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## AI1103-Assignment-8

Name: Vikhyath Sai Kothamasu, Roll Number: CS20BTECH11056



Download all latex-tikz codes from

https://github.com/Vikhyath-vec/AI1103/blob/main/Assignment-8/Assignment-8.tex

## QUESTION

Suppose  $X_1$  and  $X_2$  are independent and identically distributed random variables each following an exponential distribution with mean  $\theta$ , i.e., the common pdf is given by  $f_{\theta}(x) = \frac{1}{\theta}e^{\frac{-x}{\theta}}, 0 < x < \infty, 0 < \theta < \infty$ . Then which of the following is true? Conditional distribution of  $X_2$  given  $X_1 + X_2 = t$  is

- 1) exponential with mean  $\frac{t}{2}$  and hence  $X_1 + X_2$  is sufficient for  $\theta$
- 2) exponential with mean  $\frac{t\theta}{2}$  and hence  $X_1 + X_2$  is not sufficient for  $\theta$
- 3) uniform(0, t) and hence  $X_1 + X_2$  is sufficient for  $\theta$
- 4) uniform(0,  $t\theta$ ) and hence  $X_1+X_2$  is not sufficient for  $\theta$

## Solution

Let  $f_{X_1,X_2}(x_1,x_2)$  denote the joint probability distribution of random variables  $X_1$  and  $X_2$ . Let Z be

another random variable such that  $Z = X_1 + X_2$ . Also, given in the question,

$$0 < \theta < \infty \tag{0.0.1}$$

$$f_{X_1}(x_1) = \frac{1}{\theta} e^{\frac{-x_1}{\theta}}, 0 < x_1 < \infty$$
 (0.0.2)

$$f_{X_2}(x_2) = \frac{1}{\theta} e^{\frac{-x_2}{\theta}}, 0 < x_2 < \infty$$
 (0.0.3)

Since  $X_1$  and  $X_2$  are independent,

$$f_{X_1,X_2}(x_1,x_2) = f_{X_1}(x_1) \times f_{X_2}(x_2)$$
 (0.0.4)

$$= \frac{1}{\theta} e^{\frac{-x_1}{\theta}} \times \frac{1}{\theta} e^{\frac{-x_2}{\theta}} \tag{0.0.5}$$

$$=\frac{1}{\theta^2}e^{\frac{-(x_1+x_2)}{\theta}}$$
 (0.0.6)

$$f_Z(z) = \int_{-\infty}^{\infty} f_{X_1}(z - x_2) f_{X_2}(x_2) dx_2 \quad (0.0.7)$$

$$= \int_0^z f_{X_1}(z - x_2) f_{X_2}(x_2) dx_2 \quad (0.0.8)$$

$$= \int_0^z f_{X_1,X_2}(z-x_2,x_2) dx_2 \qquad (0.0.9)$$

$$= \int_{0}^{z} \frac{1}{\theta^{2}} e^{\frac{-(z-x_{2}+x_{2})}{\theta}} dx_{2}$$
 (0.0.10)

$$= \frac{1}{\theta^2} \int_0^z e^{\frac{-z}{\theta}} dx_2 \tag{0.0.11}$$

$$=\frac{1}{\theta^2}e^{\frac{-z}{\theta}}z\tag{0.0.12}$$

$$f_{X_2|(X_1+X_2=t)}(x_2) = \begin{cases} \frac{f_{X_1,X_2}(x_1,x_2)}{f_{Z(t)}} & x_2 \in [0,t] \\ 0 & \text{otherwise} \end{cases}$$
(0.0.13)

Let  $x_2 \in [0, t]$ .

$$f_{X_{2}|(X_{1}+X_{2}=t)}(x_{2}) = \frac{f_{X_{1},X_{2}}(x_{1},x_{2})}{f_{Z}(t)}$$

$$= \frac{\frac{1}{\theta^{2}}e^{\frac{-(x_{1}+x_{2})}{\theta}}}{\frac{1}{\theta^{2}}e^{\frac{-t}{\theta}}t}$$

$$= \frac{e^{\frac{-(t)}{\theta}}}{e^{\frac{-t}{\theta}}t}$$

$$(0.0.14)$$

$$(0.0.15)$$

$$=\frac{\frac{1}{\theta^2}e^{\frac{-(x_1+x_2)}{\theta}}}{\frac{1}{\theta^2}e^{\frac{-l}{\theta}}t}$$
(0.0.15)

$$=\frac{e^{\frac{-(t)}{\theta}}}{e^{\frac{-t}{\theta}}t}\tag{0.0.16}$$

$$= \frac{1}{t} \quad \forall x_2 \in [0, t] \tag{0.0.17}$$

The obtained pdf is uniform(0, t). And since the conditional distribution of  $X_2$  does not depend on  $\theta$  for any value of t,  $X_1 + \bar{X}_2$  is sufficient for  $\theta$ . Verifying the pdf,

total probability = 
$$\int_0^t f_{X_2|(X_1+X_2=t)}(x_2) dx_2 \quad (0.0.18)$$

$$= \int_0^t \frac{1}{t} \, dx_2 \tag{0.0.19}$$

$$= 1$$
 (0.0.20)

Hence, the correct answer is option (3)