

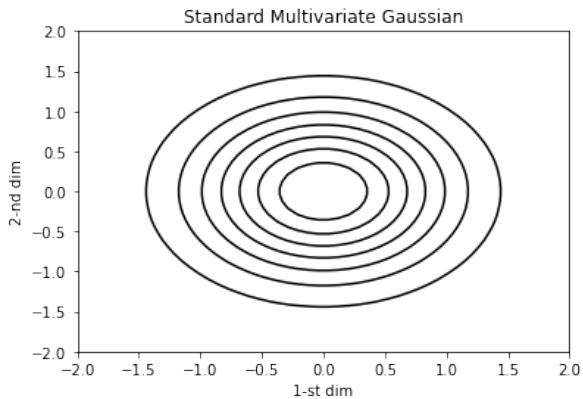
# Deep Programming Probabilistic Languages

Vangjus Komini

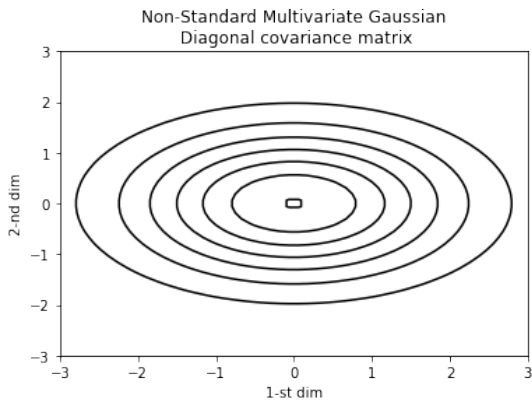
Assignment 1

August 29, 2022

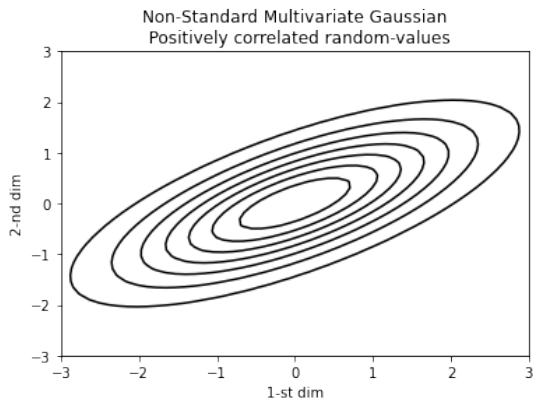
# Gaussian distribution



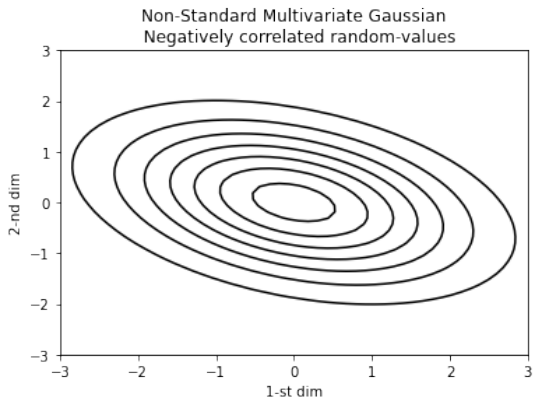
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# Likelihood Gaussian distribution

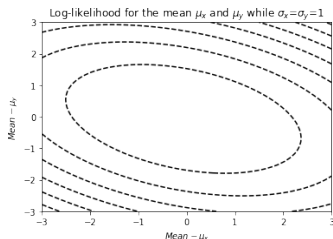
$$L(\mu, \Sigma|x) = -\frac{(x - \mu)^T \Sigma^{-1}(x - \mu)}{2} - \frac{1}{2}\{N * \ln(2\pi) + \ln(|\Sigma|)\}$$

The log-likelihood for both mean given the covariance matrix is:

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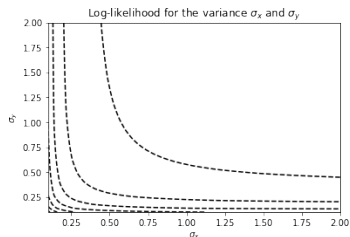
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# Gamma and Poisson conjugacy

- ▶ Gamma distribution for hyper-parameter  $\lambda$  for a given  $\alpha$  and  $\beta$  is:

$$p(\lambda|\alpha, \beta) = \frac{\beta^{-\alpha}}{\Gamma(\alpha)} \lambda^{\alpha-1} e^{-\beta\lambda}$$

- ▶ The likelihood of the hyper-parameter  $\lambda$  for a given set of independent data  $x_1, x_2, \dots, x_N$  using a Poisson distribution is:

$$P(x_1, x_2, \dots, x_N|\lambda) = \prod_{i=1}^N \frac{\lambda^{x_i} e^{-\lambda}}{x_i!} = \frac{\lambda^{\sum_{i=1}^N x_i} e^{-\sum_{i=1}^N \lambda}}{\prod_{i=1}^N x_i!} = \frac{\lambda^{N x_{mu}} e^{-N\lambda}}{\prod_{i=1}^N x_i!}$$

- ▶ Using Bayes it is possible to get the update on  $\lambda$  as:

$$p(\lambda|x_1, x_2, \dots, x_N) = \frac{p(x_1, x_2, \dots, x_N|\lambda)p(\lambda)}{\int p(x_1, x_2, \dots, x_N, \lambda)d\lambda} = \frac{p(x_1, x_2, \dots, x_N|\lambda)p(\lambda)}{p(x_1, x_2, \dots, x_N)}$$

- ▶ Since the evidence is not dependent on  $\lambda$  it is not important on the maximum a posterior:

$$p(\lambda|x_1, x_2, \dots, x_N) \propto p(x_1, x_2, \dots, x_N|\lambda)p(\lambda)$$

$$p(\lambda|x_1, x_2, \dots, x_N) \propto \frac{\lambda^{N x_{mu}} e^{-N\lambda}}{\prod_{i=1}^N x_i!} p(\lambda)$$



# Gamma and Poisson conjugacy



$$p(\lambda|x_1, x_2, \dots, x_N) \propto \frac{\lambda^{Nx_{mu}} e^{-N\lambda}}{\prod_{i=1}^N x_i!} \frac{\beta^{-\alpha}}{\Gamma(\alpha)} \lambda^{\alpha-1} e^{-\beta\lambda}$$



$$p(\lambda|x_1, x_2, \dots, x_N) \propto \frac{\beta^{-\alpha}}{\Gamma(\alpha) \prod_{i=1}^N x_i!} \lambda^{Nx_{mu}} e^{-N\lambda} \lambda^{\alpha-1} e^{-\beta\lambda}$$



$$p(\lambda|x_1, x_2, \dots, x_N) \propto \frac{\beta^{-\alpha}}{\Gamma(\alpha) \prod_{i=1}^N x_i!} \lambda^{Nx_{mu}+\alpha-1} e^{-N\lambda-\beta\lambda}$$

- ▶ Since  $\frac{\beta^{-\alpha}}{\Gamma(\alpha) \prod_{i=1}^N x_i!}$  is not dependent on  $\lambda$  can be removed away from the right hand side.



$$p(\lambda|x_1, x_2, \dots, x_N) \propto \lambda^{Nx_{mu}+\alpha-1} e^{-\alpha(N+\beta)}$$



$$p(\lambda|x_1, x_2, \dots, x_N) \propto \text{Gamma}(Nx_{mu} + \alpha, N + \beta)$$

Hence the Poisson and the Gamma distribution form a conjugacy.

# MCMC

$$p(x_1, x_2) = e^{-\frac{(2x_1 + \sin(2\pi x_1))^2}{2}} e^{-\frac{(x_2 - x_1^3)^2}{0.1}}$$

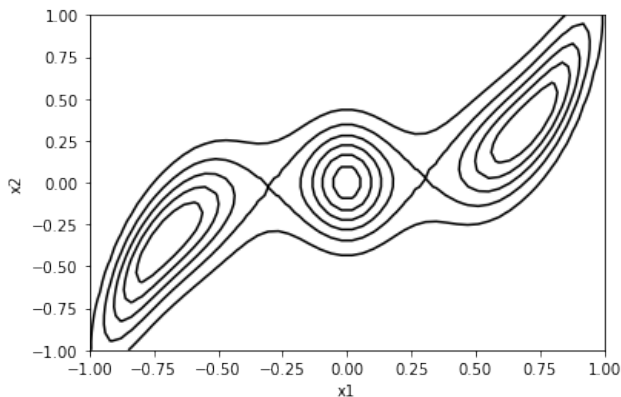


Figure: Target distribution

# Metropolis Hasting with Gibbs

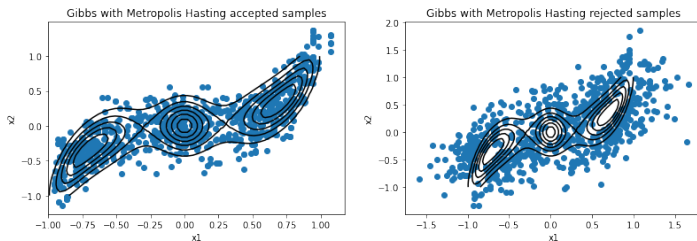


Figure: Left: Accepted samples. Right: Rejected samples

# Metropolis Hasting

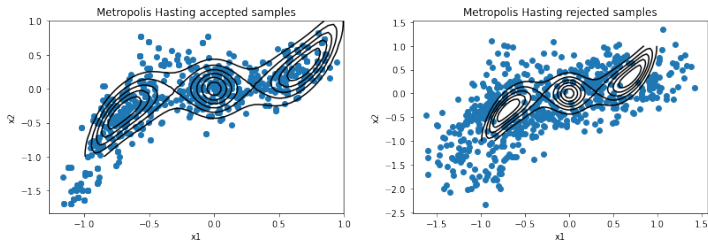


Figure: Left: Accepted samples. Right: Rejected samples

# HMC

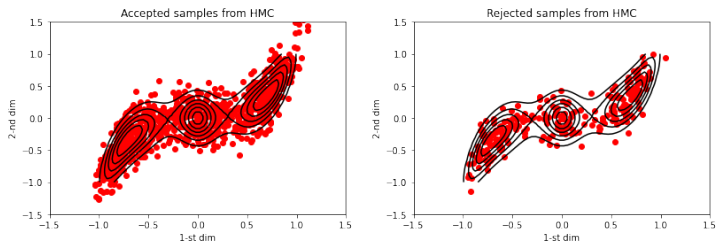
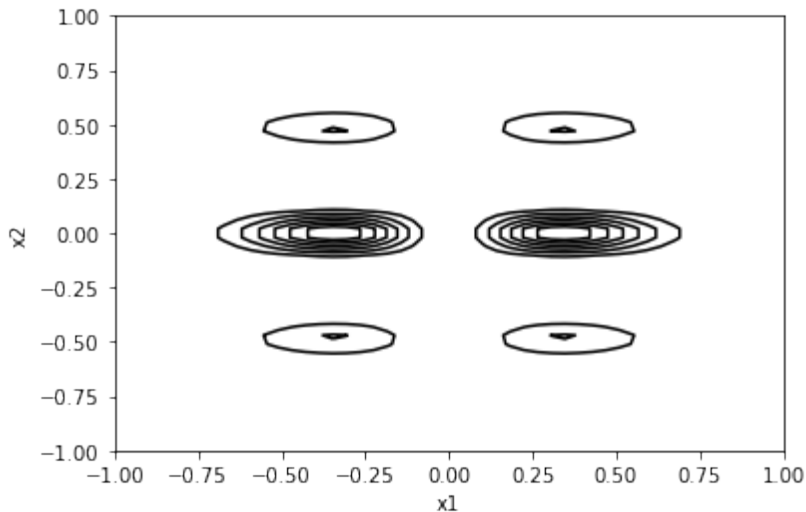


Figure: Left: Accepted samples. Right: Rejected samples

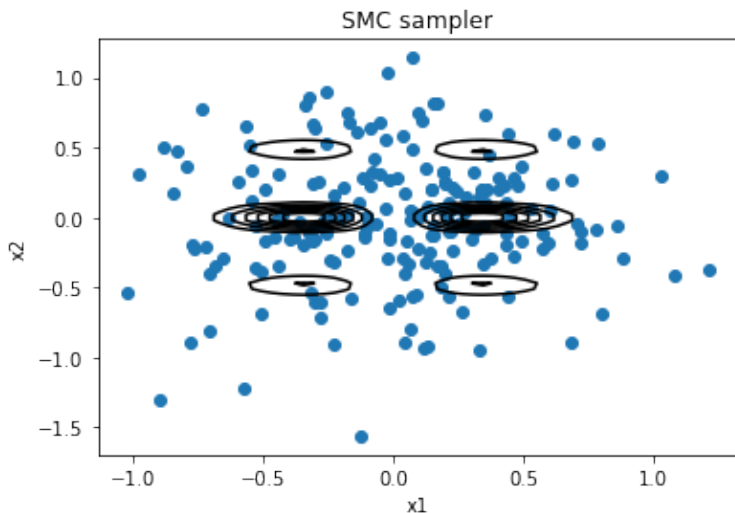
# SMC

$$\pi(x) = \frac{1}{Z} \mathbb{I}(-1 \leq x_1 \leq 1) \mathbb{I}(-1 \leq x_2 \leq 1) \sin^2(x_1 \pi) \cos^8(x_2 2\pi) \exp(-5(x_1^2 + x_2^2))$$



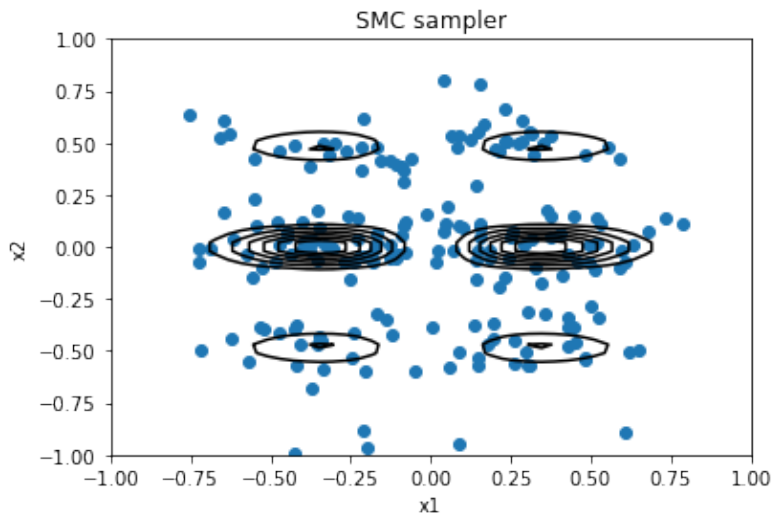
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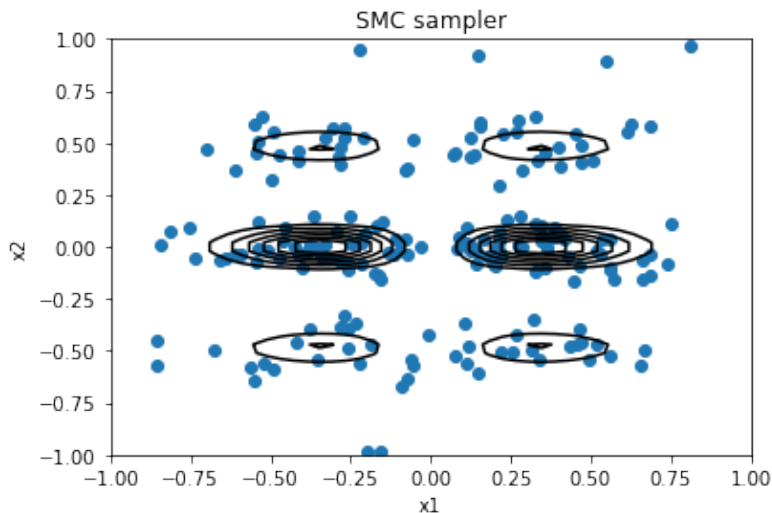
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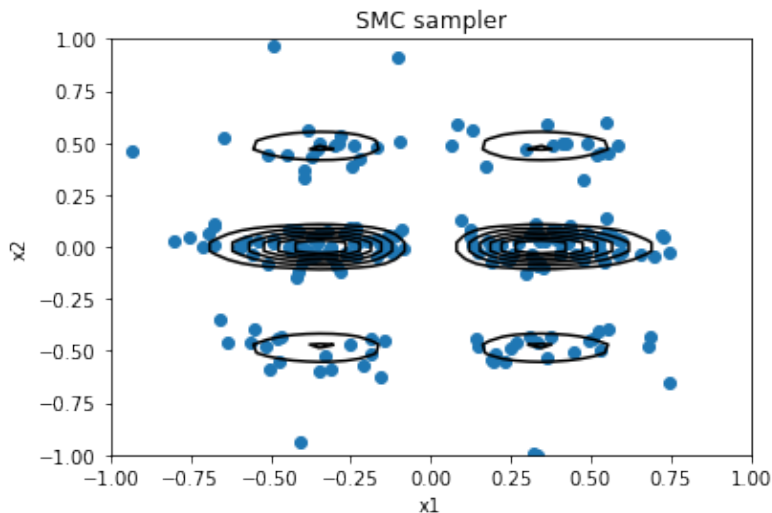
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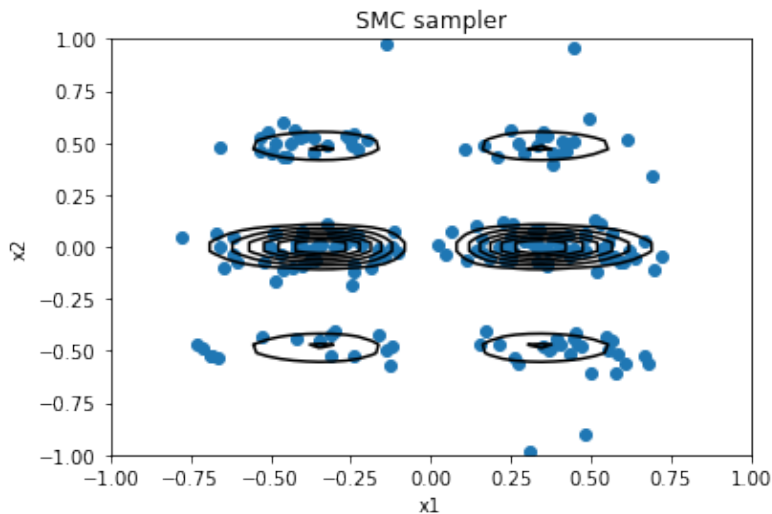
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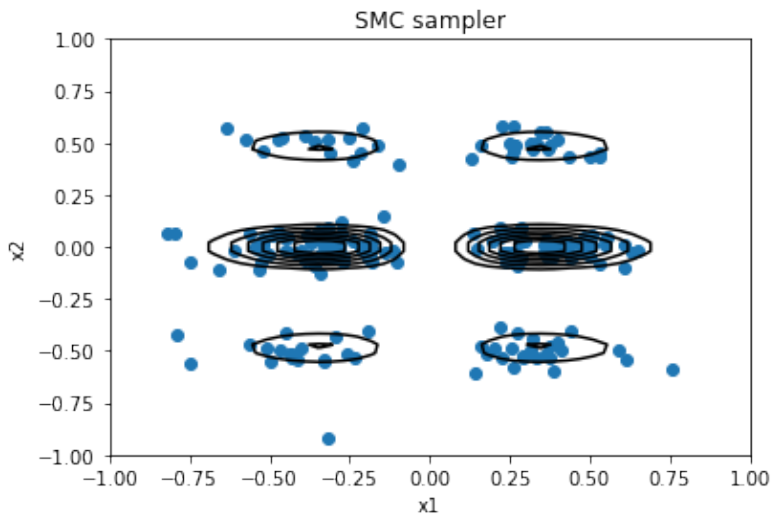
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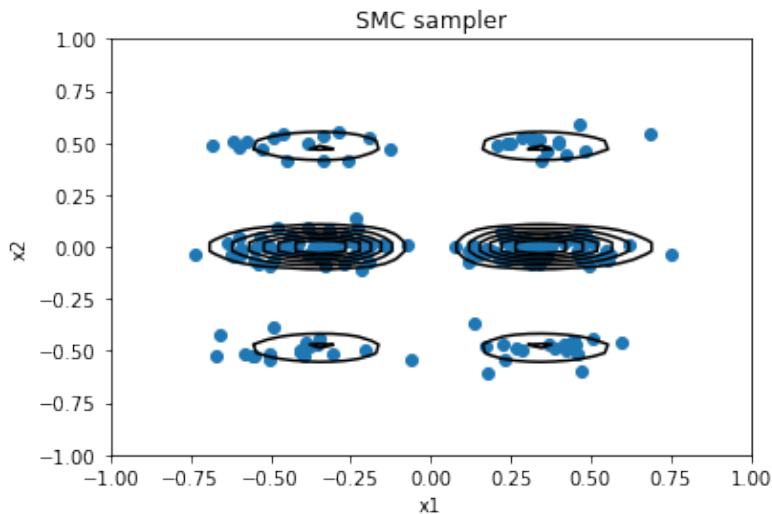
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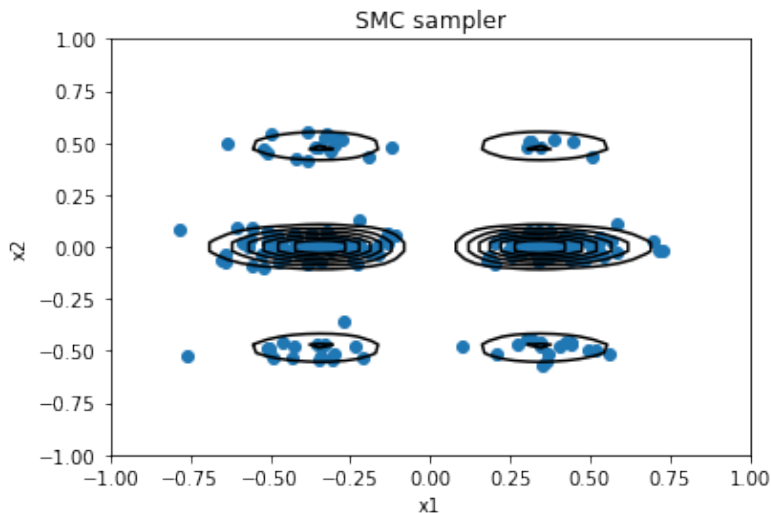
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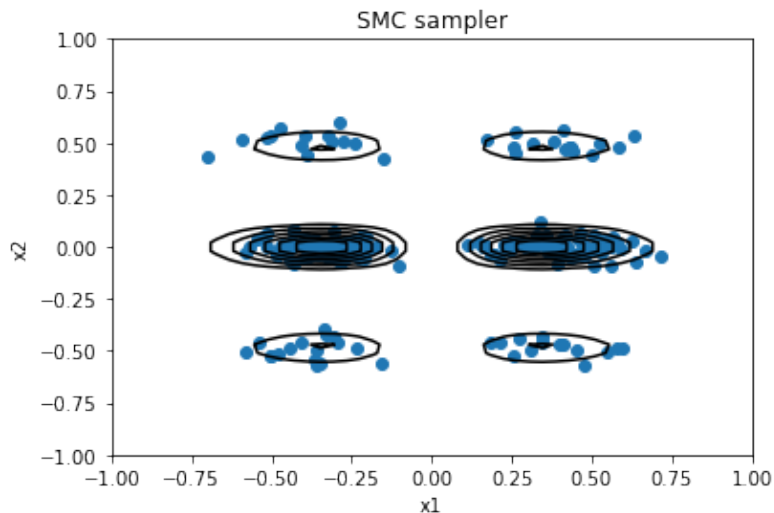
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