Hirability and Educational Prestige

John Vandivier^a

^a4400 University Dr, Fairfax, VA 22030

Abstract

Alternative credentials offer a partial solution to the skill gap and student debt

crises, supernormal returns for some students, and a tool to support diver-

sity hiring for firms. This paper tests the hypothesis that educational prestige

explains hirability better than accreditation. An original questionnaire using

repeated measures (n = 454) is investigated to determine the effects of ac-

creditation and prestige on willingness to hire. A combination of descriptive

statistics, ordinary least squares (OLS), and linear mixed models (LMM) reveal

that prestige explains a larger component of hirability variance than accredita-

tion. Accreditation remains independently important. Alternative credentials

from a Fortune 50 learning provider are competitive with a traditional degree.

Other high prestige alternative credentials demonstrate job search value, albeit

at a lower level. The conclusion includes a discussion on industry and policy

moves to improve total market surplus.

Keywords: alternative education, job search, work norms, education

economics, social economics, labor economics, skill gap, prestige, debt crisis

2010 MSC: I20, I24, J24, B55

Email address: jvandivi@masonlive.gmu.edu (John Vandivier)

1. Introduction

The accredited degree is an established means to individual-level employability, but the proliferation of the degree is associated with a variety of well-understood issues. These issues include the student debt crisis, skill gaps, grade inflation, and low social return. Alternative credentials, or non-accredited credentials, are a broad category of offerings that exhibit greater variation intensity, price, and outcomes[1]. Alternative credentials are often a signal of niche skills and expertise in a particular job family. These characteristics combine to provide the benefit of high possible value addition to the labor market with the cost of a value calculation problem shared by potential employers and education consumers.

This paper seeks to reduce the general difficulty of credential value calculation by testing a method of value normalization with heuristics to identify those credentials likely to yield meaningful benefits to the typical job search. This paper tests the lens of prestige as a tool to normalize value across accredited and alternative credentials. This study leverages an original questionnaire to identify prestige levels of various credentials. This paper tests the composite hypothesis that some level of prestige allows an alternative credential to compete with traditional credentials for employment.

Several specific lines of evidence are required to support the composite hypothesis. Statistical evidence must demonstrate significant positive effects for accreditation and prestige on hirability. The effect size for prestige must be sufficiently large to dominate the accreditation effect over the attainable range. The questionnaire allows a prestige response on a 10-point scale, so the attainable range is from 1 to 10. A vignette analysis can test whether a dominant range for prestige exists within the attainable window. An ideal result would further show that one or more actual alternative credentials fall into this dominant range.

The motivation for the lens of prestige extends from the academic work in education economics and the economics of social norms. Education economics provides two mainstream accounts of the value of a degree. One account is the human capital model, and the other is the signaling model. The human capital model explains that improved labor outcomes result from skills gained by a student in the course of education.

Stakeholders of various kinds prefer alternative credentials to the traditional degree for the attainment of specific technical skills[2]. For this reason, many college graduates supplement using alternative credentials. Some alternative learning providers specifically target this market with a special kind of alternative education called last-mile training. This presents an explanatory problem for the human capital model. If better labor outcomes arise from skill enhancement, then alternatively educated individuals should enjoy better wages, employment rates, and so on, compared to college graduates.

The signaling model holds that credentials signal a basket of applicant qualities that employers value. Proponents of the signaling model commonly argue that the college degree signals intelligence, work ethic, and conformity[3]. The signaling model presents an explanation for the correlation of weak labor outcomes and alternative credentials, even if alternative credentials endow students with better skills. The explanation is that the alternative credential signals an offsetting deficit of some kind. This paper treats prestige as a signal rather than a matter of human capital. This paper prefers the signaling approach to directly investigate prestige effects with minimal theoretical baggage and without a need to test student skill.

In a broad review of economics and norm types, hiring decisions exist within what Elster would identify as work norms[4]. Elster supports a rational model of work norms, with the caveat that social interactions may involve unobserved emotional effects. Similarly, the rational model used in this paper may not extrapolate with accuracy into abnormal emotional situations. This paper will also make use of the distinction between social and legal norms provided by Elster.

Rivera is one scholar within the economics of work norms to have recently operationalized social norms as prestige[5]. Rivera finds that prestige is important in her analysis, but her analytical scope focuses on traditional education

and a few specific industries, including health and law. The current paper extends the analysis of prestige and hiring norms across many industries and to include alternative credentials.

2. Description of Data and Methodology

This paper investigates an original set of online questionnaire responses (n = 454). Responses are cross-sectional data obtained in March of 2021. Respondents are United States citizens at or over the age of eighteen. Qualified respondents participated in the survey through the Amazon Mechanical Turk platform.

Appendix A contains the wording and response options for each question. Appendix A also contains the wording for a priming message presented at the start of the survey. The priming message lays out the definition of alternative credentials used in this study. The message also provides several concrete examples of alternative credentials, including "a Certified Project Manager certification, a portfolio of work, a Khan Academy profile, or a Nanodegree from Udacity."

The dependent variable of interest is called hirability. This variable measures individual response on a 10-point scale to the question, "For many professions, alternative credentials can qualify a person for an entry-level position." The questionnaire is composed of three sections. The first section collects respondent characteristics and baseline hirability. The second section collects prestige responses with respect to nine real-world learning providers. The third section collects hirability and prestige responses with respect to eight vignette learning providers.

Investigation of the first section of the questionnaire uses ordinary least squares analysis. Vignette data is analyzed as a panel in mixed models with individual random effects. The vignette model allows comparison between prestige and accreditation coefficients. Vignette analysis encounters a practical util-

ity problem in that the schools are only vignettes rather than actual learning

providers. A comparison of descriptive statistics across vignettes and actual schools addresses this concern.

Half of the respondents randomly received an informational message about the nine real-world learning providers. Appendix A includes the wording of this message. The message provides rating data from two leading credential aggregator websites. University ratings are US News ranking information for the 2021 school year. Course Report provides the rating data for so-called coding bootcamps as of December 2020.

As an aside, inaccurate credential category labels contribute to the knowledge and value calculation problems that inhibit social adoption. Coding bootcamps focus on roles in the information technology industry, but these roles are much broader in scope than the category label implies. Moreover, the information technology industry is a special industry that cuts across all other industries. Much of the academic, policy, and industry discussion on coding bootcamps misses that these institutions provide credentials that potentially compete with university degrees in nearly any subject.

For example, General Assembly is one of the particular coding bootcamps investigated in this study. General Assembly provides credentials for user experience design, a set of skills involving market research, and applied technical art skills. General Assembly provides credentials for product management. Product management is a job family that competes for labor among business degree graduates. The data science credential provides skills that compete with accredited labor in mathematics, statistics, economics, and even subjects in the hard sciences like computational biology. Finally, there are credentials that relate to software development and compete with accredited degrees in computer science.

Respondent characteristics are categorical variables. Hirability and prestige are 10-point Likert-type responses. Prestige takes a second representation as a stipulated boolean. Stipulating prestige enables the application of results to a real job search. If stipulated prestige is highly correlated to prestige response, and if prestige response is correlated to improved hirability, then the selection criteria for stipulated prestige can be applied in an actual job search to poten-

tially improve outcomes.

130

145

To illustrate the method of two-way prestige validation, suppose that a vignette school is stipulated as high prestige. This situation is represented in regression as a dummy variable for stipulated high prestige with a value of true. The respondent reads that the vignette school is known to be prestigious. After reading this, the respondent provides a prestige response rating on a 10-point scale. Investigation of all responses allows an analyst to determine an average prestige response level which is associated with the stipulated high prestige criteria.

To preview results, stipulated high prestige turns out to be strongly correlated with high prestige response. Interestingly, there are cases where a respondent gives a low response rating to, for example, the University of Chicago, a school with high stipulated prestige based on aggregator website ratings. This result indicates the importance of some analysis that accounts for individual effects.

Two-way representation of prestige enables the application of findings into an actual job search. In an actual job search, individuals can easily access aggregator website data. In the real world, an individual cannot readily access questionnaire results for many credentials. Results from this paper include the identification of rules of thumb that a person can use to identify actual learning providers as high prestige. To ensure clarity of results, stipulated prestige always refers to the dummy variable, and prestige response refers to the 10-point measure.

The vignette section and the section on actual schools use stipulated prestige. All other variables are either 10-point Likert-type responses or categorical variables¹. Categorical variables are exclusively respondent characteristics. Four other respondent measures are Likert-type responses. Vignette responses include responses for hirability and prestige, while actual schools only receive responses for hirability.

Respondent characteristics include eight standard controls and four questions unique to this study. The eight standard controls include age, gender, ethnicity, income, level of education, employment status, the industry of occupation, and state of residence. A unique question on work norms records whether the respondent tends "to work more closely with coworkers at your company or customers and external business partners." The motivation for this question is to test whether prestige disproportionately impacts roles that are outward or client-facing. Respondents are also directly asked whether they "prefer to hire or work with a person that has a college degree rather a person that holds a reputable certification or non-college credential."

Another unique control is support for online education. This control allows analysis to separate hirability effects due to online education preference from hirability effects due to unaccredited education preference. In practice, many alternative credentials involve online learning, but accredited learning is also increasingly taking place online.

The fourth control is expected conventionality. This variable measures whether the respondent believes that "It will soon become common for high school graduates to obtain alternative credentials instead of going to college." This is a useful correction variable for two reasons. First, it separates willingness to hire based on respondent preference from indirect willingness to hire based on perceived social norms. Individual preferences and social norms are certainly correlated,

¹It is an accepted practice to treat Likert-type responses as either categorical or continuous for regression analysis. Jaccard and Wan provide support for continuous analysis of Likert-type data. They note that severe departures from the assumptions on cardinality "do not seem to affect Type I and Type II errors dramatically," particularly when the Likert scale is five or more points[6]. This paper treats responses on a 10-point scale as continuous.

but the correlation is small enough that failure to separate these effects leads to nontrivial statistical noise.

Second, surveys sometimes overreport demand effects because of the lack of cost constraint on respondent expression. This bias is sometimes called budget constraint bias or omitted budget constraint bias[7, 8]. Without a cost constraint, respondents tend to exaggerate demand responses like the willingness to hire. Budget constraint bias affects both hirability and expected conventionality, so conventionality operates in part as a bias control.

Vignette question formatting follows Atzmüller and Steiner[9]. Each vignette stipulates whether a school is accredited, whether the respondent should imagine the school as impressive, and whether the respondent should imagine that other people consider the school impressive. Each stipulated factor can take a value of true or false, resulting in eight vignette questions.

This study uses multiple regression and descriptive statistics to generate results. Multiple regression is conducted using ordinary least squares (OLS) for baseline hirability analysis and linear mixed models (LMM) are used for vignette analysis. OLS specification of vignette data is inappropriate because repeated measures of hirability from a single participant introduce an individual-level bias into resulting coefficients. LMM models are able to account for these individual-level effects. Following Magezi[10], linear mixed models in this paper use a within-participant random factor, or individual random effects, to correct for individual-level repeated measures bias. LMM yields linear coefficients, so the interpretation of LMM coefficients is similar to OLS. One difference of note is that adjusted r-squared is not available for an LMM model.

3. Results

180

185

Results (n = 454) indicate that accredited degrees are generally higher in prestige compared to alternative credentials. Alternative credentials are meaningfully associated with hirability, and in certain situations, they are preferred to accredited degrees.

Competitive status indicates that a credential is correlated with hirability to a similar or greater extent compared to an accredited degree. Results provide evidence for three cases in which alternative credentials are competitive. First, specific alternative credentials are of particularly high prestige. This study finds that a credential from Google is sufficiently prestigious to be competitive without a requirement of supplementary conditions.

Second, some individuals award prestige preferentially to alternative learning providers. In a comparison among nine actual learning providers in this study, 71 percent of respondents prefer at least one alternative credential to at least one university degree. The proportion increases to about 75 percent when respondents view rating data from the online review aggregators Course Report and US News.

Third, certain independent factors in hiring decision models support the hirability of alternatively credentialed job candidates. Industry and state effects are two such compensating factors that can add up to overcome the average comparative preference for accredited labor to alternatively credentialed job candidates.

Baseline hirability is the institution-agnostic hirability measure. The mean response for baseline hirability is 7.58 on a 10-point scale, and the median response is 8. Table 1 gives average hirability and prestige for interesting segments of respondents. Four basic results in the table are worth noting. First, stipulated prestige always moves with prestige response as expected. Second, as expected, the hirability and prestige effects for accredited schools are generally higher than non-accredited schools.

Third, the difference in average hirability between high and low prestige providers is more than twice the difference in hirability between accredited and unaccredited providers. This supports the possibility of an actual competitive alternative credential in the attainable range of prestige. The fourth result is an initial attempt at a prestige rule of thumb. For both vignette and actual schools, if a school can obtain a prestige score of 7 or more, it will be at least as prestigious as the average accredited school.

Table 1: Average Hirability and Prestige

| | | Average Hirability | Average Prestige |
|------------------|--------------------------|--------------------|------------------|
| Actual Schools | | | 6.50 |
| | Accredited | | 7.05 |
| | Unaccredited | | 6.07 |
| | Difference | | 0.98 |
| | Stipulated High Prestige | | 6.72 |
| | Stipulated Low Prestige | | 6.23 |
| | Difference | | 0.49 |
| Vignette Schools | | 6.49 | 6.21 |
| | Accredited | 6.97 | 6.49 |
| | Unaccredited | 6.02 | 5.93 |
| | Difference | 0.95 | 0.56 |
| | Stipulated High Prestige | 7.59 | 7.69 |
| | Stipulated Low Prestige | 5.63 | 4.94 |
| | Difference | 1.96 | 2.75 |

Google is the only unaccredited learning provider to achieve a competitive status on the basis of this initial rule. The mean prestige response for Google was 7.10, and the median response was 7. Two lower bars for competitive status are interesting. First, an alternative provider can be described as moderately competitive if it fails to beat the average university, but it succeeds in beating at least one university on average. The lowest average prestige response for an accredited university is 6.34 for the University of Nebraska.

Second, an alternative provider can be described as weakly competitive if it fails to beat any university on average, but it succeeds in beating at least one university in a significant percentage of individual responses. No alternative credentials investigated in this study meet the criteria for moderate competitiveness. App Academy, General Assembly, and Google are the three alternative learning providers with stipulated high prestige. All stipulated high prestige learning providers are at least weakly competitive.

When asked directly, 41.6 percent of respondents indicated that they would not prefer to work with a person that holds an accredited credential instead of "a person that holds a reputable certification or non-college credential." When examining prestige response instead of asking directly, over 70 percent of respondents reveal a preference for at least one actual alternative credential to at least one university credential. Over half of respondents preferred at least one actual alternative credential with stipulated high prestige to at least one university credential with stipulated high prestige. After excluding Google, over one-quarter of respondents continue to prefer at least one actual alternative credential with stipulated high prestige to at least one university credential with stipulated high prestige.

Zety is an online platform that facilitates job search. Zety reports that one in six job applicants in the United States receive an interview, and the average conversion rate from interview to offer was 19.78 in 2016[11]. Assuming rejections are independent enables naive estimation that most job searches consist of at least four interviews² and dozens of applications. Given the rates at which respondents prefer alternative credentials to accredited degrees, a job search of typical length is likely to include several applications and at least one interview with one or more employers that would prefer an alternative credential with stipulated high prestige to an accredited degree.

Table 2: Table of Regression Results with Hirability as Dependent Variable

| | Model 1 Model 2 Model 3 | | | | |
|------------------------|-------------------------|--|--|--|--|
| Age, 45-60 | 0.61*** 0.10 | | | | |
| External Facing, High | 1.23*** 0.13 | | | | |
| External Facing, Low | 1.16*** 0.10 | | | | |
| Continued on Next Page | | | | | |

²Four independent games that each include an eighty percent chance of rejection yields $0.8^4 = 0.4096$. The associated probability of having at least one offer result from four interviews would be about 1 - 0.41 = 0.59, or 59 percent, which is more likely than not.

Table 2 – Continued

| | Model 1 | Model 2 | Model 3 | |
|------------------------------------|---------|----------|----------|--|
| External Facing, Medium | 1.16*** | 0.13 | | |
| Expected Conventionality | 0.32*** | 0.14*** | 0.17*** | |
| Income, 0-9999 | 0.88 | -0.87** | -1.22*** | |
| Income, 100,000-124,999 | 1.25*** | 0.47** | 0.41* | |
| Income, 175,000-199,999 | 1.58* | 0.40 | | |
| Income, $200,000+$ | 1.14 | -1.09* | | |
| Income, 25,000-49,999 | 0.57** | 0.19 | | |
| Income, 50000-74999 | 0.51** | 0.26* | 0.18 | |
| Income, 75000-99999 | 0.81*** | 0.29* | | |
| Industry, Education | 0.66** | 0.40** | | |
| Industry, Finance | 0.34 | -0.07 | | |
| Industry, Information Technology | 0.46** | 0.05 | | |
| Industry, Manufacturing | 0.34 | 0.17 | | |
| Industry, Other | 0.37 | 0.37** | | |
| Is Accredited | | 1.23*** | 1.27*** | |
| (Is Accredited)(Prestige Response) | | -0.09*** | -0.10*** | |
| Is Stipulated High Prestige | | | 0.14** | |
| Is Stipulated Other Impressed | | 0.64*** | 0.59*** | |
| Is Stipulated Self Impressed | | -0.05 | | |
| Online Ed Favorability | 0.34*** | 0.09*** | 0.07** | |
| Prefers Traditional Coworker | -0.22 | 0.19* | 0.19* | |
| Prestige Response | | 0.55*** | 0.53*** | |
| State, Arizona | 1.35** | 0.69** | | |
| State, California | 0.44** | 0.27** | 0.37** | |
| State, Connecticut | 0.72 | -0.11 | | |
| State, Florida | 0.79*** | 0.16 | | |
| Continued on Next Page | | | | |

Table 2 - Continued

| | Model 1 | Model 2 | Model 3 |
|--|---------|---------|---------|
| State, Georgia | -0.88* | -0.22 | |
| State, Kansas | 1.76 | 0.52 | |
| State, Maryland | 0.92** | 0.31 | |
| State, Massachusetts | 1.43** | 0.49 | |
| State, Michigan | 1.35*** | 0.26 | |
| State, Mississippi | 1.77*** | 0.45 | |
| State, Missouri | 0.81* | 0.34 | |
| State, Nebraska | -1.04 | -0.75 | |
| State, New Mexico | 1.76* | 0.10 | |
| State, Pennsylvania | 0.44 | 0.44** | |
| State, Tennessee | 0.74 | -0.13 | |
| State, Texas | 0.39 | -0.10 | |
| State, West Virginia | -1.31 | -0.92 | |
| Intercept | 0.30 | 0.14 | 0.50* |
| R-squared | 0.47 | | |
| R-squared Adj. | 0.42 | | |
| N | 454 | 3600 | 3600 |
| Measures Per Respondent | 1 | 8 | 8 |
| * $p < 0.10, *** p < 0.05, **** p < .01$ | | | |
| | | | |

Table 2 gives three models. The first model is an ordinary least squares model of baseline hirability. Backward elimination to the point of adjusted r-squared maximization yields Model 1. Adding factors of accreditation and prestige to Model 1, then adapting the model to a linear mixed model (LMM) yields Model 2. Model 3 results from additional backward elimination on Model 2.

Four individuals that completed the first section of the questionnaire did not complete the entire questionnaire. The remaining 450 respondents each report hirability for the eight vignette schools, yielding 3,600 observations for the mixed models.

Because LMM does not permit computation of r-squared, the termination criteria for the factor elimination process in Model 3 was to retain all factors with a p-value under 0.5. This is a permissive criterion intended to guard against overfitting. The logical basis for this rule is that each observed effect is more likely to exist than to not exist when p < 0.5. Despite permissive criteria, only one insignificant factor for income exists in Model 3.

Model 2 and Model 3 have one other interesting difference. Model 3 includes the boolean for whether a school was stipulated as high prestige. For vignette schools with high prestige, the participant viewed two statements about the vignette. The questionnaire instructs the participant to imagine a school they consider to be impressive. The questionnaire also instructs the participant to imagine that other people consider the school to be impressive. This situation is technically equivalent to an interaction of the two subcomponents. Because Model 2 includes both stipulated high prestige subcomponents and the accreditation dummy, including high prestige generates perfect multicollinearity. Backward elimination of Model 2 drops the factor for own stipulated prestige, so subsequent insertion of high prestige is nonproblematic.

Model 3 is the preferred model. Prestige and accreditation effects are positive and significant. These two effects also interact with a significant and negative coefficient. The values of these coefficients of interest are consistent across Model 2 and Model 3. The dummy variable for accreditation is about two and a half times larger than the prestige response, but the average prestige response is near seven. This indicates that the prestige response explains a larger share of hirability variance compared to accreditation.

An application of Model 3 is another approach to the identification of competitive alternative credentials. Hold factors other than accreditation and prestige constant. Let the hirability level of school k be called H_k . Let X_{ka} be

accreditation status, X_{kp} is prestige response, X_{kh} is the dummy for stipulated high prestige, and X_{ko} is the dummy for whether other people consider the school prestigious.

Let H_1 be an unaccredited school with high stipulated prestige. Let H_2 be an accredited school without high stipulated prestige. Let $X_{2p} = 6.49$, which is the prestige response equal to the average for an accredited vignette, as reported in Table 1. This system of equations is described in equations 1a through 1e:

$$H_k = 1.27X_{ka} - 0.1X_{ka}X_{kp} + 0.53X_{kp} + 0.14X_{kh} + 0.59X_{ko}$$
 (1a)

$$H_1 = 0.53X_{kp} + 0.14 + 0.59 \tag{1b}$$

$$H_2 = 1.27 - 0.1(6.49) + 0.53(6.49)$$
 (1c)

$$X_{kp} = (1.27 - 0.1(6.49) + 0.53(6.49) - 0.14 - 0.59)/0.53$$
 (1d)

$$X_{kp} \approx 6.28$$
 (1e)

Equation 1e indicates that an alternative credential with stipulated high prestige and a prestige response of 6.28 or higher is approximately competitive with the average accredited vignette. Table 1 indicates that the prestige response for the average vignette school is 6.21. This is a significant difference compared to the average actual school prestige response of 6.50. Coincidentally, additive and proportional compensation of 6.28 both yield 6.57.

315

This prestige requirement exceeds the low bar set by comparison to the University of Nebraska. Google remains the only alternative provider to obtain general competitive status without the presence of other preferential factors. App Academy and General Assembly both have average prestige responses close to 5.8. Models reveal several situations in which other factors overcome this deficit, but many of these offsetting factors are difficult to determine and leverage prior to a hiring decision. The California state effect is an interesting exception that an actual job search could exploit.

Alternative credentials provide a source of potential diverse labor to employers. Interestingly, neither ethnicity nor gender was significantly associated with hirability. There is little evidence for the thesis that client-facing roles preferentially benefit from credential prestige or accreditation. Respondent client exposure on the job was associated with a slightly larger baseline willingness to hire an alternatively educated candidate. The extent of client contact was insignificant in mixed models.

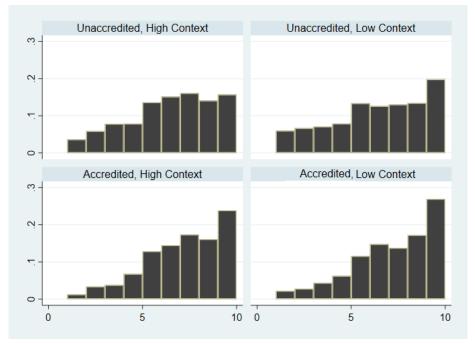


Figure 1: Prestige Response Distribution for Actual Schools

Finally, Figure 1 visualizes the prestige response distribution for actual schools. The four subplots describe whether a respondent randomly received information from review site aggregators and how they evaluated credential accreditation. Exposure to aggregated review information is associated with fewer responses at the positive and negative extrema of the response distribution for accredited and unaccredited schools. On average, alternative education prestige rose, and accredited education prestige declined when a respondent received review aggregator site information.

4. Conclusions

345

370

This study hypothesized that some level of prestige allows an alternative credential to compete with traditional credentials for employment. Results provide evidence in favor of this hypothesis. Regression results show meaningful positive correlations of prestige and accreditation on hirability. A range of hirability responses that include the average response and some below-average responses find a dominant explanation in prestige effects over accreditation alone.

While prestige explains a larger share of hirability variance than accreditation, accreditation robustly maintains a meaningful effect on its own. The robust importance of accreditation indicates that arbitrary improvements to alternative credential quality and social acceptability are not likely to displace the higher education system in expectation. This study began with the assertion that alternative credentials are a source of unexploited technical value. The study validated a partial explanation from prestige as a representation of social norms. The introduction noted an important distinction between legal and social norms from Elster. By elimination, legal norm change is an important candidate to allow alternative credentials the opportunity to fully outcompete the hirability effects of accreditation.

In 2012, The Heritage Foundation called for two policy changes that are worth considering. First, the Foundation proposed that the government should directly accredit courses rather than organizations[12]. Second, they also called for a decoupling of accreditation and federal funding. An additional option would be to replace legal requirements for formal education could be replaced with skill assessments. With a legal requirement that prefers skills to degrees, the public sector gains the ability to transfer formal accreditation duties to a market model with no loss of labor quality control.

There are several reasons to be pessimistic about the feasibility of these policy changes. Reductions to education spending are unpopular with voters in the United States. Over ninety percent of K-12 students in the United States attended a public school in 2016[13], and there is a systematized pipeline from

public school to the traditional university system. Education represents an example of an entangled political economy[14]. Robust political economy points out additional reasons to doubt rapid innovation in this space[15]. Reduced political entanglement is associated with the absence of compulsory education. However, after they exist, the elimination of compulsory laws also appears intractable. The removal of compulsory education is a qualitative change that does not appear any less subject to the path dependency, lock-in, ratchet, and other effects that inhibit contraction in the quantitative process of appropriations.

An interesting alternative to formal legislative change is the emerging model of public-private partnerships in education. In 2013, Georgia Tech formally partnered with Udacity to produce an accredited online graduate degree in Computer Science[16]. Udacity was able to facilitate an improved online learning experience at scale with an affordable price. Georgia Tech offered branding, legitimacy, and accreditation, which supported a higher price point compared to the other offerings from Udacity.

In other cases, the hybridization of traditional and alternative education is indirect and informal. Prior learning assessments and portfolio reviews are two of many processes by which a university can award credit to a student without formal requirements connected to the source of student learning[17]. University support for prior learning is an implementation pattern for course-level accreditation that does not require legislative action. Formal and informal partnerships between traditional and alternative institutions can yield increased market surplus for producers and consumers.

Finally, this paper evaluated practical alternative credential selection strategies. One strategy is to leverage credentials from industry leaders. In this study, Google represented an alternative learning provider that is also an industry leader. Fortune 50 membership is a rule of thumb used in this study to select an industry-leading firm. A credential from Google was the only alternative credential to be identified as generally competitive with an accredited degree.

The second strategy is to use credential review aggregator sites to identify high prestige credentials. This paper used Course Report as an aggregator to search for alternative credentials. App Academy and General Assembly were identified by applying search criteria that include a rating of 4.25 or better on a 5-point scale and a minimum of four hundred reviews. The combination of results with information on typical job search length from Zety indicated that these credentials provide meaningful job search benefits, albeit with significantly less efficacy than an accredited degree or a credential from Google.

References

- [1] T. Urdan, Beyond the noise: The rise of alternative credentials (Dec 2020).

 URL https://tytonpartners.com/library/
 beyond-the-noise-the-rise-of-alternative-credentials/
 - [2] R. Craig, A new U: Faster+ cheaper alternatives to college, BenBella Books, 2018.
- [3] B. Caplan, Signaling versus educational innovation (Apr 2012).
 URL https://www.econlib.org/archives/2012/04/signaling_versu.
 html
 - [4] J. Elster, Social norms and economic theory, Journal of economic perspectives 3 (4) (1989) 99–117.
 - [5] L. A. Rivera, Pedigree: How elite students get elite jobs, Princeton University Press, 2016.
 - [6] J. Jaccard, C. K. Wan, J. Jaccard, LISREL approaches to interaction effects in multiple regression, no. 114, sage, 1996.
 - [7] M. Ahlheim, Contingent valuation and the budget constraint, Ecological Economics 27 (2) (1998) 205–211.
- [8] M. J. Pachali, P. Kurz, T. Otter, Omitted budget constraint bias and implications for competitive pricing, Available at SSRN 3044553 (2020).

- [9] C. Atzmüller, P. M. Steiner, Experimental vignette studies in survey research, Methodology (2010).
- [10] D. A. Magezi, Linear mixed-effects models for within-participant psychology experiments: an introductory tutorial and free, graphical user interface (lmmgui), Frontiers in psychology 6 (2015) 2.
 - [11] B. Turczynski, 2021 hr statistics: Job search, hiring, recruiting & interviews (Feb 2021).
 URL https://zety.com/blog/hr-statistics
 - [12] L. M. Burke, S. M. Butler, Accreditation: Removing the barrier to higher
- education reform. backgrounder. no. 2728. executive summary., Heritage Foundation (2012).
 - [13] U. D. of Education, Digest of education statistics 2018, table 206.40, NCES 2020-009 (2019).
- [14] R. E. Wagner, Entangled political economy: A keynote address, in: Entangled political economy, Emerald Group Publishing Limited, 2014.
 - [15] P. J. Boettke, P. T. Leeson, Liberalism, socialism, and robust political economy, Journal of Markets and Morality 7 (1) (2004) 99–111.
- [16] R. Empson, Georgia tech teams up with udacity, at&t to offer \$6k master's
 degree in computer science, entirely online (May 2013).

URL https://techcrunch.com/2013/05/15/

[17] D. Conrad, Building knowledge through portfolio learning in prior learning assessment and recognition, Quarterly Review of Distance Education 9 (2) (2008) 139.