Linear Discriminant Analysis

B. Ravindran

Problem Setting

- $\mathcal{X} \subseteq \mathcal{R}^p$ is the input space
 - $\circ \; \mathbf{X} = (X_1,\, X_2,\, \dots,\, X_p)$ is a random variable describing the input
- $\mathcal{Y}\subseteq \Gamma$ is the output space with K number of classes.
 - **Y** is a random variable describing the output.
- ullet $\Pr(\mathbf{X}, \mathbf{Y}) = \Pr(\mathbf{Y} | \mathbf{X}) \Pr(\mathbf{X})$ is the data distribution
 - $oldsymbol{\circ} \ \Pr(\mathbf{Y} \,|\, \mathbf{X} = ar{x})$ is the predicted output probabilities given an input $ar{x} \in \mathcal{X}$

Problem Setting

We are interested in the probability of a class given the data point:

$$\Pr(\mathbf{Y} = k \,|\, \mathbf{X} = \bar{x})$$

• If the above probability is known for all K classes, we can predict the label as:

$$\hat{y} = rg \max_{k} \ \Pr(\mathbf{Y} = k \,|\, \mathbf{X} = \bar{x})$$

• We can write:

$$\Pr(\mathbf{Y}=k\,|\,\mathbf{X}=ar{x})\,=\,rac{\Pr(ar{x}\,|\,\mathbf{Y}=k)\Pr(\mathbf{Y}=k)}{\Pr(\mathbf{X}=ar{x})}$$

$$=rac{\Pr(ar{x}\,|\,\mathbf{Y}=k)\,\Pr(\mathbf{Y}=k)}{\sum_{k'=1}^K\,\Pr(ar{x}\,|\,\mathbf{Y}=k')\,\Pr(\mathbf{Y}=k')}$$

Some notation:

$$f_k(ar{x}) = \Pr(ar{x} \,|\, \mathbf{Y} = k)$$
 and $\Pi_k = \Pr(\mathbf{Y} = k)$

Therefore,

$$ext{Pr}(\mathbf{Y} = k \, | \, \mathbf{X} = ar{x}) \, = \, rac{f_k(ar{x})\Pi_k}{\sum_{k'=1}^K \, f_{k'}(ar{x})\Pi_{k'}}$$

Some notation:

$$f_k(ar{x}) = \Pr(ar{x} \,|\, \mathbf{Y} = k)$$
 and $\Pi_k = \Pr(\mathbf{Y} = k)$

Therefore,

$$\Pr(\mathbf{Y} = k \,|\, \mathbf{X} = ar{x}) \, = \, rac{f_k(ar{x})\Pi_k}{\sum_{k'=1}^K \,f_{k'}(ar{x})\Pi_{k'}} \Pi_k = \, rac{\sum_{i=1}^N \mathbb{1}_{\{y_i = k\}}}{N}$$

Some Notation:

$$f_k(ar{x}) = \Pr(ar{x} \,|\, \mathbf{Y} = k)$$
 and $\Pi_k = \Pr(\mathbf{Y} = k)$

Therefore,

$$\Pr(\mathbf{Y} = k \,|\, \mathbf{X} = ar{x}) \,=\, rac{f_k(ar{x})\Pi_k}{\sum_{k'=1}^K f_{k'}(ar{x})\Pi_{k'}}$$

Depending upon different assumptions we make on f_k we get different models.

We can assume any probabilistic form on f_k . Some of the commonly used forms are:

- Gaussian
- Mixture of Gaussian
- Non-Parametric
- Naive Bayes

LDA Assumption

We will assume f_k to be

Gaussian

$$f_k(ar{x}) \, = rac{1}{(2\pi)^{1/p} |\Sigma_k|^{1/2}} \mathrm{exp} \left(-rac{1}{2} (ar{x} - ar{\mu}_k)^T \Sigma_k^{-1} (ar{x} - ar{\mu}_k)
ight) \, .$$

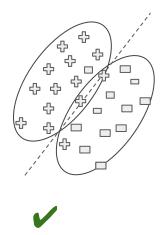
we will also assume, all covariance matrices are equal

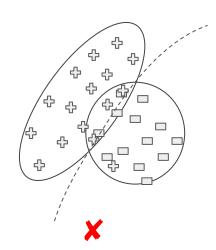
$$\forall k \;\; \Sigma_k = \Sigma$$

LDA Assumption

• The class conditional probability, $\Pr(\bar{x} \mid \mathbf{Y} = k)$, can have distinct means while sharing the same covariance matrix (shape).

Implication:





Class Boundary

• Class boundary between classes "k" and "l" is defined as:

$$\log \left(\frac{\Pr(\mathbf{Y} = k \,|\, \mathbf{X} = \bar{x})}{\Pr(\mathbf{Y} = l \,|\, \mathbf{X} = \bar{x})} \right) = 0$$

$$\Longrightarrow \log\left(rac{f_k(ar{x})\Pi_k}{f_l(ar{x})\Pi_l}
ight) = \log\left(rac{f_k(ar{x})}{f_l(ar{x})}
ight) + \log\left(rac{\Pi_k}{\Pi_l}
ight) = 0$$

$$\Longrightarrow \; \log\left(rac{\Pi_k}{\Pi_l}
ight) - rac{1}{2}(ar{\mu}_k + ar{\mu}_l)^T \Sigma^{-1}(ar{\mu}_k - ar{\mu}_l) \, + \, ar{x}^T \Sigma^{-1}(ar{\mu}_k - ar{\mu}_l) = 0$$

$$(:: \Sigma_k = \Sigma_l = \Sigma)$$

Class Boundary

• The boundary is a linear function in \bar{x}

$$\implies \log\left(rac{\Pi_k}{\Pi_l}
ight) - rac{1}{2}(ar{\mu}_k + ar{\mu}_l)^T\Sigma^{-1}(ar{\mu}_k - ar{\mu}_l) \,+\, ar{x}^T\Sigma^{-1}(ar{\mu}_k - ar{\mu}_l) = 0$$

$$\implies \delta_k(\bar{x}) - \delta_l(\bar{x}) = 0$$

where,
$$\delta_k(\bar{x}) = \log\left(\Pi_k\right) - \frac{1}{2}\bar{\mu}_k\Sigma^{-1}\bar{\mu}_k + \bar{x}^T\Sigma^{-1}\bar{\mu}_k$$

Class Prediction

ullet The data point, $ar{x}$, belongs to class "k" if

$$egin{aligned} & \Pr(\mathbf{Y}=k \,|\, \mathbf{X}=ar{x}) \,>\, \Pr(\mathbf{Y}=l \,|\, \mathbf{X}=ar{x}) \,,\,\,orall\, l
ot= k \,|\, \mathbf{X}=ar{x}) \ & \Longrightarrow \, \log \left(rac{\Pr(\mathbf{Y}=k \,|\, \mathbf{X}=ar{x})}{\Pr(\mathbf{Y}=l \,|\, \mathbf{X}=ar{x})}
ight) > 0 \,,\,\,orall\, l
ot= k \ \ & \Longrightarrow \, \delta_k(ar{x}) - \delta_l(ar{x}) > 0 \,,\,\,orall\, l
ot= k \ \end{aligned}$$

Class Prediction

ullet The data point, $ar{x}$, belongs to class "k" if

$$egin{aligned} & \Pr(\mathbf{Y} = k \,|\; \mathbf{X} = ar{x}) \,> \, \Pr(\mathbf{Y} = l \,|\; \mathbf{X} = ar{x}) \,,\; orall \, l
ot= k \,|\; \mathbf{X} = ar{x}) \ & \Longrightarrow \, \log \left(rac{\Pr(\mathbf{Y} = k \,|\; \mathbf{X} = ar{x})}{\Pr(\mathbf{Y} = l \,|\; \mathbf{X} = ar{x})}
ight) > 0 \,,\; orall \, l
ot= k \ \ & \Longrightarrow \, \delta_k(ar{x}) - \delta_l(ar{x}) > 0 \,,\; orall \, l
ot= k \ \end{aligned}$$

Therefore, class prediction for any data point:

$$\hat{y} = rg \max_{k} \, \delta_k(ar{x})$$

Estimating mean and covariance

$$\delta_k(ar{x}) = \log\left(\Pi_k
ight) - rac{1}{2}ar{\mu}_k\Sigma^{-1}ar{\mu}_k \,+\, ar{x}^T\Sigma^{-1}ar{\mu}_k$$

$$egin{aligned} \Pi_k = rac{\sum_{i=1}^N \mathbb{1}_{\{y_i = k\}}}{N} & egin{aligned} ar{\mu}_k = rac{\sum_{i=1}^N \mathbb{1}_{\{y_i = k\}} ar{x}_k}{\sum_{i=1}^N \mathbb{1}_{\{y_i = k\}}} \end{aligned}$$

$$\Sigma = rac{\sum_{k=1}^K \sum_{i=1}^N \mathbb{1}_{\{y_i = k\}} (ar{x}_i - ar{\mu}_k) (ar{x}_i - ar{\mu}_k)^T}{N - K}$$

Estimating mean and covariance

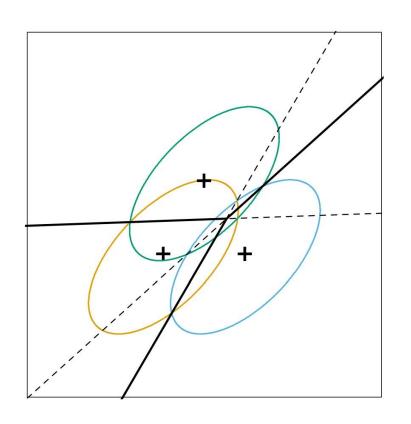
$$\delta_k(ar{x}) = \log\left(\Pi_k
ight) - rac{1}{2}ar{\mu}_k\Sigma^{-1}ar{\mu}_k \,+\, ar{x}^T\Sigma^{-1}ar{\mu}_k$$

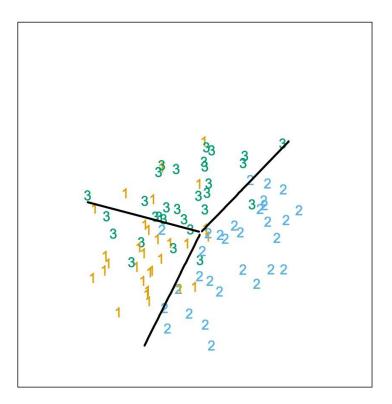
$$egin{aligned} \Pi_k = rac{\sum_{i=1}^N \mathbb{1}_{\{y_i = k\}}}{N} & ar{\mu}_k = rac{\sum_{i=1}^N \mathbb{1}_{\{y_i = k\}} ar{x}_k}{\sum_{i=1}^N \mathbb{1}_{\{y_i = k\}}} \end{aligned}$$

$$\Sigma = rac{\sum_{k=1}^K \sum_{i=1}^N \mathbb{1}_{\{y_i = k\}} (ar{x}_i - ar{\mu}_k) (ar{x}_i - ar{\mu}_k)^T}{N - K}$$

Pooled Estimate

LDA Example





Alternative View - Feature construction

By choosing a direction in which means are maximally spread apart.

LDA can also be seen as <u>maximizing the variance between the classes</u> and minimizing the variance within the classes.

Consider a 2 class classification problem, where

$$egin{array}{ll} \hat{y} = ar{w}^Tar{x} & \hat{y} > w_0 \,, & class \, 1 \, (C_1) \ & else, & class \, 2 \, (C_2) \end{array}$$

 $\bar{\mu}_1$ and $\bar{\mu}_2$ are means of class 1 and 2 respectively.

also define projected means as $\mu_1=\bar{w}^T\bar{\mu}_1$ and $\mu_2=\bar{w}^T\bar{\mu}_2$ within class variance is defined as

$$s_k^2 = \sum_{i \in \mathcal{G}} \left(ar{w}^T ar{x}_i \, - \, ar{w}^T ar{\mu}_k
ight)^2 \quad ext{, for} \quad k \in \{1, \, 2\}$$

We want to find "w" vector such that within class variance is minimized and the class means are maximally separated.

$$egin{align} ar{w}^{\star} &= rg \max_{ar{w}} \ J(ar{w}) \ := \ rac{(\mu_1 - \mu_2)^2}{s_1^2 + s_2^2} \ &= rac{ar{w}^T \, S_B \, ar{w}}{ar{w}^T \, S_W \, ar{w}} \end{aligned}$$

$$S_B = (ar{\mu}_1 - ar{\mu}_2)(ar{\mu}_1 - ar{\mu}_2)^T$$

$$S_W \ = \ \sum_{i \in C} (ar{x}_i - ar{\mu}_1) (ar{x}_i - ar{\mu}_1)^T + \sum_{i \in C} (ar{x}_i - ar{\mu}_2) (ar{x}_i - ar{\mu}_2)^T$$

$$ar{w}^\star = rg \max_{ar{w}} \, J(ar{w}) \, := \, rac{ar{w}^T \, S_B \, ar{w}}{ar{w}^T \, S_W \, ar{w}}$$

$$\left(ar{w}^{T} S_{B} \, ar{w}
ight) S_{W} \, ar{w} \, = \, \left(ar{w}^{T} S_{W} \, ar{w}
ight) S_{B} \, ar{w}$$

$$ar{w}^\star = rg \max_{ar{w}} \, J(ar{w}) \, := \, rac{ar{w}^T \, S_B \, ar{w}}{ar{w}^T \, S_W \, ar{w}}$$

$$\left(ar{w}^TS_B\,ar{w}
ight)S_W\,ar{w}\,=\,\left(ar{w}^TS_W\,ar{w}
ight)S_B\,ar{w}$$

are some scalar values

$$ar{w}^\star = rg \max_{ar{w}} \, J(ar{w}) \, := \, rac{ar{w}^T \, S_B \, ar{w}}{ar{w}^T \, S_W \, ar{w}}$$

$$\left(ar{w}^T S_B \, ar{w}
ight) S_W \, ar{w} \, = \, \left(ar{w}^T S_W \, ar{w}
ight) S_B \, ar{w}$$

$$\implies S_W \, ar w \, = \, \lambda \, S_B \, ar w$$

$$ar{w}^\star = rg \max_{ar{w}} \, J(ar{w}) \, := \, rac{ar{w}^T \, S_B \, ar{w}}{ar{w}^T \, S_W \, ar{w}}$$

$$(ar{w}^T S_B \, ar{w}) \, S_W \, ar{w} \, = \, (ar{w}^T S_W \, ar{w}) \, S_B \, ar{w}$$

$$\longrightarrow S_{rrr} a \overline{v} - \lambda S_{rr} a \overline{v}$$

$$\implies S_W \, ar w \, = \, \lambda \, S_B \, ar w$$

$$\implies S_W \, ar{w} \, = \, \lambda \, (ar{\mu}_1 - ar{\mu}_2) (ar{\mu}_1 - ar{\mu}_2)^{\, T} ar{w} \qquad \qquad :: S_B = (ar{\mu}_1 - ar{\mu}_2) (ar{\mu}_1 - ar{\mu}_2)^{\, T}$$

$$ar{w}^\star = rg \max_{ar{w}} \, J(ar{w}) \, := \, rac{ar{w}^T \, S_B \, ar{w}}{ar{w}^T \, S_W \, ar{w}}$$

$$\left(ar{w}^T S_B \, ar{w}
ight) S_W \, ar{w} \, = \, \left(ar{w}^T S_W \, ar{w}
ight) S_B \, ar{w}$$

$$\implies S_W \, \bar{w} \, = \, \lambda \, S_B \, \bar{w}$$

some scalar value

$$ar{w}^\star = rg \max_{ar{w}} \, J(ar{w}) \, := \, rac{ar{w}^T \, S_B \, ar{w}}{ar{w}^T \, S_W \, ar{w}}$$

taking derivative of
$$J(\mathbf{w})$$
 w.r.t. to \mathbf{w} and equating it to zero, we get

$$\implies S_W ar{w} = \lambda \, S_B \, ar{w}$$

$$\longrightarrow S_{rr} a \overline{v} - \lambda (\overline{u} - \overline{u}) (\overline{u} - \overline{u})^T a \overline{v}$$

 $(\bar{w}^T S_B \, \bar{w}) \, S_W \, \bar{w} \, = \, (\bar{w}^T S_W \, \bar{w}) \, S_B \, \bar{w}$

$$\implies S_W \, ar w \, = \, \lambda \, (ar \mu_1 - ar \mu_2) (ar \mu_1 - ar \mu_2)^{\,T} ar w$$

$$\implies S_W\, ar w \,=\, \lambda'\, (ar \mu_1 - ar \mu_2) \qquad \implies ar w \propto S_W^{-1} (ar \mu_1 - ar \mu_2)$$

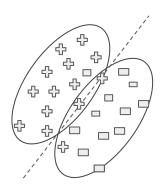
Alternative View

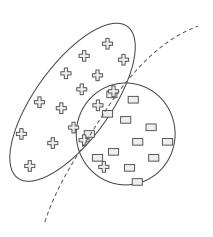
Find "w" vector such that within class variance is minimized and the class means are maximally separated.

Quadratic Discriminant

- What if the classes do not have the same covariance matrix?
- The quadratic term in the discriminant does not cancel out
- LDA -> QDA
 - Estimate covariance matrices separately

Quadratic Discriminant





Summary

- Assumes the class conditional density is a Gaussian
 - Same covariance for linear
 - Different covariance for quadratic
- Near ideal if the data comes from Gaussian distributions
 - Good approximation even if the assumption is violated
- It can be viewed as a feature selection method