

Mean and Variance of a Linear Combination of Random Variables (Section 9.5)

Results of Expectation:

1. $E(aX + bY) = aE(X) + bE(Y).$

2. For each $i = 1, 2, \dots, n$, let a_i be a constant and $E(X_i) = \mu_i$.

Then $E(\sum_{i=1}^n a_i X_i) = \sum_{i=1}^n a_i \mu_i.$

3. Let X_1, X_2, \dots, X_n be random variables which each have mean μ (i.e. $E(X_i) = \mu \forall i$).

Then, the sample mean $\bar{X} = \frac{\sum_{i=1}^n X_i}{n}$ satisfies

$$E(\bar{X}) = \mu.$$

Proof:

Results for Covariance:

1. $Cov(X, X) = Var(X)$

2. Consider $Cov(aX + bY, cU + dV)$

3. More generally, we have

$$Cov\left(\sum_{i=1}^m a_i X_i, \sum_{j=1}^n b_j Y_j\right) = \sum_{i=1}^m \sum_{j=1}^n a_i b_j Cov(X_i, Y_j).$$

Results for Variance:

1. Variance of a linear combination of 2 random variables:

$$\begin{aligned} & \text{Var}(aX + bY) \\ &= a^2 \text{Var}(X) + b^2 \text{Var}(Y) + 2ab \text{Cov}(X, Y) \end{aligned}$$

Note: X and Y are not independent, so the Cov terms are needed.

2. Variance of a sum of 2 independent random variables:

Let X and Y be **independent** random variables.

Since $Cov(X, Y) = 0$, then

$$Var(X + Y) = Var(X) + Var(Y).$$

In a similar fashion, we also have

$$Var(X - Y) = Var(X) + Var(Y).$$

This second result seems a little counter-intuitive at first, but look more closely:

$$\begin{aligned} Var(X - Y) &= Var(X) + (-1)^2 Var(Y) \\ &= Var(X) + Var(Y) \end{aligned}$$

Also, remember that Var cannot be negative!

3. Variance of a general linear combination:

For each $i = 1, 2, \dots, n$, let a_i be a constant and $Var(X_i) = \sigma_i^2$. Then,

$$\begin{aligned} & Var\left(\sum_{i=1}^n a_i X_i\right) \\ &= \sum_{i=1}^n a_i^2 \sigma_i^2 + 2 \sum_{i=1}^{n-1} \sum_{j=i+1}^n a_i a_j Cov(X_i, X_j) \end{aligned}$$

Again, the r.v.'s are not independent, so the Cov terms need to be included. The next slide will show the case for independent r.v.'s.

4. Variance of a linear combination of independent random variables:

a) If X_1, X_2, \dots, X_n are independent random variables, then $Cov(X_i, X_j) = 0$, so that:

$$Var\left(\sum_{i=1}^n a_i X_i\right) = \sum_{i=1}^n a_i^2 \sigma_i^2$$

b) If X_1, X_2, \dots, X_n are independent random variables and all have the **same** variance σ^2 , then:

$$Var\left(\sum_{i=1}^n a_i X_i\right) = \sigma^2 \sum_{i=1}^n a_i^2$$

Remark: Using the previous results, we now know that if X_1, X_2, \dots, X_n are independent random variables with the **same** mean μ and **same** variance σ^2 , then the sample mean $\bar{X} = \sum_{i=1}^n \frac{X_i}{n}$ has

$$E(\bar{X}) = \mu \text{ and } Var(\bar{X}) = \frac{\sigma^2}{n}.$$

What does this tell us about \bar{X} ?

1. If we consider $Var(\bar{X})$, we notice that it is a **scaled version** of $Var(X_i) = \sigma^2$.

In other words, \bar{X} is less variable as compared to X_i .

Does this make sense?

2. We also note that $Var(\bar{X})$ will become smaller as the sample size, n gets larger. In other words, \bar{X} becomes less variable as n gets larger.

This happens because as n increases (i.e. as we collect more data points in our sample), we are obtaining more information, so our sample average, \bar{X} , is becoming **more precise** in the sense that we have

$$Var(\bar{X}) \rightarrow 0 \text{ as } n \rightarrow \infty.$$

This implies that as $n \rightarrow \infty$, $\bar{X} \rightarrow \mu$.

This is sometimes called the “**law of averages**” or the “**law of large numbers**”.

Note: There is no mention of what the specific distribution of X_1, X_2, \dots, X_n is.

Linear Combinations of Independent Normal Random Variables (Section 9.6)

1. Let $X \sim N(\mu, \sigma^2)$ and $Y = aX + b$, where a and b are constant real numbers. Then,

$$Y \sim N(a\mu + b, a^2\sigma^2).$$

2. Let $X \sim N(\mu_1, \sigma_1^2)$ and $Y \sim N(\mu_2, \sigma_2^2)$ be independent random variables. If a and b are constants, then

$$aX + bY \sim N(a\mu_1 + b\mu_2, a^2\sigma_1^2 + b^2\sigma_2^2).$$

In general, if $X_i \sim N(\mu_i, \sigma_i^2)$, $i = 1, 2, \dots, n$, are independent random variables and a_1, a_2, \dots, a_n are constants, then

$$\sum_{i=1}^n a_i X_i \sim N \left(\sum_{i=1}^n a_i \mu_i, \sum_{i=1}^n a_i^2 \sigma_i^2 \right)$$

3. Let X_1, X_2, \dots, X_n be independent $N(\mu, \sigma^2)$ random variables.

Then:

$$\sum_{i=1}^n X_i \sim N(n\mu, n\sigma^2)$$

and

$$\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i \sim N(\mu, \sigma^2/n)$$

Example: Let $X \sim N(3, 5)$ and $Y \sim N(6, 14)$ be independent random variables.

Find $P(X > Y)$.

You Try:

Let $X \sim N(5,4)$. $Y \sim N(7,9)$. X and Y are independent random variables.

a) Find $P(2X - Y > 4)$.

b) Find the minimum number, n , of independent observations needed on X so that $P(|\bar{X} - 5| < 0.1) \geq 0.98$.

a) We want $P(2X - Y > 4)$.

Let $W = 2X - Y$. $E(W) = 2(5) - 7 = 3$. $\text{Var}(Y) = 4(4) + 9 = 25$. So, $W \sim N(3, 25)$.

So, $P(2X - Y > 4) = P(W > 4) = P(Z > \frac{4-3}{5}) = P(Z > 0.2) = 1 - 0.57926 = 0.42074$

b) $X \sim N(5,4)$. It follows that $\bar{X} \sim N(5, \frac{4}{n})$.

$$P(|\bar{X} - 5| < 0.1) \geq 0.98 \Rightarrow P\left(|Z| < \frac{0.1}{\sqrt{\frac{4}{n}}}\right) \geq 0.98$$

We need to determine the 0.99 quantile of the distribution of Z . From the quantile table, we see that $z = 2.3263$. Remember that 0.98 probability between these values implies that there is 0.01 in the left and right tails.

So, we have $2.3263 = \frac{0.1\sqrt{n}}{2}$. We can solve this for n .

This gives us $n = \left(\frac{2 \cdot 2.3263}{0.1}\right)^2 = 2,164.67$.

Therefore, a minimum of 2,165 independent observations are needed.

Indicator Random Variables (Section 9.7)

An indicator random variable is a binary variable (0 or 1) that indicates whether or not an event has occurred.

It can allow us to take more complicated scenarios and break them into simpler ones.

Example: Suppose that $X \sim \text{Bin}(n, p)$.

For $i = 1, 2, \dots, n$, define a new random variable X_i as follows:

$$X_i = \begin{cases} 0 & \text{if } i^{\text{th}} \text{ trial was a failure,} \\ 1 & \text{if } i^{\text{th}} \text{ trial was a success.} \end{cases}$$

Then:

Example: We have N letters to N different people, and N envelopes addressed to those N people. One letter is randomly put in each envelope.

Find the mean and variance of the number of letters placed in the correct envelope.

The results are fascinating! If we let X represent the number of correct matches, we showed that:

$E(X) = 1$ and $\text{Var}(X) = 1$ regardless of the value of N .

Exercise: Work this out numerically with $N = 3$. Consider the sequence A, B, C to be a perfect match (i.e. letter A is in envelope A, letter B is in envelope B, and letter C is in envelope C). So, the sequence A, C, B would only have 1 match (letter A is in envelope A).

There are 6 possible outcomes. List them, with their probabilities, then set up the pf of X , noting that it's not possible to get exactly 2 matches, as 2 matches guarantees that the third is a match!

You will see that $E(X)$ and $\text{Var}(X) = 1$!