

# 1 Kernels

1. let  $(x, y) \in \mathbb{R}^+ \times \mathbb{R}^+$ , where  $\mathbb{R}^+ = \{x \in \mathbb{R}; x \geq 0\}$ , the “french positive” real numbers.

- (a) Verify that  $\min(x, y) = \int_0^\infty \mathbb{I}_{t \leq x} \mathbb{I}_{t \leq y} dt$  where  $\mathbb{I}_A = \begin{cases} 1 & \text{if A is true} \\ 0 & \text{otherwise} \end{cases}$

When  $x \leq y$ ,

$$\begin{aligned} \int_0^\infty \mathbb{I}_{t \leq x} \mathbb{I}_{t \leq y} dt &= \int_0^x \mathbb{I}_{t \leq x} \mathbb{I}_{t \leq y} dt + \int_x^y \mathbb{I}_{t \leq x} \mathbb{I}_{t \leq y} dt + \int_y^\infty \mathbb{I}_{t \leq x} \mathbb{I}_{t \leq y} dt \\ &= \int_0^x 1 \cdot 1 dt + \int_x^y 0 \cdot 1 dt + \int_y^\infty 0 \cdot 0 dt = x \end{aligned}$$

By the same way, when  $y \leq x$ ,  $\int_0^\infty \mathbb{I}_{t \leq x} \mathbb{I}_{t \leq y} dt = y$ .

Therefore,  $\min(x, y) = \int_0^\infty \mathbb{I}_{t \leq x} \mathbb{I}_{t \leq y} dt$

- (b) Use the previous question to show that  $K(x, y) = \min(x, y)$  is a pd kernel over  $\mathbb{R}^+$

$$K(x, y) = \min(x, y) = \int_0^\infty \mathbb{I}_{t \leq x} \mathbb{I}_{t \leq y} dt = \min(y, x) = K(y, x) \text{ symmetric}$$

$$\sum_{i=1}^n \sum_{j=1}^n \alpha_i \alpha_j \min(x, y) = \int_0^\infty \sum_{i=1}^n \alpha_i \mathbb{I}_{t \leq x} \sum_{j=1}^n \alpha_j \mathbb{I}_{t \leq y} dt = \int_0^\infty (\sum_{i=1}^n \alpha_i \mathbb{I}_{t \leq x})^2 dt \geq 0$$

- (c) Show that  $\max(x, y)$  is not a pd kernel over  $\mathbb{R}^+$ .

$$\text{When } x \leq y, \int_0^\infty \mathbb{I}_{t \geq x} \mathbb{I}_{t \geq y} dt = \int_y^\infty 1 \cdot 1 dt = t|_y^\infty \neq \max(x, y)$$

The Gram Matrix

$$M(x, y) = \begin{bmatrix} \max(x, x) & \max(x, y) \\ \max(x, y) & \max(y, y) \end{bmatrix} = \begin{bmatrix} x & y \\ y & y \end{bmatrix} = xy - y^2 \leq 0$$

When  $y \leq x$  is the same.  $M(x, y) = xy - x^2 \leq 0$

Therefore,  $\max(x, y)$  can not be a pd kernel over  $\mathbb{R}^+$

2. Consider a probability space  $(\Omega, \mathcal{A}, P)$

- (a) Define for any two events  $A$  and  $B$ ,  $K_1(A, B) = P(A \cap B)$  where  $A \cap B$  is the intersection between the events A and B Verify that  $K_1$  is positive definite. Hint:  $P(A) = E[\mathbb{I}_A]$

$$K_1(A, B) = P(A \cap B) = P(B \cap A) = K_1(B, A) \text{ symmetric}$$

$$P(A) = E[\mathbb{I}_A]; P(B) = E[\mathbb{I}_B]; P(A \cap B) = E[\mathbb{I}_A \mathbb{I}_B]$$

$$\sum_{i=1}^n \sum_{j=1}^n \alpha_i \alpha_j E[\mathbb{I}_{A_i} \mathbb{I}_{A_j}] = E[\sum_{i=1}^n \sum_{j=1}^n \alpha_i \alpha_j \mathbb{I}_{A_i} \mathbb{I}_{A_j}] = E[\|\sum_{i=1}^n \alpha_i \mathbb{I}_{A_i}\|^2] \geq 0$$

- (b) Define for any two events  $A$  and  $B$ ,  $K_2(A, B) = P(A \cap B) - P(A)P(B)$  Verify that  $K_2$  is positive definite.

$$K_2(A, B) = P(A \cap B) - P(A)P(B) = E[\mathbb{I}_A \mathbb{I}_B] - E[\mathbb{I}_A]E[\mathbb{I}_B] = Cov[\mathbb{I}_A, \mathbb{I}_B]$$

$$\sum_{i=1}^n \sum_{j=1}^n \alpha_i \alpha_j Cov[\mathbb{I}_{A_i}, \mathbb{I}_{A_j}] = Cov[\sum_{i=1}^n \alpha_i \mathbb{I}_{A_i}, \sum_{j=1}^n \alpha_j \mathbb{I}_{A_j}] = Var[\sum_{i=1}^n \alpha_i \mathbb{I}_{A_i}] \geq 0$$

## 2 Kernels and RKHS

1. Define the RKHS over  $\mathbb{R}^d$   $K(x, y) = x^T y + c$  where  $c > 0$ .

(a) What is the RKHS associated with the kernel  $K$ ? no proof is required.

$$\mathcal{H} = \{f : \mathbb{R}^d \mapsto \mathbb{R}; f_{w, w_0}(x) = w^T x + w_0; \quad w \in \mathbb{R}^d, w_0 \in \mathbb{R}\}$$

(b) What is the inner product in this RKHS? no proof required.

$$\langle f_{v, v_0}, f_{w, w_0} \rangle_{\mathcal{H}} = v^T w + \frac{1}{c} v_0 w_0 \Rightarrow \langle f_{v, v_0}, f_{v, v_0} \rangle = \|f_{v, v_0}\|_{\mathcal{H}}^2 = \|v\|^2 + \frac{v_0^2}{c}$$

(c) Verify the reproducing property

$$\mathcal{H} \text{ contains all the functions } k(\cdot, x_i) : t \mapsto k(t, x) = t^T x + c = f_t(x)$$

$$\langle f_{w, w_0}, k(\cdot, x) \rangle = \langle f_{w, w_0}, f_{x, c} \rangle = x^T w + \frac{1}{c} c w_0 = w^T x + w_0 = f_w(x)$$

$$\therefore \langle f, k(\cdot, x) \rangle_{\mathcal{H}} = f(x) \text{ for each } f \in \mathcal{H}, x \in \mathcal{X}$$

2. Define the RKHS over  $\mathbb{R}^d$   $K(x, y) = (x^T y)^2$  The RKHS associated with the kernel  $K$  is  $\{f_S; f_S(x) = x^T S x\}$  where  $S$  is a symmetric  $(d, d)$  matrix. The inner product is  $\langle f_{S_1}, f_{S_2} \rangle = \langle S_1, S_2 \rangle_F$

(a) Verify the reproducing property.

$$\mathcal{H} = \{f_S : f_S(x) = x^T S x; \}$$

$\mathcal{H}$  contains all the functions

$$k(\cdot, x_i) : t \mapsto k(t, x) = (t^T x)(t^T x) = x^T \cdot (t t^T) \cdot x = f_t(x)$$

$$\langle f_{S_1}, k(\cdot, x_i) \rangle_{\mathcal{H}} = \langle f_{S_1}, f_{x x^T} \rangle_{\mathcal{H}} = \langle S_1, x x^T \rangle_{\mathcal{F}} = \text{trace}[S_1 x x^T] = \text{trace}[x^T S_1 x] = x^T S_1 x = f_{S_1}(x)$$

Frobenius Norm

$$\therefore \langle f_{S_1}, k(\cdot, x) \rangle_{\mathcal{H}} = f_{S_1}(x) \text{ for each } f \in \mathcal{H}, x \in \mathcal{X}$$

(b) Why do we require that  $S$  is symmetric?

$S_{(d, d)}$  is a symmetric Matrix,

If not, we can not complete the step of  $(t^T x)(t^T x) = x^T \cdot (t t^T) \cdot x$

3. Define the RKHS over  $\mathbb{R}^d$   $K(x, y) = (x^T y + c)^2$  where  $c > 0$ .

(a) What is the RKHS associated with the kernel  $K$ ? no proof is required.

$$\{f_{S, s_0}; f_S(x) = x^T S x + 2s^T x + s_0^2\}$$

where  $S$  is a symmetric  $(d, d)$  matrix

(b) What is the inner product in this RKHS? no proof required.

$$\langle f_{S_1}, f_{S_2} \rangle_{\mathcal{H}} = \langle S_1, S_2 \rangle_{\mathcal{F}} + \frac{s_0^2}{c}$$

(c) Verify the reproducing property

$\mathcal{H}$  contains all the functions

$$k(\cdot, x_i) : t \mapsto k(t, x) = (t^T x + c)(t^T x + c) = x^T \cdot (t t^T) \cdot x + 2c t^T x + c^2 = f_t(x)$$

$$\langle f_S, k(\cdot, x_i) \rangle_{\mathcal{H}} = \langle f_S, f_{x x^T} \rangle_{\mathcal{H}} = \langle S, x x^T \rangle_{\mathcal{F}} = x^T S x + 2s^T x + s_0^2 = f_S(x)$$

### 3 Fisher kernel

Let  $\theta \in \mathbb{R}$  be a parameter and let  $p_\theta$  be a probabilistic model (i.e a point mass function or a density) over a set  $\mathcal{X}$  indexed by  $\theta$ . Let  $\theta_0 \in \mathbb{R}$  be a specific value for  $\theta$ .

Let us define the Fisher score at  $x \in \mathcal{X}$  as  $\phi(x, \theta_0) = \frac{\delta}{\delta\theta} \ln p_\theta(x)$  evaluated at  $\theta = \theta_0$  assuming that this quantity exists.

Define  $I(\theta)$ , the Fisher information associated with the parameter  $\theta$ , i.e.,  $I(\theta) = E[\phi^2(X, \theta)]$  where  $E$  stands for expectation and  $X$  is a random variable with distribution  $p_\theta$ .

The Fisher kernel is then  $k(x, x') = \frac{\phi(x, \theta_0)\phi(x', \theta_0)}{I(\theta_0)}$  where

1. Verify that  $k(., .)$  is a positive definite kernel over  $\mathcal{X}$

$$k(x, x') = \frac{\phi(x, \theta_0)\phi(x', \theta_0)}{I(\theta_0)} = \frac{\phi(x', \theta_0)\phi(x, \theta_0)}{I(\theta_0)} = k(x', x) \text{ symmetric}$$

$$k(x, x') = \frac{1}{I(\theta_0)} \sum_{i=1}^n \alpha_i \phi(x_i, \theta_0) \sum_{j=1}^n \alpha_j \phi(x_j, \theta_0) = \frac{1}{I(\theta_0)} \left\| \sum_{i=1}^n \alpha_i \phi(x_i, \theta_0) \right\|^2 \geq 0$$

2. Consider the following model:  $x \in \{0, 1\}$ ,  $X \sim \text{Bernoulli}(\theta)$ ,  $0 < \theta < 1$ , that is  $p_\theta(x) = \theta^x(1-\theta)^{(1-x)}$   
We recall that in this case  $E[X] = \theta$  and  $\text{Var}[X] = E[(X - \theta)^2] = \theta(1 - \theta)$  Compute  $k(x, x')$

$$p_\theta(x) = \theta^x(1 - \theta)^{(1-x)}$$

$$\ln p_\theta(x) = x \ln \theta + (1 - x) \ln(1 - \theta)$$

$$\frac{d}{d\theta} \ln p_\theta(x) = \frac{x}{\theta} + \frac{1-x}{1-\theta} = \frac{x-\theta}{\theta(1-\theta)}$$

$$I(\theta) = E[\phi^2(X, \theta)] = E\left[\left(\frac{X - \theta}{\theta(1 - \theta)}\right)^2\right] = \frac{E[(X - \theta)^2]}{\theta^2(1 - \theta)^2} = \frac{V[X]}{\theta^2(1 - \theta)^2} = \frac{\theta(1 - \theta)}{\theta^2(1 - \theta)^2} = \frac{1}{\theta(1 - \theta)}$$

$$k(x, x') = \frac{\phi(x, \theta_0)\phi(x', \theta_0)}{I(\theta_0)} = \frac{(x - \theta_0)(x' - \theta_0)}{\theta_0^2(1 - \theta_0)^2} \theta_0(1 - \theta_0) = \frac{(x - \theta_0)(x' - \theta_0)}{\theta_0(1 - \theta_0)}$$

3. Assume now  $x = (x_1, x_2)$  with  $x_1 \in \{0, 1\}$  and  $x_2 \in \{0, 1\}$ . We consider the following model where  $X = (X_1, X_2)$ ,  $X_1$  and  $X_2$  are independent with the same  $\text{Bernoulli}(\theta)$  distribution. Compute  $k(x, x')$ .

$$p_\theta(\vec{x}) = p_\theta(x_1)p_\theta(x_2) = \theta^{x_1+x_2}(1 - \theta)^{2-x_1-x_2}$$

why?

$$\ln p_\theta(x) = (x_1 + x_2) \ln \theta + (2 - x_1 - x_2) \ln(1 - \theta)$$

$$\phi(\vec{x}, \theta) = \frac{d}{d\theta} \ln p_\theta(x) = \frac{x_1 + x_2}{\theta} + \frac{2 - x_1 - x_2}{1 - \theta} = \frac{x_1 + x_2 - 2\theta}{\theta(1 - \theta)}$$

$$\begin{aligned} I(\theta) &= E[\phi^2(\vec{X}, \theta)] = \frac{E[(X_1 + X_2 - 2\theta)^2]}{\theta^2(1 - \theta)^2} \\ &= \frac{E[(X_1 - \theta)^2] + E[(X_2 - \theta)^2] + 2(E[X_1] - \theta)(E[X_2] - \theta)}{\theta^2(1 - \theta)^2} \\ &= \frac{V[X_1] + V[X_2] - 0}{\theta^2(1 - \theta)^2} = \frac{2\theta(1 - \theta)}{\theta^2(1 - \theta)^2} = \frac{2}{\theta(1 - \theta)} \end{aligned}$$

$$\begin{aligned} k(x, x') &= \frac{\phi(\vec{x}, \theta_0)\phi(\vec{x}', \theta_0)}{I(\theta_0)} = \frac{(x_1 + x_2 - 2\theta_0)(x'_1 + x'_2 - 2\theta_0)}{\theta_0^2(1 - \theta_0)^2} \frac{\theta_0(1 - \theta_0)}{2} \\ &= \frac{(x_1 + x_2 - 2\theta_0)(x'_1 + x'_2 - 2\theta_0)}{2\theta_0(1 - \theta_0)} \end{aligned}$$