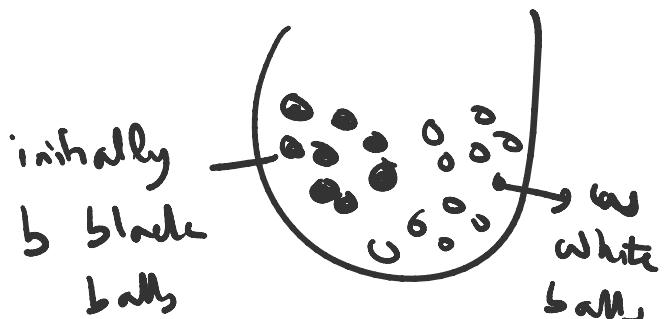


Martingales continued

Example for martingale

Polya's Urn Scheme:



Perform a sequence of random selection, where at each step the chosen ball is replaced by $c_i + 1$ balls of the same color. Let

X_i denote the fraction of black balls in the urn after the i^{th} trial.

Show that the sequence X_0, X_1, \dots , is a martingale.

Clearly $E[X_i] < \infty$ (ii)

Let Y_i be the R.V. that denotes the no. of black balls in the urn after i^{th} trial.

the i th iteration.

$$S_0, Y_0 = b.$$

$$X_i = \frac{Y_i}{b + w + c_i}.$$

X_i is fully determined by Y_0, Y_1, \dots, Y_i

To show:

$$(iii) E[X_{n+1} | Y_0, \dots, Y_n] = X_n \quad \text{for any } n \geq 0$$

$$\text{LHS} = E[X_{n+1} | Y_0, \dots, Y_n]$$

$$= E\left[\frac{Y_{n+1}}{b + w + c(n+1)} \mid Y_0, \dots, Y_n\right]$$

$$= \frac{1}{b + w + c(n+1)} E[Y_{n+1} | Y_0, \dots, Y_n]$$

$$\approx \frac{1}{b + w + c(n+1)} (Y_n + \frac{Y_n}{b + w + cn})$$

done
nr total
 $(n+1) \cdot m$
 $b + w + cn$
 $b + w + cn$

$$= \frac{1}{b + w + c(n+1)} \frac{(b + w + c(n+1)) Y_n}{b + w + cn}$$

$$= Y_n$$

$$= \frac{Y_n}{b+w+c_n}$$

$$= X_n //$$

We have concluded that X_0, X_1, \dots is a martingale w.r.t. Y_0, Y_1, \dots

• _____.

Proposition: If Z_0, Z_1, \dots, Z_n is a martingale w.r.t. X_0, \dots, X_n then Z_0, \dots, Z_n is a martingale w.r.t. itself.

Proof-Outline:

Given: $E[Z_{n+1} | X_0, \dots, X_n] = Z_n$.

To show: $E[Z_{nn} | Z_0, \dots, Z_n] = Z_n$.

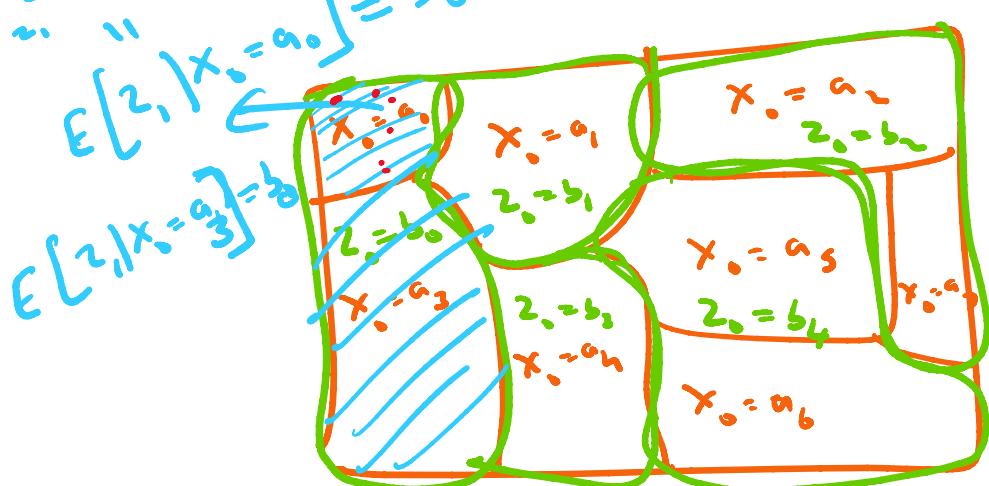
$\sigma_{Z_{nn}} = 0$

$c(x) =$

$$\rightarrow n=0.$$

Given: $E[z_1 | x_0] = z_0$, $f(x_0) = g(z_0)$

To show: $E[z_1 | z_0] = z_0$. z_0 is fully determined by x_0 .



Orange blocks show the partitioning of the sample space given by R.V. X_0 .

Green blocks .. 4

$$u \quad u \quad u \quad Z_0.$$

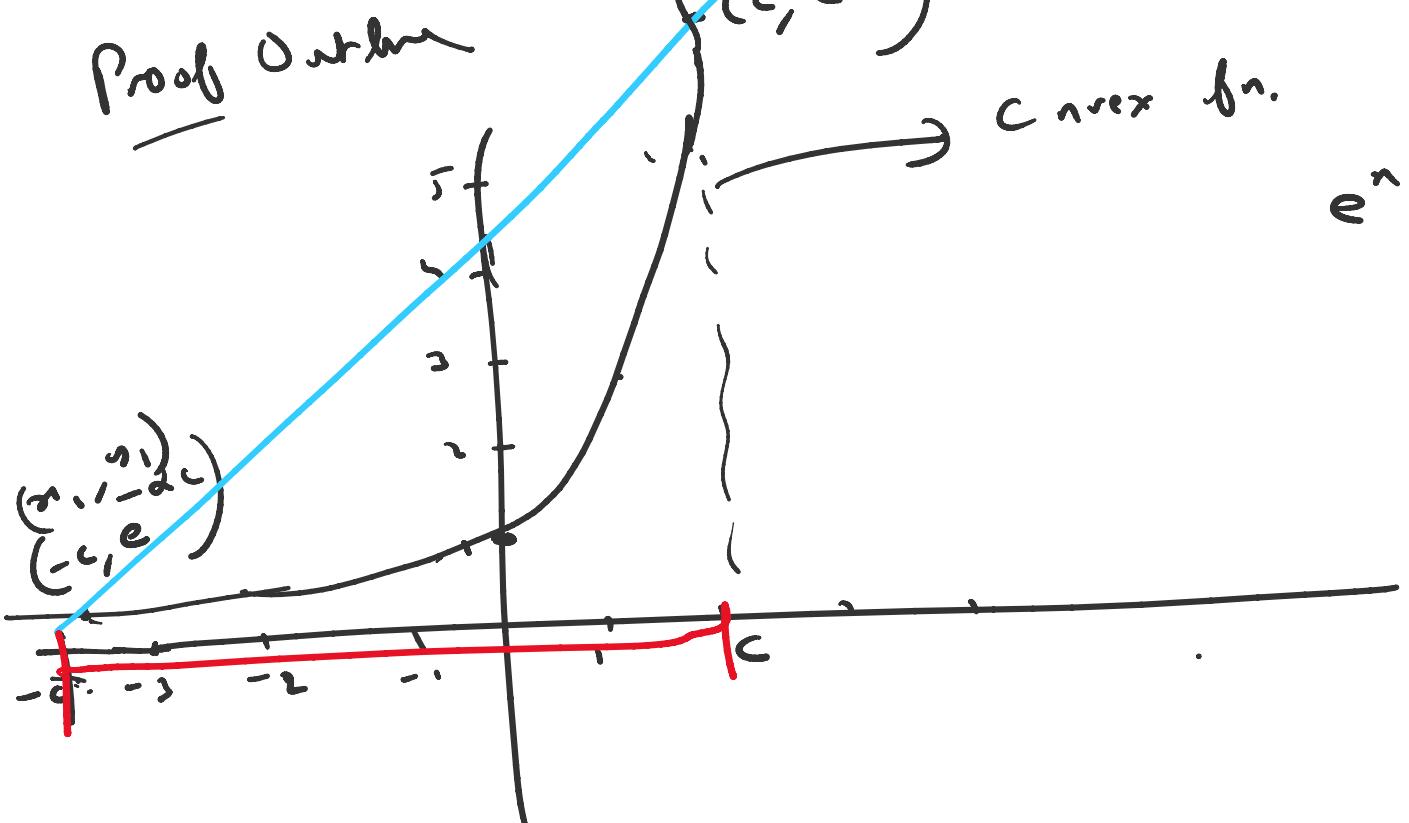
Tail Inequalities for Martingales

↳ Azuma-Hoeffding inequality

Proposition 1 : For any $x \in [-c, +c]$ and
for an arbitrary α ,

$$e^{\alpha x} \leq h(x) , \text{ where}$$

$$h(x) = \frac{e^{cx} + e^{-cx}}{2} + \frac{e^{cx} - e^{-cx}}{2c} x .$$



$$\frac{y - y_1}{x - x_1} = \frac{y_2 - y_1}{x_2 - x_1}$$

$$\frac{y - e^{-cx}}{x - x_1} = \frac{e^{cx} - e^{-cx}}{c}$$

$$\frac{y - e}{x + c} = \frac{e - e}{c + c}$$

$$y = \frac{e^{\alpha c} + e^{-\alpha c}}{2} + \frac{(e^{\alpha c} - e^{-\alpha c})_n}{2c} = h(u)$$

—————
—————

□

Proposition 2. $\cosh(x) := \frac{e^x + e^{-x}}{2} \leq e^{\frac{x^2}{2}}$

→ $1 + \frac{x^2}{2!} + \frac{x^4}{4!} + \frac{x^6}{6!} + \dots$

$1 + \frac{x^2}{2} + \frac{(x^2/2)^2}{2!} + \frac{(x^2/2)^3}{3!} \dots$

$$e^n = 1 + n + \frac{n^2}{2!} + \frac{n^3}{3!} + \frac{n^4}{4!} + \dots$$

□

Theorem [Azuma-Hoeffding Inequality]

Theorem [Azuma-Hoeffding Inequality]

Let z_0, z_1, \dots, z_n be a martingale with respect to itself with

$$|z_i - z_{i-1}| \leq c_i, \quad \forall i \in [n] \xrightarrow{\{1, \dots, n\}}$$

Then, for any $\lambda > 0$ and $t \geq 1$,

$$\Pr(|z_t - z_0| \geq \lambda) \leq \frac{2}{e^{\frac{\lambda^2}{2 \sum c_i^2}}}$$

Proof: Set $y_i = z_i - z_{i-1}, \forall i \in [n]$.

We have,

$$(i) |y_i| \leq c_i,$$

$$(ii) E[y_i | z_0, z_1, \dots, z_{i-1}]$$

$$= E[z_i - z_{i-1} | z_0, z_1, \dots, z_{i-1}]$$

$$= E[z_i | z_0, \dots, z_{i-1}] - E[z_{i-1} | z_0, \dots, z_{i-1}]$$

$$= z_{i-1} - z_{i-1}$$

$$= 0 //$$

For any α , we have

$$E[e^{\alpha Y_i}] | z_0, \dots, z_{i-1}]$$

$$\leq E[h(Y_i)] | z_0, \dots, z_{i-1}] \quad (\text{from Prop 1})$$

where

$$h(Y_i) = \frac{e^{\alpha c_i} + e^{-\alpha c_i}}{2} + \frac{e^{\alpha c_i} - e^{-\alpha c_i}}{2c} Y_i$$

$$\leq \cosh(\alpha c_i) + 0$$

$$\leq e^{\frac{\alpha^2 c_i^2}{2}} // \quad (\text{from Prop 2})$$

Now,

$$z_t - z_0 = (z_t - z_{t-1}) + (z_{t-1} - z_{t-2}) + \dots + (z_1 - z_0)$$

$$= \sum_{k=1}^t Y_k //$$

Moment generating function,

..... moment function,

$m_x(t) = E[e^{tx}]$

$$E[e^{\alpha(z_t - z_0)}] = E[e^{\alpha \sum_{k=1}^t y_k}]$$

$$\begin{aligned}
 &\approx E\left[\prod_{k=1}^t e^{\alpha y_k}\right] \quad y_k = z_k - z_{k-1} \\
 &= E\left[E\left[\prod_{k=1}^t e^{\alpha y_k} | z_0, \dots, z_{t-1}\right]\right] \\
 &= E\left[\prod_{k=1}^{t-1} e^{\frac{\alpha y_k}{2}} E\left[e^{\alpha y_t} | z_0, \dots, z_{t-1}\right]\right] \\
 &\leq E\left[e^{\frac{\alpha c_t^2}{2}} \cdot \prod_{k=1}^{t-1} e^{\alpha y_k}\right] \\
 &= e^{\frac{\alpha c_t^2}{2}} E\left[\prod_{k=1}^{t-1} e^{\alpha y_k}\right] \\
 &\quad \vdots \quad \vdots \quad \vdots \quad \ddots \\
 &\leq e^{\frac{\alpha \sum_{k=1}^t c_k^2}{2}}
 \end{aligned}$$

(repeating
the
same
approach)

Now, for any $\alpha > 0$,

use $\alpha < 0$, when proving (3)

$$\begin{aligned}
 P_x((z_t - z_0) \geq \lambda) &= P_x(e^{\alpha(z_t - z_0)} \geq e^{\alpha \lambda}) \\
 &\quad \rightarrow F(\alpha(z_t - z_0))
 \end{aligned}$$

$$\leq \frac{e^{\alpha(z_t - z_0)}}{e^{\alpha T}} \quad (\text{by Markov Ineq})$$

Substituting with $\alpha = \frac{\gamma}{\sum_{k=1}^{\infty} c_k^2}$ we get

$$\leq e^{-\frac{\gamma(\sum_{k=1}^t c_k^2)}{2}} \quad \boxed{\alpha = \frac{\gamma}{\sum_{k=1}^{\infty} c_k^2}}$$

$$= \frac{1}{e^{\frac{\gamma^2}{2 \sum_{k=1}^t c_k^2}}} \quad \boxed{A}$$

Similarly, we can show that

$$P((z_t - z_0) \leq -\gamma) \leq \frac{1}{e^{\frac{\gamma^2}{2 \sum_{k=1}^t c_k^2}}} \quad \boxed{B}$$

Combining ① and ③, we get

$$\begin{aligned} P\left[|Z_t - z_0| \geq \lambda\right] &= P\left[\left((Z_t - z_0) \geq \lambda\right) \cup \right. \\ &\quad \left.\left((Z_t - z_0) \leq -\lambda\right)\right] \\ &\leq P\left((Z_t - z_0) \geq \lambda\right) + \\ &\quad P\left((Z_t - z_0) \leq -\lambda\right) \\ &\leq ① + ③ \end{aligned}$$

$$= \frac{2}{e^{\frac{\lambda^2}{2 \sum c_i^2}}}$$


Corollary: Let z_0, \dots, z_n be a
martingale such that for all $i \geq 1$,

$$|Z_i - Z_{i-1}| \leq c.$$

Then, for all $t \geq 0$ and $\lambda > 0$,

$$P\left[|Z_t - Z_0| \geq \lambda c \sqrt{t}\right] \leq \frac{2}{e^{\frac{\lambda^2 t}{c^2}}}$$

A more general form is below.

Theorem [Azuma - Hoeffding Thm]

Let Z_0, Z_1, \dots, Z_n be a martingale such that

$$B_k \leq Z_k - Z_{k-1} \leq B_k + d_k,$$

for some constant d_k and for

some R.V.s B_k that may be functions of Z_0, \dots, Z_{k-1} . Then, for all $t > 0$ and any $\lambda > 0$,

... \dots any $\lambda > 0$,

$$P\left(|z_t - z_0| \geq \lambda\right) \leq \frac{2}{e^{\frac{2\lambda^2}{\sum_{k=1}^n d_k^2}}}$$

Putting $B_k = -c_k$ and $d_k = 2c_k$
gives the earlier theorem.

Lipschitz Condition

Let $f: D_1 \times D_2 \times \dots \times D_n \rightarrow \mathbb{R}$ be a function with n arguments. We say f satisfies the Lipschitz condition with bound C if for any $x_1 \in D_1, x_2 \in D_2, \dots, x_n \in D_n$, any $i \in \{1, \dots, n\}$, and any $y_i \in D_i$,

$$\begin{aligned} & |f(x_1, x_2, \dots, \underline{x_{i-1}}, \underline{x_i}, x_{i+1}, \dots, x_n) - \\ & f(x_1, x_2, \dots, \underline{x_{i-1}}, \underline{y_i}, x_{i+1}, \dots, x_n)| \\ & \leq C. \end{aligned}$$

$\leq c$.

Let

$$Z_0 = E[f(x_1, x_2, \dots, x_n)]$$

$$Z_k = E[f(x_1, \dots, x_n) | x_1, x_2, \dots, x_k]$$

$$Z_n = E[f(x_1, \dots, x_n) | x_1, x_2, \dots, x_n] = f(x_1, \dots, x_n).$$

This

Doob martingale.

$$|Z_i - Z_{i-1}| \leq c$$

can one
show this?
 $f: D_1 \times D_2 \times \dots \times D_n \rightarrow \mathbb{R}$

Theorem [McDiarmid's Theorem] Let f be a function that satisfies the Lipschitz condition with bound c . Let x_1, \dots, x_n be independent random variables whose ranges are in D_1, D_2, \dots, D_n . Then,

$$P\left(\left|f(x_1, \dots, x_n) - E[f(x_1, \dots, x_n)]\right| \geq \lambda\right) \leq$$

$$\frac{2}{e^{\frac{2\lambda^2}{nc^2}}}$$

Proof:

Proof:

To prove this, it is enough to show

$$B_{k+1} \leq Z_k - Z_{k+1} \leq B_k + c, \forall k \in [n]$$

We use s_k for x_1, x_2, \dots, x_k .

$$\text{That is, } E[f(\bar{x}) | x_1, \dots, x_k] = E[f(\bar{x}) | s_k]$$

$$f_k(\bar{x}, x) = f(x_1, x_2, \dots, x_{k-1}, x, x_k, \dots, x_n)$$

$$f_k(\bar{z}, n) = f(z_1, z_2, \dots, z_{k-1}, n, z_k, \dots, z_n)$$

We have

$$Z_k - Z_{k+1} = E[f(\bar{x} | s_k)] - E[f(\bar{x}) | s_{k-1}]$$

upper bound

$$\sup_x E[f(\bar{x}) | s_{k-1}, x_k = x] -$$

$$E[f(\bar{x}) | s_{k-1}]$$

lower bound

$$\Rightarrow \inf_y E[f(\bar{x})|S_{k-1}, X_k=y] - E[f(\bar{x})|S_{k-1}]$$

Here,

$$B_k = \inf_y E[f(\bar{x})|S_{k-1}, X_k=y] - E[f(\bar{x})|S_{k-1}]$$

To show:

$$\begin{aligned} & \sup_n E[f(\bar{x})|S_{k-1}, X_k=n] - \inf_y E[f(\bar{x})|S_{k-1}, X_k=y] \\ & \qquad \qquad \qquad \leq c. \\ \text{LHS} &= \sup_{x,y} \left(E[f(\bar{x})|S_{k-1}, X_k=n] - E[f(\bar{x})|S_{k-1}, X_k=y] \right) \\ &= \sup_{x,y} E[f_k(\bar{x}, x) - f_k(\bar{x}, y)]|_{S_{k-1}} \end{aligned}$$

Since, X_1, X_2, \dots, X_n are independent R.V.
for any $x_1, x_2, x_3, \dots, x_{k-1}$,

$$E[f_k(\bar{x}, x) - f_k(\bar{x}, y)]|_{X_1=x_1, \dots, X_{k-1}=x_{k-1}}$$

is equal to

is equal to

$$\sum_{z_{kn} \sim z_n} P_r((x_{kn} = z_{kn}) \cap \dots \cap (x_n = z_n)) \cdot (f_k(\bar{z}, y) - t_k(\bar{z}, y))$$
$$\leq C \sum_{z_{kn} \sim z_n} P_r((z_{kn} - z_{kn}) \cap \dots \cap (z_n = z_n))$$
$$\leq C // \leq C$$

\$\square\$.

Now, the claim follows from the
Azuma-Hoeffding bound above