

Today's Agenda

Last time:

- Central Limit Theorem
- Moment generating functions

... we finished the course material!

Today (Lec 36, 07/26):

- Practice
- Course evaluation

Suppose $X \sim Bernoulli(p)$ for some $p \in (0,1)$. Compute the mgf of X

and use it to compute
$$E(X)$$
 and $Var(X)$.

$$M(+) = E(e^{t \times x}) = e^{t \cdot 0} P(X=0) + e^{t \cdot 1} P(X=1)$$

$$= 1 - p + p e^{t} \text{ for } t \in \mathbb{R}$$

$$M(+) = E(e^{tx}) = e^{t \cdot t} P(X=0) + e^{t \cdot t} P(X=1)$$

$$= Lp + pe^{t} \text{ for } teR$$

$$E(X) = M'(0) = (p \cdot e^{t})|_{t=0} = P$$

$$E(X^{2}) = R''(0) = (p \cdot e^{t})|_{t=0} = P$$

$$= Var(X) = E(X^{2}) - E(X)^{2} = P - P^{2} = P(1-P)$$

Let X, Y be arbitrary random variables, which of the following is always correct?

- If X and Y are independent, they are uncorrelated. ✓
- If X and Y are uncorrelated, they are independent. χ
- If X and Y are dependent, they are correlated. X

In a laboratory, the temperature of an object is measured **independently** with two thermometers A and B. Let X and Y show the temperatures of the object measured by thermometers A and B, respectively. We decide to combine the two measurements as a weighted average of X and Y to report the temperature of the object, i.e. the reported temperature T is $T = \alpha X + (1 - \alpha) Y$ where $\alpha \in (0, 1)$. For what value of α is the variance of the random variable T minimized?

We compute Var(T) and minimize it wit α . $Var(T) = Var(\alpha X + (1-\alpha)Y)$ $= d^{2} \left(\operatorname{Wer}(X) + (1-a)^{2} \operatorname{Var}(Y) + 2 \alpha (1-a) \times \operatorname{Cov}(X,Y) \right)$ $= d^{2} \left(\operatorname{Wer}(X) + \operatorname{Var}(Y) \right) - 2 \alpha \operatorname{Var}(Y) + \operatorname{Var}(Y)$ $= d^{2} \left(\operatorname{Ver}(X) + \operatorname{Var}(Y) \right) - 2 \alpha \operatorname{Var}(Y) + \operatorname{Var}(Y)$

$$= \alpha^{2} \left(Var(x) + Var(y) \right) - 2 \omega Var(y) + Var(y)$$
We set the 1st derivative zero and solve for α :

 $\frac{\partial}{\partial x} Var(T) = 2x \left(Var(x) + Var(Y) \right) - 2 Var(Y) \stackrel{!}{=} 0$

We check the 2nd derivative to make sure it's a minimum , $\frac{\partial^2}{\partial x^2}$ $V_{\alpha i}(T) = 2(V_{\alpha i}(X) + V_{\alpha i}(Y)) > 0$

 $=) \quad \chi^* = \quad \frac{Var(Y)}{Var(X) + Var(Y)}$

=> $\alpha' = \frac{Vol(\gamma)}{Vol(x) + Vol(\gamma)}$ minimizes $Vol(\tau)$.

Let X be a continuous random variable with pdf

$$f(x) = e^{-x}$$
, $x \ge 0$ 50% quentile te the median of X .

and 0 otherwise. Compute the median of X.

The cdf of X is
$$x$$

$$\frac{1}{+}(x) = \int_{-\infty}^{\infty} f(t) dt = \int_{-\infty}^{\infty} e^{-t} dt = |-e^{-t}|$$
Solve $\mp(x) = 0.5$ to find median.

$$\mp(x) = \frac{1}{2} = \frac{1}{2} = \frac{1}{2} = \frac{\log(2)}{2}$$

It is known that 8% of the world's population has blue eyes. Suppose you would like to estimate this proportion by asking n people wether they have blue eyes. You would use as an estimate X/n where X is the number of people having blue eyes in a sample of n people. How many people do you need to ask so that with probability 95%, the estimate will be within 0.01 of its true value 0.08?

$$\mathbb{P}\left(\left|\frac{X}{N} - 0.08\right| \le 0.01\right) \ge 0.95$$

Suppose that $X_1,...,X_{50}$ are independent Geometric random variables with parameter 0.5. Estimate the probability that $\sum_{i=1}^{50} X_i > 55$.

Note
$$E(X_i) = \frac{1/2}{1/2} = 1$$
, $Var(X_i) = \frac{1/2}{(1/2)^2} = 2$
and the X_i are inclep.
Thus, $\sum_{i=1}^{50} X_i$ approx $N(50, 100)$ by CLT.
So that
$$P(\sum_{i=1}^{50} X_i > 55) \stackrel{\text{cont}}{=} P(\sum_{i=1}^{50} X_i \geq 55.5)$$

$$\approx P(Z \geq \frac{55.5 - 50}{100}) = 1 - P(Z \leq 0.55)$$

$$= 0.29116$$

Let X be a continuous rv with pdf

$$f(x) = \lambda x^{\lambda - 1}, \quad 0 < x < 1,$$

and 0 otherwise.

- a) Compute $E(X^k)$ for $k \in \mathbb{N}$.
- b) Let $Y = -\lambda \log(X)$. Determine the pdf of Y.

a)
$$E(X^{k}) = \int_{-\infty}^{\infty} x^{k} f(x) dx = \int_{X^{k}} x^{k-1} dx$$

$$= \int_{0}^{\infty} x^{k+d-1} dx = \int_{0}^{\infty} \frac{x^{k+d-1}}{x^{k+d}} \int_{0}^{1} = \frac{1}{x^{k+d}}$$
for keN.

b, let Y= - A log(X). We first compute the cof of X: $F_{x}(x) = \int f(t) dt = \int^{x} \lambda t^{k-1} dt = \chi^{\lambda}$ for $x \in [0,1]$ Using this we find the odf of y as $F_{y}(y) = P(yy) = P(-\lambda \log(x) + \gamma) = P(\log(x) - \frac{\gamma}{\lambda})$ $= P(x \ge e^{-\gamma/\lambda}) = |-F_{x}(e^{-\gamma/\lambda}) - e^{-\gamma/\lambda}|$

We find the pdf of
$$\gamma$$
 by taking the derivative:
 $f_{\gamma}(\gamma) = \overline{f_{\gamma}}(\gamma) = e^{-\gamma}, \quad \gamma > 0$

Course Perceptions

Please go to https://perceptions.uwaterloo.ca and fill out the survey.

when it's almost course evaluation time & my teacher starts trying me

