Chapter 4

4.1. Intro Markov Chains Chain -> Previous state 10 Bors depend कार आकारते state predict कता,

Stochastic Process:

- A collection of Random Variables.

time 10 PAZ define sign 201

- \(\lambda\((t)\), t \(\tilde{T}\) \rightarrow A state of the process at time t.

Markov Chain

Markov chain is an stochastion process {Xn, n=1,2,3,--} such that whenever the process is in state i, there is a fixed probability Pij that it will next be in state j.

$$P_{ij} = P_{0} X_{n+1} = j | X_{n} = i_{n} X_{n-1} = i_{n-1}, ..., X_{i} = i_{i}, X_{n-1} = i_{n-1}, ..., X_{i} = i_{i}, X_{n-1} = i_{n-1}, ..., X_{i} = i_{n}, X_{i} = i$$

The Pij is independent of past states and depends only on the present state.

$$\hat{O} \geq P_{ij} = 1$$
, $i = 0, 1, ...$

at each step, there must be a transition

Transition Matrix

$$P = \begin{cases} P_{00} & P_{01} & P_{02} \\ P_{10} & P_{11} & P_{12} \\ P_{20} & P_{21} & P_{22} \\ P_{10} & P_{11} & P_{22} \end{cases}$$

Transition 20 # column)

Chapman-Kolmogorov Equation State (1) R nsteps (n+k) State (j) One step transition probability Pi = n step transition probability that a process in state will reach state j, after n transitions. $P_{ij}^{n} = P_{ij} \times X_{k} = i$ $N_{ij} = i$ $P_{ij}^{(n+m)} = \underbrace{\begin{cases} P_{ik} & P_{kj} \\ Y_{i,j} \end{cases}}_{n} for \forall n, m \geqslant 0$ Pij -> i ma j to (n+m) steps IT SOTA probability

Proof:

$$P_{ij}^{(n+m)} = P_{ij}^{(n+m)} = \int_{X_{n+m}} |X_{n}| = \int_{X_{n}} |X_{n}| = \int_{X_{n}$$

$$= \sum_{k=0}^{\infty} p_{0} \times n_{+m} = j \times n_{-k} p_{0} \times n_{-k} \times k \times n_{-k} = i$$

$$\frac{(m+n)}{p_{ij}} = \sum_{k=0}^{\infty} \frac{m}{p_{kj}} \frac{n}{p_{ik}} (proved)$$

The n step transition // Basically,

probability may be obtained 11 pm+n = pmpn

by multiplying the matrix // p2 STE

by itself n times.

$$p(n+m) = p(n)p(m)$$

n steps m steps

VK

Example: 4.8

Consider example of the weather 14-1)

 $d = 0.7, \beta = 0.4.$

12 - C. MIN 19 (110) A

Calculate the probability that it

will rain 4 days from today,

given it is raining today.

Solve R

$$-\beta$$
 $=$ $0-4$

$$p = p \cdot p$$

$$= \begin{bmatrix} 0.7 & 0.3 \\ 0.4 & 0.6 \end{bmatrix} \begin{bmatrix} 0.7 & 0.3 \\ 0.4 & 0.6 \end{bmatrix}$$

$$= \begin{bmatrix} 0.61 & 0.39 \\ 0.52 & 0.48 \end{bmatrix}$$

$$(4) = p^{(2)} p^{(2)} = \begin{bmatrix} 0.61 & 0.39 \\ 0.52 & 0.48 \end{bmatrix} \begin{bmatrix} 0.61 & 0.39 \\ 0.52 & 0.48 \end{bmatrix}$$

$$= \begin{bmatrix} 0.5749 & 0-4251 \\ 0.5668 & 0.4332 \end{bmatrix}$$

there, all state communicates with each other, only 1 class.

This possible to go to 2 from 0.

Here, 20, 16 -> Recurrent State State 2 and [0,1 \D Jaar Corral of all are not on \D Jara [0,1 \D Z \ Jara] 2) -> Transient State all states are [250 state correlong accessible fromit term a ros 2002] \$ 233 -> Absorbing State (no other state [long term o country)
is accessible from forty terminate notal
it) 50,3 Classes -> 20,13, 23, 23, 23,

Example [4-4]

TW RR RS SR SS

MT

RR 0-7 0.3 0 0

RS 0 0 0-4 0-6

SR 0.5 0-5 0 0

SS 0 0 0.2 0-8]

Mt -> Monday and thesday

Tw -> Thesday and Wednesday

MT CORR means, Monday an Tuesday.

Rain Given.

Given RR on Monday, What is the probability of Raining on Thursday?

Solve MT RR RS SRE SS

RR 0.49 0.21 0.12 0.18

RR 0.49 0.21 0.12 0.48

P = RS 0.35 0.15 0.20 0.30

SR 0.35 0.15 0.20 0.64

 $P_{00} + P_{02}^{(2)} = 0.49 + 0.12 = 0.61$ (Ans)

4.4. Limiting Probability Theorem:

If lim Pij exists and is independent

of oi and $\pi_j = \lim_{n \to \infty} P_{ij}^n$, $j \ge 0$

then, Ti is the unique non-negative

solution of:

 $\pi_{ij} = \sum_{i=0}^{pq} \pi_i P_{ij}$, j > 0

1 = 1 = 1 = 1 = 1

Tij is limiting probability of a column

Example:
$$4.22$$
 R

$$P = \begin{cases} \beta & 1-\beta \\ \beta & 1-\beta \end{cases}$$

State 0 if it rains tomorrow.

If 1 is doesn't rain is

Long ran proportions π , and π , will be,

$$\pi_0 = d\pi_0 + \beta \pi_1$$

$$\pi_1 = (1-d)\pi_0 + (1-\beta)\pi_1$$

$$\pi_0 + \pi_1 = 1$$

So, $\pi_0 \Rightarrow = d\pi_0 + \beta (1-\pi_0)$

$$\Rightarrow \pi_0 - d\pi_0 + \beta \pi_0 = \beta$$

$$\therefore \pi_0 = \frac{\beta}{1-d+\beta} \Rightarrow \text{Probability of Rainy in the long run}$$

$$\pi_1 = \frac{1-d}{1-d+\beta} \Rightarrow \text{In the long run}$$

2 properties of Markov Chain

- O Period -> State (i) is said to have period of d if Pii = 0, whenever n is not divisible by d and d is the largest integer with this property
- (1) A state with period 1 is said to be aperiodic.

La delidadade de de

with the second second second

1 1 - 1 - 1

Periodic chain doesn't have a limiting probability.

Example: 4.25 Markov Chain in Grenefics Hardy-Weinberg Law: 2 types of genes {A, a} => Each individual possesses a particular pair of genes => AA, aa or Aa.

> PfAA3 = Po, PfAa3 = ro, Pfaa3 = 20

> During mating, each individuals contribute one type of his/her genes at random.

> Mating occurs at random.

Let, next generation probability,

PdAA3 = P, PdAa3 = r, Plaa3 = 2

Now, A randomly choosen gene will be type A with probability,

 $P\{A\} = P\{A|AA\}P_0 + P\{A|aa\}P_0$ $+ P\{A|Aa\}P_0 + P\{A|Aa\}P$

Now,
$$p = P \left\{AA\right\} = \left(P_0 + \frac{r_0}{2}\right)^2$$

$$Q = p daab = \left(20 + \frac{r_0}{2}\right)^{\frac{1}{2}}$$

$$V^{\circ} = P_{Q} A a^{3} = {}^{2}C_{1} \left(P_{0} + \frac{V_{0}}{2!}\right) \left(q_{0} + \frac{V_{0}}{2!}\right)$$

$$P+q+r = \left(P_0 + \frac{V_0}{2}\right)^2 + 2\left(P_0 + \frac{V_0}{2}\right)\left(2_0 + \frac{V_0}{2}\right)$$

$$+\left(Q_{-0}+\frac{V_{0}}{2}\right)^{2}$$

$$=\left(P_0+\frac{\gamma_0}{2}+Q_0+\frac{\gamma_0}{2}\right)$$

$$= (P_0 + q_0 + r_0)^2 = 1^{2/2}$$

1- + 3 = EALY

$$P + \frac{Y}{2} = \left(P_0 + \frac{Y_0}{2}\right)^2 + \left(P_0 + \frac{Y_0}{2}\right)\left(Q_0 + \frac{Y_0}{2}\right)$$

$$= \left(P_0 + \frac{Y_0}{2}\right)\left(P_0 + \frac{Y_0}{2} + Q_0 + \frac{Y_0}{2}\right)$$

$$= \left(P_0 + \frac{Y_0}{2}\right)\left(P_0 + Q_0 + Y_0\right)$$

$$= \left(P_0 + \frac{Y_0}{2}\right)\left(V_{\text{nchanged}}\right)$$

Under random mating, in all successive generations after the initial one, the percentages of the population having gene pairs AA, aa, Aa will remain same at the yalnes p, a and r.

[Hardy-Weinberg Law]

Sud to a solic of

4.5 Some Applications The Grambler's Ruin Problem P (winning one unit) = P P (losing one unit) = q) game play What is the probability that, starting with i units, the gambler's fortune will reach N before reaching 0) Goal -> Random model, STATESTO BITATET लाउ (रामाक भावंद्य मा। -> 0 TCD (TCT -) I fact som a signi -> N -> DI JOIO propapility too sign

200

Let, P; = Probability that starting from i, the Gambler will eventually reach N. Base Case: Po = 0, PN = 1 P: = pPi+1 + 2Pi-1 ⇒ (P+9) P; = PPi+1 + 9Pi-1 // P+9=1 => Pi+1 - P; = = (P; - Pi-1) So, $P_2 - P_1 = \frac{q}{p} (P_1 - P_0)$ $P_2 - P_1 = \frac{2}{P} \left(P_i \right)$ $P_3 - P_2 = \frac{q}{p} \left(P_2 - P_i \right) = \left(\frac{q}{p} \right)^p P_i$ $Pa-P_3=\frac{(q_1)}{p}P_1$ $P_i - P_{i-1} = \left(\frac{q}{P}\right)^{i-1}P_i$ $P_{1} - P_{1} = P_{1} \left[\frac{2}{P} + \left(\frac{2}{P} \right)^{2} + \left(\frac{2}{P} \right)^{2} + \cdots + \left(\frac{2}{P} \right)^{2} \right]$

$$P_{i} = \int_{0}^{\infty} \frac{1 - \left(\frac{q_{i}}{p}\right)^{2}}{1 - \frac{q_{i}}{p}} P_{i}, \text{ if } \frac{q_{i}}{q_{i}} \neq 1$$

$$P_{i} = \int_{0}^{\infty} \frac{1 - \left(\frac{q_{i}}{p}\right)^{N}}{1 - \frac{q_{i}}{p}}, \text{ p } \neq \frac{1}{2}$$

$$P_{i} = \int_{0}^{\infty} \frac{1 - \left(\frac{q_{i}}{p}\right)^{N}}{1 - \frac{q_{i}}{p}}, \text{ p } \neq \frac{1}{2}$$

$$P_{i} = \int_{0}^{\infty} \frac{1 - \left(\frac{q_{i}}{p}\right)^{N}}{1 - \left(\frac{q_{i}}{p}\right)^{N}} P_{i} = 1$$

$$P_{i} = \int_{0}^{\infty} \frac{1 - \left(\frac{q_{i}}{p}\right)^{N}}{1 - \left(\frac{q_{i}}{p}\right)^{N}} P_{i} = \frac{1}{2}$$

$$P_{i} = \int_{0}^{\infty} \frac{1 - \left(\frac{q_{i}}{p}\right)^{N}}{1 - \left(\frac{q_{i}}{p}\right)^{N}} P_{i} = \frac{1}{2}$$

$$P_{i} = \int_{0}^{\infty} \frac{1 - \left(\frac{q_{i}}{p}\right)^{N}}{1 - \left(\frac{q_{i}}{p}\right)^{N}} P_{i} = \frac{1}{2}$$

$$P_{1} = \begin{cases} \frac{1 - 2p}{1 - 2p} & ; & p \neq \frac{1}{2} \\ \frac{1}{N} & ; & p = \frac{1}{2} \end{cases}$$

$$P_{1} = \begin{cases} \frac{1 - 2p}{p} \times \frac{1 - 2p}{1 - 2p} & ; & p \neq \frac{1}{2} \\ \frac{1}{N} \times \frac{1}{N} & ; & p = \frac{1}{2} \end{cases}$$

$$P_{2} = \begin{cases} \frac{1 - 2p}{p} \times \frac{1 - 2p}{p} & ; & p \neq \frac{1}{2} \\ \frac{1 - 2p}{p} \times \frac{1 - 2p}{p} & ; & p = \frac{1}{2} \end{cases}$$

$$P_{3} = \begin{cases} \frac{1 - 2p}{p} \times \frac{1 - 2p}{p} & ; & p = \frac{1}{2} \\ \frac{1 - 2p}{p} \times \frac{1 - 2p}{p} & ; & p = \frac{1}{2} \end{cases}$$

$$P_{4} = \begin{cases} \frac{1 - 2p}{p} \times \frac{1 - 2p}{p} & ; & p = \frac{1}{2} \\ \frac{1 - 2p}{p} \times \frac{1 - 2p}{p} & ; & p = \frac{1}{2} \end{cases}$$

$$P_{5} = \begin{cases} \frac{1 - 2p}{p} \times \frac{1 - 2p}{p} & ; & p = \frac{1}{2} \\ \frac{1 - 2p}{p} \times \frac{1 - 2p}{p} & ; & p = \frac{1}{2} \end{cases}$$

$$P_{7} = \begin{cases} \frac{1 - 2p}{p} \times \frac{1 - 2p}{p} & ; & p = \frac{1}{2} \\ \frac{1 - 2p}{p} \times \frac{1 - 2p}{p} & ; & p = \frac{1}{2} \end{cases}$$

$$P_{7} = \begin{cases} \frac{1 - 2p}{p} \times \frac{1 - 2p}{p} & ; & p = \frac{1}{2} \end{cases}$$

$$P_{7} = \begin{cases} \frac{1 - 2p}{p} \times \frac{1 - 2p}{p} & ; & p = \frac{1}{2} \end{cases}$$

$$P_{7} = \begin{cases} \frac{1 - 2p}{p} \times \frac{1 - 2p}{p} & ; & p = \frac{1}{2} \end{cases}$$

$$P_{7} = \begin{cases} \frac{1 - 2p}{p} \times \frac{1 - 2p}{p} & ; & p = \frac{1}{2} \end{cases}$$

$$P_{7} = \begin{cases} \frac{1 - 2p}{p} \times \frac{1 - 2p}{p} & ; & p = \frac{1}{2} \end{cases}$$

$$P_{7} = \begin{cases} \frac{1 - 2p}{p} \times \frac{1 - 2p}{p} & ; & p = \frac{1}{2} \end{cases}$$

$$P_{7} = \begin{cases} \frac{1 - 2p}{p} \times \frac{1 - 2p}{p} & ; & p = \frac{1}{2} \end{cases}$$

$$P_{7} = \begin{cases} \frac{1 - 2p}{p} \times \frac{1 - 2p}{p} & ; & p = \frac{1}{2} \end{cases}$$

$$P_{7} = \begin{cases} \frac{1 - 2p}{p} \times \frac{1 - 2p}{p} & ; & p = \frac{1}{2} \end{cases}$$

$$P_{7} = \begin{cases} \frac{1 - 2p}{p} \times \frac{1 - 2p}{p} & ; & p = \frac{1}{2} \end{cases}$$

so, if PS & , gambler will lose everything in the long run.