

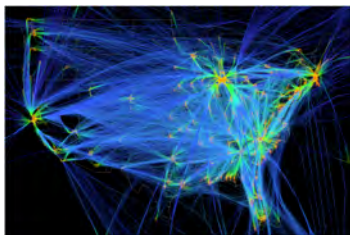
Exploring Densities of Gaussian Quadratic Forms

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Presentation outline

1. Airport delay motivation
2. Moments and densities
3. Saddlepoint approximations
4. (An attempt at) Expansion approximations



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Spatial distribution of airport delays

- Imagine two scenarios, both with a **total delay of 180 minutes**:



- Some disruption, or **off-nominal** event caused the delays ...
- ... but one scenario is **expected** given historical correlations, while the other is **unexpected**.
- Operationally, the two scenarios are very different
 - What traffic management initiatives to deploy?
 - Airline recovery strategies?



Spatial distribution of airport delays (real example)

February 3, 2014

Total delay: 2.718×10^4 min



August 8, 2016

Total delay: 2.715×10^4 min



Delay network visualization courtesy of Karthik Gopalakrishnan

Airport delays as graph-supported signals

- For our purposes, graph signals **live on vertices** \mathcal{V}
 - Formally, a graph signal is a **map** $f: \mathcal{V} \rightarrow \mathbb{R}$
- Given N vertices (airports), at time t , the **graph signal vector** \mathbf{x}_t is of the form:

$$\mathbf{x}_t = \begin{pmatrix} x_{1,t} \\ \vdots \\ x_{N,t} \end{pmatrix} \in \mathbb{R}_{\geq 0}^{N \times 1}.$$

- $x_{i,t}$ is the delay at airport i at time t



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- $x_{i,t}$ is the delay at airport i at time t
- Let us consider \mathbf{x} drawn from $\mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma})$



Definition: Total variation (TV) of a graph signal

- Given an adjacency matrix $A = [a_{ij}]$ and the corresponding **degree matrix** $D = \text{diag}(\sum_j a_{1j}, \dots, \sum_j a_{Nj})$, compute the (combinatorial) **graph Laplacian** $D - A = \mathcal{L} \in \mathbb{R}^{N \times N}$.
- Given \mathcal{L} and a graph signal vector $\mathbf{x}_t \in \mathbb{R}^{N \times 1}$, the **total variation (TV)** of \mathbf{x}_t on the graph is given by:

$$\mathbf{x}_t^\top \mathcal{L} \mathbf{x}_t = \frac{1}{2} \sum_{i \neq j} a_{ij} (x_{i,t} - x_{j,t})^2.$$



Definition: Total variation (TV) of a graph signal

- Given an adjacency matrix $A = [a_{ij}]$ and the corresponding **degree matrix** $D = \text{diag}(\sum_j a_{1j}, \dots, \sum_j a_{nj})$, compute the (combinatorial) **graph Laplacian** $D - A = \mathcal{L} \in \mathbb{R}^{N \times N}$.
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- With \mathbf{x} as m.v. Gaussian, call $\boxed{Q(\mathbf{x}) = \mathbf{x}^\top \mathcal{L} \mathbf{x}}$ a **Gaussian quadratic form**.



TV as a metric

February 3, 2014

Total delay: 2.718×10^4 min
Total variation: $2.07 \times 10^8 \text{ min}^2$



August 8, 2016

Total delay: 2.715×10^4 min
Total variation: $0.89 \times 10^8 \text{ min}^2$



Delay network visualization courtesy of Karthik Gopalakrishnan

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Moments of $Q(\mathbf{x}) = \mathbf{x}^\top \mathcal{L} \mathbf{x}$

Theorem (Theorem 3.2b.2 of Mathai & Provost (1992))

For $\mathbf{x} \sim \mathcal{N}(\boldsymbol{\mu}, \Sigma)$ with valid $m \times m$ covariance matrix Σ , denote by $Q(\mathbf{x}) = \mathbf{x}^\top \mathbf{A} \mathbf{x}$ where \mathbf{A} is a $m \times m$ symmetric, real matrix. The r^{th} moment of $Q(\mathbf{x})$ is given by

$$\mathbb{E}[Q(\mathbf{x})^r] = \sum_{r_1=0}^{r-1} \binom{r-1}{r_1} g(r-1-r_1) \sum_{r_2=0}^{r_1-1} \binom{r_1-1}{r_2} g(r_1-1-r_2) \dots$$

where $g(k) = 2^k k! (\text{tr}(\mathbf{A}\Sigma)^{k+1} + (k+1)\boldsymbol{\mu}^\top (\mathbf{A}\Sigma)^k \mathbf{A} \boldsymbol{\mu})$ for $k \in \mathbb{N}_{\geq 0}$.



Moments of $Q(\mathbf{x}) = \mathbf{x}^\top \mathcal{L} \mathbf{x}$

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where $g(k) = 2^k k! (\text{tr}(\mathbf{A}\Sigma)^{k+1} + (k+1)\boldsymbol{\mu}^\top (\mathbf{A}\Sigma)^k \mathbf{A} \boldsymbol{\mu})$ for $k \in \mathbb{N}_{\geq 0}$.

Of practical interest: **Mean** and **variance**

$$\begin{aligned}\mathbb{E}[Q(\mathbf{x})] &= \text{tr}(\mathbf{A}\Sigma) + \boldsymbol{\mu}^\top \mathbf{A} \boldsymbol{\mu}, \\ \text{Var}[Q(\mathbf{x})] &= 2 \text{tr}(\mathbf{A}\Sigma)^2 + 4\boldsymbol{\mu}^\top \mathbf{A} \Sigma \mathbf{A} \boldsymbol{\mu}.\end{aligned}$$

Yay! But what about the density for $Q(\mathbf{x})$?



$Q(x)$'s messy density

- Only series expansions, unfortunately. No general closed form ...
- *Expansions*: Power series; Laguerre series; Central χ^2 densities; Confluent Hypergeometric functions; Zonal polynomials; Densities of Gamma variates



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- *Expansions*: Power series; **Laguerre series**; Central χ^2 densities; Confluent Hypergeometric functions; Zonal polynomials; Densities of Gamma variates
- (Generalized) Laguerre polynomials $L_k^{(\alpha)}(x)$, Rodrigues formula:

$$L_k^{(\alpha)}(x) = \frac{1}{x!} e^x x^{-\alpha} \left(\frac{d^k}{dx^k} e^{-x} x^{k+\alpha} \right)$$

with $\alpha > -1$ and $k = 0, 1, \dots$

- And the density of $Q(x)$ is ...



$Q(x)$'s messy density

... not very fun to look at, or easy to poke around:

$$f_{Q(x)}(\lambda; \mathbf{b}; q) = \sum_{k=0}^{\infty} c_k^{\lambda, \mathbf{b}} \frac{k!}{2\beta \Gamma(\frac{m}{2} + k)} \left(\frac{q}{2\beta}\right)^{\frac{m}{2}-1} e^{-\frac{q}{2\beta}} L_k^{(\frac{m}{2}-1)}\left(\frac{q}{2\beta}\right),$$

for $q \in (0, \infty)$.



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for $q \in (0, \infty)$.

- β is an arbitrary positive constant
- $c_k^{\boldsymbol{\lambda}, \mathbf{b}}$ are power series expansion coefficients with $c_0 = 1$, dependent on $\boldsymbol{\lambda} = (\lambda_1, \dots, \lambda_m)$ and $\mathbf{b} = (b_1, \dots, b_m)$
- $\text{diag}(\boldsymbol{\lambda}) = \mathbf{P}^\top (\boldsymbol{\Sigma}^{1/2} \mathbf{A} \boldsymbol{\Sigma}^{1/2}) \mathbf{P}$ for an orthogonal \mathbf{P}
- Centrality parameter $\mathbf{b}^\top = \mathbf{P}^\top \boldsymbol{\Sigma}^{-1/2} \boldsymbol{\mu}$



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Saddlepoint approximations to the rescue (?)

- Unknown (univariate) density $f_X(x)$, **known** moment and cumulant generating functions $M_X(s)$, $K_X(s) = \ln M_X(s)$
- **Saddlepoint approximation** (Daniels, 1954) provide extremely accurate, closed-form approximation to $f_X(x)$, has advantages over:
 - Enumerating exact probabilities (\Rightarrow intractability issues)
 - Normal density approximation (\Rightarrow may be inaccurate)
 - Brute force simulation with kernel density estimation (\Rightarrow time consuming, KDE may be inaccurate)



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 - Enumerating exact probabilities (\Rightarrow intractability issues)
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 - Brute force simulation with kernel density estimation (\Rightarrow time consuming, KDE may be inaccurate)
- The **saddlepoint equation** $\hat{f}_X(x)$ approximates $f_X(x)$ on its support \mathcal{X} , given the **saddlepoint** $\hat{s} \triangleq \hat{s}(x)$ associated with $x \in \mathcal{X}$, where \hat{s} is the solution to $dK(\hat{s})/ds = x$.

$$\hat{f}_X(x) = \frac{1}{\sqrt{2\pi \frac{d^2}{ds^2} K_X(\hat{s})}} \exp(K(\hat{s}) - \hat{s}x).$$



The MGF for $Q(\mathbf{x})$

Theorem (Theorems 3.2a.1, 3.2a.2, and Corollary 3.2a.1 of Mathai & Provost (1992))

Let A be a real, symmetric $m \times m$ matrix, and $\mathbf{x} \in \mathbb{R}^{m \times 1}$ with $\mathbf{x} \sim \mathcal{N}(\boldsymbol{\mu}, \Sigma)$. The MGF $M_{Q(\mathbf{x})}(s)$ of $Q(\mathbf{x}) = \mathbf{x}^\top A \mathbf{x}$ can be written in a scalar form involving the eigenvalues of $\Sigma^{1/2} A \Sigma^{1/2}$ and constants that depend on the mean $\boldsymbol{\mu}$.

Specifically, let $\lambda_1, \dots, \lambda_m$ be eigenvalues of $\Sigma^{1/2} A \Sigma^{1/2}$, and define the vector of constants $\mathbf{b} = (b_1, \dots, b_m)^\top = P^\top \Sigma^{-1/2} \boldsymbol{\mu}$, where P is any $m \times m$ orthogonal matrix that diagonalizes $\Sigma^{1/2} A \Sigma^{1/2}$. Then, $M_{Q(\mathbf{x})}(s)$ can be rewritten as follows:

$$M_{Q(\mathbf{x})}(s) = \begin{cases} \exp\left(s \sum_{j=1}^m \frac{b_j^2 \lambda_j}{1 - 2s\lambda_j}\right) \prod_{j=1}^m (1 - 2s\lambda_j)^{-\frac{1}{2}}, & \text{if } \boldsymbol{\mu} \neq 0, \\ \prod_{j=1}^m (1 - 2s\lambda_j)^{-\frac{1}{2}}, & \text{if } \boldsymbol{\mu} = 0. \end{cases}$$

Saddlepoint approximation ingredients

Start with central case $\mu = 0$, we have:

$$K_Q(s) = \sum_{j=1}^m \ln \left((1 - 2s\lambda_j)^{-\frac{1}{2}} \right),$$

$$\frac{d}{ds} K_Q(s) = \sum_{j=1}^m \frac{\lambda_j}{1 - 2s\lambda_j},$$

$$\frac{d^2}{ds^2} K_Q(s) = \sum_{j=1}^m \frac{2\lambda_j^2}{(1 - 2s\lambda_j)^2}.$$

Problem: $m = 1$ and $m = 2$ are fine, $m = 3$ gets sketchy, $m \geq 4$ untenable ...

The problem is **solving for the saddlepoint** $\frac{d}{ds} K(\hat{s}) = x$.

⇒ Possibly tenable (visually) in Maple via implicit function evaluation
(Butler, 2007)



Saddlepoint approximation ingredients

Non-central case is even worse in terms of solving for the saddlepoint.

First derive the ingredients:

$$K_Q(s) = s \sum_{j=1}^m \frac{b_j^2 \lambda_j}{1 - 2s\lambda_j} + \sum_{j=1}^m \ln \left((1 - 2s\lambda_j)^{-1/2} \right),$$

$$\frac{d}{ds} K_Q(s) = \sum_{j=1}^m \frac{\lambda_j (1 + b_j^2 - 2s\lambda_j)}{(1 - 2s\lambda_j)^2},$$

$$\frac{d^2}{ds^2} K_Q(s) = 2 \sum_{j=1}^m \frac{\lambda_j^2 (-1 - 2b_j^2 + 2s\lambda_j)}{(-1 + 2s\lambda_j)^2}.$$

Problem: $m = 1$ is “fine”, $m \geq 2$ untenable ...

Let's look back at the central case.



(Central case) Saddlepoints for $m = 1, 2$

Recall that $\lambda_1, \dots, \lambda_m$ are eigenvalues of $\Sigma^{1/2} A \Sigma^{1/2}$ (**known**)

- For $m = 1$, the saddlepoint is

$$\hat{s} = \frac{1}{2} \left(\frac{1}{\lambda_1} - \frac{1}{x} \right)$$

- For $m = 2$, the saddlepoint(s) is

$$\hat{s} = \frac{1}{4\lambda_1\lambda_2x} \left(\pm \sqrt{\lambda_1^2x^2 - 2\lambda_1\lambda_2x^2 + \lambda_2^2(4\lambda_1^2 + x^2)} + \lambda_2x + \lambda_1(x - 2\lambda_2) \right)$$

- For $m = 3 \dots$



(Central case) Saddlepoint for $m = 3$ $\hat{s} =$

$$\frac{1}{12 x \lambda_1 \lambda_2 \lambda_3} \left(2 (x \lambda_1 \lambda_2 - 3 \lambda_1 \lambda_2 \lambda_3 + x (\lambda_1 + \lambda_2) \lambda_3) + \right. \\
\left. (2 \times 2^{1/3} (9 \lambda_1^2 \lambda_2^2 \lambda_3^3 + x^2 (\lambda_2^2 \lambda_3^2 - \lambda_1 \lambda_2 \lambda_3 (\lambda_2 + \lambda_3) + \lambda_1^2 (\lambda_2^2 - \lambda_2 \lambda_3 + \lambda_3^2)))) \right) / \\
\left(2 x^3 \lambda_2^3 \lambda_3^3 - 3 x^3 \lambda_1 \lambda_2^2 \lambda_3^2 (\lambda_2 + \lambda_3) - 3 x^3 \lambda_1^2 \lambda_2 \lambda_3 (\lambda_2^2 - 4 \lambda_2 \lambda_3 + \lambda_3^2) + \right. \\
\left. \lambda_1^3 (-54 \lambda_2^3 \lambda_3^3 + x^3 (\lambda_2 - 2 \lambda_3) (2 \lambda_2 - \lambda_3) (\lambda_2 + \lambda_3)) + \right. \\
\left. \frac{1}{64} \sqrt{(4096 (-54 \lambda_1^3 \lambda_2^3 \lambda_3^3 + x^3 (\lambda_1 (\lambda_2 - 2 \lambda_3) + \lambda_2 \lambda_3) (2 \lambda_1 \lambda_2 - (\lambda_1 + \lambda_2) \lambda_3) (-2 \lambda_2 \lambda_3 + \lambda_1 (\lambda_2 + \lambda_3)))^2 + \right.} \\
\left. 4 (-16 (x \lambda_1 \lambda_2 - 3 \lambda_1 \lambda_2 \lambda_3 + x (\lambda_1 + \lambda_2) \lambda_3)^2 + \right. \\
\left. 48 x \lambda_1 \lambda_2 \lambda_3 (x (\lambda_1 + \lambda_2 + \lambda_3) - 2 (\lambda_2 \lambda_3 + \lambda_1 (\lambda_2 + \lambda_3))))^3 \right)^{1/3} + \\
2^{2/3} \left(2 x^3 \lambda_2^3 \lambda_3^3 - 3 x^3 \lambda_1 \lambda_2^2 \lambda_3^2 (\lambda_2 + \lambda_3) - 3 x^3 \lambda_1^2 \lambda_2 \lambda_3 (\lambda_2^2 - 4 \lambda_2 \lambda_3 + \lambda_3^2) + \right. \\
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\left. 48 x \lambda_1 \lambda_2 \lambda_3 (x (\lambda_1 + \lambda_2 + \lambda_3) - 2 (\lambda_2 \lambda_3 + \lambda_1 (\lambda_2 + \lambda_3))))^3 \right)^{1/3} \Big) \\$$

Reduce $\left[\sum_{j=1}^3 \frac{\lambda_j}{1-2s\lambda_j} == x, t, \text{Cubics} \rightarrow \text{True}\right]$



(Non-central case) Saddlepoint for $m = 1$

Recall that $\lambda_1, \dots, \lambda_m$ are eigenvalues of $\Sigma^{1/2} A \Sigma^{1/2}$ and b_1, \dots, b_m elements of row vector $\Sigma^{-1/2} \mu$ (**both known**),

- For $m = 1$, the saddlepoint is

$$\hat{s} = \frac{\pm \sqrt{4b_1^2 \lambda_1^3 x + \lambda_1^4 - \lambda_1^2} + 2\lambda_1 x}{4\lambda_1^2 x}$$

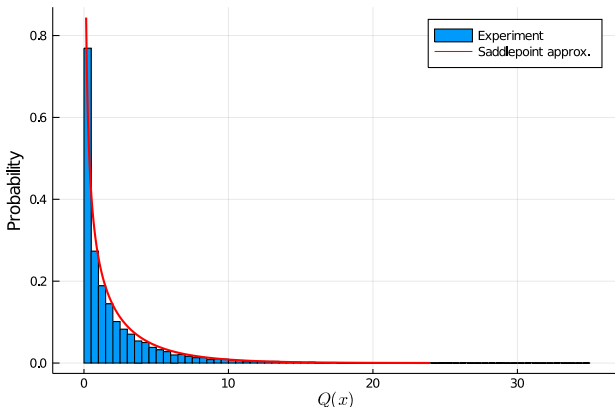
- For $m = 2$, the saddlepoint(s) **does not fit my screen**
- For $m = 3$, no thanks.



Saddlepoint approximation ($m = 2$)

Implemented in Julia

2 node, central, with $\Sigma = I_{2 \times 2}$



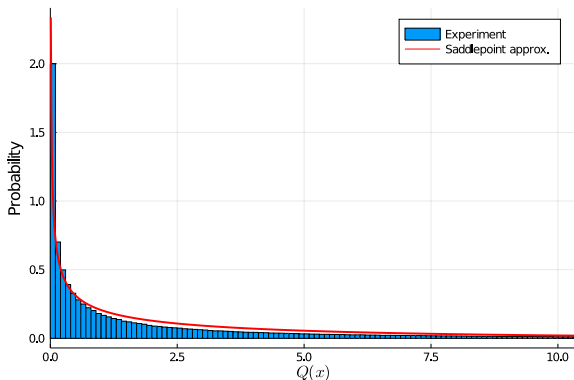
Approximation also decent for random $\Sigma = \text{randn}(2, 2)$,
then $\Sigma = \Sigma^T \Sigma$.



Saddlepoint approximation ($m = 2$), “randomized”

2 node, central, with $\Sigma = \text{randn}(2, 2)$, then $\Sigma = \Sigma^T \Sigma$

Repeat T times (including drawing $Q(x)$ samples), take *average* eigenvalues of $\Sigma^{1/2} \mathcal{L} \Sigma^{1/2}$, fit **one** saddlepoint approx. to entire histogram over T trials ...



($T = 1000$ trials)



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Numerical attempt at other expansions

- *Expansions:* Power series; **Laguerre series**; **Central χ^2 densities**; Confluent Hypergeometric functions; Zonal polynomials; Densities of Gamma variates

$$\text{Density: } f_{Q(x)}(\lambda; \mathbf{b}; q) = \sum_{k=0}^{\infty} \frac{c_k^{\lambda, \mathbf{b}}}{\beta} \underbrace{f_{\chi^2}\left(m + 2k; \frac{q}{\beta}\right)}_{\chi^2 \text{ density}}$$

Expansion constant $c_k^{\lambda, \mathbf{b}}$ computed from:

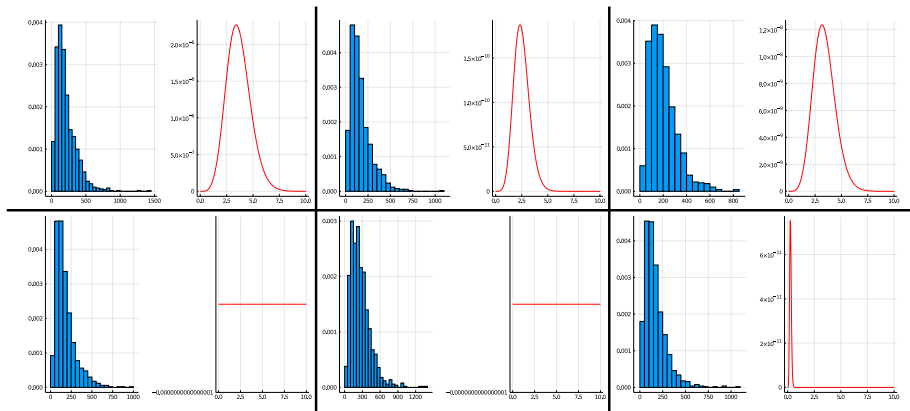
$$c_k^{\lambda, \mathbf{b}} = \begin{cases} \exp\left(-\frac{1}{2} \sum_{j=1}^m b_j^2\right) \prod_{j=1}^m \left(\frac{\beta}{\lambda_j}\right)^{1/2}, & \text{if } k = 0, \\ \frac{1}{2k} \sum_{r=0}^{k-1} \mathbf{d}_{k-r} c_r, & \text{if } k = 1, 2, \dots \end{cases}$$

$$\mathbf{d}_{k-r} = \sum_{j=1}^m \left(1 - \frac{\beta}{\lambda_j}\right)^{k-r} + \beta(k-r) \sum_{j=1}^m \left(\frac{b_j^2}{\lambda_j}\right) \left(1 - \frac{\beta}{\lambda_j}\right)^{k-r-1}$$



χ^2 density expansion

$m = 10$ node case (cycle), non-central ($\mu = \text{randn}(m)$), $\Sigma = \text{randn}(m, m)$, then $\Sigma = \Sigma^T \Sigma$, expansion $k = 10$, take $\beta = \min_{j=1, \dots, m} \lambda_j + \epsilon$



For smaller m , seems to do ... “better”, but the scaling is usually off on the value $Q(x)$. Numerical issue? My abilities at coding?
Jury is still out!



Numerical attempt at other expansions

- *Expansions*: Power series; **Laguerre series**; **Central χ^2 densities**; Confluent Hypergeometric functions; Zonal polynomials; Densities of Gamma variates

$$f_{Q(x)}(\lambda; \mathbf{b}; q) = \sum_{k=0}^{\infty} c_k^{\lambda, \mathbf{b}} \frac{k!}{2\beta \Gamma\left(\frac{m}{2} + k\right)} \left(\frac{q}{2\beta}\right)^{\frac{m}{2}-1} e^{-\frac{q}{2\beta}} L_k^{\left(\frac{m}{2}-1\right)}\left(\frac{q}{2\beta}\right)$$

Expansion constant $c_k^{\lambda, \mathbf{b}}$ (slightly different from χ^2 expansion) computed from:

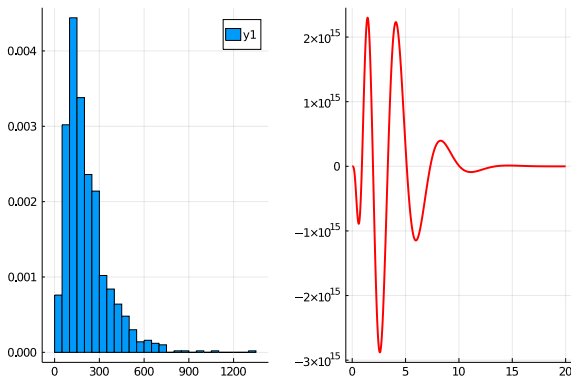
$$c_k^{\lambda, \mathbf{b}} = \begin{cases} \exp\left(-\frac{1}{2} \sum_{j=1}^m b_j^2\right) \prod_{j=1}^m (2\lambda_j)^{-1/2}, & \text{if } k = 0, \\ \frac{1}{k} \sum_{r=0}^{k-1} d_{k-r} c_r, & \text{if } k = 1, 2, \dots \end{cases}$$

$$d_{k-r} = \frac{1}{2} \sum_{j=1}^m \left(1 - \frac{\lambda_j}{\beta}\right)^{k-r} - \frac{k}{2\beta} \sum_{j=1}^m \lambda_j b_j^2 \left(1 - \frac{\lambda_j}{\beta}\right)^{k-r-1}.$$



Laguerre series expansion

$m = 10$ node case (cycle), non-central ($\mu = \text{randn}(m)$), $\Sigma = \text{randn}(m, m)$, then $\Sigma = \Sigma^T \Sigma$, expansion $k = 10$, take $\beta = \min_{j=1, \dots, m} \lambda_j + \epsilon$



A little disappointing, but the debugging continues ...



Thank you for a great semester!

