STAT 430-2 2025 Winter HW3

Tuorui Peng¹

Exercise 1 5.3.3

We prove the following: $(a) \Rightarrow (b)$, $(b) \Rightarrow (e)$, $(d) \Rightarrow (e)$, $(c) \Rightarrow$ Doob's convergence. Combined with the original condition in Doob's convergence that: $(e) \Rightarrow$ Doob's convergence, we have the equivalence of all conditions.

 $(a) \Rightarrow (b)$ If $\lim_n \mathbb{E}[|X_n|] = M$, then fix any $\varepsilon > 0$, there $\exists N_0$, such that $\forall n > N_0$, $\mathbb{E}[|X_n|] < M + \varepsilon$. Then

$$\sup_{n} \mathbb{E}\left[|X_{n}|\right] \leq \sup_{k=1}^{N_{0}} \mathbb{E}\left[|X_{k}|\right] \vee (M+\varepsilon) < \infty$$

Thus (b) holds.

 $(b) \Rightarrow (e)$ Note that $|x| = x_+ + x_-$, we thus have $\forall n$:

$$\mathbb{E}[|X_n|] = \mathbb{E}[(X_n)_+] + \mathbb{E}[(X_n)_-]$$
$$\geq \mathbb{E}[(X_n)_+]$$

thus $\sup_n \mathbb{E}[|X_n|] < \infty$ would yield $\sup_n \mathbb{E}[(X_n)_+] < \infty$, which is (e).

- $(d) \Rightarrow (e)$ Idea is the same as $(a) \Rightarrow (b)$ by just replacing $|X_n|$ with $(X_n)_+$.
- $(c) \Rightarrow \text{Doob}$ In the prove of Doob's convergence, we have in Doob's up-crossing:

$$\mathbb{E}[U_n[a,b]] \le \frac{1}{b-a} \mathbb{E}[(X_n-a)_+] \le \frac{1}{(b-a)} (|a| + \mathbb{E}[(X_n)_+])$$

then applying Fatou's lemma, we have:

$$\mathbb{E}\left[\liminf_{n} U_{n}[a,b]\right] \leq \frac{1}{(b-a)} \left(|a| + \liminf_{n} \mathbb{E}\left[(X_{n})_{+}\right]\right) < \infty$$

then the rest of the proof is the same as Doob's convergence, thus (c) implies Doob's convergence.

Exercise 2 5.3.6

We verify that $X_n := W_n + an - (K + a)N_{n-1}$ is a sup-MG as follows:

$$\mathbb{E}\left[X_{n+1} - X_n | \mathcal{F}_n\right] = \mathbb{E}\left[W_{n+1} + a(n+1) - (K+a)N_n - W_n - an + (K+a)N_{n-1} | \mathcal{F}_n\right]$$

$$= \begin{cases} \text{if } W_n \ge b, & \mathbb{E}\left[W_{n+1} - W_n + a | \mathcal{F}_n\right] \not \Vdash_{W_n \ge b} \le 0\\ \text{if } W_n < b, & \mathbb{E}\left[W_{n+1} - W_n - K | \mathcal{F}_n\right] \le 0 \end{cases}$$

 $^{^1}$ Tuorui Peng
2028@u.northwestern.edu

and finite expectation is easily verified. Thus X_n is a supermartingale. Further we notice that X_n have uniformly bounded difference 2(K+a), thus by (prop 5.3.5) we have that

$$\mathbb{P}\left(\lim_{n} X_{n} \text{ exists } \bigcup \liminf_{n} X_{n} = -\infty\right) = 1$$

Now consider the event that $\{N_{\infty} = \text{some } \nu < \infty\}$ We note that in this case, we have for large n:

$$\liminf_{n} X_n = \liminf_{n} W_n + an - (K+a)\nu \to \infty$$

since W_n is non-negative. Thus

$$\mathbb{P}\left(N_{\infty} < \infty\right) \leq \mathbb{P}\left(\liminf_{n} X_{n} = \infty\right) = \mathbb{P}\left(\left(\lim_{n} X_{n} \text{ exists } \bigcup \liminf_{n} X_{n} = -\infty\right)^{\complement}\right) = 0$$

Exercise 3 5.3.10

We verify that $W_n := \frac{X_n + 1}{\prod_{k=1}^{n-1} (Y_k + 1)}$ is a supermartingale as follows:

$$\mathbb{E}\left[W_{n+1} - W_n | \mathcal{F}_n\right] = \mathbb{E}\left[\frac{X_{n+1} + 1}{\prod_{k=1}^n (Y_k + 1)} - \frac{X_n + 1}{\prod_{k=1}^{n-1} (Y_k + 1)} | \mathcal{F}_n\right]$$

$$= \frac{1}{\prod_{k=1}^n (Y_k + 1)} \mathbb{E}\left[X_{n+1} + 1 - (X_n + 1)(Y_n + 1) | \mathcal{F}_n\right] \le 0$$

in which some technical conditions are: $\log \prod_{k=1}^{\infty} (Y_k + 1) = \sum_{k=1}^{\infty} \log(Y_k + 1) \leq \sum_{k=1}^{\infty} Y_k < \infty$, and $\mathbb{E}[|W_n|] \leq \mathbb{E}[|X_n|] < \infty$. Thus we have that W_n is a (non-negative) supermartingalem. Thus directly we have that $\lim_n W_n$ exists a.s., denote it as W_∞ .

Then note that for each $\omega \in \Omega$:

$$\prod_{k=1}^{n-1} (Y_k(\omega) + 1) < \infty, \quad \text{a.s.}$$

and is an increasing sequence, thus it has a limit $Z(\omega)$. Combine the fact that $W_n \xrightarrow{\text{a.s.}} W_{\infty}$, we have

$$X_n \xrightarrow{\text{a.s.}} W_{\infty} Z - 1$$

Exercise 4 5.3.21

Note that $\mathcal{F}_k^A := \sigma(\mathcal{A}, Z_0, \dots, Z_{k-1})$ and $\mathcal{F}_k^B := \sigma(\mathcal{B}, Y_0, \dots, Y_{k-1})$ are both filtrations in \mathcal{F} , and also we have that

$$Y_k = \mathbb{E}\left[X|\mathcal{F}_k^A\right]$$
 is U.I. M.G. $Z_k = \mathbb{E}\left[X|\mathcal{F}_k^B\right]$ is U.I. M.G.

which are easy to verify. Thus we have by Theorem 5.3.12 that there exists some Y_{∞}, Z_{∞} s.t.

$$Y_k = \mathbb{E}\left[X|\mathcal{F}_k^A\right] \xrightarrow[L_1]{\text{a.s.}} Y_{\infty}$$
$$Z_k = \mathbb{E}\left[X|\mathcal{F}_k^B\right] \xrightarrow[L_1]{\text{a.s.}} Z_{\infty}$$

Now follow similar idea as Theorem 5.3.15, we would obtain that $Y_{\infty} \in m\mathcal{F}_{\infty}^{A}$ and then $Y_{\infty} = \mathbb{E}\left[X|\mathcal{F}_{\infty}^{A}\right]$ where $\mathcal{F}_{\infty}^{A} = \sigma(\mathcal{A}, Z_{0}, \ldots)$; similarly we have $Z_{\infty} = \mathbb{E}\left[X|\mathcal{F}_{\infty}^{B}\right]$ where $\mathcal{F}_{\infty}^{B} = \sigma(\mathcal{B}, Y_{0}, \ldots)$.

Then we have:

$$\begin{aligned} Y_{\infty} = & \mathbb{E}\left[Y_{\infty}|\mathcal{F}_{\infty}^{B}\right] \\ = & \mathbb{E}\left[Y_{\infty}|\mathcal{F}_{\infty}^{A}\right] = \mathbb{E}\left[X|\mathcal{F}_{\infty}^{A}\right] \end{aligned}$$

where the first line since $Y_k \xrightarrow{\text{a.s.}} Y_{\infty}$ and the second line since $Y_{\infty} \in \mathcal{F}_{\infty}^A$. Together we have

$$Y_{\infty} = \mathbb{E}\left[Y_{\infty}|\mathcal{F}_{\infty}^{B}\right] = \mathbb{E}\left[\mathbb{E}\left[Y_{\infty}|\mathcal{F}_{\infty}^{A}\right]|\mathcal{F}_{\infty}^{B}\right] = \mathbb{E}\left[X|\sigma(\mathcal{F}_{\infty}^{A},\mathcal{F}_{\infty}^{B})\right] = Z_{\infty}$$

thus we must have $Y_{\infty} = Z_{\infty}$ a.s., then apply the U.I. property again we have $Y_n - Z_n \xrightarrow[L_1]{\text{a.s.}} 0$.

Exercise 5 5.3.24

5.(a)

We first prove for $W \in L^2\mathcal{F}_{n_W}$ fixed then we have for $n > n_W$ that

$$\mathbb{E}\left[WZ_{n}\right] = \mathbb{E}\left[\mathbb{E}\left[WZ_{n}|\mathcal{F}_{n_{W}}\right]\right]$$

$$= \mathbb{E}\left[\mathbb{E}\left[\frac{W}{\sqrt{n}}\sum_{k=1}^{n_{W}}\xi_{k}|\mathcal{F}_{n_{W}}\right] + \frac{\mathbb{E}\left[W|\mathcal{F}_{n_{W}}\right]}{\sqrt{n}}\sum_{k=n_{W}+1}^{\infty}\xi_{k}\right] \to 0$$

since the first term decay to 0 as $n \to \infty$, and the second term simply has mean 0.

Then we upgrade to $W \in L^2 \mathcal{F}_{\infty}$. Define $W_n := \mathbb{E}[W|\mathcal{F}_n]$, in this way we can write that $WZ_n = W_m Z_n + (W - W_m) Z_n$. Then we consider the two parts:

Part 1: For any fixed m we have by preceding that

$$\mathbb{E}\left[W_m Z_n\right] \to 0, \quad (n \to \infty)$$

Part 2: We have

$$\limsup_{n \to \infty} \mathbb{E}\left[\left|(W - W_m)Z_n\right|\right] \le \left(\lim_{n \to \infty} \sqrt{\mathbb{E}\left[\left|Z_n\right|^2\right] \cdot \mathbb{E}\left[\left|W - W_m\right|^2\right]}\right)$$
$$= \operatorname{const} \cdot \mathbb{E}\left[\left|W - W_m\right|^2\right]$$

since by CLT, we would have $Z_n \xrightarrow{d} N(0,1)$.

Together we have for any $m \in \mathbb{N}^+$,

$$\limsup_{n \to \infty} \mathbb{E}\left[|WZ_n|\right] \le \operatorname{const} \cdot \mathbb{E}\left[|W - W_m|^2\right]$$

further we notice that by the definition of W_m and using Doob's L_2 convergence we have $W_m \xrightarrow{L_2} W$ thus $\mathbb{E}\left[|W-W_m|^2\right] \to 0$, thus we have $\limsup_{n\to\infty} \mathbb{E}\left[|WZ_n|\right] = 0$.

5.(b)

Now for $Z \in \mathcal{F}$, we know that the above convergence works for $\mathbb{E}[Z|\mathcal{F}_{\infty}] \in \mathcal{F}_{\infty}$. Then we have

$$\mathbb{E}\left[WZ_n\right] = \mathbb{E}\left[\mathbb{E}\left[WZ_n|\mathcal{F}_{\infty}\right]\right] = \mathbb{E}\left[Z_n\mathbb{E}\left[W|\mathcal{F}_{\infty}\right]\right] \to 0$$

Now if there exists some $\tilde{Z} \in \mathcal{F}$ s.t. $Z \xrightarrow{L_2} \tilde{Z}$, which means that

$$\mathbb{E}\left[\left|Z_n - \tilde{Z}\right|^2\right] \to 0$$

we would see that

L.H.S.
$$=\mathbb{E}\left[Z_n^2\right] + \mathbb{E}\left[\tilde{Z}^2\right] - 2\mathbb{E}\left[Z_n\tilde{Z}\right] \to 0$$

however we have $\liminf_n \mathbb{E}\left[Z_n^2\right] \geq \mathbb{E}\left[\liminf_n Z_n^2\right] > 0$ by CLT, and that $\mathbb{E}\left[Z_n\tilde{Z}\right] \to 0$ by the above argument, thus we have a contradiction. Thus Z does not converge in L_2 .

Exercise 6 5.4.7

We first prove (5.4.6): Since max is a convex function, we have that $Y_n := \max\{X_n, -1\}$ is still a sub-M.G. Now assume that $\mathbb{E}\left[\sup |Y_n|\right] < \infty$ i.e. Y_n is integrable. We consider the stopping time defined as:

$$\tau := \inf\{n : Y_n < 0\}$$

and since $\inf_n X_n < 0$ a.s., we have $\tau < \infty$ a.s., for which $Y_\tau < 0$, and thus we further have by Doob's optional stopping theorem that

$$0 > \mathbb{E}\left[Y_{\tau}\right] \ge \mathbb{E}\left[Y_{0}\right] \ge 0$$

which is a contradiction. Thus we have that $\mathbb{E}[\sup |Y_n|] = \infty$. Now since $\sup |Y_n| = \max\{\sup Y_+, \sup Y_-\}$ while $\sup Y_- \le 1$, we have

$$\infty = \mathbb{E} \left[\sup |Y_n| \right] \le \mathbb{E} \left[\max \{ \sup \max \{X_n, -1\}_+, 1 \} \right]$$
$$\le \mathbb{E} \left[\max \{1, \sup X_n \} \right]$$

for this to hold, we must have $\mathbb{E}[\sup X_n] = \infty$. Thus we have proved (5.4.6).

Now using the lemma, we solve (5.4.7).

6.(a)

For random walk, we know that we have $S_n/\sqrt{n} \stackrel{\mathrm{d}}{\to} N(0,1)$ which is a non-degenerate distribution. Thus we have that $\mathbb{P}\left(\lim_n S_n \text{ exists}\right) = 0$, because for the event $\{\lim_n S_n \text{ exists}\}$, we must have that $S_n(\omega)/\sqrt{n} \to 0$.

For such M.G. with bounded difference, by (prop 5.3.5) we have that

$$\lim_{n} S_n$$
 exists , $\operatorname{or} \liminf_{n} S_n = -\infty$, $\lim_{n} \sup_{n} S_n = \infty$, a.s.

and from the above argument we have that the first case is w.p. 0, thus we have that $\limsup_n S_n = \infty$ & $\liminf_n S_n = -\infty$ a.s., which means that $\tau_b < \infty$ a.s.

6.(b)

Note that we have $S_{\tau_b} \geq b > 0$, i.e. $\sup S_{n \wedge \tau_b} > 0$. And since τ_b is a stopping time, we also have that $S_{n \wedge b}$ is a M.G. (with $S_{0 \wedge \tau_b} = 0$), thus by the lemma we have that $\mathbb{E}\left[\min\{S_n : n \leq \tau_b\}\right] = \mathbb{E}\left[\inf_n S_{n \wedge \tau_b}\right] = \infty$.

Exercise 7 5.4.10

7.(a)

First we have that

$$\mathbb{E}\left[S_{\tau}\right] = \mathbb{E}\left[\sum_{k=1}^{\infty} \xi_{k} I_{k \leq \tau}\right]$$

$$\leq \mathbb{E}\left[\sum_{k=1}^{\infty} \left|\xi_{k}\right|_{k} I_{k \leq \tau}\right]$$

$$\stackrel{\text{monotone}}{=} \sum_{k=1}^{\infty} \mathbb{E}\left[\left|\xi_{k}\right|\right] \mathbb{P}\left(k \leq \tau\right)$$

$$= \mathbb{E}\left[\left|\xi_{1}\right|\right] \sum_{k=1}^{\infty} \mathbb{P}\left(k \leq \tau\right) = \mathbb{E}\left[\left|\xi_{1}\right|\right] \mathbb{E}\left[\tau\right] < \infty$$

Then by DCT, we can apply the above argument to S_{τ} directly and obtain that $\mathbb{E}[S_{\tau}] = \mathbb{E}[\xi_1] \mathbb{E}[\tau]$.

7.(b)

Assume $\mathbb{E}\left[\xi_i\right] = 0$ (otherwise we can replace ξ_i with $\xi_i - \mathbb{E}\left[\xi_i\right]$), which would not influence the result. Now it suffice to prove $\mathbb{E}\left[S_{\tau}^2\right] = var(\xi_1)\mathbb{E}\left[\tau\right]$. To do so we consider $S_{n \wedge \tau}$ for which we have

$$S_{n \wedge \tau} = \sum_{k=1}^{n} \xi_k I_{k \le \tau}$$

and we have

$$\mathbb{E}\left[S_{n\wedge\tau}^2\right] = \mathbb{E}\left[\sum_{k=1}^n \xi_k^2 I_{k\leq\tau}\right] = var(\xi_1) \sum_{k=1}^n \mathbb{P}\left(k\leq\tau\right) = var(\xi_1) \mathbb{E}\left[\tau\right] < \infty$$

for which right hand side does not depend on n. Thus here we can apply Doob's L_p convergence theorem (p=2) to obtain

$$\mathbb{E}\left[S_{\tau}^{2}\right] = \mathbb{E}\left[\lim_{n} S_{n \wedge \tau}^{2}\right] = var(\xi_{1})\mathbb{E}\left[\tau\right]$$

7.(c)

If $\xi \geq 0$, then we have

$$\mathbb{E}\left[S_{\tau}\right] = \mathbb{E}\left[\sum_{k=1}^{\infty} \xi_{k} I_{k \leq \tau}\right]$$

$$\stackrel{\text{monotone}}{=} \sum_{k=1}^{\infty} \mathbb{E}\left[\xi_{k}\right] \mathbb{P}\left(k \leq \tau\right)$$

$$= \mathbb{E}\left[\xi_{1}\right] \sum_{k=1}^{\infty} \mathbb{P}\left(k \leq \tau\right) = \mathbb{E}\left[\xi_{1}\right] \mathbb{E}\left[\tau\right]$$

The proof of Wald's second is still the same.

Exercise 8 5.4.11

8.(a)

In 5.4.7 we already proved it (by noting that SRW is also a M.G. with bounded difference).

8.(b)

By the similar argument as in (coroll 5.4.8) in which we use the M.G. $M_n := \exp(\lambda S_n)/M(\lambda)^n$ we have

$$S_{\tau_{a,b}} = \begin{cases} -a, & w.p. \frac{e^{\lambda b} - 1}{e^{\lambda b - e^{-\lambda a}}} \\ b, & w.p. \frac{1 - e^{-\lambda a}}{e^{\lambda b - e^{-\lambda a}}} \end{cases}$$

where $\lambda = \log \frac{1-p}{p}$. Thus we have

$$\mathbb{E}\left[S_{\tau_{a,b}}\right] = -a \cdot \frac{e^{\lambda b} - 1}{e^{\lambda b} - e^{-\lambda a}} + b \cdot \frac{1 - e^{-\lambda a}}{e^{\lambda b} - e^{-\lambda a}}$$
$$= \mathbb{E}\left[\xi_1\right] \mathbb{E}\left[\tau_{a,b}\right]$$
$$= (2p - 1) \mathbb{E}\left[\tau_{a,b}\right]$$

and we have

$$\mathbb{E}\left[\tau_{a,b}\right] = \frac{1}{2p-1} \left(-a \cdot \frac{e^{\lambda b} - 1}{e^{\lambda b} - e^{-\lambda a}} + b \cdot \frac{1 - e^{-\lambda a}}{e^{\lambda b} - e^{-\lambda a}}\right)$$

where $\lambda = \log \frac{1-p}{p}$.

And for the case that p=1/2, i.e. $\lambda=0$, we use Wald's second to obtain that

$$\mathbb{E} \left[\tau_{a,b} \right] = \mathbb{E} \left[S_{\tau}^{2} \right] / var(\xi_{1})$$

$$= \mathbb{E} \left[S_{\tau}^{2} \right]$$

$$= ab$$

8.(c)

We have

$$\tau_{a,b} = \inf\{n : S_n = -a\} \land \inf\{n : S_n = b\} \underset{a \to \infty}{\longrightarrow} \inf\{n : S_n = b\} = \tau_b$$

(and the convergence is monotone, which is trivial). We thus have by Wald's first that

$$\mathbb{E}\left[\tau_{b}\right] = \mathbb{E}\left[S_{\tau_{b}}\right] / \mathbb{E}\left[\xi_{1}\right] = b/(2p-1)$$

8.(d)

For $p \ge 1/2$, it's a submartingale, thus by 5.4.7 we have $\tau_b < \infty$ a.s..

8.(e)

$$\mathbb{E}\left[Y_{n+1}|\mathcal{F}_n\right] = \left(S_n^4 + 6S_n^2 + 1\right) - 6(n+1)\left(S_n^2 + 1\right) + c_1(n^2 + 2n + 1) + c_2(n+1)$$
$$= S_n^4 - 6nS_n^2 + c_1n^2 + c_2n + \left(1 - 6(n+1) + c_1(2n+1) + c_2\right)$$

to make it a M.G., we have $c_1 = 3, c_2 = 2$, thus we have

$$Y_n := S_n^4 - 6nS_n^2 + 3n^2 + 2n$$

is a M.G.

For which we have

$$\mathbb{E}\left[Y_{\tau_{b,b}}\right] = \mathbb{E}\left[\mathbb{E}\left[Y_{k}|\tau_{b,b} = k\right]\right]$$

$$= b^{4} - 6b^{2}\mathbb{E}\left[\tau_{b,b}\right] + 3\mathbb{E}\left[\tau_{b,b}^{2}\right] + 2\mathbb{E}\left[\tau_{b,b}\right]$$

$$= b^{4} - 6b^{2} \cdot b^{2} + 3\mathbb{E}\left[\tau_{b,b}^{2}\right] + 2b^{2}$$

$$= \mathbb{E}\left[Y_{0}\right] = 0$$

where the last step is by Doob's optional stopping theorem. Now we have

$$\mathbb{E}\left[\tau_{b,b}^2\right] = \frac{5b^4 - 2b^2}{3}$$

Exercise 9 5.4.12

9.(a)

Since $M_n = \exp[\lambda S_n] / \mathbb{E}[\exp[\lambda n \xi_1]]$ is a M.G., then by Doob's optional stopping theorem we have

$$1 = \mathbb{E} [M_0]$$

$$= \mathbb{E} [M_{\tau_b}]$$

$$= \mathbb{E} [\exp [\lambda S_{\tau_b}] / \mathbb{E} [\exp [\lambda \xi_1]]^{\tau_b}]$$

$$= e^{\lambda b} \mathbb{E} [M(\lambda)^{-\tau_b}]$$

9.(b)

Denote $s = M(\lambda)^{-1} \in (0,1)$ for some λ , and use b = 1 in the above result then we have

$$\mathbb{E}\left[s^{\tau_1}\right] = e^{-\lambda}$$

and that

$$\mathbb{E}\left[\tau_{b}\right] = e^{-\lambda b} = \mathbb{E}\left[s^{\tau_{1}}\right]^{b}$$

Now we revisit $s=M(\lambda)^{-1}=\frac{1}{pe^{\lambda}+qe^{-\lambda}},$ and we have

$$e^{-\lambda} = \frac{1}{2qs} \frac{2qe^{-\lambda}}{pe^{\lambda} + qe^{-\lambda}}$$

$$= \frac{1}{2qs} \left(1 - \frac{pe^{\lambda} - qe^{-\lambda}}{pe^{\lambda} + qe^{-\lambda}} \right)$$

$$= \frac{1}{2qs} \left(1 - \sqrt{\frac{(pe^{\lambda} - qe^{-\lambda})^2}{(pe^{\lambda} + qe^{-\lambda})^2}} \right)$$

$$= \frac{1}{2qs} \left(1 - \sqrt{1 - 4\frac{pe^{\lambda} \cdot qe^{-\lambda}}{pe^{\lambda} + qe^{-\lambda}}} \right)$$

$$= \frac{1}{2qs} \left(1 - \sqrt{1 - 4pqs^2} \right)$$

substitute this back to the previous result we have

$$\mathbb{E}\left[s^{\tau_1}\right] = e^{-\lambda} = \frac{1}{2as} \left(1 - \sqrt{1 - 4pqs^2}\right), \quad s = M(\lambda)^{-1} \in (0, 1), \quad p \ge 1/2$$

9.(c)

In 5.4.11 (c) we had that $\tau_{a,b} \stackrel{a \to \infty}{\to} \tau_b$ a.s.. Thus we have

$$\mathbb{P}\left(S_{\tau_{a,b}} = b\right) \to \mathbb{P}\left(S_{\tau_b} = b\right) = \mathbb{P}\left(\tau_b < \infty\right) = \frac{1}{e^{\lambda_{\star}b}} = \exp\left[-\lambda_{\star}b\right]$$

where $\lambda_{\star} = \log \frac{1-p}{p}$.

9.(d)

Continuing from the previous result, we have

$$\mathbb{P}\left(S_{\tau_b} = b\right) = \exp\left[-\lambda_{\star} b\right], \quad b \in \mathbb{N}^+$$

which gives us for $Z := 1 + \max_{k>0} S_k$ that:

$$\mathbb{P}(Z = b + 1) = \mathbb{P}\left(\max_{k \ge 0} S_k = b\right)$$

$$= \mathbb{P}\left(\tau_b < \infty, \tau_{b+1} = \infty\right)$$

$$= \mathbb{P}\left(\tau_b < \infty, \text{ after reaching } b, \text{ never reach } b + 1\right)$$

$$= \mathbb{P}\left(\tau_b < \infty\right) \left(1 - \mathbb{P}\left(\tau_1 = \infty\right)\right)$$

$$= \exp\left[-\lambda_{\star}b\right] \left(1 - \exp\left[-\lambda_{\star}\right]\right)$$

$$\sim \text{Geo}(1 - \exp\left[-\lambda_{\star}\right])$$

Exercise 10 5.4.14

Denote $26 = \kappa$

10.(a)

We have

$$\mathbb{E}\left[M_{n+1}|\mathcal{F}_n\right] = \frac{1}{\kappa}M_n \times \kappa + \left(1 - \frac{1}{\kappa}\right) \times 0 = M_n$$

thus M_n is a martingale.

The construction of total gain would look like:

$$M_n = \sum_{n} \left(\kappa^{10} (\kappa \mathbf{1}_{PROBABILITY} - 1) + \kappa^9 (\kappa \mathbf{1}_{PROBABILIT} - 1) + \dots + \kappa^1 (\kappa \mathbf{1}_{PR} - 1) + \kappa^0 (\kappa \mathbf{1}_{P} - 1) \right)$$

10.(b)

We see that $\hat{\tau}$ is equivalent to that $\hat{\tau} := \inf\{n \geq 11 : \text{gain of a gamblers } = \kappa^{11}\}$ (i.e. when the gambler win 11 consercutive games, he gets the word 'PROBABILITY'). WHEN The gambler wins 11 consercutive games, there are other $\hat{\tau} - 1$ gamblers losing. And as we mentioned earlier, we have M_n being a martingale, thus by Doob's optional stopping theorem we have

$$0 = \mathbb{E}[M_0] = \mathbb{E}[M_{\hat{\tau}}] = \mathbb{E}[\mathbb{E}[M_k | \tau = k]] = \sum_{k=0}^{10} (\kappa^{k+1} - \kappa^k) - (\mathbb{E}[\hat{\tau}] - 1)$$

And it gives that

$$\mathbb{E}\left[\hat{\tau}\right] = \kappa^{11} = 26^{11}$$

For 'ABRACADABRA' = 'ABRA' + 'CAD' + 'ABRA', we have that

$$0 = \mathbb{E}[M_0] = \sum_{k=0}^{10} (\kappa^{k+1} - \kappa^k) + \sum_{k=0}^{3} (\kappa^{k+1} - \kappa^k) - (\mathbb{E}[\tilde{\tau}] - 2) = \kappa^{11} + \kappa^4 - \mathbb{E}[\tilde{\tau}]$$

$$\Rightarrow \mathbb{E}[\tilde{\tau}] = \kappa^{11} + \kappa^4 = 26^{11} + 26^4$$

where $+\kappa^4$ for that, when reaching 'ABRACADABRA', there is also another gambler who was winning at the first 4 letters 'ABRA'.