GeoModels Tutorial: analysis of global spatial data on the planet Earth using Gaussian random fields

Moreno Bevilacqua

October 30, 2018

Introduction

This tutorial shows how to analyze spherical data with the R package GeoModels (Bevilacqua and Morales-Oñate (2018)). Albeit the main focus of the tutorial is on Gaussian random fields (GRFs) defined over a sphere, any of the non GRF in the package GeoModels can be used for the analysis of data collected over spheres.

We denote with \mathbb{S}^2 the unit sphere of \mathbb{R}^3 , defined as

$$\mathbb{S}^2 = \{ \boldsymbol{x} \in \mathbb{R}^3 : ||\boldsymbol{x}|| = R \}.$$

Since a sphere of radius R = 6371 is a satisfactory approximation of planet Earth, the tutorial focuses on this specific sphere and shows how to simulate, estimate and predict GRFs on planet Earth.

We start by loading the libraries needed for the analysis and set the radius of the sphere:

```
>rm(list=ls())
>library(GeoModels)
>library(mapproj)
>library(globe)
>library(fields)
>library(sphereplot)
>radius=6371  # radius of the planet earth
>set.seed(1891)
```

Gaussian random fields defined on planet Earth

We consider a zero mean and unit variance GRF $Y = \{Y(\boldsymbol{x}), \boldsymbol{x} \in \mathbb{S}^2\}$ defined on a sphere of radius R = 6371. A point on \mathbb{S}^2 can be parametrized trough spherical coordinates using radians as (R, lon^*, lat^*) , where $lon^* \in [-\pi, \pi]$ is the longitude and $lat^* \in [-\pi/2, \pi/2]$ is the latitude. Alternatively, one can use decimal degree (R, lon, lat) where $lon \in [-180, 180]$ and $lat \in [-90, 90]$ (the GeoModels package uses the second parameterization).

The great circle (GC) distance corresponds to the shortest path joining two points on the spherical surface and it is the natural distance to be used when analyzing data on a sphere. Alternatively, one can compute the chordal (CH) distance, that is the segment below the arc joining any pair of points located over the spherical shell. Let us give some details on the computation of GC and CH distances. Given two location sites in longitude and latitude (expressed in decimal degrees), that is, $\mathbf{x}_i = (\text{lon}_i, \text{lat}_i)$ and $\mathbf{x}_j = (\text{lon}_j, \text{lat}_j)$, and given the radius R of the Earth, the GC distance, d_{GC} , is given by:

$$d_{GC}(\boldsymbol{x}_i, \boldsymbol{x}_j) = R\theta_{ij},$$

where

$$\theta_{ij} = \left[\arccos\{\sin a_i \sin a_j + \cos a_i \cos a_j \cos(b_i - b_j)\}\right]$$

is the GC distance on the unit sphere. Here, $a_i = (\text{lat}_i)\pi/180$, $a_j = (\text{lat}_j)\pi/180$, $b_i = (\text{lon}_i)\pi/180$, $b_j = (\text{lon}_j)\pi/180$. The CH distance is instead uniquely defined through the relation

$$d_{CH}(\boldsymbol{x}_i, \boldsymbol{x}_j) = 2Rsin(\theta_{ij}/2). \tag{1}$$

CH obviously underestimates the GC distance, and the approximation error increases with the size of the considered portion of the planet.

We assume $cor(Y(\boldsymbol{x}_i), Y(\boldsymbol{x}_j))) = \rho(d_{GC}(\boldsymbol{x}_i, \boldsymbol{x}_j))$ where $\rho : [0, R\pi] \to \mathbb{R}$ is continuous with $\rho(0) = 1$. Under this setting, the correlation function is called *geodesically isotropic* (Gneiting, 2013; Porcu et al., 2016). We then consider the location and scale transformation

$$Z(\mathbf{x}) = \mu(\mathbf{x}) + \sigma Y(\mathbf{x})$$

where $\mu(x) \in \mathbb{R}$ is a spherical varying mean and $\sigma > 0$. The package GeoModels allows a mean regression specification, that is $\mu(x) = M(x)^T \beta$ where M(x) is a vector of spherical covariates. Nevertheless, for the sake of simplicity this tutorial assumes a constant mean: $\mu(x) = \mu$.

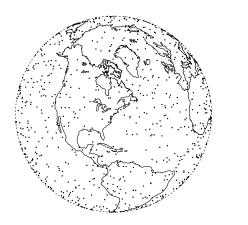
Then $\mathbb{E}(Z(\boldsymbol{x})) = \mu$ and $\operatorname{cov}(Z(\boldsymbol{x}_i), Z(\boldsymbol{x}_j))) = \sigma^2 \rho(d_{GC}(\boldsymbol{x}_i, \boldsymbol{x}_j))$ where σ^2 is the variance parameter and in order to obtain a realization from a spherical RF we need to specify a mean, a variance parameter and a valid (*i.e.* positive definite) geodesically isotropic parametric correlation function $\rho_{\gamma}(\cdot)$.

We first set the spherical coordinates in decimal degrees using the function pointsphere of the package sphereplot:

```
>NN=1500 ## number of location sites on the planet Earth
>coords=pointsphere(NN,c(-180,180),c(-90,90),c(radius,radius))[,1:2]
```

The locations on planet Earth can be visualized (Figure 1) using some functions of the package globe REF with the following code:

```
>globeearth(eye=place("newyorkcity"))
>globepoints(loc=coords,pch=20,cex=0.4)
>globeearth(eye=place("everest"))
>globepoints(loc=coords,pch=20,cex=0.4)
```



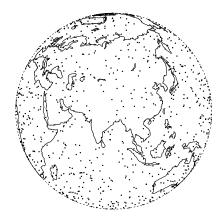


Figure 1: Location sites on the planet Earth used in this tutorial

As for the correlation function, we assume an Askey model. This is a valid correlation model in \mathbb{S}^2 (Gneiting, 2013):

$$\rho_{\alpha,\delta}(d_{GC}) = \begin{cases} (1 - d_{GC}/\alpha)^{\delta} & \theta < \alpha \\ 0 & \text{otherwise} \end{cases}$$
 (2)

This model is compactly supported provided $\alpha < R\pi$ which means that the associated covariance matrix is sparse, an interesting feature from computational point of view. We set the Askey model with associated parameters and the other parameters of the GRF.

```
>corrmodel = "Wend0" ## correlation model and parameters
>scale=radius*0.4
>power2=4
>sill=1
>nugget=0
>mean=0
```

where the scale parameter corresponds to α (the compact support of the correlation model) and sill is the variance parameter σ^2 . We are now ready to simulate a GRF on the sphere using the function GeoSim:

Note that the option distance="Geod" coupled with the location sites given in lon/lat format in degimal degree (coordx=coords) and a valid correlation model on the sphere (corrmodel) with a given radius, allows to perform a simulation based on Cholesky decomposition of the GRF using GC.

Estimation of Gaussian random fields on planet Earth

Given a realization $\mathbf{Z} = (Z(\mathbf{x}_1), Z(\mathbf{x}_2), \dots, Z(\mathbf{x}_N))^T$ from a GRF defined on \mathbb{S}^2 with correlation (2), we can perform maximum likelihood estimation maximizing the log-likelihood function

$$l(\boldsymbol{\beta}) = -0.5log(|\sigma^2 C_{\alpha,\delta}|) - 0.5 \frac{(\boldsymbol{Z} - \mu \boldsymbol{1})^T C_{\alpha,\delta}^{-1} (\boldsymbol{Z} - \mu \boldsymbol{1})}{\sigma^2}$$

with respect to $\boldsymbol{\beta} = (\mu, \sigma^2, \alpha, \delta)^T$ and where $C_{\alpha, \delta} = [\rho_{\alpha, \delta}(d_{GC}(\boldsymbol{x}_i, \boldsymbol{x}_j))]_{i,j=1}^N$ is the correlation matrix associated to (2).

We can use the GeoFit function to perform maximum likelihood estimation (we choose the fixed parameters and the parameters that must be estimated)

Note that the option sparse=TRUE allows to exploit specific algorithms for sparse matrices implemented in the spam package (Furrer and Sain (2010)) when performing maximum likelihood estimation.

Computation of the GC distances involved in the covariance matrix (given in kms) can be obtained using the function rdist.earth of the package fields:

```
>geod_ds1=rdist.earth(coords, miles=F,R=1)
>geod_ds=geod_ds1*radius
>max_geod=max(geod_ds)
```

Alternatively, pairwise likelihood estimation is obtained by maximizing the function

$$pl(\beta) = \sum_{i=1}^{N-1} \sum_{j=i+1}^{N} log(f_z(z_i, z_j; \beta)) w_{ij}$$
 (3)

with respect to β . It can be performed with the following code

In (3) $f_z(z_i, z_j; \boldsymbol{\beta})$ is the Gaussian bivariate density and c_{ij} are non-negative weights, not depending on $\boldsymbol{\beta}$, specified as:

$$w_{ij} := \begin{cases} 1 & d_{GC}(\boldsymbol{x}_i, \boldsymbol{x}_j) < d^* \\ 0 & \text{otherwise} \end{cases}$$
 (4)

Note that the option maxdist=max_geod/70 sets the compact support d^* of the weight function (4) as a fraction of the maximum GC distance (285 KM approximatively). A suitable choice of the weights allows to improve both the statistical and computational efficiency on the pairwise likelihood method (Bevilacqua and Gaetan (2015)).

The two estimates can be obtained from the objects fit_geo_ml and fit_geo as:

In the GeoModels package, estimation can be performed using also the CH distance, or the Euclidean distance after a suitable projection of the spherical coordinates.

Computation of CH distances involved in the covariance matrix can be obtained by using (1):

```
>chor_ds=2*radius*sin(geod_ds1/2)
>max_chor=max(chor_ds)
```

and pairwise likelihood estimation using chordal distances can be performed setting distance="Chor":

Finally, we consider pairwise likelihood estimation using a suitable projection of the spherical points and the Euclidean distance. We use the function mapprojec of the package mapproj in order to obtain projected coordinates using sinusoidal projection

```
>prj=mapproject(coords[,1], coords[,2], projection="sinusoidal")
>coords_prj=cbind(prj$x,prj$y)
```

and we can visualize the projected points (Figure 2) with the following code:

```
>sinusoidal.proj = map(database= "world", ylim=c(-90,90),
>xlim=c(-180,180),col="grey80", fill=TRUE,
>plot=FALSE, projection="sinusoidal")
>map(sinusoidal.proj)
>points(coords_prj,pch=20,cex=0.4)
```

Other types of projections can be chosen in the package mapproj such us, for instance, the mercator or cylindrical projections.

Finally, we can estimate with pairwise likelihood using Euclidean distances (first we need to rescale the projected coordinates with the radius)

A comparison of the three pairwise estimates with the three type of distances can be obtained as:

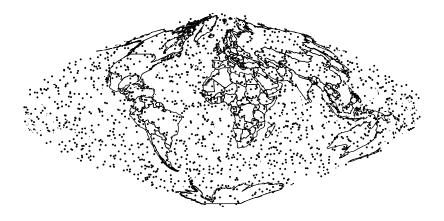


Figure 2: Sinusoidal projection

0.03336823	2328.192	1.038559
>fit_eucl\$param		
mean	scale	sill
0.04714257	2077.088	0.9466116

As expected, the parameter that is most affected by the choice of the distance is the spatial scale parameter. This is not surprising, since CH and euclidean distances are both approximations of the GC distance. Boxplots of the distances (Figure 4) can be useful to show the different behaviour of the three type of distances in this example:

Given the estimation of the mean and variance parameters, the estimated residuals

$$\widehat{Y}(\boldsymbol{x}_i) = \frac{Z(\boldsymbol{x}_i) - \hat{\mu}}{(\widehat{\sigma}^2)^{\frac{1}{2}}} \quad i = 1, \dots, N$$

can be viewed as a realization of a zero mean unit variance spherical GRF with correlation function (2). Using maximum likelihood estimates, the estimated residuals can be computed using the GeoResiduals function:

```
>res=GeoResiduals(fit_geo_ml)
```

Then the marginal distribution assumption on the residuals can be graphically checked, for instance, with a qq-plot (Figure 3, left part):

```
### checking model assumptions: marginal distribution
>qqnorm(unlist(res$data))
>abline(0,1)
```

The correlation model assumption can be graphically checked comparing the empirical and the estimated semivariogram functions using the GeoVariogram and GeoCovariogram functions (Figure 3, right part):

Prediction of Gaussian random fields on planet Earth

For a given location $x_0 \in \mathbb{S}^2$, the optimal prediction in this example is computed by:

$$\widehat{Z}(\boldsymbol{x}_0) = \widehat{\mu} + \sum_{i=1}^{N} \lambda_i [Z(\boldsymbol{x}_i) - \widehat{\mu}]$$
(5)

where $\hat{\mu}$ is a mean estimation and the vector of weights $\boldsymbol{\lambda} = (\lambda_1, \dots, \lambda_N)'$ is given by $\boldsymbol{\lambda} = \widehat{C}_{\hat{\alpha}, \delta}^{-1} \hat{c}_{\hat{\alpha}, \delta}$ where $\widehat{C}_{\hat{\alpha}, \delta}$ is the estimated correlation matrix and $\hat{c}_{\hat{\alpha}, \delta}$ is the estimated correlation between the locations and the point to predict. Moreover, the associated mean squared error is computed through

$$MSE(\widehat{Z}(\boldsymbol{x}_0)) = \hat{\sigma}^2 (1 - \hat{\boldsymbol{c}}_{\hat{\alpha},\delta}^T \widehat{C}_{\hat{\alpha},\delta}^{-1} \hat{\boldsymbol{c}}_{\hat{\alpha},\delta}).$$
 (6)

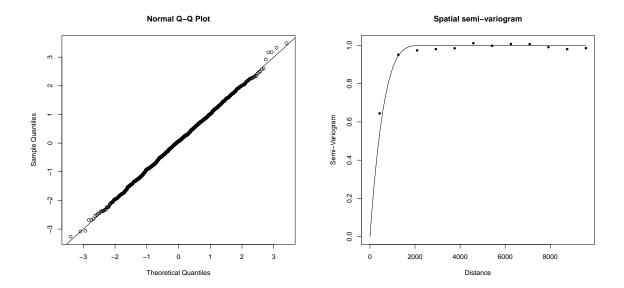


Figure 3: Left: QQ-plot for the model residuals. Right: empirical vs estimated semi-variogram function for the residuals

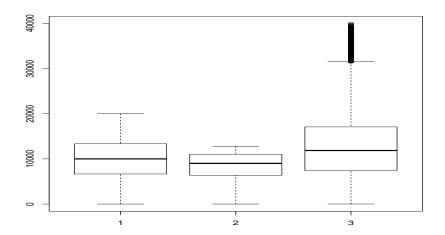


Figure 4: Boxplot of great circle, chordal and euclidean distances

Optimal prediction and associated MSE can be computed using the GeoKrig function. Suppose we want to predict our GRF around the north pole area. We need just to specify the spherical location to predict (in lon/lat format):

```
## location to predict given in lon/lat format
>loc_to_pred=as.matrix(expand.grid(seq(-180,180,2),seq(-90,90,1)))
```

Then, the optimal predictor (5) and associated MSE (6) (using the maximum likelihood estimated parameters) can be performed with the following code:

```
>param_est<-as.list(c(fixed,fit_geo_ml$param))
## computing optimal prediction. and associated MSE
>pr_geo=GeoKrig(data=data,loc=loc_to_pred,coordx=coords,
>corrmodel=corrmodel,radius=radius,distance="Geod",
>param=param_est,sparse=TRUE,mse=TRUE)
```

The matrix of locations to predict, predictions and associated MSE can be obtained as:

and a graphical visualization (Figure 5) of the prediction and associated MSE can be obtained with the following code:

```
>globeearth(eye=place("newyorkcity"))
>globepoints(loc=loc_to_pred,pch=20,
>col = heat.colors(length(pr_geo$pred),alpha=0.1)[rank(pr_geo$pred)])
>globeearth(eye=place("newyorkcity"))
>globepoints(loc=loc_to_pred,pch=20,
>col = cm.colors(length(pr_geo$mse),alpha=0.1)[rank(pr_geo$mse)])
```

References

Bevilacqua, M. and C. Gaetan (2015). Comparing composite likelihood methods based on pairs for spatial Gaussian random fields. *Statistics and Computing* 25, 877–892.

Bevilacqua, M. and V. Morales-Oñate (2018). GeoModels: A Package for Geostatistical Gaussian and non Gaussian Data Analysis. R package version 1.0.3-4.

Furrer, R. and S. R. Sain (2010). spam: a sparse matrix R package with emphasis on mcmc methods for Gaussian Markov random rields. *Journal of Statistical Software 36*, 1–25.

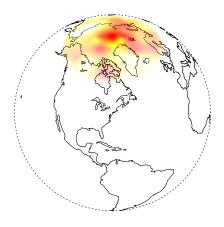




Figure 5: From left to right: prediction and associated MSE in the north pole area

Gneiting, T. (2013). Strictly and non-strictly positive definite functions on spheres. Bernoulli 19, 1327–1349.

Porcu, E., M. Bevilacqua, and M. Genton (2016). Spatio-temporal covariance and cross covariance functions of the great circle distance on a sphere. *Journal of the American Statistical Association* 111, 888–898.