

Stochastic mirror descent for nonparametric adaptive importance sampling

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Objective

Target: A probability measure admitting a density $f : \mathbb{R}^d \rightarrow \mathbb{R}$.

- Unnormalized density: f_u such that $f_u(x) / \int f_u = f(x)$.
- $f_u(x)$ is known for $x \in \mathbb{R}^d$.
- $\int f_u$ is unknown and challenging to compute.
- No access to the gradient ∇f_u .

Question

How can we evaluate the quantity

$$\mathbb{E}_{X \sim f}(\phi(X)) := \int \phi(x)f(x)dx,$$

for some function ϕ ?

In the sequel: $f = f_u$.

Applications: Variational Inference, Bayesian Inference, Reinforcement Learning, Stochastic Optimization, ...

Application: Bayesian Inference

Model

- Parameter: $\theta \sim \underbrace{p(\cdot)}_{\text{prior}}$
- Data: $x|\theta \sim \underbrace{p(\cdot|\theta)}_{\text{likelihood}}$

Objective: Find the **posterior distribution** $p(\cdot|x)$ of θ , given the data x .

$$p(\cdot|x) := \frac{p(x|\cdot)p(\cdot)}{\int p(x|\theta)p(\theta)d\theta}$$

Target: $f_u(\cdot) = p(x|\cdot)p(\cdot)$.

Context

Several types of algorithms:

- **Markov Chain Monte Carlo (MCMC)**: Construction of a Markov chain with an invariant distribution equal to the target distribution.
 - Metropolis-Hastings [Metropolis et al. 1953](#)
 - Hamiltonian Monte Carlo (HMC) [Duane et al. 1987](#)
 - No-U-Turn Sampler (NUTS) [Hoffman and Gelman 2011](#)
 - Langevin algorithm [Meyn and Tweedie 1993](#)
- **Importance Sampling Methods**: Repeated random sampling with weights assigned to each sample to approximate the target distribution.
 - Importance Sampling (IS) [Kloek and Van Dijk 1978](#)
 - Adaptive Importance Sampling [Oh and Berger 1992](#)
 - Annealed Importance Sampling [Neal 2001](#)
- **Variational Inference**: Optimization of a functional defined in the measure space (e.g., Kullback-Leibler divergence), where the minimizer corresponds to the target distribution.
 - Stein Variational Gradient Descent (SVGD) [Liu and Wang 2016](#)
 - Mollified Interaction Energy Descent (MIED) [Li et al. 2023](#)
 - Normalizing Flows [Tabak and Vanden-Eijnden 2010](#)

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Important Sampling Methods

- Generate n particles (X_1, \dots, X_n) .
- Assign a weight w_i to each particle X_i .

Importance Sampling Estimator

$$\hat{\phi}_n := \frac{1}{n} \sum_{i=1}^n w_i \phi(X_i)$$

Good estimator if:

- $\mathbb{E}(\hat{\phi}_n) = \int \phi(x) f(x) dx.$
- $\mathbb{V}(\hat{\phi}_n) \leq \frac{C}{n}.$

Vanilla Importance Sampling algorithm

q_0 : known density

Vanilla Importance Sampling

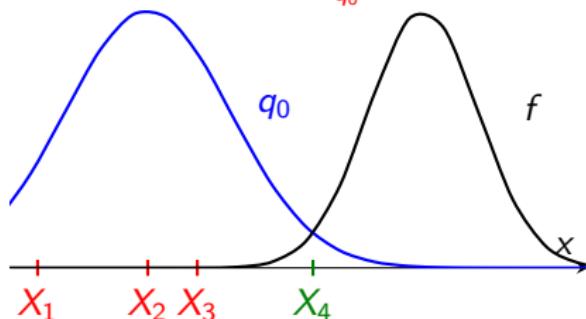
- $(X_1, \dots, X_n) \sim_{i.i.d.} q_0$
- $w_i := \frac{f}{q_0}(X_i)$
- $\mathbb{E}(\hat{\phi}_n) = \int \frac{f(x)}{q_0(x)} q_0(x) \phi(x) dx = \int f(x) \phi(x) dx.$
- $\mathbb{V}(\hat{\phi}_n) = \underbrace{\frac{1}{n} \left(\int \frac{f(x)^2}{q_0(x)} \phi(x)^2 dx - \left(\int f(x) \phi(x) dx \right)^2 \right)}_{\leq C}$

Assumption: $q_0 \geq cf$

Issue with Importance Sampling

$$\mathbb{V}(\hat{\phi}_n) = \frac{1}{n} \left(\int \frac{f(x)^2}{q_0(x)} \phi(x)^2 dx - \left(\int f(x) \phi(x) dx \right)^2 \right)$$

behaves badly when $\frac{f}{q_0}(x) \gg 1$ for some x .



$$\sum_{i=1}^3 \frac{f(X_i)}{q_0(X_i)} \ll \frac{f(X_4)}{q_0(X_4)}$$

$$\hat{\phi}_4 \simeq \frac{f(X_4)}{q_0(X_4)} \phi(X_4)$$

Optimal Importance Sampling

$$q_0 = \arg \min_{q: \text{density}} \mathbb{V} \left(\frac{1}{n} \sum_{i=1}^n \frac{f(X_i)}{q(X_i)} \phi(X_i) \right) = f$$

Adaptive strategy

Important Sampling (IS) algorithm

For $k = 0, \dots, n$,

- ① $X_{k+1} \sim q_0$
- ② $w_{k+1} = \frac{f(X_{k+1})}{q_0(X_{k+1})}$

- To minimize the variance, we want: $q_0 = f$, but f is unknown.
- $\tilde{q}_k := \frac{\sum_{i=1}^k w_i \delta_{X_i}}{\sum_{i=1}^k w_i}$ converges (in the weak-* topology) to f .
- $\tilde{q}_k = q_0$ should perform better but \tilde{q}_k has no density.

Conclusion

At step k , we replace q_0 with a probability measure q_k in the IS algorithm. Where q_k :

- has a density.
- $q_k \rightarrow f$.

Kernel Density Estimation

K : a kernel and $K_h := \frac{1}{h^d} K(\frac{x}{h})$; $(X_1, \dots, X_n) \sim_{i.i.d.} q_0$.

Kernel density estimator Parzen 1962

$$g_n := \frac{1}{n} \sum_{i=1}^n \left(\frac{f}{q_0}(X_i) \right) K_h(\cdot - X_i)$$

- $\mathbb{E}(g_n(x)) = f * K_h(x) \rightarrow_{h \rightarrow 0} f(x)$
- $\mathbb{V}(g_n(x)) = \frac{1}{n} \left(\underbrace{\int \frac{f(y)^2}{q_0(y)} K_h(x-y)^2 dy}_{\leq Ch^{-d}} - (f * K_h(x))^2 \right)$

Conditions

$$h \rightarrow 0 \text{ and } \frac{1}{nh^d} \rightarrow 0$$

Adaptive Importance sampling

$$g_0 = 0$$

Adaptive Importance Sampling algorithm

For $k = 0, \dots, n$:

① $X_{k+1} \sim q_k$

② $w_{k+1} = \frac{f(X_{k+1})}{q_k(X_{k+1})}$

③ $g_{k+1} = (1 - \frac{1}{k+1})q_k + \frac{1}{k+1} w_{k+1} K_{h_k}(\cdot - X_{k+1})$

④ $q_{k+1} = \frac{g_{k+1}}{\int g_{k+1}}$

g_k is the Kernel density estimator of f :

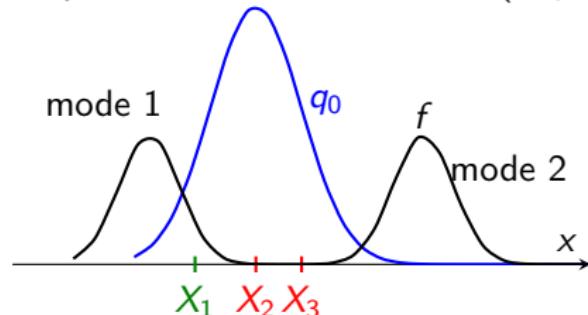
$$g_k(x) = \frac{1}{k} \sum_{i=1}^k w_i K_{h_i}(\cdot - X_i).$$

Mini-Batch Variant

At each iteration, instead of simulating a single particle, we simulate m particles according to the distribution q_k .

Challenges in Adaptive Importance Sampling

Step $k = 0$; batch-size $m = 3$, $(X_1, \dots, X_3) \sim_{i.i.d.} q_0$.



$$g_1 \simeq \frac{f}{q_0}(X_1) K_{h_0}(\cdot - X_1) \implies q_1 \propto K_{h_0}(\cdot - X_1)$$

New particles are sampled around $X_1 \rightarrow$ mode 2 cannot be recovered.

Solution

$$q_k = (1 - \lambda_k) \frac{g_k}{\int g_k} + \lambda_k q_0$$

λ_k adds a safe component to q_k .

Safe Adaptative Importance Sampling (SAIS)

(λ_k) : mixture weights satisfying $\lambda_k \rightarrow 0$.

Safe Adaptative Importance Sampling

For $k = 0, \dots, n$:

① $X_{k+1} \sim q_k$

② $w_{k+1} = \frac{f(X_{k+1})}{q_k(X_{k+1})}$

③ $g_{k+1} = q_k(1 - \frac{1}{k+1}) + \frac{1}{k+1} w_{k+1} K_{h_k}(x - X_{k+1})$

④ $q_{k+1} = (1 - \lambda_{k+1}) \frac{g_{k+1}}{\int g_{k+1}} + \lambda_{k+1} q_0$

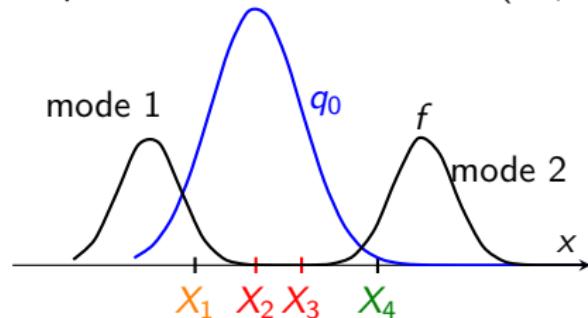
Delyon and Portier 2021

$$\lim_{n \rightarrow \infty} \sup_{x \in \mathbb{R}^d} |q_n(x) - f(x)| = 0 \quad a.s.$$

An important assumption is $\sup_{x \in \mathbb{R}^d} \frac{f}{q_0}(x) < \infty$.

Problem with SAIS

Step $k = 0$; batch-size $m = 4$, $(X_1, \dots, X_4) \sim q_0$.



$$\sum_{i=2}^3 \frac{f(X_i)}{q_0} \ll \frac{f(X_1)}{q_0} \ll \frac{f(X_4)}{q_0}$$

$$g_1 \simeq w_1 K_{h_0}(x - X_1) + w_4 K_{h_0}(x - X_4)$$

- We sample around mode 1 with probability $\frac{w_1}{w_1+w_4} \ll 1$.
- We sample around mode 2 with probability $\frac{w_4}{w_1+w_4} \simeq 1$.

Conclusion

Very low probability of sampling around mode 1, which can be forgotten by the algorithm.

Variance Problem

Reformulation of the problem

The weights (w_1, \dots, w_n) have high variance.

The probability to sample around a particle i :

$$\frac{w_i}{\sum_{k=1}^n w_{k+1}},$$

is small for most of the particles.

Idea from Korba and Portier 2022

Let $\eta \in (0, 1)$:

$$\mathbb{V}_{X \sim q} \left(\left(\frac{f}{q}(X) \right)^\eta \right) \leq \mathbb{V}_{X \sim q} \left(\frac{f}{q}(X) \right)$$

At step k ,

$$g_{k+1} = (1 - \frac{1}{k+1})g_k + \frac{1}{k+1} w_{k+1}^\eta K_{h_k}(x - X_{k+1}).$$

The Mirror Descent algorithm in \mathbb{R}^d

$$\min_{x \in \mathbb{R}^d} h(x)$$

Gradient Descent with learning rate $\eta \in (0, 1]$.

$$x_{k+1} = x_k - \eta \nabla h(x_k) = \arg \min \{x \mapsto \langle \nabla h(x_k), x - x_k \rangle + \frac{1}{\eta} \frac{\|x - x_k\|^2}{2}\}$$

Bregman divergence

Let ψ be a convex function. For $x, y \in \mathbb{R}^d$:

$$D_\psi(y|x) := \psi(y) - \psi(x) - \langle \nabla \psi(x), y - x \rangle.$$

Mirror Descent for some function ψ :

$$x_{k+1} = \arg \min \{x \mapsto \langle \nabla h(x_k), x - x_k \rangle + \frac{1}{\eta} D_\psi(x|x_k)\}$$

Depending of ψ , there are many interesting dynamics.

Mirror Descent in $\mathcal{P}_2(\mathbb{R}^d)$

$$\min_{q \in \mathcal{P}_2(\mathbb{R}^d)} \mathcal{F}(q) = KL(q|f) := \int \log\left(\frac{q}{f}(x)\right) q(x) dx$$

For $\psi = \mathcal{F}$, the Mirror Descent algorithm on \mathcal{F} with learning rate η :

$$q_{k+1} \propto q_k^{1-\eta} f^\eta$$

Chopin, Crucinio, and Korba 2024

$$KL(q_n|f) \leq \frac{(1-\eta)^n}{\eta} KL(q_0|f)$$

Problem:

$\int q^{1-\eta} f^\eta$ is unknown. The algorithm cannot be implemented.

Link between the variance reduction and Mirror Descent

Variance reduction hinted:

$$g_{k+1} = \left(1 - \frac{1}{k+1}\right)g_k + \frac{1}{k+1}w_{k+1}^\eta K_{h_k}(x - X_{k+1}).$$

Let $\mathcal{F}_k := \sigma(X_i : i \leq k)$.

$$\mathbb{E}[g_{k+1} | \mathcal{F}_k] = \left(1 - \frac{1}{k+1}\right)g_k + \frac{1}{k+1} \underbrace{f^\eta q_k^{1-\eta}}_{\text{Mirror Descent}} * K_{h_k}.$$

Conclusion

The variance reduction approach yields the same dynamics as Mirror Descent.

Mirror Descent for Adaptive Sampling (MIDAS)

$\eta \in (0, 1]$: learning rate. Let $g_0 = 0$.

MIDAS

For $k = 0, \dots, n$:

- ① $X_{k+1} \sim q_k$
- ② $w_{k+1} = \frac{f(X_{k+1})}{q_k(X_{k+1})}$
- ③ $g_{k+1} = (1 - \frac{1}{k+1})g_k + \frac{1}{k+1} w_{k+1}^{\eta} K_{h_k}(x - X_{k+1})$
- ④ $q_{k+1} = \frac{g_{k+1}}{\int g_{k+1}}(1 - \lambda_{k+1}) + \lambda_{k+1}q_0$

Bianchi et al. 2024

For every compact sets $A \subset \mathbb{R}^d$:

$$\lim_{n \rightarrow \infty} \sup_{x \in A} |q_n(x) - f(x)| = 0 \quad a.s.$$

Corrolary: $\sqrt{n}(\hat{\phi}_n - \mathbb{E}_{X \sim f}(\phi(X))) \rightarrow \mathcal{N}(0, \underbrace{\mathbb{V}_{X \sim f}(\phi(X))}_{\text{optimal variance}}).$

A spurious stationnary distribution

We recall

$$\mathbb{E}[g_{k+1} | \mathcal{F}_k] = (1 - \frac{1}{k+1})g_k + \frac{1}{k+1} f^\eta q_k^{1-\eta} * K_{h_k}.$$

Stationnary Distributions

Stationnary distributions (q_*, g_*) solve the equation (since $K_{h_k} \rightarrow \delta_0$):

$$g_*(x) = f^\eta q_*^{1-\eta}(x)$$

for every x such that $g_*(x), q_*(x) < \infty$.

When $\eta < 1$, we have two stationnary distributions:

- $q_*(x) = g_*(x) = f$, for every x .
- $g_*(x) = q_*(x) = 0$ almost everywhere $\implies q_*$ is degenerated.

Avoidance of the trap $q_* = g_* = 0$

Define (for some $v > 0$):

$$u_{k+1} = \left(1 - \frac{1}{k+1}\right)u_k + \frac{1}{k+1}u_k^{1-\eta}v^\eta.$$

When $u_k \leq v$, $u_{k+1} \geq u_k$. Therefore, when $u_0 > 0$, we can't have $u_k \rightarrow 0$. Then, by classical arguments, $u_k \rightarrow v$.

$$g_{k+1}(x) = \left(1 - \frac{1}{k+1}\right)g_k(x) + \frac{1}{k+1}f^\eta q_k^{1-\eta} * K_{h_k}(x) + \underbrace{\frac{1}{k+1}\xi_{k+1}(x)}_{\mathbb{E}[\xi_{k+1}(x)|\mathcal{F}_k]=0}.$$

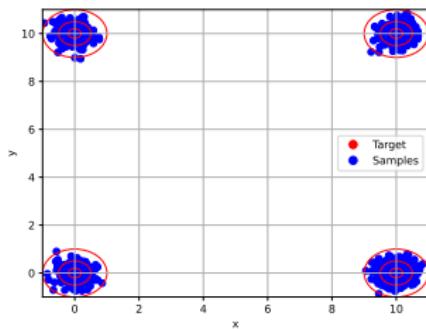
Iterating:

$$g_n(x) = \underbrace{A_n(x)}_{\text{deterministic algorithm}} + \underbrace{B_n q_0(x)}_{\text{safe component}} + \underbrace{M_n(x)}_{\text{martingale increment}}.$$

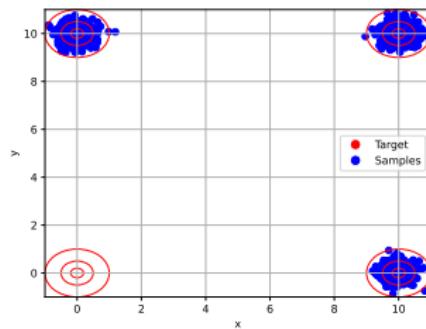
By a Freedman-type inequality, we show that for every x in a compact space $A \subset \mathbb{R}^d$ and for $n \geq n_0$, $B_n q_0(x) + M_n(x) \geq 0$. The noise cannot lead the algorithm to the unstable equilibrium.

Simulation

Target f : mixture of four gaussians centered in $(10i, 10j)$ for $i, j = 0, 1$.



(a) All the modes are found



(b) One mode is missing

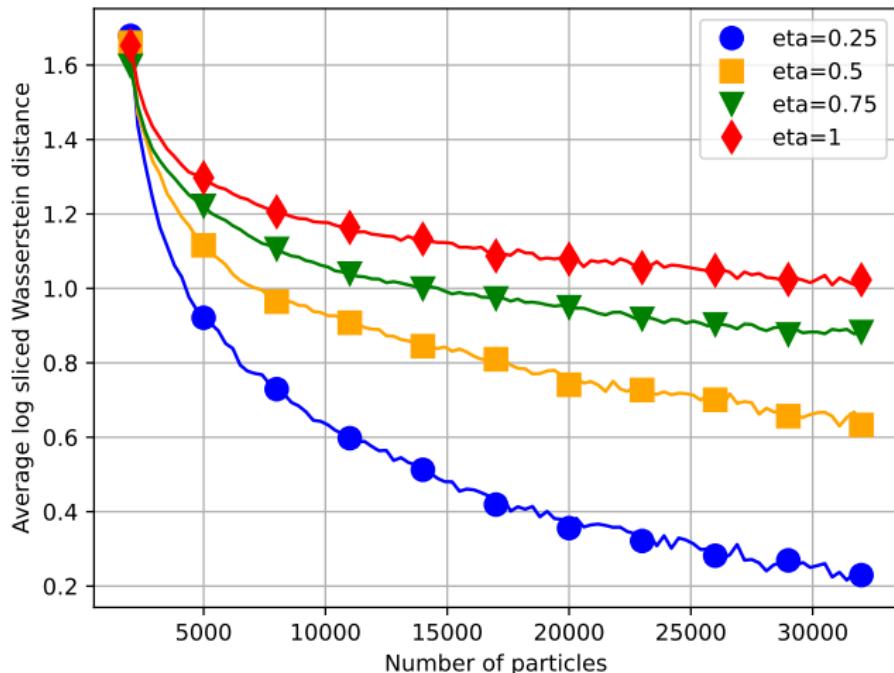


Figure: Average over 50 independent runs of the logarithm of sliced Wasserstein distance for different values of η

Conclusion

Contribution:

- Demonstrated the convergence of a novel algorithm that introduces new dynamics at the intersection of Importance Sampling methods and Variational Inference.
- Achieved improved performance compared to the Safe Adaptive Importance Sampling algorithm, particularly in identifying modes of the target distribution.

Extension/Application of this result:

- Analysis of convergence speed.
- Guidelines for selecting the parameter η .
- Established connections with the annealed importance sampling algorithm.