

Univariate Mixture model - gaussian [Without  $z_{ij}$ ]

$$P(\vec{x}_i | \theta) = \sum_{j=1}^K P_j P(x_i | \theta_j)$$

$P_j$  - mixing weight

$\theta_j$  - parameters of pdf

$$\sum_{j=1}^K P_j = 1$$

$$\theta = \{\theta_1, \theta_2, \dots, \theta_K, P_1, \dots, P_K\}$$

The PDF is a gaussian distribution

$$P(x_i | \theta_j) = \frac{1}{\sqrt{2\pi}\sigma_j} e^{-1/2 \left[ \frac{x_i - \mu}{\sigma_j} \right]^2}$$

The likelihood function is

$$P(\vec{x} | \theta_j) = \prod_{i=1}^N \sum_{j=1}^K P_j \cdot P(x_i | \theta_j)$$

The log likelihood function is

$$L(\phi(x, \theta_j, \lambda_1, \dots, \lambda_K, 1)) = \sum_{i=1}^N \log \left( \sum_{j=1}^K P_j \cdot P(x_i | \theta_j) + \lambda_1 P_1 + \lambda_2 P_2 + \dots + \lambda_K P_K + \lambda \left( 1 - \sum_{j=1}^K P_j \right) \right)$$

To derive the maximum log likelihood, differentiate w.r.t all parameters & equate to zero.

$$\textcircled{1} \quad \frac{\partial \phi(x, \theta_j, \lambda_1, \dots, \lambda_K, 1)}{\partial P_j} = 0$$

$$\textcircled{2} \quad \frac{\partial \phi(x, \theta_j, \lambda_1, \dots, \lambda_K, 1)}{\partial \theta_j} = 0$$

$$(3) \frac{\partial \phi}{\partial \lambda_j} (x, \theta_j, \lambda_1, \dots, \lambda_k, \lambda) = 0$$

$$(4) \frac{\partial \phi}{\partial \lambda} (x, \theta_j, \lambda_1, \dots, \lambda_k, \lambda) = 0$$

Take (3) and (4)

$$\text{Solve } \frac{\partial \phi}{\partial \lambda_j} (x, \theta_j, \lambda_1, \dots, \lambda_k, \lambda) = p_j$$

$\neq 0$

As it contradicts constraint

$$\lambda_1 = \lambda_2 = \dots = \lambda_k = 0$$

$$\frac{\partial \phi}{\partial \lambda} (x, \theta_j, \lambda_1, \dots, \lambda_k, \lambda) = 1 - \sum_{j=1}^k p_j = 0$$

$$\sum_{j=1}^k p_j = 1$$

→ (5)

$$\text{Take (1) } \frac{\partial \phi}{\partial p_j} (x, \theta_j, \lambda_1, \dots, \lambda_k, \lambda) = \sum_{i=1}^N \frac{1}{\sum_{j=1}^k p_j \cdot P(x_i | \theta_j)} \frac{\partial}{\partial p_j} \left\{ \sum_{j=1}^k p_j \cdot P(x_i | \theta_j) \right\} - \lambda$$

$$\Rightarrow \frac{\sum_{i=1}^N P(x_i | \theta_j)}{\sum_{j=1}^k p_j \cdot P(x_i | \theta_j)} - \lambda = 0$$

Multiplying and dividing by  $p_j$

$$\frac{1}{p_j} \left\{ \frac{\sum_{i=1}^N p_j \cdot P(x_i | \theta_j)}{\sum_{j=1}^k p_j \cdot P(x_i | \theta_j)} \right\} - \lambda$$

$$\text{By Bayes rule, } \frac{1}{p_j} \sum_{i=1}^N P(j | x_i) - \lambda = 0$$



$$\Rightarrow \hat{P}_j = \frac{1}{\Lambda} \sum_{i=1}^N P(j|x_i) \rightarrow (6)$$

Sub in (5)

$$\sum_{j=1}^K \frac{1}{\Lambda} \sum_{i=1}^N P(j|x_i) = 1$$

$$\frac{1}{\Lambda} \sum_{i=1}^N \sum_{j=1}^K P(j|x_i) = 1$$

$$\frac{1}{\Lambda} N = 1$$

$$\boxed{\Lambda = N} \rightarrow (7)$$

Sub in (6)

$$\boxed{\hat{P}_j = \frac{1}{N} \sum_{i=1}^N P(j|x_i)} \rightarrow (8)$$

Solve (2)

$$\frac{\partial \phi(x, \theta, d_1, \dots, d_K, \Lambda)}{\partial \theta_j} = \frac{\sum_{i=1}^N \frac{1}{\sum_{j=1}^K P_j \cdot P(x_i|\theta_j)}}{\partial \theta_j} \sum_{j=1}^K P_j \cdot P(x_i|\theta_j)$$

$$= \frac{\sum_{i=1}^N P_j \frac{\partial}{\partial \theta_j} \sum_{j=1}^K P(x_i|\theta_j)}{\sum_{j=1}^K P_j P(x_i|\theta_j)}$$

Multiplying and dividing by  $P(x_i | \theta_j)$

$$= \frac{\sum_{i=1}^N p_j \cdot P(x_i | \theta_j)}{\sum_{j=1}^K p_j \cdot P(x_i | \theta_j)} \cdot \frac{\frac{\partial}{\partial \theta_j} \sum_{i=1}^N P(x_i | \theta_j)}{P(x_i | \theta_j)}$$

$$= \sum_{i=1}^N P(j | x_i) \frac{\partial}{\partial \theta_j} \log P(x_i | \theta_j)$$

$$\log P(x_i | \theta_j) = \log \left\{ \frac{1}{\sqrt{2\pi} \sigma_j} e^{-1/2 \left( \frac{x_i - \mu_j}{\sigma_j} \right)^2} \right\}$$

$$= \log (2\pi)^{-1/2} + \log (\sigma_j)^{-1/2} + \log \left[ -\frac{1}{2} \frac{(x_i - \mu_j)^2}{\sigma_j^2} \right]$$

$$= -\frac{1}{2} \log 2\pi - \frac{1}{2} \log \sigma_j - \frac{1}{2} \frac{(x_i - \mu_j)^2}{\sigma_j^2}$$

Deriving mean,  $\theta_j = (\mu_j, \sigma_j^2)$

$$\frac{\partial}{\partial \mu_j} \log P(x_i | \theta_j) = \frac{\partial}{\partial \mu_j} \left[ -\frac{1}{2} \frac{(x_i - \mu_j)^2}{\sigma_j^2} \right]$$

$$= \frac{1}{\sigma_j^2} (x_i - \mu_j)$$

$$\frac{\partial \phi}{\partial \mu_j} = \sum_{i=1}^N P(j | x_i) \left[ \frac{(x_i - \mu_j)}{\sigma_j^2} \right]$$

$$0 = \sum_{i=1}^N P(j | x_i) x_i - \sum_{i=1}^N P(j | x_i) \mu_j$$

$$\sum_{i=1}^N P(j | x_i) x_i = \sum_{i=1}^N P(j | x_i) \mu_j$$

$$\hat{\mu}_j = \frac{\sum_{i=1}^N P(j | x_i) x_i}{P(j | x_i)}$$



$$\frac{\partial}{\partial \sigma_j} \log P(x_i | \theta_j) = -\frac{1}{2\sigma_j} + \left( \frac{-1}{2\sigma_j^3} (x_i - \mu_j)^2 \right)$$

$$\frac{\partial \phi}{\partial \sigma_j} = \sum_{i=1}^N P(j | x_i) \left\{ \frac{-1}{\sigma_j} + \frac{1}{\sigma_j^3} (x_i - \mu_j)^2 \right\}$$

$$0 = \sum_{i=1}^N P(j | x_i) \left\{ -\sigma_j^2 + (x_i - \mu_j)^2 \right\}$$

$$\sum_{i=1}^N P(j | x_i) \sigma_j^2 = \sum_{i=1}^N P(j | x_i) (x_i - \mu_j)^2$$

$$\sigma_j^2 = \frac{\sum_{i=1}^N P(j | x_i) (x_i - \mu_j)^2}{\sum_{i=1}^N P(j | x_i)}$$

② Multivariate Mixture model (without  $z_{ij}$ )

③ Binomial Mixture model - With and without  $z_{ij}$   
 - univariate  
 - multivariate

④ Bernoulli mixture model

⑤ Multinomial mixture model