Machine Learning 2 - Homework 5

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May 6, 2018

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Problem 1

1. For deriving the update rules, we take the derivative of the expected value of the completedata log-likelihood with regard to the parameter to be updated and set it to zero. Meanwhile, we keep the term $\gamma(z_{nk})$ fixed and treat it as a constant.

For deriving the update rule for π , we need to introduce a Lagrangian Multiplier in order to satisfy the constraint $\sum_{k=1}^{K} \pi_k = 1$:

$$\mathbb{E}_{\text{posterior}} = \sum_{n=1}^{N} \sum_{k=1}^{K} \gamma(z_{nk}) \ln \pi_k + \lambda (\sum_{k=1}^{K} \pi_k - 1) + \text{const}$$

$$\frac{\partial \mathbb{E}_{\text{posterior}}}{\partial \pi_k} = \sum_{n=1}^{N} \frac{\gamma(z_{nk})}{\pi_k} + \lambda = 0$$

$$N_k = -\lambda \pi_k$$

$$\sum_{k=1}^{K} N_k = -\sum_{k=1}^{K} \lambda \pi_k$$

$$N = -\lambda$$

$$\pi_k = \frac{N_k}{N}$$

where $N_k = \sum_{n=1}^N \gamma(z_{nk})$ is the effective number of data points associated with component k.

For μ we get:

$$\mathbb{E}_{\text{posterior}} = \sum_{n=1}^{N} \sum_{k=1}^{K} \gamma(z_{nk}) \left[-\frac{1}{2} (\mathbf{x}_{n}^{T} \boldsymbol{\Sigma}_{k}^{-1} \mathbf{x}_{n} - 2\boldsymbol{\mu}_{k}^{T} \boldsymbol{\Sigma}_{k}^{-1} \mathbf{x}_{n} + \boldsymbol{\mu}_{k}^{T} \boldsymbol{\Sigma}_{k}^{-1} \boldsymbol{\mu}_{k}) \right] + \text{const}$$

$$\frac{\partial \mathbb{E}_{\text{posterior}}}{\partial \boldsymbol{\mu}_{k}} = -\frac{1}{2} \sum_{n=1}^{N} \gamma(z_{nk}) (-2\boldsymbol{\Sigma}_{k}^{-1} \mathbf{x}_{n} + 2\boldsymbol{\Sigma}_{k}^{-1} \boldsymbol{\mu}_{k}) = 0$$

$$\sum_{n=1}^{N} \gamma(z_{nk}) \boldsymbol{\Sigma}_{k}^{-1} \mathbf{x}_{n} = \sum_{n=1}^{N} \gamma(z_{nk}) \boldsymbol{\Sigma}_{k}^{-1} \boldsymbol{\mu}_{k}$$

$$\sum_{n=1}^{N} \gamma(z_{nk}) \boldsymbol{\Sigma}_{k}^{-1} \mathbf{x}_{n} = N_{k} \boldsymbol{\Sigma}_{k}^{-1} \boldsymbol{\mu}_{k}$$

$$\boldsymbol{\mu}_{k} = \frac{\sum_{n=1}^{N} \gamma(z_{nk}) \mathbf{x}_{n}}{N_{k}}$$

And for Σ (using equations (57) and (61) from the Matrix Cookbook and $\Sigma_k \Sigma_k^{-1} = I$):

$$\mathbb{E}_{\text{posterior}} = -\frac{1}{2} \sum_{n=1}^{N} \sum_{k=1}^{K} \gamma(z_{nk}) \left[\ln(|\mathbf{\Sigma}_{k}|) + (\mathbf{x}_{n} - \boldsymbol{\mu}_{k})^{T} \mathbf{\Sigma}_{k}^{-1} (\mathbf{x}_{n} - \boldsymbol{\mu}_{k}) \right] + \text{const}$$

$$\frac{\partial \mathbb{E}_{\text{posterior}}}{\partial \mathbf{\Sigma}_{k}} = -\frac{1}{2} \sum_{n=1}^{N} \gamma(z_{nk}) \left[\mathbf{\Sigma}_{k}^{-1} - \mathbf{\Sigma}_{k}^{-1} (\mathbf{x}_{n} - \boldsymbol{\mu}_{k}) (\mathbf{x}_{n} - \boldsymbol{\mu}_{k})^{T} \mathbf{\Sigma}_{k}^{-1} \right] = 0$$

$$\sum_{n=1}^{N} \gamma(z_{nk}) \mathbf{\Sigma}_{k}^{-1} = \sum_{n=1}^{N} \gamma(z_{nk}) \mathbf{\Sigma}_{k}^{-1} (\mathbf{x}_{n} - \boldsymbol{\mu}_{k}) (\mathbf{x}_{n} - \boldsymbol{\mu}_{k})^{T} \mathbf{\Sigma}_{k}^{-1}$$

$$\mathbf{\Sigma}_{k} N_{k} \mathbf{\Sigma}_{k}^{-1} \mathbf{\Sigma}_{k} = \sum_{n=1}^{N} \gamma(z_{nk}) \mathbf{\Sigma}_{k} \mathbf{\Sigma}_{k}^{-1} (\mathbf{x}_{n} - \boldsymbol{\mu}_{k}) (\mathbf{x}_{n} - \boldsymbol{\mu}_{k})^{T} \mathbf{\Sigma}_{k}^{-1} \mathbf{\Sigma}_{k}$$

$$\mathbf{\Sigma}_{k} = \frac{\sum_{n=1}^{N} \gamma(z_{nk}) (\mathbf{x}_{n} - \boldsymbol{\mu}_{k}) (\mathbf{x}_{n} - \boldsymbol{\mu}_{k})^{T}}{N_{k}}$$

2. When constraining all covariance matrices to have a common value Σ , the update rules for π and μ remain the same as they are not dependent on Σ . For the update rule of Σ , we get:

$$\mathbb{E}_{\text{posterior}} = -\frac{1}{2} \sum_{n=1}^{N} \sum_{k=1}^{K} \gamma(z_{nk}) \left[\ln(|\mathbf{\Sigma}|) + (\mathbf{x}_n - \boldsymbol{\mu}_k)^T \mathbf{\Sigma}^{-1} (\mathbf{x}_n - \boldsymbol{\mu}_k) \right] + \text{const}$$

$$\frac{\partial \mathbb{E}_{\text{posterior}}}{\partial \mathbf{\Sigma}} = -\frac{1}{2} \sum_{n=1}^{N} \sum_{k=1}^{K} \gamma(z_{nk}) \left[\mathbf{\Sigma}^{-1} - \mathbf{\Sigma}^{-1} (\mathbf{x}_n - \boldsymbol{\mu}_k) (\mathbf{x}_n - \boldsymbol{\mu}_k)^T \mathbf{\Sigma}^{-1} \right] = 0$$

$$\sum_{n=1}^{N} \sum_{k=1}^{K} \gamma(z_{nk}) \mathbf{\Sigma}^{-1} = \sum_{n=1}^{N} \sum_{k=1}^{K} \gamma(z_{nk}) \mathbf{\Sigma}^{-1} (\mathbf{x}_n - \boldsymbol{\mu}_k) (\mathbf{x}_n - \boldsymbol{\mu}_k)^T \mathbf{\Sigma}^{-1}$$

$$\mathbf{\Sigma} = \frac{\sum_{n=1}^{N} \sum_{k=1}^{K} \gamma(z_{nk}) (\mathbf{x}_n - \boldsymbol{\mu}_k) (\mathbf{x}_n - \boldsymbol{\mu}_k)^T}{N}$$

Where $\sum_{n=1}^{N} \sum_{k=1}^{K} \gamma(z_{nk}) = \sum_{k=1}^{K} N_k = N$.

Problem 2

Using the dependencies as indicated in the graphical model, the posterior distribution $p(\theta|\mathbf{X})$ can be rewritten as:

$$\begin{aligned} p(\boldsymbol{\theta}|\mathbf{X}) &\propto p(\mathbf{X}|\boldsymbol{\theta})p(\boldsymbol{\theta}) \\ &= \sum_{\mathbf{Z}} p(\mathbf{X}, \mathbf{Z}|\boldsymbol{\theta})p(\boldsymbol{\theta}) \\ &\propto \ln \sum_{\mathbf{Z}} p(\mathbf{X}, \mathbf{Z}|\boldsymbol{\theta}) + \ln p(\boldsymbol{\theta}) \end{aligned}$$

When applying the EM-Algorithm, we maximize the posterior over the latent variables \mathbf{Z} in the E-Step, while keeping the parameters $\boldsymbol{\theta}$ fixed. This gives us:

$$\arg \max_{\mathbf{Z}} p(\boldsymbol{\theta}|\mathbf{X}) \propto \arg \max_{\mathbf{Z}} \ln \sum_{\mathbf{Z}} p(\mathbf{X}, \mathbf{Z}|\boldsymbol{\theta}) + \ln p(\boldsymbol{\theta})$$
$$\propto \arg \max_{\mathbf{Z}} \ln \sum_{\mathbf{Z}} p(\mathbf{X}, \mathbf{Z}|\boldsymbol{\theta})$$

Since the prior $p(\theta)$ is independent of **Z**, we can drop the term and end up with a maximization over the log-likelihood. Therefore, the E-Step remains the same as in the maximum likelihood case.

In the M-Step, we maximize the posterior over the parameters θ , while keeping the latent variables **Z** fixed. This gives us:

$$\mathop{\arg\max}_{\boldsymbol{\theta}} p(\boldsymbol{\theta}|\mathbf{X}) \propto \mathop{\arg\max}_{\boldsymbol{\theta}} \ln \sum_{\mathbf{Z}} p(\mathbf{X},\mathbf{Z}|\boldsymbol{\theta}) + \ln p(\boldsymbol{\theta})$$

Here, we want to maximize over the complete-data log-likelihood. In practice, however, we are not given the complete data set $\{\mathbf{X}, \mathbf{Z}\}$. The only knowledge that we have about the latent variables \mathbf{Z} is given only by the posterior distribution $p(\mathbf{Z}|\mathbf{X}, \boldsymbol{\theta})$. As we can not make use of the complete-data log-likelihood directly, we use its expected value under the posterior distribution of the latent variables instead. This gives us the quantity to be maximized in the M-Step as follows:

$$\begin{split} \arg\max_{\boldsymbol{\theta}} p(\boldsymbol{\theta}|\mathbf{X}) &\propto \arg\max_{\boldsymbol{\theta}} \underset{p(\mathbf{Z}|\mathbf{X},\boldsymbol{\theta}^{\text{old}})}{\mathbb{E}} \left[\ln\sum_{\mathbf{Z}} p(\mathbf{X},\mathbf{Z}|\boldsymbol{\theta}) \right] + \ln p(\boldsymbol{\theta}) \\ &\approx \arg\max_{\boldsymbol{\theta}} \sum_{\mathbf{Z}} p(\mathbf{Z}|\mathbf{X},\boldsymbol{\theta}^{\text{old}}) \ln p(\mathbf{X},\mathbf{Z}|\boldsymbol{\theta}) + \ln p(\boldsymbol{\theta}) \end{split}$$

Problem 3

For the M-Step, we need to derive the update rules for π and μ . For this, we use the log-posterior:

$$\ln p(\boldsymbol{\mu}, \boldsymbol{\pi} | \{x_n\}_{n=1}^N) \propto \sum_{n=1}^N \sum_{k=1}^K \gamma(z_{nk}) \left(\ln \boldsymbol{\pi}_k + \sum_{i=1}^D \left[x_{ni} \ln \mu_{ki} + (1 - x_{ni}) \ln(1 - \mu_{ki}) \right] \right)$$

$$+ \sum_{k=1}^K \ln p(\boldsymbol{\mu}_k | a_k, b_k) + \ln p(\boldsymbol{\pi} | \boldsymbol{\alpha})$$

$$= \sum_{n=1}^N \sum_{k=1}^K \gamma(z_{nk}) \left(\ln \boldsymbol{\pi}_k + \sum_{i=1}^D \left[x_{ni} \ln \mu_{ki} + (1 - x_{ni}) \ln(1 - \mu_{ki}) \right] \right)$$

$$+ \sum_{k=1}^K (a_k - 1) \ln \boldsymbol{\mu}_k + (b_k - 1) \ln(1 - \boldsymbol{\mu}_k) + \sum_{k=1}^K (\alpha_k - 1) \ln \boldsymbol{\pi}_k + \text{const}$$

For the update rule for π , we need to introduce a Lagrangian Multiplier in order to fulfill the constraint $\sum_{k=1}^{K} \pi_k = 1$:

$$\frac{\partial \ln p(\mu, \pi | \{x_n\}_{n=1}^N) - \lambda(\sum_{k=1}^K \pi_k - 1)}{\partial \pi_k} = \frac{\sum_{n=1}^N \gamma(z_{nk})}{\pi_k} + \frac{\alpha_k - 1}{\pi_k} - \lambda = 0$$

$$N_k + \alpha_k - 1 = \lambda \pi_k$$

$$\sum_{k=1}^K (N_k + \alpha_k - 1) = \lambda$$

$$N + \sum_{k=1}^K \alpha_k - K = \lambda$$

$$\pi_k = \frac{N_k + \alpha_k - 1}{N + \sum_{k=1}^K \alpha_k - K}$$

For the update rule of μ_k we get:

$$\frac{\partial \ln p(\boldsymbol{\mu}, \boldsymbol{\pi} | \{x_n\}_{n=1}^N)}{\partial \boldsymbol{\mu}_k} = \sum_{n=1}^N \gamma(z_{nk}) \left(\sum_{i=1}^D \frac{x_{ni}}{\mu_{ki}} - \frac{1 - x_{ni}}{1 - \mu_{ki}} \right) + \frac{a_k - 1}{\boldsymbol{\mu}_k} - \frac{b_k - 1}{1 - \boldsymbol{\mu}_k} = 0$$

$$\sum_{n=1}^N \gamma(z_{nk}) \left(\frac{\mathbf{x}_n}{\boldsymbol{\mu}_k} \right) + \frac{a_k - 1}{\boldsymbol{\mu}_k} = \sum_{n=1}^N \gamma(z_{nk}) \left(\frac{1 - \mathbf{x}_n}{1 - \boldsymbol{\mu}_k} \right) + \frac{b_k - 1}{1 - \boldsymbol{\mu}_k}$$

$$(1 - \boldsymbol{\mu}_k) \left(\sum_{n=1}^N \gamma(z_{nk}) \mathbf{x}_n + a_k - 1 \right) = \boldsymbol{\mu}_k \left(\sum_{n=1}^N \gamma(z_{nk}) (1 - \mathbf{x}_n) + b_k - 1 \right)$$

$$\sum_{n=1}^N \gamma(z_{nk}) \mathbf{x}_n + a_k - 1 = \boldsymbol{\mu}_k \left(\sum_{n=1}^N \gamma(z_{nk}) + b_k - 1 + a_k - 1 \right)$$

$$\boldsymbol{\mu}_k = \frac{\sum_{n=1}^N \gamma(z_{nk}) \mathbf{x}_n + a_k - 1}{N_k + b_k + a_k - 2}$$