

## 4) Analysis of Variance (ANOVA): Completely Randomized Designs

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# Adaptation vs Mutation

**Fact:** Strains of bacteria die if exposed to certain virus, but some survives and reproduce fast

- In 1940s, both theories predict same average numbers of resistant bacteria
- But, Mutation Theory predicts a much higher variance
- 1969 Nobel Prize in Physiology/Medicine for Luria and Delbruck

# Log(Lifetime) of Resin in Integrated Circuits

Temperature (°C)									
175		194		213		231		250	
2.04	1.85	1.66	1.66	1.53	1.35	1.15	1.21	1.26	1.02
1.91	1.96	1.71	1.61	1.54	1.27	1.22	1.28	.83	1.09
2.00	1.88	1.42	1.55	1.38	1.26	1.17	1.17	1.08	1.06
1.92	1.90	1.76	1.66	1.31	1.38	1.16			

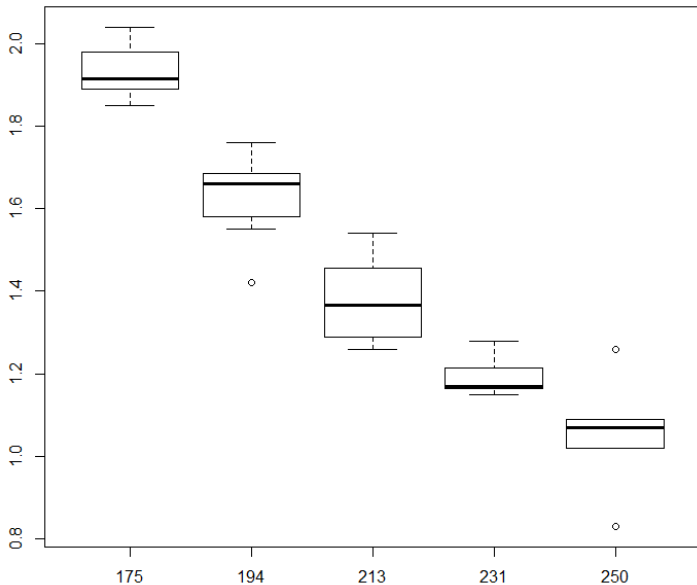
**summary(resin)**

**attach(resin)**

Statistic	N	Mean	St. Dev.	Min	Max
temp	37	210.081	26.144	175	250
y	37	1.465	0.326	0.830	2.040

Nelson (1990)

# boxplot(y~temp)



# Mechanics of ANOVA

$$y_{ij} - \mu = \alpha_i + \epsilon_{ij}$$

$$y_{ij} - \bar{y}_{\bullet\bullet} = (\bar{y}_{i\bullet} - \bar{y}_{\bullet\bullet}) + (y_{ij} - \bar{y}_{i\bullet})$$

$$y_{ij} - \bar{y}_{\bullet\bullet} = \hat{\alpha}_i + r_{ij}$$

$$(y_{ij} - \bar{y}_{\bullet\bullet})^2 = \hat{\alpha}_i^2 + r_{ij}^2 + 2\hat{\alpha}_i r_{ij}$$

$$SS_T = SS_{Trt} + SS_E + 2 \sum_{i=1}^g \sum_{j=1}^{n_i} \hat{\alpha}_i r_{ij}$$

# Generic ANOVA Table

Source	DF	SS	MS	F
Treatments	$g - 1$	$SS_{Trt}$	$\frac{SS_{Trt}}{g-1}$	$\frac{MS_{Trt}}{MS_E}$
Error	$N - g$	$SS_E$	$\frac{SS_E}{N-g}$	

$$MS_{Trt} = \frac{1}{g-1} \sum_{i=1}^g \sum_{j=1}^{n_i} (\bar{y}_{i\bullet} - \bar{y}_{\bullet\bullet})^2 = \sum_{i=1}^g n_i \hat{\alpha}_i^2$$

$$MS_E = \frac{1}{N-g} \sum_{i=1}^g \sum_{j=1}^{n_i} (y_{ij} - \bar{y}_{i\bullet})^2 = \hat{\sigma}^2$$

# ANOVA Table

```
Dummy <- with(resin,as.factor(temp))
```

```
Result <- lm(y~Dummy)
```

```
anova(Result)
```

## Analysis of Variance Table

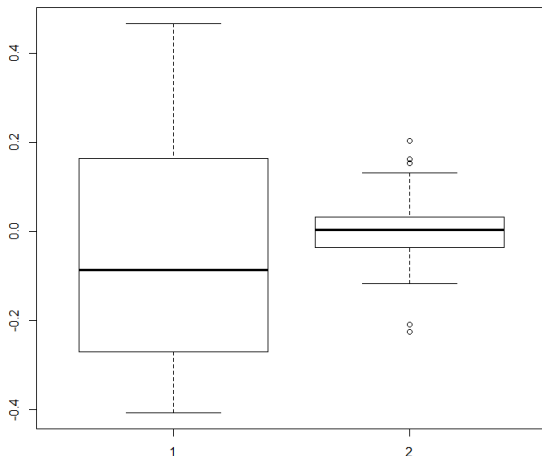
Response: y

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
Dummy	4	3.5376	0.88441	96.363	< 2.2e-16 ***
Residuals	32	0.2937	0.00918		

# Side-by-Side Plots

```
yhat <- predict(Result); alpha <- yhat - 1.465
```

```
Residuals <- resid(Result); boxplot(alpha, Residuals)
```





# summary(Result)

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )	
(Intercept)	1.93250	0.03387	57.055	< 2e-16	***
Dummy194	-0.30375	0.04790	-6.341	4.06e-07	***
Dummy213	-0.55500	0.04790	-11.586	5.49e-13	***
Dummy231	-0.73821	0.04958	-14.889	6.13e-16	***
Dummy250	-0.87583	0.05174	-16.928	< 2e-16	***

# Dose-Response Modeling

$$\mu + \alpha_i = f(z_i; \theta)$$

$$\mu + \alpha_i = \theta_0 + \theta_1 z_i + \theta_2 z_i^2 + \dots + \theta_{g-1} z_i^{g-1}$$

```
p1 <- lm(y~temp)
```

```
p2 <- lm(y~temp+l(temp^2))
```

```
p3 <- lm(y~temp+l(temp^2)+l(temp^3))
```

```
p4 <- lm(y~temp+l(temp^2)+l(temp^3)+l(temp^4))
```

```
stargazer(p1,p2,p3,p4, omit.stat=c("ser","f"),
```

```
type="text", out="Reg.txt")
```

# Regression Results

	<i>Dependent variable:</i>			
	Lifetime (in hours)			
	(1)	(2)	(3)	(4)
temp	-0.012*** (0.001)	-0.045*** (0.011)	-0.037 (0.187)	0.076 (3.750)
l(temp^2)		0.0001*** (0.00003)	0.00004 (0.001)	-0.001 (0.027)
l(temp^3)			0.00000 (0.00000)	0.00000 (0.0001)
l(temp^4)				-0.000 (0.00000)
Constant	3.956*** (0.139)	7.418*** (1.156)	6.827 (12.987)	0.970 (195.724)
Observations	37	37	37	37
R <sup>2</sup>	0.903	0.923	0.923	0.923
Adjusted R <sup>2</sup>	0.900	0.919	0.916	0.914

Note:

\*p<0.1; \*\*p<0.05; \*\*\*p<0.01

# anova(p1,p2,p3,p4)

## Analysis of Variance Table

Model 1:  $y \sim \text{temp}$

Model 2:  $y \sim \text{temp} + \text{I}(\text{temp}^2)$

Model 3:  $y \sim \text{temp} + \text{I}(\text{temp}^2) + \text{I}(\text{temp}^3)$

Model 4:  $y \sim \text{temp} + \text{I}(\text{temp}^2) + \text{I}(\text{temp}^3) + \text{I}(\text{temp}^4)$

	Res.Df	RSS	Df	Sum of Sq	F	Pr(>F)
1	35	0.37206				
2	34	0.29372	1	0.078343	8.5361	0.006338 **
3	33	0.29370	1	0.000019	0.0020	0.964399
4	32	0.29369	1	0.000008	0.0009	0.976258

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Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1