



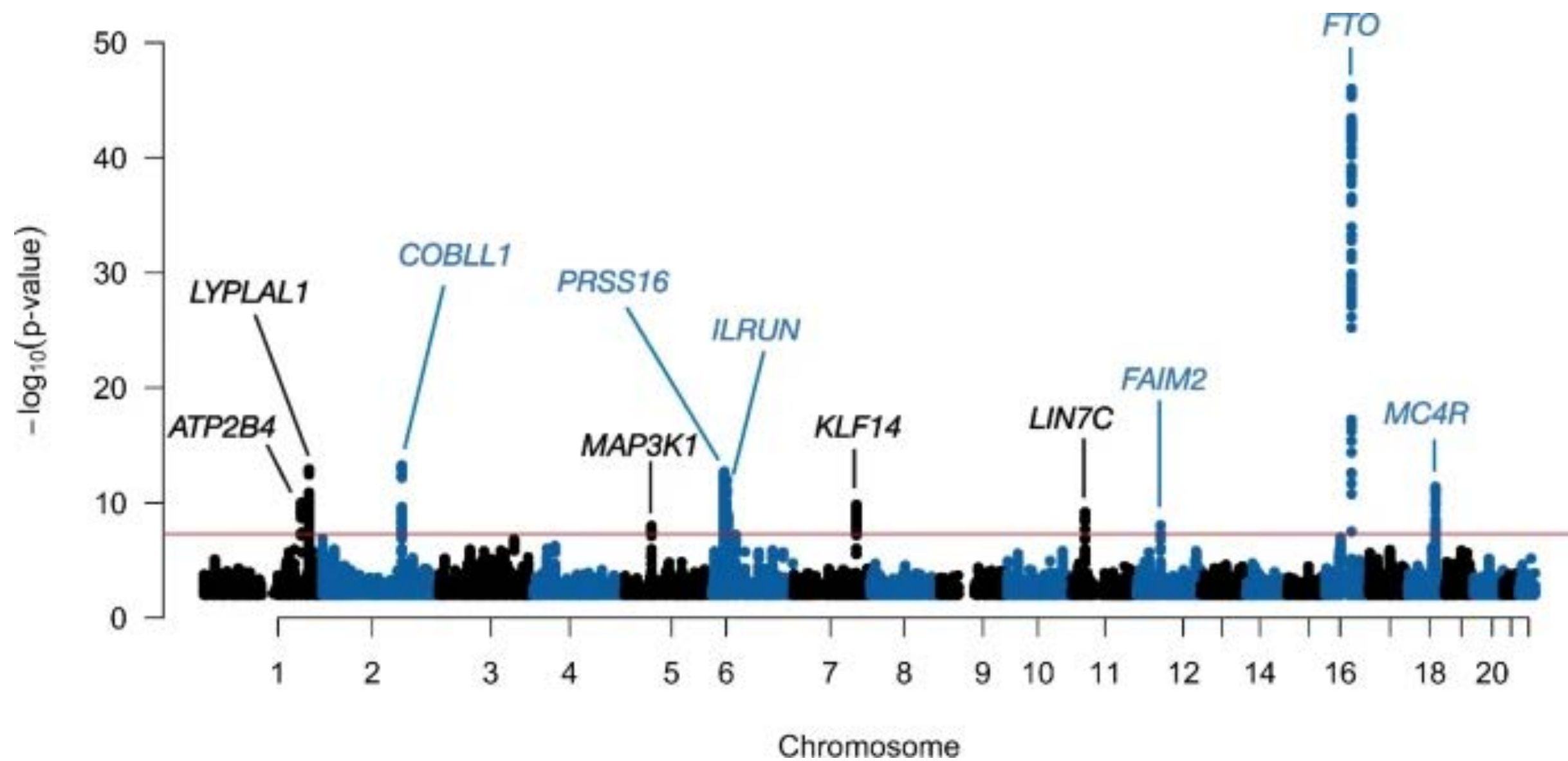
Post-GWAS

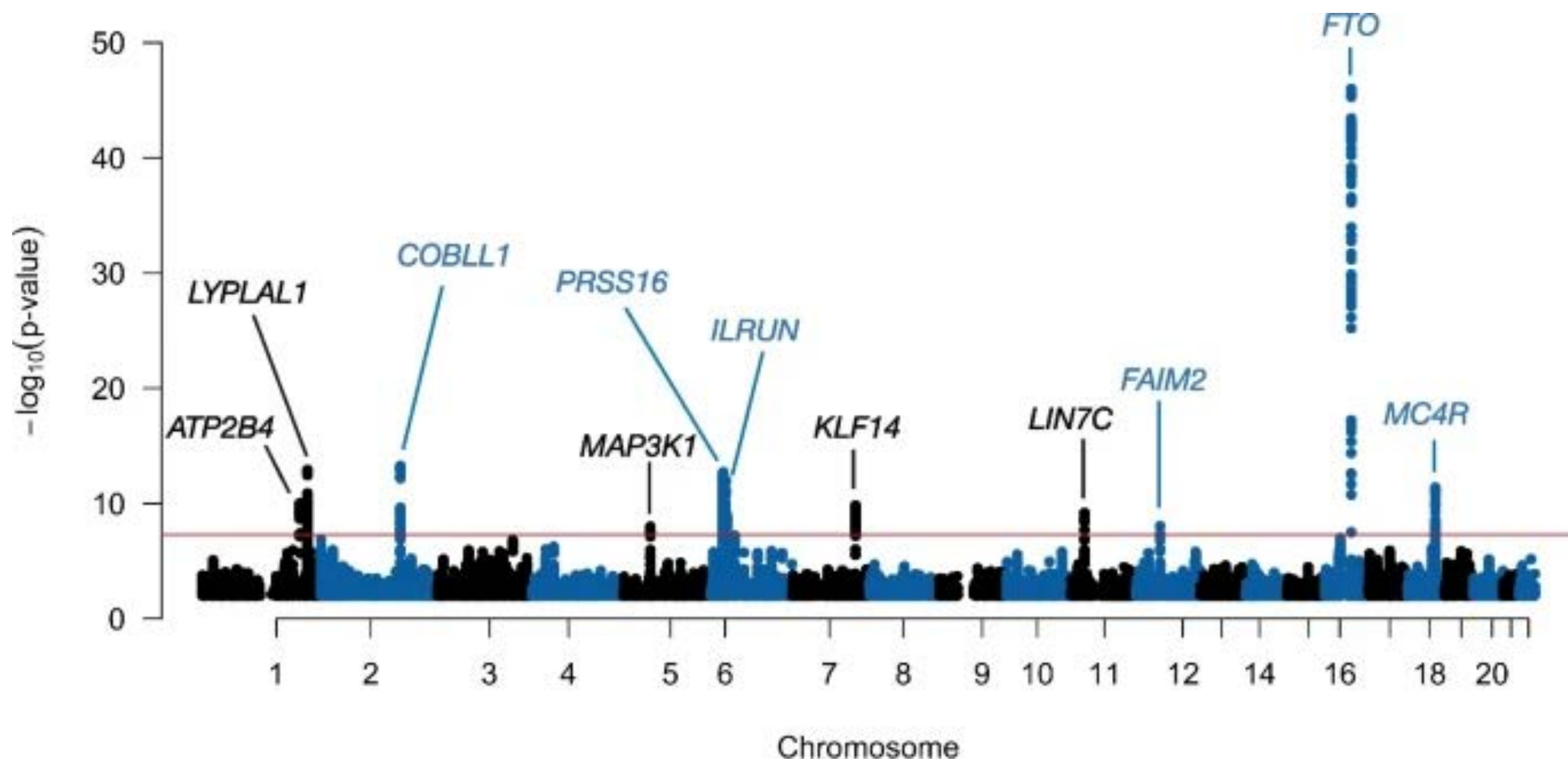
Семинар наставника 30.09.2024

Александр Ракитько

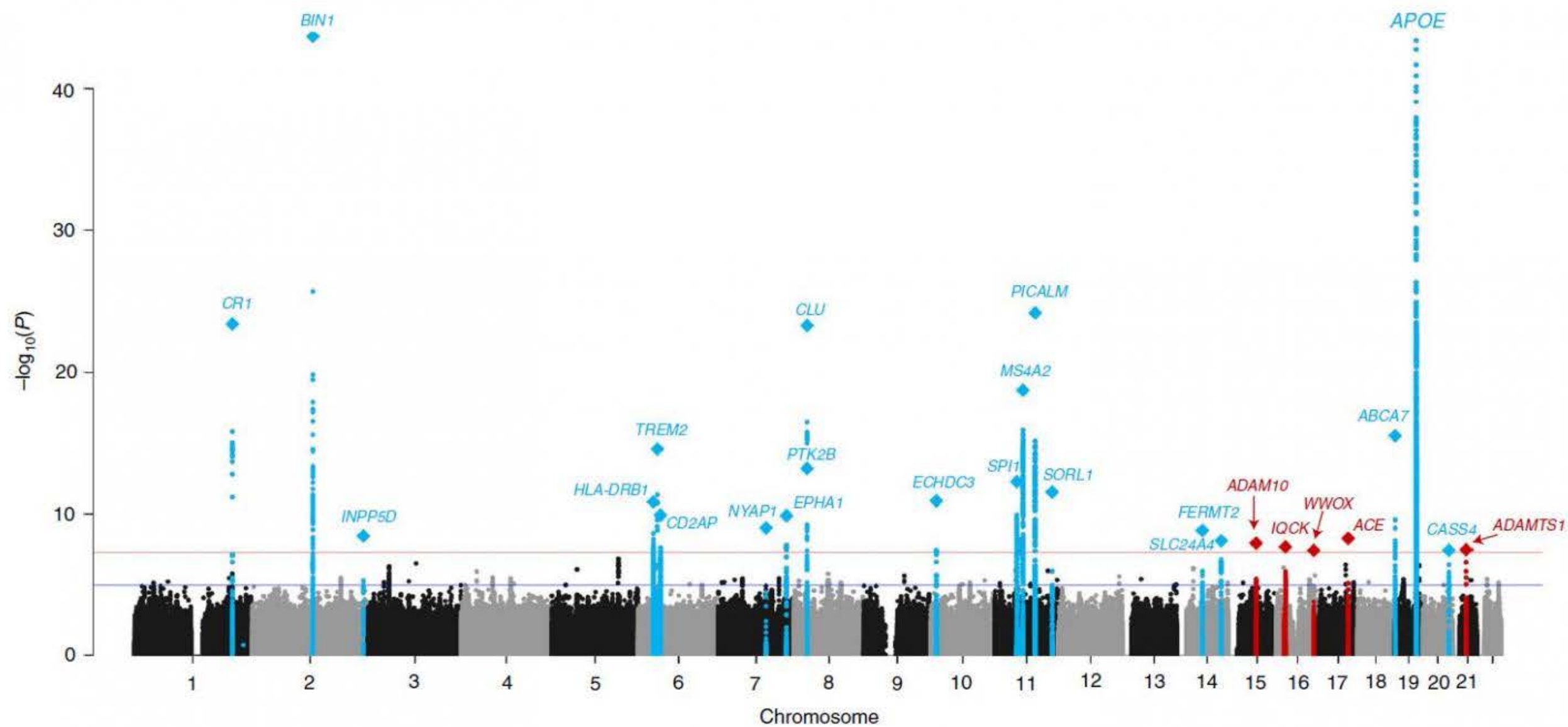
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1	rs3094315	792429	G	0.1562	0.08537	A	2.051	0.1521	1.984
1	rs6672353	817376	A	0.0102	0	G	0.8209	0.3649	NA
1	rs4040617	819185	G	0.1429	0.08537	A	1.432	0.2315	1.786
1	rs4075116	1043552	C	0.04082	0.07317	T	0.8907	0.3453	0.539
1	rs9442385	1137258	T	0.3646	0.4268	G	0.7181	0.3968	0.7705
1	rs11260562	1205233	A	0.02128	0.03659	G	0.3719	0.542	0.5725
1	rs6685064	1251215	C	0.3854	0.439	T	0.5253	0.4686	0.8013
1	rs3766180	1563420	T	0.1735	0.09756	C	2.151	0.1425	1.941
1	rs6603791	1586208	A	0.1735	0.08537	G	2.999	0.08332	2.249
1	rs7519837	1596068	C	0.1667	0.08537	T	2.598	0.107	2.143
1	rs3737628	1755094	T	0.5102	0.4756	C	0.2137	0.6438	1.149
1	rs7511905	1825948	A	0.08333	0.1098	C	0.3574	0.5499	0.7374
1	rs3855951	1836464	C	0.1224	0.2125	T	2.619	0.1056	0.5171
1	rs6603803	1844850	A	0.4896	0.5122	G	0.09045	0.7636	0.9135
1	rs2803285	1920531	A	0.1354	0.08537	G	1.111	0.2919	1.678
1	rs7513222	2060063	G	0.4592	0.3415	A	2.566	0.1092	1.637
1	rs3107146	2079746	T	0.03061	0.08537	C	2.551	0.1102	0.3383
1	rs3107157	2094131	T	0.1979	0.1951	C	0.002187	0.9627	1.018
1	rs3753242	2101843	C	0.3469	0.3902	T	0.3605	0.5482	0.8301
1	rs385039	2109571	G	0.2041	0.1463	A	1.018	0.3129	1.496
1	rs2292857	2138600	A	0.06122	0.06098	G	4.82e-005	0.9945	1.004
1	rs626479	2142422	A	0.2083	0.1585	G	0.7261	0.3941	1.397
1	rs262680	2199311	C	0.3438	0.4024	T	0.6529	0.4191	0.7778
1	rs16824948	2218382	T	0.09184	0.125	C	0.508	0.476	0.7079
1	rs12084736	2221742	T	0.3878	0.4146	C	0.1344	0.7139	0.8941
1	rs12045693	2237743	C	0.4082	0.4756	A	0.8247	0.3638	0.7604
1	rs2132303	2255420	T	0.2041	0.1098	C	2.939	0.08647	2.08
1	rs1496555	2266413	A	0.2292	0.122	G	3.448	0.06334	2.141
1	rs2645072	2312585	A	0.07143	0.122	C	1.332	0.2484	0.5538
1	rs7527871	2313888	C	0.4388	0.4024	A	0.2416	0.623	1.161
1	rs2840528	2316058	G	0.4255	0.4756	A	0.444	0.5052	0.8167

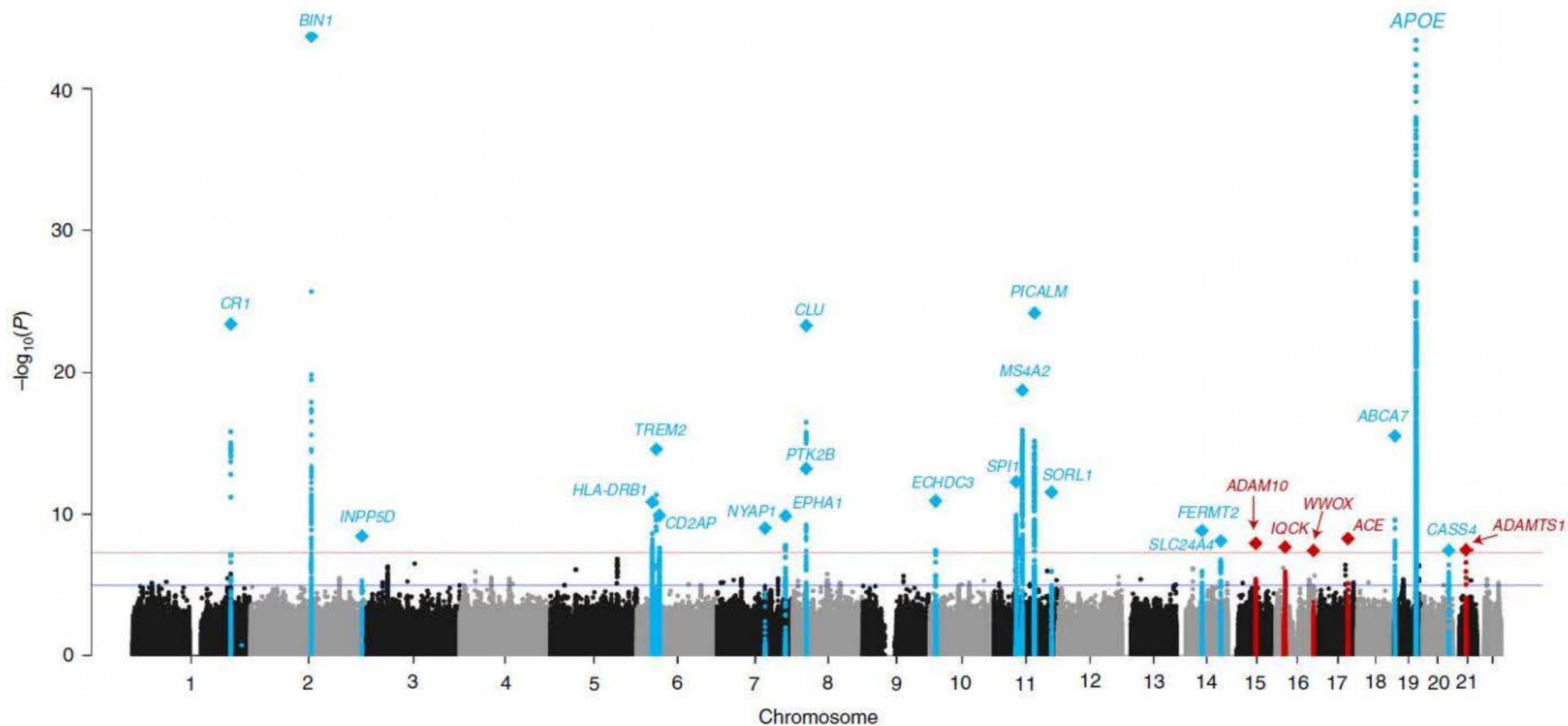
GWAS Summary Statistics





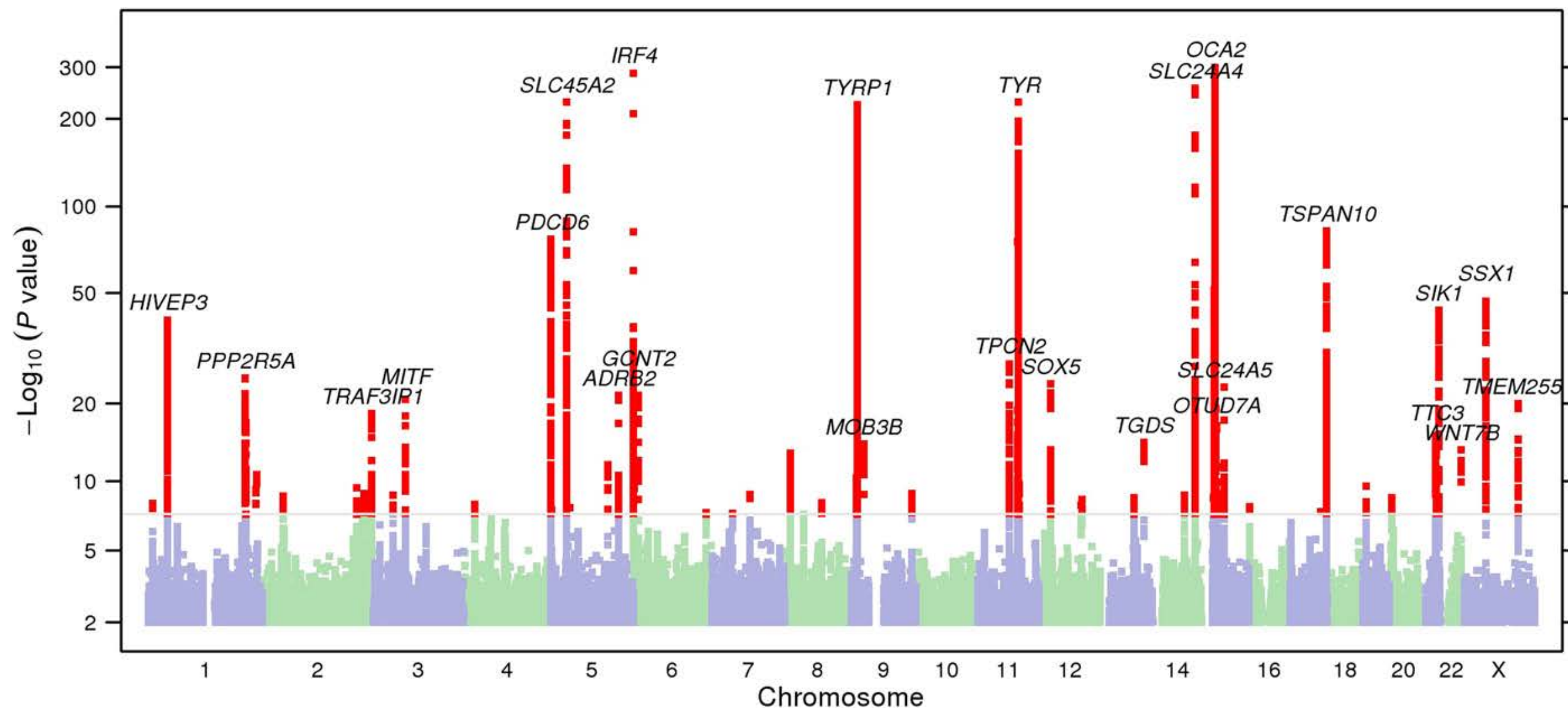
Ожирение

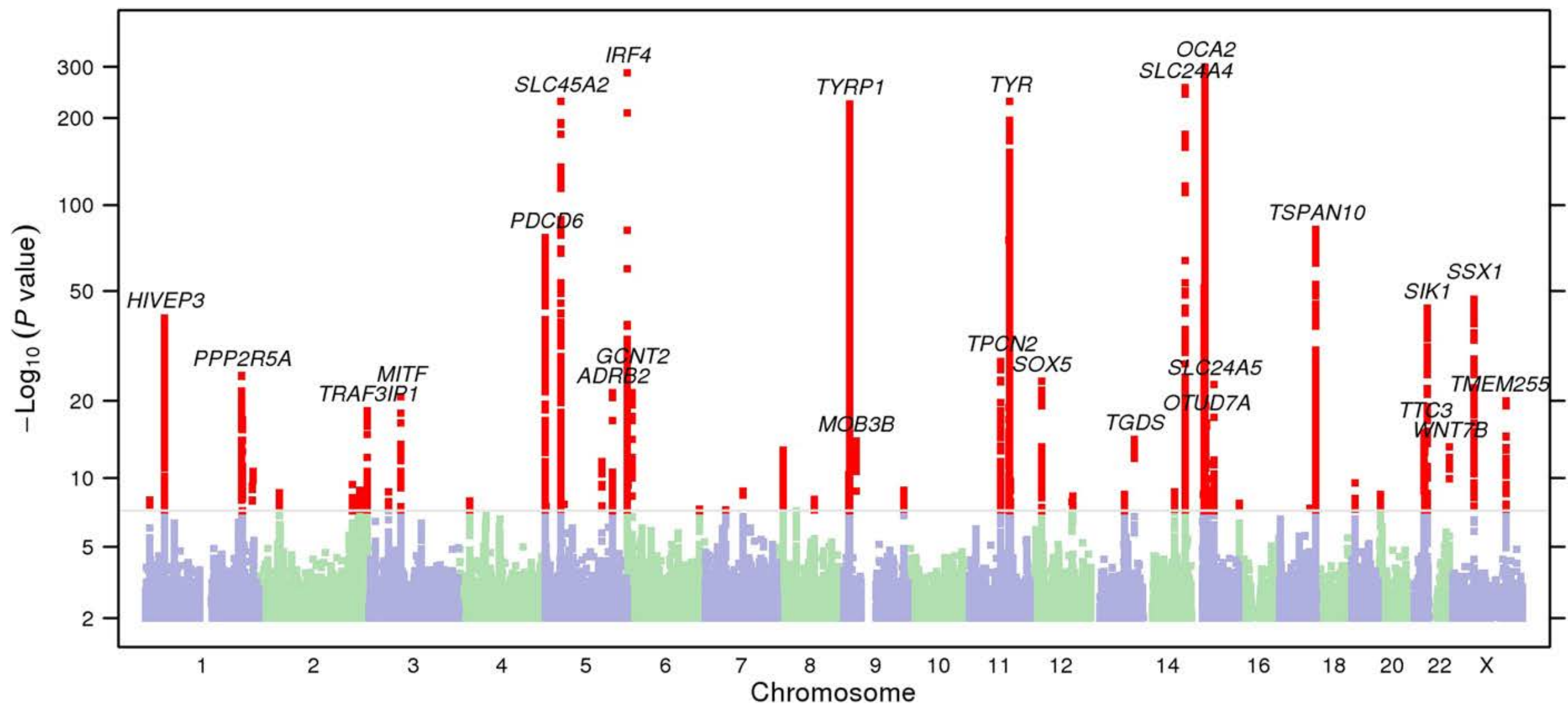




Болезнь Альцгеймера

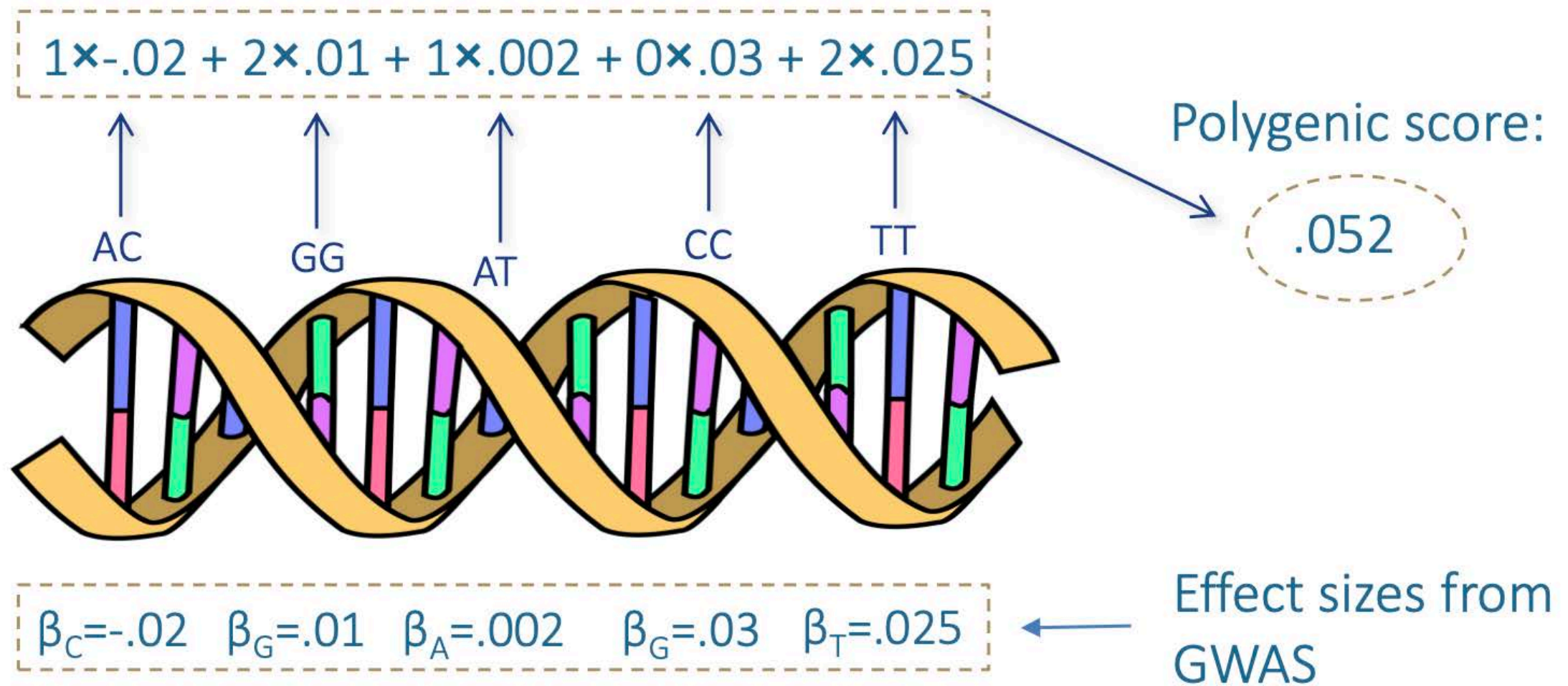
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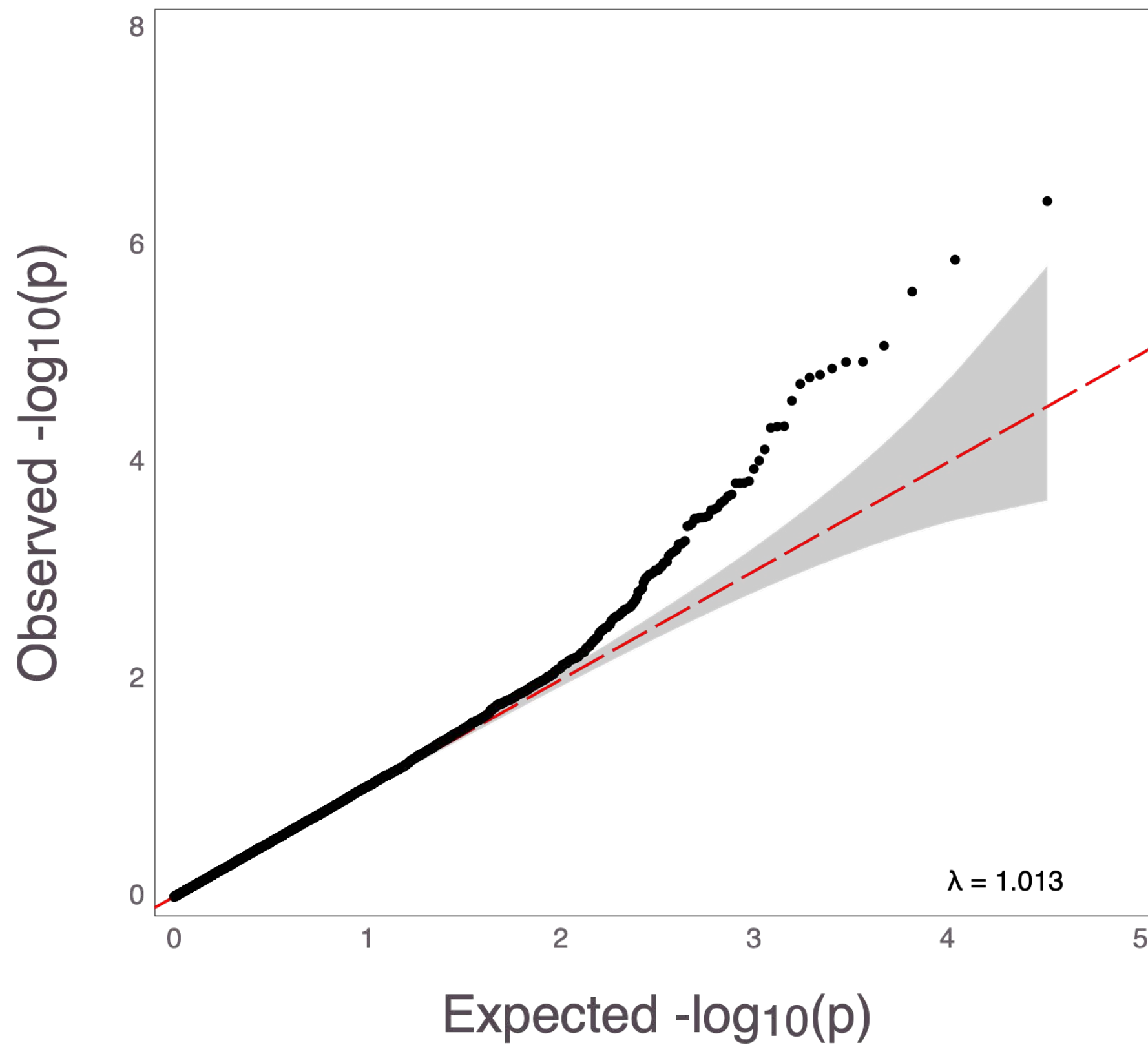


Цвет глаз

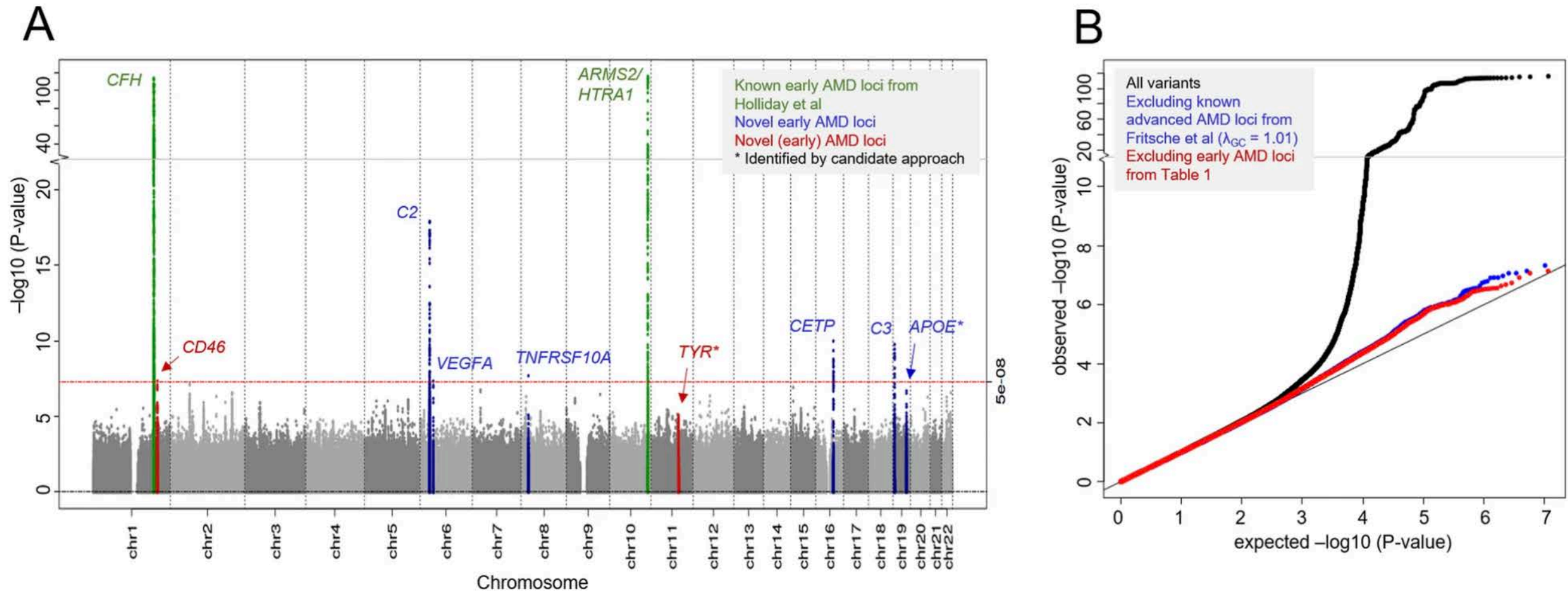
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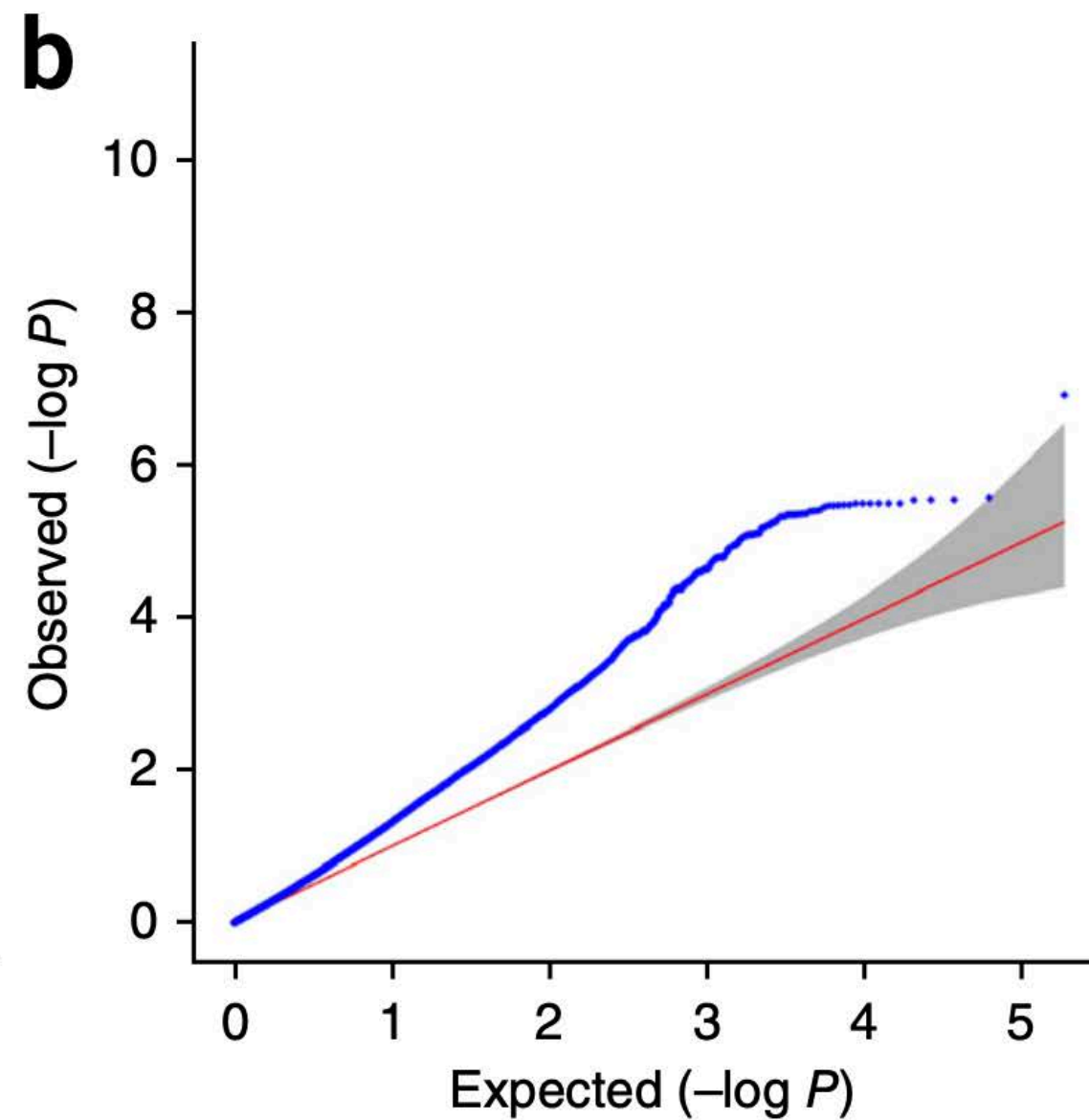
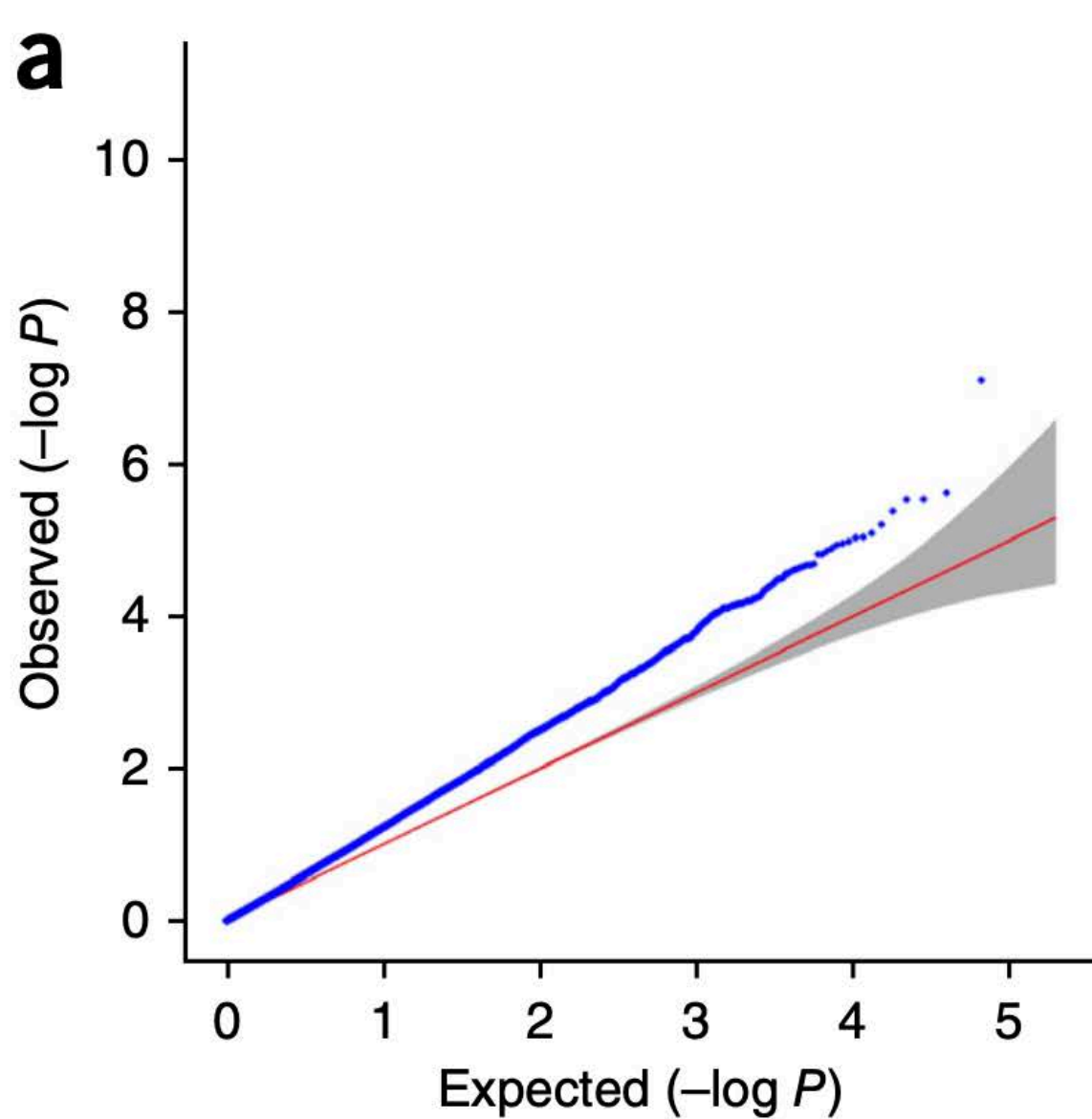
QQ Plot of GWAS p-values



From: [Genome-wide association meta-analysis for early age-related macular degeneration highlights novel loci and insights for advanced disease](#)



Early AMD meta-analysis. Shown are the association results of the meta-analysis for early AMD: **a** by their position on the genome (Manhattan plot) with color indicating whether the locus was previously identified by Holliday et al. [12] (blue), novel for early AMD (red), or among the other advanced AMD loci identified by Fritsche et al. [9] (green); and **b** their distribution (QQ plot)



На одном из QQ-графиков симуляция с популяционной стратификацией, а на другом с полигонной наследуемостью. Коэффициент инфляции в обоих случаях равен 1.32

LD Score regression distinguishes confounding from polygenicity in genome-wide association studies

Brendan K Bulik-Sullivan¹⁻³, Po-Ru Loh^{1,4}, Hilary K Finucane^{4,5}, Stephan Ripke^{2,3}, Jian Yang⁶, Schizophrenia Working Group of the Psychiatric Genomics Consortium⁷, Nick Patterson¹, Mark J Daly¹⁻³, Alkes L Price^{1,4,8} & Benjamin M Neale¹⁻³

Both polygenicity (many small genetic effects) and confounding biases, such as cryptic relatedness and population stratification, can yield an inflated distribution of test statistics in genome-wide association studies (GWAS). However, current methods cannot distinguish between inflation from a true polygenic signal and bias. We have developed an approach, LD Score regression, that quantifies the contribution of each by examining the relationship between test statistics and linkage disequilibrium (LD). The LD Score regression intercept can be used to estimate a more powerful and accurate correction factor than genomic control. We find strong evidence that polygenicity accounts for the majority of the inflation in test statistics in many GWAS of large sample size.

of this equation is provided in the **Supplementary Note**). This relationship holds for meta-analyses and also for ascertained studies of binary phenotypes, in which case h^2 is on the observed scale. Consequently, if we regress the χ^2 statistics from GWAS against LD Score (LD Score regression), the intercept minus one is an estimator of the mean contribution of confounding bias to the inflation in the test statistics.

RESULTS

Overview of methods

We estimated LD Scores from the European-ancestry samples in the 1000 Genomes Project⁷ (EUR) using an unbiased estimator⁸ of r^2 with 1-cM windows, singletons excluded (MAF > 0.13%) and no r^2 cutoff. Standard errors were estimated by jackknifing over blocks of

LD Estimation

Suppose we have two SNPs whose alleles are A/a and B/b .

The haplotype frequencies are:

Haplotype	Frequency
AB	p_{AB}
Ab	p_{Ab}
aB	p_{aB}
ab	p_{ab}

The allele frequencies are:

Allele	Frequency
A	$p_A = p_{AB} + p_{Ab}$
a	$p_a = p_{aB} + p_{ab}$
B	$p_B = p_{AB} + p_{aB}$
b	$p_b = p_{Ab} + p_{ab}$

D : the level of LD between A and B can be estimated using **coefficient of linkage disequilibrium (D)**, which is defined as:

$$D_{AB} = p_{AB} - p_A p_B$$

If A and B are in **linkage equilibrium**, we can get

$$D_{AB} = p_{AB} - p_A p_B = 0$$

which means the coefficient of linkage disequilibrium is 0 in this case.

D can be calculated for each pair of alleles and their relationships can be expressed as:

$$D_{AB} = -D_{Ab} = -D_{aB} = D_{ab}$$

So we can simply denote $D = D_{AB}$, and the relationship between haplotype frequencies and allele frequencies can be summarized in the following table.

Allele	A	a	Total
B	$p_{AB} = p_A p_B + D$	$p_{aB} = p_a p_B - D$	p_B
b	$p_{Ab} = p_A p_b - D$	$p_{ab} = p_a p_b + D$	p_b
Total	p_A	p_a	1



The range of possible values of D depends on the allele frequencies, which is not suitable for comparison between different pairs of alleles.

Lewontin suggested a method for the normalization of D :

$$D_{normalized} = \frac{D}{D_{max}}$$

where

$$D_{max} = \begin{cases} \max\{-p_A p_B, -(1-p_A)(1-p_B)\} & \text{when } D < 0 \\ \min\{p_A(1-p_B), p_B(1-p_A)\} & \text{when } D > 0 \end{cases}$$

It measures how much proportion of the haplotypes had undergone recombination.

In practice, the most commonly used alternative metric to $D_{normalized}$ is r^2 , the correlation coefficient, which can be obtained by:

$$r^2 = \frac{D^2}{p_A(1-p_A)p_B(1-p_B)}$$

Reference: Slatkin, M. (2008). Linkage disequilibrium—understanding the evolutionary past and mapping the medical future. *Nature Reviews Genetics*, 9(6), 477-485.

Welcome to **LDlink**

LDlink is a suite of web-based applications designed to easily and efficiently interrogate linkage disequilibrium in population groups. Each included application is specialized for querying and displaying unique aspects of linkage disequilibrium.

LDassoc

Interactively visualize association p-value results and linkage disequilibrium patterns for a genomic region of interest.

LDexpress

Search if a list of variants (or variants in LD with those variants) is associated with gene expression in multiple tissue types.

LDhap

Calculate population specific haplotype frequencies of all haplotypes observed for a list of query variants.

LDmatrix

Create an interactive heatmap matrix of pairwise linkage disequilibrium statistics.

LDpair

Investigate correlated alleles for a pair of variants in high LD.

LDpop

Investigate allele frequencies and linkage disequilibrium patterns across 1000G populations.

LDproxy

Interactively explore proxy and putatively functional variants for a query variant.

LDtrait

Search if a list of variants (or variants in LD with those variants) have previously been associated with a trait or disease.

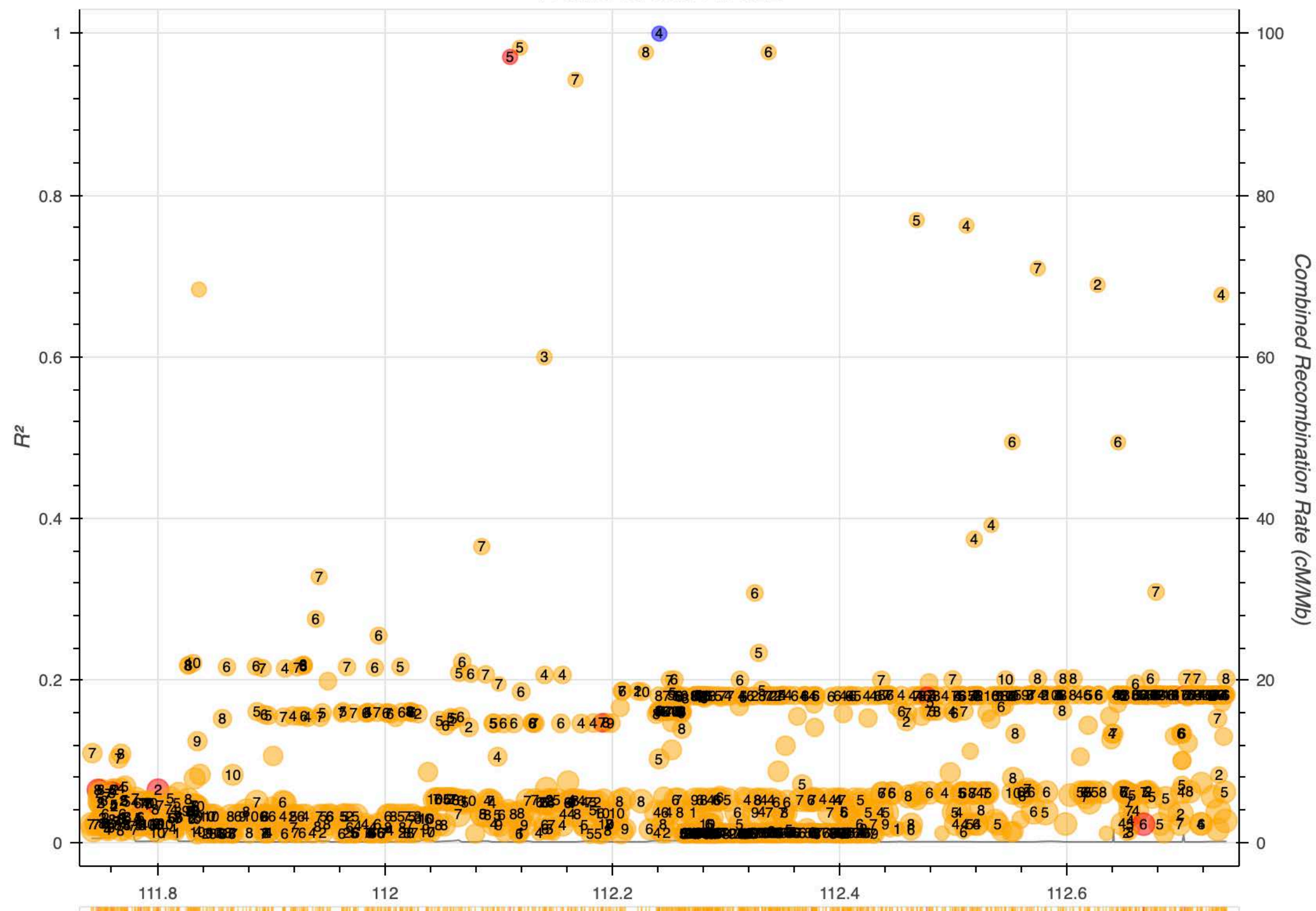
SNPchip

Find commercial genotyping platforms for variants.

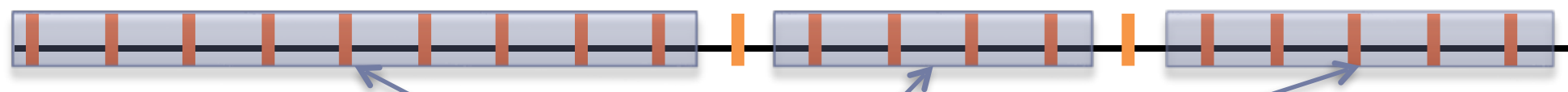
SNPclip

Prune a list of variants by linkage disequilibrium.

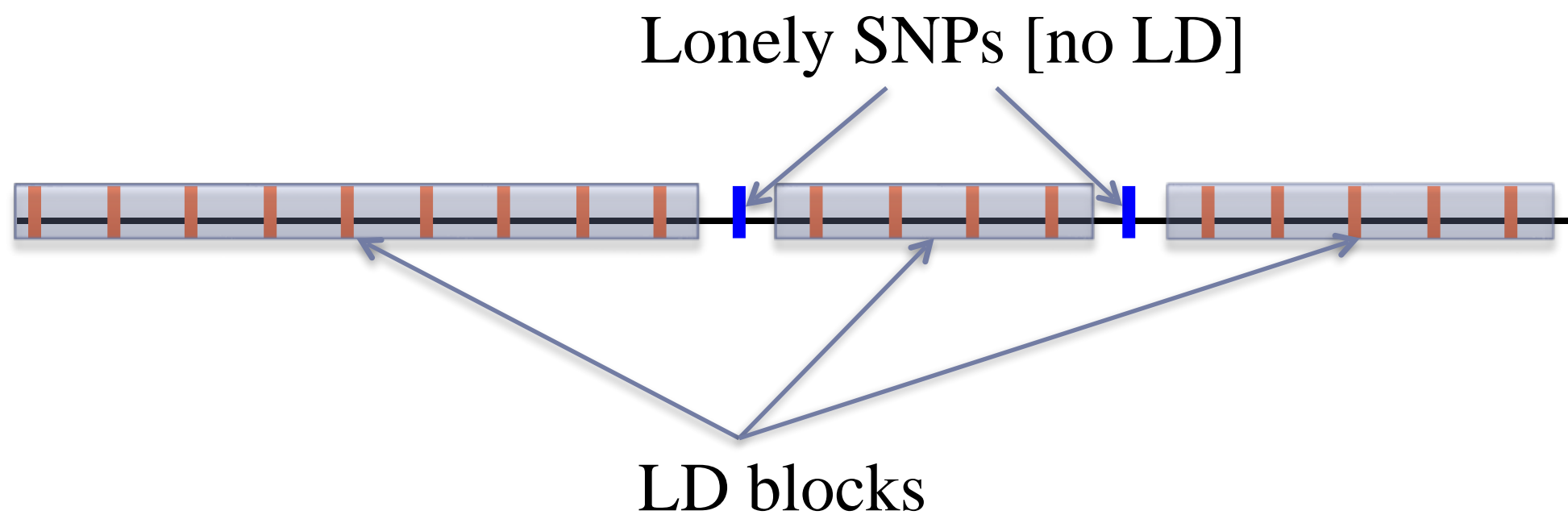
Proxies for rs671 in ALL







LD blocks



LD Block
Lonely SNP



- LD Block
- Lonely SNP
- Causal SNP



- LD Block
- Lonely SNP
- * Causal SNP



Sim 1

- LD Block
- Lonely SNP
- Causal SNP



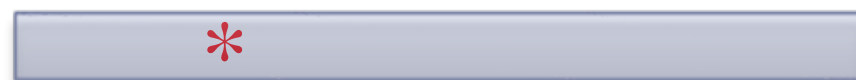
Sim 1



- LD Block
- Lonely SNP
- Causal SNP



Sim 1



- LD Block
- Lonely SNP
- Causal SNP



Sim 1



Sim 2



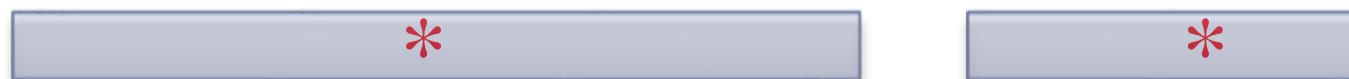
LD Block
Lonely SNP
Causal SNP



Sim 1



Sim 2



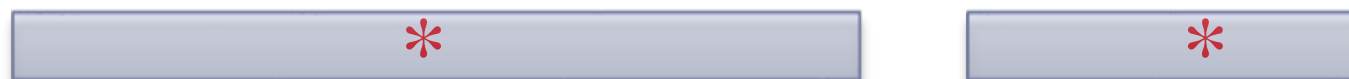
LD Block
Lonely SNP
Causal SNP



Sim 1



Sim 2



Sim 3



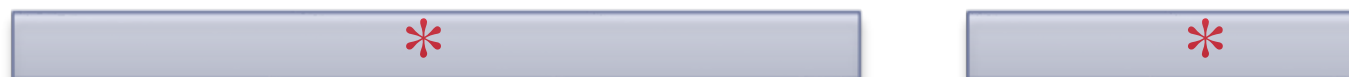
LD Block
Lonely SNP
Causal SNP



Sim 1

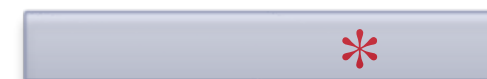


Sim 2

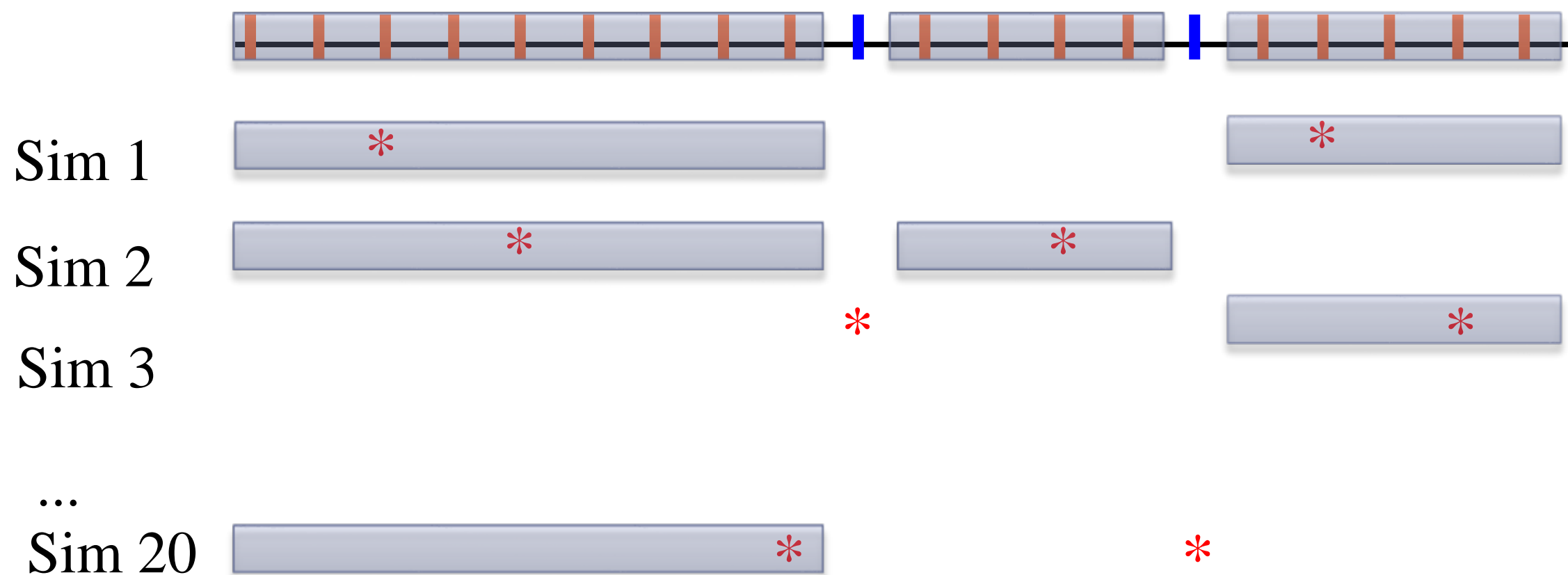


Sim 3

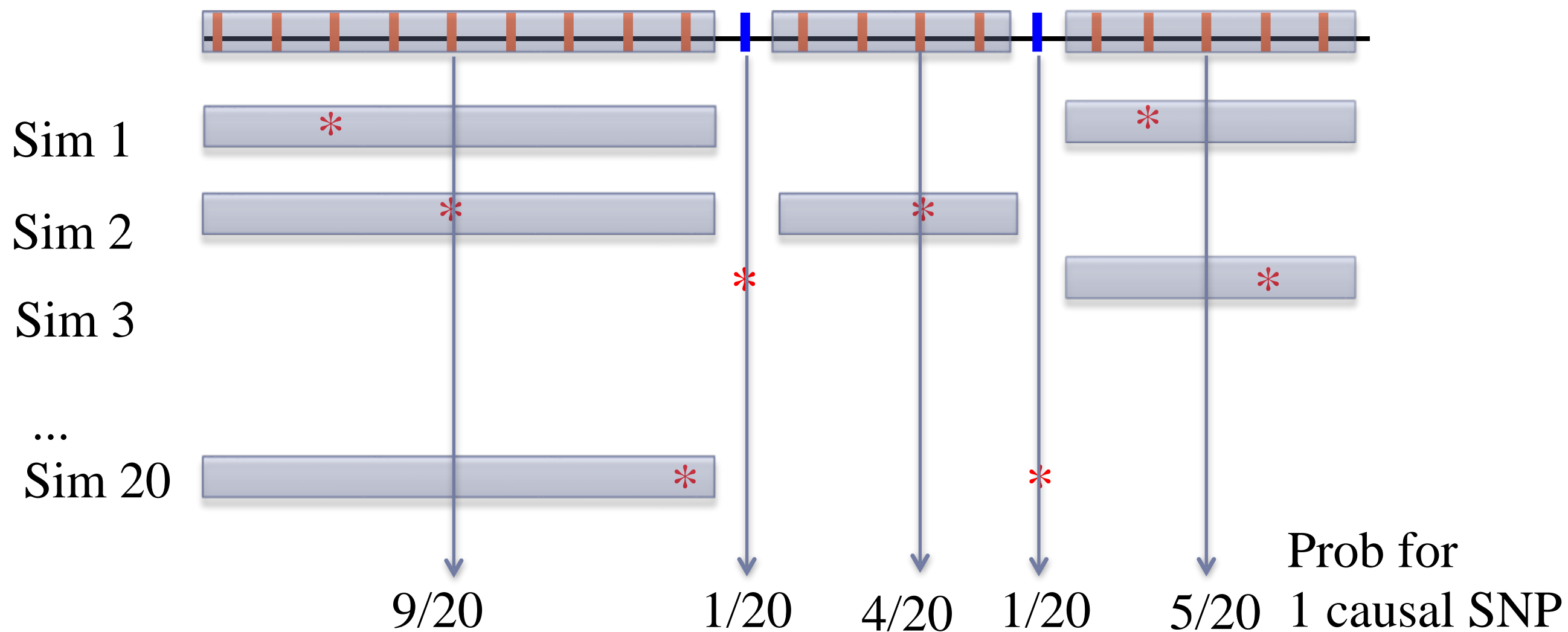
*



LD Block
Lonely SNP
Causal SNP



- LD Block
- Lonely SNP
- Causal SNP



Heritability

Heritability is a term used in genetics to describe how much phenotypic variation can be explained by genetic variation.

For any phenotype, its variation $Var(P)$ can be modeled as the combination of **genetic effects** $Var(G)$ and **environmental effects** $Var(E)$.

$$Var(P) = Var(G) + Var(E)$$

Broad-sense Heritability

The **broad-sense heritability** $H^2_{broad-sense}$ is mathematically defined as :

$$H^2_{broad-sense} = \frac{Var(G)}{Var(P)}$$

LD: Linkage disequilibrium

Linkage disequilibrium (LD) : non-random association of alleles at different loci in a given population. ([Wiki](#))

LD score

LD score l_j for a SNP j is defined as the sum of r^2 for the SNP and other SNPs in a region.

$$l_j = \sum_k r_{j,k}^2$$

LD score regression

Key idea: A variant will have higher test statistics if it is in LD with causal variant, and the elevation is proportional to the correlation (r^2) with the causal variant.

$$E[\chi^2 | l_j] = \frac{Nh^2 l_j}{M} + Na + 1$$

- N : sample size.
- M : number of SNPs.
- h^2 : observed-scale heritability
- a : the effect of confounding factors, including cryptic relatedness and population stratification.

For more details of LD score regression, please refer to : - Bulik-Sullivan, Brendan K., et al. "LD Score regression distinguishes confounding from polygenicity in genome-wide association studies." *Nature genetics* 47.3 (2015): 291-295.

In words:

Test statistic = average causal effect per SNP * LD score + inflation due to population stratification + 1

$$E(\chi^2 | l_j) = \frac{Nh^2}{M} l_j + Na + 1$$

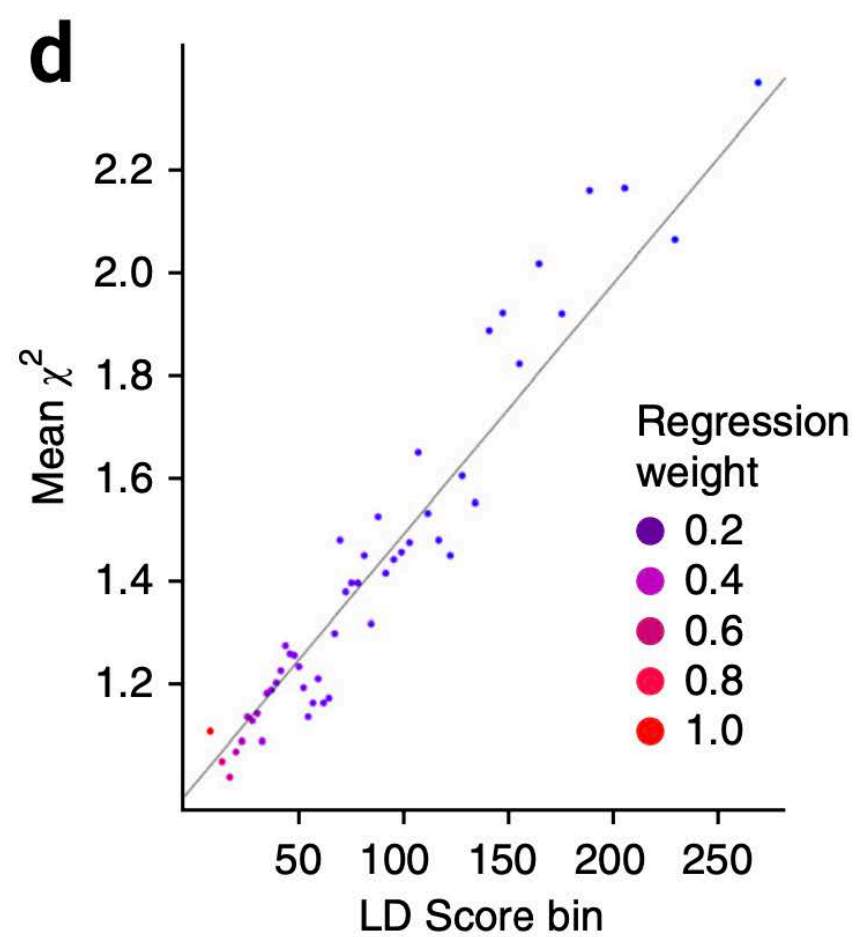
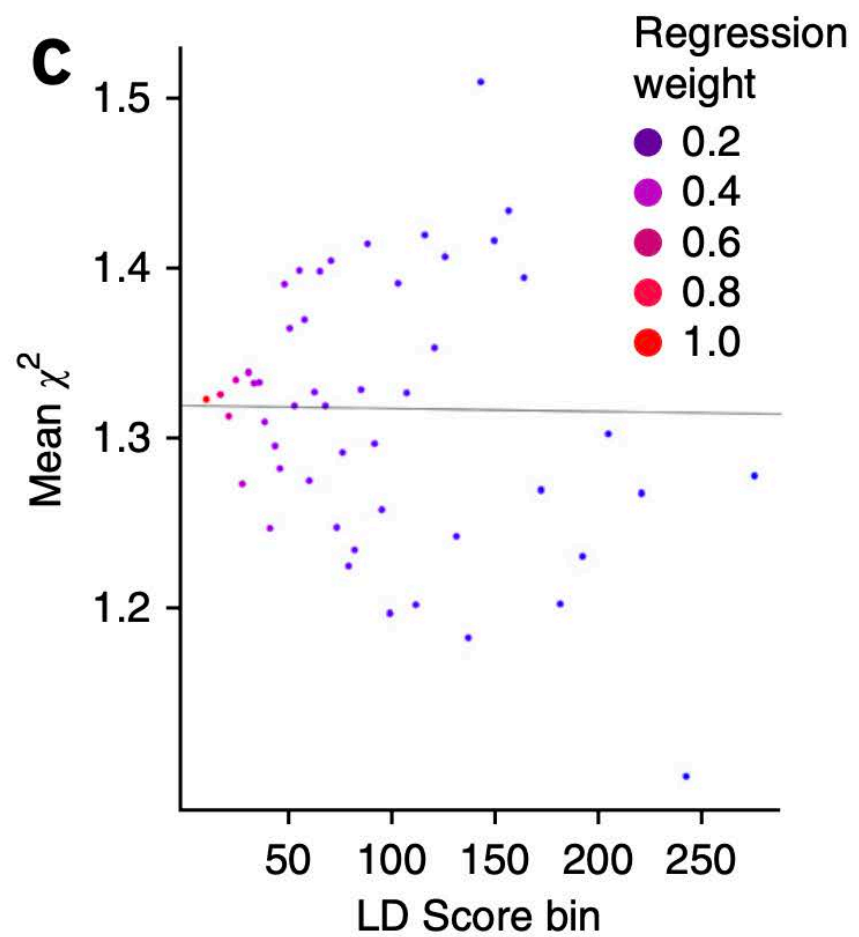
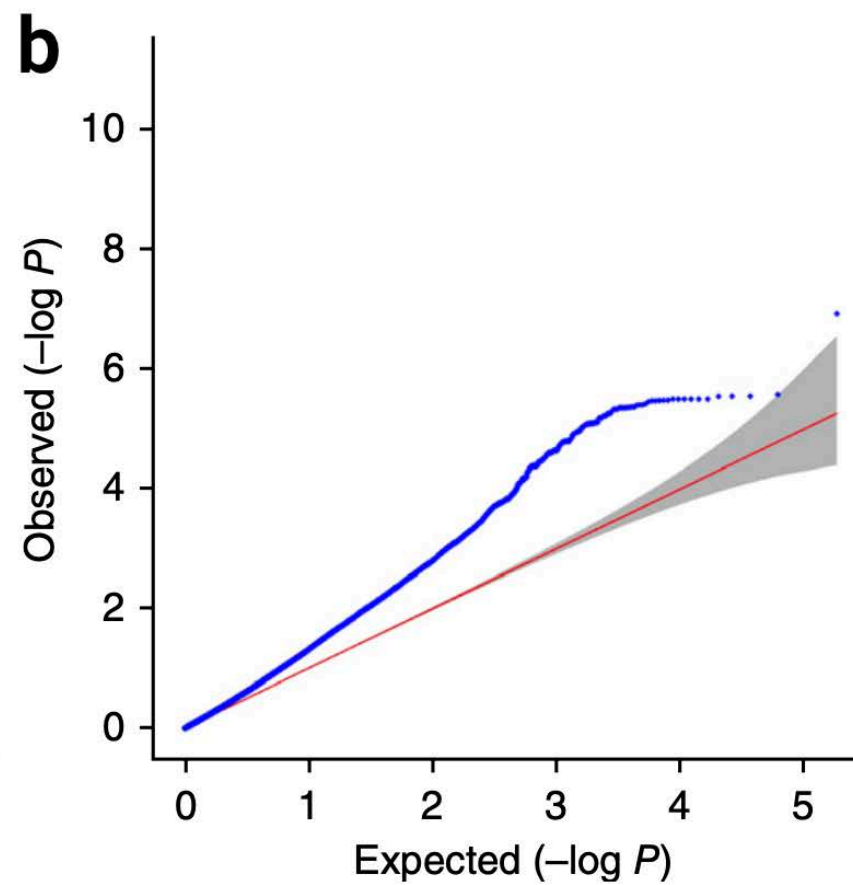
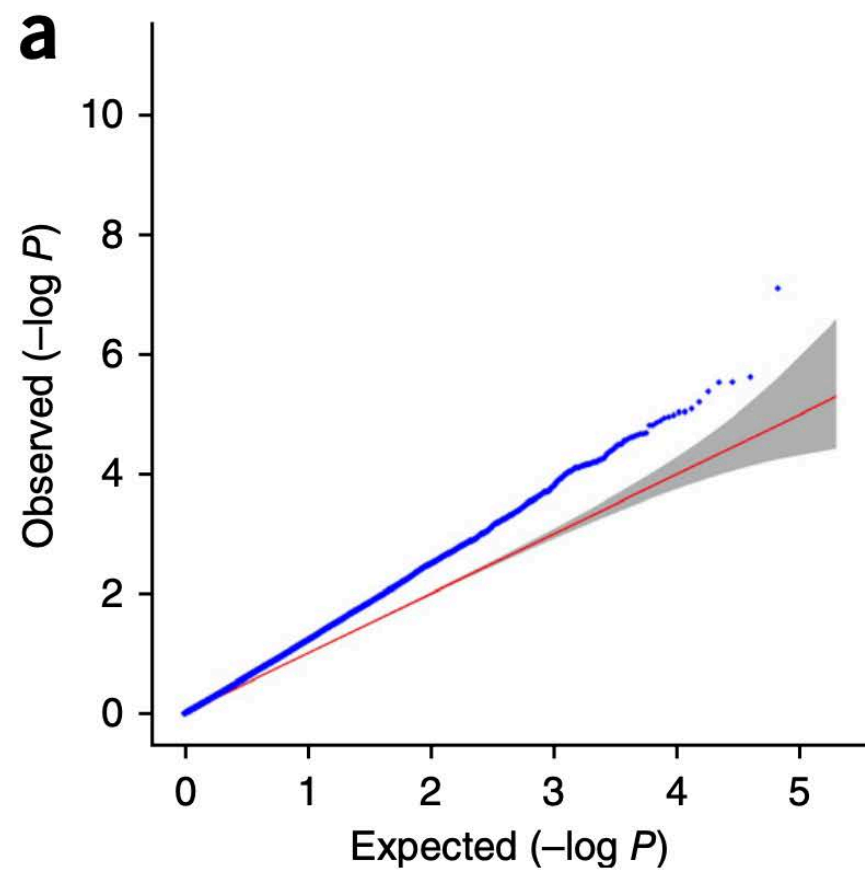
$$l_j = \sum_k r_{jk}^2 \quad (\text{LD Score})$$

N = sample size

M = number of SNPs

h^2 / M = average heritability per SNP

a = Population structure /
cryptic relatedness



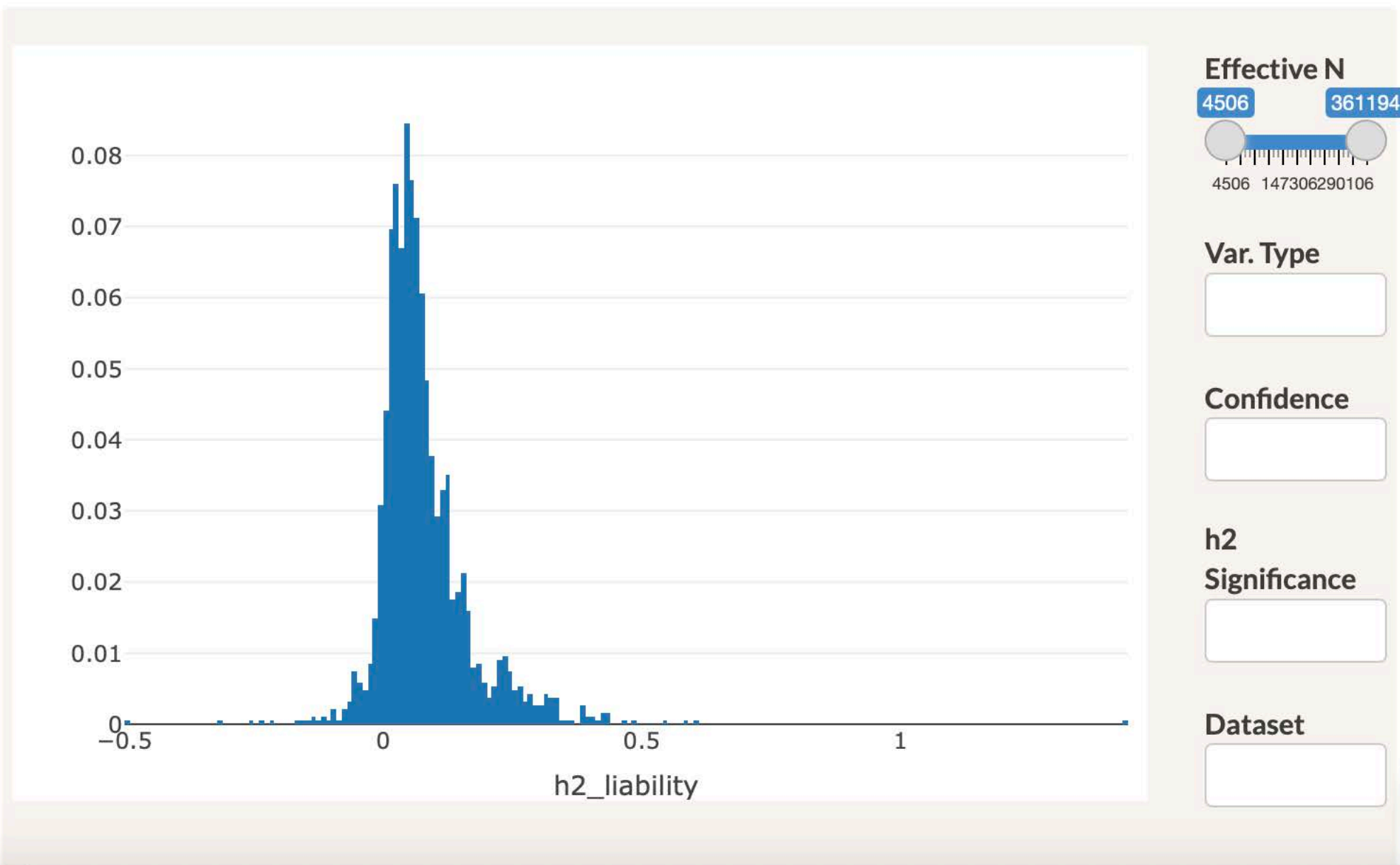
SNP Heritability

The average SNP heritability estimate across all 4178 primary GWAS is 0.065 (median = 0.058). Average estimates increase when restricting to any **confidence** (mean = 0.078, median = 0.059) or **high confidence** (mean = 0.088, median = 0.062). Strong departure from the null hypothesis of $h_g^2 = 0$ is observed across phenotypes, especially among high confidence results.

SNP Heritability Estimates

SNP Heritability p-values

p-values by Confidence



UKB SNP-Heritability Browser

Results from the [Neale Lab](#)

Last updated 2022-10-11

Select Columns ▾

Search:

Phenotype	h2	h2 p	h2 sig?	Confidence	Int.	Int. p	Neff	Source
Leg fat percentage (left)	0.221	8.97e-262	z7	high	1.109	8.73e-11	354,791	phesant
Leg fat percentage (right)	0.221	8.89e-261	z7	high	1.103	6.57e-10	354,811	phesant
Body fat percentage	0.230	1.4e-249	z7	high	1.095	2.25e-08	354,628	phesant
Arm fat percentage (left)	0.225	1.33e-240	z7	high	1.094	5.43e-08	354,707	phesant
Trunk fat percentage	0.221	2.96e-233	z7	high	1.087	2.24e-07	354,619	phesant
Arm fat percentage (right)	0.222	3.41e-227	z7	high	1.100	1.33e-08	354,760	phesant
Impedance of arm (right)	0.248	2.43e-209	z7	high	1.084	5.98e-05	354,792	phesant
Leg fat mass (left)	0.234	4.75e-208	z7	high	1.095	4.41e-07	354,788	phesant
Qualifications: College or University degree	0.286	1.54e-207	z7	high	1.119	5.59e-15	313,437	phesant
Leg fat mass (right)	0.232	1.03e-203	z7	high	1.095	5.15e-07	354,807	phesant
Impedance of arm (left)	0.244	2.09e-198	z7	high	1.099	2.99e-06	354,807	phesant
Trunk fat mass	0.239	8.95e-198	z7	high	1.094	9.4e-07	354,597	phesant
Whole body fat mass	0.239	9.55e-197	z7	high	1.096	7.6e-07	354,244	phesant
Body mass index (BMI)	0.249	2.52e-194	z7	high	1.103	6.28e-07	354,831	phesant

An atlas of genetic correlations across human diseases and traits

Brendan Bulik-Sullivan^{1-3,9}, Hilary K Finucane^{4,9}, Verner Anttila¹⁻³, Alexander Gusev^{5,6}, Felix R Day⁷, Po-Ru Loh^{1,5}, ReproGen Consortium⁸, Psychiatric Genomics Consortium⁸, Genetic Consortium for Anorexia Nervosa of the Wellcome Trust Case Control Consortium 3⁸, Laramie Duncan¹⁻³, John R B Perry⁷, Nick Patterson¹, Elise B Robinson¹⁻³, Mark J Daly¹⁻³, Alkes L Price^{1,5,6,10} & Benjamin M Neale^{1-3,10}

Identifying genetic correlations between complex traits and diseases can provide useful etiological insights and help prioritize likely causal relationships. The major challenges preventing estimation of genetic correlation from genome-wide association study (GWAS) data with current methods are the lack of availability of individual-level genotype data and widespread sample overlap among meta-analyses. We circumvent these difficulties by introducing a technique—cross-trait LD Score regression—for estimating genetic correlation that requires only GWAS summary statistics and is not biased by sample overlap. We use this method to estimate 276 genetic correlations among 24 traits. The results include genetic correlations between anorexia nervosa and schizophrenia, anorexia and obesity, and educational attainment and several diseases. These results highlight the power of genome-wide analyses, as there currently are no significantly associated SNPs for anorexia nervosa and only three for educational attainment.

inferences from such studies can be challenging because of issues such as confounding and reverse causation, which can lead to spurious associations and mask the effects of real risk factors^{1,2}. Genetics can help elucidate cause and effect, as inherited genetic risks cannot be subject to reverse causation and are correlated with a smaller list of confounders.

The first methods to test for genetic overlap were family studies³⁻⁷. To estimate the genetic overlap for many pairs of phenotypes, family study designs require the measurement of multiple traits for the same individuals. Consequently, it is challenging to scale these designs to a large number of traits, especially traits that are difficult or costly to measure (for example, low-prevalence diseases). More recently, GWAS have allowed effect size estimates to be obtained for specific genetic variants, so it is possible to test for shared genetics by looking for correlations in effect sizes across traits, which does not require measuring multiple traits per individual.

There exists a large class of methods for interrogating genetic overlap via GWAS that focus only on genome-wide significant SNPs. One of the most influential methods in this class is Mendelian randomiza-

Cross-trait LD score regression

Cross-trait LD score regression is employed to estimate the genetic correlation between a pair of traits.

Key idea: replace χ^2 in univariate LD score regression and the relationship (SNPs with high LD) still holds.

$$E[z_{1j}z_{2j}] = \frac{\sqrt{N_1N_2}\rho_g}{M}l_j + \frac{\rho N_s}{\sqrt{N_1N_2}}$$

- z_{ij} : z score of trait i for SNP j
- N_i : sample size of trait i
- ρ : phenotypic correlation
- ρ_g : genetic covariance
- l_j : LD score
- M : number of SNPs

Then we can get the genetic correlation by :

$$r_g = \frac{\rho_g}{\sqrt{h_1^2 h_2^2}}$$

- Reference: Bulik-Sullivan, Brendan, et al. "An atlas of genetic correlations across human diseases and traits." Nature genetics 47.11 (2015): 1236-1241.
-

