# 20) Nested and Ordered Logit

Vitor Kamada

July 2018

## Independence of Irrelevant Alternatives (IIA)

## Multinomial Logit and Conditional Logit

- any two pairs of alternatives is a binary logit model
- ullet  $\epsilon_{ij}$  are iid (independent and identically distributed)

## Car Blue Bus Red Red

Blue Bus Error  $(\epsilon_{i2})$  are very correlated Red Red Error  $(\epsilon_{i3})$ 

Vitor Kamada ECO 7110 Econometrics II July 2018

## **Limb (level 1):** j = 1, 2

**Branch (level 2):** k = 1, 2, 3, 4

## **Generalized Extreme Value (GEV)**

$$p_{jk} = p_j \times p_{k|j}$$

$$F(\epsilon) = exp[-G(e^{-\epsilon_{11}},...,e^{-\epsilon_{1k_1}};...e^{-\epsilon_{j1}},...,e^{-\epsilon_{jk_j}})]$$

$$G(Y) = \sum_{j=1}^{J} \left(\sum_{k=1}^{K_j} Y_{jk}^{\frac{1}{\tau_j}}\right)^{\tau_j}$$

Scale Parameters:  $\tau_j = \sqrt{1 - Cor[\epsilon_{jk}, \epsilon_{jl}]}$ 

 $au_j = 1 o \mathsf{Multinomial\ Logit}$ 

## **Nested Logit**

$$p_{jk} = p_j \times p_{k|j}$$

$$\frac{\exp(z_j'\alpha + \tau_j I_j)}{\sum\limits_{m=1}^{J} \exp(z_m'\alpha + \tau_m I_m)} \times \frac{\exp(x_{jk}'\beta_j/\tau_j)}{\sum\limits_{l=1}^{K_j} \exp(x_{jl}'\beta_j/\tau_j)}$$

$$I_j = In\{\sum_{l=1}^{K_j} exp(x'_{jl}\beta_j/ au_j)\}$$

Vitor Kamada ECO 7110 Econometrics II July 2018 5/17

## Full Information Maximum Likelihood (FIML)

$$f(y_i) = \prod_{j=1}^{J} \prod_{k=1}^{K_j} [p_{ik|j} \times p_{ij}]^{y_{ijk}}$$

$$f(y_i) = \prod_{j=1}^{J} (p_{ij}^{y_{ij}} \prod_{k=1}^{K_j} p_{ik|j}^{y_{ijk}})$$

$$InL = \sum_{i=1}^{N} \sum_{j=1}^{J} y_{ij} Inp_{ij} + \sum_{i=1}^{N} \sum_{j=1}^{J} \sum_{k=1}^{K_j} y_{ijk} Inp_{ik|j}$$

 Vitor Kamada
 ECO 7110 Econometrics II
 July 2018 6 / 17

## **Define the Tree for Nested Logit**

nlogitgen type = fishmode(shore: pier | beach, boat:
private | charter)
nlogittree fishmode type, choice(d)

type	N	fishmode	N	k
shore	2364	beach	1182	134
		beach pier	1182	178
boat	2364	charter private	1182	452
		$^{ddash}$ private	1182	418

k = number of times alternative is chosen

N = number of observations at each level

### **Nested Logit Model**

nlogit d p q || type:, base(shore) || fishmode: income, case(id) notree nolog

```
RUM-consistent nested logit regression
                                                 Number of obs
                                                                              4728
Case variable: id
                                                 Number of cases
                                                                              1182
Alternative variable: fishmode
                                                 Alts per case: min =
                                                                               4.0
                                                                  avg =
                                                                  max =
                                                     Wald chi2(5)
                                                                            212.37
Log likelihood = -1192.4236
                                                     Prob > chi2
                                                                            0.0000
           d
                     Coef.
                             Std. Err.
                                                  P>|z|
                                                             [95% Conf. Interval]
fishmode
                             .0018937
                                         -14.13
                                                            -.0304741
                 -.0267625
                                                  0.000
                                                                          -.023051
           р
                  1.340091
                             .3080519
                                           4.35
                                                             .7363199
                                                  0.000
                                                                          1.943861
           q
```

8 / 17

## fishmode equations

beach	I					
income	(base)					
_cons	(base)					
	(base)					
charter						
income	-8.40204	78.35484	-0.11	0.915	-161.9747	145.1706
_cons	69.96985	558.8914	0.13	0.900	-1025.437	1165.377
pier						
income	-9.458105	80.30173	-0.12	0.906	-166.8466	147.9304
_cons	58.94372	500.7334	0.12	0.906	-922.4757	1040.363
private						
income	-1.634919	8.588459	-0.19	0.849	-18.46799	15.19815
_cons	37.52542	230.9007	0.16	0.871	-415.0317	490.0825
dissimilarity	parameters					
type						
/shore_tau	83.46915	718.5287			-1324.821	1491.76
/boat_tau	52.55972	542.8935			-1011.492	1116.611
LR test for I	IA (tau = 1):		chi2(2) =	45.43	Prob > chi	2 = 0.0000

July 2018

#### **Predicted Probabilities**

estimates store NL predict plevel1 plevel2, pr tabulate fishmode, summarize(plevel2)

		y of Pr(fishmode Alternatives)	е
fishmode	Mean	Std. Dev.	Freq.
beach	.11323521	. 1333593	1182
charter	.38070949	.15724226	1182
pier	.15072734	.16982064	1182
private	.35532796	.16444334	1182
Total	. 25	.19690015	4728

10 / 17

## **AME** of Beach Price Change

```
quietly summarize p generate delta = r(sd)/1000 quietly replace p = p + delta if fishmode == "beach" predict pnew1 pnew2, pr generate dpdbeach = (pnew2 - plevel2)/delta tabulate fishmode, summarize(dpdbeach)
```

	- 1	Summ	ary of dpdbea	ch
fishmo	de	Mean	Std. Dev.	Freq.
bea	ch	00053326	.0004792	1182
chart	er	.00063591	.00054938	1182
pi	er	00065944	.00057603	1182
priva	te	.00055682	.00051133	1182
Tot	al	8.815e-09	.00079968	4728

11 / 17

#### **Ordered Outcomes**

$$y_i^* = x_i' eta + u_i$$
 $y_i = j ext{ if } lpha_{j-1} < y_i^* \le lpha_j,$ 
for  $j = 1, ..., m$ 
where  $lpha_0 = -\infty$  and  $lpha_m = \infty$ 
 $Pr(y_i = j) = Pr(lpha_{j-1} < y_i^* \le lpha_j)$ 
 $Pr(lpha_{j-1} - x_i' eta < u_i \le lpha_j - x_i' eta)$ 
 $F(lpha_j - x_i' eta) - F(lpha_{j-1} - x_i' eta)$ 

## **Ordered Logit and Probit for** j = 3

$$F(\alpha_j - x_i'\beta) - F(\alpha_{j-1} - x_i'\beta)$$

$$u \sim F(z) = \frac{e^z}{1 + e^z} \text{ or } \Phi(z)$$

$$\frac{\partial Pr(y_i=1)}{\partial x_{ri}} = -F'(\alpha_1 - x_i'\beta)\beta_r$$

$$\frac{\partial Pr(y_i=2)}{\partial x_{ri}} = \{F'(\alpha_1 - x_i'\beta) - F'(\alpha_2 - x_i'\beta)\}\beta_r$$

The term in braces can be + or -

$$\frac{\partial Pr(y_i=3)}{\partial x_{ri}} = F'(\alpha_2 - x_i'\beta)\beta_r$$

 ✓ □ ▷ ✓ ⓓ ▷ ✓ ඕ ▷ ✓ ඕ ▷ ☑
 ✓ ℚ ҈

 Vitor Kamada
 ECO 7110 Econometrics II
 July 2018
 13 / 17

### Rand Health Insurance Experiment

health status	Freq.	Percent	Cum.
poor_or_fair good excellent	523 2,034 3,017	9.38 36.49 54.13	9.38 45.87 100.00
Total	5,574	100.00	

Variable	0bs	Mean	Std. Dev.	Min	Max
hlthstat	5574	2.447435	.659524	1	3
age	5574	25.57613	16.73011	.0253251	63.27515
linc	5574	8.696929	1.220592	0	10.28324
ndisease	5574	11.20526	6.788959	0	58.6

### **Ordered Logit Model**

## ologit hlthstat age linc ndisease, nolog

hlthstat	Coef.	Std. Err.	z	P> z	[95% Conf.	Interval]
age linc ndisease	0292944 .2836537 0549905	.001681 .0231098 .0040692	-17.43 12.27 -13.51	0.000 0.000 0.000	0325891 .2383593 0629661	0259996 .3289481 047015
/cut1 /cut2	-1.39598 .9513097	.2061301 .2054301			-1.799987 .5486741	9919722 1.353945

#### **Predicted Probabilities**

predict p1ologit p2ologit p3ologit, pr summarize hlthpf hlthg hlthe p1ologit p2ologit p3ologit, separator(0)

Variable	0bs	Mean	Std. Dev.	Min	Max
hlthpf	5574	.0938285	.2916161	0	1
hlthg	5574	.3649085	.4814477	0	1
hlthe	5574	.541263	.4983392	0	1
p1ologit	5574	.0946903	.0843148	.0233629	.859022
p2ologit	5574	.3651672	.0946158	.1255265	.5276064
p3ologit	5574	.5401425	.1640575	.0154515	.7999009

#### Marginal Effect at Mean for Health Status Excellent

## margins, dydx(\*) predict(outcome(3)) atmean

Conditional marginal effects Number of obs = 5574

Model VCE : OIM

Expression : Pr(hlthstat==3), predict(outcome(3))

dy/dx w.r.t. : age linc ndisease

	dy/dx	Delta-method Std. Err.	l z	P> z	[95% Conf.	Interval]
age linc ndisease	0072824 .070515 0136704	.0004179 .0057527	-17.43 12.26 -13.50	0.000 0.000 0.000	0081014 .05924 015655	0064634 .0817901 0116858

Vitor Kamada ECO 7110 Econometrics II July 2018 17,