

*Course: Computer and Communication Networks*

*Topic: Basic Queueing Theory - I (Analysis of M/M/-/- Type Queues  
and M/G/1 Queues)*

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# Basic Queueing Theory

## Kendall's Notations for Queues

A is the Inter-arrival time distribution

B is the Service time distribution

C is the Number of servers

D is the Maximum number of jobs that can be there in the system (waiting and in service)

*Default  $\infty$  for infinite number of waiting positions*

E is the Queueing models (FCFS, LCFS, SIRO, PS) *Default is FCFS*

M: exponential

D: deterministic

E<sub>k</sub>: Erlangian (order k)

G: general

M/M/1 or M/M/1/ $\infty$

Single server queue with Poisson arrivals, exponentially distributed service times and infinite number of waiting positions

The Poisson distribution with parameter  $\lambda$  is given by

$$\frac{(\lambda t)^n e^{-\lambda t}}{n!}$$

Where n is the number of arrivals. We find that if we set n= 0, the Poisson distribution gives us

$$e^{-\lambda t}$$

which is equal to  $P(T > t)$  from the exponential distribution.

The exponential distribution with parameter  $\lambda$  is given by

$$\lambda e^{-\lambda t} \text{ for } t \geq 0$$

If T is a random variable that represents inter arrival times with the exponential distribution, then

$$P(T \leq t) = 1 - e^{-\lambda t} \text{ and } P(T > t) = e^{-\lambda t}.$$



# *Basic Queueing Theory*

Little's Theorem states that  
 $N = \lambda W$

Result holds in general for virtually all types of queueing situations where  $\lambda$  = Mean arrival rate of jobs that actually enter the system

Note: Jobs blocked and refused entry into the system will not be counted in  $\lambda$

## The Poisson Arrivals See Time Average (PASTA) Property

$p_k(t) = P\{\text{system is in state } k \text{ at time } t\}$

$q_k(t) = P\{\text{an arrival at time } t \text{ finds the system in state } k\}$

$N(t)$  be the actual number in the system at time  $t$

$A(t, t+\Delta t)$  be the event of an arrival in the time interval  $(t, t+\Delta t)$

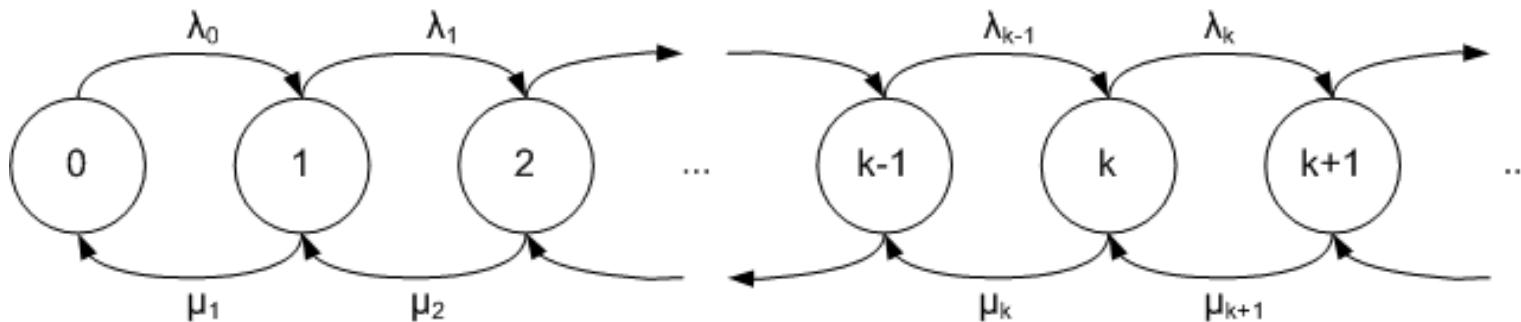
We have

$$\begin{aligned} q_k(t) &= \lim_{\Delta t \rightarrow 0} P\{N(t) = k \mid A(t, t + \Delta t)\} \\ &= \lim_{\Delta t \rightarrow 0} \frac{P\{A(t, t + \Delta t) \mid N(t) = k\} P\{N(t) = k\}}{P\{A(t, t + \Delta t)\}} = p_k(t) \end{aligned}$$

Because  $P\{A(t, t + \Delta t) \mid N(t) = k\} = P\{A(t, t + \Delta t)\}$

# Basic Queueing Theory: Equilibrium Solutions for M/M/-/- Queues

## M/M/1 (or M/M/1/ $\infty$ ) Queue



$$\left. \begin{array}{ll} \lambda_k = \lambda & \forall k \\ \mu_k = 0 & k=0 \\ = \mu & k=1, 2, 3, \dots \end{array} \right\}$$

For  $\rho < 1$

$$p_k = p_0 \left( \frac{\lambda}{\mu} \right)^k = p_0 \rho^k$$

$$p_0 = (1 - \rho)$$

$$N = \sum_{i=0}^{\infty} i p_i = \sum_{i=0}^{\infty} i \rho^i (1 - \rho) = \frac{\rho}{1 - \rho}$$

$$W_q = W - \frac{1}{\mu} = \frac{\rho}{\mu(1 - \rho)}$$

$$N_q = \lambda W_q = \frac{\rho^2}{(1 - \rho)}$$

## Birth-death process

The values in the circles represent the state of the birth-death process,  $k$ .

The system transitions between values of  $k$  by "births" and "deaths" which occur at rates given by various values of  $\lambda_i$  and  $\mu_i$ , respectively.

For a queueing system,  $k$  is the number of jobs in the system. Furthermore, for a queue, the arrival rates and departure rates are generally considered not to vary with the number of jobs in the queue.

Hence, to consider a single average rate of arrivals/departures per unit time to the queue.

Therefore, for a queue, the figure shown above has arrival rates of  $\lambda = \lambda_1, \lambda_2, \dots, \lambda_k$  and departure rates of  $\mu = \mu_1, \mu_2, \dots, \mu_k$

## Basic Queueing Theory: M/M/1/∞ Queue with Discouraged Arrivals

$$\left. \begin{array}{ll} \lambda_k = \frac{\lambda}{k+1} & \forall k \\ \mu_k = 0 & k=0 \\ = \mu & k=1, 2, 3, \dots \end{array} \right\} \text{For } \rho = \lambda/\mu < \infty$$

$$p_k = p_0 \prod_{i=0}^{k-1} \frac{\lambda}{\mu(i+1)} = p_0 \left( \frac{\lambda}{\mu} \right)^k \frac{1}{k!}$$

$$p_0 = \exp(-\frac{\lambda}{\mu})$$

$$N = \sum_{k=0}^{\infty} kp_k = \frac{\lambda}{\mu}$$

$$W = \frac{N}{\lambda_{eff}} = \frac{\lambda}{\mu^2 \left[ 1 - \exp(-\frac{\lambda}{\mu}) \right]} \quad \left. \right\} \lambda_{eff} = \sum_{k=0}^{\infty} \lambda_k p_k = \mu \left[ 1 - \exp(-\frac{\lambda}{\mu}) \right]$$



## Basic Queueing Theory: M/M/m/∞ Queue (m servers, infinite number of waiting positions)

$$\begin{aligned}\lambda_k &= \lambda & \forall k \\ \mu_k &= k\mu & 0 \leq k \leq (m-1) \\ &= m\mu & k \geq m\end{aligned}$$

$$\text{For } \rho = \lambda/\mu < m \quad p_k = p_0 \frac{\rho^k}{k!} \quad \text{for } k \leq m$$

$$= p_0 \frac{\rho^k}{m!m^{k-m}} \quad \text{for } k > m$$

$$p_0 = \left( \sum_{k=0}^{m-1} \frac{\rho^k}{k!} + \frac{m\rho^m}{m!(m-\rho)} \right)^{-1}$$

$$P\{\text{queueing}\} = \sum_{k=m}^{\infty} p_k = C(m, \rho) = p_0 \frac{m\rho^m}{m!(m-\rho)}$$



## *Basic Queueing Theory: M/M/m/m Queue (m servers loss system and no waiting)*

$$\begin{aligned}\lambda_k &= \lambda(K - k) & k < K \\ &= 0 & \text{otherwise}\end{aligned}$$

$$\begin{aligned}\mu_k &= \mu & k \leq K \\ &= 0 & \text{otherwise}\end{aligned}$$

For

$$p_k = p_0 \rho^k \frac{K!}{(K - k)!} \quad k=1, \dots, K$$

$$\rho = \frac{\lambda}{\mu} < \infty \quad p_0 = \frac{1}{\sum_{k=0}^K \rho^k \frac{K!}{(K - k)!}}$$



# *Basic Queueing Theory: M/G/1 queue*

In queueing theory, an M/G/1 queue is a queue model where arrivals are Markovian, service times have a general distribution and it has a single server.

It is an extension of the M/M/1 queue, where service times must be exponentially distributed.

M(memoryless): Poisson arrival process, intensity  $\lambda$   
G(general): general holding time distribution, mean  $\bar{S} = 1/\mu$   
1 : single server, load  $\rho = \lambda \bar{S}$   
(in a stable queue one has  $\rho < 1$ )

Here, the model name is written in Kendall's notation.

Example: fixed head hard disk.

The probability per time unit for a transition from the state {N=n} to the state {N=n-1}, i.e. for a departure of a customer, depends also on the time the customer in service has already spent in the server;

We start with the derivation of the expectation of the waiting time W. W is the time the customer has to wait for the service (time in the “waiting room”, i.e. in the actual queue), R is residual service time

$$E[W] = \underbrace{\frac{E[N_q]}{\text{number of waiting customers}} \cdot \frac{E[S]}{\text{mean service time}}}_{\text{mean time needed to serve the customers ahead in the queue}} + E[R]$$

# *Basic Queueing Theory*

R is the remaining service time of the customer in the server (unfinished work expressed as the time needed to discharge the work).

*If the server is idle (i.e. the system is empty), then R= 0.*

- In order to calculate the mean waiting time of an arriving customer one needs the expectation of Nq(number of waiting customers) at the instant of arrival.
- Due to the PASTA property of Poisson process, the distributions seen by the arriving customer are the same as those at an arbitrary instant. The key observation is that by Little's result the mean queue length E[Nq] can be expressed in terms of the waiting time.

$$E[N_q] = \lambda E[W]$$

It remains to determine E[R].

$$E[W] = \frac{E[R]}{1 - \rho}$$

$$\rho = \lambda E[S]$$

