

## Discrete Distribution Families

### Binomial( $n, p$ ) \*binom

successes in  $n$  Bernoulli trials.

PMF:  $P(X = x) = \binom{n}{x} p^x (1 - p)^{n-x}$

Mean:  $np$ , Var:  $np(1 - p)$

### Geometric( $p$ ) \*geom

Failures before 1st success.

PMF:  $P(X = x) = (1 - p)^x p, x \geq 0$

Mean:  $(1 - p)/p$ , Var:  $(1 - p)/p^2$

### NegBin( $k, p$ ) \*nbinom

Failures before  $k$ -th success.

PMF:  $P(X = x) = \binom{x+k-1}{x} p^k (1 - p)^x$

Mean:  $k(1 - p)/p$ , Var:  $k(1 - p)/p^2$

### Poisson( $\lambda$ ) \*pois

Counts in interval, rate  $\lambda$ .

PMF:  $P(X = x) = \frac{\lambda^x e^{-\lambda}}{x!}$

Mean:  $\lambda$ , Var:  $\lambda$

## Continuous Distribution Families

### Uniform( $a, b$ ) - \*unif

All values equally likely on  $[a, b]$ .

PDF:  $f(x) = \frac{1}{b-a}, a \leq x \leq b$

Mean:  $(a + b)/2$ , Var:  $(b - a)^2/12$

For a uniform distribution between  $a$  and  $b$ ,

$$P(X > x) = \frac{b - x}{b - a}, \quad a \leq x \leq b$$

### Normal( $\mu, \sigma^2$ ) - \*norm

Bell-shaped,  $-\infty < \mu < \infty, \sigma^2 > 0$

PDF:  $f(x) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-(x-\mu)^2/(2\sigma^2)}$

Mean:  $\mu$ , Var:  $\sigma^2$

Z score:  $\mu = x - Z\sigma$

### Lognormal( $\mu, \sigma^2$ ) - \*lnorm

If  $\ln X \sim N(\mu, \sigma^2)$ .

PDF:  $f(x) = \frac{1}{x\sigma\sqrt{2\pi}} e^{-(\ln x - \mu)^2/(2\sigma^2)}$

Mean:  $e^{\mu + \sigma^2/2}$ , Var:  $(e^{\sigma^2} - 1)e^{2\mu + \sigma^2}$

Skewness:  $(e^{\sigma^2} + 2)\sqrt{e^{\sigma^2} - 1}$

### Exponential( $\lambda$ ) - \*exp

Time between Poisson events, Positive RV, wait time, memoryless

( $\lambda$ ) is average rate ( $\beta$ ) is mean wait time

PDF:  $f(x) = \lambda e^{-\lambda x}, x \geq 0$

Mean:  $1/\lambda$ , Var:  $1/\lambda^2$

Skewness: 2

### Beta( $\alpha, \beta$ ) - \*beta

On  $[0, 1]$ , Bayesian priors. Uniform distribution special case

( $\Gamma$ ) function is a generalization of factorial function to non-integer numbers.

PDF:  $f(x) = \frac{1}{B(\alpha, \beta)} x^{\alpha-1} (1 - x)^{\beta-1}$

Mean:  $\alpha/(\alpha + \beta)$ , Var:  $\frac{\alpha\beta}{(\alpha + \beta)^2(\alpha + \beta + 1)}$

Skewness:  $\frac{2(\beta - \alpha)\sqrt{\alpha + \beta + 1}}{(\alpha + \beta + 2)\sqrt{\alpha\beta}}$

### Weibull( $k, \lambda$ ) - \*weibul

Lifetimes, survival, reliability.

Exponential family (when  $k = 1$ ), longer you wait

Survival Analysis

PDF:  $f(x) = \frac{k}{\lambda} (x/\lambda)^{k-1} e^{-(x/\lambda)^k}$

Mean:  $\lambda\Gamma(1 + 1/k)$ , Var:  $\lambda^2[\Gamma(1 + 2/k) - \Gamma(1 + 1/k)^2]$

Skewness:  $\frac{\Gamma(1 + 3/k)\lambda^3 - 3\mu\sigma^2 - \mu^3}{\sigma^3}$

## Continuous RV (numerical integration)

$f_X(x)$  can take on values larger than 1

$f_X(x)$  cannot give probability, area under the graph does

the density of  $f_X(x)$  cannot take negative

values even if total area is 1 because

probability cannot be negative

the support of  $f_X(x)$  must be finite

f <- function(x) 2\*x

EV <- integrate(function(x) x\*f(x), 0, 1)\$value

Define the PDF function

f <- function(x) 2\*x

prob <- integrate(f, lower = 0.5, upper =

0.75)\$value # Integrate over [0.5, 0.75]

## Conditional Distributions

Conditional dist is just a segment of marginal dist, then re-normalized to have an area under the curve equal to 1 If X and Y are independent, in continuous case,

$$f_{Y|X}(y) = f_Y(y)$$

This means conditional PDF of Y and X is marginal PDF of Y.

## Random Sample

It is independent and identically distributed (iid). Each pair of observations are independent, and each observation comes from the same distribution.

## MLE

Way to find estimators. Mean - univariate, linear regression - multivariate

## Steps for MLE

- 1.(discrete or contin) distribution
- 2.Find parameters of a theoretical distribution (eg  $\lambda$  in a Poisson distribution)
- 3.Find sample and parameters for the distribution to find the most likely estimates
- 4.To get estimates, we use the likelihood function of our observed random sample.

## R: Generating Sample and Likelihood

#Generate 30 random samples from Exponential(beta=20)

```
set.seed(123)
```

```
sample_n30 <- tibble(values = rexp(30, rate = 1/20))
```

```
#Compute likelihood and log-likelihood for candidate
```

```
exp_values <- tibble(
```

```
  possible_betas = seq(5, 50, 0.5),
```

```
  likelihood = map_dbl(1 / possible_betas,
```

```
    ~ prod(dexp(sample_n30$values, .))),
```

```
  log_likelihood = map_dbl(1 / possible_betas,
```

```
    ~ log(prod(dexp(sample_n30$values, .))))
```

```
)
```

```
empirical_MLE <- exp_values |>
```

```
  arrange(desc(likelihood)) |>
```

```
  slice(1)
```

```
analytical_MLE <- mean(sample_n30$values)
```

```
#We use the sample mean() function from R!
```

```
round(analytical_MLE, 2)
```

```
LL <- function(l) log(prod(dexp(sample_n30$values,
```

```
  rate = 1 / l)))
```

```
optimize(LL, c(5, 50), maximum = TRUE)
```

stochastic - having some uncertain outcome.

deterministic - an outcome that will be known

with 100% certainty.

Computers cannot actually generate truly

random outcomes, they can generate only

pseudorandom numbers.

Although this sequence is deterministic,

it behaves like a random sample.

One pitfall is that neighbouring pairs are not

independent of each other.

The empirical approach (using observed data),

resulting in approx. values that improve

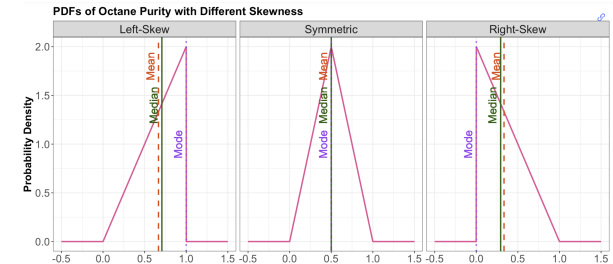
as the sample size incre (frequentist paradigm)

The Law of Large of no. states that, as we

increase our sample size  $n$ , our empirical mean

converges to the true mean we want to estimate.

## Skewness



Left-skewed: long tail to the left, mean < median < mode

Right-skewed: long tail to the right, mean > median > mode