

Time Series Analysis Forecasting

Dr. Ludovic Amruthalingam

ludovic.amruthalingam@hslu.ch

Informatik



Outline

- Forecasting
- Forecasting with baselines, ARIMA
- Underfitting and overfitting
- Evaluation workflow
- Splitting time series data
- Residual analysis
- Confidence intervals
- Information criteria
- Performance metrics
- Back-transforms

Forecasting

Extrapolating past observations to predict future data.

- Works well provided future data follows past patterns.
- Strong signals (low noise) can lead to accurate forecasts.
- Noise increases uncertainty, making predictions reliable only for the short term.

Sources of uncertainty

- Data: unexpected disruption from past patterns.
- Model: chosen model may not represent the true data-generating process.
- Parameters: even with the correct model, estimated parameters may be inaccurate.
- Forecasts: model typically yield an estimate of the **conditional mean** of future instances, which may be strongly influenced by future **unpredictable innovations**.

Forecasts must be complemented with a measure of the model uncertainty, typically prediction intervals.

Forecasting

Given a time series realization $\{x_1, x_2, ... x_T, ... x_n\}$,

The h-step forecast of x_{T+h} based on the data $\{x_1, x_2, ... x_T\}$ is represented as $\hat{x}_{T+h|T}$.

- T is the forecast time.
- h is the **forecast horizon** i.e., how far into the future the forecast is made.
- T + h is the **target time** i.e., the time point of the forecast.

Considering a monthly time series and a 1-year forecast horizon,

- Point forecast is $\hat{x}_{T+12|T}$.
- Multi-step forecast is $\{\hat{x}_{T+1|T}, ... \hat{x}_{T+12|T}\}$.

HSLU

Forecasting baselines

Mean: forecasts are equal to the average of the observed data, $\hat{x}_{T+h|T} = \frac{1}{T}\sum_{i=1}^{T} x_i$

Naïve: forecasts are equal to the last observed value of the series, $\hat{x}_{T+h|T} = x_T$

Naïve is optimal for random walk process.

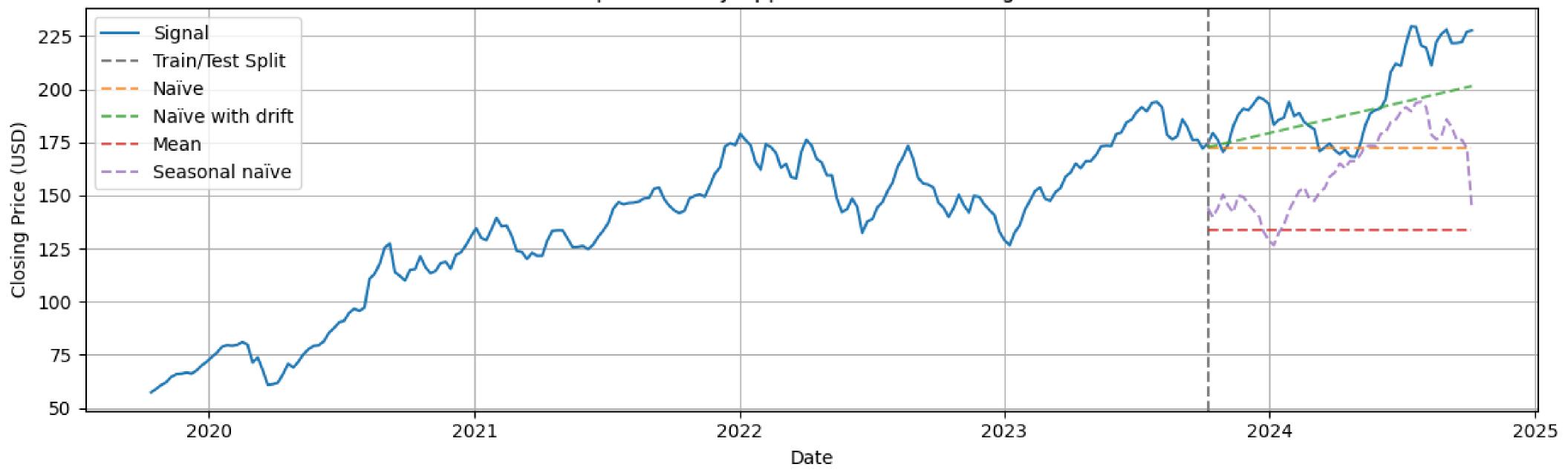
Seasonal naïve: forecasts are equal to the last observed value from the same season, $\hat{x}_{T+h|T} = x_{T+h-[h/P]P}$

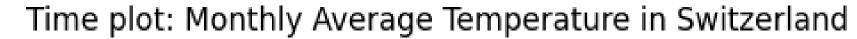
• Setting P = 1 (non-seasonal data) results in the naïve forecast

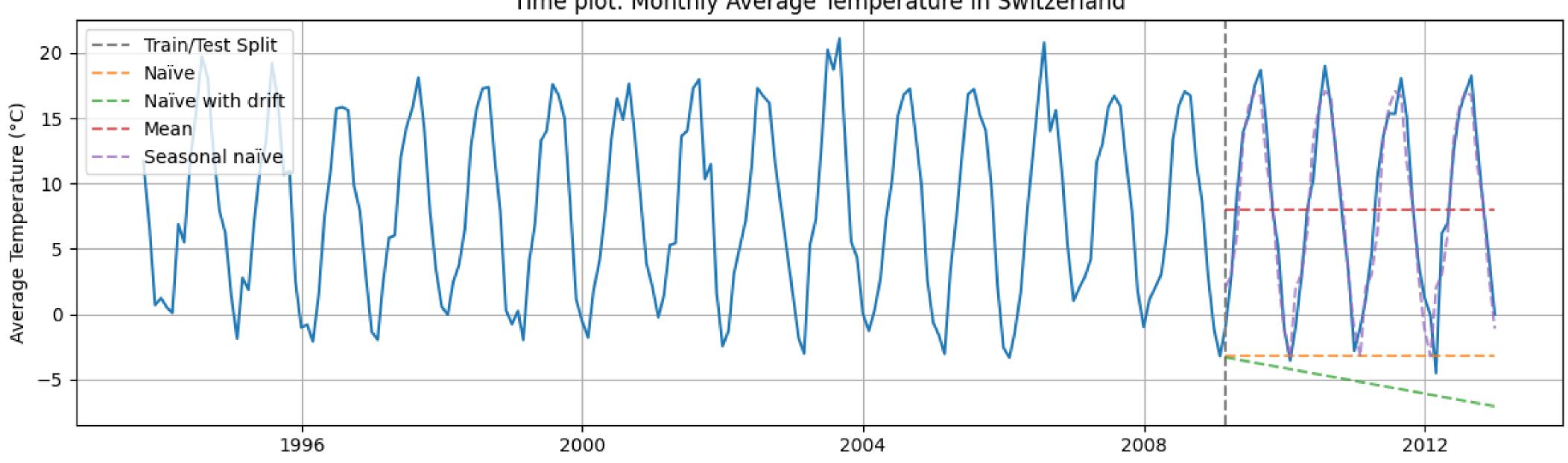
Naïve with drift: naïve forecast with linear drift, $\hat{x}_{T+h|T} = x_T + h \frac{x_T - x_1}{T-1}$

Baselines serve as benchmark to evaluate the added value of more complex methods.









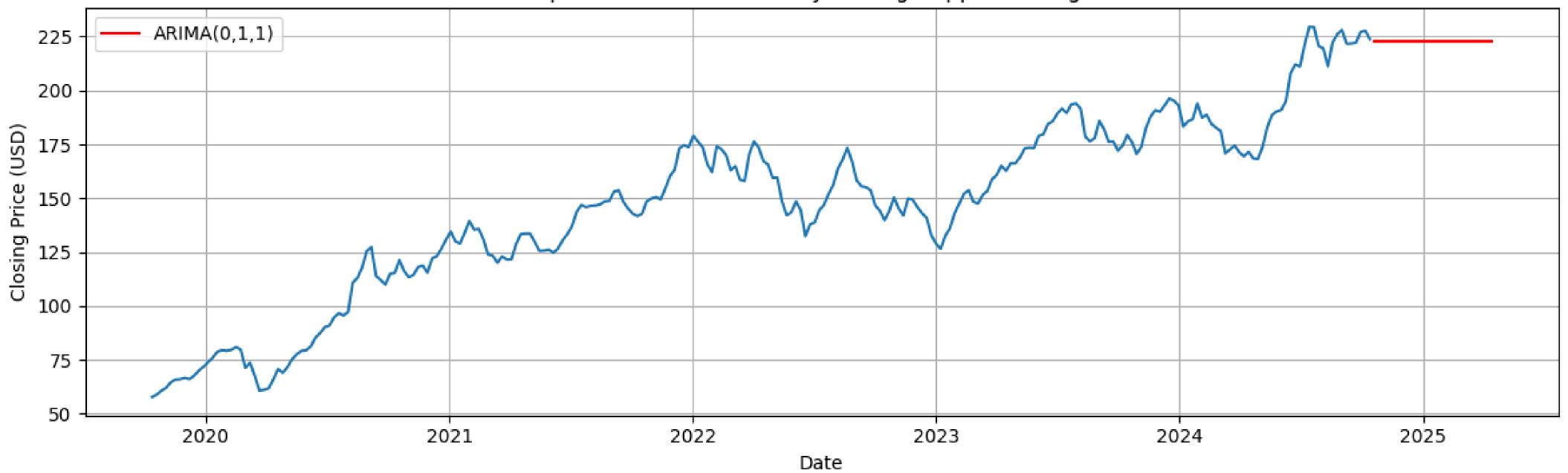
Forecasting with ARIMA

The forecast $\hat{x}_{T+h|T}$ from an ARIMA model can be computed as follows:

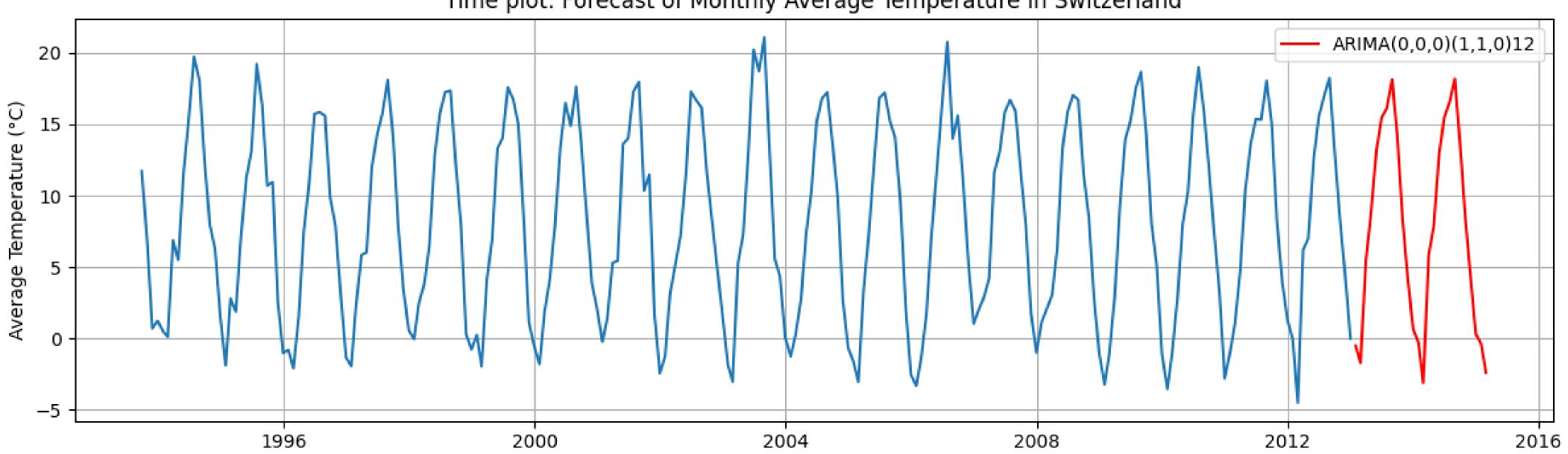
- 1. Rewrite the ARIMA equation $\Phi(B)\nabla^d X_t = \Phi(B)(1-B)^d X_t = c + \Theta(B)W_t$ with x_t on the left-hand side.
 - For an ARIMA(1,1,1): $(1-\hat{\phi}B)(x_T-x_{T-1})=\hat{c}+(1+\hat{\theta}B)w_T \Leftrightarrow x_T=\hat{c}+w_T+\hat{\theta}w_{T-1}+(\hat{\phi}+1)x_{T-1}-\hat{\phi}x_{T-2}$
- 2. Replace future observations with their forecast, future errors with zero, and past errors with the ARIMA residuals.
 - $T + h \rightarrow x_{T+h} = \hat{c} + w_{T+h} + \hat{\theta} w_{T+h-1} + (\hat{\phi} + 1) x_{T+h-1} \hat{\phi} x_{T+h-2}$
 - $h = 1 \longrightarrow \hat{x}_{T+1|T} = \hat{c} + 0 + \hat{\theta} \hat{w}_T + (\hat{\phi} + 1)x_T \hat{\phi} x_{T-1}$
 - $h = 2 \longrightarrow \hat{x}_{T+2|T} = \hat{c} + 0 + 0 + (\hat{\phi} + 1)\hat{x}_{T+1|T} \hat{\phi}x_T$
 - $h = 3 \longrightarrow \hat{x}_{T+3|T} = \hat{c} + 0 + 0 + (\hat{\phi} + 1)\hat{x}_{T+2|T} \hat{\phi}\hat{x}_{T+1|T}$

•

Time plot: Forecast of Monthly Average Apple Closing Prices







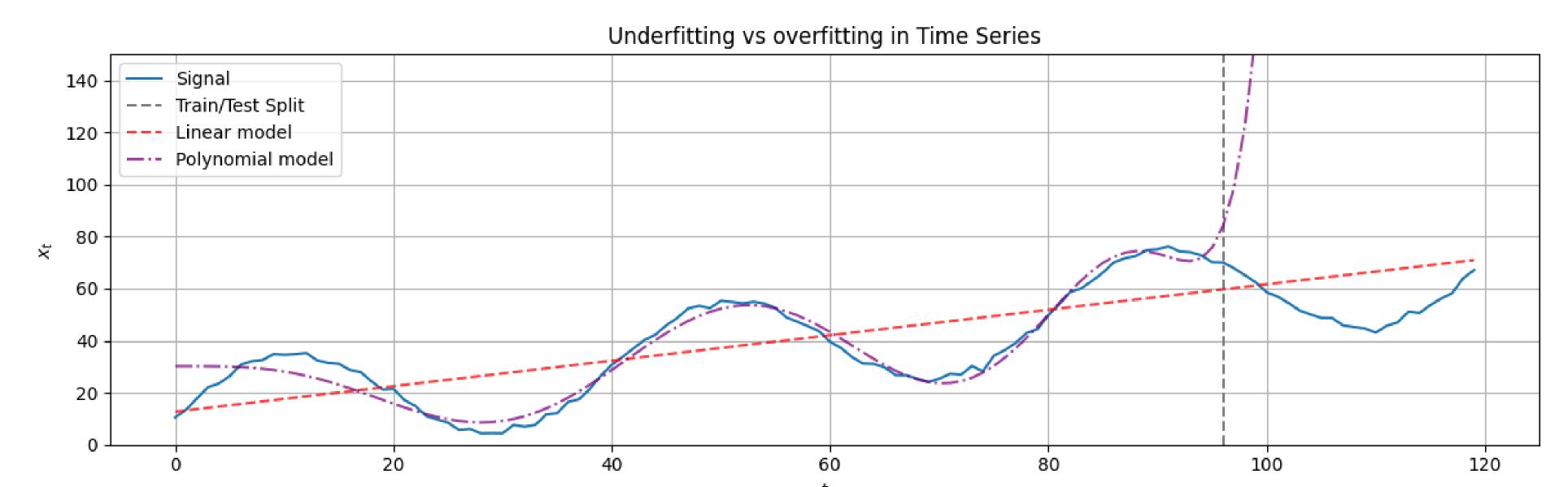
Underfitting and overfitting

Underfitting occurs when a model fits training data poorly and fails to generalize to new data.

Model is too simple to capture the underlying patterns in the data.

Overfitting occurs when a model fits training data very well but fails to generalize to new data.

- Model is too complex and learns to reproduce the training data exactly.
- Training data is too small.
- Training procedure does not involve regularization (discussed in following lectures).



Evaluation workflow

Given a time series realization $\{x_1, x_2, ... x_T, ... x_V, ... x_n\}$,

- 1. Split dataset into training $\{x_1, ... x_T\}$, validation $\{x_{T+1}, ... x_V\}$, and test $\{x_{V+1}, ... x_n\}$ sets.
 - Validation set is required either when model fitting involves hyperparameters tuning or when model selection is based on performance metrics.
 - · Multiple versions of training and validation sets can be considered with cross-validation.
- 2. Train candidate models on training set.
 - · Tune hyperparameters using the validation set.
- 3. Select model based on model fit, complexity and performance on the validation set.
 - Information criteria, residual analysis, uncertainty, performance metrics
- 4. Train selected model on training + validation sets then evaluate performance on test set.
 - Metrics provide an indication of how well the model will forecast new data.

Splitting time series data & rolling cross-validation

Split dataset into training $\{x_1, ... x_T\}$, validation $\{x_{T+1}, ... x_V\}$, and test $\{x_{V+1}, ... x_n\}$ sets

- Time series must be split chronologically → no random splits.
- Seasonality and trends: ensure the splits account for any patterns in the data.
- Validation and test set should be at least as large as the **forecast horizon** i.e., $h \le V T$ and $h \le n V$.

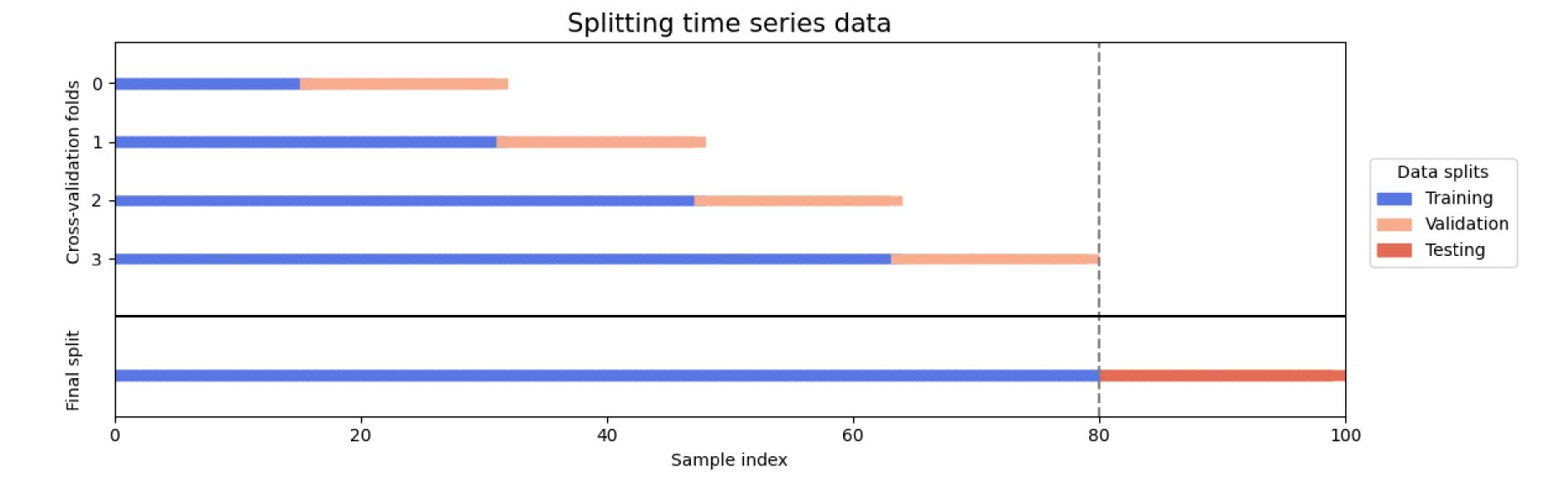
Use test data once for final evaluation, otherwise risk of over-estimating performance on new data

→ compare candidate models performance on the validation set.

Rolling Cross-Validation: sequentially increase the training set, while moving the validation set forward.

Expanding Window Cross Validation

HSLU



Residual analysis

Residuals are the difference between observed values and predicted values: $e_i = x_i - \hat{x}_{i|T}$ for i = 1, ... T.

- Also called training set errors, it is an estimate of the noise/innovation component of the data.
- Residuals are expected to be normal, uncorrelated, zero-mean, and homoscedastic.
- Analyse standardized residuals $\tilde{e}_i = e_i/\hat{\sigma}_e$

Identify patterns or autocorrelations that the model did not capture.

- Time plot, correlogram
- Ljung-Box test: null hypothesis (H_0) states that residuals are **uncorrelated** up to a certain lag.

Validate model assumptions, typically $E_{\rm t} \sim \mathcal{N}(0, \sigma^2)$.

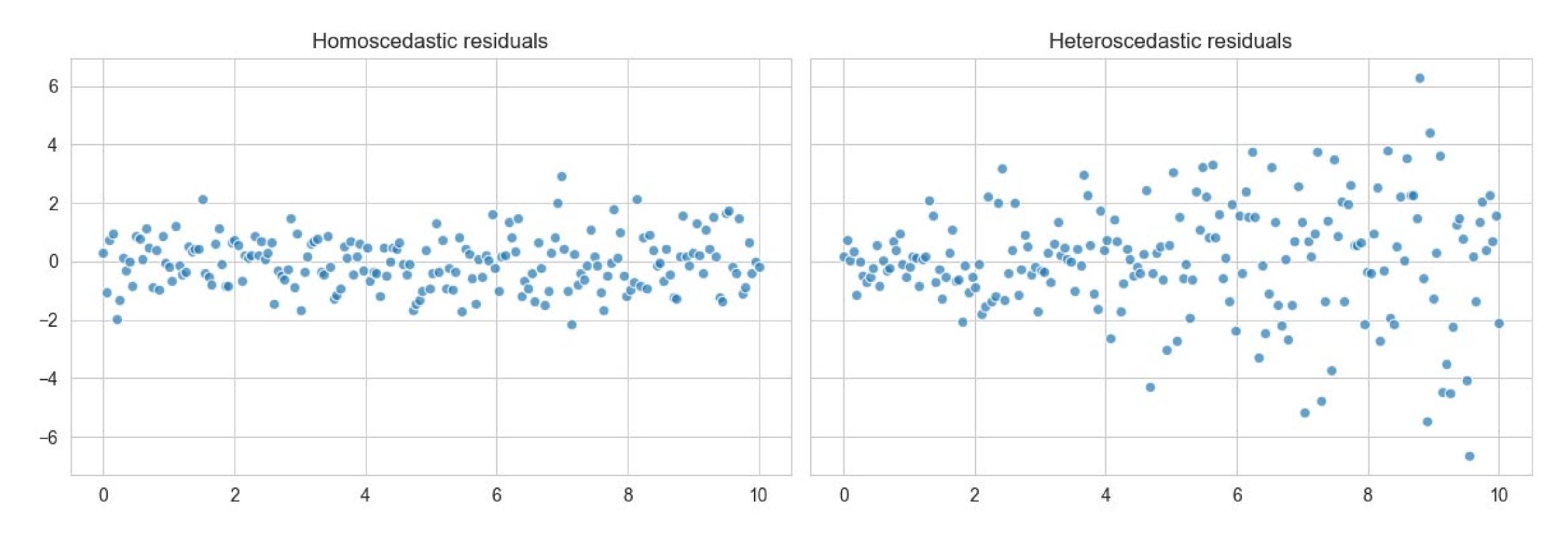
- Q-Q plot: compare the residuals quantiles with normal quantiles.
- Histogram: visual representation of the residuals distribution.

Evaluate how well a model utilizes available signal in the data but does not help with model selection.

HSLU

Recap – Homoscedasticity

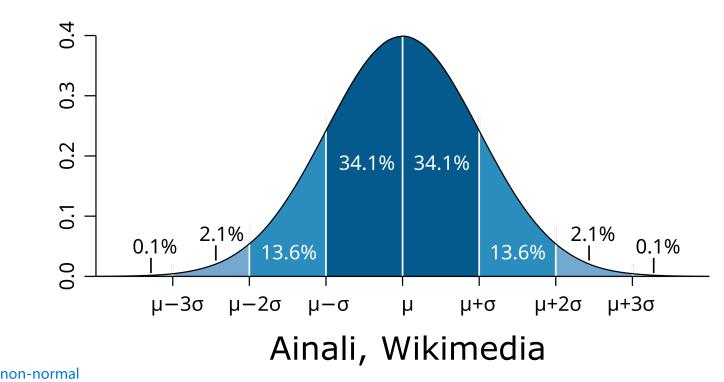
The variance of the error is **constant** over the entire feature space.



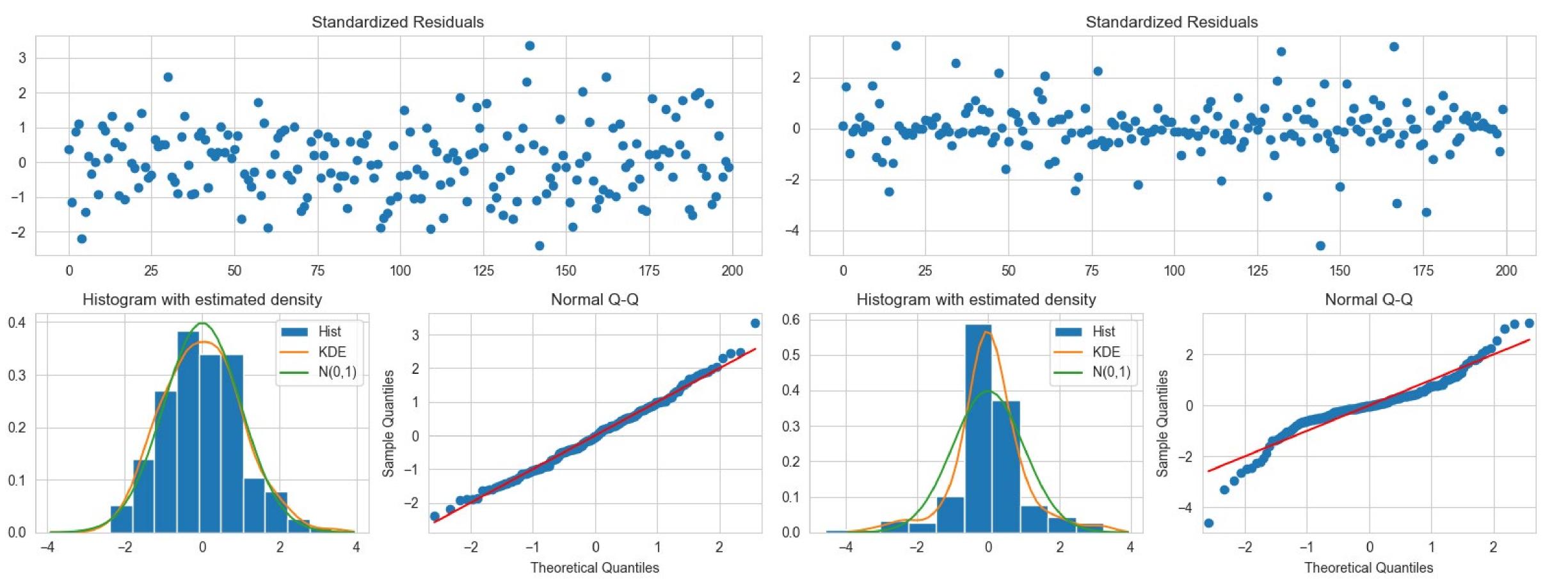
Recap - Normality

HSLU

The model error given the features follows a **normal distribution** $\epsilon \mid X \sim \mathcal{N}(0, \sigma^2)$.



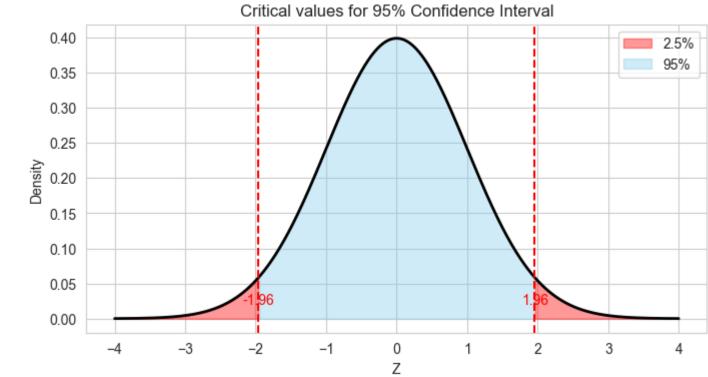
Seite 14



15

Example of the entire residual analysis

Confidence intervals (CI) for normal residuals



Confidence intervals provide a measure of **forecast uncertainty**.

When the residuals are **normally distributed**, the CI of $\hat{x}_{T+h|T}$ is $\hat{x}_{T+h|T} \pm z_{1-\alpha/2}\hat{\sigma}_{h|T}$

- $z_{1-\alpha/2}$ is the critical value of the normal distribution e.g., for 95% CI $z_{1-\alpha/2}=1.96$
- $\hat{\sigma}_{h|T}$ is an **estimate of the standard deviation** of the h-step forecast with $\hat{\sigma}_1 = \hat{\sigma}_e$ (residuals std)

$$\hat{\sigma}_e = \sqrt{\frac{1}{T - k - m} \sum_{i=m+1}^{T} e_i^2}$$

with k the number of model parameters and m the number of missing residuals due to initialization.

For h > 1, $\hat{\sigma}_h$ depends on the forecasting method and how the **innovations accumulate**:

Mean	Naïve	Seasonal naïve	Naïve with drift
$\hat{\sigma}_{h T} = \hat{\sigma}_e \sqrt{1 + \frac{1}{T}}$	$\hat{\sigma}_{h T} = \hat{\sigma}_e \sqrt{h}$	$\hat{\sigma}_{h T} = \hat{\sigma}_e \sqrt{\lceil h/p \rceil}$	$\hat{\sigma}_{h T} = \hat{\sigma}_e \sqrt{h\left(1 + \frac{h}{T - 1}\right)}$

Confidence intervals (CI) for non-normal residuals

non-parametric way of calculating the confidence interval

Use bootstrapping when the residuals are uncorrelated and have constant variance.

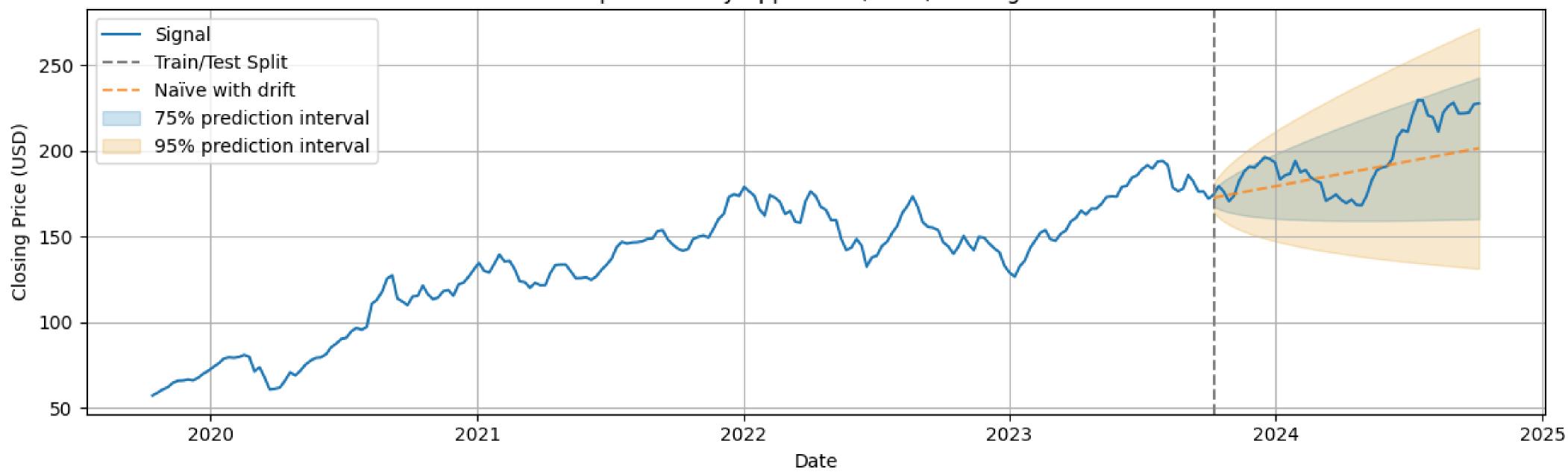
with (strongly) correlated residuals --> problem correlated residuals mean remaining signal that the model hasn't picked up.

Assuming future and past errors will be similar, generate possible futures:

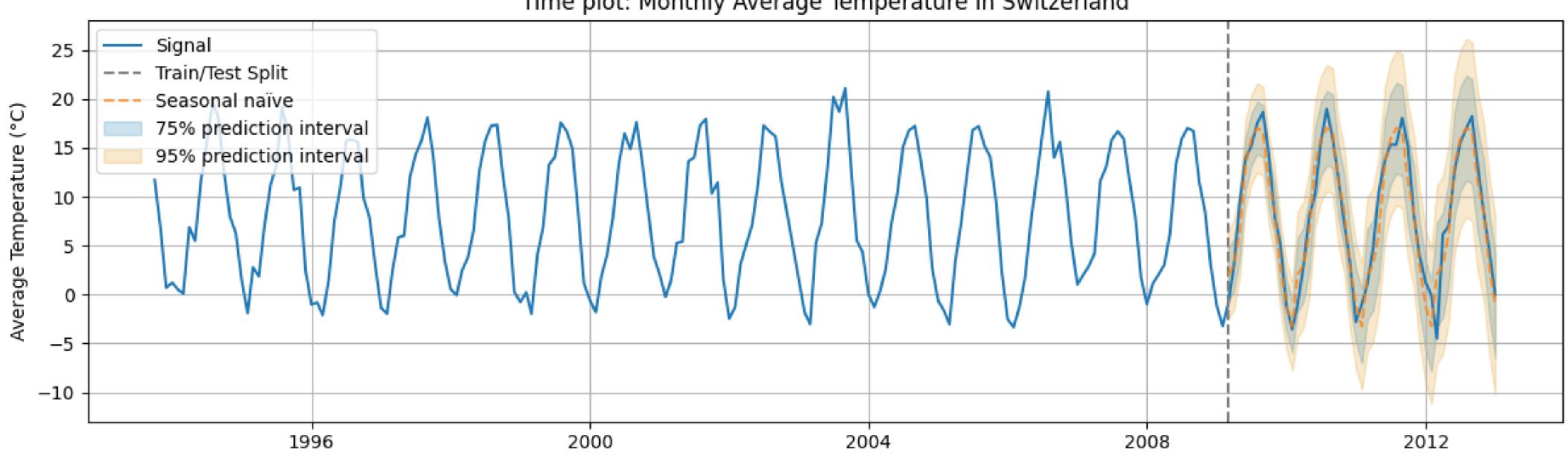
- 1. Fit forecasting model on $\{x_1, x_2, ... x_T\}$ and compute residuals $\{e_1, e_2, ... e_T\}$.
- 2. Resample residuals with replacement to simulate future forecast errors $\{e_{k_1}, e_{k_2}, \dots e_{k_h}\}$.
- 3. Generate futures by adding resampled residuals to model forecasts $\{\hat{x}_{T+1|T} + e_{k_1}, \hat{x}_{T+2|T} + e_{k_2}, ... \hat{x}_{T+h|T} + e_{k_h}\}$.
- 4. Repeat step 2, 3 multiple times e.g., 1000 iterations.
- 5. Derive confidence intervals by computing the percentiles.
 - e.g., for 95% CI use the 2.5th and 97.5th percentiles.

Note that the bootstrapped confidence intervals are not symmetric.

Time plot: Weekly Apple Inc. (AAPL) Closing Prices



Time plot: Monthly Average Temperature in Switzerland



Information criteria

Measure of the goodness of fit of a model while penalizing for model complexity.

Goodness of fit is measured by the likelihood of the data under the model:

$$L = \prod_{i=1}^{T} \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{e_i^2}{2\sigma^2}\right) \text{ with } e_i = x_i - \hat{x}_{i|T} \text{ and assuming } e_i \sim \mathcal{N}(0, \sigma^2)$$

Akaike's Information Criterion: AIC = 2k - 2log(L) with k the number of model parameters.

Bayesian Information Criterion: BIC = log(T) k - 2log(L)

For ARIMA
$$e_i = \widehat{w}_i$$
 and $k = \begin{cases} p+q & \text{if } c=0\\ p+q+1 & \text{if } c\neq 0 \end{cases}$

Select model minimizing either AIC or BIC (for models in the same class)

- BIC tends to favor simpler models than AIC due to a larger penalty term.
- AIC prioritizes **model fit** (potentially better performance), while BIC emphasizes **model simplicity** (faster inference, simpler model interpretation).

Seite 19

Performance metrics

Given a validation set $\{x_{T+1}, ... x_V\}$, the forecast performance can be evaluated based on

- Scale-dependent errors: $e_i = x_i \hat{x}_{i|T}$ (same unit as the data \rightarrow not comparable for TS with different units)
- Scaled errors: $e_i = (x_i \hat{x}_{i|T})/(\frac{1}{T}\sum_{j=1}^T |x_j \hat{x}'_{j|T}|)$ with $\hat{x}'_{j|T}$ a baseline training forecasts.

Considering the **multi-step** forecasts $\{\hat{x}_{T+1|T}, ... \hat{x}_{V|T}\}$, the errors can be aggregated as follows:

• Mean absolute (scaled) error MAE/MASE: $\frac{1}{T-V}\sum_{i=T+1}^{V}|e_i|$

→ robust to outliers

• Root mean squared (scaled) error RMSE/RMSSE: $\sqrt{\frac{1}{T-V}\sum_{i=T+1}^{V}e_i^2}$

→ sensitive to outliers

When units has a **meaningful zero**, consider the mean absolute **percentage errors** MAPE $\frac{1}{T-V}\sum_{i=T+1}^{V}\left|\frac{100(x_i-\hat{x}_{i|T})}{x_i}\right|$

Rolling forecast performance

When the validation set encompasses multiple forecast horizons $\{x_{T+1}, ..., x_h, ... x_V\}$,

- 1. With i=0 for the first iteration, forecast $\{\hat{x}_{T+i+1|T},...\hat{x}_{T+i+h|T}\}$.
- 2. Compute performance metrics of the forecast.
- 3. Increment i = i + k with k the chosen step-size, typically k = h.
- 4. Refit model with the newly available values $\{x_{T+i+1}, ..., ..., x_{T+i+h}\}$.
 - · On the test set choosing between refit vs update strategy depends on the training objective.
 - Update: recalculate model internal state given new data points without refitting its parameters.
- 5. Repeat until the end of the validation set and then aggregate performance.

HSLU

Back-transforms

("just be aware of this, but library will do it for you")

To obtain forecasts on the original scale, we need to reverse transformations applied to the data.

The back-transform for differencing is

- First-order: $\hat{x}_{T+h|T} = x_T + \sum_{i=1}^h \hat{y}_{T+i|T}$ with $y_t = \nabla x_t$
- Seasonal differencing: $\hat{x}_{T+(kP+n)|T} = x_{T+n} + \sum_{i=1}^k \hat{y}_{T+iP+n|T}$ with $y_t = \nabla_P x_t$

Reversing non-linear transforms does **not** preserve the **mean** from the transformed scale but the **median** (assuming the distribution on the transformed scale is symmetric)

- Considering a log-normal distribution $y_t = \log(x_t)$, the mean of y_t corresponds to the median of x_t .
- Bias correction is needed to account for the variance in the transformed space.

$\hat{x}_{T+h|T} = \begin{cases} \exp(\hat{y}_{T+h|T}) & \text{if } \lambda = 0 \\ sign(\lambda \hat{y}_{T+h|T} + 1)|\lambda \hat{y}_{T+h|T} + 1|^{\frac{1}{\lambda}} & \text{if } \lambda \neq 0 \end{cases}$ $\hat{x}_{T+h|T}^* = \begin{cases} \hat{x}_{T+h|T} \left(1 + \sigma_h^2/2\right) & \text{if } \lambda = 0 \\ \hat{x}_{T+h|T}^* \left(1 + \frac{\sigma_h^2(1-\lambda)}{2(\lambda \hat{y}_{T+h|T} + 1)^2}\right) & \text{if } \lambda \neq 0 \end{cases}$

Exercise

Forecasting

- Split real-world time series into train/test sets.
- Fit ARIMA on training, forecast test observations.
- Plot forecasts vs. actuals, what patterns does your model capture or miss?
- Perform residual analysis
- Generate 80% / 95% forecast intervals. Are values within the intervals? What does this imply?

Evaluation workflow

- Use cross-validation to generate different validation folds
- Compute rolling forecast performance. Which metric is best suited for your data?
- Compare ARIMA with baseline models (mean, naïve, seasonal naïve).
- Review <u>sktime forecasting notebook</u> and test different <u>forecasting approaches</u>.
- Select best model and evaluate performance on the test set.