

# SOLUTIONS FOR PATTERN CLASSIFICATION CH.3

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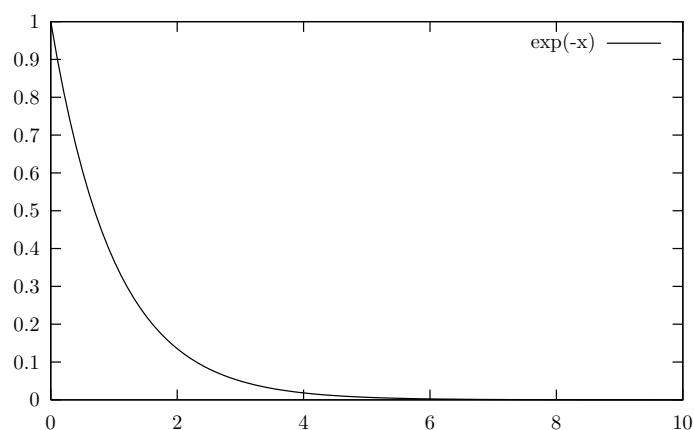
## Problem 1

- a

When  $\theta = 1$

$$p(x|\theta) = \begin{cases} e^{-x}, & x \geq 0 \\ 0 & \text{others} \end{cases}$$

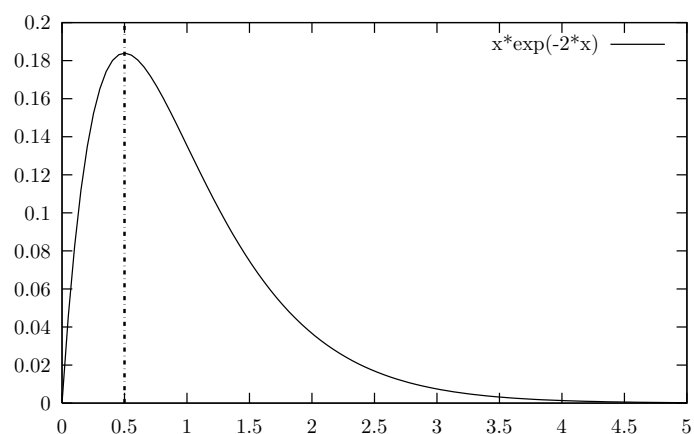
Plot as follows:



When  $x = 2$  we get:

$$p(x|\theta) = \begin{cases} \theta e^{-2\theta}, & x \geq 0 \\ 0 & \text{others} \end{cases}$$

Plot as follows:



- b

$x_1, \dots, x_n$   $p(x|\theta)$ , and  $x_i$  are independently. Then we define the log-likelihood as follows:

$$l(\theta) = \sum_{i=1}^n \ln p(x_i|\theta) = \sum_{i=1}^n (\ln \theta - \theta x_i)$$

Then we get:

$$\frac{dl(\theta)}{d\theta} = \sum_{i=1}^n \left(\frac{1}{\theta} - x_i\right) = \frac{n}{\theta} - \sum_{i=1}^n x_i = 0$$

Thus:

$$\hat{\theta} = \frac{1}{n \sum_{i=1}^n x_i}$$

- c

$\theta = 1$ , when  $n$  is large enough we get:

$$\begin{aligned} \lim_{n \rightarrow \infty} \left(\frac{1}{n} \sum_{i=1}^n x_i\right) &= \int_{-\infty}^{\infty} x e^{-x} dx \\ &= -(x+1)e^{-x} \Big|_0^{\infty} = 1 \end{aligned}$$

## Problem 2

- a  $x$  have a uniform density, when  $\theta \geq \max(\mathcal{D})$  define the likelihood as:

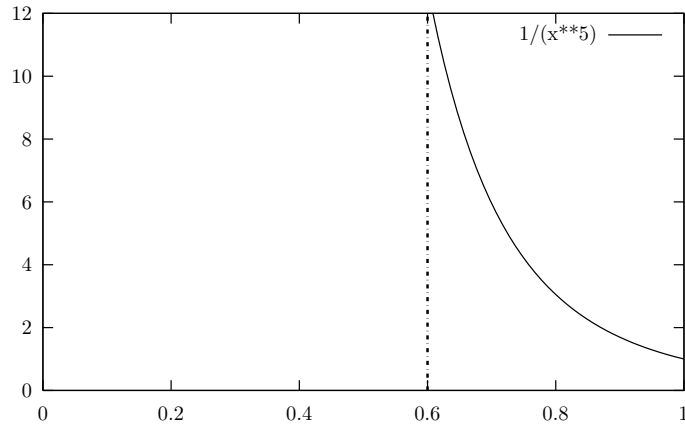
$$\begin{aligned} l(\mathcal{D}|\theta) &= \prod_{i=1}^n p(x_i|\theta) \\ &= \prod_{i=1}^n \frac{1}{\theta} = \frac{1}{\theta^n} \end{aligned}$$

$l(\theta)$  is monotonically decreasing, the  $\hat{\theta}$  that maximizes the  $l(\theta)$  minimize the  $\theta$ . As  $\theta \geq \max(\mathcal{D})$ , the  $l(\theta)$  is maximized in  $\hat{\theta} = \max(\mathcal{D})$

- b

When  $n = 5$ , From a we know that the likelihood of  $p(\mathcal{D}|\theta)$  is:  $\frac{1}{\theta^5}$ , the  $\hat{\theta} = \max(\mathcal{D}) = \max(x_i)$ . Thus we can get  $\hat{\theta}$ , and need not to know the values of the other four points as soon as we know the  $\max x_i = 0.6$ . When  $\theta < 0.6$ ,  $\max(x_i) > \theta$ , so  $p(\mathcal{D}|\theta) = 0$ .

Plot:



## Problem 4

- a

Because  $X$  is a  $d$ -dimensional binary vector with a multivariate Bernoulli distribution:

$$P(X|\theta) = \prod_{i=1}^d \theta_i^{x_i} (1 - \theta_i)^{(1 - x_i)}$$

hence,

$$P(X^1, \dots, X^n|\theta) = \prod_{k=1}^n \prod_{i=1}^d \theta_i^{x_i^k} (1 - \theta_i)^{(1 - x_i^k)}$$

And the likelihood function:

$$l(\theta) = \sum_{k=1}^n \sum_{i=1}^d [x_i^k \ln \theta_i + (1 - x_i^k) \ln(1 - \theta_i)]$$

$$\frac{\partial l(\theta)}{\partial \theta_i} = \sum_{k=1}^n \left[ \frac{x_i^k}{\theta_i} - \frac{1 - x_i^k}{1 - \theta_i} \right] = \sum_{k=1}^n \frac{x_i^k - \theta_i}{\theta_i(1 - \theta_i)}$$

thus, we get

$$\sum_{k=1}^n (x_i^k - \theta_i) = 0$$

$$\theta_i = \frac{1}{n} \sum_{k=1}^n x_i^k$$

Then we can get its vector form as follows:

$$\theta_i = \frac{1}{n} \sum_{k=1}^n X_k$$

## Problem 16

- a

Since A and B are nonsingular matrices of the same order,

$$\begin{aligned}
 (A^{-1} + B^{-1})[A(A + B)^{-1}B] &= (1 + B^{-1}A)(A + B)^{-1}B \\
 &= (A + B)^{-1}B + B^{-1}A(A + B)^{-1}B \\
 &= B^{-1}B(A + B)^{-1}B + B^{-1}A(A + B)^{-1}B \\
 &= B^{-1}(B + A)(A + B)^{-1}B \\
 &= B^{-1}B \\
 &= I
 \end{aligned}$$

Left multiply  $(A^{-1} + B^{-1})^{-1}$  on both sides of the above equation, we get:

$$(A^{-1} + B^{-1})^{-1} = A(A + B)^{-1}B$$

Analogously,

$$\begin{aligned}
 B(A + B)^{-1}A(A^{-1} + B^{-1}) &= B(A + B)^{-1}(I + AB^{-1}) \\
 &= B(A + B)^{-1} + B(A + B)^{-1}AB^{-1} \\
 &= B(A + B)^{-1}BB^{-1} + B(A + B)^{-1}AB^{-1} \\
 &= B(A + B)^{-1}(B + A)B^{-1} \\
 &= BB^{-1} \\
 &= I
 \end{aligned}$$

Right multiply  $(A^{-1} + B^{-1})^{-1}$  on both sides of the above equation, we get:

$$(A^{-1} + B^{-1})^{-1} = B(A + B)^{-1}A$$

- b

If these matrices are not square, then they are both nonsingular,  $A^{-1}$  and  $B^{-1}$  won't exist, which means the above equation won't hold, so, A and B must be square.

- c

From (41),  $\Sigma_n^{-1} = n\Sigma^{-1} + \Sigma_0^{-1}$ , we can get

$$\Sigma_n = (n\Sigma^{-1} + \Sigma_0^{-1})^{-1} = (\Sigma_0^{-1} + n\Sigma^{-1})^{-1}$$

Making use of the matrix identity in (a), we have

$$\Sigma_n = \Sigma_0(\Sigma_0 + \frac{1}{n}\Sigma)^{-1}(\frac{1}{n}\Sigma)$$

Thus, we get Eqs.(45) similarly, we have

$$\Sigma_n = (n\Sigma^{-1} + \Sigma_0^{-1})^{-1} = (\frac{1}{n}\Sigma)(\Sigma_0 + \frac{1}{n}\Sigma)^{-1}\Sigma_0$$

From (42),  $\Sigma_n^{-1}\mu_n = n\Sigma^{-1}\hat{\mu}_n + \Sigma_0^{-1}\mu_0$ , Left multiply the above equation by  $\Sigma_n$ :

$$\begin{aligned}
\mu_n &= \Sigma_n(n\Sigma^{-1}\hat{\mu}_n + \Sigma_0^{-1}\mu_0) \\
&= \Sigma_n n\Sigma^{-1}\hat{\mu}_n + \Sigma_n \Sigma_0^{-1}\mu_0 \\
&= \Sigma_0(\Sigma_0 + \frac{1}{n}\Sigma)^{-1}(\frac{1}{n}\Sigma)n\Sigma^{-1}\hat{\mu}_n + \frac{1}{n}\Sigma(\frac{1}{n}\Sigma + \Sigma_0)^{-1}\Sigma_0\Sigma_0^{-1}\mu_0 \\
&= \Sigma_0(\Sigma_0 + \frac{1}{n}\Sigma)^{-1}(\frac{1}{n}\Sigma)\hat{\mu}_n + \frac{1}{n}\Sigma(\frac{1}{n}\Sigma + \Sigma_0)^{-1}\mu_0
\end{aligned}$$

Which is Eqs.46.