# RSM8512 Assignment - Linear Regression

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### Question 1 [25 marks]

Describe the null hypotheses to which the p-values given in Table 3.4 correspond. Explain what conclusions you can draw based on these p-values. Your explanation should be phrased in terms of sales, TV, radio, and newspaper, rather than in terms of the coefficients of the linear model.

**Answer:** The null hypothesis for 'TV' is that in the presence of radio ads and newspaper ads, TV has no effect on sales. Similarly, the null hypothesis for radio is that in the presence of TV and newspaper, radio has no effect on sales. The similar null hypothesis for newspaper too.

The p-value of TV and radio is <0.0001, which means that we should reject null hypothesis for these two factors. However, the p-value of newspaper is bigger that 0.05, which means we should accept null hypothesis and think that newspaper ads has no effect on sales.

#### Question 3 [25 marks]

Suppose we have a data set with five predictors, X1 = GPA, X2 = IQ, X3 = Level (1 for College and 0 for High School), X4 = Interaction between GPA and IQ, and X5 = Interaction between GPA and Level. The response is starting salary after graduation (in thousands of dollars). Suppose we use least squares to fit the model, and get

$$\hat{\beta}_0 = 50, \quad \hat{\beta}_1 = 20, \quad \hat{\beta}_2 = 0.07, \quad \hat{\beta}_3 = 35, \quad \hat{\beta}_4 = 0.01, \quad \hat{\beta}_5 = -10$$

- (a) Which answer is correct, and why?
- i. For a fixed value of IQ and GPA, high school graduates earn more on average than college graduates.
- ii. For a fixed value of IQ and GPA, college graduates earn more on average than high school graduates.
- iii. For a fixed value of IQ and GPA, high school graduates earn more on average than college graduates provided that the GPA is high enough.
- iv. For a fixed value of IQ and GPA, college graduates earn more on average than high school graduates provided that the GPA is high enough.

**Answer:** We can generate a model based on the info above:

$$Y = 50 + 20 * X1 + 0.07 * X2 + 35 * X3 + 0.01 * X1 * X2 - 10 * X1 * X3$$

$$salary(x_3 = 0) = 50 + 20*X1 + 0.07*X2 + 0.01*X1*X2 - 10*X1*X3$$
 
$$salary(x_3 = 1) = 50 + 20*X1 + 0.07*X2 + 35 + 0.01*X1*X2 - 10*X1*X3$$
 
$$salary = salary(x_3 = 0) - salary(x_3 = 1) = 35$$

Therefore, iii is correct, when GPA is high enough, college graduates earn more on average than high school graduates.

(b) Predict the salary of a high school graduates with IQ of 110 and a GPA of 4.0.

#### Answer:

$$salary = 50 + 20 * 4.0 + 0.07 * 110 + 35 * 1 + 0.01 * 4.0 * 110 - 10 * 4.0 * 1 = 137.1$$

The salary of a high school graduates with IQ of 110 and a GPA of 4.0 is 137.1.

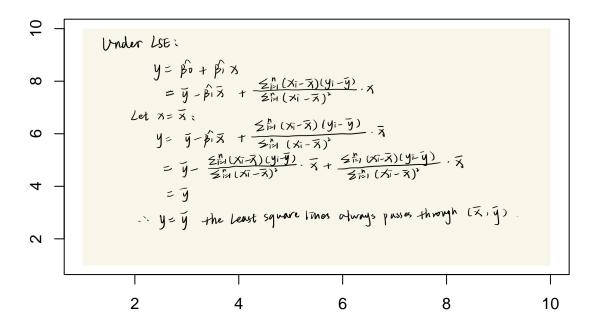
(c) True or false: Since the coefficient for the GPA/IQ interaction term is very small, there is very little evidence of an interaction effect. Justify your answer.

**Answer:** False. We must examine the p-value of the regression coefficient to determine if the interaction term is statistically significant or not.

#### Question 6 [25 marks]

Using (3.4), argue that in the case of simple linear regression, the least squares line always passes through the point  $(\bar{x}, \bar{y})$ .

```
img <- image_read("Linear Regression Q3.png")
plot(1:10, type="n", ann=FALSE)
rasterImage(as.raster(img), 1, 1, 10, 10)</pre>
```



# Question 8 [25 marks]

##

This question involves the use of simple linear regression on the Auto data set.

(a) Use the lm() function to perform a simple linear regression with mpg as the response and horsepower as the predictor. Use the summary() function to print the results. Comment on the output. For example:

```
model1 = lm(mpg ~ horsepower, data = Auto)
summary(model1)
##
## Call:
   lm(formula = mpg ~ horsepower, data = Auto)
##
## Residuals:
##
        Min
                   1Q
                        Median
                                     3Q
                                              Max
## -13.5710 -3.2592
                      -0.3435
                                 2.7630
                                         16.9240
```

```
## Coefficients:
##
                Estimate Std. Error t value Pr(>|t|)
## (Intercept) 39.935861
                           0.717499
                                      55.66
                                              <2e-16 ***
## horsepower
              -0.157845
                           0.006446
                                    -24.49
                                              <2e-16 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 4.906 on 390 degrees of freedom
## Multiple R-squared: 0.6059, Adjusted R-squared: 0.6049
## F-statistic: 599.7 on 1 and 390 DF, p-value: < 2.2e-16
```

#### #cor(Auto\$mpg, Auto\$horsepower)

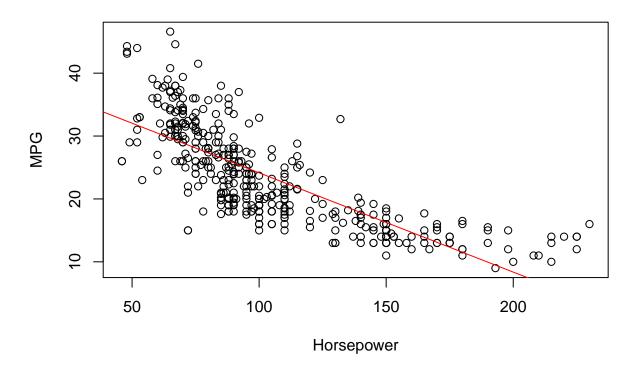
- i. Is there a relationship between the predictor and the response? **Answer:** Yes, there is relationship between the predictor and the response.
- ii. How strong is the relationship between the predictor and the response? **Answer:** Not too strong for the  $R^2$  is 0.6059, and just under two-thirds of the variability in mpg is explained by a linear regression on horsepower.
- iii. Is the relationship between the predictor and the response positive or negative? Answer: Negative
- iv. What is the predicted mpg associated with a horsepower of 98? What are the associated 95% confidence and prediction intervals? **Answer:** According to model 1, we have

$$mpg = 39.935861 - 0.157845 * 98 = 24.46705$$

(b) Plot the response and the predictor. Use the abline() function to display the least squares regression line.

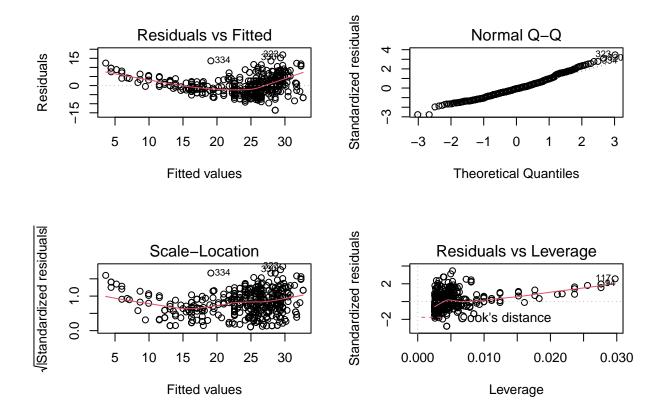
```
plot(x = Auto$horsepower,y = Auto$mpg, main="Scatterplot with Regression Line", xlab="Horsepower
abline(model1, col = 'red')
```

# **Scatterplot with Regression Line**



(c) Use the plot() function to produce diagnostic plots of the least squares regression.

```
par(mfrow=c(2,2)) # To display 4 plots in a 2x2 grid
plot(model1)
```



# Question 11 [25 marks]

In this problem we will investigate the t-statistic for the null hypothesis  $H_0:\beta=0$  in simple linear regression without an intercept. To begin, we generate a predictor x and a response y as follows.

```
set.seed (1)
x=rnorm (100)
y=2*x+rnorm (100)
```

(a) Perform a simple linear regression of y onto x, without an intercept. Report the coefficient estimate  $\hat{\beta}$ , the standard error of this coefficient estimate, and the t-statistic and p-value associated with the null hypothesis  $H_0: \beta=0$ . Comment on these results. (You can perform regression without an intercept using the command  $lm(y\sim x+0)$ .

```
lm1 = lm(y~x+0)
#summary_fit = summary(lm1)
#coef(lm1)
#summary_fit$coefficients["x", "Std. Error"]
```

```
#summary_fit$coefficients["x", "t value"]
#summary_fit$coefficients["x", "Pr(>/t/)"]
```

**Answer:** The coefficient estimate  $\hat{\beta}$  is 1.9939, the standard error of this coefficient estimate is 0.1065, and the t-statistic and p-value associated with the null hypothesis  $H_0:\beta=0$  are 18.73 and 2.642197e-34 separately. The p-value of the t-statistic is near to zero so we need to reject the null hypothesis.

(b) Now perform a simple linear regression of x onto y without an intercept, and report the coefficient estimate, its standard error, and the corresponding t-statistic and p-values associated with the null hypothesis  $H_0: \beta=0$ . Comment on these results.

```
lm2 = lm(x~y+0)
summary_fit2 = summary(lm2)
#summary_fit$coefficients["x", "Pr(>/t/)"]
```

**Answer:** The coefficient estimate is 0.3911, the standard error is 0.02089, and the corresponding t-statistic and p-values are 18.73 and 2.642197e-34 separately. The p-value of the t-statistic is near to zero so we need to reject the null hypothesis.

- (c) What is the relationship between the results obtained in (a) and (b)? **Answer:** Both results in (a) and (b) reflect the same line created in 11a. In other words,  $y = 2x + \epsilon$  could also be written  $x=0.5(y-\epsilon)$ .
- (d) For the regression of Y onto X without an intercept, the t statistic for  $H_0: \beta = 0$  takes the form  $\frac{\hat{\beta}}{SE(\hat{\beta})}$ , where  $\hat{\beta}$  is given by (3.38), and where

$$\mathrm{SE}(\hat{\beta}) = \sqrt{\frac{\sum_{i=1}^{n} (y_i - x_i \hat{\beta})^2}{(n-1)\sum_{i'=1}^{n} x_{i'}^2}}$$

(These formulas are slightly different from those given in Sections 3.1.1 and 3.1.2, since here we are performing regression without an intercept.) Show algebraically, and confirm numerically in R, that the t-statistic can be written as

$$t = \frac{(\sqrt{n-1}) \sum_{i=1}^{n} x_i y_i}{\sqrt{(\sum_{i=1}^{n} x_i^2)(\sum_{i'=1}^{n} y_{i'}^2) - (\sum_{i'=1}^{n} x_{i'} y_{i'})^2}}$$

Answer: Based on the two formula above, we have

$$t = \frac{\beta}{\mathrm{SE}(\beta)} = \frac{\sqrt{n-1}\sum x_i y_i}{\sqrt{\sum x_i^2 \sum y_i^2 - (\sum x_i y_i)^2}}$$

```
(\operatorname{sqrt}(\operatorname{length}(x)-1) * \operatorname{sum}(x*y)) / (\operatorname{sqrt}(\operatorname{sum}(x*x) * \operatorname{sum}(y*y) - (\operatorname{sum}(x*y))^2))
```

The t-statistic is 18.72593.

## [1] 18.72593

(e) Using the results from (d), argue that the t-statistic for the regression of y onto x is the same as the t-statistic for the regression of x onto y.

```
(\operatorname{sqrt}(\operatorname{length}(y)-1)*\operatorname{sum}(y*x))/(\operatorname{sqrt}(\operatorname{sum}(y*y)*\operatorname{sum}(x*x)-(\operatorname{sum}(y*x)^2)))
```

```
## [1] 18.72593
```

The t-statistics for x onto y is 18.7253.

(f) In R, show that when regression is performed with an intercept, the t-statistic for  $H_0: \beta_1 = 0$  is the same for the regression of y onto x as it is for the regression of x onto y.

```
lm3 = lm(y~x)
lm4 = lm(x~y)
summary(lm3)
```

```
##
## Call:
## lm(formula = y \sim x)
## Residuals:
##
      Min
               10 Median
                                3Q
                                       Max
## -1.8768 -0.6138 -0.1395 0.5394 2.3462
##
## Coefficients:
              Estimate Std. Error t value Pr(>|t|)
##
## (Intercept) -0.03769
                           0.09699 -0.389
                                              0.698
## x
               1.99894
                           0.10773 18.556
                                             <2e-16 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.9628 on 98 degrees of freedom
## Multiple R-squared: 0.7784, Adjusted R-squared: 0.7762
## F-statistic: 344.3 on 1 and 98 DF, p-value: < 2.2e-16
```

#### summary(1m4)

```
##
## Call:
## lm(formula = x \sim y)
##
## Residuals:
##
        Min
                  1Q
                       Median
                                     3Q
                                             Max
## -0.90848 -0.28101 0.06274 0.24570 0.85736
##
## Coefficients:
##
               Estimate Std. Error t value Pr(>|t|)
## (Intercept) 0.03880
                           0.04266
                                      0.91
                                               0.365
## y
                0.38942
                           0.02099
                                      18.56
                                              <2e-16 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.4249 on 98 degrees of freedom
## Multiple R-squared: 0.7784, Adjusted R-squared: 0.7762
## F-statistic: 344.3 on 1 and 98 DF, p-value: < 2.2e-16
```

**Answer:** Based on the result above, we can see that the t-statistic value is the same for the two linear regression.

#### Question 14 [20 marks]

This problem focuses on the collinearity problem.

(a) Perform the following commands in R:

```
set.seed (1)
x1=runif (100)
x2 =0.5* x1+rnorm (100) /10
y=2+2* x1 +0.3* x2+rnorm (100)
```

The last line corresponds to creating a linear model in which y is a function of x1 and x2. Write out the form of the linear model. What are the regression coefficients?

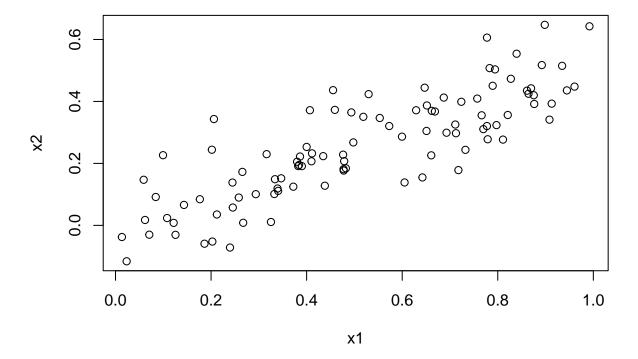
**Answer:** The form of the modle is multiple linear regression model. The regression coefficients are:  $\beta_0=2$ ,  $\beta_1=2,\,\beta_2=0.3$ 

(b) What is the correlation between x1 and x2? Create a scatterplot displaying the relationship between the variables.

cor(x1,x2)

## [1] 0.8351212

plot(x1,x2)



(c) Using this data, fit a least squares regression to predict y using x1 and x2. Describe the results obtained. What are  $\hat{\beta}_0$ ,  $\hat{\beta}_1$ , and  $\hat{\beta}_2$ ? How do these relate to the true  $\beta_0$ ,  $\beta_1$ , and  $\beta_2$ ? Can you reject the null hypothesis  $H_0:\beta_1=0$ ? How about the null hypothesis  $H_0:\beta_2=0$ ?

```
lm.fit = lm(y~x1+x2)
summary(lm.fit)
```

##

## Call:

```
## lm(formula = y \sim x1 + x2)
##
## Residuals:
##
       Min
                1Q Median
                                 3Q
                                        Max
## -2.8311 -0.7273 -0.0537 0.6338
                                     2.3359
##
## Coefficients:
               Estimate Std. Error t value Pr(>|t|)
##
## (Intercept)
                 2.1305
                             0.2319
                                      9.188 7.61e-15 ***
                 1.4396
                             0.7212
                                               0.0487 *
## x1
                                      1.996
## x2
                 1.0097
                             1.1337
                                      0.891
                                               0.3754
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 1.056 on 97 degrees of freedom
## Multiple R-squared: 0.2088, Adjusted R-squared: 0.1925
## F-statistic: 12.8 on 2 and 97 DF, p-value: 1.164e-05
 (d) Now fit a least squares regression to predict y using only x1. Comment on your results. Can you reject
     the null hypothesis H_0: \beta_1 = 0?
lm.fit1 = lm(y~x1)
summary(lm.fit1)
##
## Call:
## lm(formula = y \sim x1)
##
## Residuals:
                                      3Q
##
        Min
                  1Q
                        Median
                                              Max
## -2.89495 -0.66874 -0.07785 0.59221 2.45560
##
## Coefficients:
               Estimate Std. Error t value Pr(>|t|)
                                      9.155 8.27e-15 ***
## (Intercept)
                 2.1124
                             0.2307
## x1
                 1.9759
                             0.3963
                                     4.986 2.66e-06 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
```

```
##
## Residual standard error: 1.055 on 98 degrees of freedom
## Multiple R-squared: 0.2024, Adjusted R-squared: 0.1942
## F-statistic: 24.86 on 1 and 98 DF, p-value: 2.661e-06
```

lm.fit2 = lm(y~x2)

## ---

##

**Answer:** the p-value of t-statistic is 2.661e-06, which is less than 0.05, so we can reject the null hypothesis.

(e) Now fit a least squares regression to predict y using only x2. Comment on your results. Can you reject the null hypothesis  $H_0: \beta_1 = 0$ ?

```
summary(lm.fit2)
##
## Call:
## lm(formula = y \sim x2)
## Residuals:
##
        Min
                   10
                        Median
                                              Max
                                      3Q
## -2.62687 -0.75156 -0.03598 0.72383 2.44890
##
## Coefficients:
               Estimate Std. Error t value Pr(>|t|)
##
  (Intercept)
                 2.3899
                             0.1949
                                       12.26 < 2e-16 ***
                             0.6330
                                        4.58 1.37e-05 ***
## x2
                 2.8996
```

**Answer:** the p-value of t-statistic is 1.366e-05, which is less than 0.05, so we can reject the null hypothesis.

(f) Do the results obtained in (c)–(e) contradict each other? Explain your answer.

## Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

## Residual standard error: 1.072 on 98 degrees of freedom
## Multiple R-squared: 0.1763, Adjusted R-squared: 0.1679
## F-statistic: 20.98 on 1 and 98 DF, p-value: 1.366e-05

**Answer:** not contradict with each other because x1 and x2 have collinearity, it is hard to distinguish their effects when regressed upon together. When they are regressed upon separately, the linear relationship between y and each predictor is indicated more clearly.

(g) Now suppose we obtain one additional observation, which was unfortunately mismeasured.

```
x1=c(x1, 0.1)

x2=c(x2, 0.8)

y=c(y,6)
```

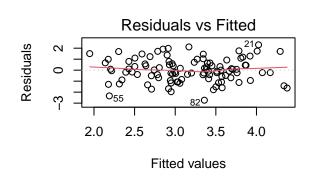
Re-fit the linear models from (c) to (e) using this new data. What effect does this new observation have on the each of the models? In each model, is this observation an outlier? A high-leverage point? Both? Explain your answers.

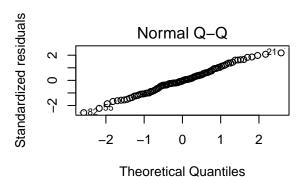
```
model_c = lm(y~x1+x2)
summary(model_c)
```

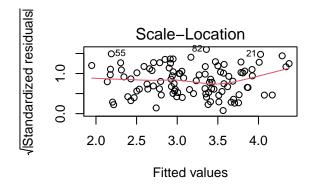
```
##
## Call:
## lm(formula = y \sim x1 + x2)
##
## Residuals:
##
        Min
                  1Q
                       Median
                                    3Q
                                             Max
## -2.73348 -0.69318 -0.05263 0.66385 2.30619
##
## Coefficients:
##
               Estimate Std. Error t value Pr(>|t|)
                 2.2267
                            0.2314
                                     9.624 7.91e-16 ***
## (Intercept)
## x1
                 0.5394
                            0.5922
                                     0.911 0.36458
                                     2.801 0.00614 **
## x2
                 2.5146
                            0.8977
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 1.075 on 98 degrees of freedom
## Multiple R-squared: 0.2188, Adjusted R-squared: 0.2029
## F-statistic: 13.72 on 2 and 98 DF, p-value: 5.564e-06
```

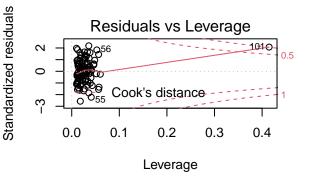
In model1, x1 is not significant but x2 is significant.

```
par(mfrow=c(2,2))
plot(model_c)
```







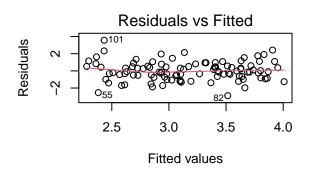


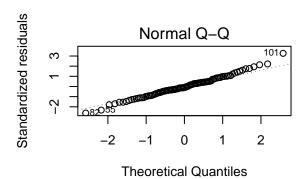
```
model_d = lm(y~x1)
summary(model_d)
```

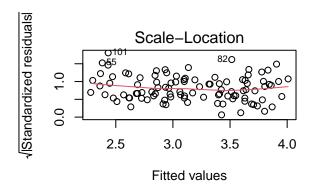
```
##
## Call:
## lm(formula = y \sim x1)
##
## Residuals:
##
       Min
                 1Q
                    Median
                                  3Q
                                         Max
##
   -2.8897 -0.6556 -0.0909 0.5682
                                      3.5665
##
## Coefficients:
##
               Estimate Std. Error t value Pr(>|t|)
  (Intercept)
                  2.2569
                             0.2390
                                       9.445 1.78e-15 ***
                                       4.282 4.29e-05 ***
## x1
                  1.7657
                             0.4124
## ---
                    0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
## Signif. codes:
##
```

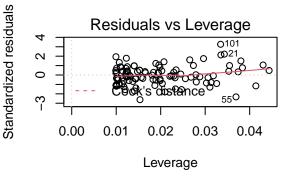
```
## Residual standard error: 1.111 on 99 degrees of freedom
## Multiple R-squared: 0.1562, Adjusted R-squared: 0.1477
## F-statistic: 18.33 on 1 and 99 DF, p-value: 4.295e-05
```

```
par(mfrow=c(2,2))
plot(model_d)
```







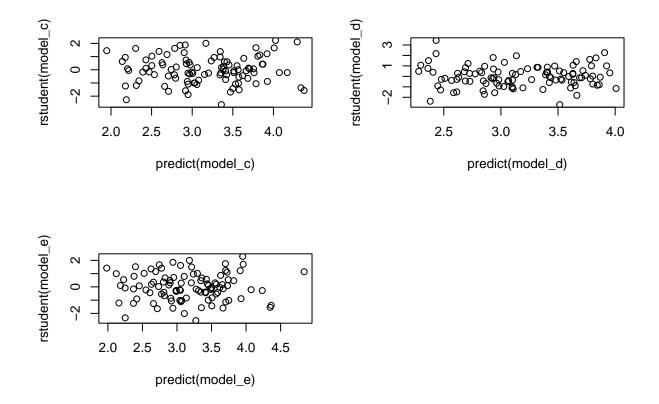


```
model_e = lm(y~x2)
summary(model_e)
```

```
##
## Call:
## lm(formula = y ~ x2)
##
## Residuals:
## Min 1Q Median 3Q Max
## -2.64729 -0.71021 -0.06899 0.72699 2.38074
##
```

```
## Coefficients:
                  Estimate Std. Error t value Pr(>|t|)
##
   (Intercept)
                     2.3451
                                  0.1912
                                            12.264 < 2e-16 ***
                                              5.164 1.25e-06 ***
## x2
                     3.1190
                                  0.6040
##
                       0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
## Signif. codes:
##
## Residual standard error: 1.074 on 99 degrees of freedom
## Multiple R-squared: 0.2122, Adjusted R-squared: 0.2042
## F-statistic: 26.66 on 1 and 99 DF, p-value: 1.253e-06
par(mfrow=c(2,2))
plot(model_e)
                                                   Standardized residuals
                                                                       Normal Q-Q
                 Residuals vs Fitted
                                                                                      1000000 O<sup>210</sup>
Residuals
                                                        ^{\circ}
      0
                                                        0
                                                        7
      က
                                                                                          2
                2.5
                           3.5
                                 4.0
                                      4.5
                                                                              0
                      3.0
                                                                 -2
                      Fitted values
                                                                    Theoretical Quantiles
/|Standardized residuals
                                                   Standardized residuals
                   Scale-Location
                                                                 Residuals vs Leverage
                                                                                           1010
                                                                            o
distance
      0.0
                2.5
                                                                  0.02 0.04 0.06
          2.0
                      3.0
                           3.5
                                 4.0
                                      4.5
                                                            0.00
                                                                                     0.08 0.10
                      Fitted values
                                                                          Leverage
```

```
par(mfrow=c(2,2))
plot(predict(model_c), rstudent(model_c))
plot(predict(model_d), rstudent(model_d))
plot(predict(model_e), rstudent(model_e))
```



By checking the residual plots above, we find only model  $c(y\sim x1+x2)$  has outlier observation.

BONUS [10 marks]: If you were to conduct the simulation in a) many times (with different draws for the random variables) so that you had a large number of datasets, what you would expect the average values (over all simulations) for the coefficients from the regression in c) to be? Explain your answer.