Chapter 13

Concentration of sub-Gaussian random variables

Definition 13.1 (Sub-gaussian random variables¹). A random variable m is called sub-gaussian if its MGF is dominated² by a mean zero normal random variable with variance σ^2 , i.e.,

$$\mathbb{E}\left[e^{\lambda m}\right] \le e^{\frac{\lambda^2 \sigma^2}{2}}.\tag{13.1}$$

m is sometimes called a σ -sub-gaussian or a sub-gaussian with proxy σ^2 . A direct consequence of the definition 1 is that a sub-gaussian random variable has zero mean and its variance is bounded above by σ^2 .

Proposition 13.1. *If* m *is* a σ -sub-gaussian, then $\mathbb{E}[m] = 0$ and $\mathbb{V}ar[m] \leq \sigma^2$.

Proof. For any λ , using Taylor expansion for both sides of (13.1) we have

Why the first equality is true?

$$\sum_{n=0}^{\infty} \frac{\lambda^n}{n!} \mathbb{E}\left[m^n\right] = \mathbb{E}\left[e^{\lambda m}\right] \le e^{\frac{\lambda^2 \sigma^2}{2}} = \sum_{n=0}^{\infty} \frac{\lambda^{2n} \sigma^{2n}}{2^n n!}.$$

Now dividing both sides by $\lambda > 0$ and taking the limit $\lambda \to 0$, we conclude that $\mathbb{E}[m] = 0$. Next, dividing both sides by λ^2 and again taking the limit $\lambda \to 0$ we can conclude that $\mathbb{V}ar[m] \le \sigma^2$.

The following theorem characterizes sub-gaussian random variables.

Theorem 13.1 (Sub-gaussian properties). Let m be a random variable. Then the following are equivalent:

i) There exists a constant c_1 such that the tail of m satisfies

$$\mathbb{P}[|m| \ge t] \le 2e^{-\frac{t^2}{2c_1^2}}, \quad \forall t \ge 0.$$
 (13.2)

$$\mathbb{P}[|m| \ge t] \le 2e^{-t^2/c^2},$$

for some constant c.

 $^{^{1}}$ A weaker definition is based on the tail bound, i.e., m is a sub-gaussian random variable if

² That is, the Laplace transform of *m* is dominated by the Laplace transform of $\mathcal{N}(0, \sigma^2)$.

ii) There exists a constant c_2 such that the moments of m satisfy

$$||m||_p \stackrel{def}{=} (\mathbb{E}[|m|^p])^{1/p} \le c_2 \sqrt{p}, \quad \forall p \ge 1.$$

iii) There exists a constant c_3 such that the MGF of m^2 satisfies

$$\mathbb{E}\left[e^{\lambda^2 m^2}\right] \leq e^{c_3^2 \lambda^2}, \quad \forall \, |\lambda| \leq \frac{1}{c_3}.$$

Moreover, if $\mathbb{E}[m] = 0$ *, then all the above are equivalent to*

iv) There exists a constant c4 such that the MGF of m satisfies

$$\mathbb{E}\left[e^{\lambda m}\right] \leq e^{\sigma^2 \lambda^2/2}, \quad \forall \lambda \in \mathbb{R}.$$

Proof. We prove $1 \Rightarrow 2, 2 \Rightarrow 3, 3 \Rightarrow 1$, and $4 \Rightarrow 1$.

• $1 \Rightarrow 2$): Using Proposition 12.1 we have

$$\mathbb{E}[|X|^{p}] = \int_{0}^{\infty} \mathbb{P}[|m|^{p} > t] dt = \int_{0}^{\infty} \mathbb{P}[|m| > u] p u^{p-1} du \le 2 \int_{0}^{\infty} p u^{p-1} e^{-\frac{u^{2}}{2c_{1}^{2}}} du$$

$$= \left(\sqrt{2}c_{1}\right)^{p} 2 \int_{0}^{\infty} p t^{p-1} e^{-t^{2}} dt \le \left(\sqrt{2}c_{1}\right)^{p} p \Gamma(p/2) \le \left(\sqrt{2}c_{1}\right)^{p} p \left(\frac{p}{2}\right)^{p/2},$$

where we have used a change of variable $t = u^p$ in the second equality, i) in the first inequality, definition of the Gamma function in the third equality, and a Stirling's approximation in the second inequality. Taking the pth-root both side and using the fact that $p^{1/p} < 2$ ends the proof with $c_2 = 2c_1$.

• $2 \Rightarrow 3$): By Taylor expansion and the monotone convergence theorem we have

$$\mathbb{E}\left[e^{\lambda^2 m^2}\right] = 1 + \sum_{k=1}^{\infty} \frac{\lambda^{2k}}{k!} \mathbb{E}\left[m^{2k}\right] \le 1 + \sum_{k=1}^{\infty} \frac{\lambda^{2k}}{\left(k/e\right)^k} \left(8c_1^2\right)^k k^k = \frac{1}{1 - 8c_1^2 e \lambda^2},$$

where we have used ii) and a Stirling approximation $k! \ge (k/e)^k$ in the first inequality. The last equality holds provided that $8c_1^2e\lambda^2 \le 1$. We can further bound the right hand side if we take λ such that $8c_1^2e\lambda^2 \le 1/2$ and use the inequality $(1-x)^{-1} \le e^{2x}$ for $x \in [0,1/2]$, i.e.,

$$\mathbb{E}\left[e^{\lambda^2 m^2}\right] \le e^{16c_1^2 e \lambda^2},$$

which ends the proof by taking $c_3^2 = 16c_1^2e$.

• $3 \Rightarrow 1$): we have

$$\mathbb{P}[|m| \ge t] = \mathbb{P}\left[e^{m^2} \ge e^{t^2}\right] \stackrel{Markov}{\le} \frac{\mathbb{E}\left[e^{m^2}\right]}{e^{t^2}},$$

which ends the proof by applying *iii*) with $\lambda = c_3 = 1$.

• $4 \Rightarrow 1$): Using Chernoff inequality (12.1) we have

$$\mathbb{P}[m \ge t] \le \min_{\lambda > 0} \frac{\mathbb{E}\left[e^{\lambda m}\right]}{e^{\lambda t}} \le \min_{\lambda > 0} e^{\sigma^2 \lambda^2 / 2 - \lambda t} = e^{-\frac{t^2}{2\sigma^2}},$$

where we have used iv) in the second inequality.

Remark 13.1. Note that the first three assertions are equivalent for any random variable. The equivalence says that if a random variable has a exponential decaying tail bound of the form (13.2), not only its expectation (see Proposition 12.1) is bounded, but its L^p -norm grows like $\mathcal{O}(\sqrt{p})$. It also says that the exponential decaying tail bound (13.2) is necessary and sufficient for the integrability of fast growing function $e^{\lambda^2 m^2}$. Assertion iv) tells us that sub-gaussian random variables have all these properties, which is not surprising since a Gaussian random variable is also a sub-gaussian random variable.

Similar to Gaussian distributions, a finite sum of sub-gaussian random variables is sub-gaussian.

Proposition 13.2 (Sum of independent sub-gaussians). Assume m_1, \ldots, m_N are independent, sub-gaussian random variables with proxy c_i^2 , then $\sum_{i=1}^N \mathbf{a}_i m_i$ is also a sub-gaussian random variable with proxy $\sum_{i=1}^N c_i^2 \mathbf{a}_i^2$.

Proof. The proof is straightforward:

$$\mathbb{E}\left[e^{\lambda \sum_{i=1}^{N} \mathbf{a}_i m_i}\right] = \Pi_{i=1}^{N} \mathbb{E}\left[e^{\lambda \mathbf{a}_i m_i}\right] \leq e^{\lambda^2 \sum_{i=1}^{N} c_i^2 \mathbf{a}_i^2}.$$