GeoModels Tutorial: analysis of (large) spatio-temporal data using Gaussian random fields

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Introduction

In this tutorial we show how to analyze geo-referenced spatio temporal data using Gaussian random fields (RFs) with the R package GeoModels when the number of space-time location is relatively large.

We first load the R libraries needed for the analysis and set the name of the model in the GeoModels package:

```
rm(list=ls())
require(devtools)
install_github("vmoprojs/GeoModels-OCL")
require(GeoModels)
require(fields)
model="Gaussian" # model name in the GeoModels package
set.seed(12)
```

Simulation of a space-time Gaussian random field

Let us consider a space-time Gaussian RF $Z = \{Z(s,t), s \in S, t \in B\}$, where s represents a location in the domain S and t represents a temporal instant the domain B. We assume that Z is stationary with zero mean, unit variance and correlation function given by $\rho(\mathbf{h}, u) = \text{cor}(Z(s + \mathbf{h}, t + u), Z(s, t))$.

Then we consider the RF $Y = \{Y(s,t), s \in S, t \in T\}$ defined by the location and scale transformation:

$$Y(s,t) = \mu(s,t) + \sigma Z(s,t)$$
(1)

where $\mu(s,t) = X(s,t)^T \boldsymbol{\beta}$ and X(s,t) is a k-dimensional vector of covariates and $\boldsymbol{\beta} = (\beta_1, \dots, \beta_k)^T$ is a k-dimensional vector of (unknown) parameters (in this tutorial we fix k = 2). Then $\mathbb{E}(Y(s,t)) = X(s,t)^T \boldsymbol{\beta}$, $\operatorname{var}(Y(s,t)) = \sigma^2$ and $\operatorname{cov}(Y(s+h,t+u),Y(s,t)) = \sigma^2 \rho(\boldsymbol{h},u)$.

Suppose we want to simulate a realization of Y at $t_1 = 1, t_2 = 2, ..., t_T = 25, T = 25$ temporal instants and N = 400 spatial locations uniformly distributed in the unit square. The total number of space-time locations is given by NT = 10000.

We first set the temporal instants and then the spatial coordinates with associated covariates.

```
coordt=1:25  # number of temporal instants
T=length(coordt)
NN=400  # number of spatial locations
x = runif(NN, 0, 1); y = runif(NN, 0, 1)
coords=cbind(x,y)
X=cbind(rep(1,NN*T),runif(NN*T))
```

We then specify the mean, variance and nugget parameters

```
mean = 0.5; mean1= -0.25
sill=2; nugget=0
```

where mean, mean1 and sill are respectively β_1 , β_2 and σ^2 .

A possible approach to simulate a large (exact) realization from a Gaussian spacetime RF is to consider a compactly supported correlation space-time correlation function combined with cholesky decomposition algorithms for sparse matrices.

In this tutorial we assume a simple spatially isotropic and symmetric in time separable Wendland model

$$\rho((\boldsymbol{h}, u); \alpha_s, \alpha_t, \mu_s, \mu_t) = \left(1 - \frac{||\boldsymbol{h}||}{\alpha_s}\right)_+^{\mu_s} \left(1 - \frac{|u|}{\alpha_t}\right)_+^{\mu_t}$$
(2)

Then we set the name of the correlation model and the associated parameters:

```
corrmodel="Wend0_Wend0";
scale_s=0.2; scale_t=2
```

where scale_s and scale_t corresponds to α_s and α_t , the compact supports of the correlation model. We are now ready to simulate the space time Gaussian RF using the function GeoSim:

Note that the option sparse=TRUE allows to consider algorithms for sparse matrices when performing Cholesky decomposition, as described in the package spam (Gerber et al. (2017)). Informations about the sparsity of the covariance matrix can be obtained though the function GeoCovmatrix with the following code

This means that (approximatively) 98% of the covariance matrix are zeros *i.e* the matrix is highly sparsed.

Estimation of a Gaussian space-time random field

Given a space-time realization $\{Y(s_i, t_l), l = 1...T, i = 1,...N\}$, let $f_Y(y_{il}, y_{jk})$ the Gaussian density of a pair of observations $Y(s_i, t_l)$ and $Y(s_j, t_k)$. Then, the pairwise likelihood function is defined as (Bevilacqua et al. (2012); Bevilacqua and Gaetan (2015)):

$$pl(\boldsymbol{\theta}) = \sum_{i,j,l,k \in D} log(f_Y(y_{il}, y_{jk})) w_{ijlk}$$
(3)

where

$$D = \begin{cases} l = 1 \dots T, & i = 1, \dots, N, & k = l, \dots, T \\ j = i + 1, \dots, N & \text{if} \quad l = k \\ j = 1, \dots, N & \text{if} \quad l > k \end{cases}$$

and w_{ijlk} are non-negative weights, not depending on $\boldsymbol{\theta}$, specified as:

$$w_{ijlk} = \begin{cases} 1 & ||\boldsymbol{s}_i - \boldsymbol{s}_j|| < d_s, |t_l - t_k| < d_t \\ 0 & \text{otherwise} \end{cases}$$
 (4)

and in this case $\boldsymbol{\theta} = (\mu, \sigma^2, \alpha_s, \alpha_t)^T$. The pairwise likelihood estimator $\hat{\boldsymbol{\theta}}_{pl}$ is obtained maximizing (3) with respect to $\boldsymbol{\theta}$. In the GeoModels package we can choose the fixed parameters and the parameters that must be estimated. Pairwise likelihood estimation is performed with the function GeoFit:

```
## estimation with pairwise likelihood
start=list(mean=mean, mean1=mean1,scale_s=scale_s,scale_t=scale_t,sill=sill)
fixed=list(nugget=nugget,power2_s=4,power2_t=4)
```

Note that the option maxdist=0.04 and maxtime=1 set the (arbitrary) compact supports of the weight function (4) i.e. $d_s = 0.04$ and $d_t = 1$. A suitable choice of the weights allows to improve both the statistical and computational efficiency (Bevilacqua and Gaetan (2015))

The option GPU=0,local=c(1,1) allows to speed up time performance. The first parameter (GPU=0) sets the computing device. You can call the available devices in your computer through DeviceInfo() and choose the associated number. The second argument (local=c(1,1)) lets you set the number of local work-items of the OpenCL setup.

The object fit include informations about the pairwise likelihood estimation

```
fit
Maximum Composite-Likelihood Fitting of Gaussian Random Fields
Setting: Marginal Composite-Likelihood
Model associated to the likelihood objects: Gaussian
Type of the likelihood objects: Pairwise
Covariance model: Wend0_Wend0
Number of spatial coordinates: 400
Number of dependent temporal realisations: 25
Type of the random field: univariate
Number of estimated parameters: 5
Type of convergence: Successful
Maximum log-Composite-Likelihood value: -144544.23
Estimated parameters:
  mean
        mean1 scale_s scale_t
                               sill
0.5556 -0.2681
             0.1983
                      2.0416
                              1.9484
```

Checking model assumptions

Given the estimation of the mean regression and sill parameters, the estimated residuals

$$\widehat{Z(s,t)} = \frac{Y(s,t) - X(s,t)^T \widehat{\boldsymbol{\beta}}}{(\widehat{\sigma}^2)^{\frac{1}{2}}}$$

can be viewed as a realization of a zero mean stationary Gaussian RF with correlation function $\rho(h, u)$. The residuals can be computed using the GeoResiduals function:

```
res=GeoResiduals(fit) # computing residuals
```

Then the marginal distribution assumption on the residuals can be graphically checked for instance with a qq-plot (Figure 1, left part):

```
### checking model assumptions: marginal distribution
qqnorm(unlist(res$data))
abline(0,1)
```

The correlation model assumption can be checked comparing the empirical and the estimated space-time semivariogram functions using the GeoVariogram and GeoCovariogram functions (Figure 1, right part):

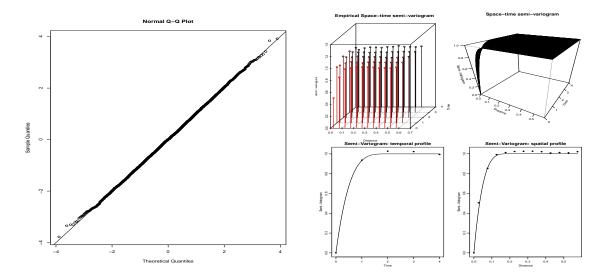


Figure 1: Left: QQ-plot for the residuals of the space-time Gaussian RF. Right: space-time empirical vs estimated semi-variogram function for the residuals

Prediction of space-time Gaussian random fields

For a given space time location (s_0, t_0) with associated covariates $X(s_0, t_0)$, the optimal prediction of Gaussian RF is computed as:

$$\widehat{Y}(\boldsymbol{s}_0, t_0) = X(\boldsymbol{s}_0, t_0)^T \widehat{\boldsymbol{\beta}} + \sum_{l=1}^T \sum_{i=1}^N \lambda_{l,i} [Y(\boldsymbol{s}_i, t_l) - X(\boldsymbol{s}_i, t_l)^T \widehat{\boldsymbol{\beta}}]$$
 (5)

where the vector of weights $\boldsymbol{\lambda} = (\lambda_{1,1}, \dots, \lambda_{T,N})'$ is given by $\boldsymbol{\lambda} = R^{-1}\boldsymbol{c}$ and

- $c = (cor(Y(s_0, t_0), Y(s_1, t_1)), \dots, cor(Y(s_0, t_0), Y(s_N, t_T)))^T$.
- $R = [[cor(Y(s_i, t_l), Y(s_j, t_k)]_{l,k=1}^T]]_{i,j=1}^N$ is the correlation matrix.

Kriging can be performed using the GeoKrig function. We need just to specify the spatial location and temporal instants to predict.

In this example we consider a spatial regular grid and one temporal instant:

```
xx=seq(0,1,0.03)
loc_to_pred=as.matrix(expand.grid(xx,xx))  # locations to predict
n_loc=nrow(loc_to_pred)
times=c(2)  #time to predict
```

Moreover we need to specify the associated covariates:

```
Xloc=cbind(rep(1,n_loc),runif(n_loc))
```

Then the optimal linear prediction (5), using the estimated parameters, can be performed using the GeoKrig function:

A kriging map for the temporal instant t=2 with associate mean square error (Figure 2) can be obtained with the following code:

```
par(mfrow=c(1,3))
colour <- rainbow(100)
## observed data at time 2
quilt.plot(coords[,1],coords[,2],ss1[2,],col=colour,main ="Time=2")</pre>
```

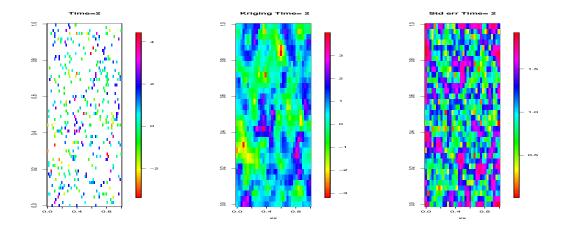


Figure 2: From left to right: observed spatial data at time t = 2, associated kriging map and mean square error map.

References

Bevilacqua, M. and C. Gaetan (2015). Comparing composite likelihood methods based on pairs for spatial Gaussian random fields. *Statistics and Computing* 25, 877–892.

Bevilacqua, M., C. Gaetan, J. Mateu, and E. Porcu (2012). Estimating space and spacetime covariance functions for large data sets: a weighted composite likelihood approach. *Journal of the American Statistical Association* 107, 268–280.

Gerber, F., K. Moesinger, and R. Furrer (2017). Extending R packages to support 64-bit compiled code: An illustration with spam64 and GIMMS NDVI3g data. Computer & Geoscience 104, 109–119.