

# STAT 430-2 2025 Winter

## HW3

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### Exercise 1 5.3.3

We prove the following:  $(a) \Rightarrow (b)$ ,  $(b) \Rightarrow (e)$ ,  $(d) \Rightarrow (e)$ ,  $(c) \Rightarrow$  Doob's convergence. Combined with the original condition in Doob's convergence that:  $(e) \Rightarrow$  Doob's convergence, we have the equivalence of all conditions.

$(a) \Rightarrow (b)$  If  $\lim_n \mathbb{E}[|X_n|] = M$ , then fix any  $\varepsilon > 0$ , there  $\exists N_0$ , such that  $\forall n > N_0$ ,  $\mathbb{E}[|X_n|] < M + \varepsilon$ . Then

$$\sup_n \mathbb{E}[|X_n|] \leq \sup_{k=1}^{N_0} \mathbb{E}[|X_k|] \vee (M + \varepsilon) < \infty$$

Thus  $(b)$  holds.

$(b) \Rightarrow (e)$  Note that  $|x| = x_+ + x_-$ , we thus have  $\forall n$ :

$$\begin{aligned} \mathbb{E}[|X_n|] &= \mathbb{E}[(X_n)_+] + \mathbb{E}[(X_n)_-] \\ &\geq \mathbb{E}[(X_n)_+] \end{aligned}$$

thus  $\sup_n \mathbb{E}[|X_n|] < \infty$  would yield  $\sup_n \mathbb{E}[(X_n)_+] < \infty$ , which is  $(e)$ .

$(d) \Rightarrow (e)$  Idea is the same as  $(a) \Rightarrow (b)$  by just replacing  $|X_n|$  with  $(X_n)_+$ .

$(c) \Rightarrow$  Doob In the prove of Doob's convergence, we have in Doob's up-crossing:

$$\mathbb{E}[U_n[a, b]] \leq \frac{1}{b-a} \mathbb{E}[(X_n - a)_+] \leq \frac{1}{(b-a)} (|a| + \mathbb{E}[(X_n)_+])$$

then applying Fatou's lemma, we have:

$$\mathbb{E}\left[\liminf_n U_n[a, b]\right] \leq \frac{1}{(b-a)} (|a| + \liminf_n \mathbb{E}[(X_n)_+]) < \infty$$

then the rest of the proof is the same as Doob's convergence, thus  $(c)$  implies Doob's convergence.

### Exercise 2 5.3.6

We verify that  $X_n := W_n + an - (K + a)N_{n-1}$  is a sup-MG as follows:

$$\begin{aligned} \mathbb{E}[X_{n+1} - X_n | \mathcal{F}_n] &= \mathbb{E}[W_{n+1} + a(n+1) - (K+a)N_n - W_n - an + (K+a)N_{n-1} | \mathcal{F}_n] \\ &= \begin{cases} \text{if } W_n \geq b, & \mathbb{E}[W_{n+1} - W_n + a | \mathcal{F}_n] \mathbb{1}_{W_n \geq b} \leq 0 \\ \text{if } W_n < b, & \mathbb{E}[W_{n+1} - W_n - K | \mathcal{F}_n] \leq 0 \end{cases} \end{aligned}$$

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and finite expectation is easily verified. Thus  $X_n$  is a supermartingale. Further we notice that  $X_n$  have uniformly bounded difference  $2(K + a)$ , thus by (prop 5.3.5) we have that

$$\mathbb{P} \left( \lim_n X_n \text{ exists } \bigcup_n \liminf X_n = -\infty \right) = 1$$

Now consider the event that  $\{N_\infty = \text{some } \nu < \infty\}$  We note that in this case, we have for large  $n$ :

$$\liminf_n X_n = \liminf_n W_n + an - (K + a)\nu \rightarrow \infty$$

since  $W_n$  is non-negative. Thus

$$\mathbb{P}(N_\infty < \infty) \leq \mathbb{P} \left( \liminf_n X_n = \infty \right) = \mathbb{P} \left( \left( \lim_n X_n \text{ exists } \bigcup_n \liminf X_n = -\infty \right)^c \right) = 0$$

### Exercise 3 5.3.10

We verify that  $W_n := \frac{X_n + 1}{\prod_{k=1}^{n-1} (Y_k + 1)}$  is a supermartingale as follows:

$$\begin{aligned} \mathbb{E}[W_{n+1} - W_n | \mathcal{F}_n] &= \mathbb{E} \left[ \frac{X_{n+1} + 1}{\prod_{k=1}^n (Y_k + 1)} - \frac{X_n + 1}{\prod_{k=1}^{n-1} (Y_k + 1)} \middle| \mathcal{F}_n \right] \\ &= \frac{1}{\prod_{k=1}^n (Y_k + 1)} \mathbb{E}[X_{n+1} + 1 - (X_n + 1)(Y_n + 1) | \mathcal{F}_n] \leq 0 \end{aligned}$$

in which some technical conditions are:  $\log \prod_{k=1}^\infty (Y_k + 1) = \sum_{k=1}^\infty \log(Y_k + 1) \leq \sum_{k=1}^\infty Y_k < \infty$ , and  $\mathbb{E}[|W_n|] \leq \mathbb{E}[|X_n|] < \infty$ . Thus we have that  $W_n$  is a (non-negative) supermartingale. Thus directly we have that  $\lim_n W_n$  exists a.s., denote it as  $W_\infty$ .

Then note that for each  $\omega \in \Omega$ :

$$\prod_{k=1}^{n-1} (Y_k(\omega) + 1) < \infty, \quad \text{a.s.}$$

and is an increasing sequence, thus it has a limit  $Z(\omega)$ . Combine the fact that  $W_n \xrightarrow{\text{a.s.}} W_\infty$ , we have

$$X_n \xrightarrow{\text{a.s.}} W_\infty Z - 1$$

### Exercise 4 5.3.21

Note that  $\mathcal{F}_k^A := \sigma(\mathcal{A}, Z_0, \dots, Z_{k-1})$  and  $\mathcal{F}_k^B := \sigma(\mathcal{B}, Y_0, \dots, Y_{k-1})$  are both filtrations in  $\mathcal{F}$ , and also we have that

$$\begin{aligned} Y_k &= \mathbb{E}[X | \mathcal{F}_k^A] \text{ is U.I. M.G.} \\ Z_k &= \mathbb{E}[X | \mathcal{F}_k^B] \text{ is U.I. M.G.} \end{aligned}$$

which are easy to verify. Thus we have by Theorem 5.3.12 that there exists some  $Y_\infty, Z_\infty$  s.t.

$$\begin{aligned} Y_k &= \mathbb{E}[X | \mathcal{F}_k^A] \xrightarrow[L_1]{\text{a.s.}} Y_\infty \\ Z_k &= \mathbb{E}[X | \mathcal{F}_k^B] \xrightarrow[L_1]{\text{a.s.}} Z_\infty \end{aligned}$$

Now follow similar idea as Theorem 5.3.15, we would obtain that  $Y_\infty \in m\mathcal{F}_\infty^A$  and then  $Y_\infty = \mathbb{E}[X|\mathcal{F}_\infty^A]$  where  $\mathcal{F}_\infty^A = \sigma(\mathcal{A}, Z_0, \dots)$ ; similarly we have  $Z_\infty = \mathbb{E}[X|\mathcal{F}_\infty^B]$  where  $\mathcal{F}_\infty^B = \sigma(\mathcal{B}, Y_0, \dots)$ .

Then we have:

$$\begin{aligned}\mathbb{E}[Y_\infty|\mathcal{F}_\infty^B] &= \mathbb{E}[\mathbb{E}[X|\mathcal{F}_\infty^A]|\mathcal{F}_\infty^B] \\ &= \mathbb{E}[\mathbb{E}[X|\mathcal{F}_\infty^B]|\mathcal{F}_\infty^A] \\ &= \mathbb{E}[Z_\infty|\mathcal{F}_\infty^A] = Z_\infty\end{aligned}$$

and vice versa  $\mathbb{E}[Z_\infty|\mathcal{F}_\infty^A] = Y_\infty$ . Together we have

$$Y_\infty = \mathbb{E}[Z_\infty|\mathcal{F}_\infty^A] = \mathbb{E}[\mathbb{E}[Y_\infty|\mathcal{F}_\infty^B]|\mathcal{F}_\infty^A]$$

thus we must have  $Y_\infty = Z_\infty$  a.s., then apply the U.I. property again we have  $Y_n - Z_n \xrightarrow[L_1]{\text{a.s.}} 0$ .

### Exercise 5 5.3.24

5.(a)

For any fixed  $W$ , there exists some  $n_W$  such that  $W \in L^2\mathcal{F}_{n_W}$ , then we have for  $n > n_W$  that

$$\begin{aligned}\mathbb{E}[WZ_n] &= \mathbb{E}[\mathbb{E}[WZ_n|\mathcal{F}_{n_W}]] \\ &= \mathbb{E}\left[\mathbb{E}\left[\frac{W}{\sqrt{n}}\sum_{k=1}^{n_W}\xi_k|\mathcal{F}_{n_W}\right] + \frac{\mathbb{E}[W|\mathcal{F}_{n_W}]}{\sqrt{n}}\sum_{k=n_W+1}^{\infty}\xi_k\right] \rightarrow 0\end{aligned}$$

since the first term decay to 0 as  $n \rightarrow \infty$ , the second term simply has mean 0.

5.(b)

Define  $W_n := \mathbb{E}[W|\mathcal{F}_n]$ , in this way we can write that  $WZ_n = W_mZ_n + (W - W_m)Z_n$ . Then we consider the two parts:

Part 1: For any fixed  $m$  we have by (a) that

$$\mathbb{E}[W_mZ_n] \rightarrow 0, \quad (n \rightarrow \infty)$$

Part 2: We have

$$\begin{aligned}\limsup_{n \rightarrow \infty} \mathbb{E}[|(W - W_m)Z - n|] &\leq (\lim_{n \rightarrow \infty} \mathbb{E}[|Z_n|]) \cdot \mathbb{E}[|W - W_m|] \\ &= \text{const} \cdot \mathbb{E}[|W - W_m|]\end{aligned}$$

since by CLT, we would have  $Z_n \xrightarrow{d} N(0, 1)$ .

Together we have for any  $m \in \mathbb{N}^+$ ,

$$\limsup_{n \rightarrow \infty} \mathbb{E}[|WZ_n|] \leq \text{const} \cdot \mathbb{E}[|W - W_m|]$$

further we notice that by the definition of  $W_m$  we have  $W_m \xrightarrow[L_1]{\text{a.s.}} W$  thus  $\mathbb{E}[|W - W_m|] \rightarrow 0$ , thus we have  $\limsup_{n \rightarrow \infty} \mathbb{E}[|W Z_n|] = 0$ .

Now if there exists some  $\tilde{Z} \in \mathcal{F}$  s.t.  $Z \xrightarrow{L_2} \tilde{Z}$ , which means that

$$\mathbb{E}\left[\left|Z_n - \tilde{Z}\right|^2\right] \rightarrow 0$$

we would see that

$$\text{L.H.S.} = \mathbb{E}[Z_n^2] + \mathbb{E}[\tilde{Z}^2] - 2\mathbb{E}[Z_n \tilde{Z}] \rightarrow 0$$

however we have  $\liminf_n \mathbb{E}[Z_n^2] \geq \mathbb{E}[\liminf_n Z_n^2] > 0$  by CLT, and that  $\mathbb{E}[Z_n \tilde{Z}] \rightarrow 0$  by the above argument, thus we have a contradiction. Thus  $Z$  does not converge in  $L_2$ .

### Exercise 6 5.4.7

We first prove (5.4.6): Since  $\max$  is a convex function, we have that  $Y_n := \max\{X_n, -1\}$  is still a sub-M.G. Now assume that  $\mathbb{E}[\sup |Y_n|] < \infty$  i.e.  $Y_n$  is integrable. We consider the stopping time defined as:

$$\tau := \inf\{n : Y_n < 0\}$$

and since  $\inf_n X_n < 0$  a.s., we have  $\tau < \infty$  a.s., for which  $Y_\tau < 0$ , and thus we further have by Doob's optional stopping theorem that

$$0 > \mathbb{E}[Y_\tau] \geq \mathbb{E}[Y_0] \geq 0$$

which is a contradiction. Thus we have that  $\mathbb{E}[\sup |Y_n|] = \infty$ .

Now using the lemma, we solve (5.4.7).

6.(a)

For random walk, we know that we have  $S_n/\sqrt{n} \xrightarrow{d} N(0, 1)$  which is a non-degenerate distribution. Thus we have that  $\mathbb{P}(\lim_n S_n \text{ exists}) = 0$ , because for the event  $\{\lim_n S_n \text{ exists}\}$ , we must have that  $S_n(\omega)/\sqrt{n} \rightarrow 0$ .

For such M.G. with bounded difference, by (prop 5.3.5) we have that

$$\lim_n S_n \text{ exists, or } \liminf_n S_n = -\infty, \limsup_n S_n = \infty, \quad \text{a.s.}$$

and from the above argument we have that the first case is w.p. 0, thus we have that  $\limsup_n S_n = \infty$  &  $\liminf_n S_n = -\infty$  a.s., which means that  $\tau_b < \infty$  a.s.

6.(b)

Note that we have  $S_{\tau_b} \geq b > 0$ , i.e.  $\sup_{n \wedge \tau_b} S_n > 0$ . And since  $\tau_b$  is a stopping time, we also have that  $S_{n \wedge b}$  is a M.G. (with  $S_{0 \wedge \tau_b} = 0$ ), thus by the lemma we have that  $\mathbb{E}[\min\{S_n : n \leq \tau_b\}] = \mathbb{E}[\inf_n S_{n \wedge \tau_b}] = \infty$ .

**Exercise 7 5.4.10**

7.(a)

First we have that

$$\begin{aligned}
\mathbb{E}[S_\tau] &= \mathbb{E}\left[\sum_{k=1}^{\infty} \xi_k I_{k \leq \tau}\right] \\
&\leq \mathbb{E}\left[\sum_{k=1}^{\infty} |\xi_k| I_{k \leq \tau}\right] \\
&\stackrel{\substack{\text{monotone} \\ \text{converge}}}{=} \sum_{k=1}^{\infty} \mathbb{E}[|\xi_k|] \mathbb{P}(k \leq \tau) \\
&= \mathbb{E}[|\xi_1|] \sum_{k=1}^{\infty} \mathbb{P}(k \leq \tau) = \mathbb{E}[|\xi_1|] \mathbb{E}[\tau] < \infty
\end{aligned}$$

Then by DCT, we can apply the above argument to  $S_\tau$  directly and obtain that  $\mathbb{E}[S_\tau] = \mathbb{E}[\xi_1] \mathbb{E}[\tau]$ .

7.(b)

Assume  $\mathbb{E}[\xi_i] = 0$  (otherwise we can replace  $\xi_i$  with  $\xi_i - \mathbb{E}[\xi_i]$ ), which would not influence the result. Now it suffice to prove  $\mathbb{E}[S_\tau^2] = \text{var}(\xi_1) \mathbb{E}[\tau]$ . To do so we consider  $S_{n \wedge \tau}$  for which we have

$$S_{n \wedge \tau} = \sum_{k=1}^n \xi_k I_{k \leq \tau}$$

and we have

$$\mathbb{E}[S_{n \wedge \tau}^2] = \mathbb{E}\left[\sum_{k=1}^n \xi_k^2 I_{k \leq \tau}\right] = \text{var}(\xi_1) \sum_{k=1}^n \mathbb{P}(k \leq \tau) = \text{var}(\xi_1) \mathbb{E}[\tau] < \infty$$

for which right hand side does not depend on  $n$ . Thus here we can apply Doob's  $L_p$  convergence theorem ( $p = 2$ ) to obtain

$$\mathbb{E}[S_\tau^2] = \mathbb{E}\left[\lim_n S_{n \wedge \tau}^2\right] = \text{var}(\xi_1) \mathbb{E}[\tau]$$

7.(c)

If  $\xi \geq 0$ , then we have

$$\begin{aligned}
\mathbb{E}[S_\tau] &= \mathbb{E}\left[\sum_{k=1}^{\infty} \xi_k I_{k \leq \tau}\right] \\
&\stackrel{\substack{\text{monotone} \\ \text{converge}}}{=} \sum_{k=1}^{\infty} \mathbb{E}[\xi_k] \mathbb{P}(k \leq \tau) \\
&= \mathbb{E}[\xi_1] \sum_{k=1}^{\infty} \mathbb{P}(k \leq \tau) = \mathbb{E}[\xi_1] \mathbb{E}[\tau]
\end{aligned}$$

The proof of Wald's second is still the same.

**Exercise 8 5.4.11**

8.(a)

In 5.4.7 we already proved it (by noting that SRW is also a M.G. with bounded difference).

8.(b)

By the similar argument as in (coroll 5.4.8) in which we use the M.G.  $M_n := \exp(\lambda S_n)/M(\lambda)^n$  we have

$$S_{\tau_{a,b}} = \begin{cases} -a, & w.p. \frac{e^{\lambda b} - 1}{e^{\lambda b} - e^{-\lambda a}} \\ b, & w.p. \frac{1 - e^{-\lambda a}}{e^{\lambda b} - e^{-\lambda a}} \end{cases}$$

where  $\lambda = \log \frac{1-p}{p}$ . Thus we have

$$\begin{aligned} \mathbb{E}[S_{\tau_{a,b}}] &= -a \cdot \frac{e^{\lambda b} - 1}{e^{\lambda b} - e^{-\lambda a}} + b \cdot \frac{1 - e^{-\lambda a}}{e^{\lambda b} - e^{-\lambda a}} \\ &= \mathbb{E}[\xi_1] \mathbb{E}[\tau_{a,b}] \\ &= (2p - 1) \mathbb{E}[\tau_{a,b}] \end{aligned}$$

and we have

$$\mathbb{E}[\tau_{a,b}] = \frac{1}{2p - 1} \left( -a \cdot \frac{e^{\lambda b} - 1}{e^{\lambda b} - e^{-\lambda a}} + b \cdot \frac{1 - e^{-\lambda a}}{e^{\lambda b} - e^{-\lambda a}} \right)$$

where  $\lambda = \log \frac{1-p}{p}$ .

And for the case that  $p = 1/2$ , i.e.  $\lambda = 0$ , we use Wald's second to obtain that

$$\begin{aligned} \mathbb{E}[\tau_{a,b}] &= \mathbb{E}[S_\tau^2] / \text{var}(\xi_1) \\ &= \mathbb{E}[S_\tau^2] \\ &= ab \end{aligned}$$

8.(c)

We have

$$\tau_{a,b} = \inf\{n : S_n = -a\} \wedge \inf\{n : S_n = b\} \xrightarrow{a \rightarrow \infty} \inf\{n : S_n = b\} = \tau_b$$

(and the convergence is monotone, which is trivial). We thus have by Wald's first that

$$\mathbb{E}[\tau_b] = \mathbb{E}[S_{\tau_b}] / \mathbb{E}[\xi_1] = b / (2p - 1)$$

8.(d)

For  $p \geq 1/2$ , it's a submartingale, thus by 5.4.7 we have  $\tau_b < \infty$  a.s..

8.(e)

$$\begin{aligned}\mathbb{E}[Y_{n+1}|\mathcal{F}_n] &= (S_n^4 + 6S_n^2 + 1) - 6(n+1)(S_n^2 + 1) + c_1(n^2 + 2n + 1) + c_2(n+1) \\ &= S_n^4 - 6nS_n^2 + c_1n^2 + c_2n + (1 - 6(n+1) + c_1(2n+1) + c_2)\end{aligned}$$

to make it a M.G., we have  $c_1 = 3, c_2 = 2$ , thus we have

$$Y_n := S_n^4 - 6nS_n^2 + 3n^2 + 2n$$

is a M.G.

For which we have

$$\begin{aligned}\mathbb{E}[Y_{\tau_{b,b}}] &= \mathbb{E}[\mathbb{E}[Y_k|\tau_{b,b} = k]] \\ &= b^4 - 6b^2\mathbb{E}[\tau_{b,b}] + 3\mathbb{E}[\tau_{b,b}^2] + 2\mathbb{E}[\tau_{b,b}] \\ &= b^4 - 6b^2 \cdot b^2 + 3\mathbb{E}[\tau_{b,b}^2] + 2b^2 \\ &= \mathbb{E}[Y_0] = 0\end{aligned}$$

where the last step is by Doob's optional stopping theorem. Now we have

$$\mathbb{E}[\tau_{b,b}^2] = \frac{5b^4 - 2b^2}{3}$$

## Exercise 9 5.4.12

9.(a)

Since  $M_n = \exp[\lambda S_n] / \mathbb{E}[\exp[\lambda n \xi_1]]$  is a M.G., then by Doob's optional stopping theorem we have

$$\begin{aligned}1 &= \mathbb{E}[M_0] \\ &= \mathbb{E}[M_{\tau_b}] \\ &= \mathbb{E}[\exp[\lambda S_{\tau_b}] / \mathbb{E}[\exp[\lambda \xi_1]]^{\tau_b}] \\ &= e^{\lambda b} \mathbb{E}[M(\lambda)^{-\tau_b}]\end{aligned}$$

9.(b)

Denote  $s = M(\lambda)^{-1} \in (0, 1)$  for some  $\lambda$ , and use  $b = 1$  in the above result then we have

$$\mathbb{E}[s^{\tau_1}] = e^{-\lambda}$$

and that

$$\mathbb{E}[\tau_b] = e^{-\lambda b} = \mathbb{E}[s^{\tau_1}]^b$$

Now we revisit  $s = M(\lambda)^{-1} = \frac{1}{pe^\lambda + qe^{-\lambda}}$ , and we have

$$\begin{aligned}
 e^{-\lambda} &= \frac{1}{2qs} \frac{2qe^{-\lambda}}{pe^\lambda + qe^{-\lambda}} \\
 &= \frac{1}{2qs} \left( 1 - \frac{pe^\lambda - qe^{-\lambda}}{pe^\lambda + qe^{-\lambda}} \right) \\
 &= \frac{1}{2qs} \left( 1 - \sqrt{\frac{(pe^\lambda - qe^{-\lambda})^2}{(pe^\lambda + qe^{-\lambda})^2}} \right) \\
 &= \frac{1}{2qs} \left( 1 - \sqrt{1 - 4 \frac{pe^\lambda \cdot qe^{-\lambda}}{pe^\lambda + qe^{-\lambda}}} \right) \\
 &= \frac{1}{2qs} \left( 1 - \sqrt{1 - 4pqs^2} \right)
 \end{aligned}$$

substitute this back to the previous result we have

$$\mathbb{E}[s^{\tau_1}] = e^{-\lambda} = \frac{1}{2qs} \left( 1 - \sqrt{1 - 4pqs^2} \right), \quad s = M(\lambda)^{-1} \in (0, 1), \quad p \geq 1/2$$

9.(c)

In 5.4.11 (c) we had that  $\tau_{a,b} \xrightarrow{a \rightarrow \infty} \tau_b$  a.s.. Thus we have

$$\mathbb{P}(S_{\tau_{a,b}} = b) \rightarrow \mathbb{P}(S_{\tau_b} = b) = \mathbb{P}(\tau_b < \infty) = \frac{1}{e^{\lambda_\star b}} = \exp[-\lambda_\star b]$$

where  $\lambda_\star = \log \frac{1-p}{p}$ .

9.(d)

Continuing from the previous result, we have

$$\mathbb{P}(S_{\tau_b} = b) = \exp[-\lambda_\star b], \quad b \in \mathbb{N}^+$$

which gives us for  $Z := 1 + \max_{k \geq 0} S_k$  that:

$$\begin{aligned}
 \mathbb{P}(Z = b + 1) &= \mathbb{P}\left(\max_{k \geq 0} S_k = b\right) \\
 &= \mathbb{P}(\tau_b < \infty, \tau_{b+1} = \infty) \\
 &= \mathbb{P}(\tau_b < \infty, \text{after reaching } b, \text{ never reach } b + 1) \\
 &= \mathbb{P}(\tau_b < \infty) (1 - \mathbb{P}(\tau_1 = \infty)) \\
 &= \exp[-\lambda_\star b] (1 - \exp[-\lambda_\star]) \\
 &\sim \text{Geo}(1 - \exp[-\lambda_\star])
 \end{aligned}$$

## Exercise 10 5.4.14

Denote  $26 = \kappa$



10.(a)

We have

$$\mathbb{E}[M_{n+1}|\mathcal{F}_n] = \frac{1}{\kappa}M_n \times \kappa + \left(1 - \frac{1}{\kappa}\right) \times 0 = M_n$$

thus  $M_n$  is a martingale.

The construction of total gain would look like:

$$M_n = \sum_n (\kappa^{10}(\kappa \mathbf{1}_{\text{PROBABILITY}} - 1) + \kappa^9(\kappa \mathbf{1}_{\text{PROBABILITY}} - 1) + \cdots + \kappa^1(\kappa \mathbf{1}_{\text{PR}} - 1) + \kappa^0(\kappa \mathbf{1}_{\text{P}} - 1))$$

10.(b)

We see that  $\hat{\tau}$  is equivalent to that  $\hat{\tau} := \inf\{n \geq 11 : \text{gain of a gamblers} = \kappa^{11}\}$  (i.e. when the gambler win 11 consecutive games, he gets the word ‘PROBABILITY’). WHEN The gambler wins 11 consecutive games, there are other  $\hat{\tau} - 1$  gamblers losing. And as we mentioned earlier, we have  $M_n$  being a martingale, thus by Doob’s optional stopping theorem we have

$$0 = \mathbb{E}[M_0] = \mathbb{E}[M_{\hat{\tau}}] = \mathbb{E}[\mathbb{E}[M_k|\tau = k]] = \sum_{k=0}^{10} (\kappa^{k+1} - \kappa^k) - (\mathbb{E}[\hat{\tau}] - 1)$$

And it gives that

$$\mathbb{E}[\hat{\tau}] = \kappa^{11} = 26^{11}$$

For ‘ABRACADABRA’ = ‘ABRA’ + ‘CAD’ + ‘ABRA’, we have that

$$\begin{aligned} 0 = \mathbb{E}[M_0] &= \sum_{k=0}^{10} (\kappa^{k+1} - \kappa^k) + \sum_{k=0}^3 (\kappa^{k+1} - \kappa^k) - (\mathbb{E}[\tilde{\tau}] - 2) = \kappa^{11} + \kappa^4 - \mathbb{E}[\tilde{\tau}] \\ \Rightarrow \mathbb{E}[\tilde{\tau}] &= \kappa^{11} + \kappa^4 = 26^{11} + 26^4 \end{aligned}$$

where  $+\kappa^4$  for that, when reaching ‘ABRACADABRA’, there is also a gambler who was winning at the first 4 letters ‘ABRA’.