

The objective of estimation:

\* Find values of parameters of model/distribution using data. These values of parameter are called parameter estimates.

\* Estimation Techniques

- Method of moments
- maximum likelihood Estimation
- Bayesian Estimation

\* Maximum Likelihood Estimation

consider a random variable  $X \sim f(x; \theta)$ , Probability distribution with parameters  $\theta \in \mathbb{R}^m$ ;  $\theta = \begin{bmatrix} \theta_1 \\ \theta_2 \\ \vdots \\ \theta_m \end{bmatrix}$ .

Simplicity, let us assume  $m=1$ .

\* Data:  $x_1, x_2, \dots, x_N$  (Random samples)

\* Objective: Estimate a value of parameter,  $\theta$ .

\* Likelihood function:

$$\begin{aligned} L(\theta) &= L(x_1, \dots, x_N | \theta) = L(\theta | x_1, \dots, x_N) \\ &= f(x_1; \theta) \cdot f(x_2; \theta) \cdot \dots \cdot f(x_N; \theta) \\ &= \prod_{i=1}^N f(x_i; \theta) \end{aligned}$$

Data
Unknown parameter

\* Maximum Likelihood Estimation: finds a value of  $\theta$  that maximizes  $L(\theta)$

\* Example: Data:  $x_1, \dots, x_N$ , Exponential distribution;  $\text{Exp}(x; \lambda)$   
Estimate  $\lambda$  using data.

$$\begin{aligned} f(x; \lambda) &= \lambda e^{-\lambda x}, \quad x \geq 0 \\ &= 0, \quad x < 0 \end{aligned}$$

\* Likelihood function

$$\begin{aligned} L(\lambda) &= L(\lambda | \theta) = \prod_{i=1}^N f(x_i, \lambda) = \prod_{i=1}^N \lambda e^{-\lambda x_i} \\ &= \lambda^N e^{-\lambda \sum_{i=1}^N x_i} \end{aligned}$$

\*  $\ln(L(\lambda))$ : Log likelihood function  
 $= n \ln \lambda - \lambda \sum_{i=1}^N x_i$

$$\frac{\partial \ln(L(\lambda))}{\partial \lambda} = 0 \Leftrightarrow \frac{n}{\lambda} - \sum_{i=1}^N x_i \Leftrightarrow \lambda = \frac{1}{\bar{x}}$$

$$\frac{\partial^2 \ln L(\lambda)}{\partial \lambda^2} = -\frac{n}{\lambda^2} < 0$$

\* **Example 2**: Data:  $x_1, \dots, x_n$

Distribution: Normal with  $\mu, \sigma^2$  parameters

Unknown Parameters:  $\mu, \sigma^2$

\* Likelihood function:

$$\begin{aligned} L(\mu, \sigma^2) &= \prod_{i=1}^n f(x_i, \mu, \sigma^2) \\ &= \prod_{i=1}^n \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x_i - \mu)^2}{2\sigma^2}} \\ &= \frac{1}{(2\pi\sigma^2)^{n/2}} e^{-\frac{1}{2\sigma^2} \sum_{i=1}^n (x_i - \mu)^2} \end{aligned}$$

$$* \ln(L(\cdot)) = -\frac{n}{2} \ln(2\pi\sigma^2) - \frac{1}{2\sigma^2} \sum_{i=1}^n (x_i - \mu)^2$$

$$\begin{aligned} * \frac{\partial \ln(L(\cdot))}{\partial \mu} &= 0 \Leftrightarrow +\frac{1}{2\sigma^2} \sum_{i=1}^n (x_i - \mu) = 0 \\ &\Leftrightarrow \hat{\mu} = \frac{\sum_{i=1}^n x_i}{n} = \bar{x} \end{aligned}$$

$$\begin{aligned} * \frac{\partial \ln(L(\cdot))}{\partial \sigma^2} &= 0 \Leftrightarrow -\frac{n}{2\sigma^2} + \frac{1}{2\sigma^4} \sum_{i=1}^n (x_i - \mu)^2 = 0 \\ &\Leftrightarrow +n\hat{\sigma}^2 = \sum_{i=1}^n (x_i - \mu)^2 \\ &\Leftrightarrow \hat{\sigma}^2 = \frac{1}{n} \sum_{i=1}^n (x_i - \mu)^2 \end{aligned}$$

\* MLE for mean is the sample mean and variance is a biased version of the sample variance.  $= \frac{1}{n-1} \sum_{i=1}^n (x_i - \mu)^2$

\* Properties of MLE

→ How good are the estimates?

Minimum Variance Unbiased estimator (MVUE)

\* MLE with large enough samples

$n \rightarrow \infty$

i)  $\hat{\theta}$  (estimate of parameters) is approximately unbiased estimator  
 $\square(\hat{\theta}) \sim N \rightarrow T \dots \text{value.}$

1)  $\theta$  (estimate of parameters) is approximately unbiased estimator

$$E(\hat{\theta}) \approx \theta \rightarrow \text{True value}$$

2)  $\text{var}(\hat{\theta})$  : Nearly as small as the variance that could be obtained with any other estimators

$$\therefore \boxed{\text{MLE} \approx \text{MVUE}}$$

\* Large sample property:  $s^2 = \frac{1}{n-1} \sum_{i=1}^n (x_i - \mu)^2$

$$E[\hat{\sigma}^2] = \frac{n-1}{n} \sigma^2$$

$$\text{Bias} = E[\hat{\sigma}^2] - \sigma^2 = -\frac{\sigma^2}{n}$$

$n \rightarrow \infty$  (Large sample); Bias  $\rightarrow 0$

MLE: Underestimate of  $\sigma^2$

3) Invariance Property:

objective: Given sample,  $x_1, \dots, x_n$ , find an estimate for a function of  $\theta$ ,  $h(\theta)$ .

We can use the invariance property of MLE as follows.

If  $\hat{\theta}$  is an MLE of  $\theta$ , then  $h(\hat{\theta})$  is an MLE of  $h(\theta)$  when  $h(\cdot)$  is one-to-one function.

Note that  $y = f(x)$  is a one-to-one function when  $x = f^{-1}(y)$  or Existence of Inverse of  $f$ .

\* Consider  $h(\hat{\theta}) = \hat{\beta} \rightarrow \hat{\theta} = h^{-1}(\hat{\beta})$

\* Construct Likelihood function for  $\beta$

$$\begin{aligned} \max_{\beta} L^*(\beta) &= \max_{\beta} \prod_{i=1}^n f(\beta, x_i) \\ &= \max_{\beta} \prod_{i=1}^n f(h^{-1}(\beta), x_i) \\ &= \max_{\theta} \prod_{i=1}^n f(\theta, x_i) = \max_{\theta} L(\theta) \end{aligned}$$

\* Then, the value of  $\hat{\beta}$  that maximizes  $L^*(\beta)$  will be called an MLE of  $\beta = h(\theta)$ .

\* The invariance Property of MLE in

\* The invariance Property of MLE in multivariate cases.

If  $\hat{\theta}_1, \dots, \hat{\theta}_k$  are MLEs of  $\theta_1, \dots, \theta_k$ .

Then,  $h(\hat{\theta}_1, \dots, \hat{\theta}_k)$  is MLE of  $h(\theta_1, \dots, \theta_k)$ .