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AI1103-Assignment-4

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Download all python codes from

https://github.com/Vikhyath-vec/AI1103/tree/main/ Assignment-4/codes

and latex-tikz codes from

https://github.com/Vikhyath-vec/AI1103/blob/main/Assignment-4/Assignment-4.tex

QUESTION

Let $\{X_j\}$ be a sequence of independent Bernoulli random variables with $\mathbb{P}(X_j = 1) = \frac{1}{4}$ and let $Y_n = \frac{1}{n} \sum_{j=1}^n X_j^2$. Then Y_n converges, in probability, to ______.

SOLUTION

A sequence of random variables $Y_1, Y_2, Y_3...$ converges, in probability, to a random variable Y if

$$\lim_{n \to \infty} \Pr(|Y_n - Y| \ge \epsilon) = 0 \quad \forall \epsilon > 0$$
 (0.0.1)

Similarly, a sequence of random variables $Y_1, Y_2, Y_3...$ converges, in mean square, to a random variable Y if

$$\lim_{n \to \infty} E(|Y_n - Y|^2) = 0 \tag{0.0.2}$$

A random variable converges, in probability, to a value if it converges, in mean square, to the same particular value by Markov's Inequality.

$$Y_n \xrightarrow{\mu_s} c \Rightarrow Y_n \xrightarrow{p} c$$
 (0.0.3)

Proof for (0.0.3): For any $\epsilon > 0$

$$\Pr(|Y_n - Y| \ge \epsilon) = \Pr(|Y_n - Y|^2 \ge \epsilon^2)$$
 (0.0.4)

$$\Pr(|Y_n - Y| \ge \epsilon) \le \frac{E|Y_n - Y|^2}{\epsilon^2}$$
 (by Markov's Inequality) (0.0.5)

$$\lim_{n \to \infty} E(|Y_n - Y|^2) = 0 \tag{0.0.6}$$

$$0 \le \lim_{n \to \infty} \Pr(|Y_n - Y| \ge \epsilon) \le \frac{0}{\epsilon^2}$$
 (0.0.7)

$$\lim_{n \to \infty} \Pr(|Y_n - Y| \ge \epsilon) = 0 \quad \forall \epsilon > 0$$
 (0.0.8)

Given in the question that $\{X_j\}$ is a sequence of random variables with

$$\Pr(X_j = 1) = \frac{1}{4} \tag{0.0.9}$$

$$Pr(X_i = 0) + Pr(X_i = 1) = 1$$
 (0.0.10)

$$\Pr(X_j = 0) = 1 - \frac{1}{4} = \frac{3}{4}$$
 (0.0.11)

$$X_j \in \{0, 1\} \tag{0.0.12}$$

Since $0^2 = 0$ and $1^2 = 1$,

$$X_j^2 = X_j \quad \forall j \in \{1, 2, \dots, n\}$$
 (0.0.13)

Thus,

$$Y_n = \frac{1}{n} \sum_{j=1}^n X_j^2 \tag{0.0.14}$$

$$=\frac{1}{n}\sum_{i=1}^{n}X_{i}$$
 (0.0.15)

$$\Pr(Y_n = y) = {}^{n}C_{ny} \left(\frac{1}{4}\right)^{ny} \left(\frac{3}{4}\right)^{n-ny}$$
 (0.0.16)

Let us assume

$$k = ny \tag{0.0.17}$$

$$k \in \{0, 1, 2, \dots, n - 1, n\}$$
 (0.0.18)

$$\Pr(Y_n = \frac{k}{n}) = {}^{n}C_k \left(\frac{1}{4}\right)^k \left(\frac{3}{4}\right)^{n-k}$$
 (0.0.19)

$$E\left(|Y_n - \frac{1}{4}|^2\right) = E\left(Y_n^2 - \frac{1}{2}Y_n + \frac{1}{16}\right) \tag{0.0.20}$$

$$= E(Y_n^2) - \frac{1}{2}E(Y_n) + \frac{1}{16} \qquad (0.0.21)$$

$$E(Y_n^2) = \sum_{k=0}^n \left(\frac{k}{n}\right)^2 \Pr\left(Y_n = \frac{k}{n}\right) \tag{0.0.22}$$

$$= \sum_{k=0}^{n} \left(\frac{k^2}{n^2}\right)^n C_k \left(\frac{1}{4}\right)^k \left(\frac{3}{4}\right)^{n-k}$$
 (0.0.23)

$$E(Y_n^2) = 0 + \frac{1}{n^2} \times n \left(\frac{1}{4}\right)^1 \left(\frac{3}{4}\right)^{n-1} + \sum_{k=2}^n \left(\frac{k}{n}\right)^2 \times \frac{n(n-1)}{k(k-1)} \times {n-2 \choose k-2} \left(\frac{1}{4}\right)^k \left(\frac{3}{4}\right)^{n-k}$$
 (0.0.24)

$$E(Y_n^2) = \frac{1}{4n} \left(\frac{3}{4}\right)^{n-1} + \frac{n-1}{n}$$

$$\times \sum_{k=2}^n \left(\frac{k}{k-1}\right)^{n-2} C_{k-2} \left(\frac{1}{4}\right)^k \left(\frac{3}{4}\right)^{n-k} \quad (0.0.25)$$

$$E(Y_n^2) = \frac{1}{4n} \left(\frac{3}{4}\right)^{n-1} + \frac{n-1}{n} \left(\sum_{k=2}^n {}^{n-2}C_{k-2} \left(\frac{1}{4}\right)^k \left(\frac{3}{4}\right)^{n-k}\right) + \frac{n-1}{n} \left(\sum_{k=2}^n \frac{1}{k-1} {}^{n-2}C_{k-2} \left(\frac{1}{4}\right)^k \left(\frac{3}{4}\right)^{n-k}\right)$$
(0.0.26)

$$E(Y_n^2) = \frac{1}{4n} \left(\frac{3}{4}\right)^{n-1} + \frac{n-1}{n}$$

$$\times \frac{1}{16} \left(\sum_{k=2}^n {n-2 \choose k-2} \left(\frac{1}{4}\right)^{k-2} \left(\frac{3}{4}\right)^{(n-2)-(k-2)}\right)$$

$$+ \frac{1}{n} \left(\sum_{k=2}^n \frac{n-1}{k-1} {n-2 \choose k-2} \left(\frac{1}{4}\right)^k \left(\frac{3}{4}\right)^{n-k}\right) \quad (0.0.27)$$

$$E(Y_n^2) = \frac{1}{4n} \left(\frac{3}{4}\right)^{n-1} + \frac{n-1}{16n} \left(\sum_{j=0}^{n-2} {n-2 \choose j} \left(\frac{1}{4}\right)^j \left(\frac{3}{4}\right)^{(n-2)-j}\right) + \frac{1}{4n} \left(\sum_{j=0}^{n-1} {n-1 \choose k-1} \left(\frac{1}{4}\right)^{k-1} \left(\frac{3}{4}\right)^{(n-1)-(k-1)}\right)$$
(0.0.28)

$$E(Y_n^2) = \frac{1}{4n} \left(\frac{3}{4}\right)^{n-1} + \frac{n-1}{16n} \left(\frac{1}{4} + \frac{3}{4}\right)^{n-2} + \frac{1}{4n} \left(\sum_{j=1}^{n-1} {n-1 \choose j} \left(\frac{1}{4}\right)^j \left(\frac{3}{4}\right)^{(n-1)-j}\right) \quad (0.0.29)$$

$$E(Y_n^2) = \frac{1}{4n} \left(\frac{3}{4}\right)^{n-1} + \frac{n-1}{16n} + \frac{1}{4n} \left(\left(\frac{1}{4} + \frac{3}{4}\right)^{n-1} - \left(\frac{3}{4}\right)^{n-1} \right) \quad (0.0.30)$$

$$E(Y_n^2) = \frac{1}{4n} \left(\frac{3}{4}\right)^{n-1} + \frac{n-1}{16n} + \frac{1}{4n} - \frac{1}{4n} \left(\frac{3}{4}\right)^{n-1}$$

$$= \frac{1}{16} + \frac{3}{16n}$$
(0.0.32)

$$E(Y_n) = \sum_{k=0}^{n} \frac{k}{n} \Pr\left(Y_n = \frac{k}{n}\right)$$

$$= \sum_{k=0}^{n} \left(\frac{k}{n}\right)^n C_k \left(\frac{1}{4}\right)^k \left(\frac{3}{4}\right)^{n-k}$$

$$= 0 + \sum_{k=1}^{n} \frac{k}{n} \times \frac{n}{k} \times {n-1 \choose k-1} \left(\frac{1}{4}\right)^k \left(\frac{3}{4}\right)^{n-k}$$

$$= \frac{1}{4} \sum_{j=0}^{n-1} {n-1 \choose j} \left(\frac{1}{4}\right)^j \left(\frac{3}{4}\right)^{(n-1)-j}$$

$$= \frac{1}{4} \left(\frac{1}{4} + \frac{3}{4}\right)^{n-1}$$

$$(0.0.35)$$

Using equations (0.0.32) and (0.0.38) in (0.0.21),

$$E\left(|Y_n - \frac{1}{4}|^2\right) = \frac{1}{16} + \frac{3}{16n} - \frac{1}{2} \times \frac{1}{4} + \frac{1}{16} \quad (0.0.39)$$
$$= \frac{3}{16n} \quad (0.0.40)$$

(0.0.38)

$$\lim_{n \to \infty} E\left(|Y_n - \frac{1}{4}|^2\right) = \lim_{n \to \infty} \frac{3}{16n}$$

$$= \frac{3}{16} \lim_{n \to \infty} \frac{1}{n}$$

$$= 0$$
(0.0.41)

Thus, Y_n converges, in mean square, to $\frac{1}{4}$ and hence Y_n converges, in probability, to $\frac{1}{4}$.

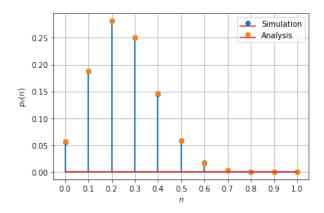


Fig. 0: The PMF distribution of Y_n for n=10

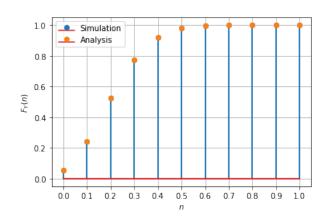


Fig. 0: The CDF distribution of Y_n for n=10