

проблема !!

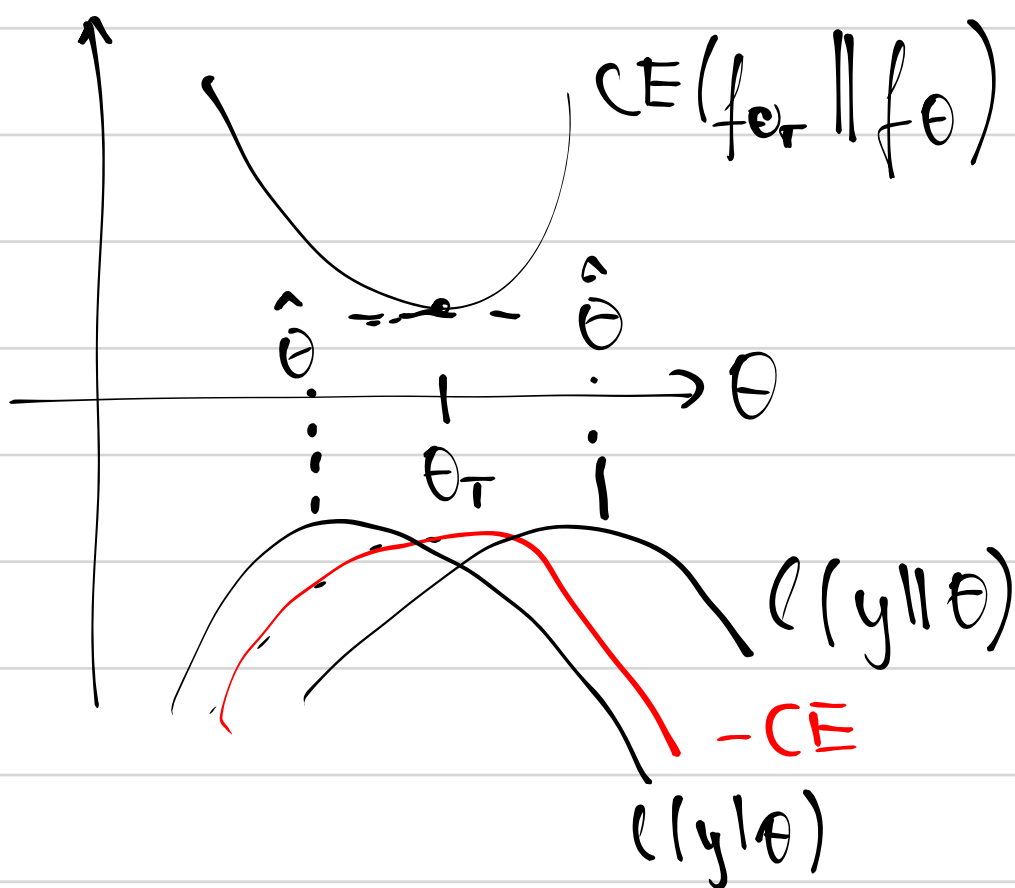
$\xi_{\text{min}}$  [перезарядка],  $\tau$  [характеристическое время];  
 $\hookrightarrow$  pdf

$$\ell(y|\theta) = \ln f(y|\theta), \quad \text{-- лог. ф. правдоподобия.}$$

$$s(y|\theta) = \left( \frac{\partial \ell}{\partial \theta} \right) \quad \text{s-score function}$$

$$\theta \text{ -- параметр} \Rightarrow \eta \text{ -- скаляр}$$

$$\theta = \begin{pmatrix} \theta_1 \\ \vdots \\ \theta_p \end{pmatrix} \Rightarrow \eta = \begin{pmatrix} \partial \ell / \partial \theta_1 \\ \vdots \\ \partial \ell / \partial \theta_p \end{pmatrix}$$



теорема

$$\xi_{\text{min}} [\text{перезарядка}]$$

$$\tau \text{ -- } E_{\theta} (s(y|\theta_{\tau})) = 0$$

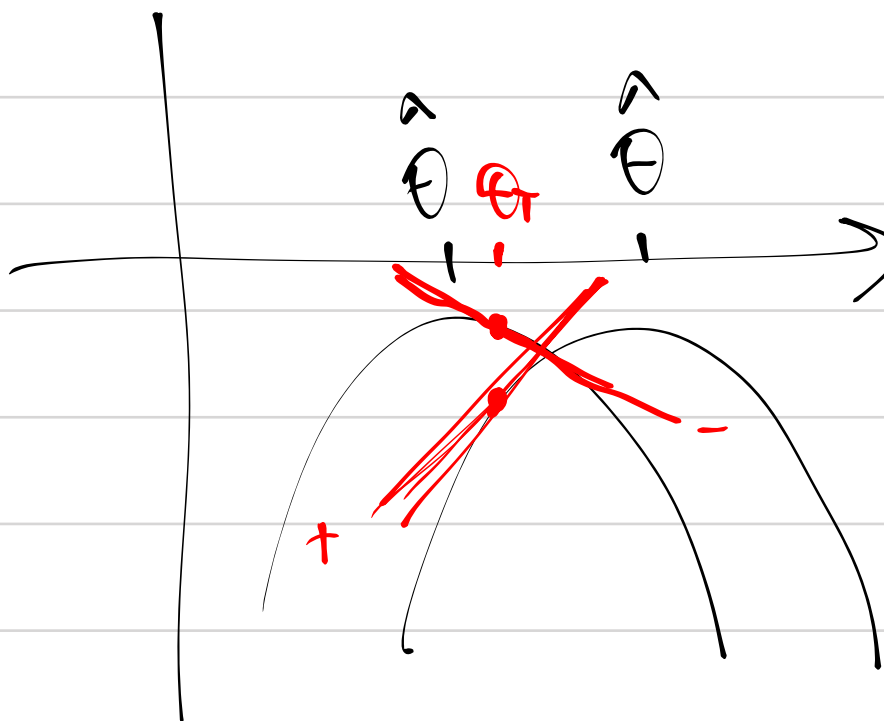
гол-во:

$$\theta_{\tau} \text{ -- минимизирует } CE(f_{\theta_{\tau}} || f_{\theta})$$

$$\frac{\partial CE(f_{\theta_{\tau}} || \theta)}{\partial \theta} \bigg|_{\theta_{\tau}} = 0$$

$$\frac{\partial}{\partial \theta} E_{\theta_{\tau}} (-\ln f(y|\theta)) \bigg|_{\theta_{\tau}} = 0$$

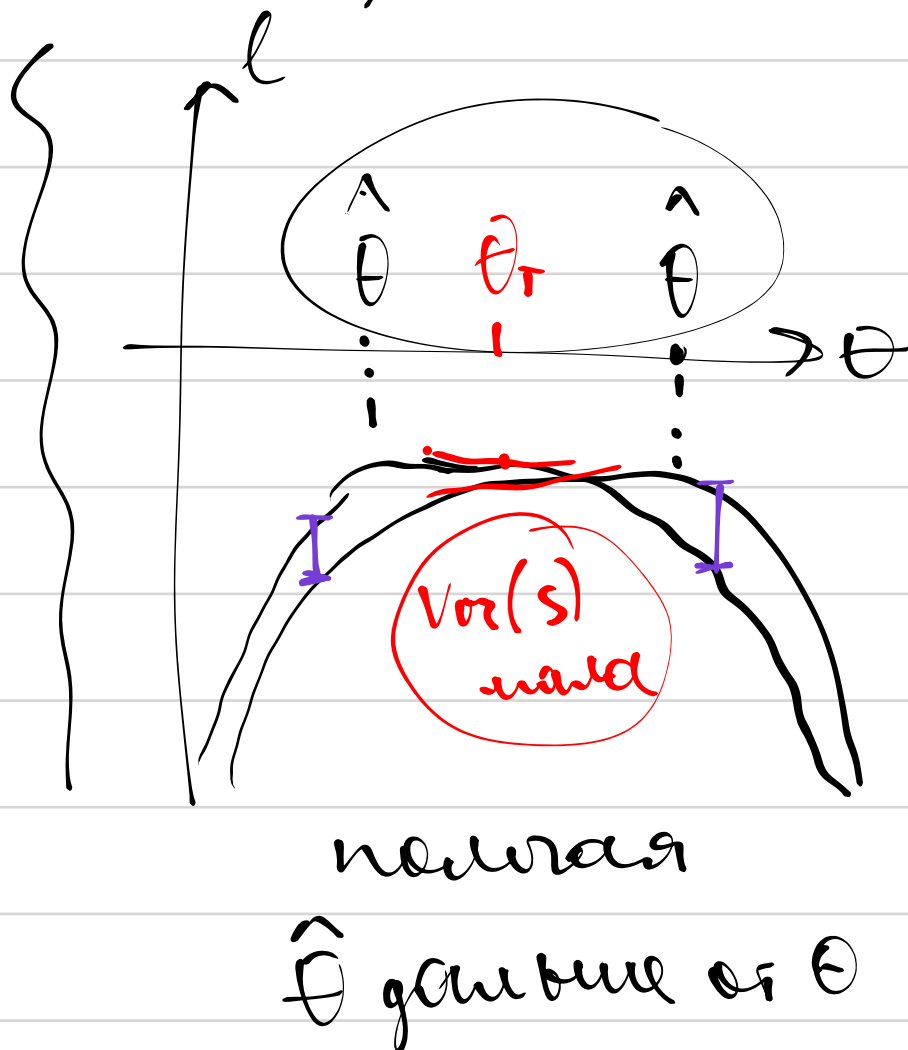
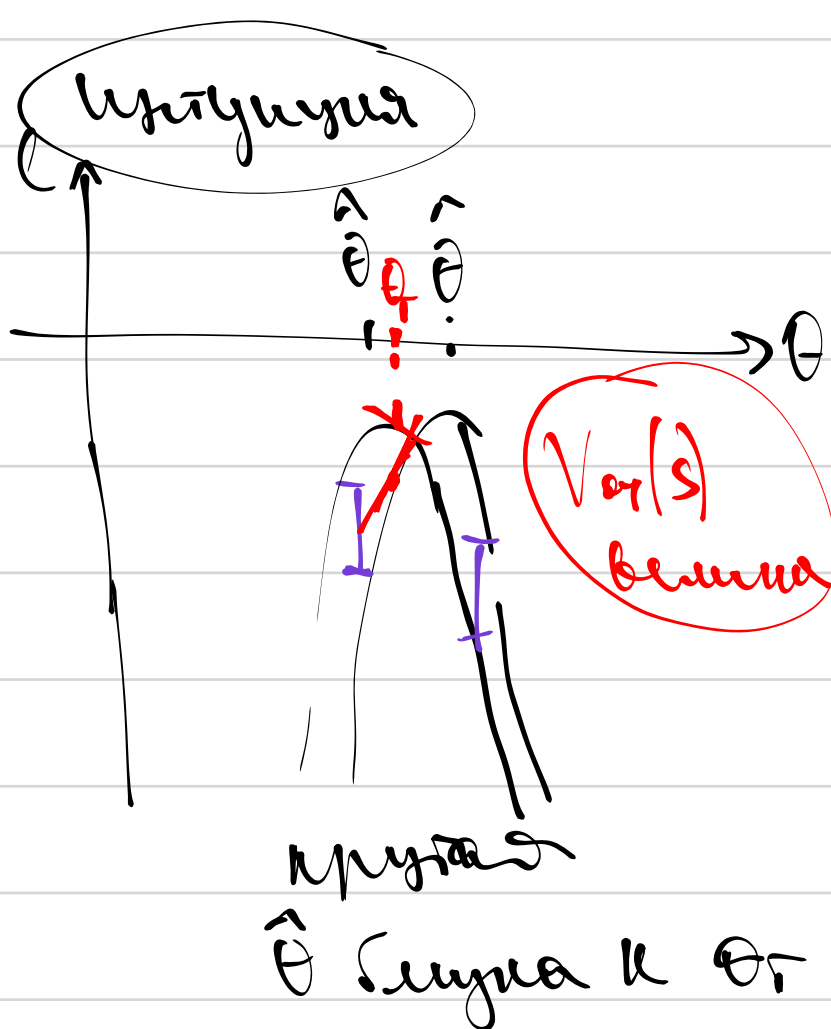
$$E_{\theta_{\tau}} \left( -\frac{\partial \ell}{\partial \theta} \right) \bigg|_{\theta_{\tau}} = 0$$



$$\nabla E(\hat{\theta}) = 0$$

def информационная Рунера. [о крив-м  $\theta$ ]

$$I_F = \text{Var}_\tau(s(y|\theta)) = \text{Var}_\tau\left(\frac{\partial \ell(y|\theta)}{\partial \theta}\right)$$



Теор.

$$I_F = \text{Var}\left(\frac{\partial \ell}{\partial \theta}\right) = E\left(\left(\frac{\partial \ell}{\partial \theta}\right)^2\right) - \left(E\left(\frac{\partial \ell}{\partial \theta}\right)\right)^2 = E\left(\left(\frac{\partial \ell}{\partial \theta}\right)^2\right)$$

Var-variance генерация

E/Var

n/D

$$I_F = -E_{\theta_T}\left(H(y|\theta)\right)$$

$H(y|\theta)$ -м-ча  
тесса для  
 $\ell(y|\theta)$

(м-ча бранных  
м-ч)

\* для вектора

$$I_F = E_{\theta_T}\left[\left(\frac{\partial \ell}{\partial \theta}\right) \cdot \left(\frac{\partial \ell}{\partial \theta}\right)^T\right]$$

слова.

$$\theta = \begin{pmatrix} \theta_1 \\ \vdots \\ \theta_p \end{pmatrix}$$

[ген. бр] по имени

! Имяны:

если  $\theta$ -вектор, то  $(I_F)$

$$\text{Var}(s) =$$

$$\begin{bmatrix} \text{Var}\left(\frac{\partial \ell}{\partial \theta_1}\right) & \text{Cor}\left(\frac{\partial \ell}{\partial \theta_1}, \frac{\partial \ell}{\partial \theta_2}\right) \\ \vdots & \vdots \end{bmatrix}$$

$$y = \begin{pmatrix} y_1 \\ \vdots \\ y_n \end{pmatrix}$$

$$I_n^F(\theta_T)$$

← ссб заб тб от и и от  $\theta_T$

# Теор [Чер-бо Крамера-Рао]

Если  $\hat{\theta}$  - не см-бная оценка для  $\theta$ ,  
 $(E(\hat{\theta}) = \theta)$  и [непрерывная плотность]  
 то [минимум возможных дисперсий для не см-бной оценки]  
 $Var(\hat{\theta}) \geq I_F^{-1}$  ←  $I_F$  — матрица Ф

$$\Delta = I_F \cdot Var(\hat{\theta}) - I$$

инв. матрица

← вект.  $\theta$   
 полн. непрер-на.  
 непрерыв. см-на /  
 identity matrix.

для до:

$$Corr^2(\hat{\theta}, s) \leq 1$$

$$Corr(X, Y) = \frac{Cov(X, Y)}{\sqrt{Var(X)Var(Y)}} \in [-1; 1] \quad \rho(X, Y)$$

$$\frac{1}{Var(\hat{\theta}) \cdot I_F} \leq 1$$

$$Var(\hat{\theta}) \geq \frac{1}{I_F}$$

$\frac{\partial \ell}{\partial \theta}$

$$\frac{Corr^2(\hat{\theta}, s)}{Var(\hat{\theta}) \cdot Var(s)} \leq 1$$

$\Rightarrow I_F$

$$E(\hat{\theta}) = \int_{\mathbb{R}^n} \hat{\theta} f(y|\theta) dy$$

$$Cov(\hat{\theta}, s) = E(\hat{\theta} \cdot s) - \underbrace{E(s) \cdot E(\hat{\theta})}_{=0} = E(\hat{\theta} \cdot s) =$$

$$= \int_{\mathbb{R}^n} \hat{\theta} \cdot \frac{\partial \ell}{\partial \theta} \cdot f(y|\theta) dy =$$

кто тут?  
 б. непрерыв.  
 А.  $y_1, \dots, y_n$

$$= \int_{\mathbb{R}^n} \hat{\theta} \cdot \frac{\partial f(y|\theta)}{\partial \theta} dy \stackrel{①}{=} \frac{\partial}{\partial \theta} \left( \int_{\mathbb{R}^n} \hat{\theta} \cdot f(y|\theta) dy \right) \stackrel{②}{=} \frac{\partial}{\partial \theta} E(\hat{\theta}) \stackrel{③}{=} \frac{\partial \theta}{\partial \theta} = 1$$

① —  $\frac{\partial}{\partial \theta}$   
 ② —  $\frac{\partial}{\partial \theta}$   
 ③ —  $\frac{\partial \theta}{\partial \theta}$

$$y = \begin{pmatrix} y_1 \\ \vdots \\ y_n \end{pmatrix}$$

$$h(y)$$

$$E(h(y)) = \int_{\mathbb{R}^n} h(y) \cdot f(y|\theta) dy$$

вариант:  $\left( \text{Cor}(\hat{\theta}, S) = \dots = 1 \right)$

$$(1) \quad \text{Cor}^2(\hat{\theta}, S) \leq 1$$

$$(2) \quad \frac{\text{Cor}^2(\hat{\theta}, S)}{\text{Var}(\hat{\theta}) \cdot \text{Var}(S)} \leq 1$$

$$(3) \quad \frac{1}{\text{Var}(\hat{\theta}) \cdot \text{Var}(S)} \leq 1$$

$$(4) \quad \text{Var}(\hat{\theta}) \geq \frac{1}{\text{Var}(S)} = \frac{1}{IF}$$

① Когда у нас-то лучше всего  
лучше  $\text{Var}(\hat{\theta})$ ?

$$\text{Cor}^2(S, \hat{\theta}) = 1$$

Менее ст. берем  
 $S(y|\theta)$  и  $\hat{\theta}$  есть  
мин. дисп. ст.  
(асс)

!

sum (params) to

$$\begin{bmatrix} 1/I_n^F \rightarrow 0 \\ \theta \end{bmatrix} \begin{pmatrix} 0 \\ \vdots \\ 0 \end{pmatrix}$$

$\theta$ -best

$$(I_n^F)^{\frac{1}{2}} \cdot (\hat{\theta}_n - \theta) \xrightarrow[n \rightarrow \infty]{\text{dist}} N(0; I)$$

↑ eq. var.

$\theta$ -max

$$(I_n^F)^{\frac{1}{2}} \cdot (\hat{\theta}_n - \theta) \xrightarrow[n \rightarrow \infty]{\text{dist}} N(0; I)$$

$$\theta = \begin{pmatrix} \theta_1 \\ \vdots \\ \theta_p \end{pmatrix}$$

p-params

$$y = \begin{pmatrix} y_1 \\ \vdots \\ y_n \end{pmatrix}$$

$n \rightarrow \infty$

$\hat{\theta}_n$  - asymptotically normal

$\hat{\theta}_n$  - asymptotically efficient

$\hat{\theta}_n$  - asymptotically unbiased

$$\frac{\hat{\theta}_n - \theta}{(1/I_n^F)^{\frac{1}{2}}} \xrightarrow[n \rightarrow \infty]{\text{dist}} N(0; I)$$

$$\frac{\text{Var}(\hat{\theta}_n)}{1/I_n^F} \rightarrow 1$$

т. кр-ра-Рав:

$$\sqrt{\text{Var}(\hat{\theta}_n)} \geq \sqrt{1/I_n^F}$$

т. кр-ра-Рав

эффективность

$$\hat{\theta}_A, \hat{\theta}_B$$

def

$\hat{\theta}_{\text{best}}$  - best estimator among  $\{\hat{\theta}_1, \hat{\theta}_2, \dots\}$   
 $E[(\hat{\theta}_{\text{best}} - \theta)^2] \leq E[(\hat{\theta} - \theta)^2]$   
 for each  $\hat{\theta}$   
 из заданного набора.

т. Параллель:

$$E[(\hat{\theta} - \theta)^2] = \text{Var}(\hat{\theta}) + (E(\hat{\theta}) - \theta)^2$$

елин = 0



Теорема ④

эсм (перу)

$$\underbrace{I_n^F(\theta_T)}_{\substack{\text{зависит от} \\ \text{от перу } \theta_T}}$$

тут важно  
была  $I^F$

$$\sqrt{\hat{I}_n^F} \cdot (\hat{\theta}_n^{ML} - \theta) \xrightarrow[n \rightarrow \infty]{\text{dist}} N(0;1)$$

как вычисл  $I^F$ ?

① способ А:

$$I^F = -E\left(\frac{\partial^2 \ell}{\partial \theta^2}\right)$$

$$\hat{I}^F = -\frac{\partial^2 \ell}{\partial \theta^2} \Big|_{\hat{\theta}_{ML}}$$

способ Б:

$$\hat{I}^F = I_n^F(\hat{\theta})$$

## Теорема

перу  $f(y|\theta)$

$$\frac{\partial^2 \ell(y|\theta)}{\partial \theta^2} \Big|_{\hat{\theta}_n^{ML}}$$

$se(\hat{\theta})$  - оценка  
для  $\sqrt{Var(\hat{\theta})}$

Ува 1.  $\hat{\theta}_{ML}$  - точка

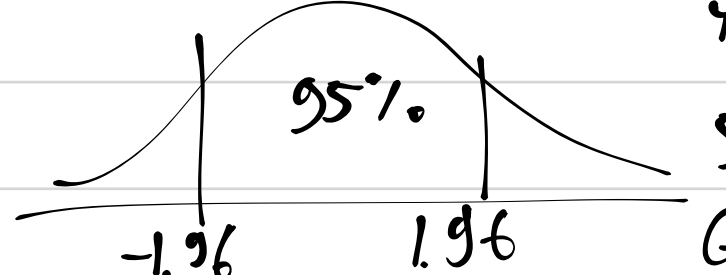
Ува 2.  $\hat{I}_F = -\frac{\partial^2 \ell}{\partial \theta^2} \Big|_{\hat{\theta}_{ML}}$

Ува 3.  $Var(\hat{\theta}_{ML}) = 1/\hat{I}_F$

$$se(\hat{\theta}_{ML}) = \sqrt{Var(\hat{\theta}_{ML})}$$

$se$  - standard  
error

$$\frac{\hat{\theta} - \theta}{se(\hat{\theta})} \xrightarrow{\text{dist}} N(0;1)$$



CI: Ува 4.

$$-1.96 \leq \frac{\hat{\theta} - \theta}{se(\hat{\theta})} \leq 1.96$$

$$\left[ \hat{\theta} - 1.96 \cdot se(\hat{\theta}) \leq \theta \leq \hat{\theta} + 1.96 \cdot se(\hat{\theta}) \right]$$

ува 5.

bootstrap не подходит -1.96  
и 1.96

