UNIVERSIDADE FEDERAL FLUMINENSE

Programa de Mestrado e Doutorado em Engenharia de Produção

Multivariate Data Analysis

Binary Logistic Regression

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Outline

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1. Definition

A regression technique that has a dichotomous dependent variable, and metric or dichotomous independent variables is known as **Binary Logistic Regression**, with the following formulation:

$$Y_i \in \{0; 1\}$$

$$Z_i = ln\left(\frac{p_i}{1 - p_i}\right) = B_0 + B_1 X_{1i} + \dots + B_k X_{kn}$$

Where:

i = Each case of a sample size n;

 Y_i = Dependent Variable Dichotomous (Occurrence = 1 and Non-Occurrence = 0);

 Z_i = Logit;

 p_i = Probability of Occurrence $[\mu(Y) = p_i \ e \ \sigma^2(Y) = p_i \times (1 - p_i)];$

 $1 - p_i$ = Probability of non-occurrence;

 B_0 = Constant;

 B_k = Regression coefficients;

 X_{ki} = Independent Variable k (Predictor k).

The logit, which is a continuous variable, is calculated as the natural logarithm of chance, and chance is defined as the ratio between the occurrence and non-occurrence of an event. For example, a chance 3:1 means that for every 4 events, 3 events occurs and 1 do not.

$$ln\left(\frac{p_i}{1-p_i}\right) = Z_i$$

$$\frac{p_i}{1 - p_i} = e^{(Z_i)}$$

$$chance_{Y_i=1} = e^{Z_i}$$

The output of a logit model is the probability of a case (i) to belong to an occurrence group ($Y_i = 1$) or a non-occurrence group ($Y_i = 0$).

$$\frac{p_i}{1 - p_i} = e^{(Z_i)}$$

$$p_i = \left(\frac{e^{(Z_i)}}{1 + e^{(Z_i)}}\right) = \left(\frac{1}{1 + e^{-(B_0 + B_1 X_{1i} + \dots + B_k X_{kn})}}\right)$$

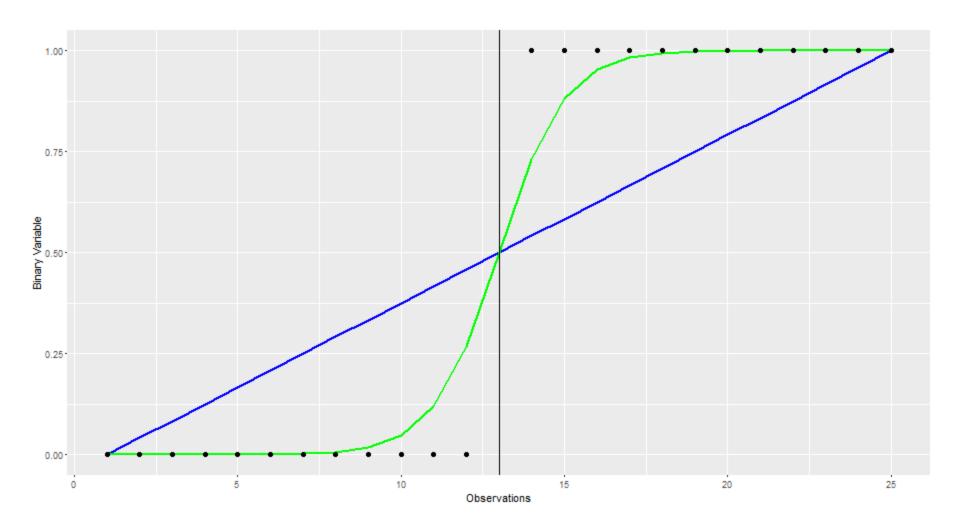
$$1 - p_i = \left(\frac{1}{1 + e^{(Z_i)}}\right) = \left(\frac{1}{1 + e^{(B_0 + B_1 X_{1i} + \dots + B_k X_{kn})}}\right)$$

In order to properly model a dataset in which the dependent variable is non-metric, the multiple linear regression cannot be used, because the assumption of homoscedasticity is violated. This violation is very severe and invalidate the results of the multiple linear regression model.

But when the variables do not meet the assumptions of:

- Normality,
- Linearity,
- Homoscedasticity.

The logit model is the technique of choice, since it does not make these assumptions.



Assumptions

ASSUMPTIONS

- The dependent variable must be dichotomous;
- The independent variables must be metric or dichotomous;
- The Ratio $\frac{n}{k} \ge 10 \to \text{at least } 10 \text{ observations } (n) \text{ for each predictor } (k). The higher the ratio <math>\frac{n}{k}$ better;
- Absence of collinearity or multicollinearity;
- Outliers Verification.

In order to explain a **Logit** approach, the following dataset case study will be used: A high school needs to know if students that go to school by car are more likely to arrive late $(Y_i = 1)$ or not $(Y_i = 0)$ in the classroom. A sample of 100 students was collected and in addition to the indication of, if the student arrived late or not, the following information was also collected:

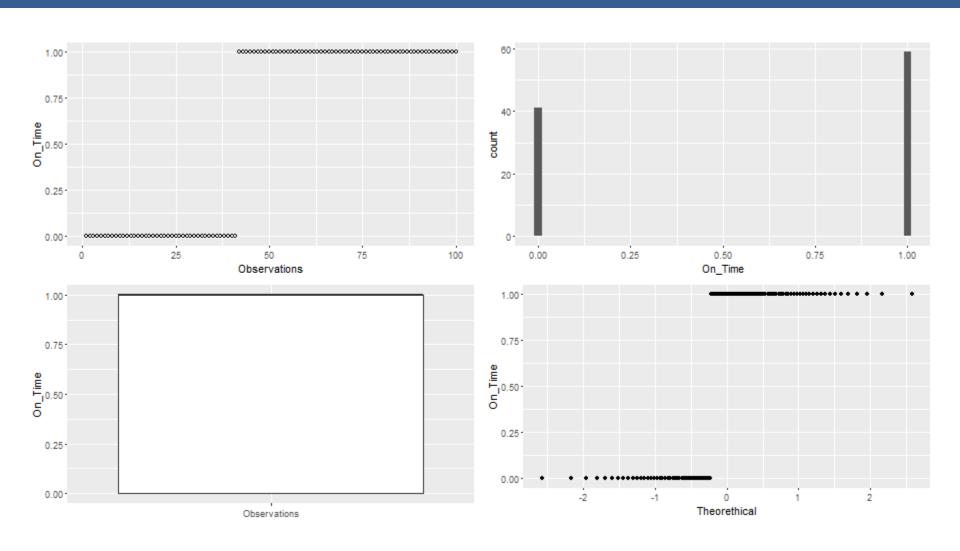
- Distance traveled (km);
- Quantity of traffic lights (discrete variable);
- Period day (categorical variable: Morning or Afternoon*);
- Profile of the driver (categorical variable: Calm*, Moderate or Aggressive).

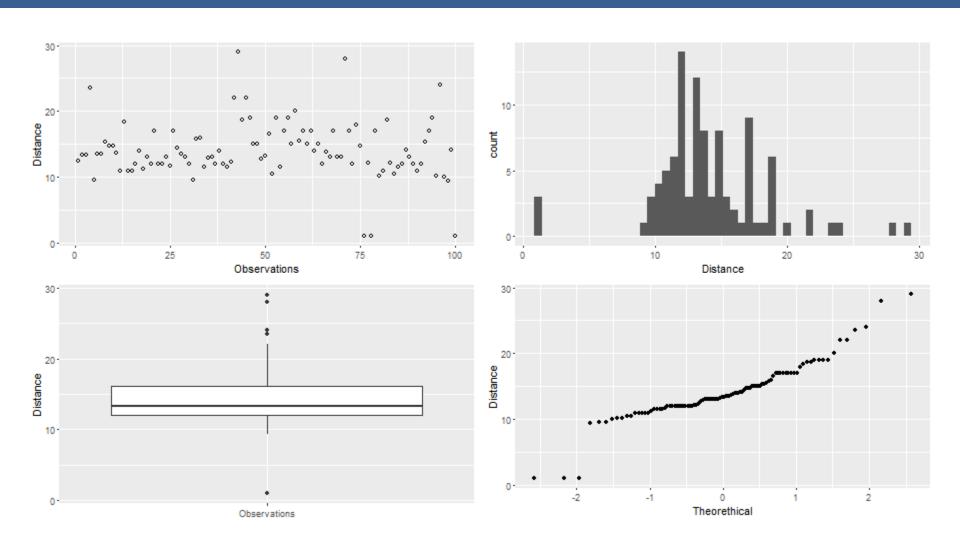
^{*} Reference Category.

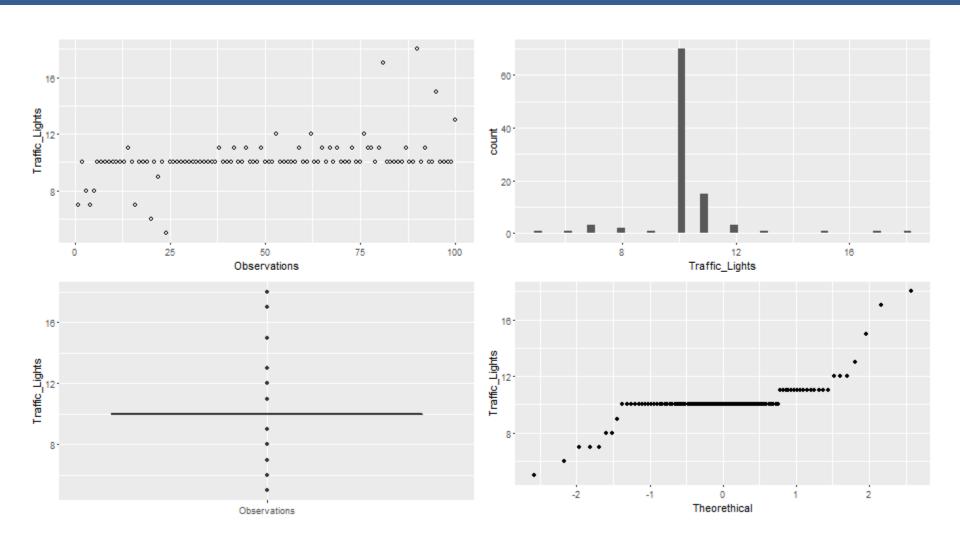
Id	Student	Y (Late?) Yes = 1; No = 0	Distance (X ₁)	Traffic L. (X ₂)	Period (X ₃)	Profile (X ₄)
1	Gabriela	0	12.5	7	Morning	Calm
2	Patrícia	0	13.3	10	Morning	Calm
3	Gustavo	0	13.4	8	Morning	Aggressive
4	Letícia	0	23.5	7	Morning	Calm
5	Luiz Ovídio	0	9.5	8	Morning	Calm
6	Leonor	0	13.5	10	Morning	Calm
7	Dalila	0	13.5	10	Morning	Calm
8	Antônio	0	15.4	10	Morning	Calm
9	Júlia	0	14.7	10	Morning	Calm
10	Mariana	0	14.7	10	Morning	Calm
34	Cintia	0	11.5	10	Afternoon	Calm
99	Leandro	1	14.2	10	Morning	Moderate
100	Estela	1	1	13	Morning	Calm

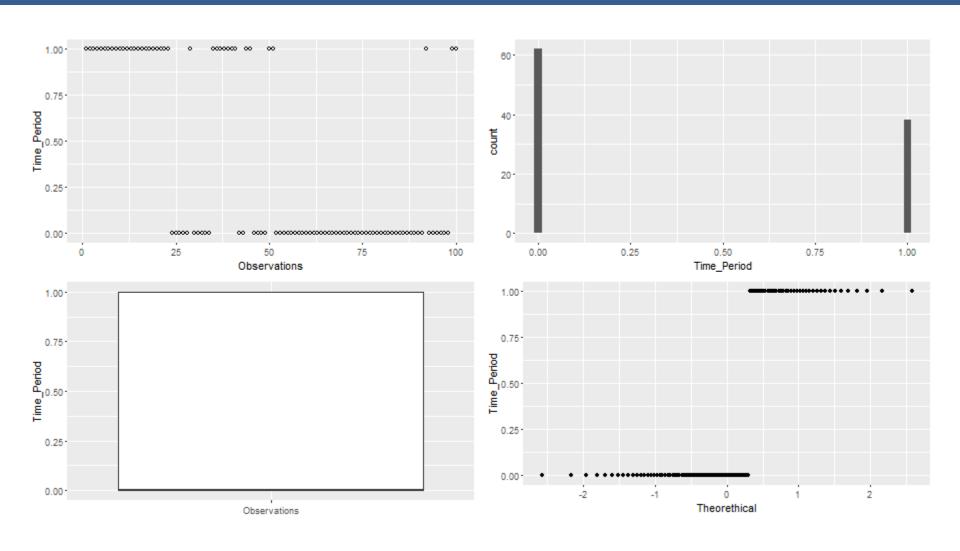
Id	Student	Y (Late?) Yes = 1; No = 0	Distance (X ₁)	Traffic L. (X ₂)	Period (X ₃)	Profile A (X ₄)	Profile B(X ₅)
1	Gabriela	0	12.5	7	1	0	0
2	Patrícia	0	13.3	10	1	0	0
3	Gustavo	0	13.4	8	1	1	0
4	Letícia	0	23.5	7	1	0	0
5	Luiz Ovídio	0	9.5	8	1	0	0
6	Leonor	0	13.5	10	1	0	0
7	Dalila	0	13.5	10	1	0	0
8	Antônio	0	15.4	10	1	0	0
9	Júlia	0	14.7	10	1	0	0
10	Mariana	0	14.7	10	1	0	0
34	Cintia	0	11.5	10	0	0	0
99	Leandro	1	14.2	10	1	0	1
100	Estela	1	1	13	1	0	0

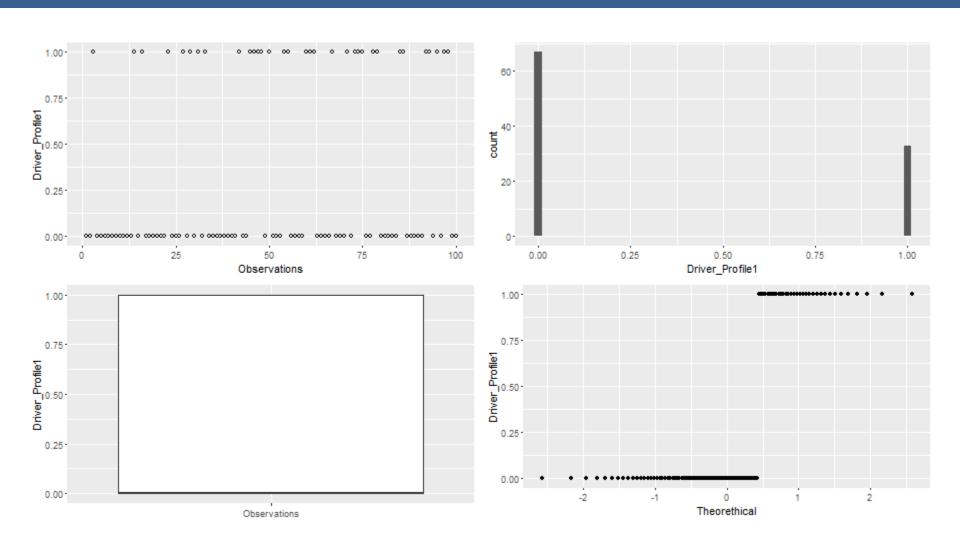
```
library("ggplot2")
ggplot(data = my_data, aes(x = Observations, y = On_Time)) + geom_point(shape = 1)
ggplot(data = my_data, aes(x = "", y = On_Time)) + geom_boxplot() + theme(axis.title.x = element_blank())
ggplot(data = my_data, aes(On_Time)) + geom_histogram(bins = 50)
ggplot(data = my_data, aes(sample = On_Time)) + stat_qq()+ xlab("Theorethical") + ylab("On_Time")
ggplot(data = my_data, aes(x = Observations, y = Distance)) + geom_point(shape = 1)
ggplot(data = my_data, aes(x = "", y = Distance)) + geom_boxplot() + theme(axis.title.x = element_blank())
ggplot(data = my_data, aes(Distance)) + geom_histogram(bins = 50)
ggplot(data = my_data, aes(sample = Distance)) + stat_qq()+ xlab("Theorethical") + ylab("Distance")
ggplot(data = my_data, aes(x = Observations, y = Traffic_Lights)) + geom_point(shape = 1)
ggplot(data = my_data, aes(x = "", y = Traffic_Lights)) + geom_boxplot() + theme(axis.title.x = element_blank())
ggplot(data = my_data, aes(Traffic_Lights)) + geom_histogram(bins = 50)
ggplot(data = my_data, aes(sample = Traffic_Lights)) + stat_qq()+ xlab("Theorethical") + ylab("Traffic_Lights")
ggplot(data = my_data, aes(x = Observations, y = Time_Period)) + geom_point(shape = 1)
ggplot(data = my_data, aes(x = "", y = Time_Period)) + geom_boxplot() + theme(axis.title.x = element_blank())
ggplot(data = my_data, aes(Time_Period)) + geom_histogram(bins = 50)
ggplot(data = my_data, aes(sample = Time_Period)) + stat_qq()+ xlab("Theorethical") + ylab("Time_Period")
qqplot(data = my data, aes(x = Observations, y = Driver Profile1)) + qeom point(shape = 1)
ggplot(data = my_data, aes(x = "", y = Driver_Profile1)) + geom_boxplot() + theme(axis.title.x = element_blank())
ggplot(my_data, aes( Driver_Profile1)) + geom_histogram(bins = 50)
ggplot(data = my_data, aes(sample = Driver_Profile1)) + stat_qq()+ xlab("Theorethical") + ylab("Driver_Profile1")
ggplot(data = my_data, aes(x = Observations, y = Driver_Profile2)) + geom_point(shape = 1)
ggplot(data = my_data, aes(x = "", y = Driver_Profile2)) + geom_boxplot() + theme(axis.title.x = element_blank())
ggplot(my_data, aes( Driver_Profile2)) + geom_histogram(bins = 50)
ggplot(data = my_data, aes(sample = Driver_Profile2)) + stat_qq()+ xlab("Theorethical") + ylab("Driver_Profile2")
```

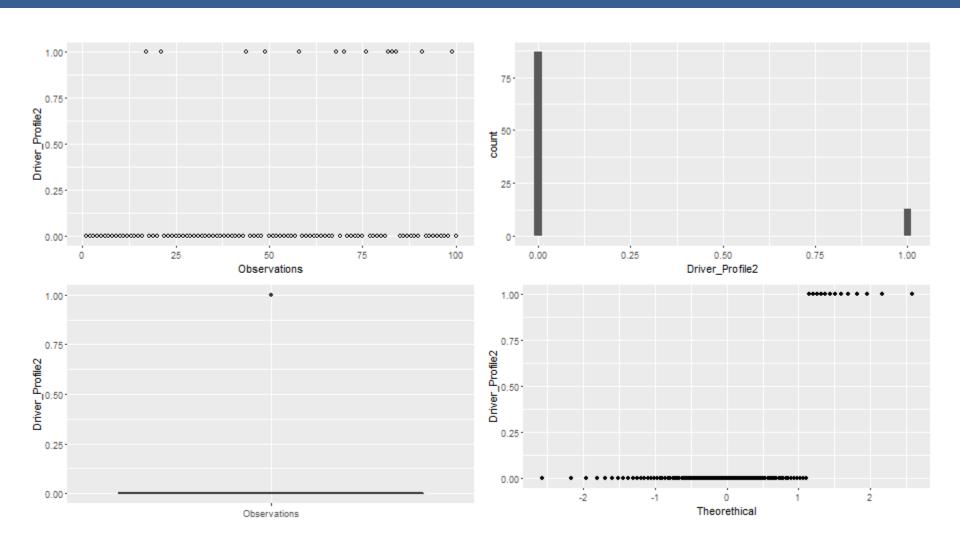




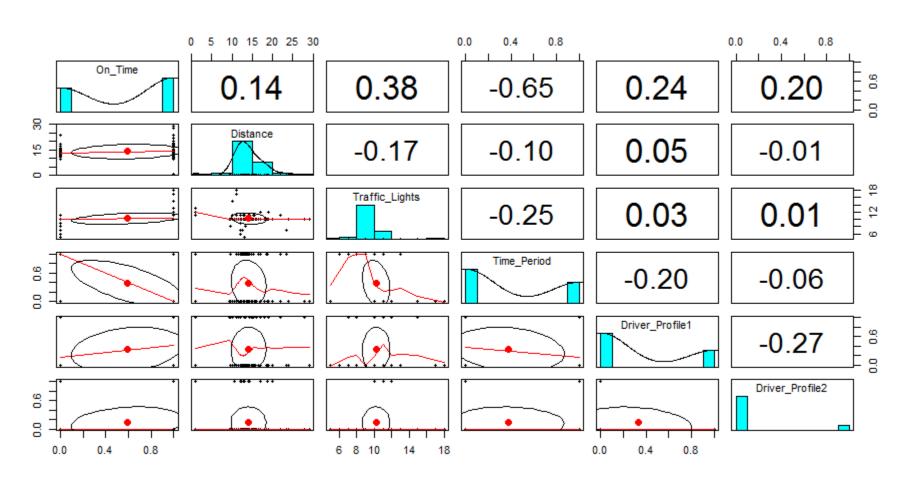








library ("psych")
pairs.panels(my_data)



Binary Logistic Regression

```
# Logit
logit_01 <- glm(On_Time ~ ., data = my_data, family = "binomial")
summary(logit 01)
Deviance Residuals:
  Min
         1Q Median
                        3Q
                               Max
-2.2037 -0.2638 0.1231 0.4419 2.2928
Coefficients:
                            Std. Error
             Estimate
                                            z value
                                                          Pr(>|z|)
              -30.2003
                              9.9806
                                            -3.026
                                                         0.00248 **
(Intercept)
Distance
          0.2202
                              0.1097
                                             2.007
                                                         0.04474 *
Traffic_Lights 2.7667
                                             3.002
                              0.9216
                                                         0.00268 **
                                                         3.18e-05 ***
Time Period -3.6534
                              0.8781
                                            -4.160
                                            1.800
Driver_Profile1 1.3460
                              0.7477
                                                         0.07184.
Driver Profile2 2.9145
                              1.1788
                                             2.472
                                                         0.01342 *
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
(Dispersion parameter for binomial family taken to be 1)
Null deviance: 135.372 on 99 degrees of freedom
```

Residual deviance: 58.131 on 94 degrees of freedom

AIC: 70.131

```
logit_02 <- glm(On_Time ~ Distance + Traffic_Lights + Time_Period + Driver_Profile2, data = my_data, family = "binomial") summary(logit_02)
```

Deviance Residuals:

Min 1Q Median 3Q Max -1.9495 -0.3169 0.1364 0.5692 2.3967

Coefficients:

	Estimate	Std. Error	z value	Pr(> z)
(Intercept)	-30.9333	10.6344	-2.909	0.00363 **
Distance	0.2041	0.1012	2.018	0.04358 *
Traffic_Lights	2.9201	1.0106	2.890	0.00386 **
Time_Period	-3.7763	0.8466	-4.461	8.18e-06 ***
Driver_Profile	2 2.4591	1.1394	2.158	0.03091 *

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

(Dispersion parameter for binomial family taken to be 1)

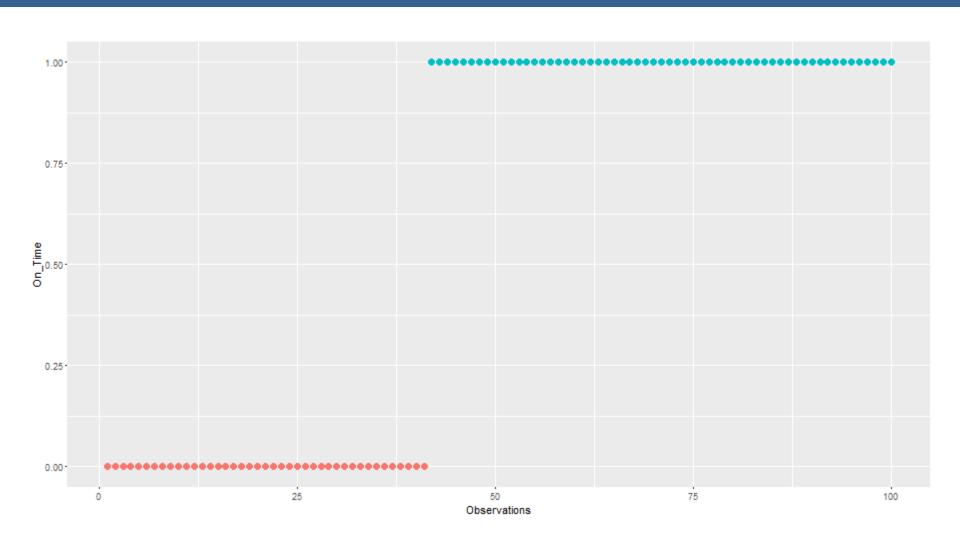
Null deviance: 135.372 on 99 degrees of freedom Residual deviance: 61.602 on 95 degrees of freedom

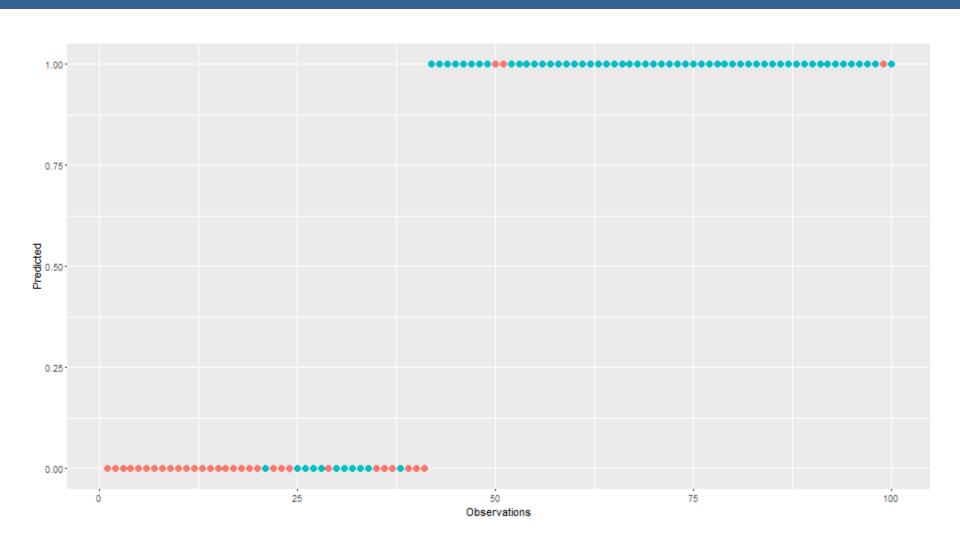
AIC: 71.602

```
# Prediction
prob <- predict(logit_02,type = c("response"))

ggplot(my_data, aes(x = 1:100, y = On_Time)) + geom_point(aes(colour = ifelse(my_data$On_Time >= 0.5, "red", "blue")), size = 3) + theme(legend.position = "none") + xlab ("Observations") + ylab("On_Time")

ggplot(my_data, aes(x = 1:100, y = On_Time)) + geom_point(aes(colour = ifelse(prob >= 0.5, "red", "blue")), size = 3) + theme(legend.position = "none") + xlab ("Observations") + ylab("Predicted")
```





The probability of Y_i is given by:

$$p(Y_i) = (p_i)^{Y_i} \times (1 - p_i)^{1 - Y_i}$$

For a sample with n cases, we can define the likelihood function as :

$$L = \prod_{i=1}^{n} [(p_i)^{Y_i} \times (1 - p_i)^{1 - Y_i}]$$

$$L = \prod_{i=1}^{n} \left[\left(\frac{e^{(Z_i)}}{1 + e^{(Z_i)}} \right)^{Y_i} \times \left(\frac{1}{1 + e^{(Z_i)}} \right)^{1 - Y_i} \right]$$

In practice it is more convenient to work with the maximum estimation of the log likelihood function:

$$LL = \sum_{i=1}^{n} \left\{ \left[(Y_i) \ln \left(\frac{e^{(Z_i)}}{1 + e^{(Z_i)}} \right) \right] + \left[(1 - Y_i) \ln \left(\frac{1}{1 + e^{(Z_i)}} \right) \right] \right\} = m \acute{a} x$$

```
# Betas logit_02$coefficients
```

(Intercept) -30.9333453 Distance 0.2041463

Traffic_Lights 2.9201140

Time_Period -3.7763010

Driver_Profile2 2.4590675

Maximum LogLikelyhood LLmax <- logLik(logit_02)

'log Lik.' -30.80079 (df=5)

The confidence interval is given by $B_i \pm z_{\alpha/2} \times SE_{B_i}$. Therefore for a 95% (z=1.96) confidence interval:

```
# Cls Using Standard Errors
confint.default(logit_02,level = 0.95)
                  2.5 %
                                   97.5 %
(Intercept)
             -51.776290960
                                -10.0903996
         0.005886105
Distance
                                  0.4024066
Traffic Lights 0.939388292
                            4.9008396
Time Period -5.435617700
                                 -2.1169843
Driver Profile2 0.225909956
                                  4.6922250
```

Goodness of Fit

We need to calculate the adequacy of the model. To measure the adequacy we use the null model (LL_0) and compared with our final model ($LL_{m\acute{a}x}$) through the likelihood ratio test. The null model has only the intercept (B_0). Therefore the following hypothesis can be tested:

 H_0 : The model is not adequate H_1 : The model is adequate

```
# Likelihood Ratio Test with(logit_02, pchisq(null.deviance - deviance, df.null - df.residual, lower.tail = FALSE))
```

[1] 3.6265e-15

There are several measures of association designed to mimic the r^2 analysis of, like the $pseudo\ r^2$. Values between 0.2 and 0.4 are considered highly satisfactory.

$$pseudo(r^{2})_{MacFadden} = \left(\frac{-2LL_{0} + 2LL_{max}}{-2LL_{0}}\right)$$

$$pseudo(r^{2})_{cox \& Snell} = 1 - \left(\frac{e^{LL_{0}}}{e^{LL_{max}}}\right)^{\frac{L}{N}}$$

$$pseudo(r^{2})_{Nagelkerke} = \frac{1 - \left(\frac{e^{LL_{0}}}{e^{LL_{max}}}\right)^{\frac{2}{N}}}{1 - (e^{LL_{0}})^{\frac{2}{N}}}$$

```
# Pseudo-r2
LL0 <- logLik(glm(On_Time ~ 1, data = my_data, family = "binomial"))
LLmax <- logLik(logit_02)

psd_r_McFadden = (-2*LL0 + 2*LLmax)/(-2*LL0)

[1] 'log Lik.' 0.544945 (df=1)

psd_r_Cox_and_Snell = 1 - (exp(LL0)/exp(LLmax))^(2/nrow(my_data))

[1] 'log Lik.' 0.5217881 (df=1)

psd_r_Nagelkerke = (1 - (exp(LL0)/exp(LLmax))^(2/nrow(my_data)))/(1 - (exp(LL0))^(2/nrow(my_data)))

[1] 'log Lik.' 0.7034824 (df=1)
```

Interpretation

$$Z_i = -30.933 + 0.204X_{1i} + 2.920X_{2i} - 3.776X_{3i} + 2.459X_{5i}$$

$$p_i = \left(\frac{e^{-30.933 + 0.204X_{1i} + 2.920X_{2i} - 3.776X_{3i} + 2.459X_{5i}}}{1 + e^{-30.933 + 0.204X_{1i} + 2.920X_{2i} - 3.776X_{3i} + 2.459X_{5i}}}\right)$$

 e^{B_i} ; $i \neq 0 \rightarrow$ Average change in the chance of arriving late (Y = 1) all other conditions remain constant.

Odds Ratios

(Intercept) Distance Traffic_Lights Time_Period Driver_Profile2
3.679754e-14 1.226478e+00 1.854340e+01 2.290727e-02 1.169390e+01

 $e^{B_1} = e^{0.204} = 1.226$: Chance to arrive late increases 22.6% if the distance increase in 1 km.

 $e^{B_2}=e^{2.920}=18.543$. Chance to arrive late increases 1754.3% if the quantity of traffic ligths increases in 1 unity.

 $e^{B_3}=e^{-3.776}=0.023$. Chance to arrive late decreases 97.7% in the morning period.

 $e^{B_5}=e^{2.459}=11.693$: Chance to arrive late increases 1069.3% if the driver has an aggressive profile

$$Z_i = -30.933 + 0.204 X_{1i} + 2.920 X_{2i} - 3.776 X_{3i} + 2.459 X_{5i}$$

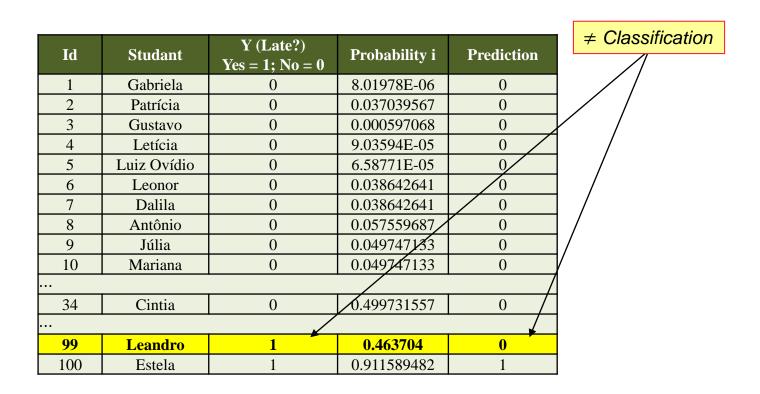
$$p_{i} = \left(\frac{e^{-30.933 + 0.204X_{1i} + 2.920X_{2i} - 3.776X_{3i} + 2.459X_{5i}}}{1 + e^{-30.933 + 0.204X_{1i} + 2.920X_{2i} - 3.776X_{3i} + 2.459X_{5i}}}\right)$$

$$p_{Gabriela} = \left(\frac{e^{-30.933 + 0.204(\mathbf{12.5}) + 2.920(\mathbf{7}) - 3.776(\mathbf{1}) + 2.459(\mathbf{0})}}{1 + e^{-30.933 + 0.204(\mathbf{12.5}) + 2.920(\mathbf{7}) - 3.776(\mathbf{1}) + 2.459(\mathbf{0})}}\right)$$

$$p_{Gabriela} = \left(\frac{e^{-13.819}}{1 + e^{-13.819}}\right) = \left(\frac{0.00000996}{1.00000996}\right) = 0.0009\%$$

Gabriela has a change of 0.009% to arrive late (Y = 1)

Cutoff de 0.5



Confusion Matrix:

OP = TP + FN;

TN	FN	PN
(True Negative)	(False Negative)	(Predicted Negative)
FP	TP	PP
(False Positive)	(True Positive)	(Predicted Positive)
ON	OP	TNI ED EN TD
(Observed Negative)	(Observed Positive)	n = TN + FP + FN + TP

```
TP = Case: 1 & Prediction: 1;

FN = Case: 0 & Prediction: 1; Type II Error

FP = Case: 1 & Prediction: 0; Type I Error

PN = TN + FN;

PP = TP + FP;

ON = TN + FP;
```

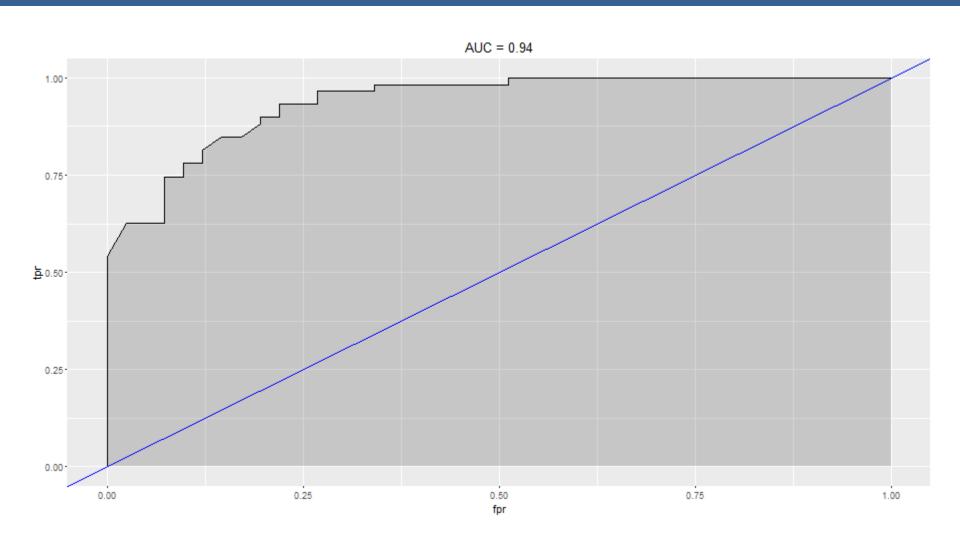
TN = Case: 0 & Prediction: 0;

- **TPR** (True Positive Rate) = $\frac{TP}{OP}$
- TNR (True Negative Rate) = $\frac{TN}{ON}$
- ACC (Accuracy) = $\frac{TP+TN}{n}$;
- **FPR** (False Positive Rate) = 1 TNR ou $\frac{FP}{ON}$
- **PPV** (Positive Prediticed Value = Sensitivity) = $\frac{TP}{PP}$;
- **NPV** (Negative Prediticed Value = Specificity) = $\frac{TN}{PN}$;

TN = 30	FN = 3	PN = 41
FP = 11	TP = 56	PP = 59
ON = 44	OP = 59	n = 100

The values of *Sensitivity* and 1 - *Specificity* are used to plot a ROC curve (Receiver Operating Characteristic). The more distant the ROC curve is in relation to a reference curve, the better. A curve very close to the reference shows that the model's ability to discriminate between the occurrence and non-occurrence is due to chance.

```
# ROC Curve
library("ggplot2")
library("ROCR")
pred <- prediction(prob, my_data$On_Time)
perf <- performance(pred, measure = "tpr", x.measure = "fpr")
auc <- performance(pred, measure = "auc")
auc <- auc@y.values[[1]]
roc.data <- data.frame(fpr = unlist(perf@x.values), tpr = unlist(perf@y.values), model = "GLM")
ggplot(roc.data, aes(x = fpr, ymin = 0, ymax = tpr)) + geom_ribbon(alpha = 0.2) + geom_line(aes(y = tpr)) +
geom_abline(colour = "blue", intercept = 0, slope = 1) + ggtitle(paste0("AUC = ", round(auc, digits = 2)))</pre>
```



Prediction

```
# Prediction
new_data <- as.data.frame(cbind(9.5, 8, 1, 0))
colnames(new_data) <- c("Distance", "Traffic_Lights", "Time_Period", "Driver_Profile2")
predict(logit_02, new_data, type = c("response"))
```

MVDA

https://github.com/Valdecy/Multivariate_Data_Analysis

```
# Created by: Prof. Valdecy Pereira. D.Sc.
```

- # UFF Universidade Federal Fluminense (Brazil)
- # email: valdecypereira@yahoo.com.br
- # Course: Multivariate Data Analysis
- # Lesson: Binary Logistic Regression (Logit)

Citation:

PEREIRA. V. (2016). Project: Multivariate Data Analysis. File: R-MVDA-08-LR-B.pdf. GitHub

repository: https://github.com/Valdecy/Multivariate_Data_Analysis

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