

## Problem 7

- a) To show that the Poisson distribution belongs to the single parameter exponential family, we need to express it's probability mass function (PMF) in the following form:

$$p(y; \eta) = b(y) \exp(\eta^T T(y) - a(\eta))$$

where:

- $\eta$  is the natural parameter.
- $T(y)$  is the sufficient statistic.
- $a(\eta)$  is the log-partition function.
- $b(y)$  is the base measure.

PMF of the Poisson distribution:

$$P(Y = y) = \frac{e^{-\lambda} \lambda^y}{y!}, \quad y = 0, 1, 2, \dots$$

**Rewriting the PMF:**

$$\lambda^y = e^n$$

where ,

$$n = \log(\lambda^y) \implies n = y \log(\lambda)$$

We can rewrite the PMF as:

$$P(Y = y) = \frac{e^{-\lambda} e^{y \log(\lambda)}}{y!}, \quad y = 0, 1, 2, \dots$$

$$\implies P(Y = y) = \frac{1}{y!} \exp(-\lambda + y \log(\lambda))$$

**Comparing with exponential family form**

- $\eta = \log(\lambda)$  (natural parameter)
- $T(y) = y$  (sufficient statistic)
- $a(\eta) = \lambda = \exp(\eta)$  (log-partition function)
- $b(y) = \frac{1}{y!}$  (base measure)

**Expressing the PMF in exponential family form:**

Substituting these components into the general form, we get:

$$\begin{aligned} P(Y = y) &= \frac{1}{y!} \exp(\log(\lambda) \cdot y - \exp(\log(\lambda))) \\ &= b(y) \exp(\eta^T T(y) - a(\eta)) \end{aligned}$$

b) (a) **The log-likelihood function of  $\theta$ , given  $\mathbf{S}$ .**

Given the dataset  $S = \{(x^{(i)}, y^{(i)}), x^{(i)} \in R^k, y^{(i)} \in Z^+, i = 1, 2, \dots, n\}$ , and the assumption that given  $x$ ,  $y$  follows a Poisson distribution with rate  $e^{\theta^T x}$ , the likelihood function is:

$$\begin{aligned} L(\theta) &= \prod_{i=1}^n P(y^{(i)} | x^{(i)}; \theta) \\ &= \prod_{i=1}^n \frac{e^{-e^{\theta^T x^{(i)}}} (e^{\theta^T x^{(i)}})^{y^{(i)}}}{y^{(i)}!} \end{aligned}$$

The log-likelihood function is then:

$$\begin{aligned} l(\theta) &= \log L(\theta) \\ &= \sum_{i=1}^n \log \left( \frac{e^{-e^{\theta^T x^{(i)}}} (e^{\theta^T x^{(i)}})^{y^{(i)}}}{y^{(i)}!} \right) \\ &= \sum_{i=1}^n \left( -e^{\theta^T x^{(i)}} + y^{(i)} \theta^T x^{(i)} - \log(y^{(i)}!) \right) \end{aligned}$$

(b) The gradient of the log-likelihood with respect to  $\theta$  is:

$$\begin{aligned} \nabla l(\theta) &= \frac{\partial l(\theta)}{\partial \theta} \\ &= \sum_{i=1}^n \left( -e^{\theta^T x^{(i)}} x^{(i)} + y^{(i)} x^{(i)} \right) \\ &= \sum_{i=1}^n \left( y^{(i)} - e^{\theta^T x^{(i)}} \right) x^{(i)} \end{aligned}$$

(c) Since we want to maximize the log-likelihood, we will use gradient ascent, which is the same algorithm but with a positive step size.

**1.Initialization:** Choose an initial value for  $\theta$ , say  $\theta^{(0)}$ .

**2.Iteration:** Update  $\theta$  using the following rule:

$$\theta^{(t+1)} = \theta^{(t)} + \alpha \nabla l(\theta^{(t)})$$

where  $\alpha > 0$  is the learning rate (step size) and  $\nabla l(\theta^{(t)})$  is the gradient of the log-likelihood evaluated at  $\theta^{(t)}$ .

**3.Convergence:** Repeat step 2 until convergence, which can be determined by monitoring the change in  $l(\theta)$  or in  $\theta$  itself. For example, stop if  $\|\theta^{(t+1)} - \theta^{(t)}\| < \epsilon$  for some small tolerance  $\epsilon$ .

Substituting the gradient we derived earlier, the update rule for gradient ascent becomes:

$$\theta^{(t+1)} = \theta^{(t)} + \alpha \sum_{i=1}^n \left( y^{(i)} - e^{\theta^T x^{(i)}} \right) x^{(i)}$$