

# Eigenvalues of Random Matrices under Finite-Rank Perturbations

18.338 Project Presentation

---

Xiaomin Li <sup>1</sup>   Yi Tian <sup>2</sup>

<sup>1</sup>School of Engineering and Applied Sciences, Harvard University

<sup>2</sup>EECS Department, MIT

# Random matrices under finite-rank perturbations

- Let  $X_n$  be an  $n \times n$  Hermite ensemble (Wigner, Wishart, Jacobi...).
- For a fixed  $r \geq 1$ , let  $\theta_1 \geq \dots \geq \theta_s > 0 > \theta_{s+1} \geq \dots \geq \theta_r$  be deterministic nonzero real numbers.
- Let  $P_n$  be an  $n \times n$  Hermite ensemble that has rank  $r$  and  $\theta_1, \dots, \theta_r$  as its nonzero eigenvalues.
- $X_n$  and  $P_n$  are independent.
- How are the eigenvalues of  $X_n + P_n$  and  $X_n(I_n + P_n)$  distributed as  $n \rightarrow \infty$ ?

Benaych-Georges, Florent, and Raj Rao Nadakuditi. “The eigenvalues and eigenvectors of finite, low rank perturbations of large random matrices.” *Advances in Mathematics* 227.1 (2011): 494-521.

- Further assume that either  $X_n$  or  $P_n$  is unitarily invariant.
- Also studied the eigenvectors.
- Phase transition: generalization of [BBAP05].

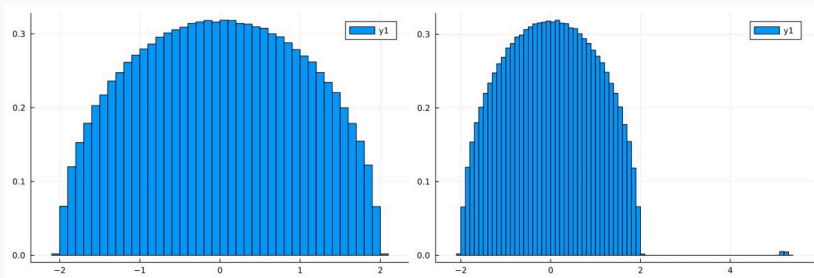
### Theorem (informal, largest eigenvalues)

Let  $\tilde{X}_n$  be either  $X_n + P_n$  or  $X_n(I_n + P_n)$ . Let  $b$  be the supremum of the support of the limiting eigenvalue distribution of  $X_n$ . For each  $1 \leq i \leq s$ , there exists a threshold  $\theta_c$  and some function  $f$  such that

$$\lambda_i(\tilde{X}_n) \xrightarrow{\text{a.s.}} \begin{cases} f(\theta_i) & \text{if } \theta_i > \theta_c \\ b & \text{otherwise,} \end{cases}$$

while for fixed  $i > s$ ,  $\lambda_i(\tilde{X}_n) \xrightarrow{\text{a.s.}} b$ , as  $n \rightarrow \infty$ .

# Experiments



**Figure 1:** Histograms of the eigenvalues of real Hermite ensembles  $X_n$  perturbed by additive  $\theta u_n u_n^\top$ . Left:  $\theta = 0.5$ . Right:  $\theta = 5$ .

- Confirm the known theoretical results with experiments.

- Confirm the known theoretical results with experiments.
- Understand the proof techniques in [BGN11].
- Study a new question: perturbing random matrices whose eigenvalue distributions have multiple bulks.

## Main idea in the analysis in [BGN11] (1/3)

By unitary invariance, w.l.o.g., suppose  $X_n = \text{diag}(\lambda_1, \dots, \lambda_n)$ .

Consider rank-2 perturbation  $P_n = \theta_1 u_n u_n^\top + \theta_2 v_n v_n^\top$ .

## Main idea in the analysis in [BGN11] (1/3)

By unitary invariance, w.l.o.g., suppose  $X_n = \text{diag}(\lambda_1, \dots, \lambda_n)$ .

Consider rank-2 perturbation  $P_n = \theta_1 u_n u_n^\top + \theta_2 v_n v_n^\top$ .

The eigenvalues of  $X_n + P_n$  are solutions to the equation

$$\det(zI - (X_n + P_n)) = 0.$$



## Main idea in the analysis in [BGN11] (1/3)

By unitary invariance, w.l.o.g., suppose  $X_n = \text{diag}(\lambda_1, \dots, \lambda_n)$ .

Consider rank-2 perturbation  $P_n = \theta_1 u_n u_n^\top + \theta_2 v_n v_n^\top$ .

The eigenvalues of  $X_n + P_n$  are solutions to the equation

$$\det(zI - (X_n + P_n)) = 0.$$

If  $zI - X_n$  is invertible, we have

$$zI - (X_n + P_n) = (zI - X_n)(I - (zI - X_n)^{-1}P_n).$$

$$\det(zI - (X_n + P_n)) = \det(zI - X_n) \cdot \det(I - (zI - X_n)^{-1}P_n).$$

## Main idea in the analysis in [BGN11] (1/3)

By unitary invariance, w.l.o.g., suppose  $X_n = \text{diag}(\lambda_1, \dots, \lambda_n)$ .

Consider rank-2 perturbation  $P_n = \theta_1 u_n u_n^\top + \theta_2 v_n v_n^\top$ .

The eigenvalues of  $X_n + P_n$  are solutions to the equation

$$\det(zI - (X_n + P_n)) = 0.$$

If  $zI - X_n$  is invertible, we have

$$zI - (X_n + P_n) = (zI - X_n)(I - (zI - X_n)^{-1}P_n).$$

$$\det(zI - (X_n + P_n)) = \det(zI - X_n) \cdot \det(I - (zI - X_n)^{-1}P_n).$$

This means  $z$  is an eigenvalue of  $X_n + P_n$  and not an eigenvalue of  $X_n$  if and only if  $\det(I - (zI - X_n)^{-1}P_n) = 0$ .

## Main idea in the analysis in [BGN11] (2/3)

$$\begin{aligned} & \det(I - (zI - X_n)^{-1}P_n) \\ &= \det(I - (zI - X_n)^{-1}(\theta_1 u_n u_n^\top + \theta_2 v_n v_n^\top)) \\ &\stackrel{(i)}{=} \det(I - [u_n, v_n]^\top \text{diag}((z - \lambda_1)^{-1}, \dots, (z - \lambda_n)^{-1})[\theta_1 u_n, \theta_2 v_n]) \\ &= \det \left( \begin{bmatrix} 1 - \theta_1 \sum_{i=1}^n u_i^2 (z - \lambda_i)^{-1} & \theta_2 \sum_{i=1}^n u_i v_i (z - \lambda_i)^{-1} \\ \theta_1 \sum_{i=1}^n u_i v_i (z - \lambda_i)^{-1} & 1 - \theta_2 \sum_{i=1}^n v_i^2 (z - \lambda_i)^{-1} \end{bmatrix} \right), \end{aligned}$$

where (i) is due to Sylvester's determinant identity.

## Main idea in the analysis in [BGN11] (2/3)

$$\begin{aligned} & \det(I - (zI - X_n)^{-1}P_n) \\ &= \det(I - (zI - X_n)^{-1}(\theta_1 u_n u_n^\top + \theta_2 v_n v_n^\top)) \\ &\stackrel{(i)}{=} \det(I - [u_n, v_n]^\top \text{diag}((z - \lambda_1)^{-1}, \dots, (z - \lambda_n)^{-1})[\theta_1 u_n, \theta_2 v_n]) \\ &= \det \left( \begin{bmatrix} 1 - \theta_1 \sum_{i=1}^n u_i^2 (z - \lambda_i)^{-1} & \theta_2 \sum_{i=1}^n u_i v_i (z - \lambda_i)^{-1} \\ \theta_1 \sum_{i=1}^n u_i v_i (z - \lambda_i)^{-1} & 1 - \theta_2 \sum_{i=1}^n v_i^2 (z - \lambda_i)^{-1} \end{bmatrix} \right), \end{aligned}$$

where (i) is due to Sylvester's determinant identity. In the limit,

$$\sum_{i=1}^n u_i v_i (z - \lambda_i)^{-1} \xrightarrow{\text{a.s.}} 0, \quad \sum_{i=1}^n u_i^2 (z - \lambda_i)^{-1} \xrightarrow{\text{a.s.}} n^{-1} \sum_{i=1}^n (z - \lambda_i)^{-1}.$$

## Main idea in the analysis in [BGN11] (2/3)

$$\begin{aligned} & \det(I - (ZI - X_n)^{-1}P_n) \\ &= \det(I - (ZI - X_n)^{-1}(\theta_1 u_n u_n^\top + \theta_2 v_n v_n^\top)) \\ &\stackrel{(i)}{=} \det(I - [u_n, v_n]^\top \text{diag}((z - \lambda_1)^{-1}, \dots, (z - \lambda_n)^{-1})[\theta_1 u_n, \theta_2 v_n]) \\ &= \det \left( \begin{bmatrix} 1 - \theta_1 \sum_{i=1}^n u_i^2 (z - \lambda_i)^{-1} & \theta_2 \sum_{i=1}^n u_i v_i (z - \lambda_i)^{-1} \\ \theta_1 \sum_{i=1}^n u_i v_i (z - \lambda_i)^{-1} & 1 - \theta_2 \sum_{i=1}^n v_i^2 (z - \lambda_i)^{-1} \end{bmatrix} \right), \end{aligned}$$

where (i) is due to Sylvester's determinant identity. In the limit,

$$\sum_{i=1}^n u_i v_i (z - \lambda_i)^{-1} \xrightarrow{\text{a.s.}} 0, \quad \sum_{i=1}^n u_i^2 (z - \lambda_i)^{-1} \xrightarrow{\text{a.s.}} n^{-1} \sum_{i=1}^n (z - \lambda_i)^{-1}.$$

Then,  $\det(I - (ZI - X_n)^{-1}P_n) = 0$  if and only if

$$1 - \theta_1 n^{-1} \sum_{i=1}^n (z - \lambda_i)^{-1} = 0, \text{ or } 1 - \theta_2 n^{-1} \sum_{i=1}^n (z - \lambda_i)^{-1} = 0.$$

## Main idea in the analysis in [BGN11] (3/3)

$z$  is an eigenvalue of  $X_n + P_n$  and not an eigenvalue of  $X_n$

$$\iff \det(I - (zI - X_n)^{-1}P_n) = 0$$

$$\iff 1 - \theta_1 n^{-1} \sum_{i=1}^n (z - \lambda_i)^{-1} = 0, \text{ or } 1 - \theta_2 n^{-1} \sum_{i=1}^n (z - \lambda_i)^{-1} = 0.$$

## Main idea in the analysis in [BGN11] (3/3)

$z$  is an eigenvalue of  $X_n + P_n$  and not an eigenvalue of  $X_n$

$$\iff \det(I - (zI - X_n)^{-1}P_n) = 0$$

$$\iff 1 - \theta_1 n^{-1} \sum_{i=1}^n (z - \lambda_i)^{-1} = 0, \text{ or } 1 - \theta_2 n^{-1} \sum_{i=1}^n (z - \lambda_i)^{-1} = 0.$$

Let  $\mu_{X_n} := n^{-1} \sum_{i=1}^n \delta_{\lambda_i(X_n)}$ . Then, its Cauchy transform is precisely

$$G_{\mu_{X_n}}(z) = \int (z - t)^{-1} d\mu_{X_n}(t) = n^{-1} \sum_{i=1}^n (z - \lambda_i)^{-1}.$$

## Main idea in the analysis in [BGN11] (3/3)

$z$  is an eigenvalue of  $X_n + P_n$  and not an eigenvalue of  $X_n$

$$\iff \det(I - (zI - X_n)^{-1}P_n) = 0$$

$$\iff 1 - \theta_1 n^{-1} \sum_{i=1}^n (z - \lambda_i)^{-1} = 0, \text{ or } 1 - \theta_2 n^{-1} \sum_{i=1}^n (z - \lambda_i)^{-1} = 0.$$

Let  $\mu_{X_n} := n^{-1} \sum_{i=1}^n \delta_{\lambda_i(X_n)}$ . Then, its Cauchy transform is precisely

$$G_{\mu_{X_n}}(z) = \int (z - t)^{-1} d\mu_{X_n}(t) = n^{-1} \sum_{i=1}^n (z - \lambda_i)^{-1}.$$

In the limit of  $n \rightarrow \infty$ ,  $G_{\mu_{X_n}} \rightarrow G_{\mu_X}$ , where  $\mu_X$  is the limiting eigenvalue distribution of  $X_n$ . Hence, new limiting eigenvalue  $z$  of  $X_n + P_n$  satisfy

$$G_{\mu_X}(z) = \theta_1^{-1}, \text{ or } G_{\mu_X}(z) = \theta_2^{-1}.$$



# Results in the additive case in [BGN11]

## Theorem (informal, largest eigenvalues)

Let  $\tilde{X}_n$  be either  $X_n + P_n$ . Let  $b$  be the supremum of the support of the limiting eigenvalue distribution of  $X_n$ . For each  $1 \leq i \leq s$ ,

$$\lambda_i(\tilde{X}_n) \xrightarrow{\text{a.s.}} \begin{cases} G_{\mu_X}^{-1}(\theta_i^{-1}) & \text{if } \theta_i > (G_{\mu_X}(b^+))^{-1}, \\ b & \text{otherwise,} \end{cases}$$

while for fixed  $i > s$ ,  $\lambda_i(\tilde{X}_n) \xrightarrow{\text{a.s.}} b$ , as  $n \rightarrow \infty$ .

# Results in the additive case in [BGN11]

## Theorem (informal, largest eigenvalues)

Let  $\tilde{X}_n$  be either  $X_n + P_n$ . Let  $b$  be the supremum of the support of the limiting eigenvalue distribution of  $X_n$ . For each  $1 \leq i \leq s$ ,

$$\lambda_i(\tilde{X}_n) \xrightarrow{\text{a.s.}} \begin{cases} G_{\mu_X}^{-1}(\theta_i^{-1}) & \text{if } \theta_i > (G_{\mu_X}(b^+))^{-1}, \\ b & \text{otherwise,} \end{cases}$$

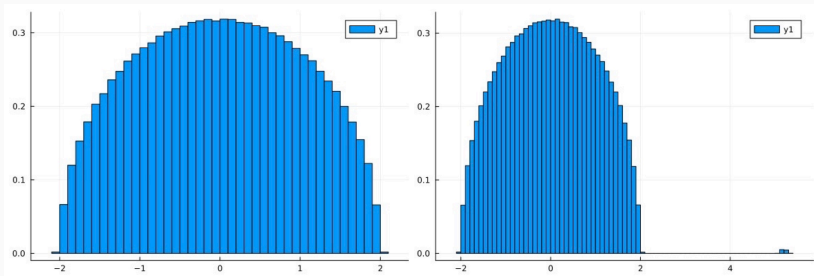
while for fixed  $i > s$ ,  $\lambda_i(\tilde{X}_n) \xrightarrow{\text{a.s.}} b$ , as  $n \rightarrow \infty$ .

For Wigner matrices,  $\mu_X$  is the semicircle distribution, and

$$G_{\mu_X}(z) = \frac{z - \text{sign}(z)\sqrt{z^2 - 4\sigma^2}}{2\sigma^2}.$$

For  $0 < z \leq 2\sigma$ ,  $G_{\mu_X}^{-1}(z) = \sigma^2 z + z^{-1}$ .

# Experiments



**Figure 2:** Histograms of the eigenvalues of real Hermite ensembles  $X_n$  perturbed by additive  $\theta u_n u_n^\top$ . Left:  $\theta = 0.5$  (theory: no spike). Right:  $\theta = 5$  (theory: spike at 5.2).

## Perturbing random matrices with two bulks

Consider  $X_{2n} = \text{diag}(Y_n, Z_n)$  where  $Y_n, Z_n$  are diagonal and two bulks.

Rank-2 perturbation  $P_{2n} = \theta_1 u_{2n} u_{2n}^\top + \theta_2 v_{2n} v_{2n}^\top$ .

## Perturbing random matrices with two bulks

Consider  $X_{2n} = \text{diag}(Y_n, Z_n)$  where  $Y_n, Z_n$  are diagonal and two bulks.

Rank-2 perturbation  $P_{2n} = \theta_1 u_{2n} u_{2n}^\top + \theta_2 v_{2n} v_{2n}^\top$ .

$z$  is a limit eigenvalue of  $X_n + P_n$  and not a limit eigenvalue of  $X_n$  if and only if  $G_{\mu_X}(z) = \theta_1^{-1}$ , or  $G_{\mu_X}(z) = \theta_2^{-1}$ . So, reduces to finding  $G_{\mu_X}$ .

# Perturbing random matrices with two bulks

Consider  $X_{2n} = \text{diag}(Y_n, Z_n)$  where  $Y_n, Z_n$  are diagonal and two bulks.

Rank-2 perturbation  $P_{2n} = \theta_1 u_{2n} u_{2n}^\top + \theta_2 v_{2n} v_{2n}^\top$ .

$z$  is a limit eigenvalue of  $X_n + P_n$  and not a limit eigenvalue of  $X_n$  if and only if  $G_{\mu_X}(z) = \theta_1^{-1}$ , or  $G_{\mu_X}(z) = \theta_2^{-1}$ . So, reduces to finding  $G_{\mu_X}$ .

Let  $\mu_{Y_n} = \frac{1}{n} \sum_{i=1}^n \delta_{\lambda_i(X_{2n})}$  and  $\mu_{Z_n} = \frac{1}{n} \sum_{i=n+1}^{2n} \delta_{\lambda_i(X_{2n})}$  Then,

$$\mu_{X_n} = \frac{1}{2n} \sum_{i=1}^{2n} \delta_{\lambda_i(X_{2n})} = \frac{1}{2} (\mu_{Y_n} + \mu_{Z_n}),$$

# Perturbing random matrices with two bulks

Consider  $X_{2n} = \text{diag}(Y_n, Z_n)$  where  $Y_n, Z_n$  are diagonal and two bulks.

Rank-2 perturbation  $P_{2n} = \theta_1 u_{2n} u_{2n}^\top + \theta_2 v_{2n} v_{2n}^\top$ .

$z$  is a limit eigenvalue of  $X_n + P_n$  and not a limit eigenvalue of  $X_n$  if and only if  $G_{\mu_X}(z) = \theta_1^{-1}$ , or  $G_{\mu_X}(z) = \theta_2^{-1}$ . So, reduces to finding  $G_{\mu_X}$ .

Let  $\mu_{Y_n} = \frac{1}{n} \sum_{i=1}^n \delta_{\lambda_i(X_{2n})}$  and  $\mu_{Z_n} = \frac{1}{n} \sum_{i=n+1}^{2n} \delta_{\lambda_i(X_{2n})}$ . Then,

$$\mu_{X_n} = \frac{1}{2n} \sum_{i=1}^{2n} \delta_{\lambda_i(X_{2n})} = \frac{1}{2}(\mu_{Y_n} + \mu_{Z_n}),$$

Let  $\mu_X, \mu_Y, \mu_Z$  be the limiting distributions of  $\mu_{X_{2n}}, \mu_{Y_n}, \mu_{Z_n}$ . We have

$$G_{\mu_X} = \frac{1}{2}(G_{\mu_Y} + G_{\mu_Z}).$$

## Numerical example: Wigner matrices

Suppose  $\mu_Z, \mu_Y$  are two semicircle distributions in  $[-2, 2], [6, 10]$ . Then

$$G_{\mu_Z}(z) = \frac{z - \text{sign}(z)\sqrt{z^2 - 4}}{2}, \quad z \leq -2, \text{ or } z \geq 2,$$

$$G_{\mu_Y}(z) = \frac{z - 8 - \text{sign}(z - 8)\sqrt{(z - 8)^2 - 4}}{2}, \quad z \leq 6, \text{ or } z \geq 10.$$



## Numerical example: Wigner matrices

Suppose  $\mu_Z, \mu_Y$  are two semicircle distributions in  $[-2, 2], [6, 10]$ . Then

$$G_{\mu_Z}(z) = \frac{z - \text{sign}(z)\sqrt{z^2 - 4}}{2}, \quad z \leq -2, \text{ or } z \geq 2,$$

$$G_{\mu_Y}(z) = \frac{z - 8 - \text{sign}(z - 8)\sqrt{(z - 8)^2 - 4}}{2}, \quad z \leq 6, \text{ or } z \geq 10.$$

Hence,

$$G_{\mu_X}(z) = \frac{2z - 8 - \text{sign}(z)\sqrt{z^2 - 4} - \text{sign}(z - 8)\sqrt{(z - 8)^2 - 4}}{4},$$

where  $z \leq -2, 2 \leq z \leq 6, \text{ or } z \geq 10$ .

## Numerical example: Wigner matrices

Suppose  $\mu_Z, \mu_Y$  are two semicircle distributions in  $[-2, 2], [6, 10]$ . Then

$$G_{\mu_Z}(z) = \frac{z - \text{sign}(z)\sqrt{z^2 - 4}}{2}, \quad z \leq -2, \text{ or } z \geq 2,$$

$$G_{\mu_Y}(z) = \frac{z - 8 - \text{sign}(z - 8)\sqrt{(z - 8)^2 - 4}}{2}, \quad z \leq 6, \text{ or } z \geq 10.$$

Hence,

$$G_{\mu_X}(z) = \frac{2z - 8 - \text{sign}(z)\sqrt{z^2 - 4} - \text{sign}(z - 8)\sqrt{(z - 8)^2 - 4}}{4},$$

where  $z \leq -2, 2 \leq z \leq 6, \text{ or } z \geq 10$ . In particular,

$$G(-2) = -G(10) = \sqrt{6} - 3 \approx -0.55, \quad G(2) = -G(6) = \sqrt{2} - 1 \approx 0.41.$$

## Numerical example: Wigner matrices

Suppose  $\mu_Z, \mu_Y$  are two semicircle distributions in  $[-2, 2], [6, 10]$ . Then

$$G_{\mu_Z}(z) = \frac{z - \text{sign}(z)\sqrt{z^2 - 4}}{2}, \quad z \leq -2, \text{ or } z \geq 2,$$

$$G_{\mu_Y}(z) = \frac{z - 8 - \text{sign}(z - 8)\sqrt{(z - 8)^2 - 4}}{2}, \quad z \leq 6, \text{ or } z \geq 10.$$

Hence,

$$G_{\mu_X}(z) = \frac{2z - 8 - \text{sign}(z)\sqrt{z^2 - 4} - \text{sign}(z - 8)\sqrt{(z - 8)^2 - 4}}{4},$$

where  $z \leq -2, 2 \leq z \leq 6$ , or  $z \geq 10$ . In particular,

$$G(-2) = -G(10) = \sqrt{6} - 3 \approx -0.55, \quad G(2) = -G(6) = \sqrt{2} - 1 \approx 0.41.$$

- $|\theta_i|^{-1} \geq 0.55$ : no new eigenvalue;
- $0.41 \leq |\theta_i|^{-1} < 0.55$ : a spike outside  $[-2, 10]$  but none in  $[2, 6]$ ;
- $|\theta_i|^{-1} < 0.41$ , both a spike outside  $[-2, 10]$  and one in  $[2, 6]$ .

Jupyter notebook



Jinho Baik, Gérard Ben Arous, and Sandrine Péché.

**Phase transition of the largest eigenvalue for nonnull complex sample covariance matrices.**

*The Annals of Probability*, 33(5):1643–1697, 2005.



Florent Benaych-Georges and Raj Rao Nadakuditi.

**The eigenvalues and eigenvectors of finite, low rank perturbations of large random matrices.**

*Advances in Mathematics*, 227(1):494–521, 2011.