE	A	34	4	8.5	$\frac{8.5}{2.3} = 3.70$.009
$\beta_{j(i)}$	1	0	4	2		$8\sigma_\beta^2 + \sigma^2$	MSE	B(A)	44	10	4.4	$\frac{4.4}{2.3} = 1.91$.06
γ_k	5	3	0	2	$\frac{30}{3} \sum_k \gamma_k^2 + 6\sigma_\gamma^2 + \sigma^2$	MSAC	C	24	3	8	$\frac{8}{2.5} = 3.2$.06	
$\tau\gamma_{ik}$	1	3	0	2		$6\sigma_\tau^2 + \sigma^2$	MSE	AC	30	12	2.5	$\frac{2.5}{2.3} = 1.09$.385
$\beta\gamma_{jk(i)}$	1	0	0	2		$2\sigma_\beta^2 + \sigma^2$	MSE	BC(A)	54	30	1.8	$\frac{1.8}{2.3} = 0.78$.768
$\varepsilon_{(ijk)l}$	1	1	1	1		σ^2	Total	Err	138	60	2.3		
								Total	324	119			

(df)	SS_F	$[-3, 4]$	SS_{12}	$[-2, 4]$	$SS_{1,3,5}$	α	2	110	110	0	110	110	0	SS_{γ}	110	110	0	SS_{γ}	110	110	0
R	36(5)	-96(3)	271(2)	-153(2)	214(3)	β	11	160	160	0	160	160	0	$SS_{\alpha\gamma}$	160	160	0	SS_{γ}	160	160	0
E	336(69)	+96(3)	432(72)	+153(2)	489(71)	β	250	10	240	240	0	240	0	$SS_{\alpha\gamma}$	240	240	0	SS_{γ}	240	240	0
T	703(74)					β	2230	120	420	420	250	550	110	$SS_{\alpha\gamma}$	550	550	110	SS_{γ}	550	550	110

$$F = \frac{\frac{(SSE_{red} - SSE_{full})}{(dfe_{red} - dfe_{full})}}{MSE_{full}} = \frac{\frac{432 - 336}{72 - 69}}{\frac{336}{69}} = \frac{32}{4.87} = 6.57 \sim F_{3,69}$$

Calculate R^2 and R^2_{adj} for $Y_i = \beta_0 + \beta_1 X_{1i} + \beta_2 X_{2i} + \varepsilon_i$.

$$R^2 = \frac{SSR}{SST} = \frac{272}{703} = .39$$

$$R^2_{adj} = 1 - \frac{MSE}{MST} = 1 - \frac{SSE}{SST} \frac{n-1}{n-p} = 1 - \frac{431}{703} \frac{74}{72} = .37$$

Calculate the F-stat for testing $H_0 : \beta_2 = \beta_4 = 0$.

The model under the null hypothesis is given by Model B. That makes the F-stat:

$$F = \frac{\frac{489 - 336}{336}}{\frac{69}{4.87}} = \frac{76.5}{14.27} = 5.36 \sim F_{2,69}$$

$$\begin{aligned}
y_{ijk} &= \mu + \tau_{i,2} + \beta_{j(i)1,2,1,2,3} + \varepsilon_{(ij)k} \quad k=1,3; i=1,2; j=1,3; i=1,3; i=1,3 \\
y_{ijk} &= \mu + \tau_i + \gamma_{ij} + \varepsilon_{ijk} \quad 27 \times 9 \\
y_{ijk} &= \mu + \tau_i + \beta_j + (\tau\beta)_{ij} + \varepsilon_{ijk}; i,j=1,3; k=1,2,3; n_{ij} = 20 \times 9
\end{aligned}$$

$$\begin{aligned} ssa &= 5(\bar{y}_{1..} - \bar{y}_{..})^2 + 10(\bar{y}_{2..} - \bar{y}_{..})^2; \text{sst} = \sum (y_{ijk} - \bar{y}_{..})^2 \\ ssb &= 3(\bar{y}_{11.} - \bar{y}_{1..})^2 + 2(\bar{y}_{12.} - \bar{y}_{1..})^2 + 4(\bar{y}_{21.} - \bar{y}_{2..})^2 + 3(\bar{y}_{22.} - \bar{y}_{2..})^2 + 3(\bar{y}_{23.} - \bar{y}_{2..})^2 \\ ssc &= \sum_k [(y_{11k} - \bar{y}_{11.})^2 + (y_{12k} - \bar{y}_{12.})^2 + (y_{21k} - \bar{y}_{21.})^2 + (y_{22k} - \bar{y}_{22.})^2 + (y_{23k} - \bar{y}_{23.})^2] \\ dft = n - 1; \text{dfE} &= n - b; \text{dfB(A)} = b - 1; \text{dFA} = a - 1 \\ 0.0.0.0.0.1.1.1.1.0.0.0. -1. -1. -1. &= c1_{in} a2; 1.1.1.1.1.0.0.0.0.0.0.0.0.0.0. = a1 \\ 0.0.0.0.0.0.0.0.0.1.1. -1. -1. -1. &= c2_{in} a2; 1.1.1.0.0.0.0.0.0.0.0.0.0.0.0. = b1_{in} a1 \\ E[MS_{A_f}] &= \frac{bn}{a-1} \sum_i E[(\bar{y}_{i..} - \bar{y}_{..})^2] = \frac{bn}{a-1} \sum_i E[\tau_i^2 + V] = \frac{bn}{a-1} \sum_i \tau_i^2 + \sigma^2 \\ E(MS_{A_r B_r}) &= E\left[\frac{\sum_i \sum_j \sum_k (\bar{\tau}_{\beta ij})^2}{(a-1)(b-1)}\right] = \frac{n \sum_i \sum_j E[(\bar{y}_{ij.} - \bar{y}_{i..} - \bar{y}_{.j.} + \bar{y}_{..})^2]}{(a-1)(b-1)} = \frac{n \sum_i \sum_j E[V] + 0}{(a-1)(b-1)} = n\sigma_{\tau\beta}^2 + \sigma^2 \end{aligned}$$