

3) Generalized Method of Moments (GMM)

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Wooldridge (2010). **Econometric Analysis of Cross Section and Panel Data.** Ch 8

<https://ebookcentral.proquest.com/lib/wayne/detail.action?docID=3339196&>

Method of Moments

$$E[y - \mu] = \frac{1}{N} \sum_{i=1}^N (y_i - \mu) = 0$$

$$\hat{\mu}_{MM} = \frac{1}{N} \sum_{i=1}^N y_i = \bar{y}$$

$$E[u|\mathbf{x}] \rightarrow E[\mathbf{x}u] = 0$$

$$E[\mathbf{x}(y - \mathbf{x}'\beta)] = \frac{1}{N} \sum_{i=1}^N x_i(y_i - \mathbf{x}_i'\beta) = 0$$

$$\hat{\beta}_{MM} = (\sum_i x_i x_i')^{-1} (\sum_i x_i y_i) = 0$$

Generalized Method of Moments (GMM)

$$Q_N(\beta) = \left\{ \frac{1}{N}(y - X\beta)'Z \right\} W_N \left\{ \frac{1}{N}Z'(y - X\beta) \right\}$$

$$\frac{\partial Q_N(\beta)}{\partial \beta} = -2 \left[\frac{1}{N}X'Z \right] W_N \left[\frac{1}{N}Z'(y - X\beta) \right] = 0$$

$$\hat{\beta}_{GMM} = (X'ZW_NZ'X)^{-1}X'ZW_NZ'y$$

$$\hat{\beta}_{IV} = (Z'X)^{-1}Z'y$$

$$\hat{\beta}_{2SLS} = \{X'Z(Z'Z)^{-1}Z'X\}^{-1}X'Z(Z'Z)^{-1}Z'y$$

Optimal GMM

$$\hat{\beta}_{OGMM} = (X'Z\hat{S}^{-1}Z'X)^{-1}X'Z\hat{S}^{-1}Z'y$$

$$\hat{S} = \frac{1}{N} \sum_{i=1}^N \hat{u}^2 z_i z_i' = \frac{Z'DZ}{N}$$

$$\text{If } E[u_i^2 | z_i] = \sigma^2, \text{ then } \hat{S} = \frac{s^2 Z'Z}{N}$$

Medical Expenditure Panel Survey (MEPS): Individuals over the age of 65 years

ldrugexp: the log of total out-of-pocket expenditures on prescribed medications

hi_empunion: indicator for whether the individual holds either employer or union-sponsored health insurance

totchr: # of chronic conditions

sociodemographic variables: age, female, blhisp, and linc

ssiratio: ratio of an individual's social security income to the individual's income from all sources

multlc: if the firm is a large operator with multiple locations

pip install linearmodels

```
import numpy as np
import pandas as pd
import matplotlib.pyplot as plt
from linearmodels import IV2SLS, IVLIML, IVGMM
```

```
file="https://github.com/VitorKamada/ECO7110/raw/master/Data/mus06data.dta"
data = pd.read_stata(file)
data.head()
```

	ssiratio	age	educyr	drugexp	private	female	hisp
0	0.149877	74	16	595	1	1	0
1	0.395856	73	8	1783	1	0	0
2	1.000000	80	12	176	0	1	0
3	0.206639	70	17	2437	1	0	0
4	0.537192	91	16	330	0	0	0

Summary Statistics

```
controls = ['totchr', 'female', 'age', 'linc', 'blhisp']  
print(data[['ldrugexp', 'hi_empunion'] +  
        controls].describe(percentiles=[]))
```

	ldrugexp	hi_empunion	linc	blhisp
count	10391.000000	10391.000000	10089.000000	10391.000000
mean	6.479664	0.379655	2.743271	0.170340
std	1.363393	0.485324	0.913144	0.375956
min	0.000000	0.000000	-6.907755	0.000000
50%	6.677083	0.000000	2.743160	0.000000
max	10.180172	1.000000	5.744476	1.000000


```
instruments = ['ssiratio', 'lowincome', 'multlc', 'firmsz']
print(data[instruments].describe(percentiles=[]))
```

	ssiratio	lowincome	multlc	firmsz
count	10391.000000	10391.000000	10391.000000	10391.000000
mean	0.520628	0.185641	0.060341	0.136501
std	0.374589	0.388836	0.238128	2.138914
min	-2.100647	0.000000	0.000000	0.000000
50%	0.490139	0.000000	0.000000	0.000000
max	9.250620	1.000000	1.000000	50.000000

```
data[['hi_empunion'] + instruments].corr()
```

	hi_empunion	ssiratio	lowincome	multlc	firmsz
hi_empunion	1.000000	-0.196284	-0.114425	0.119054	0.037173
ssiratio	-0.196284	1.000000	0.249820	-0.174394	-0.040481
lowincome	-0.114425	0.249820	1.000000	-0.060701	-0.007802
multlc	0.119054	-0.174394	-0.060701	1.000000	0.187394
firmsz	0.037173	-0.040481	-0.007802	0.187394	1.000000

```
from statsmodels.api import OLS, add_constant
data['const'] = 1
controls = ['const'] + controls
```

```
ivolsmod = IV2SLS(data.ldrugexp,
                   data[['hi_empunion'] + controls], None, None)
res_ols = ivolsmod.fit()
print(res_ols)
```

	Parameter	Std. Err.	T-stat	P-value
hi_empunion	0.0739	0.0260	2.8441	0.0045
const	5.8611	0.1570	37.320	0.0000
totchr	0.4404	0.0094	47.049	0.0000
female	0.0578	0.0254	2.2797	0.0226
age	-0.0035	0.0019	-1.8228	0.0683
linc	0.0105	0.0137	0.7646	0.4445
blhisp	-0.1513	0.0341	-4.4353	0.0000

2SLS

```
ivmod = IV2SLS(data.ldrugexp, data[controls],  
               data.hi_empunion, data[['ssiratio', 'multlc']])  
res_2sls = ivmod.fit()  
print(res_2sls.summary)
```

	Parameter	Std. Err.	T-stat	P-value

const	6.8752	0.2579	26.660	0.0000
totchr	0.4512	0.0103	43.769	0.0000
female	-0.0278	0.0322	-0.8653	0.3869
age	-0.0141	0.0029	-4.8753	0.0000
linc	0.0943	0.0219	4.3079	0.0000
blhisp	-0.2237	0.0396	-5.6514	0.0000
hi_empunion	-0.9899	0.2046	-4.8386	0.0000

OGMM

```
ivmod = IVGMM(data.ldrugexp, data[controls],  
               data.hi_empunion, data[['ssiratio', 'multlc']])  
res_gmm = ivmod.fit()  
print(res_gmm)
```

	Parameter	Std. Err.	T-stat	P-value

const	6.8778	0.2580	26.658	0.0000
totchr	0.4510	0.0103	43.738	0.0000
female	-0.0282	0.0322	-0.8752	0.3815
age	-0.0142	0.0029	-4.8773	0.0000
linc	0.0945	0.0219	4.3142	0.0000
blhisp	-0.2231	0.0396	-5.6344	0.0000
hi_empunion	-0.9933	0.2047	-4.8530	0.0000

```
print(res_gmm.first_stage)
```

const	0.9834 (16.780)
totchr	0.0133 (3.6234)
female	-0.0727 (-7.5644)
age	-0.0081 (-11.311)
linc	0.0444 (6.7838)
blhisp	-0.0679 (-5.5484)
ssiratio	-0.1823 (-7.8326)
multlc	0.1209 (5.8212)

```

ivmod = IVLIML(data.ldrugexp, data[controls],
               data.hi_empunion, data[['ssiratio', 'multlc']])
res_liml = ivmod.fit(cov_type='robust')
from linearmodels.iv.results import compare
print(compare({'2SLS': res_2sls, 'LIML': res_liml, 'GMM': res_gmm}))

```

	2SLS	GMM	LIML
const	6.8752 (26.660)	6.8778 (26.658)	6.8807 (26.577)
totchr	0.4512 (43.769)	0.4510 (43.738)	0.4513 (43.730)
female	-0.0278 (-0.8653)	-0.0282 (-0.8752)	-0.0283 (-0.8776)
age	-0.0141 (-4.8753)	-0.0142 (-4.8773)	-0.0142 (-4.8781)
linc	0.0943 (4.3079)	0.0945 (4.3142)	0.0947 (4.3114)
blhisp	-0.2237 (-5.6514)	-0.2231 (-5.6344)	-0.2241 (-5.6531)
hi_empunion	-0.9899 (-4.8386)	-0.9933 (-4.8530)	-0.9957 (-4.8361)

Overidentified Test (OID), Hansen's Test, and Sargan's Test

$$Q(\hat{\beta}) = \left\{ \frac{1}{N} (y - X\hat{\beta})' Z \right\} \hat{S}^{-1} \left\{ \frac{1}{N} Z' (y - X\hat{\beta}) \right\}$$

$$Z' (y - X\hat{\beta}) \simeq 0, \text{ so } Q(\hat{\beta}) \simeq 0$$

$$Q(\hat{\beta}) \stackrel{a}{\sim} \chi_r^2,$$

r is the # of overidentifying restrictions

$$H_0 : E\{Z'(y - X\beta)\} = 0$$

Rejection means that at least one of the instruments is not valid

Test of Overidentifying Restrictions

```
res_gmm.j_stat
```

H_0 : Overidentifying Restriction is Valid

H0: Expected moment conditions are equal to 0

Statistic: 1.0475

P-value: 0.3061

Distributed: chi2(1)

WaldTestStatistic, id: 0x7f1d6b251160

Four Available Instruments

```
ivmod = IVGMM(data.ldrugexp, data[controls],  
              data.hi_empunion, data[instruments])  
res_gmm_all = ivmod.fit()  
res_gmm_all.j_stat
```

H0: Expected moment conditions are equal to 0

Statistic: 11.5903

P-value: 0.0089

Distributed: chi2(3)

WaldTestStatistic, id: 0x7f1d6af14710

```
print(res_gmm_all)
```

	Parameter	Std. Err.	T-stat	P-value
const	6.7126	0.2426	27.670	0.0000
totchr	0.4495	0.0100	44.738	0.0000
female	-0.0105	0.0307	-0.3406	0.7334
age	-0.0125	0.0027	-4.5364	0.0000
linc	0.0797	0.0203	3.9162	0.0001
blhisp	-0.2061	0.0383	-5.3828	0.0000
hi_empunion	-0.8124	0.1846	-4.3999	0.0000

```
print(res_gmm_all.first_stage)
```

totchr	0.0133 (3.6494)
female	-0.0724 (-7.5497)
age	-0.0080 (-11.206)
linc	0.0410 (6.3552)
blhisp	-0.0676 (-5.5369)
ssiratio	-0.1690 (-7.3289)
lowincome	-0.0637 (-5.1947)
multlc	0.1151 (5.4799)
firmsz	0.0037 (1.9286)