

Autocovariance function

Autocovariance function gives a statistical characterization of the dependence between the sequence of random variables $X_t, X_{t-1}, X_{t-2}, \dots$

Distribution Properties of a_t

Fixed t : a_t is a random variable

all t : a_t is a stochastic process

$a_t \sim \text{NID}(0, \sigma_a^2)$ $E(a_t) = 0$; $\text{Var}(a_t) = \sigma_a^2$; $\text{Cov}(a_j, a_i) = 0$ if $i \neq j$

Covariance of a stochastic process with itself at different values of t , called **autocovariance**.

In general,

$$\text{Cov}(a_t, a_{t-k}) = E(a_t - \mu)(a_{t-k} - \mu)$$

since $\mu = 0$

$$\text{Cov}(a_t, a_{t-k}) = E(a_t a_{t-k}) = \begin{cases} \sigma_a^2 & k=0 \\ 0 & k \neq 0 \end{cases}$$

$$E(a_t a_{t-k}) = \delta_k \sigma_a^2 \quad \delta_k = \begin{cases} 0 & k \neq 0 \\ 1 & k = 0 \end{cases}$$

Autocovariance and autocorrelation function

$$E(a_t a_{t-k}) = \delta_k \sigma_a^2 \quad \delta_k = \begin{cases} 0 & k \neq 0 \\ 1 & k = 0 \end{cases}$$

$$\gamma_k = E(X_t X_{t-k}) \quad \rho_k = \frac{\gamma_k}{\gamma_0} \quad \text{with } \rho_0 = 1$$

covariance between a and X :

$$X_t = \sum_{j=0}^{\infty} G_j a_{t-j}$$

$$E(a_t X_t) = \sigma_a^2 \quad \text{since } G_0 = 1$$

$$E(a_t X_{t-k}) = 0, \quad \text{for } k > 0 \quad X_{t-k} = \sum_{j=0}^{\infty} G_j a_{t-k-j}$$

$$E(a_t X_{t-k}) = \delta_k \sigma_a^2$$

$$\gamma_k = E(X_t X_{t-k}) = E(X_{t-k} X_t) = E(X_t X_{t+k}) = \gamma_{-k}$$

1). For AR(1) model:

$$X_t = \phi_1 X_{t-1} + a_t$$

$$\gamma_k = E(X_t X_{t-k}) = \phi_1 E(X_{t-1} X_{t-k}) + E(a_t X_{t-k})$$

$$\gamma_0 = \phi_1 \gamma_1 + \sigma_a^2$$

$$\gamma_k = \phi_1 \gamma_{k-1} \quad k > 0$$

$$\gamma_0 = \frac{\sigma_a^2}{1 - \phi_1^2} = \text{Var}(X_t)$$

$$\rho_0 = 1 \quad \rho_k = \phi_1 \rho_{k-1}$$

implicit method

$$G_j = \phi_1^j$$

$$\gamma_k = E(X_t X_{t-k}) = E\left[\left(\sum_{i=0}^{\infty} G_i a_{t-i}\right)\left(\sum_{j=0}^{\infty} G_j a_{t-(j+k)}\right)\right] = \left(\sum_{j=0}^{\infty} G_{j+k} G_j\right) \sigma_a^2 = \left(\phi_1^k \sum_{j=0}^{\infty} \phi_1^{2j}\right) \sigma_a^2$$

$$= \frac{\sigma_a^2}{1 - \phi_1^2} \phi_1^k$$

$$\rho_0 = 1 \quad \rho_k = \phi_1^k$$

explicit method

2). For MA(1) model:

$$X_t = a_t - \theta_1 a_{t-1}$$

$$E(X_t X_{t-k}) = E(a_t X_{t-k}) - \theta_1 E(a_{t-1} X_{t-k})$$

since $G_0=1$, $G_1=-\theta_1$, $G_j=0$ for $j \geq 2$

$$\gamma_0 = \sigma_a^2 - \theta_1 G_1 \sigma_a^2 = \sigma_a^2 + \theta_1^2 \sigma_a^2$$

$$\gamma_1 = -\theta_1 \sigma_a^2$$

$$\gamma_k = 0 \quad \text{for } k \geq 2$$

$$\rho_0 = 1 \quad \rho_1 = -\frac{\theta_1}{1 + \theta_1^2} \quad \rho_k = 0 \quad k \geq 2$$

$$X_t = \phi_1 X_{t-1} + \phi_2 X_{t-2} - \theta_1 a_{t-1} + a_t$$

$$E(X_t X_{t-k}) = \phi_1 E(X_{t-1} X_{t-k}) + \phi_2 E(X_{t-2} X_{t-k}) - \theta_1 E(a_{t-1} X_{t-k}) + E(a_t X_{t-k})$$

$$k=0: \gamma_0 = \phi_1 \gamma_1 + \phi_2 \gamma_2 - \theta_1 G_1 \sigma_a^2 + \sigma_a^2$$

$$k=1: \gamma_1 = \phi_1 \gamma_0 + \phi_2 \gamma_1 - \theta_1 \sigma_a^2$$

$$k: \gamma_k = \phi_1 \gamma_{k-1} + \phi_2 \gamma_{k-2}, \quad k \geq 2$$

implicit method, recursive

To find γ_k , γ_0 and γ_1 must be found first by simultaneously solving for γ_0 and γ_1 . (p.120)

$$G_j = g_1 \lambda_1^j + g_2 \lambda_2^j$$

$$\begin{aligned} \gamma_k &= E(X_t X_{t-k}) = E\left[\left(\sum_{i=0}^{\infty} G_i a_{t-i}\right)\left(\sum_{j=0}^{\infty} G_j a_{t-(j+k)}\right)\right] = \left(\sum_{j=0}^{\infty} G_{j+k} G_j\right) \sigma_a^2 \\ &= \sigma_a^2 \sum_{j=0}^{\infty} (g_1 \lambda_1^{j+k} + g_2 \lambda_2^{j+k}) (g_1 \lambda_1^j + g_2 \lambda_2^j) \\ &= \sigma_a^2 \left[\frac{g_1^2}{1-\lambda_1^2} \lambda_1^k + \frac{g_2^2}{1-\lambda_2^2} \lambda_2^k + \frac{g_1 g_2}{1-\lambda_1 \lambda_2} (\lambda_1^k + \lambda_2^k) \right] \end{aligned}$$

$$\gamma_0 = \sigma_a^2 \left[\frac{g_1^2}{1-\lambda_1^2} + \frac{g_2^2}{1-\lambda_2^2} + \frac{2g_1 g_2}{1-\lambda_1 \lambda_2} \right]$$

$$\gamma_k = \sigma_a^2 \left[\frac{g_1^2}{1-\lambda_1^2} + \frac{g_1 g_2}{1-\lambda_1 \lambda_2} \right] \lambda_1^k + \sigma_a^2 \left[\frac{g_2^2}{1-\lambda_2^2} + \frac{g_1 g_2}{1-\lambda_1 \lambda_2} \right] \lambda_2^k$$

$$\gamma_0 = d_1 + d_2$$

$$\gamma_k = d_1 \lambda_1^k + d_2 \lambda_2^k$$

Remarks:

- The total variance γ_0 in an ARMA(2,1) system is decomposed into two components d_1 and d_2 . d_1 and d_2 are respectively associated with the characteristic roots λ_1 and λ_2 .
- There are similarities between Green's function and the autocovariance function.

Given: a_t 's $X_t = \sum_{j=0}^{\infty} G_j a_{t-j}$ convolution

Given: $E(a_t a_{t-k})$ $\gamma_k = \sigma_a^2 \sum_{j=0}^{\infty} G_j G_{j+k}$ double convolution

4). General Results

ARMA(n,m) model:

$$X_t = \phi_1 X_{t-1} + \phi_2 X_{t-2} + \dots + \phi_n X_{t-n} + a_t - \theta_1 a_{t-1} - \dots - \theta_m a_{t-m}$$

$$\gamma_0 = \phi_1 \gamma_1 + \phi_2 \gamma_2 + \dots + \phi_n \gamma_n + (1 - \theta_1 G_1 - \theta_2 G_2 - \dots - \theta_m G_m) \sigma_a^2$$

$$\gamma_1 = \phi_1 \gamma_0 + \phi_2 \gamma_1 + \dots + \phi_n \gamma_{n-1} + (-\theta_1 G_0 - \theta_2 G_1 - \dots - \theta_m G_{m-1}) \sigma_a^2$$

.....

$$\gamma_m = \phi_1 \gamma_{m-1} + \phi_2 \gamma_{m-2} + \dots + \phi_n \gamma_{m-n} - \theta_m \sigma_a^2$$

$$\gamma_k = \phi_1 \gamma_{k-1} + \phi_2 \gamma_{k-2} + \dots + \phi_n \gamma_{k-n} \quad k \geq m+1$$

$$\Phi(B) \gamma_k = 0, \quad k \geq m+1$$

For an ARMA(n,n-1) model,

$$\gamma_k = d_1 \lambda_1^k + d_2 \lambda_2^k + \dots + d_n \lambda_n^k$$

d can be found in Eqs.(3.3.26) and (3.1.26)

$$\gamma_0 = d_1 + d_2 + \dots + d_n$$

Applications in machine condition monitoring and diagnostics.

Theoretical/Sample Autocovariance/Autocorrelation function

$$\gamma_k = E(X_t X_{t-k}) \quad \rho_k = \frac{\gamma_k}{\gamma_0} \quad \text{with } \rho_0 = 1$$

$$\hat{\gamma}_k = \frac{1}{N} \sum_{t=k+1}^N X_t X_{t-k} = \frac{1}{N} \sum_{t=k+1}^N (\dot{X}_t - \bar{X})(\dot{X}_{t-k} - \bar{X})$$

$$\hat{\rho}_k = \frac{\hat{\gamma}_k}{\hat{\gamma}_0}$$

Remarks:

- Sample autocovariance/autocorrelation functions can be obtained directly from the data before fitting a model.
- The use of autocorrelation functions for modeling and estimation of time series would be appropriate if a good estimate can be obtained.
- Sample autocorrelation functions often are very poor estimates and have large variances and present a distorted version of the true autocorrelations.