

Problem set - IV

Theory: $\underline{X} = (x_1, x_2, \dots, x_n)$: $S \rightarrow \mathbb{R}^n$
 \underline{X} is Random vector iff x_1, x_2, \dots, x_n are random variables.

c.d.f.

Joint

c.d.f.

p.m.f.

Joint

p.m.f.

p.d.f.

Joint

p.d.f.

$$F_x(x) = P(x \leq x)$$

$$F_{\underline{x}}(x_1, x_2, \dots, x_n) = P(x_1 \leq x_1, x_2 \leq x_2, \dots, x_n \leq x_n)$$

$$\boxed{\frac{x, y}{F_x(x, y)}} = P(x \leq x, y \leq y)$$

$$F_x(x, y) = P(y \leq y) = F_y(y)$$

$$\lim_{x \rightarrow +\infty} F_x(x, y) = F_x(x)$$

$$\lim_{y \rightarrow \infty} F_x(x, y) = F_x(x)$$

$$\lim_{x \rightarrow -\infty} F_x(x, y) = 0 = \lim_{y \rightarrow -\infty} F_x(x, y)$$

$$\lim_{x \rightarrow \infty} F_x(x, y) = 1$$

$$y \rightarrow \infty$$

for each $(a_1, b_1] \times (a_2, b_2]$ in \mathbb{R}^2

$$\Delta = F_{\underline{x}}(b_1, b_2) - F_{\underline{x}}(a_1, b_2) - F_{\underline{x}}(a_2, b_1) + F_{\underline{x}}(a_1, a_2)$$

why?

Proof:-

$$P(a_1 < x_1 \leq b_1, a_2 < x_2 \leq b_2)$$

$$= P(x \leq b_1, a_2 < x_2 \leq b_2) - P(x \leq a_1, a_2 \leq x_2 \leq b_2)$$

↓

$$P(x \leq b_1, x_2 \leq b_2) - P(x \leq b_1, x_2 \leq a_2)$$

$$- P(x_1 \leq a_1, x_2 \leq b_2) + P(x_1 \leq a_1, x_2 \leq a_2)$$

$$\Rightarrow F_{\underline{x}}(b_1, b_2) - F_{\underline{x}}(a_1, b_2) - F_{\underline{x}}(a_1, b_1) + F_{\underline{x}}(a_1, a_2) \geq 0$$

P.m.f. or P.d.f.

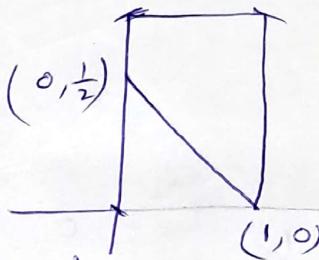
$$f_x(x) = P(x = x)$$

Here it becomes like this
 $f_{\underline{x}}(x_1, x_2, \dots, x_n) = P(x_1 = x_1, x_2 = x_2, \dots, x_n = x_n)$

~~classmate~~

$$F(x, y) = \begin{cases} \frac{2}{3} & \text{if } x+2y \geq 1 \\ 0 & \text{if } x+2y < 1 \end{cases}$$

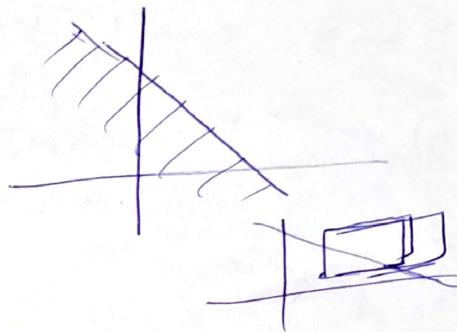
$$[0, 1] \times [0, 1]$$



$$F(1,1) - F(0,1) - F(1,0) + F(0,0) \\ = 1 - 1 - 1 + 0 = -1 < 0 \Rightarrow \underline{\text{Not}}$$

ii)

$$F(x, y) = \begin{cases} 0 & \text{if } x < 0 \text{ or } y < 0 \\ 1 & \text{o.w.} \end{cases}$$



$$[\frac{1}{3}, 1] \times [\frac{1}{3}, 1]$$

$$F(1,1) - F(\frac{1}{3}, 1) - F(1, \frac{1}{3}) + F(\frac{1}{3}, \frac{1}{3})$$

$$1 - 1 - 1 + 0 = -1 < 0 \quad \underline{\text{Not}} \text{ a c.d.f.}$$

(2)

Suppose a fair coin is tossed three times

$x = \# \text{ of Heads}$ is a discrete type Random vector

$y = \# \text{ absolute difference of Heads & tails}$

To prove: $Z = (x, y)$ is a discrete type Random vector

$$x(\omega) = \begin{cases} 0 & \text{if } \omega \in \{TTT\} \\ 1 & \text{if } \omega \in \{TTH, HTT, THT\} \\ 2 & \text{if } \omega \in \{HHT, HTH, THH\} \\ 3 & \text{if } \omega \in \{HHH\} \end{cases}$$

O.w.

$$y(\omega) = \begin{cases} 1 & \text{if } \omega \in \{HHH, TTT\} \\ 3 & \text{if } \omega \in \{HHT, HTH, THH\} \end{cases}$$

$x \backslash y$	0	1	2	3	$P(Y = \cdot)$
0	0	$\frac{3}{8}$	$\frac{3}{8}$	0	$\frac{6}{8}$
1	$\frac{1}{8}$	0	0	$\frac{1}{8}$	$\frac{2}{8}$
$P(X = x)$	$\frac{1}{8}$	$\frac{3}{8}$	$\frac{3}{8}$	$\frac{1}{8}$	1

$(0, 1) \rightarrow 0 \rightarrow TTT \notin \Sigma$

$(1, 1) \rightarrow HTT, THT, TTH$

Joint p.m.f.

$$f_{\Sigma}(x, y) = \begin{cases} 0 & \text{o.w.} \\ \frac{1}{8} & \text{if } (x, y) \in \{(0, 0), (3, 3)\} \\ \frac{3}{8} & \text{if } (x, y) \in \{(1, 1), (2, 1)\} \end{cases}$$

Q Let A & B be two events
 & two R.V. X & Y be
 $X(\omega) = \begin{cases} 1 & \text{if } \omega \in A} \\ 0 & \text{o.w.} \end{cases}$

$$Y(\omega) = \begin{cases} 1 & \text{if } \omega \in B} \\ 0 & \text{o.w.} \end{cases}$$

To prove: X & Y are independent $\Leftrightarrow A$ & B are independent.

Soln: Joint p.m.f. $f(x, y) = \begin{cases} P(A^c \cap B^c) & \text{if } (x, y) = (0, 0) \\ P(A \cap B^c) & \text{if } (x, y) = (1, 0) \\ P(A^c \cap B) & \text{if } (x, y) = (0, 1) \\ P(A \cap B) & \text{if } (x, y) = (1, 1) \\ 0 & \text{o.w.} \end{cases}$

$$\text{p.m.f. of } x \text{ & } y \text{ are}$$

$$P_x(x) = \begin{cases} P(A^c) & \text{if } x=0 \\ P(A) & \text{if } x=1 \\ 0 & \text{o.w.} \end{cases}$$

$$P_y(y) = \begin{cases} P(B^c) & \text{if } y=0 \\ P(B) & \text{if } y=1 \\ 0 & \text{o.w.} \end{cases}$$

\Rightarrow Let x & y are Independent then

$$p(x, y) = p_x(x) p_y(y)$$

$$p(1, 1) = p_x(1) p_y(1)$$

$$p(A \cap B) = P(A) P(B)$$

\Leftarrow let A & B are Independent
 $\Rightarrow A^c$ & B , A & B^c & A^c & B^c are independent
 $\Rightarrow p(x, y) = p_x(x) p_y(y)$
 $\Rightarrow x$ & y are Independent

(5) x & y be two R.V. with Joint p.d.f.

$$f(x, y) = \begin{cases} \frac{1+xy}{4} & \text{if } |x| < 1, |y| < 1 \\ 0 & \text{o.w.} \end{cases}$$

To prove: x & y are not independent & x^2, y^2 are independent

$$f_x(x) = \int_{-\infty}^{\infty} f(x, y) dy$$

$$= \begin{cases} \int_{-1}^1 \frac{1+xy}{4} dy & \text{if } -1 < x < 1 \\ 0 & \text{o.w.} \end{cases}$$

$$= \begin{cases} \frac{1}{2} & \text{if } -1 < x < 1 \\ 0 & \text{o.w.} \end{cases}$$

$$\text{if } -1 < y < 1$$

similarly $f_y(y) = \begin{cases} \frac{1}{2} & \text{if } -1 < y < 1 \\ 0 & \text{o.w.} \end{cases}$

$f(x, y) \neq f_x f_y \Rightarrow x$ & y are not independent.



x^2 & y^2 are Independent

$$P(x^2 \leq u, y^2 \leq v) = P(-u^{1/2} \leq x \leq u^{1/2}, -v^{1/2} \leq y \leq v^{1/2})$$

$$= \int_{-v^{1/2}}^{v^{1/2}} \int_{-u^{1/2}}^{u^{1/2}} \frac{1+xy}{4} dx dy$$

$$= u^{1/2} v^{1/2}$$

for $0 \leq u \leq 1$
 $0 \leq v \leq 1$

$$P(x^2 \leq u) = P(-u^{1/2} \leq x \leq u^{1/2})$$

$$= \int_{-u^{1/2}}^{u^{1/2}} \frac{1}{2} dx = u^{1/2}$$

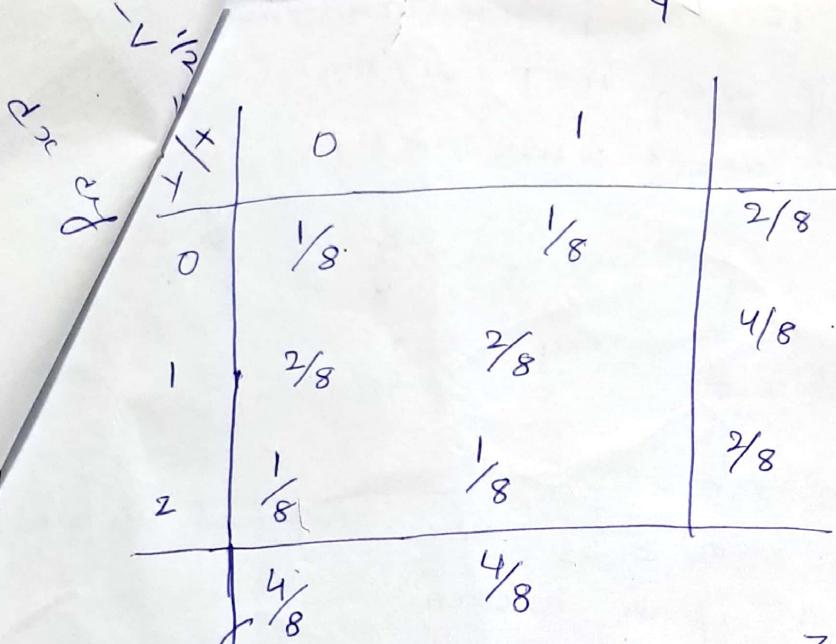
$$F_{x^2}(x) = \begin{cases} u^{1/2} & 0 \leq u \leq 1 \\ 0 & \text{o.w.} \end{cases}$$

$$F_{y^2}(y) = \begin{cases} v^{1/2} & 0 \leq v \leq 1 \\ 0 & \text{o.w.} \end{cases}$$

$$\Rightarrow F_z(x, y) = F_{x^2}(x) F_{y^2}(y)$$

- ⑥ Toss a fair coin 3 times
- $x = \# \text{ of Heads on first toss}$
- $y = \# \text{ of Heads on last two tosses}$
- $z = \# \text{ of Heads on first two tosses}$

- a) Find Joint p.m.f. for x & y . Compute $\text{cov}(x, y)$
- b) Find Joint p.m.f. for x & z compute $\text{cov}(x, z)$
- Soln: $x(\omega) = \begin{cases} 0 & \text{if } \omega \in \{\text{THT, TTH, THH, TTT}\} \\ 1 & \text{if } \omega \in \{\text{HTT, HHT, HTH, HHTH}\} \end{cases}$
- $y(\omega) = \begin{cases} 0 & \text{if } \omega \in \{\text{TTT, HTT}\} \\ 1 & \text{if } \omega \in \{\text{TTH, THT, HTH, HTHT}\} \\ 2 & \text{if } \omega \in \{\text{HHH, THH}\} \end{cases}$



Joint p.m.f. $Z = (x, y)$

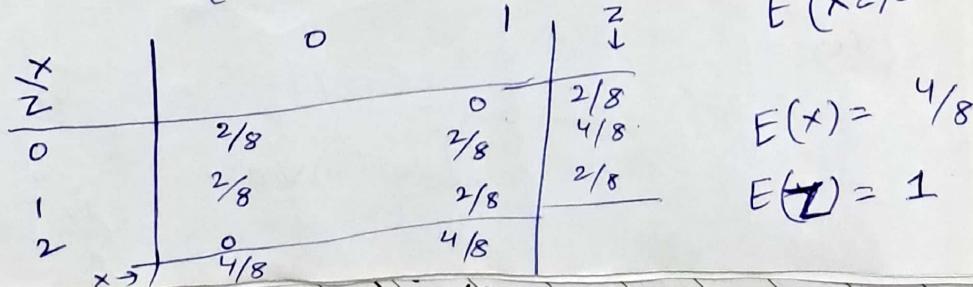
$$p_Z(x, y) = \begin{cases} \frac{1}{8} & \text{if } (x, y) \in \{(0, 0), (1, 1)\} \\ \frac{2}{8} & \text{if } (x, y) \in \{(0, 1), (1, 0)\} \\ 0 & \text{o.w.} \end{cases}$$

$$p_X(x) = \begin{cases} \frac{1}{2} & \text{if } x=0 \\ \frac{1}{2} & \text{if } x=1 \\ 0 & \text{o.w.} \end{cases}$$

$$p_Y(y) = \begin{cases} \frac{2}{8} & \text{if } y=0, 2 \\ \frac{1}{2} & \text{if } y=1 \\ 0 & \text{o.w.} \end{cases}$$

$$p_Z = p_X p_Y \Rightarrow x \text{ & } y \text{ are indep.} \Rightarrow \text{cov}(x, y) = 0$$

$$(b) z(\omega) = \begin{cases} 0 & \text{if } \omega \in \{\text{TTT}, \text{TTH}\} \\ 1 & \text{if } \omega \in \{\text{HTH}, \text{THH}, \text{HTT}, \text{HTH}\} \\ 2 & \text{if } \omega \in \{\text{HHH}, \text{HHT}\} \end{cases}$$



$$E(XZ) =$$

$$E(X) = 4/8$$

$$E(Z) = 1$$

back

Joint
 $p(x, y)$

$$x \cdot z(\omega) = \begin{cases} 0 & \text{if } \omega \in \{\text{THT, TTH, THH}\} \\ 1 & \text{if } \omega \in \{\text{HTT, HTH}\} \\ 2 & \text{if } \omega \in \{\text{HHH, HHT}\} \end{cases}$$

$$E(xz) = 0 \times \frac{1}{8} + 1 \times \frac{2}{8} + 2 \times \frac{1}{8} = \frac{3}{8}$$

$$\text{Cov}(x, z) = \frac{6}{8} - \frac{9}{8} = \frac{1}{4}$$

⑦ Roll a dice ($n=1, 2, \dots, 6$)
Two events s_1 & s_2 are defined as follows

$$s_1 = \begin{cases} 1 & \text{if } n \text{ is even} \\ 0 & \text{otherwise} \end{cases}$$

$$s_2 = \begin{cases} 1 & \text{if } n \text{ is prime} \\ 0 & \text{o.w.} \end{cases}$$

find Joint p.m.f. $P(s_1, s_2)$

Also covariance & correlation coefficient between s_1 & s_2 .

s_1	0	1	
s_2	$\frac{1}{6}$	$\frac{2}{6}$	$\frac{3}{6}$
0	$\frac{1}{6}$	$\frac{1}{6}$	$\frac{3}{6}$
1	$\frac{2}{6}$	$\frac{1}{6}$	$\frac{3}{6}$

$$E(s_1, s_2) = \frac{1}{6}$$

$$E(s_1) = \frac{3}{6}$$

$$E(s_2) = \frac{3}{6}$$

Joint p.m.f.
we can write

$$E(s_1^2) = \frac{3}{6} \quad E(s_2^2) = \frac{3}{6}$$

$$\text{Cov}(s_1, s_2) = \frac{1}{6} - \frac{9}{36} = -\frac{1}{12}$$

$$\text{Var}(s_1) = \frac{3}{6} - \frac{9}{36} = \frac{9}{36} \Rightarrow \sqrt{\text{Var}(s_1)} = \frac{1}{2}$$

$$\sqrt{\text{Var}(s_2)} = \frac{1}{2}$$

$$\Rightarrow \rho(s_1, s_2) = \frac{\text{Cov}(s_1, s_2)}{\sqrt{\text{Var}(s_1)\text{Var}(s_2)}} = \frac{-\frac{1}{12}}{\frac{1}{4}} = -\frac{1}{3}$$

Joint p.m.f. of x & y is

$$p(x, y) = \begin{cases} c(2x+y) & \text{if } (x, y) \in \{(0, 0), (0, 1), (1, 0), (1, 1), (2, 0)\} \\ 0 & \text{o.w.} \end{cases}$$

- a) find c & $P(X \geq 1, Y \leq 2)$
- b) $P(X=2, Y=1)$
- c) find Marginal p.m.f. of x & y
- d) Are x & y independent

Soln: $c = \frac{1}{42}$

$$P(X=2, Y=1) = \frac{5}{42}$$

$$\begin{aligned} P(X \geq 1, Y \leq 2) &= P(X=1, Y \leq 2) + P(X=2, Y \leq 2) \\ &= P(X=1, Y=0) + P(X=1, Y=1) + P(X=1, Y=2) \\ &\quad + P(X=2, Y=0) + P(X=2, Y=1) + P(X=2, Y=2) \\ &= \frac{4}{7} \end{aligned}$$

9) $p_x(x) = \begin{cases} p(0,0) + p(0,1) + p(0,2) & \text{if } x=0 \\ p(1,0) + p(1,1) + p(1,2) & \text{if } x=1 \\ 0 & \text{if } x=2 \end{cases}$ o.w.

similarly we can find p_y

product $\not\propto \Rightarrow$ Not Ind

⑨ $\text{Cov}(ax+b, y) = a \text{Cov}(x, y)$
 $\text{Var}(ax+b) = a^2 \text{Var}(x)$

$$P(ax+b, y) = \frac{\text{cov}(ax+b, y)}{\sqrt{\text{var}(ax+b), y}} = \frac{a \text{cov}(x, y)}{a \sqrt{\text{var}(x) \text{var}(y)}} = P(x, y)$$

$$T = \frac{5x}{9} - \frac{160}{9} \quad S = \frac{5y}{9} - \frac{160}{9}$$

$$\rho(x, y) = 0.8$$

$$\text{cov}(x, y) = 4 \quad \Rightarrow \rho(T, S) = 0.8$$

~~cov(T, S)~~ formula:

$$\text{cov}(a_1x + b_1, a_2y + b_2) = a_1a_2 \text{cov}(x, y)$$

$$\Rightarrow \text{cov}(T, S) = \frac{25}{81} \cos(x, y) = \frac{100}{81}$$

10) $x = \begin{cases} 4 & \text{if drawn card is heart} \\ \text{---} & \end{cases}$

11) ~~12) 13) 14)~~

12) ~~13)~~ Let x, y be Random vector with Joint p.d.f.

$$f(x, y) = \begin{cases} cx+1 & \text{if } xy \geq 0 \text{ & } x+y < 1 \\ 0 & \text{o.w.} \end{cases}$$

a) find the value of c

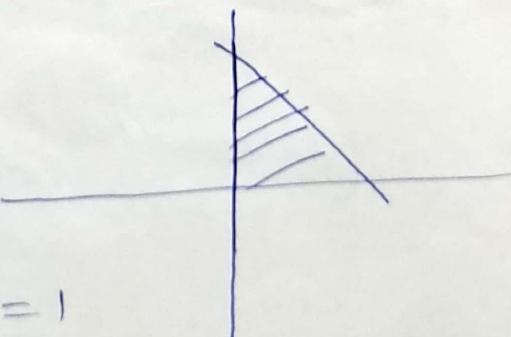
b) find the Marginal p.d.f. of x & y

c) $P(X < 2x^2)$

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(x, y) = 1$$

$$\Rightarrow \int_0^1 \int_0^{1-x} (cx+1) dy dx = 1$$

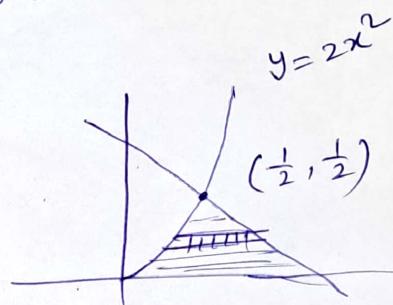
$$\Rightarrow c = 3$$



$$f_x(x) = \begin{cases} \int_{-\infty}^6 f(x,y) dy & \text{if } 0 \leq x < 1 \\ \int_0^{1-x} (3x+1) dy & \text{o.w.} \end{cases}$$

$$= \begin{cases} (3x+1)(1-x) & \text{if } 0 \leq x \leq 1 \\ 0 & \text{o.w.} \end{cases}$$

$$\text{Q} \quad P(Y < 2x^2)$$



~~$$= \iint_A f(x,y) dA = \frac{53}{96}$$~~

AND

(19)

Probability distribution for no. of eggs in a clutch is $P(\lambda)$ & probability that each egg will hatch is p . Show by calculation that probability distribution of no. of chicks that hatch is $P(\lambda p)$.

λ = no. of eggs ~ $P(\lambda)$

y = no. of chicks hatching

~~$f(y=y)$~~ Probability of k chicks hatching from a clutch of size n is $\binom{n}{k} p^k (1-p)^{n-k}$

Probability of clutch of size n is $\frac{e^{-\lambda} \lambda^n}{n!}$

$$T(\alpha) \cdot x \quad T(\alpha) \cdot x \quad \dots \quad x_{-1}$$

Probability of K chicks hatching from a clutch
is

$$Pr(K \text{ chicks}) = \sum_{n=K}^{\infty} \frac{e^{-\lambda} \lambda^n}{n!} n_C K p^K q^{n-K}$$

$$= \frac{e^{-\lambda} \lambda^P}{K!} (x P)^K$$

$$\Rightarrow Y \sim P(\lambda P)$$

Soln

(12) Let X denote the time until the miner reaches safety, and let Y denote the door he initially chooses. Now,

$$E[X] = E[E[X|Y]]$$

$$= \sum_y E[X|Y=y] P(Y=y)$$

$$= E[X|Y=1] P(Y=1) + E[X|Y=2] P(Y=2) \\ + E[X|Y=3] P(Y=3).$$

$$= \frac{1}{3} (E[X|Y=1] + E[X|Y=2] + E[X|Y=3])$$

↓ ↓ ↓
 2 3 + E[X] 5 + E[X]

$$\Rightarrow \boxed{E[X] = 10}$$

To understand why $E[X|Y=2] = 3 + E[X]$:

If miner chooses the second door, then he spends three hours in the tunnel and then returns to the mine.

But once he returned to the mine the problem is as before, and hence his expected additional time under safety is just $E[X]$. Hence $E[X|Y=2] =$

$$3 + E[X].$$

(14) At a party N men throw their hats into the center of a room. The hats are mixed up and each man randomly select one. Let X be the # of men those get their own hat. Find $E(X)$ & $\text{var}(X)$.

Solⁿ: $X = \# \text{ of men who got own hat}$

For $i = 1, 2, \dots, N$

$$X_i = \begin{cases} 1 & \text{if } i\text{th man get their own hat} \\ 0 & \text{o.w.} \end{cases}$$

Then $X = X_1 + X_2 + \dots + X_N$

For each i ,

$$\begin{aligned} E(X_i) &= 1 \cdot P(X_i=1) + 0 \cdot P(X_i=0) \\ &= \sum_{x=0}^1 x P(X_i=x) = \frac{1}{N} = E(X_i^2) \end{aligned}$$

$$\begin{aligned} E(X) &= E(X_1 + \dots + X_N) \\ &= E(X_1) + \dots + E(X_N) \\ &= N \times \frac{1}{N} = 1. \quad \checkmark \end{aligned}$$

$$\begin{aligned} \text{Var}(X) &= E(X^2) - E(X)^2 = 1. \quad \checkmark \\ E((X_1 + X_2 + \dots + X_N)^2) &= E(X_1^2) + \dots + E(X_N^2) \\ &\quad + 2(E(X_1 X_2) + \dots +) \end{aligned}$$

$$E(X_i X_j) = \sum_{(x,y) \in \{(0,0), (1,0), (0,1), (1,1)\}} x y P(X_i=x, X_j=y) = \frac{1}{N(N-1)}$$

$$E(X|Y=\frac{1}{2}) = \int_0^{\frac{1}{2}} x f_{X|Y}(x|y=\frac{1}{2}) dx$$

$$= \int_0^{\frac{1}{2}} x \frac{2x}{(\frac{1}{2})^2} dx = 8 \int_0^{\frac{1}{2}} x^2 dx$$

$$= \frac{8x^3}{3} \Big|_0^{\frac{1}{2}} = \frac{1}{3}$$

$$E(X^2|Y=\frac{1}{2}) = 8 \int_0^{\frac{1}{2}} x^2 \cdot 2x dx$$

$$= 8 \left(\frac{x^4}{4} \right) \Big|_0^{\frac{1}{2}} = \frac{2}{16} = \frac{1}{8}$$

Then $\text{Var}(X|Y=\frac{1}{2}) = \frac{1}{8} - \left(\frac{1}{3}\right)^2 = \frac{1}{8} - \frac{1}{9} = \frac{1}{72}$

$$P(0 < Y < \frac{1}{3}) = \int_0^{\frac{1}{3}} f_Y(y) dy = \int_0^{\frac{1}{3}} 4y^3 dy$$

$$= 4 \left(\frac{y^4}{4} \right) \Big|_0^{\frac{1}{3}} = \frac{1}{81}$$

$$P\left(\frac{1}{3} < Y < \frac{2}{3} | X=\frac{1}{2}\right) = \int_{\frac{1}{2}}^{\frac{2}{3}} f_{Y|X=\frac{1}{2}}(y) dy = \int_{\frac{1}{2}}^{\frac{2}{3}} \frac{2y}{1-(\frac{1}{2})^2} dy$$

$$= \frac{4}{3} \times 2 \left(\frac{y^2}{2} \right) \Big|_{\frac{1}{2}}^{\frac{2}{3}} = \frac{4}{3} \left(\frac{4}{9} - \frac{1}{4} \right)$$

$$= \frac{4}{3} \left(\frac{16-9}{36} \right) = \frac{4}{3} \times \frac{7}{36} = \frac{7}{27}$$

4) For $i=1, 2, \dots, N$ y_i i-th man get their own hat

$$X_i = \begin{cases} 1 & \text{if } y_i \text{ i-th man get their own hat} \\ 0 & \text{o.w.} \end{cases}$$

$$\text{Then } X = X_1 + X_2 + \dots + X_N$$

$$\text{For each } i, E(X_i) = \sum_{x=0}^1 x P(X_i=x) = 0 \cdot P(X_i=0) + 1 \cdot P(X_i=1)$$

$$\text{Thus } E(X) = E(X_1) + \dots + E(X_N) = N \cdot \frac{1}{N} = 1$$

$$\text{For } i \neq j \sum_{x,y} xy P(X_i=x, X_j=y)$$

$$E(X_i X_j) = \sum_{(x,y) \in \{(0,0), (1,0), (0,1), (1,1)\}}$$

$$= P(X_i=1, X_j=1) = \frac{1}{N(N-1)}$$

thus for $i=1$,

$$\text{Cov}(X_i, X_j) = E(X_i X_j) - E(X_i) E(X_j)$$

$$= \frac{1}{N(N-1)} - \frac{1}{N} \cdot \frac{1}{N} = \frac{1}{N} \left(\frac{1}{N(N-1)} \right)$$

Also $\forall i=1, 2, \dots, N$

$$\text{Var}(X_i) = E(X_i^2) - (E(X_i))^2$$

$$= P(X_i=1) - \frac{1}{N^2} = \frac{1}{N} - \frac{1}{N^2} = \frac{N-1}{N^2}$$

Hence

$$\text{Var}(X) = \sum_{i=1}^N \text{Var}(X_i) + 2 \sum_{\substack{1 \leq i < j \leq N}} \text{Cov}(X_i, X_j)$$

$$= N \cdot \left(\frac{(N-1)}{N^2} \right) + 2 \binom{N}{2} \frac{1}{N} \left(\frac{1}{N(N-1)} \right)$$

$$= \frac{N-1}{N} + \cancel{\frac{N(N-1)}{N}} \frac{1}{N} \left(\frac{1}{N(N-1)} \right)$$

$$= \frac{N-1}{N} + \frac{1}{N} = \frac{N-1+1}{N} = \frac{N}{N} = 1$$

Thus $\boxed{\text{Var}(X)=1}$

(3) The 10 randomly selected pages have independent distributions of errors per page. \rightarrow Also $\sum_{i=1}^m P(\lambda_i) \sim P\left(\sum_{i=1}^m \lambda_i\right)$.

For $i=1, 2, \dots, 10$ if the error in i -th page

$$X_i = \begin{cases} 1 & \text{if the error in } i\text{-th page} \\ 0 & \text{o.w.} \end{cases}$$

Then $X = \text{no. of errors in 10 randomly selected pages}$ where $X_i \sim P(\lambda)$
 $= X_1 + X_2 + \dots + X_{10}$ $\forall 1 \leq i \leq 10$

Thus $X \sim P(10\lambda)$.

Hence $P(X=10) = \frac{e^{-10\lambda} (10\lambda)^{10}}{10!}$

(b) X : - expected number of rolls

$$W.Y = \begin{cases} 1 & A \text{ wins on his 1st roll} \\ 2 & B \text{ wins on his 1st roll} \\ 3 & \text{otherwise neither A nor B wins their 1st roll} \end{cases}$$

Let X be the length after we break for the first time.
 Since the breakpoint is chosen uniformly over a piece of length Y ,
 we have $E[X|Y] = \frac{Y}{2}$.

For a similar reason, we have $E(Y) = \frac{l}{2}$

$$\text{Thus } E(X) = E(E(X|Y)) = E\left(\frac{Y}{2}\right) = \frac{1}{2}E(Y) = \frac{l}{4}.$$

(16) The c.d.f of Y_1 is

$$\begin{aligned} F_{Y_1}(y) &= P(Y_1 \leq y) \\ &= P(\max(X_1, X_2, \dots, X_n) \leq y) \\ &= P(X_1 \leq y, X_2 \leq y, \dots, X_n \leq y) \quad (\because X_i \text{ is continuous}) \\ &= \prod_{i=1}^n P(X_i \leq y) = \prod_{i=1}^n F_{X_i}(y) \\ &= (F(y))^n \end{aligned}$$

The c.d.f of Y_2 is

$$\begin{aligned} F_{Y_2}(y) &= P(Y_2 \leq y) = 1 - P(Y_2 > y) \\ &= 1 - P(\min(X_1, X_2, \dots, X_n) > y) \\ &= 1 - P(X_1 > y, X_2 > y, \dots, X_n > y) \\ &= 1 - \prod_{i=1}^n P(X_i > y) \\ &= 1 - \prod_{i=1}^n (1 - P(X_i \leq y)) \\ &= 1 - \prod_{i=1}^n (1 - F(y)) \\ F_{Y_2}(y) &= 1 - (1 - F(y))^n \end{aligned}$$

$$(19) \quad \text{The m.g.f of } Y \text{ is} \\ \phi_{X_i}(t) = e^{a_i t + \frac{\sigma_i^2 t^2}{2}}$$

Let $Y = \sum_{i=1}^n a_i X_i$.
Then the m.g.f of Y is

$$\begin{aligned} \phi_Y(t) &= E(e^{tY}) \\ &= E\left(e^{t(\sum_{i=1}^n a_i X_i)}\right) \\ &= E\left(\prod_{i=1}^n e^{ta_i X_i}\right) = \prod_{i=1}^n E(e^{ta_i X_i}) \\ &= \prod_{i=1}^n \phi_{X_i}(ta_i) \\ &= \prod_{i=1}^n e^{a_i \mu_i t + \frac{a_i^2 \sigma_i^2 t^2}{2}} \\ &= \prod_{i=1}^n \left((\sum a_i \mu_i) t + \frac{(\sum a_i^2 \sigma_i^2) t^2}{2} \right) \\ &= e^{\sum a_i \mu_i t + \frac{\sum a_i^2 \sigma_i^2 t^2}{2}} \end{aligned}$$

Hence $Y \sim N\left(\sum a_i \mu_i, \sum a_i^2 \sigma_i^2\right)$

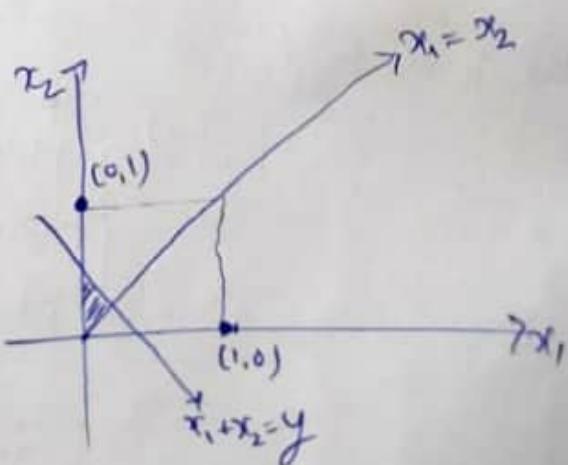
(20) Same as (19)

$$\begin{aligned} (21) \quad F_Y(y) &= P(X_1 + X_2 \leq y) \\ &= \int_0^1 \int_0^1 f(x_1, x_2) dx_1 dx_2 \\ &\quad \text{where } 0 \leq x_1 \leq x_2 \leq 1 \\ &\quad \{(x_1, x_2) : x_1 + x_2 \leq y\} \end{aligned}$$

(Clearly for $y < 0$ $F_Y(y) = 0$)

For $0 \leq y < 1$

$$\begin{aligned} F_Y(y) &= \int_0^y \left(\int_{x_1}^{y-x_1} 2 dx_2 \right) dx_1 \\ &= \int_0^y \left(2(y - 2x_1) \right) dx_1 \\ &= 2 \left[yx_1 - x_1^2 \right]_0^y = 2 \left(\frac{y^2}{2} - \frac{y^2}{4} \right) = \frac{y^2}{2} \end{aligned}$$



(21) Let (X, Y) be random vector with joint p.d.f.

$$f(x, y) = \begin{cases} cx+1 & \text{if } x, y \geq 0 \text{ and } x+y < 1 \\ 0 & \text{otherwise.} \end{cases}$$

(a) Find the value of c .

Soln:- To find the value of c , we write

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(x, y) dx dy = 1$$

$$\Rightarrow \int_{-\infty}^{\infty} \int_0^{1-x} (cx+1) dy dx = 1$$

$$\Rightarrow \int_{-\infty}^{\infty} (cx+1)(1-x) dx = 1$$

$$\Rightarrow \int_0^1 (cx+1)(1-x) dx = 1$$

$$\Rightarrow \frac{1}{2} + \frac{1}{6} c = 1 \Rightarrow \boxed{c = 3}$$

(b) Find the marginal pdf of X and Y .

$$f_X(x) = \int_{-\infty}^{\infty} f(x, y) dy$$

$$= \int_0^{1-x} (3x+1) dy = (3x+1)(1-x) \text{ for } x \in [0, 1]$$

$$\text{Thus, } f_x(x) = \begin{cases} (3x+1)(1-x) & 0 \leq x \leq 1 \\ 0 & \text{otherwise.} \end{cases}$$

Similarly, we obtain

$$\begin{aligned} f_y(y) &= \int_{-\infty}^{\infty} f(x, y) dx \\ &= \int_0^{1-y} (3x+1) dx \\ &= \frac{1}{2} (1-y)(5-3y) : y \in [0, 1]. \end{aligned}$$

$$\text{Thus, } f_y(y) = \begin{cases} \frac{1}{2} (1-y)(5-3y) : 0 \leq y \leq 1 \\ 0, \quad \text{o.w.} \end{cases}$$

$$\begin{aligned} (\text{C}) \text{ Find } P(Y < 2X^2) &= \int_{-\infty}^{\infty} \int_{-\infty}^{2x^2} f(x, y) dy dx \\ &= \int_0^1 \int_0^{\min(2x^2, 1-x)} (3x+1) dy dx \\ &= \int_0^1 (3x+1) \min(2x^2, 1-x) dx \\ &= \int_0^{1/2} (2x^2)(3x+1) dx + \int_{1/2}^1 (3x+1)(1-x) dx \\ &= \frac{53}{96}, \text{ ut} \end{aligned}$$

(22) Let (X, Y) be a random vector with joint pdf

$$f(x, y) = \begin{cases} 6e^{-(2x+3y)} & \text{if } x, y \geq 0 \\ 0 & \text{otherwise} \end{cases}$$

(a) Are X and Y independent? (\checkmark).

Soln:-

$$f_X(x) = \int_{-\infty}^{\infty} f(x, y) dy$$

$$= \int_0^{\infty} f(x, y) dy = \int_0^{\infty} 6e^{(-2x+3y)} dy$$

$$= \int_0^{\infty} 6e^{-2x} e^{3y} dy =$$

$$= 6e^{-2x} \int_0^{\infty} e^{3y} dy = 2 \cdot 6e^{-2x} \cdot \frac{e^{3y}}{3} \Big|_0^{\infty}$$
$$= -2e^{-2x} [e^{3y}]_0^{\infty}$$

$$= -2e^{-2x} [0 - 1]$$

$$= 2e^{-2x}, x \geq 0$$

Similarly, $f_Y(y) = 3e^{-3y}, y \geq 0$

(b) Find $E(Y|X>2)$.

Since, X and Y are independent, we have

$$E(Y|X>2) = E(Y).$$

Also, $Y \sim \text{Exp}(3) \Rightarrow E(Y) = \frac{1}{\lambda} = \frac{1}{3}$.

(c) Find $P(X>Y) = \int_0^{\infty} \int_y^{\infty} 6e^{-(2x+3y)} dx dy$

$$= \int_0^{\infty} 3e^{-5y} dy = \frac{3}{5}.$$

(23) Let X be a random variable with pdf

$$f_X(x) = \begin{cases} 2x & \text{if } 0 \leq x \leq 1 \\ 0 & \text{otherwise.} \end{cases}$$

We know that given $X=x$, the r.v. Y uniformly distributed on $[-x, x]$.

(a) Find the joint pdf of X and Y .

Soln:- Since, $f_{Y/X}(y/x) = \begin{cases} \frac{1}{x-(-x)} = \frac{1}{2x} & -x \leq y \leq x \\ 0 & \text{o.w.} \end{cases}$

Thus, we have

$$f_{X,Y}(x,y) = f_{Y|X}(y|x) f_X(x)$$

$$= \begin{cases} 1, & 0 \leq x \leq 1, -x \leq y \leq x \\ 0, & \text{otherwise.} \end{cases}$$

$$\Rightarrow f_{X,Y}(x,y) = \begin{cases} 1, & |y| \leq x \leq 1 \\ 0, & \text{otherwise.} \end{cases}$$

(b) Find pdf of Y.

$$f_Y(y) = \int_{-\infty}^{\infty} f_{X,Y}(x,y) dx$$

$$= \int_{|y|}^1 1 dx$$

$$= 1 - |y|, \quad -1 \leq y \leq 1$$

$$\therefore f_Y(y) = \begin{cases} 1 - |y|, & -1 \leq y < 1 \\ 0, & \text{otherwise.} \end{cases}$$

$$(c) \text{ Find } P(|Y| < X^3).$$

Soln :-

$$A = \left\{ (x, y) : |y| \leq x^3 \right\}$$

$$P(|Y| < x^2) = \iint_A f(x,y) dx dy$$

$$= \int_0^1 \left(\int_{x^3}^x 1 \, dy \right) dx$$

$$= \int_0^1 (2x^3) dx$$

(24) Let X be random variable with pdf

$$f_x(x) = \begin{cases} 4x(1-x^2), & 0 < x < 1 \\ 0, & \text{o.w.} \end{cases}$$

For a fixed $x \in (0, 1)$, the conditional pdf of Y given $X=x$ is

$$f_{Y|X}(y|x) = \begin{cases} \frac{2y}{1-x^2} & \text{if } x < y < 1 \\ 0 & \text{o.w.} \end{cases}$$

(a) Find the conditional pdf of X given $Y=y$ for appropriate values of y .

Soln:- $f_{X|Y}(x|y) = \frac{f(x,y)}{f_Y(y)}$

$$\text{Now, } f(x,y) = f_{Y|X}(y|x) f_X(x)$$

$$= \begin{cases} 8xy & \text{if } 0 < x < y < 1 \\ 0 & \text{o.w.} \end{cases}$$

$$\begin{aligned} f_Y(y) &= \int_{-\infty}^{\infty} f(x,y) dx \\ &= \int_0^y 8xy dx && 0 < y < 1 \\ &= 4y^3, && 0 < y < 1 \end{aligned}$$

$$\therefore f_{X|Y}(x|y) = \begin{cases} \frac{8xy}{9y^3} & \text{if } 0 < x < y \\ 0 & \text{otherwise} \end{cases}, \quad 0 < y < 1$$

(b) Find $E(X|Y=0.5)$ and $\text{var}(X|Y=0.5)$.

$$E(X|Y=0.5) = \int_{x \in E_X} x \cdot f_{X|Y}(x|y) dx$$

$$= \int_0^{y=0.5} x \cdot \frac{2x}{y^2} dx$$

$$= \frac{2}{3} \left[\frac{x^3}{y^2} \right]_0^y$$

$$= \frac{2}{3} y = \frac{2}{3} \times 0.5 = \frac{1}{3} \quad (\text{for } y=0.5)$$

$$\text{var}(X|Y=0.5) = E(X^2|Y=0.5) - E(X|Y=0.5)^2$$

$$= \int_{-\infty}^{\infty} x^2 f_{X|Y}(x|y) dx - \frac{1}{9}$$

$$= \int_0^y x^2 \cdot \frac{2x}{y^2} dx - \frac{1}{9}$$

$$= \frac{1}{2y^2} y^4 - \frac{1}{9} = \frac{1}{2} \times \left(\frac{1}{2}\right)^2 - \frac{1}{9} = \frac{1}{72}$$

$$(c) P(0 < Y < \frac{1}{3}) = \int_0^{1/3} f_Y(y) dy$$

$$= \int_0^{1/3} 9y^3 dy$$

$$= \left(\frac{1}{3}\right)^9 = \frac{1}{19683}$$

$$P\left(\frac{1}{3} < Y < \frac{2}{3} \mid X = 0.5\right) = P\left(\frac{1}{2} < Y < \frac{2}{3}\right)$$

$$= \int_{1/2}^{2/3} \frac{2y}{1-x^2} dy$$

$$= \left[\frac{y^2}{1-x^2} \right]_{1/2}^{2/3}$$

$$= \frac{1}{1-x^2} \left[\frac{4}{9} - \frac{1}{4} \right]$$

$$(X = 0.5) = \frac{4}{3} \times \frac{7}{9 \times 4} = \frac{7}{27}$$

$$= \frac{7}{27}$$

for $1 \leq y \leq 2$

$$F_Y(y) = 1 - P(X_1 + X_2 > y)$$

$$= 1 - 2 \cdot \frac{1}{2} (1 - \frac{y}{2})(2-y)$$

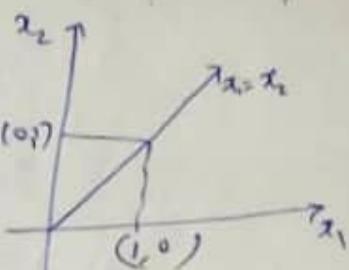
$$= 1 - \frac{1}{2} (2-y)^2$$

For $y > 2$

$$F_Y(y) = 2 \cdot \frac{1}{2} = 1$$

Hence the c.d.f of Y is

$$F_Y(y) = \begin{cases} 0 & \text{if } y < 0 \\ \frac{y^2}{2} & \text{if } 0 \leq y < 1 \\ 1 - \frac{(y-2)^2}{2} & \text{if } 1 \leq y < 2 \\ 1 & \text{if } y \geq 2 \end{cases}$$



So the p.d.f of Y is

$$f_Y(y) = \begin{cases} y & \text{if } 0 \leq y < 1 \\ 2-y & \text{if } 1 \leq y < 2 \\ 0 & \text{o.w.} \end{cases}$$

27) $y_1 = \frac{x_1}{x_2}, \quad y_2 = x_2$

$$y_1 = \frac{x_1}{x_2} + y_2 = x_2 \Rightarrow x_1 = y_1 y_2$$

$$J = \begin{vmatrix} \frac{\partial x_1}{\partial y_1} & \frac{\partial x_1}{\partial y_2} \\ \frac{\partial x_2}{\partial y_1} & \frac{\partial x_2}{\partial y_2} \end{vmatrix} = \begin{vmatrix} y_2 & y_1 \\ 0 & 1 \end{vmatrix} = y_2 \neq 0 \text{ on the range of transformation}$$

Thus the joint p.d.f of y_1, y_2 is

$$f_{Y_1 Y_2}(y_1, y_2) = f_X(y_1 y_2, y_2) |J|$$

$$= \begin{cases} 8y_1 y_2^3 & \text{if } 0 < y_1 < 1, 0 < y_2 < 1 \\ 0 & \text{o.w.} \end{cases}$$

$$\text{As } 0 < x_1 < x_2 < 1$$

$$\Rightarrow 0 < y_1 y_2 < y_2 < 1$$

$$\Rightarrow 0 < y_1 < 1 \text{ and } 0 < y_2 < 1$$

Thus the marginal p.d.f of Y_1 is

$$\begin{aligned}f_{Y_1}(y_1) &= \int_0^{\infty} f_{Y_1,Y_2}(y_1, y_2) dy_2 \\&= \int_0^1 8y_1 y_2^3 dy_2 \quad \text{if } 0 < y_1 < 1 \\&= \begin{cases} 2y_1, & \text{if } 0 < y_1 < 1 \\ 0, & \text{otherwise.} \end{cases}\end{aligned}$$

(28) (a) $\int_0^{\infty} \left(\int_0^{\infty} \frac{e^{-(y+x)}}{y} dx \right) dy$ Put $\frac{x}{y} = t$

$$\begin{aligned}&= \int_0^{\infty} \left(\frac{e^{-y}}{y} \left(-\frac{1}{y} e^{-y/t} \right) \right)_0^{\infty} dy = \int_0^{\infty} \frac{e^{-y}}{y} dy \\&\Rightarrow \int_0^{\infty} \int_0^{\infty} f(x,y) dx dy = 1. \quad = (-e^{-y})_0^{\infty} = 1\end{aligned}$$

(b) $E(X) = \int_{-\infty}^{\infty} \int_0^{\infty} x f(x,y) dx dy$

$$= \int_0^{\infty} e^{-y} \left(\int_0^{\infty} \frac{x}{y} e^{-x/y} dx \right) dy$$

$$= \int_0^{\infty} y e^{-y} dy = 1$$

$E(Y) = \int_{-\infty}^{\infty} \int_0^{\infty} y f(x,y) dx dy$

$$= \int_0^{\infty} e^{-y} \int_0^{\infty} e^{-x/y} dx dy$$

$$= \int_0^{\infty} y e^{-y} dy = 1$$

$E(XY) = \int_0^{\infty} y e^{-y} \int_0^{\infty} \frac{x}{y} e^{-x/y} dx dy$

$$= \int_0^{\infty} y^2 e^{-y} dy = 2$$

Hence $\text{Cov}(X,Y) = E(XY) - E(X)E(Y)$
 $= 2 - 1 = 1$

(29) (a)

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(x, y) dx dy = 1$$

$$\Rightarrow \int_0^1 \left(\int_0^1 (x + cy^2) dy \right) dx = 1$$

$$\Rightarrow \int_0^1 \left(xy + \frac{cy^3}{3} \right)_0^1 dx = 1$$

$$\Rightarrow \int_0^1 \left(x + \frac{c}{3} \right) dx = 1 \Rightarrow \left(\frac{x^2}{2} + \frac{cx}{3} \right)_0^1 = 1$$

$$\Rightarrow \frac{1}{2} + \frac{c}{3} = 1 \Rightarrow \frac{c}{3} = \frac{1}{2} \Rightarrow c = \frac{3}{2}$$

$$F_{(X,Y)}(x, y) = \int_{-\infty}^x \int_{-\infty}^y f(x_1, x_2) dx_1 dx_2$$

$$= \int_{-\infty}^x \int_{-\infty}^y (x_1 + \frac{3}{2}x_2^2) dx_1 dx_2$$

if $x < 0$ or $y < 0$

$$= \begin{cases} 0 & \text{if } x < 0, y < 0 \\ \int_0^x \int_0^y (x_1 + \frac{3}{2}x_2^2) dx_2 dx_1 & \text{if } 0 \leq x, y \leq 1 \\ \int_0^x \int_0^1 (x_1 + \frac{3}{2}x_2^2) dx_2 dx_1 & \text{if } 0 \leq x \leq 1, y \geq 1 \\ \int_0^1 \int_0^y (x_1 + \frac{3}{2}x_2^2) dx_2 dx_1 & \text{if } x \geq 1, 0 \leq y \leq 1 \\ \int_0^1 \int_0^1 (x_1 + \frac{3}{2}x_2^2) dx_2 dx_1 & \text{if } x \geq 1, y \geq 1 \\ & \text{if } x < 0 \text{ or } y \neq 0 \end{cases}$$

$$= \begin{cases} 0 & \text{if } x < 0, y < 0 \\ \frac{1}{2}x^2y + \frac{1}{2}xy^3 & \text{if } 0 \leq x, y \leq 1 \\ \frac{1}{2}x^2 + \frac{1}{2}x & \text{if } 0 \leq x \leq 1, y \geq 1 \\ \frac{1}{2}y^2 + \frac{1}{2}y^3 & \text{if } 0 \leq y \leq 1, x \geq 1 \\ & \text{if } x \geq 1, y \geq 1 \end{cases}$$



$$\begin{aligned}
 \text{c) } P(0 \leq X \leq \frac{1}{2}, 0 \leq Y \leq \frac{1}{2}) &= \int_0^{\frac{1}{2}} \int_0^{\frac{1}{2}} (x + \frac{3}{2}y^2) dx dy \\
 &= \int_0^{\frac{1}{2}} \left(\frac{x^2}{2} + \frac{3}{2}xy^2 \right) \Big|_0^{\frac{1}{2}} dy \\
 &= \int_0^{\frac{1}{2}} \left(\frac{1}{8} + \frac{3}{4}y^2 \right) dy \\
 &= \left(\frac{1}{8}y + \frac{3}{4}\frac{y^3}{3} \right) \Big|_0^{\frac{1}{2}} = \frac{1}{16} + \frac{1}{32} = \frac{3}{32}
 \end{aligned}$$

(d) The marginal p.d.f. of X & Y is $\forall x \in \mathbb{R}$

$$\begin{aligned}
 f_X(x) &= \int_0^{\infty} f(x, y) dy & 0 \leq x \leq 1 \\
 &= \begin{cases} \int_0^{\infty} (x + \frac{3}{2}y^2) dy & 0 \leq x \leq 1 \\ 0 & \text{o.w.} \end{cases} \\
 &= \begin{cases} \left(xy + \frac{3}{2}y^3 \right) \Big|_0^{\infty} & 0 \leq x \leq 1 \\ 0 & \text{o.w.} \end{cases} = \begin{cases} x + \frac{1}{2} & 0 \leq x \leq 1 \\ 0 & \text{o.w.} \end{cases}
 \end{aligned}$$

$$\begin{aligned}
 f_Y(y) &= \int_{-\infty}^{\infty} f(x, y) dx & \forall y \in \mathbb{R} \\
 &= \begin{cases} \int_{-\infty}^{\infty} (x + \frac{3}{2}y^2) dx & 0 \leq y \leq 1 \\ 0 & \text{o.w.} \end{cases} \\
 &= \begin{cases} \left(\frac{x^2}{2} + \frac{3}{2}y^2x \right) \Big|_{-\infty}^{\infty} & 0 \leq y \leq 1 \\ 0 & \text{o.w.} \end{cases} = \begin{cases} \frac{3}{2}y^2 + \frac{1}{2} & 0 \leq y \leq 1 \\ 0 & \text{o.w.} \end{cases}
 \end{aligned}$$

$$(30) \quad y_1 = x_1 + x_2 \quad \text{and} \quad y_2 = \frac{x_2}{x_1 + x_2}$$

$$\Rightarrow x_2 = y_1 y_2 \quad \text{and} \quad x_1 = y_1(1-y_2)$$

$$J = \begin{vmatrix} \frac{\partial x_1}{\partial y_1} & \frac{\partial x_1}{\partial y_2} \\ \frac{\partial x_2}{\partial y_1} & \frac{\partial x_2}{\partial y_2} \end{vmatrix} = \begin{vmatrix} 1-y_2 & -y_1 \\ y_2 & y_1 \end{vmatrix} = y_1 - y_1 y_2 + y_1 y_2 = y_1 \neq 0$$

in the range
of transformation

Hence the joint p.d.f of y_1 and y_2

$$f_{(y_1, y_2)}(y_1, y_2) = f_x(y_1(1-y_2), y_1 y_2) |J|$$

$$= \begin{cases} \frac{1}{2} y_1 e^{-y_1(1-y_2)} & \text{if } 0 < y_1 y_2 < y_1(1-y_2) < \infty \\ \frac{1}{2} y_1 e^{-y_1 y_2} & \text{if } 0 < y_1(1-y_2) < y_1 y_2 < \infty \\ 0 & \text{o.w.} \end{cases}$$

The p.d.f of x_i is

$$f_{x_i}(x_i) = \begin{cases} e^x & \text{if } x_i > 0 \\ 0 & \text{o.w.} \end{cases}$$

The joint p.d.f of x_1, x_2, x_3 is

$$f_x(x_1, x_2, x_3) = f_{x_1}(x_1) f_{x_2}(x_2) f_{x_3}(x_3) \quad \forall (x_1, x_2, x_3) \in \mathbb{R}^3$$

$$= \begin{cases} e^{-(x_1+x_2+x_3)} & \text{if } x_1 > 0, x_2 > 0, x_3 > 0 \\ 0 & \text{o.w.} \end{cases}$$

Let $y_1 = x_1 + x_2 + x_3, y_2 = \frac{x_1 + x_2}{x_1 + x_2 + x_3}, \text{ and } y_3 = \frac{x_1}{x_1 + x_2}$

$$\Rightarrow y_1 y_2 = x_1 + x_2 \quad \text{and} \quad y_3 = \frac{x_1}{x_1 + x_2} \Rightarrow \begin{aligned} x_1 &= y_1 y_2 y_3 \\ x_2 &= y_1 y_2 (1-y_3) \\ x_3 &= y_1 (1-y_2) \end{aligned}$$



$$J = \begin{vmatrix} \frac{\partial x_1}{\partial y_1} & \frac{\partial x_1}{\partial y_2} & \frac{\partial x_1}{\partial y_3} \\ \frac{\partial x_2}{\partial y_1} & \frac{\partial x_2}{\partial y_2} & \frac{\partial x_2}{\partial y_3} \\ \frac{\partial x_3}{\partial y_1} & \frac{\partial x_3}{\partial y_2} & \frac{\partial x_3}{\partial y_3} \end{vmatrix} = \begin{vmatrix} y_2 y_3 & y_1 y_3 & y_1 y_2 \\ y_2(1-y_3) & y_1(1-y_3) & -y_1 y_2 \\ 1-y_2 & -y_1 & 0 \end{vmatrix} = -y_1^2 y_2$$

as $x_i > 0$

Since $y_1 = x_1 + x_2 + x_3, \quad 0 < y_1 < \infty$

$$y_1 y_2 (1-y_3) > 0 \quad \& \quad y_1 (1-y_2) > 0$$

Since $y_1 > 0, \quad 1-y_2 > 0 \Rightarrow y_2 < 1$

Clearly $0 < y_2$. Hence $0 < y_2 < 1$

As $y_1, y_2 > 0 \Rightarrow 1-y_3 > 0 \Rightarrow y_3 < 1$

$$\Rightarrow 0 < y_3 < 1$$

The joint p.d.f of y_1, y_2, y_3 is

$$f(y_1, y_2, y_3) = f_X(y_1, y_2, y_3, y_1 y_2 (1-y_3), y_1 (1-y_2)) | J |$$

$$= \begin{cases} y_1^2 y_2 e^{-y_1} & \text{if } 0 < y_1 < \infty, 0 < y_2, y_3 < 1 \\ 0 & \text{o.w.} \end{cases}$$

The marginal p.d.f of y_1 is

$$f_{Y_1}(y_1) = \int_0^\infty \int_0^\infty f(y_1, y_2, y_3) dy_2 dy_3$$

$$= \begin{cases} \int_0^1 \int_0^1 y_1^2 y_2 e^{-y_1} dy_2 dy_3 & \text{if } 0 < y_1 < \infty \\ 0 & \text{o.w.} \end{cases}$$

$$= \begin{cases} \frac{y_1^2}{2} e^{-y_1} & \text{if } 0 < y_1 < \infty \\ 0 & \text{o.w.} \end{cases}$$

The marginal p.d.f of y_2 is

$$f_{Y_2}(y_2) = \int_0^\infty \int_0^\infty f(y_1, y_2, y_3) dy_1 dy_3$$

$$f_{Y_3}(y_3) = \int_0^\infty \int_0^1 \int_0^1 y_1^2 y_2 e^{-y_1} dy_1 dy_2 dy_3, \quad 0 < y_2 < 1$$

$$= \begin{cases} \int_0^\infty \int_0^1 y_1^2 y_2 e^{-y_1} dy_1 dy_2 & \text{if } 0 < y_2 < 1 \\ 0 & \text{o.w.} \end{cases}$$

$$= \begin{cases} 2y_2 & \text{if } 0 < y_2 < 1 \\ 0 & \text{o.w.} \end{cases}$$

