

STAT 661: Project

LS v.s. EM

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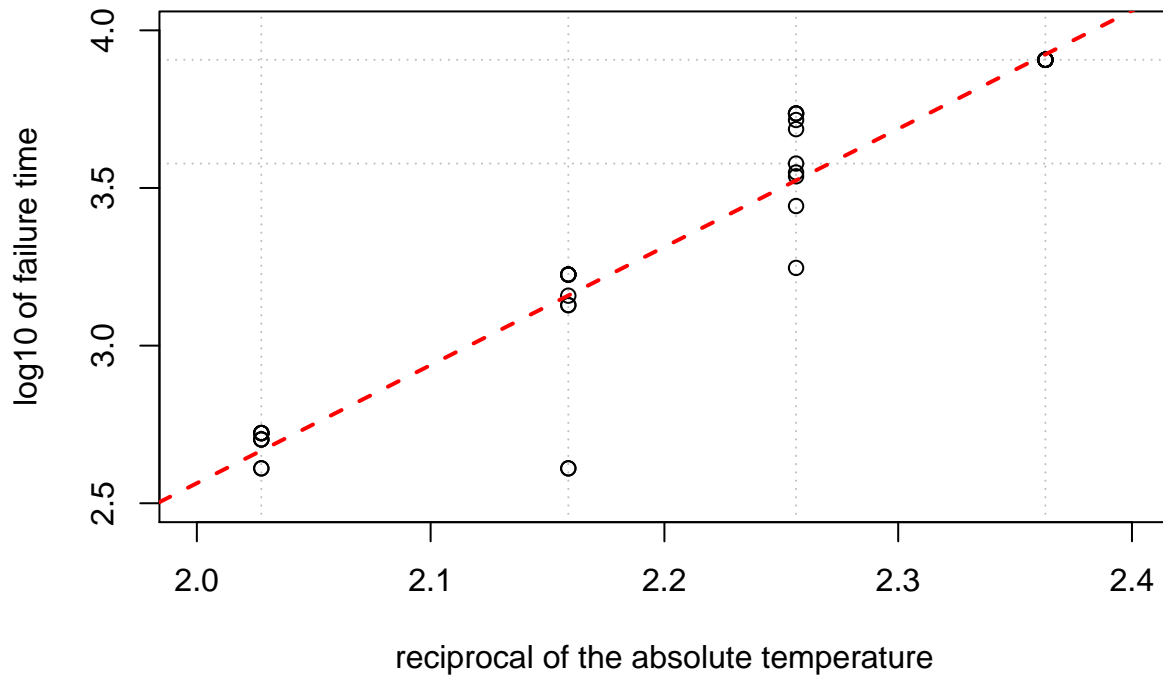
1 Least Square Method v.s. EM Method

2 Appendix

2.1 LS Code

```
temp <- c(150,170,190,220) #temperature levels
trec <- 1000/(temp+273.2) #reciprocal of the absolute temperature T
x <- c(rep(trec[1],10),rep(trec[2],10),rep(trec[3],10),rep(trec[4],10))
cen <- c(8064,5448,1680,528) #censoring times
logcen <- log10(cen) #log10 censoring times
y_uncensored <- log10(c(rep(1,10),
                        1764,2772,3444,3542,3780,4860,5196,rep(1,3),
                        408,408,1344,1344,1440,rep(1,5),
                        408,408,504,504,504,rep(1,5)))
y_censored <- c(rep(logcen[1],10),
                rep(0,7),rep(logcen[2],3),
                rep(0,5),rep(logcen[3],5),
                rep(0,5),rep(logcen[4],5))
S <- 23; Y<-matrix(nrow=S,ncol=40)
Y[1,] <- y_0 <- y_uncensored+y_censored
fit0 <- lm(y_0~x) #linear model between log10 of observed life time
```

Scatterplot



```
# Iteration 0
sigma_0 <- sigma(fit0) #standard error of residuals
beta_00 <- coef(fit0)[1] #intercept
beta_10 <- coef(fit0)[2] #slope
mu_0 <- beta_00 + beta_10*trec #mean log time to failure
z <- (logcen-mu_0)/sigma_0 #z-vector
ex_mu_0 <- mu_0 + sigma_0*dnorm(z)/(1-pnorm(z)) #new expected mean log times to failure

delta = 1e-006; iteration <- 1

PHI<-matrix(nrow=S,ncol=8,dimnames=list(NULL,
  c('mu150','mu170','mu190','mu220','Intercept','Slope','Sigma','Iteration'))))

PHI[1,]<-phi<-c(ex_mu_0, beta_00, beta_10,sigma_0,iteration)
# Subsequent iteration
repeat {
  phi[8] <- phi[8]+1
  y_censored <- c(rep(phi[1],10),
    rep(0,7),rep(phi[2],3),
    rep(0,5),rep(phi[3],5),
    rep(0,5),rep(phi[4],5))

  y<- y_ uncensored+y_censored
  Y[phi[8],]<-y # Replace the new censored values
  fit <- lm(y~x) # fit a new model
  phi[5] <- coef(fit)[1] #intercept
  phi[6] <- coef(fit)[2] #slope
  phi[7] <- sigma(fit) #standard error of residuals
  mu <- phi[5] + phi[6]*trec
```

```

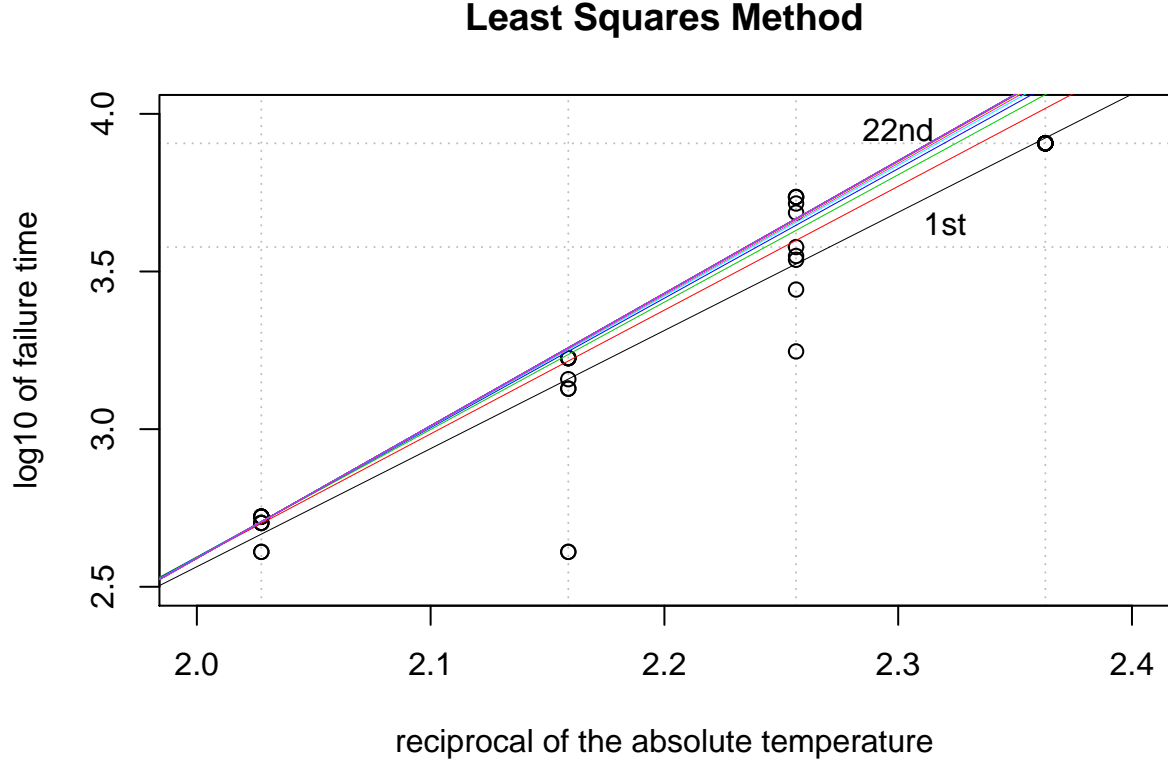
z <- (logcen-mu)/phi[7] #z-vector
#new expected mean log times to failure
phi[1:4] <- mu + phi[7]*dnorm(z)/(1-pnorm(z))
conv <- dist(rbind(PHI[phi[8]-1,1:4],phi[1:4]))
if(conv < delta) break
PHI[phi[8],]<-phi
}

```

2.2 LS Results

mu150	mu170	mu190	mu220	Intercept	Slope	Sigma	Iteration
4.038	3.809	3.329	2.83	-4.931	3.747	0.1572	1
4.099	3.84	3.366	2.858	-5.26	3.926	0.1799	2
4.131	3.856	3.382	2.869	-5.486	4.04	0.1911	3
4.149	3.865	3.39	2.874	-5.623	4.108	0.1969	4
4.159	3.87	3.395	2.877	-5.704	4.148	0.2001	5
4.165	3.873	3.397	2.878	-5.752	4.172	0.2019	6
4.168	3.874	3.399	2.879	-5.779	4.185	0.2029	7
4.17	3.875	3.4	2.879	-5.795	4.193	0.2035	8
4.171	3.876	3.4	2.879	-5.805	4.198	0.2039	9
4.172	3.876	3.401	2.88	-5.81	4.2	0.2041	10
4.172	3.877	3.401	2.88	-5.814	4.202	0.2042	11
4.173	3.877	3.401	2.88	-5.815	4.203	0.2042	12
4.173	3.877	3.401	2.88	-5.817	4.203	0.2043	13
4.173	3.877	3.401	2.88	-5.817	4.204	0.2043	14
4.173	3.877	3.401	2.88	-5.818	4.204	0.2043	15
4.173	3.877	3.401	2.88	-5.818	4.204	0.2043	16
4.173	3.877	3.401	2.88	-5.818	4.204	0.2043	17
4.173	3.877	3.401	2.88	-5.818	4.204	0.2043	18
4.173	3.877	3.401	2.88	-5.818	4.204	0.2043	19
4.173	3.877	3.401	2.88	-5.818	4.204	0.2043	20
4.173	3.877	3.401	2.88	-5.818	4.204	0.2043	21
4.173	3.877	3.401	2.88	-5.818	4.204	0.2043	22
NA	NA	NA	NA	NA	NA	NA	NA

2.3 LS figure



2.4 EM Method

- E-step

$$Q(\vec{\theta}, \vec{\theta}^*) = -\frac{n}{2} \ln(2\pi) - n \ln(\sigma) - \frac{1}{2\sigma^2} \sum_{j=1}^m (t_j - \beta_0 - \beta_1 \nu_j)^2 - \frac{1}{2\sigma^2} \sum_{i=m+1}^n E[(T_i - \beta_0 - \beta_1 \nu_i)^2 | T_i > w_i, \vec{\theta}^*]$$

$$E[T_i | T_i > w_i, \vec{\theta}^*] = \mu_i^* + \sigma^* H\left(\frac{w_i - \mu_i^*}{\sigma^*}\right)$$

$$E[T_i^2 | T_i > w_i, \vec{\theta}^*] = \mu_i^{*2} + \sigma^{*2} + \sigma^* (w_i + \mu_i^*) H\left(\frac{w_i - \mu_i^*}{\sigma^*}\right)$$

$$E[(T_i - \beta_0 - \beta_1 \nu_i)^2 | T_i > w_i, \vec{\theta}^*] = \mu_i^{*2} + \sigma^{*2} + \sigma^* (w_i + \mu_i^*) H\left(\frac{w_i - \mu_i^*}{\sigma^*}\right) - 2(\beta_0 + \beta_1 \nu_i) [\mu_i^* + \sigma^* H\left(\frac{w_i - \mu_i^*}{\sigma^*}\right)] + (\beta_0 + \beta_1 \nu_i)^2$$

- M-step

$$\frac{\partial Q}{\partial \beta_0} = -\frac{1}{\sigma^2} \left\{ \sum_{j=1}^m [t_j - \beta_0 - \beta_1 \nu_j] + \sum_{i=m+1}^n [\mu_i^* + \sigma^* H\left(\frac{w_i - \mu_i^*}{\sigma^*}\right) - \beta_0 - \beta_1 \nu_i] \right\} = 0$$

$$\begin{aligned}\frac{\partial Q}{\partial \beta_1} &= -\frac{1}{\sigma^2} \left\{ \sum_{j=1}^m [t_j - \beta_0 - \beta_1 \nu_j] \nu_j + \sum_{i=m+1}^n [\mu_i^* + \sigma^* H(\frac{w_i - \mu_i^*}{\sigma^*}) - \beta_0 - \beta_1 \nu_i] \nu_i \right\} = 0 \\ \frac{\partial Q}{\partial \sigma^2} &= \frac{1}{2\sigma^2} \left\{ -n + \frac{1}{\sigma^2} \sum_{j=1}^m [t_j - \beta_0 - \beta_1 \nu_j]^2 + \frac{1}{\sigma^2} \sum_{i=m+1}^n E[(T_i - \beta_0 - \beta_1 \nu_i)^2 | T_i > w_i, \vec{\theta}^*] \right\} = 0 \\ \sum_{j=1}^m [t_j - \beta_0 - \beta_1 \nu_j]^2 + \sum_{i=m+1}^n \left\{ \mu_i^{*2} + \sigma^{*2} + \sigma^* (w_i + \mu_i^* - 2\mu_i) H(\frac{w_i - \mu_i^*}{\sigma^*}) - 2\mu_i \mu_i^* + \mu_i^2 \right\} &= n\sigma^2\end{aligned}$$

2.5 the EM algorithm's pseudo code

Algorithm 1: EM algorithm

input : observed data $\mathcal{D} = \{\vec{x}_1, \vec{x}_2, \dots, \vec{x}_N\}$, joint distribution $P(\vec{x}, \vec{z} | \vec{\theta})$

output: model's parameters $\vec{\theta}$

// 1. identify hidden variables \vec{z} , write out the log likelihood function $\ell(\vec{x}, \vec{z} | \vec{\theta})$

$\vec{\theta}^{(0)} = \dots$ // initialize

while (!convergency) **do**

 // 2. E-step: plug in $P(\vec{x}, \vec{z} | \vec{\theta})$, derive the formula of $Q(\vec{\theta}, \vec{\theta}^{t-1})$

$Q(\vec{\theta}, \vec{\theta}^{t-1}) = \mathbb{E} [\ell_c(\vec{\theta}) | \mathcal{D}, \vec{\theta}^{t-1}]$

 // 3. M-step: find $\vec{\theta}$ that maximizes the value of $Q(\vec{\theta}, \vec{\theta}^{t-1})$

$\vec{\theta}^t = \arg \max_{\vec{\theta}} Q(\vec{\theta}, \vec{\theta}^{t-1})$

2.6 EM Code

```
temp <- c(150,170,190,220) #temperature levels
trec <- 1000/(temp+273.2) #reciprocal of the absolute temperature T
nu <- c(rep(trec[1],10),rep(trec[2],10),rep(trec[3],10),rep(trec[4],10))
# index_nu <- c(rep(-2,7),rep(-3,5),rep(-4,5),rep(1,10),rep(2,3),rep(3,5),rep(4,5))

cen <- c(8064,5448,1680,528) #censoring times
w <- log10(cen) #log10 censoring times: last time still working
y_uncensored <- log10(c(rep(1,10),
                        1764,2772,3444,3542,3780,4860,5196,rep(1,3),
                        408,408,1344,1344,1440,rep(1,5),
                        408,408,504,504,504,rep(1,5)))
y_censored <- c(rep(w[1],10),
                rep(0,7),rep(w[2],3),
                rep(0,5),rep(w[3],5),
                rep(0,5),rep(w[4],5))

S <- 30
Y<-matrix(nrow=S,ncol=40)
Y[1,] <- y <- (y_uncensored+y_censored)

index_censored <- which(y %in% w)
```

```

index_uncensored <- which(!(y %in% w))

m <- length(index_uncensored)
n<-length(index_censored)+m

# initial value
fit0 <- lm(y~nu) #linear model
sigma_0 <- sigma(fit0) #standard error of residuals
beta0_0 <- as.numeric(coef(fit0)[1]) #intercept
beta1_0 <- as.numeric(coef(fit0)[2]) #slope

mu_i0 <- beta0_0 + beta1_0 * nu[index_censored]
mu_j0 <- beta0_0 + beta1_0 * nu[index_uncensored]
D_j <- y[index_uncensored]-mu_j0

beta0_star <- beta0_0
beta1_star<- beta1_0
sigma_star<- sigma_0 # should be replaced by solutions

par=c(beta0_star,beta1_star,sigma_star)

# define mu_i function
mu_i_F <- function (beta0,beta1) {
  beta0+beta1*nu[index_censored]
}

k=1

# define H function
HF <- function (beta0,beta1,sigma) {
  dnorm((y_censored[index_censored]-mu_i_F(beta0,beta1))/sigma) / (
    1-pnorm((y_censored[index_censored]-mu_i_F(beta0,beta1))/sigma)
  )
}

ET <- mu_i_F(par[1],par[2])+ par[3]*HF(par[1],par[2],par[3])
ET_sq <- mu_i_F(par[1],par[2])^2+
  sigma_star^2+
  sigma_star*(y[index_censored]+mu_i_F(par[1],par[2]))*HF(par[1],par[2],par[3])
ER <- ET_sq-2*mu_i0*ET+mu_i0^2

Q <- 0
Q[1] <- -n*log(2*pi)/2-n*log(sigma_0)-(1/(2*sigma_0^2))*(sum((D_j)^2)+sum(ER))

THETA<-matrix(nrow=S,ncol=8,dimnames=list(NULL,
  c('Intercept','Slope','Sigma','mu150','mu170','mu190','mu220','Iteration')))
THETA[1,]<-c(par,unique(ET),k)
delta = 1e-6
repeat {
  # load THETA
  beta0 <- THETA[k,1]

```

```

beta1 <- THETA[k,2]
sigma <- THETA[k,3]

# Update y
y[index_censored] <- ET
y[index_uncensored] <- y_uncensored[index_uncensored]

# E step

mu_i <- beta0 + beta1 * nu[index_censored]
mu_j <- beta0 + beta1 * nu[index_uncensored]
D_j <- y[index_uncensored] - mu_j # difference between uncensored y and mean_j

Q_theta <- function(par){
  -n*log(2*pi)/2-
  n*log(sigma)-
  (1/(2*sigma^2))*sum((D_j)^2)-
  (1/(2*sigma^2))*sum(
    (mu_i_F(par[1],par[2]))^2+
    par[3]*(
      y[index_censored]+
      mu_i_F(par[1],par[2])-
      2*mu_i
    )*HF(par[1],par[2],par[3])-
    2*mu_i*mu_i_F(par[1],par[2])+
    mu_i^2
  )+
  (1/(2*sigma^2))*(n-m)*(par[3]^2)
}

# Get Q
Q[k]<- Q_theta(par)

# M step

##### from Jacob
# par <- optim(par,
#             Q_theta,
#             control=list(fnscale=-1),
#             method="L-BFGS-B",
#             lower=c(-10,2,0.1),
#             upper=c(-2,10,1))$par

##### cheat
fit <- lm(y~nu) # fit a new model
par[1] <- coef(fit)[1] #intercept
par[2] <- coef(fit)[2] #slope
par[3] <- sigma(fit)
#####

k <- k+1

```

```

# Update ET
ET <- mu_i_F(par[1],par[2])+ par[3]*HF(par[1],par[2],par[3])

# Update Q
Q[k]<- Q_theta(par)

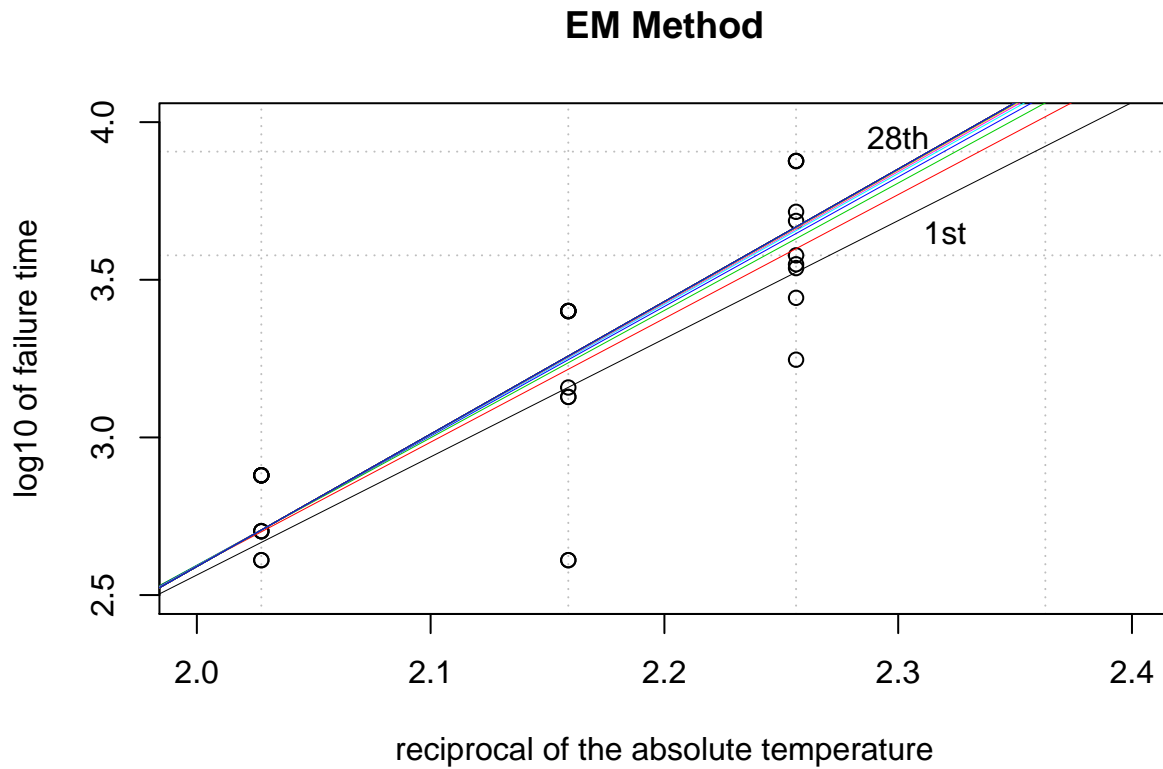
# Update THETA
THETA[k,1] <- par[1]
THETA[k,2] <- par[2]
THETA[k,3] <- par[3]
THETA[k,4:7]<-unique(ET)
THETA[k,8] <- k

# Update Y
Y[k,] <- y

# par=c(beta0,beta1,sigma)
if(abs(Q[k]-Q[k-1])<=delta) break
}

```

2.7 EM figure



2.8 The results of THETA

Intercept	Slope	Sigma	mu150	mu170	mu190	mu220	Iteration
-4.931	3.747	0.1572	4.038	3.809	3.329	2.83	1

Intercept	Slope	Sigma	mu150	mu170	mu190	mu220	Iteration
-5.26	3.926	0.1799	4.099	3.84	3.366	2.858	2
-5.486	4.04	0.1911	4.131	3.856	3.382	2.869	3
-5.623	4.108	0.1969	4.149	3.865	3.39	2.874	4
-5.704	4.148	0.2001	4.159	3.87	3.395	2.877	5
-5.752	4.172	0.2019	4.165	3.873	3.397	2.878	6
-5.779	4.185	0.2029	4.168	3.874	3.399	2.879	7
-5.795	4.193	0.2035	4.17	3.875	3.4	2.879	8
-5.805	4.198	0.2039	4.171	3.876	3.4	2.879	9
-5.81	4.2	0.2041	4.172	3.876	3.401	2.88	10
-5.814	4.202	0.2042	4.172	3.877	3.401	2.88	11
-5.815	4.203	0.2042	4.173	3.877	3.401	2.88	12
-5.817	4.203	0.2043	4.173	3.877	3.401	2.88	13
-5.817	4.204	0.2043	4.173	3.877	3.401	2.88	14
-5.818	4.204	0.2043	4.173	3.877	3.401	2.88	15
-5.818	4.204	0.2043	4.173	3.877	3.401	2.88	16
-5.818	4.204	0.2043	4.173	3.877	3.401	2.88	17
-5.818	4.204	0.2043	4.173	3.877	3.401	2.88	18
-5.818	4.204	0.2043	4.173	3.877	3.401	2.88	19
-5.818	4.204	0.2043	4.173	3.877	3.401	2.88	20
-5.818	4.204	0.2043	4.173	3.877	3.401	2.88	21
-5.818	4.204	0.2043	4.173	3.877	3.401	2.88	22
-5.818	4.204	0.2043	4.173	3.877	3.401	2.88	23
-5.818	4.204	0.2043	4.173	3.877	3.401	2.88	24
-5.818	4.204	0.2043	4.173	3.877	3.401	2.88	25
-5.818	4.204	0.2043	4.173	3.877	3.401	2.88	26
-5.818	4.204	0.2043	4.173	3.877	3.401	2.88	27
-5.818	4.204	0.2043	4.173	3.877	3.401	2.88	28
NA	NA	NA	NA	NA	NA	NA	NA
NA	NA	NA	NA	NA	NA	NA	NA

2.9 The results of Q

20.08, 21.4, 20.64, 20.09, 19.75, 19.55, 19.43, 19.36, 19.32, 19.29, 19.28, 19.27, 19.27, 19.26, 19.26, 19.26, 19.26, 19.26, 19.26, 19.26, 19.26, 19.26, 19.26, 19.26, 19.26, 19.26 and 19.26

3 Reference

Schmee, J., & Hahn, G. (1979). A Simple Method for Regression Analysis with Censored Data. *Technometrics*, 21(4), 417-432. doi:10.2307/1268280

Aitkin, M. (1981). A Note on the Regression Analysis of Censored Data. *Technometrics*, 23(2), 161-163. doi:10.2307/1268032