

# Mathematical Expectation & its Properties (1)

Discrete Case

$$* E(x) = \sum_i x_i p_i$$

$$* E(x^r) = \sum_i x_i^r p_i$$

\* "  $r^{\text{th}}$  moment about origin "  $= M_r'$

\* Properties

$$① E(c) = c$$

$$② E(x+y) = E(x) + E(y) \quad (\because \text{Addition Theorem})$$

$$③ E\left(\sum_{i=1}^n x_i\right) = \sum_{i=1}^n E(x_i)$$

④ Multiplication Theorem

\* If  $x$  &  $y$  are independent then  $E(xy) = E(x) \cdot E(y)$

\* In General,

$$E\left(\sum_{i=1}^n x_i\right) = \sum_{i=1}^n E(x_i)$$

⑤ If  $x$  is a random variable & 'a' is a constant then

$$i) E(a \psi(x)) = a E(\psi(x))$$

Continuous Case

$$E(x) = \int_{-\infty}^{\infty} x f(x) dx$$

$$E(x^r) = \int_{-\infty}^{\infty} x^r f(x) dx = M_r'$$

$$ii) E(\psi(x)+a) = E(\psi(x)) + a$$

### Corollary

$$i) \text{ If } \psi(x) = x \text{ then } E(ax) = aE(x)$$

$$\& E(x+a) = E(x) + a$$

$$ii) E(ax+b) = aE(x) + b$$

$$iii) E(g(x)) = g(E(x))$$

### Note

$$i) E(1/x) \neq \frac{1}{E(x)}$$

$$iii) E(\log x) \neq \log(E(x))$$

$$ii) E(x^{1/2}) \neq (E(x))^{1/2}$$

$$iv) E(x^2) \neq (E(x))^2$$

$$v) E\left(\sum_{i=1}^n a_i x_i\right) = \sum_{i=1}^n a_i E(x_i)$$

$$vi) \text{ If } x \geq 0 \text{ then } E(x) \geq 0$$

$$vii) \text{ Let } x \& y \text{ be two Random Variables}$$

$$y \leq x \text{ then } E(y) \leq E(x)$$

$$viii) |E(x)| \leq E(|x|)$$

### Properties

$$(i) \text{ If } x \text{ is a Random Variable then}$$

$$V(ax+b) = a^2 V(x)$$

$$\text{Since } V(b) = 0$$

(2) If  $b=0$  then  $v(ax) = a^2 v(x)$  (2)  
 $\Rightarrow$  Variance is not independent of change of scale.

(3) If  $a=0$  then  $v(b) = 0$   
 $\Rightarrow$  Variance of constant is zero

(4) If  $a=1$ ,  $v(x+b) = v(x)$   
 $\Rightarrow$  Variance is independent of change of origin.

### Covariance

\* If  $X$  &  $Y$  are two Random Variables then covariance between them is defined as

$$\text{Cov}(X, Y) = E(XY) - E(X) \cdot E(Y)$$

### Note

(1) If  $X$  &  $Y$  are Independent then

$$E(XY) = E(X)E(Y)$$

$$\therefore \text{Cov}(X, Y) = 0$$

$$(2) \text{Cov}(ax, by) = ab \text{Cov}(x, y)$$

$$(3) \text{Cov}(x+a, x+b) = \text{Cov}(x, y)$$

$$(4) \text{Cov}\left(\frac{x-\bar{x}}{\sigma_x}, \frac{y-\bar{y}}{\sigma_y}\right) = \frac{1}{\sigma_x \sigma_y} \text{Cov}(x, y)$$

$$(5) \text{Cov}(ax+b, cx+d) = ac \text{Cov}(x, y)$$

$$(6) \text{Cov}(ax+by, cx+dy) = a c \sigma_x^2 + b d \sigma_y^2 + (ad+bc) \text{Cov}(x, y)$$

Correlation Coefficient: (Karl Pearson's)

The correlation coefficient between the variables  $x$  &  $y$  is defined as,

$$P_{xy} = \frac{\text{Cov}(x, y)}{\sigma_x \sigma_y}$$

Note

$$\text{Var}(ax \pm by) = a^2 \text{Var}(x) + b^2 \text{Var}(y) \pm 2ab \text{Cov}(x, y)$$

Problems

(1) Let  $x$  be a random variable with the following probability distribution

|          |               |               |               |
|----------|---------------|---------------|---------------|
| $x$      | -3            | 6             | 9             |
| $P(X=x)$ | $\frac{1}{6}$ | $\frac{1}{2}$ | $\frac{1}{3}$ |



Then find i)  $E(x)$  ii)  $E(x^2)$  iii)  $E((2x+1)^2)$

Sol

$$E(x) = \sum_i x_i p_i = -3 \times \frac{1}{6} + 6 \times \frac{1}{2} + 9 \times \frac{1}{3} \\ = \frac{11}{2}$$

$$E(x^2) = \sum_i x_i^2 p_i = 9 \times \frac{1}{6} + 36 \times \frac{1}{2} + 81 \times \frac{1}{3} \\ = \frac{93}{2}$$

$$E(2x+1)^2 = E(4x^2 + 4x + 1) = 4E(x^2) + 4E(x) + 1 \\ = 4\left(\frac{93}{2}\right) + 4 \times \frac{11}{2} + 1 = 209$$

Note

$$\text{Var}(x) = E(x^2) - (E(x))^2 \\ = \frac{93}{2} - \frac{121}{4}$$

$$= \frac{186 - 121}{4} = \frac{45}{4}$$

$$V(2x \pm 3) = 4V(x) = 4\left(\frac{45}{4}\right) = 45$$

## Moment Generating function (MGF)

The MGF of a random variable  $X$  (about origin) having the probability function  $f(n)$  is given by

$$M_x(t) = E(e^{tx}) = \begin{cases} \sum_n e^{tx} f(n) & \text{(for discrete case)} \\ \int_{-\infty}^{\infty} e^{tx} f(n) dn & \text{(for continuous case)} \end{cases}$$

### Remarks

$$* M_x(t) = E(e^{tx})$$

$$= E\left(1 + tx + \frac{t^2 x^2}{2!} + \frac{t^3 x^3}{3!} + \dots + \frac{t^r x^r}{r!} + \dots\right)$$

$$= 1 + t E(x) + \frac{t^2}{2!} E(x^2) + \frac{t^3}{3!} E(x^3) + \dots + \frac{t^r}{r!} E(x^r) + \dots$$

$\therefore M_r' = r^{\text{th}}$  moment about the origin

= Coefficient of  $\frac{t^r}{r!}$  in  $M_x(t)$

$$M_r' = \left[ \frac{d^r}{dt^r} M_x(t) \right]_{t=0}$$

$M_r' = E[(X-a)^r]$  is the  $r^{\text{th}}$  moment about  $x=a$

### Some limitation of MGF

① A Random Variable  $X$  may have no moments although mgf exists.

② A Random Variable  $X$  can have mgf & some moments, yet the mgf does not generate the moments.

③

### Properties of MGF

①  $M_{cX}(t) = M_X(ct)$  where  $c$  is a Constant.

② If  $X_1, X_2, \dots, X_n$  are independent random variables then the MGF of their sum is equal to product of their respective MGFs.

i.e.,  $M_{X_1+X_2+\dots+X_n}(t) = M_{X_1}(t) M_{X_2}(t) \dots M_{X_n}(t)$

③ Effect of change of origin & scale on MGF.

Sums

① Let  $x$  be the random variable with  $P(x=r) = q^{r-1} \cdot p$ ,  $r=1, 2, \dots$  then find the

i) MGF of  $x$

ii) mean & variance of  $x$ .

Sol

$$M_x(t) = E(e^{tx}) = \sum_{r=1}^{\infty} e^{tr} P(x=r)$$

$$= \sum_{r=1}^{\infty} e^{tr} q^{r-1} \cdot p = p/q \sum_{r=1}^{\infty} (q e^t)^r$$

$$= p/q (q e^t) \sum_{r=1}^{\infty} (q e^t)^{r-1}$$

$$= p e^t [1 + q e^t + (q e^t)^2 + \dots] = p e^t (1 - q e^t)^{-1}$$

$$M_x(t) = \frac{p e^t}{1 - q e^t}$$

$$M_x'(t) = \frac{(1 - q e^t) p e^t - p e^t (-q e^t)}{(1 - q e^t)^2} = \frac{p e^t (1 - q e^t + q e^t)}{(1 - q e^t)^2}$$



(5)

$$M_x(t) = \frac{pe^t}{(1-qe^t)^2}$$

$$M_x''(t) = \frac{(1-qe^t)^2 pe^t - 2pe^t(1-qe^t)(-qe^t)}{(1-qe^t)^4}$$

$$= \frac{pe^t(1-qe^t)(1-qe^t+2qe^t)}{(1-qe^t)^4}$$

$$= \frac{pe^t(1+qe^t)}{(1-qe^t)^3}$$

$$\therefore M_x'(0) = \frac{p}{(1-q)^2} = \frac{p}{p^2} = \frac{1}{p} = \mu_1' = \text{mean}$$

$$M_x''(0) = \frac{p(1+q)}{(1-q)^3} = \frac{p(1+q)}{p^3}$$

$$= \frac{1+q}{p^2} = \mu_2'$$

$$\text{Variance} = \mu_2' - (\mu_1')^2 = \frac{1+q}{p^2} - \frac{1}{p^2}$$

$$= \frac{q}{p^2}$$

② Find the MGF of a R.V.  $x$  whose pdf is given by

$$f(x) = \begin{cases} e^{-x}, & x > 0 \\ 0 & \text{otherwise} \end{cases} \quad \& \text{ also find}$$

its mean & variance

Sol

$$\text{Let } M_x(t) = E(e^{tx}) = \int_0^{\infty} e^{tx} f(x) dx$$

$$M_x(t) = E(e^{tx}) = \int_0^{\infty} e^{tx} e^{-x} dx$$

$$= \int_0^{\infty} e^{-(1-t)x} dx = \left[ \frac{e^{-(1-t)x}}{-(1-t)} \right]_0^{\infty}$$

$$= \frac{-1}{1-t} (e^{-\infty} - e^0)$$

$$= \frac{1}{1-t} \quad \text{if } |t| < 1$$

$$M_x(t) = \frac{1}{1-t} = (1-t)^{-1} = 1 + t + t^2 + t^3 + \dots + t^r + \dots$$

$$= 1 + \frac{t}{1!} + \frac{2!}{2!} t^2 + \frac{3!}{3!} t^3 + \dots + \frac{r!}{r!} t^r + \dots$$

(6)

$$E(x) = M_1' = \text{Coefficient of } t^1/1! = 1 = \text{mean}$$

$$E(x^2) = M_2' = \text{Coefficient of } t^2/2! = 2$$

$$M_2 = \text{Variance} = M_2' - M_1'^2 = 2 - 1 = 1$$

② Find the MGF of a R.V.  $x$  whose pdf is given by

$$f(n) = \frac{1}{2} e^{-|n|}, \quad -\infty < n < \infty \text{ \& find its mean \& variance. (If it possible)}$$

Sol

$$M_x(t) = \int_{-\infty}^{\infty} e^{tn} f(n) dn$$

$$\text{Since } |n| = \begin{cases} n & \text{if } n > 0 \\ -n & \text{if } n < 0 \end{cases}$$

$$= \frac{1}{2} \int_{-\infty}^{\infty} e^{tn} e^{-|n|} dn$$

$$= \frac{1}{2} \int_{-\infty}^{\infty} e^{tn} \cdot e^{-n} dn + \frac{1}{2} \int_{-\infty}^{\infty} e^{tn} e^{-n} dn$$

$$= \frac{1}{2} \int_{-\infty}^0 e^{(1+t)n} dn + \frac{1}{2} \int_0^{\infty} e^{-(1+t)n} dn$$

$$= \frac{1}{2} \left\{ \left( \frac{e^{(1+t)x}}{(1+t)} \right)_{-\infty}^0 + \left( \frac{e^{-(1-t)x}}{-(1-t)} \right)_0^{\infty} \right\}$$

$$= \frac{1}{2} \left[ \frac{1}{1+t} + \frac{1}{1-t} \right]$$

$$= \frac{1}{2} \left( \frac{2}{1-t^2} \right) = \frac{1}{1-t^2}$$

\* Mean & Variance - H.W

Characteristic function

$$\phi_X(t) = E(e^{itx}) = \begin{cases} \int e^{itx} f(x) dx & \text{for Continuous} \\ \sum_n e^{itx} p(n) & \text{for discrete.} \end{cases}$$

It is known as the characteristic function of the random variable  $X$  & where  $\phi_X(t)$  is a complex valued function of real variable 't'.

Remarks

i)  $\phi_X(t)$  always exists only when  $|\phi(t)| \leq 1$

ii)  $\phi_X(0) = 1$