

## MATH-DATA 363 FINALEXAM

(8) (a)  $H_0: p_{wb} = 0.55, p_o = 0.16, p_{ca} = 0.12, p_m = 0.06, p_l = 0.06$   
 $\& p_a = 0.05$

$H_1$ : At least one of the  $p_i$  values is not equal to the given values, where  $i$  is the species of trees

(b) tree species	WB	O	CA	M	L	A
proport of forest	110	32	24	12	12	10

$$\begin{aligned}
 (c) \chi^2 &= \frac{(127-110)^2}{110} + \frac{(38-32)^2}{32} + \frac{(10-24)^2}{24} + \frac{(9-12)^2}{12} + \frac{(10-12)^2}{12} \\
 &\quad + \frac{(10-6)^2}{10} \\
 &= 14.6023 \quad [DF = 6 - 1 = 5]
 \end{aligned}$$

(d)  $p\text{-value} = \chi^2 \text{cdf}(14.6023, \infty, 5)$  [calculator]  
 $= 0.012204$

Since our  $p$ -value  $< 0.05$ , we can reject the null hypothesis at the 5% significance level.

This means that there were tree species ~~where~~ that marsh tits preferred at a different probability than what was predicted.

$$(7) (a) H_0 : \mu_s \geq 78,000$$

$$H_1 : \mu_s < 78,000$$

Where  $\mu_s$  is the mean salary of the division.

$$(b) \bar{x} = 71500 \quad s = 18000$$

$$t = \frac{\bar{x} - \mu_s}{s / \sqrt{n}} = \frac{71500 - 78000}{18000 / \sqrt{21}}$$

$$= \frac{71500 - 78000}{18000 / \sqrt{21}} = -1.6548$$

$$(c) p = \text{t.cdf}(-\infty, -1.6548, 20)$$

$$= 0.05678 \quad [\text{calculator}]$$

This means that there is a 5.678% probability that this salary is not significantly lower than the mean.

(d) When we change the no. of observations to 31, we get a t score of -2.01 which gives us a p-value of 0.0267. This means we have more evidence to reject the null hypothesis.

This is because by <sup>increasing</sup> ~~decreasing~~ the no. of observations, we are reducing the spread of the data. Due to the ~~larger~~ spread, the critical region increases, meaning it is easier to reject  $H_0$ . <sub>smaller</sub>



Here  $x$  = Rank &  $y$  = MCAT score

① (a) The std of 4.21 is for the mean MCAT Score since its spread is quite small. The scores remain roughly around 511, whereas rank ranges from 0 all the way to 67, both of which are far from the mean.

$$(b) \quad r = \frac{\text{cov}(x, y)}{S_x S_y} = -0.5135$$

This just means that the variables are negatively associated somewhat strongly. ~~We get no~~ ~~This association is negative.~~ It tells us nothing about cause & effect.

$$(c) \quad \hat{\beta} = \frac{\text{cov}(x, y)}{S_x^2} = -0.1034$$

$$\hat{\alpha} = \bar{y} - \hat{\beta} \bar{x} \Rightarrow 511.33 + (0.1034) 25.89 \\ \Rightarrow 514.007$$

$$\text{Eq} \Rightarrow \hat{y} = 514.007 + 0.1034 x_i$$

$$(d) \quad \text{Washington} \Rightarrow \hat{y} = 514.007 + 0.1034 (12) \\ = 515.248$$

$$\text{Residual} = 507 - 515.248 \\ = -8.248$$

This means that the avg score at Washington is -8.248 points less than our prediction.

$$\textcircled{5} \text{ (a) } L(\lambda | t) = (\lambda^2)^8 [t_1 e^{-\lambda t_1} \cdot t_2 e^{-\lambda t_2} \cdots t_8 e^{-\lambda t_8}]$$

$$= (\lambda^{16}) [t_1 t_2 t_3 t_4 t_5 t_6 t_7 t_8 e^{-\lambda(t_1 + t_2 + \cdots + t_8)}]$$

$$\textcircled{5} \text{ (b) } \ln L(\lambda | t) = 16 \ln(\lambda) + \ln(t_1) + \cdots + \ln(t_8) - \lambda(t_1 + t_2 + \cdots + t_8)$$

$$\frac{\partial \ln L}{\partial \lambda} = \frac{16}{\lambda} - (t_1 + t_2 + t_3 + \cdots + t_8)$$

$$\text{But } \frac{\partial \ln L}{\partial \lambda} = 0$$

$$\Rightarrow \frac{16}{\lambda} - (t_1 + t_2 + \cdots + t_8) = 0 \quad \parallel \Rightarrow \hat{\lambda} = \frac{16}{t_1 + t_2 + \cdots + t_8}$$

$$\frac{\partial^2 (\ln L)}{\partial \lambda^2} = -\frac{16}{\lambda^2} \quad \text{which is -ve, so this is a maximum}$$

$$\textcircled{5} \text{ (c) } \hat{\lambda} = \frac{16}{0.68 + 12.00 + \cdots + 1.72} = 0.3468$$

$$\textcircled{5} \text{ (d) } P(\text{Wait} \leq 10) = F_T(10 | \hat{\lambda})$$

$$= 1 - (1 + (0.3468) 10) e^{-0.3468 \cdot 10}$$

$$= \boxed{0.861}$$



$$(a) H_0: p_w - p_m \leq 0$$

$$H_1: p_w - p_m > 0$$

Where  $p_w$  is ~~popula~~ proportion of women  
&  $p_m$  is proportion of men agreeing with the  
statement.

$$(b) \hat{p}_m = 91/914 = 0.09956 \quad \hat{p}_w = 135/921 = 0.14658 \quad \hat{p} = 221/1835 = 0.12316$$

$$z = \frac{0.14658 - 0.09956}{\sqrt{0.12316(1-0.12316)\left(\frac{1}{914} + \frac{1}{921}\right)}} = 3.06459$$

$$(c) p\text{value} = \text{normcdf}(3.06459, \infty, 0, 1) \text{ [calc]} = 0.00109$$

This means that we can expect to see this  
diff in proportion 0.109% of the time ~~which~~ if  
the null hypothesis is true. Therefore, we  
can reject it at the ~~1~~% level.

1  
→ 1.96 from calculator

$$(d) C.I = 0.14658 \pm z_{0.975}^* \sqrt{\frac{0.14658(1-0.14658)}{921}}$$

$$= (0.123737, 0.169423)$$

SAME  
↕

$$\textcircled{3} \text{ (a) } F_x^{-1}(x|\theta) = (x)^{1/\theta} \quad \left| \quad F_x(x|\theta) = 1/2 = x^\theta \right.$$

$$F_x^{-1}(1/2|\theta) = (1/2)^{1/\theta} \quad \left| \quad x = (1/2)^{1/\theta} \right.$$

which is the median

$$\text{(b) } \text{density} = \frac{\partial F_x(x|\theta)}{\partial \theta} \rightarrow \frac{\partial x^\theta}{\partial \theta} = \theta x^{\theta-1}$$

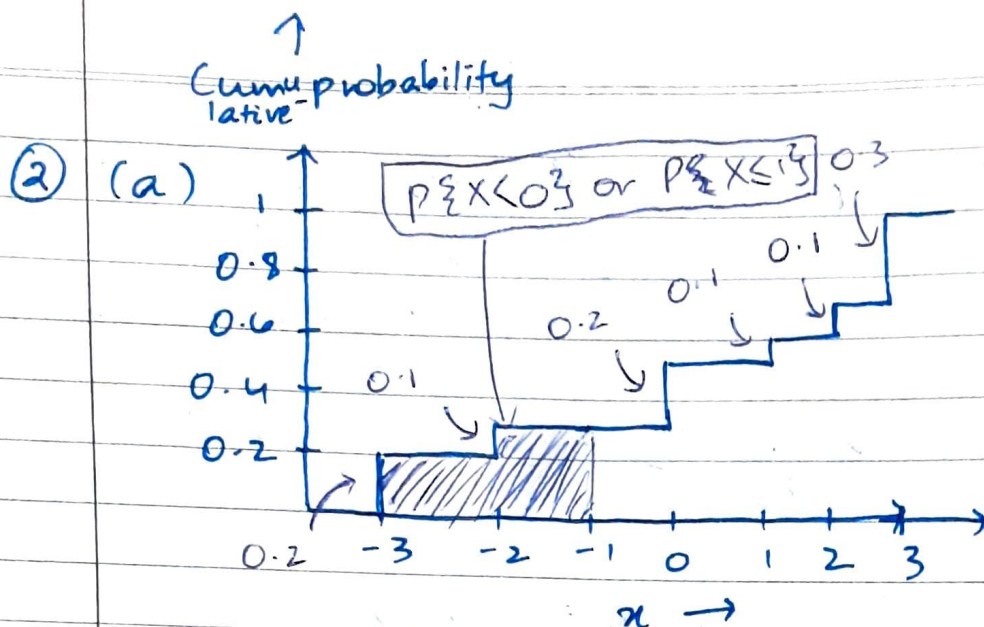
$$\text{(c) } EX = \int_0^1 f_x(x|\theta) \cdot x \, dx = \theta \int_0^1 x^\theta \, dx$$

$$= \theta \left[ \frac{x^{\theta+1}}{\theta+1} \right]_0^1 \Rightarrow \frac{\theta}{\theta+1} (1)^{\theta+1} = \frac{\theta}{\theta+1} = \mu$$

$$\text{(d) } \bar{x} = \frac{\theta}{\theta+1} \quad \Rightarrow \bar{x}\theta + \bar{x} = \theta$$

$$\Rightarrow \bar{x} = \theta - \bar{x}\theta = \theta(1 - \bar{x})$$

$$\hat{\theta} = \frac{\bar{x}}{1 - \bar{x}}$$



(b)  $E[X] = \sum f_x(x) \cdot x$

$$\Rightarrow (-3)(0.2) + (-2)(0.1) + (-1)(0) + 0(0.2) + (1)(0.1) + (2)(0.1) + (3)(0.3)$$

$$= 0.4$$

(c)  $\text{Var}(X) = E[X^2] - (E[X])^2$

$$E[X^2] = (-3)^2(0.2) + (-2)^2(0.1) + 1^2(0.1) + (2)^2(0.1) + (3)^2(0.3)$$

$$= 5.4$$

$$\text{Var}(X) = 5.4 - (0.4)^2 = 5.24$$

(d)  $E[5 - 3X] = E[5] - 3E[X] = 5 - 3(0.4)$

$$= 3.8$$

$$\text{Var}[5 - 3X] = \text{Var}[5] - \text{Var}[3X]$$

$$= 5 - 3^2 \text{Var}[X]$$

$$= 5 - (9)(5.24)$$

$$= -42.16$$



$$P(A|B) = \frac{P(A \cap B)}{P(B)}$$

$$P(C|A) = \frac{P(A|C)P(C)}{P(A|C)P(C) + P(A|C^c)P(C^c)}$$

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Mean with  $\sigma$  known

Mean with  $\sigma$  unknown

1 sample  $\Rightarrow z = \frac{\bar{x} - \mu_0}{\sigma/\sqrt{n}} \parallel \frac{\bar{x} + z^* \sigma}{\sqrt{n}}$

1 sam  $\Rightarrow t = \frac{\bar{x} - \mu_0}{s/\sqrt{n}} \parallel \frac{\bar{x} + t^* s}{\sqrt{n}}$

2 sample  $z = \frac{\bar{x}_1 - \bar{x}_2}{\sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}}$

2 sam  $\Rightarrow t = \frac{\bar{x}_1 - \bar{x}_2}{\sqrt{\frac{s_1^2}{n_1} + \frac{s_2^2}{n_2}}}$

Regression

$\hat{\beta} = \frac{\text{cov}(x,y)}{\text{Var}(x)} = \frac{r s_y}{s_x}$

$\frac{\sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}}{\sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}}$

$\hat{\alpha} = \bar{y} - \hat{\beta} \bar{x}$

$\bar{x}_1 - \bar{x}_2 \pm z^* \sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}$

$\bar{x}_1 - \bar{x}_2 \pm t^* \sqrt{\frac{s_1^2}{n_1} + \frac{s_2^2}{n_2}}$

$r = \frac{\text{cov}(x,y)}{s_x s_y}$

$\chi^2 = \sum \frac{(\theta_{ij} - E_{ij})^2}{E_{ij}} \quad df = (n-1) \times (m-1)$

Matched pairs:

$d = x_1 - x_2 \quad \bar{d} = \bar{x}_1 - \bar{x}_2 \quad sd = sd(x_1 - x_2)$

Proportions (Only z)

1 sample:  $z = \frac{\hat{p} - p_0}{\sqrt{\frac{p_0(1-p_0)}{n}}}$

CI:  $\hat{p} \pm z^* \sqrt{\frac{\hat{p}(1-\hat{p})}{n}}$

2 Sample:  $z = \frac{\hat{p}_1 - \hat{p}_2}{\sqrt{\hat{p}(1-\hat{p})(\frac{1}{n_1} + \frac{1}{n_2})}}$

Here,  $\hat{p} = \frac{x_1 + x_2}{n_1 + n_2}$

ANOVA =  $\bar{y} - \bar{y} \pm t \frac{s_{res}}{\sqrt{1/n + 1/n}}$

Calc commands  $\Rightarrow p \text{ normcdf}, invcdf$

CI:  $\hat{p}_1 - \hat{p}_2 \pm z^* \sqrt{\frac{\hat{p}_1(1-\hat{p}_1)}{n_1} + \frac{\hat{p}_2(1-\hat{p}_2)}{n_2}}$

$\alpha = (1-\alpha)/2$

Conf  $\Rightarrow \bar{y}_j \pm t_{(1-\alpha/2), n-q} \frac{s_{resid}}{\sqrt{n_j}}$

$F = \frac{SS_{\text{between}} / (q-1)}{SS_{\text{resid}} / (n-q)}$   
 $\uparrow$  within

$SS_{\text{resid}} = \sum n_j s_x^2 (n-1) \Rightarrow SSQ$   
 $SS_{\text{betw}} = \sum n(\bar{y} - \bar{\bar{y}})^2 \Rightarrow SSE$