# High-Dimensional Probability

December 13, 2023

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**Reference** High-Dimensional Probability: An Introduction with Applications in Data Science (Roman Vershynin)

### 0 Appetizer

- Convex combination: For  $z_1, z_2, \dots, z_m \in \mathbb{R}^n$ , the form of  $\sum_{i=1}^m \lambda_i z_i$  with  $\lambda_i \geq 0, \sum_{i=1}^m \lambda_i = 1$ . Convex hull of  $T \subset \mathbb{R}^n$ : conv $(T) = \{\text{convex combinations of } z_1, \dots, z_m \in T, m \in \mathbb{N}\}.$
- Caratheodory's theorem: Every point in the convex hull of a set  $T \subset \mathbb{R}^n$  can be expressed as a convex combination of at most n+1 points from T.
- Approximate Caratheodory's theorem: Consider  $T \subset \mathbb{R}^n$ , diam $(T) = \sup\{\|s t\|_2, s, t \in T\} < 1$ . Then for any  $x \in \text{conv}(T)$  and any k, one can find points  $x_1, x_2, \dots, x_k \in T$  such that  $\|x \frac{1}{k} \sum_{i=1}^k x_i\|_2 \leq \frac{1}{\sqrt{k}}$  (repetition is allowed).

Proof WLOG assume 
$$||t||_2 \le 1, \forall t \in T$$
. Fix  $x \in \text{conv}(T), x = \sum_{i=1}^m \lambda_i z_i, z_i \in T$ . Define  $Z, \mathbb{P}(Z = z_i) = \lambda_i, \mathbb{E}Z = x$ . Consider i.i.d.  $Z_1, Z_2, \cdots$  of  $Z, \frac{1}{n} \sum_{j=1}^n Z_j \to x$  a.s.  $n \to +\infty$ .  $\mathbb{E}||x - \frac{1}{k} \sum_{j=1}^k Z_j||_2^2 = \frac{1}{k^2} \sum_{j=1}^k \mathbb{E}||Z_j - x||_2^2 = \frac{1}{k^2} \sum_{j=1}^k (\mathbb{E}||Z_j||^2 - \|\mathbb{E}Z_j\||_2^2) \le \frac{1}{k} \Rightarrow \exists \text{ a realization of } Z_1, \cdots, Z_k \text{ such that } ||x - \frac{1}{k} \sum_{j=1}^k Z_j||_2 \le \frac{1}{\sqrt{k}}$ .

• Corollary (Covering polytopes by balls): P is a polytope in  $\mathbb{R}^n$  with N vertices, diam $(P) \leq 1$ . Then P can be covered by at most  $N^{\lfloor 1/\epsilon^2 \rfloor}$  Euclidean balls of radii  $\epsilon > 0$ .

#### 1 Preliminaries on random variables

- Jensen's inequality: convex  $\phi$ ,  $\phi(\mathbb{E}X) \leq \mathbb{E}\phi(X)$ .  $\Rightarrow \|X\|_{L^p} \leq \|X\|_{L^q}$  for  $p \leq q$ .
- Minkowski inequality:  $p \ge 1, ||X + Y||_{L^p} \le ||X||_{L^p} + ||Y||_{L^p}.$
- Cauchy-Schwarz inequality:  $\mathbb{E}|XY| \leq ||X||_{L^2}||Y||_{L^2}$ .
- Holder inequality:  $p, q \in (1, +\infty), \frac{1}{p} + \frac{1}{q} = 1 \text{ or } p = 1, q = \infty, \mathbb{E}||XY|| \le ||X||_{L^p}||Y||_{L^q}.$
- $X \ge 0$ , then  $\mathbb{E}X = \int_0^{+\infty} \mathbb{P}(X > t) dt$ .
- Markov inequality:  $X \ge 0, t > 0, \mathbb{P}(X \ge t) \le \frac{\mathbb{E}X}{t}$ .
- LLN:  $X_1, \dots, X_n, \dots$  i.i.d.,  $\mathbb{E}X_i = \mu, \operatorname{Var}(X_i) = \sigma^2, S_N = X_1 + X_2 + \dots + X_N$ . Then: (WLLN)  $\mathbb{P}(|\frac{S_N}{N} \mu| > \epsilon) \to 0, \forall \epsilon > 0$ ; (SLLN)  $\mathbb{P}(\frac{S_N}{N} \to \mu, N \to +\infty) = 1$ .
- CLT:  $Z_N = \frac{S_N \mathbb{E}S_N}{\sqrt{\operatorname{Var}(S_N)}} \xrightarrow{d} \mathcal{N}(0, 1).$
- $X_{N,i}, 1 \leq i \leq N$  independent  $\operatorname{Ber}(p_{N,i}), \max_{i \leq N} p_{N,i} \to 0, S_N = \sum_{i=1}^N X_{N,i}, \mathbb{E}S_N \to \lambda < +\infty$ . Then  $S_N \xrightarrow{d} \operatorname{Poisson}(\lambda)$ .

# 2 Concentration of sums of independent random variables

- Question: N times,  $\mathbb{P}(\text{head} \geq \frac{3}{4}N) = ?$  Let  $S_N$  be the number of heads,  $\mathbb{E}S_N = \frac{N}{2}$ ,  $\text{Var}(S_N) = \frac{N}{4}$ . (1) Chebyshev's inequality:  $\mathbb{P}(S_N \geq \frac{3}{4}N) \leq \mathbb{P}(|S_N \frac{N}{2}| \geq \frac{N}{4}) \leq \frac{4}{N}$ ; (2)  $Z_N = \frac{S_N \frac{N}{2}}{\sqrt{N/4}}$ , expect:  $\mathbb{P}(Z_N \geq \sqrt{N/4}) \approx \mathbb{P}(g \geq \sqrt{N/4}) \leq \frac{1}{\sqrt{2\pi}}e^{-N/8}$  where  $g \sim \mathcal{N}(0, 1)$ .
- For all t > 0,  $(\frac{1}{t} \frac{1}{t^3}) \frac{1}{\sqrt{2\pi}} e^{-t^2/2} \le \mathbb{P}(g \sim \mathcal{N}(0, 1) \ge t) \le \frac{1}{t} \frac{1}{\sqrt{2\pi}} e^{-t^2/2}$ .
- Berry-Esseen bound:  $|\mathbb{P}(Z_N \geq t) \mathbb{P}(g \geq t)| \leq \frac{\rho}{\sqrt{N}}$  where  $\rho = \mathbb{E}|X_1 \mu|^3/\sigma^3$ . And in general, no improvement since  $\mathbb{P}(S_N = \frac{N}{2}) \sim \frac{1}{\sqrt{N}} \Rightarrow \mathbb{P}(Z_N = 0) \sim \frac{1}{\sqrt{N}}$  but  $\mathbb{P}(g = 0) = 0$ .
- Hoeffding's inequality:  $X_1, \dots, X_N$  i.i.d. symmetric Bernoulli  $(\mathbb{P}(X=1) = \mathbb{P}(X=-1) = \frac{1}{2}), a = (a_1, \dots, a_N) \in \mathbb{R}^N$ . Then  $\forall t \geq 0, \mathbb{P}(\sum_{i=1}^N a_i X_i \geq t) \leq e^{-t^2/2\|a\|_2^2}, \mathbb{P}(|\sum_{i=1}^N a_i X_i| \geq t) \leq 2e^{-t^2/2\|a\|_2^2}$ .

Proof WLOG, 
$$||a||_{2}^{2} = 1$$
. For  $\lambda > 0$ ,  $\mathbb{P}(\sum_{i=1}^{N} a_{i}X_{i} \geq t) = \mathbb{P}(e^{\lambda \sum a_{i}X_{i}} \geq e^{\lambda t}) \leq e^{-\lambda t}\mathbb{E}e^{\lambda \sum_{i=1}^{N} a_{i}X_{i}} = e^{-\lambda t}\prod_{i=1}^{N}\mathbb{E}e^{\lambda a_{i}X_{i}} = e^{-\lambda t}\prod_{i=1}^{N}e^{\lambda^{2}a_{i}^{2}/2} = e^{-\lambda t + \frac{\lambda^{2}}{2}} \Rightarrow \mathbb{P}(\sum_{i=1}^{N} a_{i}X_{i} \geq t) \leq \inf_{\lambda \geq 0}e^{-\lambda t + \frac{\lambda^{2}}{2}} = e^{-\frac{t^{2}}{2}}(\lambda = t)$ .

#### CONCENTRATION OF SUMS OF INDEPENDENT RANDOM VARIABLES

- Bounded r.v.s:  $X_1, \dots, X_N$  independent,  $X_i \in [m_i, M_i]$ . Then  $\forall t \geq 0, \mathbb{P}(\sum_{i=1}^N (X_i \mathbb{E}X_i) \geq t) \leq e^{-\frac{2t^2}{\sum_{i=1}^N (M_i m_i)^2}}$ .
- Chernoff's inequality:  $X_i \sim \text{Ber}(p_i)$  independent,  $S_N = \sum_{i=1}^N X_i, \mu = \mathbb{E}S_N \Rightarrow \forall t > \mu, \mathbb{P}(S_N \geq t) \leq e^{-\mu} (\frac{e\mu}{t})^t$ .  $Proof \ \mathbb{P}(S_N \geq t) \leq e^{-\lambda t} \prod_{i=1}^N \mathbb{E}e^{\lambda X_i}. \ \mathbb{E}e^{\lambda X_i} = e^{\lambda}p_i + (1-p_i) = 1 + (e^{\lambda}-1)p_i \leq e^{(e^{\lambda}-1)p_i} \Rightarrow \mathbb{P}(S_N \geq t) \leq e^{-\lambda t}e^{(e^{\lambda}-1)\mu}.$  Take  $\lambda^* = \log(t/\mu)$ .
- d = (n-1)p is the expected degree. There is an absolute constant C s.t. for G(n,p),  $d \ge C \log n$ . Then with high prob (for example 0.9), all vertices of G have degrees between 0.9d and 1.1d.

Proof Ex  $2.3.5 \Rightarrow \mathbb{P}(|d_i - d| \ge \delta d) \le 2e^{-c\delta^2 d}$ . Union bound:  $\mathbb{P}(\exists i, |d_i - d| \ge \delta d) \le n \cdot 2e^{-c\delta^2 d} \le n \cdot 2 \cdots n^{-Cc\delta^2} = 2n^{1-Cc\delta^2} \le 1-p^*$  (let  $Cc\delta^2 > 1$ ).

• Sub-gaussian properties: The following are equivalent: (i)  $\mathbb{P}(|X| \geq t) \leq 2e^{-t^2/k_1^2}$  for all  $t \geq 0$ ; (ii)  $\|X\|_{L^p} = (\mathbb{E}|X|^p)^{1/p} \leq k_2\sqrt{p}$  for all  $p \geq 1$ ; (iii)  $\mathbb{E}e^{\lambda^2X^2} \leq e^{k_3^2\lambda^2}$  for all  $\lambda$  s.t.  $|\lambda| \leq \frac{1}{k_3}$ ; (iv)  $\mathbb{E}e^{X^2/k_4^2} \leq 2$ ; (v)  $\mathbb{E}e^{\lambda X} \leq e^{k_5^2\lambda^2}$ , for all  $\lambda \in \mathbb{R}$  (if  $\mathbb{E}X = 0$ ).

 $Proof \text{ (i)} \Rightarrow \text{ (ii): WLOG } k_1 = 1. \ \mathbb{E}|X|^p = \int_0^{+\infty} \mathbb{P}(|X| \geq t) p t^{p-1} \mathrm{d}t \leq \int_0^{+\infty} 2e^{-t^2} p t^{p-1} \mathrm{d}t = p \Gamma(\frac{p}{2})^{\Gamma(x) \leq 3x^x \text{ for } x \geq \frac{1}{2}} 3p(\frac{p}{2})^{p/2} \Rightarrow \|X\|_p \leq \frac{1}{\sqrt{2}} (3p)^{1/p} p^{1/2} \leq 3\sqrt{p}.$ 

(ii)  $\Rightarrow$  (iii): WLOG  $k_2 = 1$ .  $\mathbb{E}e^{\lambda^2 X^2} = \mathbb{E}[1 + \sum_{p=1}^{+\infty} \frac{(\lambda^2 X^2)^p}{p!}]$ .  $\mathbb{E}|X|^{2p} \leq (2p)^p, p! \geq (\frac{p}{e})^p \Rightarrow \mathbb{E}e^{\lambda^2 X^2} \leq 1 + \sum_{p=1}^{+\infty} \frac{(2\lambda^2 p)^p}{(\frac{p}{e})^p} = \frac{1}{1-2e\lambda^2}$  (if  $2e\lambda^2 < 1$ )  $\leq e^{4e\lambda^2}$  (if  $2e\lambda^2 \leq 1/2 \Leftrightarrow \lambda \leq \frac{1}{2\sqrt{e}}$ ).

(iii)  $\Rightarrow$  (iv): trivial.

- $(\mathrm{iv}) \Rightarrow (\mathrm{i}) \colon \mathbb{P}(|X| \geq t) = \mathbb{P}(e^{X^2} \leq e^{t^2}) \leq e^{-t^2} \mathbb{E}e^{X^2} \leq 2e^{-t^2}.$
- (iii)  $\Rightarrow$  (v): WLOG  $k_3 = 1$ . If  $|\lambda| \le 1$ , then  $\mathbb{E}e^{\lambda X} \le \mathbb{E}(\lambda X + e^{\lambda^2 X^2}) = \mathbb{E}e^{\lambda^2 X^2} \le e^{\lambda^2}$ . If  $|\lambda| \ge 1$ , then  $\mathbb{E}e^{\lambda X} \le e^{\frac{\lambda^2}{2}} \mathbb{E}e^{\frac{X^2}{2}} e^{\frac{\lambda^2}{2}} e^{\frac{\lambda^2}{2}$
- $(v) \Rightarrow (i)$ : mimic the proof of  $(iv) \Rightarrow (i)$ .
- Sub-gaussian r.v.: satisfy the above sub-gaussian properties.  $||X||_{\psi_2} = \inf\{t > 0 : \mathbb{E}e^{X^2/t^2} \le 2\}$ . Thus  $\mathbb{P}(|X| \ge t) \le 2e^{-ct^2/||X||_{\psi_2}^2}$ ;  $||X||_{L^p} \le C||X||_{\psi_2}\sqrt{p}$ ; if  $\mathbb{E}X = 0$  then  $\mathbb{E}e^{\lambda X} \le e^{C\lambda^2||X||_{\psi_2}^2}$ .
- Maximum of sub-gaussians:  $K = \max_{i \leq N} \|X_i\|_{\psi_2}$ . Then  $\mathbb{E} \max_{i \leq N} X_i \leq CK\sqrt{\log N}$ .
- Let  $X_1, \dots, X_N$  be independent and mean zero sub-gaussian, then  $\|\sum_{i=1}^N X_i\|_{\psi_2}^2 \leq C \cdot \sum_{i=1}^N \|X_i\|_{\psi_2}^2$ . Proof  $\forall \lambda \in \mathbb{R}, \mathbb{E}e^{\lambda \sum_{i=1}^N X_i} = \prod_{i=1}^N \mathbb{E}e^{\lambda X_i} \leq \prod_{i=1}^N e^{c\lambda^2 \|X_i\|_{\psi_2}^2} = e^{c\lambda^2 \sum_{i=1}^n \|X_i\|_{\psi_2}^2}$
- Centering: X is sub-gaussian  $\Rightarrow X \mathbb{E}X$  is sub-gaussian and  $\|X \mathbb{E}X\|_{\psi_2} \le C\|X\|_{\psi_2}$ .

 $Proof \|\mathbb{E}X\|_{\psi_2} \le C_1 \|\mathbb{E}X\| \le C_1 \mathbb{E}|X| = C_1 \|X\|_{L^1} \le C_1 C_2 \|X\|_{\psi_2}.$ 

• Sub-exponential properties: The following are equivalent: (1)  $\mathbb{P}(|X| \geq t) \leq 2e^{-t/k_1}, \forall t \geq 0$ ; (2)  $||X||_{L^p} \leq k_2 p, p \geq 1$ ; (3)  $\mathbb{E}e^{\lambda|X|} \leq e^{k_3\lambda}$  for all  $0 \leq \lambda \leq \frac{1}{k_3}$ ; (4)  $\mathbb{E}e^{|X|/k_4} \leq 2$ ; (5) if  $\mathbb{E}X = 0, \mathbb{E}e^{\lambda X} \leq e^{k_5^2\lambda^2}$  for  $|\lambda| \leq \frac{1}{k_5}$ .

Proof (2)  $\Rightarrow$  (5):  $k_2 = 1$ ,  $\mathbb{E}e^{\lambda X} = 1 + \sum_{p=2}^{+\infty} \frac{\lambda^p \mathbb{E}X^p}{p!} \le 1 + \sum_{p=2}^{+\infty} \frac{\lambda^p p^p}{(p/e)^p} = 1 + \frac{(e\lambda)^2}{1-e\lambda} (|e\lambda| < 1)$ . If  $|e\lambda| \le \frac{1}{2}$ ,  $1 + \frac{(e\lambda)^2}{1-e\lambda} \le 1 + 2e^2\lambda^2 \le e^{2e^2\lambda^2}$ , i.e.  $k_5 = 2e$ .

$$(5) \Rightarrow (1): k_5 = 1, |x|^p \le p^p(e^x + e^{-x}) \Rightarrow \mathbb{E}|X|^p \le p^p(\mathbb{E}e^X + \mathbb{E}e^{-X}) \le 2ep^p.$$

- $||X||_{\psi_1} = \inf\{t > 0 : \mathbb{E}e^{|X|/t} \le 2\}$ . X is sub-gaussian  $\Leftrightarrow X^2$  is sub-exponential.  $||X^2||_{\psi_1} = ||X||_{\psi_2}^2$ .
- X, Y are sub-gaussian  $\Rightarrow XY$  is sub-exponential and  $||XY||_{\psi_1} \leq ||X||_{\psi_2} ||Y||_{\psi_2}$ .

Proof WLOG  $||X||_{\psi_2} = ||Y||_{\psi_2} = 1$ .  $\mathbb{E}e^{XY} \le \mathbb{E}e^{\frac{X^2 + Y^2}{2}} = \mathbb{E}[e^{\frac{X^2}{2} + \frac{Y^2}{2}}] \le \frac{1}{2}(\mathbb{E}e^{X^2} + \mathbb{E}e^{Y^2}) = 2$ .

• Orlicz function/space:  $\psi: [0, +\infty) \to [0, +\infty)$ , convex, increasing,  $\psi(0) = 0$ ,  $\psi(x) \to +\infty$ ,  $x \to +\infty$ .  $||X||_{\psi} := \inf\{t > 0 : \mathbb{E}\psi(|X|/t) \le 1\}$ .  $L_{\psi} := \{X : ||X||_{\psi} < +\infty\}$  is Banach space. Examples: (1)  $L_p : \psi(x) = x^p, p \ge 1$ ; (2)  $L_{\psi_2} : \psi_2(x) = e^{x^2} - 1, L_{\infty} \subset L_{\psi_2} \subset L_p$ .

#### RANDOM VECTORS IN HIGH DIMENSIONS

• Bernstein's inequality:  $X_1, \dots, X_N$  independent, mean zero and sub-exponential. Then for  $t \geq 0, \mathbb{P}(|\sum_{i=1}^N X_i| \geq t) \leq 2e^{-c\min(\frac{t^2}{\sum_{i=1}^N \|X_i\|_{\psi_1}^2}, \frac{t}{\max_i \|X_i\|_{\psi_1}})}$ .

Proof  $S = \sum_{i=1}^{N} X_i$ .  $\mathbb{P}(S \ge t) \le e^{-\lambda t} \prod_{i=1}^{N} \mathbb{E} e^{\lambda X_i}$ .  $\mathbb{E} e^{\lambda X_i} \le e^{c\lambda^2 \|X_i\|_{\psi_1}^2}$  if  $|\lambda| \le \frac{c}{\max \|X_i\|_{\psi_1}}$ . Then  $\mathbb{P}(S \ge t) \le e^{-\lambda t + c\lambda^2 \sigma^2}$  where  $\sigma^2 := \sum_{i=1}^{N} \|X_i\|_{\psi_1}^2$ . The following is to find the minimum of a quadratic function with the restriction  $|\lambda| \le \frac{c}{\max \|X_i\|_{\psi_1}}$ .

- Corollary 1:  $\mathbb{P}(|\sum_{i=1}^{N} a_i X_i| \ge t) \le 2e^{-c \min(\frac{t^2}{K^2 ||a||_2^2}, \frac{t}{K ||a||_\infty})}$  where  $K = \max_i ||X_i||_{\psi_1}$ .
- Corollary 2:  $|X_i| \le K$ , then  $\mathbb{P}(|\sum_{i=1}^N X_i| \ge t) \le 2 \exp(-\frac{t^2/2}{\sigma^2 + Kt/3})$  where  $\sigma^2 = \sum_{i=1}^N \mathbb{E} X_i^2$ .

### 3 Random vectors in high dimensions

- $X \in \mathbb{R}^n$ , independent sub-gaussian coordinate  $X_i$ ,  $\mathbb{E}X_i^2 = 1$ . Then  $\|\|X\|_2 \sqrt{n}\|_{\psi_2} \le CK^2$ ,  $K = \max_i \|X_i\|_{\psi_2}$ .  $Proof \ \mathbb{E}X_i^2 = 1 \Rightarrow K \ge 1$ .  $\|X_i^2 - 1\|_{\psi_1} \le C\|X_i^2\|_{\psi_1} = C\|X_i\|_{\psi_2}^2 \le CK^2$ . Bernstein's inequality:  $\mathbb{P}(|\frac{1}{n}\|X\|_2^2 - 1| \ge u) = \mathbb{P}(|\frac{1}{n}\sum_{i=1}^n (X_i^2 - 1)| \ge u) \le 2e^{-cn\min(\frac{u^2}{K^4}, \frac{u}{K^2})} \le 2e^{-\frac{cn}{K^4}\min(u^2, u)}$ . For any  $\delta > 0$ ,  $\mathbb{P}(|\frac{1}{\sqrt{n}}\|X\|_2 - 1| \ge \delta) \le \mathbb{P}(|\frac{1}{n}\|X\|_2^2 - 1| \ge \max(\delta, \delta^2)) \le 2e^{-\frac{cn}{K^4}\delta^2} \Rightarrow \mathbb{P}(|\|X\|_2 - \sqrt{n}| \ge t) \le 2e^{-ct^2/K^4}$ .
- Isotropy:  $\Sigma(X) = \mathbb{E}XX^T = I$ . If  $\Sigma \neq I_n$ , then let  $Z = \Sigma^{-1/2}X$ . X is isotropic  $\Leftrightarrow \mathbb{E}\langle X, x \rangle^2 = ||x||_2^2$  for any  $x \in \mathbb{R}^n$ .

$$Proof \ \mathbb{E}\langle X, x \rangle^2 = \mathbb{E}(x^T X X^T x) = x^T (\mathbb{E}X X^T) x. \ \|x\|_2^2 = x^T I_n x. \ \Rightarrow \mathbb{E}X X^T = I_n.$$

• X is isotropic  $\Rightarrow \mathbb{E}||X||_2^2 = n$ . If X, Y are independent and isotropic  $\Rightarrow \mathbb{E}\langle X, Y \rangle^2 = n$ .

Proof 
$$\mathbb{E}||X||_2^2 = \mathbb{E}(X^T X) = \mathbb{E}(\operatorname{tr}(X^T X)) = \operatorname{tr}(\mathbb{E}XX^T) = n.$$

$$\mathbb{E}\langle X,Y\rangle^2 = \mathbb{E}(X^TYY^TX) = \mathbb{E}(\operatorname{tr}(X^TYY^TX)) = \mathbb{E}(\operatorname{tr}(XX^TYY^T)) = \operatorname{tr}((\mathbb{E}XX^T)(\mathbb{E}YY^T)) = n.$$

- Examples:  $X \sim U(\sqrt{n}\mathbb{S}^{n-1}), X \sim U(\{-1,1\}^n), X = (X_1, \cdots, X_n) \text{ i.i.d.}, \mathbb{E}X_i = 0, \text{Var}(X_i) = 1 \text{ are all isotropic.}$
- $g \sim \mathcal{N}(0, I_n)$ , then  $\mathbb{P}(|\|g\|_2 \sqrt{n}| \ge t) \le 2e^{-ct^2}$ .
- Frame:  $\{u_i\}_{i=1}^N \subset \mathbb{R}^n$ , Approximate Parseval's identity:  $A\|x\|_2^2 \leq \sum_{i=1}^N \langle u_i, x \rangle^2 \leq B\|x\|_2^2$ . A, B: frame bounds. A = B: tight frame  $(\Leftrightarrow \sum_{i=1}^N u_i u_i^T = AI_n)$  and in this case,  $\sum_{i=1}^N \langle u_i, x \rangle u_i = Ax$ .
- (a) Tight frame  $\{u_i\}_{i=1}^N$ ,  $A = B, X = \text{Unif}\{u_i, i = 1, 2, \dots, N\}$ , then  $(\frac{N}{A})^{1/2}X$  is isotropic. (b) X is isotropic,  $\mathbb{P}(X = x_i) = p_i, i = 1, 2, \dots, N$ . Then  $u_i = \sqrt{p_i}x_i$  form a tight frame with A = B = 1.
- Isotropic convex sets:  $X \sim \mathrm{Unif}(K), K \subset \mathbb{R}^n$  convex, bounded, non-empty interior (convex body). Assume  $\mathbb{E}X = 0, \Sigma = \mathrm{Cov}(X)$ . Then  $Z = \Sigma^{-1/2}X$  is isotropic and  $Z \sim \mathrm{Unif}(\Sigma^{-1/2}K)$ .
- $X \in \mathbb{R}^n$  is sub-gaussian  $\Leftrightarrow \forall X \in \mathbb{R}^n, \langle X, x \rangle$  are sub-gaussian.  $\|X\|_{\psi_2} = \sup_{x \in \mathbb{S}^{n-1}} \|\langle X, x \rangle\|_{\psi_2}$ .
- $X = (X_1, \dots, X_n)$  independent, mean zero, sub-gaussian coordinate. Then X is sub-gaussian with  $||X||_{\psi_2} \le C \max_{i \le n} ||X_i||_{\psi_2}$ .

$$Proof \ \|\langle X, x \rangle\|_{\psi_2}^2 = \|\sum X_i x_i\|_{\psi_2}^2 \le C \sum_{i=1}^n x_i^2 \|X_i\|_{\psi_2}^2 \le C \max_{i \le n} \|X_i\|_{\psi_2}^2.$$

- Gaussian dist:  $X \sim \mathcal{N}(0, I_n), ||X||_{\psi_2} \leq C$ .
- Discrete dist:  $X \sim \text{Unif}\{\sqrt{n}e_i, i = 1, 2, \cdots, n\}, \|X\|_{\psi_2} \simeq \sqrt{\frac{n}{\log n}}$ .
- Uniform dist:  $X \sim \text{Unif}\{\sqrt{n}\mathbb{S}^{n-1}\}, ||X||_{\psi_2} \leq C.$

Proof  $g \sim \mathcal{N}(0, I_n), X \stackrel{d}{=} \frac{\sqrt{n}g}{\|g\|_2}$ .  $p(t) := \mathbb{P}(|X_1| \geq t) = \mathbb{P}(\frac{|g|}{\|g\|_2} \geq \frac{t}{\sqrt{n}})$ .  $\|\|g\|_2 - \sqrt{n}\|_{\psi_2} \leq C \Rightarrow \mathbb{P}(\|g\|_2 \leq \frac{\sqrt{n}}{2}) \leq 2e^{-cn}$ . Need to show that all one-dimensional marginals  $\langle X, x \rangle$  are sub-gaussian. By rotation invariance, we may assume that  $x = (1, 0, \dots, 0)$ . Let  $\mathcal{E} = \{\|g\|_2 \geq \frac{\sqrt{n}}{2}\} \Rightarrow p(t) \leq \mathbb{P}(\frac{|g|}{\|g\|_2} \geq \frac{t}{\sqrt{n}}, \mathcal{E}) + \mathbb{P}(\mathcal{E}) \leq \mathbb{P}(|g_1| \geq \frac{t}{2}, \mathcal{E}) + 2e^{-cn} \leq 2e^{-t^2/8} + 2e^{-cn} \stackrel{t \leq \sqrt{n}}{\leq} 4e^{-ct^2}$ .

• Grothendieck's inequality:  $A = \{a_{ij}\}_{m \times n}$  of real numbers. Assume  $\forall x_i, y_i \in \{-1, 1\}$ , we have  $|\sum_{i,j} a_{ij} x_i y_j| \leq 1$ . Then for any Hilbert space  $\mathscr{H}$ , any  $u_i, v_j \in \mathscr{H}$  satisfying  $||u_i|| = ||v_j|| = 1$ , we have  $|\sum_{i,j} a_{ij} \langle u_i, v_j \rangle| \leq K$  with  $K \leq 1.783$ .

#### RANDOM MATRICES

- *Proof* (1) Reduction. For any  $u_i, v_j \in \mathbb{R}^N$  s.t.  $||u_i||_2 = ||v_j||_2 = 1, K_{u,v} := \sum_{i,j} a_{ij} \langle u_i, v_j \rangle, K = \sup_{||u||_2 = ||v||_2 = 1} K_{u,v}$ .
- (2) Introduce randomness.  $g \sim \mathcal{N}(0, I_N), U_i = \langle g, u_i \rangle, V_j = \langle g, v_j \rangle, \mathbb{E}U_iV_j = \langle u_i, v_j \rangle.$   $K_{u,v} = \mathbb{E}(\sum_{i,j} a_{ij}U_iV_j) \Rightarrow K_{u,v} \leq R^2$  if  $|U_i| \leq R, |V_j| \leq R$ .
- (3) Truncation. Given  $R \ge 1, U_i = U_i^- + U_i^+, U_i^- = U_i \mathbb{1}_{\{|U_i| \le R\}}, V_j = V_j^- + V_j^+, |U_i^-| \le R, |V_j^-| \le R.$   $||U_i^+||_{L^2}^2 \le 2(R + \frac{1}{R}) \frac{1}{\sqrt{2\pi}} e^{-R^2/2} < \frac{4}{R^2} (R > 1).$
- (4) Breaking up the sum.  $\mathbb{E} \sum a_{ij} U_i V_j = \mathbb{E} \sum a_{ij} U_i^- V_j^- + \mathbb{E} \sum a_{ij} U_i^+ V_j^- + \mathbb{E} \sum a_{ij} U_i^- V_j^+ + \mathbb{E} \sum a_{ij} U_i^+ V_j^+ := S_1 + S_2 + S_3 + S_4.$   $S_1 \leq R^2, S_2 \leq K \max \|U_i^+\|_2 \max \|V_j^-\|_2 \leq K \frac{2}{R}, S_3 \leq K \frac{2}{R}, S_4 \leq K \frac{4}{R^2}.$
- (5) Putting everything together.  $K_{u,v} \leq R^2 + \frac{4K}{R} + \frac{4K}{R^2} \Rightarrow K \leq R^2 + \frac{4K}{R} + \frac{4K}{R^2} \Rightarrow K \leq \frac{R^2}{1 \frac{4}{R} \frac{4}{R^2}}$ .
- Remark: The assumption can be equivalently stated as  $|\sum_{i,j} a_{ij} x_i y_j| \le \max_i |x_i| \max_j |y_j|$ . The conclusion can be equivalently stated as  $|\sum_{i,j} a_{ij} \langle u_i, v_j \rangle| \le K \max_i ||u_i|| \max_j ||v_j||$ .
- Semidefinite programming:  $\max \langle A, X \rangle$  s.t.  $X \succeq 0, \langle B_i, X \rangle = b_i, A, B_i \, n \times n, b_i$  real number,  $\langle A, X \rangle = \operatorname{tr}(A^T X) = \sum_{i,j=1}^n A_{ij} X_{ij}$ .
- Semidefinite relaxation:  $\max \sum_{i,j=1}^n A_{ij} x_i x_j, x_i = \pm 1, A n \times n$  symmetric matrix. Relax to  $\max \sum_{i,j=1}^n A_{ij} \langle X_i, X_j \rangle, \|X_i\|_2 = 1, X_i \in \mathbb{R}^n.$
- A positive semidefinite,  $INT(A) := \max \sum_{i,j=1}^n A_{ij} x_i x_j, x_i = \pm 1, SDP(A) := \max \sum_{i,j=1}^n A_{ij} \langle X_i, X_j \rangle, ||X_i||_2 = 1.$ Then  $INT(A) \leq SDP(A) \leq 2K \cdot INT(A)$ .
- Maximum cut: G = (V, E) finite simple,  $V \to V_1 + V_2$ , cut number of edges crossing between  $V_1$  and  $V_2$ . MAX-CUT(G): NP-hard. Adjacency matrix  $A = \{A_{ij}\}_{n \times n}, A_{ij} = \begin{cases} 1, & i \leftrightarrow j \\ 0, \text{ otherwise} \end{cases}$ . Partition:  $X = (x_i)_{n \times 1}, x_i = \pm 1$ . CUT $(G, X) = \frac{1}{4} \sum_{i,j=1}^{n} A_{ij} (1 x_i x_j)$ . MAX-CUT $(G) = \frac{1}{4} \max\{\sum_{i,j} A_{ij} (1 x_i x_j) : x_i = \pm 1\}$ .
- 0.5-approximation algorithm: Partition at random,  $\mathbb{E}CUT(G, X) = 0.5|E| \ge 0.5MAX-CUT(G)$ .
- 0.878-approximation algorithm: SDP(G) =  $\frac{1}{4}$  max{ $\sum_{i,j=1}^{n} A_{ij}(1 \langle X_i, X_j \rangle), x_i \in \mathbb{R}^n, \|X_i\|_2 = 1$ }.  $X_1, \dots, X_n \to x_1, \dots, x_n : g \sim \mathcal{N}(0, I_n), x_i := \operatorname{sgn}(\langle X_i, g \rangle)$ .  $\mathbb{E}\operatorname{CUT}(G, X) \geq 0.878\operatorname{SDP}(G) \geq 0.878\operatorname{MAX-CUT}(G)$ . Proof  $\mathbb{E}\operatorname{CUT}(G, X) = \frac{1}{4}\sum_{i,j=1}^{n} A_{ij}(1 - \mathbb{E}x_i x_j)$  and  $1 - \mathbb{E}x_i x_j = 1 - \mathbb{E}\operatorname{sgn}\langle g, X_i \rangle \operatorname{sgn}\langle g, X_j \rangle = 1 - \frac{2}{\pi}\operatorname{arcsin}\langle X_i, X_j \rangle \geq 0.878(1 - \langle X_i, X_j \rangle)$ .
- $u, v \in \mathbb{S}^{n-1}$ ,  $\mathbb{E}\operatorname{sgn}(\langle g, u \rangle)\operatorname{sgn}(\langle g, v \rangle) = \frac{2}{\pi} \arcsin\langle u, v \rangle$ .
- There exists a Hilbert space  $\mathcal{H}$  and  $\phi, \psi : \mathbb{S}^{n-1} \to \mathbb{S}(\mathcal{H})$  s.t.  $\frac{2}{\pi} \arcsin\langle \phi(u), \psi(v) \rangle = \beta \langle u, v \rangle$  for all  $u, v \in \mathbb{S}^{n-1}$  and  $\beta = \frac{2}{\pi} \log(1 + \sqrt{2})$ .

 $Proof \ \langle \phi(u), \psi(v) \rangle = \sin(\frac{\beta \pi}{2} \langle u, v \rangle). \ \text{Ex } 3.7.6 \Rightarrow \exists \mathcal{H}, \phi, \psi. \ \sin t = t - \frac{t^3}{3!} + \frac{t^5}{5!} + \cdots, \sinh t = t + \frac{t^3}{3!} + \frac{t^5}{5!}, \|\phi(u)\|_2^2 = \|\psi(u)\|_2^2 = \sinh(\frac{\beta \pi}{2}) = 1 \ \text{for all } u \in \mathbb{S}^{n-1} \Rightarrow \beta = \frac{2}{\pi} \log(1 + \sqrt{2}).$ 

Proof of Grothendieck's inequality with  $K \leq \frac{1}{\beta} \approx 1.783$  WLOG  $u_i, v_j \in \mathbb{S}^{N-1}$ , then  $\frac{2}{\pi} \arcsin \langle u_i', v_j' \rangle = \beta \langle u_i, v_j \rangle$ ,  $\mathcal{H} = \mathbb{R}^M$ ,  $g \sim \mathcal{N}(0, I_M)$ .  $\beta \sum a_{ij} \langle u_i, v_j \rangle = \sum_{i,j} a_{ij} \frac{2}{\pi} \arcsin \langle u_i', v_j' \rangle = \sum_{i,j} a_{ij} \mathbb{E} \operatorname{sgn} \langle g, u_i' \rangle \operatorname{sgn} \langle g, v_j' \rangle \leq 1$ .

#### 4 Random matrices

- Singular vector decomposition:  $A_{m \times n} = U_{m \times m} \Sigma_{m \times n} V_{n \times n}^T = \sum_{i=1}^n s_i U_i V_i^T, \Sigma = \text{diag}(s_1, s_2, \dots, s_r), s_i \ge 0$  sigular values.  $s_i = \sqrt{\lambda_i(AA^T)} = \sqrt{\lambda_i(A^TA)}$ . If A is symmetric,  $s_i = |\lambda_i(A)|$ .
- Courant-Fisher's min-max theorem:  $\lambda_i(A) = \max_{\dim E = i} \min_{x \in \mathbb{S}(E)} \langle Ax, x \rangle, s_i(A) = \max_{\dim E = i} \min_{x \in \mathbb{S}(E)} \|Ax\|_2.$
- Operator norm/spectral norm:  $||A|| = \max_{x \in \mathbb{R}^n \setminus \{0\}} \frac{||Ax||_2}{||x||_2} = \max_{x \in \mathbb{S}^{n-1}} ||Ax||_2 = s_1(A)$ . Or equivalently,  $||A|| = \max_{x \in \mathbb{S}^{n-1}, y \in \mathbb{S}^{m-1}} \langle Ax, y \rangle$ .
- $s_n(A) > 0 \Leftrightarrow m \ge n = \operatorname{rank}(A), s_n(A) = \frac{1}{\|A^+\|}$  where  $A^+$  is pseudo-inverse (the norm of  $A^{-1}$  restriction to the image of A).

#### RANDOM MATRICES

- Frobenius norm:  $||A||_F = (\sum_{i=1}^n \sum_{j=1}^n A_{ij}^2)^{\frac{1}{2}} = (\sum_{i=1}^n s_i^2(A))^{\frac{1}{2}}$ .
- Low-rank approximation:  $\operatorname{rank}(A) = r, k < r, A_k := \sum_{i=1}^k s_i u_i v_i^T, ||A A_k|| = \min_{\operatorname{rank}(A') \le k} ||A A'|| \text{ (holds for } ||\cdot||, ||\cdot||_F).$
- Approximate isometries:  $m||x||_2 \le ||Ax||_2 \le n||x||_2$  where  $m = s_n(A)$ ,  $n = s_1(A)$ , or  $s_n||x y||_2 \le ||Ax Ay||_2 \le s_1||x y||_2$ .
- $A_{m \times n}, \delta > 0$ . If  $||A^T A I_n|| \le \max(\delta, \delta^2)$ , then  $(1 \delta)||x||_2 \le ||Ax||_2 \le (1 + \delta)||x||_2$  for all x. Proof WLOG  $||x||_2 = 1$ .  $|||Ax||_2^2 - 1| = |\langle (A^T A - I_n)x, x \rangle| \le \max(\delta, \delta^2) \Rightarrow \max(|||Ax||_2 - 1|, (||Ax||_2 - 1)^2) \le \max(\delta, \delta^2) \Rightarrow |||Ax||_2 - 1| \le \delta$ .
- $Q_{n \times m}, QQ^T = I_n \Leftrightarrow P = Q^TQ$  is an orthogonal proj in  $\mathbb{R}^m$  onto a subspace with dim n.
- $\epsilon$ -net: (T, d) a metric space,  $K \subset T$ ,  $\epsilon > 0$ .  $\mathcal{N} \subset K$  is an  $\epsilon$ -net of K if  $\forall x \in K, \exists x_0 \in \mathcal{N}$  s.t.  $d(x, x_0) \leq \epsilon$ . Covering number: smallest  $|\mathcal{N}| = |\mathcal{N}(K, d, \epsilon)|$ .
- Compactness:  $\mathcal{N}(K, d, \epsilon) < +\infty$  for all  $\epsilon > 0$ .
- $\epsilon$ -separated:  $\mathcal{P} \subset T$  is  $\epsilon$ -separated if  $d(x,y) > \epsilon$  for all  $x,y \in \mathcal{P}$ . Packing number: largest  $|\mathcal{P}| = |\mathcal{P}(K,d,\epsilon)|$ .
- $\mathcal{P}$  is a maximal  $\epsilon$ -separated subset  $\Rightarrow \mathcal{P}$  is a  $\epsilon$ -net of K.
- $\mathcal{P}(K, d, 2\epsilon) \le \mathcal{N}(K, d, \epsilon) \le \mathcal{P}(K, d, \epsilon)$ .

Proof The upper bound follows from the previous lemma. For the lower bound, choose an  $2\epsilon$ -separated subset  $\mathcal{P} = \{x_i\}$  in K and an  $\epsilon$ -net  $\mathcal{N} = \{y_j\}$  of K.  $\forall x_i, \exists y_j \in \mathcal{N}$ , s.t.  $|x_i - y_j| < \epsilon$ .  $\forall y_j$ , there exists at most a  $x_j \in \mathcal{P}$  s.t.  $|x_i - y_j| < \epsilon$ .

- Minkowski sum:  $A, B \in \mathbb{R}^n, A + B := \{a + b, a \in A, b \in B\}.$
- $K \subset \mathbb{R}^n$ ,  $\epsilon > 0$ ,  $\frac{|K|}{|\epsilon B_2^n|} \le \mathcal{N}(K, \epsilon) \le \mathcal{P}(K, \epsilon) \le \frac{|K + \frac{\epsilon}{2} B_2^n|}{|\frac{\epsilon}{2} B_2^n|}$  where  $|\cdot|$  denotes the volume in  $\mathbb{R}^n$ ,  $B_2^n$  denotes the unit Euclidean ball in  $\mathbb{R}^n$ .
- Corollary: Let  $K = B_2^n$ .  $|\epsilon B_2^n| = \epsilon^n |K|, |K + \frac{\epsilon}{2} B_2^n| = (1 + \frac{\epsilon}{2})^n |K| \Rightarrow (\frac{1}{\epsilon})^n \leq \mathcal{N}(B_2^n, \epsilon) \leq (1 + \frac{2}{\epsilon})^n$ .  $\epsilon \in (0, 1] \Rightarrow (\frac{1}{\epsilon})^n \leq \mathcal{N}(B_2^n, \epsilon) \leq (\frac{3}{\epsilon})^n$ .
- Hamming cubde:  $x, y \in \{0, 1\}^n, d_H(x, y) := \#\{i : x(i) \neq y(i)\}.$
- (T,d) a metric space,  $K \subset T$ ,  $\mathcal{C}(K,d,\epsilon)$  the smallest number of bits sufficient specify every points  $x \in K$  with accuracy  $\epsilon$  in the metric d. Then  $\log_2 \mathcal{N}(K,d,\epsilon) \leq \mathcal{C}(K,d,\epsilon) \leq \log_2 \mathcal{N}(K,d,\frac{\epsilon}{2})$ .  $\log_2 \mathcal{N}(K,\epsilon)$  is often called the metric entropy of K.

*Proof* Lower bound. Assume  $C(K, d, \epsilon) \leq N$ . There exists a transformation of  $x \in K$  into bit strings of length N. A partition of K into at most  $2^N$  subsets.

Upper bound. Assume  $\log_2 \mathcal{N}(K, d, \frac{\epsilon}{2}) \leq N$ . There exists an  $\frac{\epsilon}{2}$ -net  $\mathcal{N}$  with  $|\mathcal{N}| \leq 2^N$ . To every point  $x \in K$ , assign a point  $x_0 \in \mathcal{N}$  that is closest to x. The encoding  $x \mapsto x_0$  represents points in K with accuracy  $\epsilon$ .

- Error correcting code: Fix integers k, n and r. Encoder  $\{0, 1\}^k \to \{0, 1\}^n$ , Decoder  $\{0, 1\}^n \to \{0, 1\}^k$ , D(y) = x if  $x \in \{0, 1\}^k$ ,  $y \in \{0, 1\}^n$  and  $d_H(E(x), y) \le r$ .
- If  $\log_2 \mathcal{P}(\{0,1\}^n, d_H, 2r) \geq k$ , then there exists an error correcting code, k bits  $\to n$  bits, correct r error.  $Proof \ \exists \mathcal{P} \in \{0,1\}^n, |\mathcal{P}| = 2^k \text{ s.t closed balls centered at } \mathcal{P} \text{ with radii } r \text{ are disjoint. } E: \{0,1\}^k \to \mathcal{N} \text{ one to one; } D: \{0,1\}^n \to \{0,1\}^k \text{ nearest-neighbor decodes.}$
- If  $n \ge k + 2r \log_2(\frac{en}{2r})$ , then there exists an error correcting code that encodes k-bit strings into n-bit strings and can correct r errors.

Proof 
$$\mathcal{P}(\{0,1\}^n, d_H, 2r) \ge \mathcal{N}(\{0,1\}^n, d_H, 2r) \ge \frac{2^n}{\sum_{k=0}^{2r} C_n^k} \ge 2^n (\frac{2r}{en})^{2r} \ge 2^k$$
.

#### CONCENTRATION WITHOUT INDEPENDENCE

•  $A_{m \times n}, \epsilon \in [0, 1)$ . Then for any  $\epsilon$ -set  $\mathcal{N}$  of  $\mathbb{S}^{n-1}$ ,  $\sup_{x \in \mathcal{N}} ||Ax||_2 \le ||A|| \le \frac{1}{1-\epsilon} \sup_{x \in \mathcal{N}} ||Ax||_2$ .

Proof Fix  $x \in \mathbb{S}^{n-1}$ ,  $||A|| = ||Ax||_2$ .  $\exists x_0 \in \mathcal{N}$ ,  $||x - x_0||_2 \le \epsilon$ ,  $||Ax - Ax_0||_2 \le ||A|| ||x - x_0||_2 \le \epsilon ||A|| \Rightarrow ||Ax_0||_2 \ge ||Ax||_2 - ||A(x - x_0)||_2 \ge ||A|| - \epsilon ||A||$ .

•  $A_{m\times n} = \{A_{ij}\}, A_{ij}$  independent mean zero sub-gaussian,  $K = \max_{i,j} \|A_{ij}\|_{\psi_2}$ . Then for any t > 0,  $\mathbb{P}(\|A\| \le CK(\sqrt{m} + \sqrt{n} + t)) \ge 1 - 2e^{-t^2}$ .

Proof Step 1: Approximation. Choose  $\epsilon = 1/4$  and  $\epsilon$ -net  $\mathcal{N}$  of  $\mathbb{S}^{n-1}$ ,  $\epsilon$ -net  $\mathcal{M}$  of  $\mathbb{S}^{m-1}$  with  $|\mathcal{N}| \leq 9^n$ ,  $|\mathcal{M}| \leq 9^m$ . Ex 4.4.3  $\Rightarrow ||A|| \leq 2 \max_{x \in \mathcal{N}, y \in \mathcal{M}} \langle Ax, y \rangle$ .

Step 2: Concentration.  $\langle Ax,y\rangle = \sum_{i,j} A_{ij}x_iy_j, \|\langle Ax,y\rangle\|_{\psi_2}^2 \leq C\sum_{i,j} \|A_{ij}\|_{\psi_2}^2 x_i^2 y_j^2 \leq CK^2 \Rightarrow \mathbb{P}(\langle Ax,y\rangle \geq u) \leq 2e^{-cu^2/K^2}$ .

Step 3: Union bound.  $\mathbb{P}(\max_{x \in \mathcal{N}, y \in \mathcal{M}} \langle Ax, y \rangle \ge u) \le \sum_{x \in \mathcal{N}, y \in \mathcal{M}} \mathbb{P}(\langle Ax, y \rangle \ge u) \le 9^{n+m} 2e^{-cu^2/K^2}$ . Take  $u = CK(\sqrt{m} + \sqrt{n} + t)$ ,  $u^2 \ge C^2K^2(m+n+t^2)$ . C sufficiently large s.t.  $cu^2/K^2 \ge 3(n+m+t^2)$ .

•  $A_{n\times n}$  symmetric,  $A_{ij}, i \leq j$  independent mean zero sub-gaussian. Then for  $t \geq 0, \mathbb{P}(\|A\| \leq CK(\sqrt{n} + t)) \geq 1 - 4e^{-t^2}$ .

 $\begin{aligned} & \textit{Proof } \ \ A = \underbrace{A^+ + A^-}_{\text{upper + lower triangular matrix}}, \mathbb{P}(\|A^+\| \leq CK(\sqrt{n} + t)) \geq 1 - 2e^{-4t^2} \geq 1 - 2e^{-t^2}, \mathbb{P}(\|A^-\| \leq CK(\sqrt{n} + t)) \geq 1 - 2e^{-4t^2} \geq 1 - 2e^{-t^2}, \mathbb{P}(\|A^-\| \leq CK(\sqrt{n} + t)) \geq 1 - 2e^{-4t^2} \geq 1 - 2e^{-t^2}, \mathbb{P}(\|A^-\| \leq CK(\sqrt{n} + t)) \geq 1 - 2e^{-4t^2} \geq 1 - 2e^{-t^2}, \mathbb{P}(\|A^-\| \leq CK(\sqrt{n} + t)) \geq 1 - 2e^{-4t^2} \geq 1 - 2e^{-t^2}, \mathbb{P}(\|A^-\| \leq CK(\sqrt{n} + t)) \geq 1 - 2e^{-4t^2} \geq 1 - 2e^{-t^2}, \mathbb{P}(\|A^-\| \leq CK(\sqrt{n} + t)) \geq 1 - 2e^{-4t^2} \geq 1 - 2e^{-t^2}, \mathbb{P}(\|A^-\| \leq CK(\sqrt{n} + t)) \geq 1 - 2e^{-4t^2} \geq 1 - 2e^{-t^2}, \mathbb{P}(\|A^-\| \leq CK(\sqrt{n} + t)) \geq 1 - 2e^{-4t^2} \geq 1 - 2e^{-t^2}, \mathbb{P}(\|A^-\| \leq CK(\sqrt{n} + t)) \geq 1 - 2e^{-4t^2} \geq 1 - 2e^{-t^2}, \mathbb{P}(\|A^-\| \leq CK(\sqrt{n} + t)) \geq 1 - 2e^{-4t^2} \geq 1 - 2e^{-t^2}, \mathbb{P}(\|A^-\| \leq CK(\sqrt{n} + t)) \geq 1 - 2e^{-4t^2} \geq 1 - 2e^{-t^2}, \mathbb{P}(\|A^-\| \leq CK(\sqrt{n} + t)) \geq 1 - 2e^{-4t^2} \geq 1 - 2e^{-t^2}, \mathbb{P}(\|A^-\| \leq CK(\sqrt{n} + t)) \geq 1 - 2e^{-4t^2} \geq 1 - 2e^{-t^2}, \mathbb{P}(\|A^-\| \leq CK(\sqrt{n} + t)) \geq 1 - 2e^{-4t^2} \geq 1 - 2e^{-t^2}, \mathbb{P}(\|A^-\| \leq CK(\sqrt{n} + t)) \geq 1 - 2e^{-t^2}, \mathbb{P}(\|A^-\| \leq CK(\sqrt{n} + t)) \geq 1 - 2e^{-t^2}, \mathbb{P}(\|A^-\| \leq CK(\sqrt{n} + t)) \geq 1 - 2e^{-t^2}, \mathbb{P}(\|A^-\| \leq CK(\sqrt{n} + t)) \geq 1 - 2e^{-t^2}, \mathbb{P}(\|A^-\| \leq CK(\sqrt{n} + t)) \geq 1 - 2e^{-t^2}, \mathbb{P}(\|A^-\| \leq CK(\sqrt{n} + t)) \geq 1 - 2e^{-t^2}, \mathbb{P}(\|A^-\| \leq CK(\sqrt{n} + t)) \geq 1 - 2e^{-t^2}, \mathbb{P}(\|A^-\| \leq CK(\sqrt{n} + t)) \geq 1 - 2e^{-t^2}, \mathbb{P}(\|A^-\| \leq CK(\sqrt{n} + t)) \geq 1 - 2e^{-t^2}, \mathbb{P}(\|A^-\| \leq CK(\sqrt{n} + t)) \geq 1 - 2e^{-t^2}, \mathbb{P}(\|A^-\| \leq CK(\sqrt{n} + t)) \geq 1 - 2e^{-t^2}, \mathbb{P}(\|A^-\| \leq CK(\sqrt{n} + t)) \geq 1 - 2e^{-t^2}, \mathbb{P}(\|A^-\| \leq CK(\sqrt{n} + t)) \geq 1 - 2e^{-t^2}, \mathbb{P}(\|A^-\| \leq CK(\sqrt{n} + t)) \geq 1 - 2e^{-t^2}, \mathbb{P}(\|A^-\| \leq CK(\sqrt{n} + t)) \geq 1 - 2e^{-t^2}, \mathbb{P}(\|A^-\| \leq CK(\sqrt{n} + t)) \geq 1 - 2e^{-t^2}, \mathbb{P}(\|A^-\| \leq CK(\sqrt{n} + t)) \geq 1 - 2e^{-t^2}, \mathbb{P}(\|A^-\| \leq CK(\sqrt{n} + t)) \geq 1 - 2e^{-t^2}, \mathbb{P}(\|A^-\| \leq CK(\sqrt{n} + t)) \geq 1 - 2e^{-t^2}, \mathbb{P}(\|A^-\| \leq CK(\sqrt{n} + t)) \geq 1 - 2e^{-t^2}, \mathbb{P}(\|A^-\| \leq CK(\sqrt{n} + t)) \geq 1 - 2e^{-t^2}, \mathbb{P}(\|A^-\| \leq CK(\sqrt{n} + t)) \geq 1 - 2e^{-t^2}, \mathbb{P}(\|A^-\| \leq CK(\sqrt{n} + t)) \geq 1 - 2e^{-t^2}, \mathbb{P}(\|A^-\| \leq CK(\sqrt{n} + t)) \geq 1 - 2e^{-t^2},$ 

- Stochastic block model (SBM): G(n, p, q), p > q, n vertices, two community of size n/2,  $x, y \in$  same community  $\Rightarrow \mathbb{P}(x \sim y) = p$ , otherwise  $\mathbb{P}(x \sim y) = q$ .  $A = \{A_{ij}\}, A_{ij} = 1 \text{ if } i \sim j \text{ otherwise } 0$ .  $A = \mathbb{E}A + R := D + R, ||D|| = \frac{p+q}{2} \cdot n, \mathbb{P}(||R|| \leq C\sqrt{n}) \geq 1 4e^{-n}$ .
- Weyl's inequality: Symmetric matrices S and T with same dim,  $\max_i |\lambda_i(S) \lambda_i(T)| \le ||S T||$ .
- Davis-Kahan: Fix i,  $\min_{j \neq i} |\lambda_i(S) \lambda_j(S)| = \delta > 0$ . Then  $\sin \angle (v_i(S), v_i(T)) \le \frac{2\|S T\|}{\delta} \Rightarrow \exists \theta \in \{-1, 1\}, \|v_i(S) \theta v_i(T)\|_2 \le \frac{\|S T\|}{\delta} \cdot 2^{3/2}$ .
- Spectual clustering: Recall SBM A = D + R and let S = D, T = A = D + R in Davis-Kahan.  $\delta = \min(\lambda_2, \lambda_2 \lambda_1) = \min(\frac{p-q}{2}, q)n := \mu n$ .  $\mathbb{P}(\|R\| = \|T S\| \le C\sqrt{n}) \ge 1 4e^{-n} \Rightarrow \exists \theta \in \{\pm 1\}, \|v_2(D) \theta v_2(A)\| \le \frac{C}{\mu\sqrt{n}}$ . Let  $u_2(D) = (1, 1, \dots, 1, -1, -1, \dots, -1) \Rightarrow \|u_2(D) \theta u_2(A)\| \le \frac{C}{\mu} \Rightarrow \sum_{j=1}^n |u_2(D)_j \theta u_2(A)_j|^2 \le \frac{C}{\mu^2}$ . Thus the number of disagreeing signs between  $u_2(D)$  and  $u_2(A)$  must be bounded by  $\frac{C}{\mu^2}$ .
- $A_{m \times n}$ , rows  $A_i$  independent mean zero sub-gaussian, isotropic. Then for any  $t \geq 0$ ,  $\sqrt{m} CK^2(\sqrt{n} + t) \leq s_n(A) \leq s_1(A) \leq \sqrt{m} + CK^2(\sqrt{n} + t)$  with prob  $\geq 1 2e^{-t^2}$ . Here  $K = \max_i ||A_i||_{\psi_2}$ .

Proof Only need to prove  $\|\frac{1}{m}A^TA - I_n\| \le \epsilon := K^2 \max\{\delta, \delta^2\}, \delta = C(\frac{\sqrt{n}}{\sqrt{m}} + \frac{t}{\sqrt{m}}).$ 

Step 1: Approximation. Find an  $\frac{1}{4}$ -net  $\mathcal{N}$  of the unit space  $\mathbb{S}^{n-1}$ ,  $|\mathcal{N}| \leq 9^n$ .  $\|\frac{1}{m}A^TA - I_n\| \leq 2 \max_{x \in \mathcal{N}} |\langle (\frac{1}{m}A^TA - I_n)x, x \rangle| = 2 \max_{x \in \mathcal{N}} |\frac{1}{m}\|Ax\|_2^2 - 1|$ .

Step 2: Concentration.  $X_i := \langle A_i, x \rangle$  independent, mean zero,  $||X_i||_{\psi_2} \leq K$ ,  $\mathbb{E}X_i^2 = 1$ .  $\mathbb{P}(|\frac{1}{m}||Ax||_2^2 - 1| \geq \frac{\epsilon}{2}) \leq 2e^{-c_1\delta^2 m} \leq 2e^{-c_1C^2(n+t^2)}$ .

Step 3: Union bound.  $\mathbb{P}(|\frac{1}{m}||Ax||_2^2 - 1| \ge \frac{\epsilon}{2}) \le 9^n \cdot 2e^{-c_1C^2(n+t^2)} \le 2e^{-t^2}$ .

•  $X \in \mathbb{R}^n$  sub-gaussian.  $\mathbb{E}X = 0, \Sigma = \mathbb{E}XX^T, X_i \stackrel{\mathrm{d}}{=} X \text{ i.i.d.}, \Sigma_m = \frac{1}{m} \sum_{i=1}^m X_i X_i^T$ . Assume there exists  $K \ge 1$  s.t.  $\|\langle X, x \rangle\|_{\psi_2}^2 \le K^2 \|\langle X, x \rangle\|_{L^2}^2$ . Then for m,  $\mathbb{E}\|\Sigma_m - \Sigma\| \le CK^2(\sqrt{\frac{n}{m}} + \frac{n}{m})\|\Sigma\|$ .

 $Proof \ \ Z_i = \Sigma^{-1/2} X_i, Z = \Sigma^{-1/2} X, \\ \mathbb{E} Z_i Z_i^T = I_n, \\ \|Z\|_{\psi_2} \leq K, \\ \|Z_i\|_{\psi_2} \leq K. \ \ \text{Then} \ \|\Sigma_m - \Sigma\| = \|\Sigma^{1/2} R_m \Sigma^{1/2}\| \leq \|R_m\| \|\Sigma\| \ \ \text{where} \\ R_m = \frac{1}{m} \sum_{i=1}^m Z_i Z_i^T - I. \ \ \text{Consider an } m \times n \text{ random matrix } A \text{ whose rows are } Z_i^T. \\ \mathbb{E} \|R_m\| = \mathbb{E} \|\frac{1}{m} A^T A - I\| \leq C K^2 (\sqrt{\frac{n}{m}} + \frac{n}{m}). \\ \square$ 

# 5 Concentration without independence

- $(X, d_X) \xrightarrow{f} (Y, d_Y), d_Y(f(u), f(v)) \leq L \cdot d_x(u, v), \forall u, v \in X$ . The infimum of all L in this definition is called the Lipschitz norm of f and is denoted  $||f||_{\text{Lip}}$ .
- $\epsilon > 0, A_{\epsilon} = A + \epsilon B_2^n, A \subset \mathbb{R}^n, \min_A \text{ volume of } A_{\epsilon} \text{ with column } A \text{ fixed is achieved when } A \text{ is a ball.}$

#### CONCENTRATION WITHOUT INDEPENDENCE

- $\sigma_{n-1}(A)$  normalized area on  $\mathbb{S}^{n-1}$ ,  $\epsilon > 0$ . With given  $\sigma_{n-1}(A)$ ,  $\min_A \sigma_{n-1}(A_{\epsilon})$  is achieved when A is a spherical cap.
- $A \subset \sqrt{n}\mathbb{S}^{n-1}$ . If  $\sigma(A) \geq \frac{1}{2}$ , then  $\forall t \geq 0, \sigma(A_t) \geq 1 2e^{-ct^2}$ .

  Proof Let  $H = \{x \in \sqrt{n}\mathbb{S}^{n-1}, x_1 \leq 0\}$ .  $\sigma(A) \geq \frac{1}{2}\sigma(H)$ . Thm  $5.1.6 \Rightarrow \sigma(A_t) \geq \sigma(H_t) = \mathbb{P}(X \in H_t)$ .  $H_t \supset \{x \in \sqrt{n}\mathbb{S}^{n-1}, x_1 \leq \frac{t}{\sqrt{2}}\} \Rightarrow \sigma(H_t) \geq \mathbb{P}(X_1 \leq \frac{t}{\sqrt{2}})$ .  $\|X_1\|_{\psi_2} \leq C \Rightarrow \mathbb{P}(X_1 \leq \frac{t}{\sqrt{2}}) \geq 1 2e^{-ct^2}$ .
- $X \sim \text{Unif}(\sqrt{n}\mathbb{S}^{n-1}), f: \sqrt{n}\mathbb{S}^{n-1} \to \mathbb{R}$ . Then  $||f(X) \mathbb{E}f(X)||_{\psi_2} \leq C||f||_{\text{Lip}}$ . Proof WLOG  $||f||_{\text{Lip}} = 1$ ,  $\mathbb{P}(f(X) \geq M) \geq \frac{1}{2}$ ,  $\mathbb{P}(f(X) \leq M) \geq \frac{1}{2}$ .  $A := \{x \in \sqrt{n}\mathbb{S}^{n-1} : f(x) \leq M\}$ .  $\mathbb{P}(X \in A) \geq \frac{1}{2} \Rightarrow \mathbb{P}(A_t) \geq 1 - 2e^{-ct^2} \Rightarrow \mathbb{P}(f(X) \leq M + t) \geq 1 - 2e^{-ct^2}$ . By centering,  $f(X) - \mathbb{E}f(X) = f(X) - M - (\mathbb{E}f(X) - M)$  is sub-gaussian.  $\square$
- $X \sim \mathcal{N}(0, I_n), \gamma_n(A) = \mathbb{P}(X \in A), \epsilon > 0, \gamma_n(A)$  given, half spaces minimize  $\gamma_n(A_{\epsilon})$ .
- $X \sim \mathcal{N}(0, I_n), f : \mathbb{R}^n \to \mathbb{R}, ||f||_{\text{Lip}} < \infty.$  Then  $||f(X) \mathbb{E}f(X)||_{\psi_2} \le C||f||_{\text{Lip}}.$
- Hamming cube,  $d(x,y) = \frac{1}{n} |\{i : x_i \neq y_i\}|, \mathbb{P}(A) = \frac{|A|}{2^n}. \|f(X) \mathbb{E}f(X)\|_{\psi_2} \leq \frac{C\|f\|_{\text{Lip}}}{\sqrt{n}}.$
- $S_n: n!$  permutation of n symbols.  $d(\pi, \rho) = \frac{1}{n} |\{i: \pi(i) \neq \rho(i)\}|, \mathbb{P}(A) = \frac{|A|}{n!}. \|f(X) \mathbb{E}f(X)\|_{\psi_2} \leq \frac{C\|f\|_{\text{Lip}}}{\sqrt{n}}.$
- Special orthogonal group  $\mathrm{SO}(n)$ , determinant = 1,  $d = \|\cdot\|_F$ ,  $\mathbb P$  is uniform measure.  $\|f(X) \mathbb E f(X)\|_{\psi_2} \leq \frac{C\|f\|_{\mathrm{Lip}}}{\sqrt{n}}$ .
- $G_{n,m}$  all m-dim subspaces of  $\mathbb{R}^n$  ( $\simeq \mathcal{P}_{G_{n,m}}$  orthogonal projections),  $d(E,F) = \|\mathcal{P}_E \mathcal{P}_F\|$ ,  $\mathbb{P}$  is uniform measure. A random subspace E can be constructed by computing the column span (i.e. the image) of a random  $n \times m$  Gaussian random matrix G with i.i.d.  $\mathcal{N}(0,1)$  entries.  $\|f(X) \mathbb{E}f(X)\|_{\psi_2} \leq \frac{C\|f\|_{\text{Lip}}}{\sqrt{n}}$ .
- A random vector X in  $\mathbb{R}^n$  with density  $p(x) = e^{-U(x)}$ , Hess  $U(x) \succeq \kappa I_n$ .  $||f(X) \mathbb{E}f(X)||_{\psi_2} \leq \frac{C||f||_{\text{Lip}}}{\sqrt{\kappa}}$ .
- $X = (X_1, \dots, X_n)$  independent coordinates,  $|X_i| \le 1$  a.s., f convex and Lipschitz.  $||f(X) \mathbb{E}f(X)||_{\psi_2} \le C||f||_{\text{Lip}}$ .
- $E \sim \text{Unif}(G_{n,m}), z \in \mathbb{R}^n, \epsilon > 0$ . Then (a)  $(\mathbb{E}\|P_E z\|_2^2)^{\frac{1}{2}} = \sqrt{\frac{m}{n}}\|z\|_2$ ; (b)  $\mathbb{P}(\|P_E z\|_2 \sqrt{\frac{m}{n}}\|z\|_2) \le \epsilon \sqrt{\frac{m}{n}}\|z\|_2$ )  $\ge 1 2e^{-c\epsilon^2 m}$ .

Proof (a): WLOG  $||z||_2 = 1$ . Rotational invariance:  $\mathbb{P}(E \in A) = \mathbb{P}(U(E) \in A)$  where U is  $n \times n$  orthogonal  $\Rightarrow$  The dist. of  $P_E z$  is the same if we fix  $E, z \in \mathrm{Unif}(\mathbb{S}^{n-1})$ . WLOG  $Pz = (z_1, \dots, z_m, 0, \dots, 0)$ .  $\mathbb{E}||Pz||_2^2 = m\mathbb{E}z_i^2 = \frac{m}{n}$ .

(b): 
$$f: z \to ||Pz||_2, ||f||_{\text{Lip}} = 1 \Rightarrow |||Pz||_2 - \mathbb{E}||Pz||_2||_{\psi_2} \leq \frac{C}{\sqrt{n}}$$

• Johnson-Lindenstrauss lemma:  $\mathcal{X}$  a set of N points in  $\mathbb{R}^n$ ,  $\epsilon > 0$ ,  $m \ge \frac{C}{\epsilon^2} \log N$ ,  $E \sim \mathrm{Unif}(G_{n,m})$ ,  $Q = \sqrt{\frac{n}{m}} \mathcal{P}_E$ . Then  $\mathbb{P}(|||Qx - Qy||_2 - ||x - y||_2| \le \epsilon ||x - y||_2)$  for any  $x, y \in \mathcal{X} \ge 1 - 2e^{-c\epsilon^2 m}$ .

Proof Let  $\mathcal{X} - \mathcal{X} := \{x - y : x, y \in \mathcal{X}\}$ . The latest lemma  $\Rightarrow \forall z, \mathbb{P}\left((1 - \epsilon)\sqrt{\frac{m}{n}}\|z\|_2 \le \|Pz\|_2 \le (1 + \epsilon)\sqrt{\frac{m}{n}}\|z\|_2\right) \ge 1 - 2e^{-c\epsilon^2 m}$ . Union bound:  $\mathbb{P}(\cdots, for any z \in \mathcal{X} - \mathcal{X}) \ge 1 - N^2 \cdot 2e^{-c\epsilon^2 m} \ge 1 - 2e^{-c'\epsilon^2 m}$ .

- $f: \mathbb{R} \to \mathbb{R}, X = \sum_{i=1}^n \lambda_i u_i u_i^T$ , define  $f(X) = \sum_{i=1}^n f(\lambda_i) u_i u_i^T$ .
- P.S.D. order:  $X \succeq 0, X \succeq Y$  if  $X Y \succeq 0$ .
- Golden-Thompson inequality:  $\operatorname{tr}(e^{A+B}) \leq \operatorname{tr}(e^A e^B)$ .
- Lieb's inequality:  $H: n \times n$  symmetric matrix, X P.D.,  $f(X) = \operatorname{tr}(e^{H + \log X})$ . Then f is concave.
- X is a random P.D. matrix  $\Rightarrow \mathbb{E}f(X) \le f(\mathbb{E}X)$ .  $X = e^Z$ , Z symmetric. Then  $\mathbb{E}\operatorname{tr}(e^{H+Z}) \le \operatorname{tr}(e^{H+\log \mathbb{E}e^Z})$ .
- $X_1, \dots, X_N$  independent mean zero  $n \times n$  symmetric random matrices,  $||X_i|| \leq K$  a.s. for all i. Then for  $\forall t \geq 0$ ,  $\mathbb{P}(||\sum_{i=1}^N X_i|| \geq t) \leq 2ne^{-\frac{t^2/2}{\sigma^2 + Kt/3}}$  where  $\sigma^2 = ||\sum_{i=1}^N \mathbb{E}X_i^2||$ .

Proof Step 1: Reduction to MGF.  $S := \sum_{i=1}^{N} X_i$ .  $||S|| = \max_i |\lambda_i(S)| = \max(\lambda_{\max}(S), \lambda_{\max}(-S))$ .  $\mathbb{P}(\lambda_{\max}(S) \geq t) \leq e^{-\lambda t} \mathbb{E} e^{\lambda \lambda_{\max}(S)}$ .  $E := \mathbb{E} e^{\lambda \lambda_{\max}(S)} = \mathbb{E} \lambda_{\max}(e^{\lambda S}) \Rightarrow E \leq \mathbb{E} \operatorname{tr}(e^{\lambda S})$ .

 $\text{Step 2: Apply Lieb's inequality. } \mathbb{E} \text{tr}(e^{\lambda S}) = \mathbb{E} \text{tr}(e^{\sum_{i=1}^{N-1} \lambda X_i + \lambda X_N}) \leq \mathbb{E} \text{tr}(e^{\sum_{i=1}^{N-1} \lambda X_i + \log \mathbb{E} e^{\lambda X_N}}) \leq \text{tr}(e^{\sum_{i=1}^{N} \log \mathbb{E} e^{\lambda X_i}}).$ 

Step 3: Lemma: X is an  $n \times n$  symmetric mean zero random matrix,  $||X|| \leq K$  a.s. Then  $e^{\lambda X} \leq e^{g(\lambda)\mathbb{E}X^2}$  where  $g(\lambda) = \frac{\lambda^2/2}{1-|\lambda|K/3}, |\lambda| < 3/K$ .

#### QUADRADIC FORMS, SYMMETRIZATION, CONTRACTION

Proof  $e^z \le 1 + z + \frac{1}{1 - |z|/3} \frac{z^2}{2}$  if |z| < 3.  $z = \lambda x$ . If  $|x| \le K, |\lambda| < \frac{3}{K}, e^{\lambda x} \le 1 + \lambda x + g(\lambda)x^2$ . (b) of Ex. 5.4.5  $\Rightarrow$  If  $||X|| \le K, |\lambda| < 3/K, \mathbb{E}e^{\lambda X} \le I + g(\lambda)\mathbb{E}X^2 \text{ (since } \mathbb{E}X = 0) \le e^{g(\lambda)\mathbb{E}X^2}.$ Step 4:  $E \leq \operatorname{tr}(e^{\sum_{i=1}^{N} \log \mathbb{E} e^{\lambda X_i}})$ . The latest lemma + (g) of Ex.5.4.5  $\Rightarrow \log \mathbb{E} e^{\lambda X_i} \leq g(\lambda) \mathbb{E} X_i^2 \Rightarrow \sum_{i=1}^{N} \log \mathbb{E} e^{\lambda X_i} \leq g(\lambda) \cdot Z$ where  $Z := \sum_{i=1}^{N} \mathbb{E}X_i^2$  and  $\sigma^2 = ||Z||$ . (e) of Ex.5.4.5  $\Rightarrow \operatorname{tr}(e^{\sum_{i=1}^{N} \log \mathbb{E}e^{\lambda X_i}}) \leq \operatorname{tr}(e^{g(\lambda Z)}) \Rightarrow E \leq \operatorname{tr}(e^{g(\lambda Z)}) \leq n\lambda_{\max}(e^{g(\lambda Z)}) = n\lambda_{\max}(e^{g(\lambda Z)})$  $ne^{g(\lambda)||Z||} = ne^{g(\lambda)\sigma^2}$ . Minimize for  $\lambda$  as a function of t with  $0 < \lambda < 3/K$ .

•  $X \in \mathbb{R}^n, \Sigma = \mathbb{E}XX^T, \Sigma_m = \frac{1}{m} \sum_{i=1}^m X_i X_i^T, X_i \overset{\text{i.i.d.}}{\sim} X, \|X\|_2 \le K(\mathbb{E}\|X\|_2^2)^{\frac{1}{2}} \text{ a.s.. Then } \mathbb{E}\|\Sigma_m - \Sigma\| \le C(\sqrt{\frac{K^2 n \log n}{m}} + \frac{1}{2})^{\frac{1}{2}} \mathbb{E}(X_i - X_i) + \frac{1}{2} \mathbb{E}($ 

 $Proof \ \mathbb{E}||X||_2^2 = \mathbb{E}X^TX = \mathbb{E}\operatorname{tr}(X^TX) = \mathbb{E}\operatorname{tr}(XX^T) = \operatorname{tr}(\Sigma) \Rightarrow ||X||_2^2 \leq K^2\operatorname{tr}(\Sigma) \text{ a.s..} \quad \text{Ex } 5.4.11 \Rightarrow \mathbb{E}||\Sigma_m - \Sigma|| = \mathbb{E}\operatorname{tr}(XX^T) = \mathbb{$  $\frac{1}{m}\mathbb{E}\|\sum_{i=1}^{m}(X_iX_i^T-\Sigma)\|\lesssim \frac{1}{m}(\sigma\sqrt{\log n}+M\log n) \text{ where } \sigma^2=\|\sum_{i=1}^{m}\mathbb{E}(X_iX_i^T-\Sigma)^2\|=m\|\mathbb{E}(XX^T-\Sigma)^2\| \text{ and } M \text{ is chosen s.t.}$  $\|XX^T - \Sigma\| \leq M \text{ a.s.. Then } \mathbb{E}(XX^T - \Sigma)^2 = \mathbb{E}(XX^T)^2 - \Sigma^2 \leq \mathbb{E}(XX^T)^2 = \mathbb{E}(\|X\|_2^2 XX^T) \leq K^2 \mathrm{tr}(\Sigma)\Sigma \Rightarrow \sigma^2 \leq K^2 m \mathrm{tr}(\Sigma)\|\Sigma\|.$  $||XX^T - \Sigma|| \le ||X||_2^2 + ||\Sigma|| \le K^2 \operatorname{tr}(\Sigma) + ||\Sigma|| \le 2K^2 \operatorname{tr}(\Sigma) := M \text{ (since } K \ge 1 \text{ and } ||\Sigma|| \le \operatorname{tr}(\Sigma)).$  Substitute our bounds for  $\sigma^2$ and M into the previous bound  $\frac{1}{m}(\sigma\sqrt{\log n} + M\log n)$ .

## Quadradic forms, symmetrization, contraction

•  $Y \perp Z, \mathbb{E}Z = 0$ , then  $\mathbb{E}F(Y) \leq \mathbb{E}F(Y + Z)$ .

 $Proof \ \ F(y) = F(\mathbb{E}(y+Z)) \leq \mathbb{E}F(y+Z) \Rightarrow \mathbb{E}F(Y) = \mathbb{E}(\mathbb{E}(F(Y+\mathbb{E}Z)|Y)) = \mathbb{E}(\mathbb{E}(F(\mathbb{E}(Y+Z))|Y)) \leq \mathbb{E}(\mathbb{E}(F(Y+Z)|Y)) = \mathbb{E}(\mathbb{E}(F(Y+Z)|Y)) =$  $\mathbb{E}F(Y+Z)$ .

• Decoupling:  $A_{n\times n}$  diagonal-free(i.e. the diagonal entries of A equal zero),  $X=(X_1,\cdots,X_n)$  independent mean zero. Then for every convex function  $F: \mathbb{R} \to \mathbb{R}$ ,  $\mathbb{E}F(X^TAX) \leq \mathbb{E}F(4X^TAX')$  where  $X' \stackrel{\mathrm{d}}{=} X, X' \perp X$ .

 $Proof \ \delta_1, \dots, \delta_n \overset{\text{i.i.d.}}{\sim} \text{Ber}(1, \frac{1}{2}), I = \{i : \delta_i = 1\}, \mathbb{E}\delta_i(1 - \delta_j) = \frac{1}{4}, X^T A X = \sum_{i \neq j} a_{ij} X_i X_j = 4\mathbb{E}_{\delta} \sum_{i,j} \delta_i(1 - \delta_j) a_{ij} X_i X_j = 4\mathbb{E}_{\delta} \sum_{i,j} \delta_i(1 - \delta_j) a_{ij} X_i X_j = 4\mathbb{E}_{\delta} \sum_{i,j} \delta_i(1 - \delta_j) a_{ij} X_i X_j = 4\mathbb{E}_{\delta} \sum_{i,j} \delta_i(1 - \delta_j) a_{ij} X_i X_j = 4\mathbb{E}_{\delta} \sum_{i,j} \delta_i(1 - \delta_j) a_{ij} X_i X_j = 4\mathbb{E}_{\delta} \sum_{i,j} \delta_i(1 - \delta_j) a_{ij} X_i X_j = 4\mathbb{E}_{\delta} \sum_{i,j} \delta_i(1 - \delta_j) a_{ij} X_i X_j = 4\mathbb{E}_{\delta} \sum_{i,j} \delta_i(1 - \delta_j) a_{ij} X_i X_j = 4\mathbb{E}_{\delta} \sum_{i,j} \delta_i(1 - \delta_j) a_{ij} X_i X_j = 4\mathbb{E}_{\delta} \sum_{i,j} \delta_i(1 - \delta_j) a_{ij} X_i X_j = 4\mathbb{E}_{\delta} \sum_{i,j} \delta_i(1 - \delta_j) a_{ij} X_i X_j = 4\mathbb{E}_{\delta} \sum_{i,j} \delta_i(1 - \delta_j) a_{ij} X_i X_j = 4\mathbb{E}_{\delta} \sum_{i,j} \delta_i(1 - \delta_j) a_{ij} X_i X_j = 4\mathbb{E}_{\delta} \sum_{i,j} \delta_i(1 - \delta_j) a_{ij} X_i X_j = 4\mathbb{E}_{\delta} \sum_{i,j} \delta_i(1 - \delta_j) a_{ij} X_i X_j = 4\mathbb{E}_{\delta} \sum_{i,j} \delta_i(1 - \delta_j) a_{ij} X_i X_j = 4\mathbb{E}_{\delta} \sum_{i,j} \delta_i(1 - \delta_j) a_{ij} X_i X_j = 4\mathbb{E}_{\delta} \sum_{i,j} \delta_i(1 - \delta_j) a_{ij} X_i X_j = 4\mathbb{E}_{\delta} \sum_{i,j} \delta_i(1 - \delta_j) a_{ij} X_i X_j = 4\mathbb{E}_{\delta} \sum_{i,j} \delta_i(1 - \delta_j) a_{ij} X_i X_j = 4\mathbb{E}_{\delta} \sum_{i,j} \delta_i(1 - \delta_j) a_{ij} X_i X_j = 4\mathbb{E}_{\delta} \sum_{i,j} \delta_i(1 - \delta_j) a_{ij} X_i X_j = 4\mathbb{E}_{\delta} \sum_{i,j} \delta_i(1 - \delta_j) a_{ij} X_i X_j = 4\mathbb{E}_{\delta} \sum_{i,j} \delta_i(1 - \delta_j) a_{ij} X_i X_j = 4\mathbb{E}_{\delta} \sum_{i,j} \delta_i(1 - \delta_j) a_{ij} X_i X_j = 4\mathbb{E}_{\delta} \sum_{i,j} \delta_i(1 - \delta_j) a_{ij} X_i X_j = 4\mathbb{E}_{\delta} \sum_{i,j} \delta_i(1 - \delta_j) a_{ij} X_i X_j = 4\mathbb{E}_{\delta} \sum_{i,j} \delta_i(1 - \delta_j) a_{ij} X_i X_j = 4\mathbb{E}_{\delta} \sum_{i,j} \delta_i(1 - \delta_j) a_{ij} X_i X_j = 4\mathbb{E}_{\delta} \sum_{i,j} \delta_i(1 - \delta_j) a_{ij} X_i X_j = 4\mathbb{E}_{\delta} \sum_{i,j} \delta_i(1 - \delta_j) a_{ij} X_i X_j = 4\mathbb{E}_{\delta} \sum_{i,j} \delta_i(1 - \delta_j) a_{ij} X_i X_j = 4\mathbb{E}_{\delta} \sum_{i,j} \delta_i(1 - \delta_j) a_{ij} X_i X_j = 4\mathbb{E}_{\delta} \sum_{i,j} \delta_i(1 - \delta_j) a_{ij} X_i X_j = 4\mathbb{E}_{\delta} \sum_{i,j} \delta_i(1 - \delta_j) a_{ij} X_i X_j = 4\mathbb{E}_{\delta} \sum_{i,j} \delta_i(1 - \delta_j) a_{ij} X_i X_j = 4\mathbb{E}_{\delta} \sum_{i,j} \delta_i(1 - \delta_j) a_{ij} X_i X_j = 4\mathbb{E}_{\delta} \sum_{i,j} \delta_i(1 - \delta_j) a_{ij} X_i X_j = 4\mathbb{E}_{\delta} \sum_{i,j} \delta_i(1 - \delta_j) a_{ij} X_i X_j = 4\mathbb{E}_{\delta} \sum_{i,j} \delta_i(1 - \delta_j) a$  $4\mathbb{E}_{I} \sum_{(i,j) \in I \times I^{c}} a_{ij} X_{i} X_{j} \Rightarrow \mathbb{E}_{X} F(X^{T} A X) = \mathbb{E}_{X} F(4\mathbb{E}_{I} \sum_{(i,j) \in I \times I^{c}} a_{ij} X_{i} X_{j}) \leq \mathbb{E}_{I} \mathbb{E}_{X} F(4\sum_{(i,j) \in I \times I^{c}} a_{ij} X_{i} X_{j}). \text{ There exists an } \mathbb{E}_{X} F(X^{T} A X) = \mathbb{E}_{X} F(X^{T}$  $I \text{ s.t. } \mathbb{E}_X F(4\sum_{(i,j)\in I\times I^c} a_{ij}X_iX_j') = \mathbb{E}_X F(4\sum_{(i,j)\in I\times I^c} a_{ij}X_iX_j) \geq \mathbb{E}F(X^TAX). \text{ LHS} \leq \mathbb{E}_X F(4\sum_{i,j} a_{ij}X_iX_j') \text{ by the latest}$ lemma since  $\mathbb{E}[(\sum_{(i,j)\in I\times I} + \sum_{(i,j)\in I^c\times I^c} + \sum_{(i,j)\in I^c\times I})a_{ij}X_iX_j']\big|\{X_i,i\in I\},\{X_j',j\in I^c\}=0.$ 

•  $X, X' \sim \mathcal{N}(0, I_n), X \perp X'$ , then  $\mathbb{E}e^{\lambda X^T A X'} \leq e^{C\lambda^2 \|A\|_F^2}, |\lambda| \leq \frac{c}{\|A\|}$ .

 $Proof \ A = \sum_{i} s_{i} u_{i} v_{i}^{T}, X^{T} A X' = \sum_{i} s_{i} \underbrace{\langle u_{i}, X \rangle}_{:=g_{i}} \underbrace{\langle v_{i}, X' \rangle}_{:=g_{i}} \cdot (g_{1}, \cdots, g_{n}) \perp (g'_{1}, \cdots, g'_{n}) \sim \mathcal{N}(0, I_{n}) \Rightarrow \mathbb{E}e^{\lambda X^{T} A X} = \prod_{i=1}^{n} \mathbb{E}e^{\lambda s_{i} g_{i} g'_{i}} = \prod_{i=1}^{n} \mathbb{E}e^{\lambda^{2} s_{i}^{2} g'_{i}/2} \leq \prod_{i=1}^{n} e^{C\lambda^{2} s_{i}^{2}} (\lambda^{2} s_{i}^{2} \leq c) \leq e^{C\lambda^{2} \|A\|_{F}^{2}} (\lambda^{2} \leq \frac{c}{\max_{i} s_{i}^{2}} = \frac{c}{\|A\|^{2}}).$ 

• X, X' independent sub-gaussian mean zero,  $\|X\|_{\psi_2} \leq K, \|X'\|_{\psi_2} \leq K, g, g' \sim \mathcal{N}(0, I_n), g \perp g'$ . Then  $\mathbb{E}e^{\lambda X^T A X'} \leq K$  $\mathbb{E}e^{CK^2\lambda g^TAg'}$ 

 $Proof \ \ \text{Conditioned on } X', \mathbb{E}_X e^{\lambda X^A X'} \leq e^{C\lambda^2 K^2 \|AX'\|_2^2}, \mathbb{E}_g e^{\mu g^T A X'} = e^{\frac{\mu^2 \|AX'\|_2^2}{2}}. \ \ \mu = \sqrt{2c} K\lambda \Rightarrow \mathbb{E}_X e^{\lambda X^T A X'} \leq \mathbb{E}_g e^{\sqrt{2c} K\lambda g^T A X'} \Rightarrow \mathbb{E}_g e^{\lambda X^T A X'} \leq \mathbb{E}_g e^{\lambda X^T A$  $\mathbb{E}e^{\lambda X^T A X'} < \mathbb{E}e^{\sqrt{2c}K\lambda g^T A X'} < \mathbb{E}e^{2cK^2\lambda g^T A g'}$ 

• Hanson-Wright inequality:  $X = (X_1, \dots, X_n)$  independent mean zero sub-gaussian, then  $\mathbb{P}(|X^TAX - \mathbb{E}X^TAX| \geq$  $t) \le 2e^{-c\min(\frac{t^2}{K^4 \|A\|_F^2}, \frac{t}{K^2 \|A\|})}$ 

Proof WLOG K=1.  $X^TAX=\sum_{i,j}a_{ij}X_iX_j$ ,  $\mathbb{E}X^TAX=\sum_ia_{ii}\mathbb{E}X_i^2$ ,  $X^TAX-\mathbb{E}X^TAX=\sum_ia_{ii}(X_i^2-\mathbb{E}X_i^2)+\sum_{i\neq j}a_{ij}X_iX_j$ .  $p := \mathbb{P}(X^T A X - \mathbb{E}X^T A X \ge t) \le \mathbb{P}(\sum_{i=1}^{\infty} a_{ii}(X_i^2 - \mathbb{E}X_i^2) \ge \frac{t}{2}) + \mathbb{P}(\sum_{i \ne j} a_{ij} X_i X_j \ge \frac{t}{2}) := p_1 + p_2.$ 

 $\text{Step 1: } \|X_i^2 - \mathbb{E}X_i^2\|_{\psi_1} \lesssim 1. \text{ Bernstein} \Rightarrow p_1 \leq e^{-c\min(\frac{t^2}{\sum a_{ii}^2}, \frac{t}{\max_i |a_{ii}|})} \leq e^{-c\min(\frac{t^2}{\|A\|_F^2}, \frac{t}{\|A\|})}.$ 

Step 2:  $S := \sum_{i \neq j} a_{ij} X_i X_j$ .  $p_2 \leq e^{-\frac{\lambda t}{2}} \mathbb{E} e^{\lambda S}$ ,  $\mathbb{E} e^{\lambda S} \leq \mathbb{E} e^{4\lambda X^T A X'} \leq \mathbb{E} e^{c_1 \lambda g^T A g'} \leq e^{C\lambda^2 \|A\|_F^2}$  (with  $\lambda \leq \frac{c}{\|A\|}$ ).

•  $B_{m \times n}, X \in \mathbb{R}^n, \{X_i\}$  independent mean-zero, unit-variance, sub-gaussian. Then  $\|\|BX\|_2 - \|B\|_F\|_{\psi_2} \le CK^2\|B\|$ ,  $K = \max_i \|X\|_{\psi_2}.$ 

 $Proof \ A = B^T B, X^T A X = \|BX\|_2^2, \mathbb{E}X^T A X = \|B\|_F^2, \|A\| = \|B\|^2, \|A\|_F = \|B^T B\|_F \le \|B^T\| \|B\|_F = \|B\| \|B\|_F. \text{Thus } \forall u \ge 0,$  $\mathbb{P}(|\|BX\|_2^2 - \|B\|_F^2| \ge u) \le e^{-\frac{c}{K^4}\min(\frac{u^2}{\|B\|^2\|B\|_F^2}, \frac{u}{\|B\|^2})}. \text{ Let } u = \epsilon \|B\|_F^2, \mathbb{P}(|\|BX\|_2^2 - \|B\|_F^2) \le 2e^{-c\min(\epsilon^2, \epsilon)\frac{\|B\|_F^2}{K^4\|B\|^2}}. \text{ Let } u = \epsilon \|B\|_F^2, \mathbb{P}(|\|BX\|_2^2 - \|B\|_F^2) \le 2e^{-c\min(\epsilon^2, \epsilon)\frac{\|B\|_F^2}{K^4\|B\|^2}}.$  $\delta^2 = \min(\epsilon^2, \epsilon), \text{ then } \epsilon = \max(\delta, \delta^2), |||BX|| - ||B||_F| \ge \delta ||B||_F \Rightarrow |||BX||_2^2 - ||B||_F^2| \ge \epsilon ||B||_F^2 \Rightarrow \mathbb{P}(|||BX||_2 - ||B||_F| \ge \delta ||B||_F) \le \delta ||B||_F \Rightarrow ||BX||_2^2 - ||B||_F^2| \ge \epsilon ||B||_F^2 \Rightarrow ||BX||_2^2 - ||B||_F^2| \ge \delta ||B||_F \Rightarrow |$  $2e^{-c\delta^2 \frac{\|B\|_F^2}{K^4 \|B\|^2}}$ 

#### RANDOM PROCESSES

- Symmetrization:  $X_1, X_2, \dots, X_N$  independent, mean zero in a normed space,  $\epsilon_1, \epsilon_2, \dots, \epsilon_N$  a sequence of independent symmetric Bernoulli random variables. Then  $\frac{1}{2}\mathbb{E}\|\sum_{i=1}^N \epsilon_i X_i\| \leq \mathbb{E}\|\sum_{i=1}^N X_i\| \leq 2\mathbb{E}\|\sum_{i=1}^N \epsilon_i X_i\|$ .

  Proof Upper bound.  $X' \perp X, X' \stackrel{\text{d}}{=} X$ .  $p = \mathbb{E}\|\sum_{i=1}^N X_i\| \leq \mathbb{E}\|\sum_{i=1}^N X_i \sum_{i=1}^N X_i'\| = \mathbb{E}\|\sum_{i=1}^N \epsilon_i (X_i X_i')\| \leq \mathbb{E}\|\sum_{i=1}^N \epsilon_i X_i\| + \mathbb{E}\|\sum_{i=1}^N \epsilon_i X_i'\| = 2\mathbb{E}\|\sum_{i=1}^N \epsilon_i X_i\|.$
- $A_{n \times n}$  symmetric independent mean zero. Then  $\mathbb{E}||A|| \leq C\sqrt{\log n}\mathbb{E}\max||A_i||_2$  where  $A_i$  is *i*-th row of A.

$$Proof \ A = \sum_{i \leq j} Z_{ij} \text{ independent mean zero symmetric where } Z_{ij} = \begin{cases} A_{ij} (e_i e_j^T + e_j e_i^T), & i \leq j \\ A_{ii} e_i e_i^T & i = j \end{cases} \Rightarrow \mathbb{E} \|A\| \leq 2\mathbb{E} \|\sum_{i \leq j} \epsilon_{ij} Z_{ij}\|.$$

$$\text{Ex 5.4.3(a)} \Rightarrow \text{Conditioned on } \{Z_{ij}\}, \mathbb{E}_{\epsilon} \|\sum_{i \leq j} \epsilon_{ij} Z_{ij}\| \leq C\sqrt{\log n} \|\sum_{i \leq j} Z_{ij}^2\|^{\frac{1}{2}} \Rightarrow \mathbb{E} \|\sum_{i \leq j} \epsilon_{ij} Z_{ij}\| \leq C\sqrt{\log n} \mathbb{E} \|\sum_{i \leq j} Z_{ij}^2\|^{\frac{1}{2}},$$

$$\sum_{i \leq j} Z_{ij}^2 = \sum_{i=1}^n (\sum_{j=1}^n A_{ij}^2) e_i e_i^T = \sum_{i=1}^n \|A_i\|_2^2 e_i e_i^T \Rightarrow \|\sum_{i \leq j} Z_{ij}^2\|^{\frac{1}{2}} = \max \|A_i\|_2.$$

- Matrix completion:  $X_{n\times n}$ ,  $\operatorname{rank}(X) = r << n, Y_{ij} = \delta_{ij}X_{ij}, \delta_{ij} \sim \operatorname{Ber}(p), p = \frac{m}{n^2}, \hat{X} = \arg\min_{\operatorname{rank}(A') \leq r} \|p^{-1}Y A'\|$ . Then  $\mathbb{E}\frac{1}{n}\|\hat{X} X\|_F \leq C\sqrt{\frac{rn\log n}{m}}\|X\|_{\infty}$ .
  - $\begin{aligned} & Proof \ \ \text{Step 1.} \ \|\hat{X} X\| \leq \|\hat{X} p^{-1}Y\| + \|p^{-1}Y X\| \leq 2\|p^{-1}Y X\| = \frac{2}{p}\|Y pX\|. \ \ (Y pX)_{ij} = (\delta_{ij} p)X_{ij} \ \text{independent mean zero, Ex } 6.5.2 \Rightarrow \mathbb{E}\|Y pX\| \leq C\sqrt{\log n}(\mathbb{E}\max_i \|(Y pX)_i\|_2 + \mathbb{E}\max_j \|(T pX)^j\|_2). \ \ \|(Y pX)_i\|_2^2 = \sum_{j=1}^n (\delta_{ij} p)^2 X_{ij}^2 \leq \sum_{j=1}^n (\delta_{ij} p)^2 \|X\|_\infty^2. \end{aligned}$   $\sum_{j=1}^n (\delta_{ij} p)^2 \|X\|_\infty^2. \ \ \text{Ex } 6.6.2 \Rightarrow \mathbb{E}\max_i \sum_{j=1}^n (\delta_{ij} p)^2 \leq Cpn \Rightarrow \frac{2}{p}\|Y pX\| \leq C\sqrt{\frac{n\log n}{p}}\|X\|_\infty.$

Step 2. 
$$\operatorname{rank}(X) \leq r, \operatorname{rank}(\hat{X}) \leq r, \operatorname{rank}(\hat{X} - X) \leq 2r. \|\hat{X} - X\|_F \leq \sqrt{2r} \|\hat{X} - X\| \Rightarrow \mathbb{E} \|\hat{X} - X\|_F \leq C \sqrt{\frac{rn \log n}{p}} \|X\|_{\infty}.$$

• Contraction principle:  $X_1, \dots, X_N$  vectors in some normed space,  $a = (a_1, \dots, a_N) \in \mathbb{R}^N$ . Then  $\mathbb{E}\|\sum_{i=1}^N a_i \epsilon_i X_i\| \le \|a\|_{\infty} \mathbb{E}\|\sum_{i=1}^N \epsilon_i X_i\|$ .

Proof WLOG 
$$||a||_{\infty} \leq 1$$
,  $f(a) = \mathbb{E}||\sum_{i=1}^{N} a_i \epsilon_i X_i||$  is convex, which implies the maximum of  $f$  is attained at the boundary. Thus  $f(a) \leq f(a^*) = \mathbb{E}||\sum_{i=1}^{N} \epsilon_i X_i||$  with  $a_i^* = 1$  or  $-1$ .

• Symmetrization with gaussians:  $X_1, \dots, X_N$  independent mean zero,  $g_1, \dots, g_N \overset{\text{i.i.d.}}{\sim} \mathcal{N}(0, 1), \frac{C}{\sqrt{\log N}} \mathbb{E} \| \sum_{i=1}^N g_i X_i \| \le \mathbb{E} \| \sum_{i=1}^N X_i \| \le 3 \mathbb{E} \| \sum_{i=1}^N g_i X_i \|.$ 

 $Proof \ \ \text{Upper: } \mathbb{E}\|\sum_{i=1}^{N}X_i\| \leq 2\mathbb{E}\|\sum_{i=1}^{N}\epsilon_iX_i\| = 2\sqrt{\frac{\pi}{2}}\mathbb{E}_{X,\epsilon}\|\sum_{i=1}^{N}\epsilon_i\mathbb{E}_g|g_i|X_i\| \leq 2\sqrt{\frac{\pi}{2}}\mathbb{E}\|\sum_{i=1}^{N}\epsilon_i|g_i|X_i\| = 2\sqrt{\frac{\pi}{2}}\mathbb{E}\|\sum_{i=1}^{N}g_iX_i\|.$   $\text{Lower: } \mathbb{E}\|\sum_{i=1}^{N}g_iX_i\| = \mathbb{E}\|\sum_{i=1}^{N}\epsilon_ig_iX_i\| \leq \mathbb{E}_g\mathbb{E}_X(\|g\|_{\infty}\mathbb{E}_{\epsilon}\|\sum_{i=1}^{N}\epsilon_iX_i\|) = \mathbb{E}_g\|g\|_{\infty}\mathbb{E}_{X,\epsilon}\|\sum_{i=1}^{N}\epsilon_iX_i\| \leq 2\mathbb{E}_g\|g\|_{\infty}\mathbb{E}_X\|\sum_{i=1}^{N}X_i\| \leq C\sqrt{\log N}\mathbb{E}_X\|\sum_{i=1}^{N}X_i\|.$ 

# 7 Random processes

- Basic concepts:  $\{X_t\}_{t \in T \subset \mathbb{R}^n}$ ,  $\mathbb{E}X_t = 0$ ,  $\forall t \in T$ ,  $\Sigma(t, s) = \text{Cov}(X_t, X_s) = \mathbb{E}X_t X_s$ ,  $d(t, s) = \|X_t X_s\|_{L^2} = (\mathbb{E}(X_t X_s)^2)^{\frac{1}{2}}$  (increments).
- Gaussian process:  $T_0 \subset T, |T_0| < \infty, \{X_t\}_{t \in T_0}$  has normal distribution.
- Y is a mean zero Gaussian r.v. in  $\mathbb{R}^n$ . Then there exists  $t_1, \dots, t_n \in \mathbb{R}^n$  s.t.  $Y \stackrel{\mathrm{d}}{=} (\langle g, t_i \rangle)_{i=1}^n, g \sim \mathcal{N}(0, I_n)$ .
- Gaussian integration by parts:  $X \sim \mathcal{N}(0,1)$ . Then for f differentiable,  $\mathbb{E}f'(X) = \mathbb{E}Xf(X)$ . Proof f has bounded support:  $\mathbb{E}f(X) = \int_{\mathbb{R}} f'(x)\phi(x)dx = -\int_{\mathbb{R}} f(x)\phi'(x)dx$ . General  $f: f_n \to f$ .
- $X \sim \mathcal{N}(0, \Sigma), f : \mathbb{R}^n \to \mathbb{R}$ . Then  $\mathbb{E}X f(X) = \Sigma \cdot \mathbb{E}\nabla f(X)$ .
- $X \sim \mathcal{N}(0, \Sigma^X), Y \sim \mathcal{N}(0, \Sigma^Y), X \perp Y, Z(u) = \sqrt{u}X + \sqrt{1 u}Y, u \in [0, 1]$ . Then  $f : \mathbb{R}^n \to \mathbb{R}$  twice-differentiable,  $\frac{d}{du}\mathbb{E}f(Z(u)) = \frac{1}{2}\sum_{i,j=1}^n (\Sigma_{ij}^X \Sigma_{ij}^Y)\mathbb{E}[\frac{\partial^2 f}{\partial x_i \partial x_j}(Z(u))].$

$$Proof \frac{d}{du}\mathbb{E}f(Z(u)) = \frac{1}{2}\sum_{i=1}^{n}\mathbb{E}\frac{\partial f}{\partial x_{i}}(Z(u))(\frac{X_{i}}{\sqrt{u}} - \frac{Y_{i}}{\sqrt{1-u}}). \sum_{i=1}^{n}\frac{1}{\sqrt{u}}\mathbb{E}X_{i}\frac{\partial f}{\partial x_{i}}(Z(u)) := \sum_{i=1}^{n}\frac{1}{\sqrt{u}}\mathbb{E}X_{i}g_{i}(X) \text{ (conditioned on } Y) \text{ where } g_{i}(X) := \frac{\partial f}{\partial x_{i}}(\sqrt{u}X + \sqrt{1-u}Y). \quad \mathbb{E}X_{i}g_{i}(X) = \sum_{j=1}^{n}\sum_{ij}^{X}\mathbb{E}\frac{\partial g_{i}}{\partial x_{j}}(X) = \sum_{j=1}^{n}\sum_{ij}^{X}\mathbb{E}\frac{\partial^{2} f}{\partial x_{i}x_{j}}(\sqrt{u}X + \sqrt{1-u}Y)\sqrt{u}. \quad \Box$$

- $\mathbb{E}X_i^2 = \mathbb{E}Y_i^2$ ,  $\mathbb{E}(X_i X_j)^2 \le \mathbb{E}(Y_i Y_j)^2$ ,  $\frac{\partial^2 f}{\partial x_i \partial x_j} \ge 0$ ,  $\forall i \ne j$ . Then  $\mathbb{E}f(X) \ge \mathbb{E}f(Y)$ .
- Slepian's inequality: Let  $\{X_t\}_{t\in T}$  and  $\{Y_t\}_{t\in T}$  be two mean zero Gaussian processes. Assume  $\mathbb{E}X_t^2 = \mathbb{E}Y_t^2$ ,  $\mathbb{E}(X_t X_s)^2 \leq \mathbb{E}(Y_t Y_s)^2$ . Then for every  $t \in \mathbb{R}$ ,  $\mathbb{P}(\sup_{t \in T} X_t \geq t) \leq \mathbb{P}(\sup_{t \in T} Y_t \geq t)$  and  $\mathbb{E}\sup_{t \in T} X_t \leq \mathbb{E}\sup_{t \in T} Y_t$ .

*Proof* Let  $f(x) \approx 1_{\{\max x_i < t\}} = \prod_{i=1}^n 1_{\{x_i < t\}}$  and use the latest lemma.

• Sudakov-Fernique's inequality: Let  $\{X_t\}_{t\in T}$  and  $\{Y_t\}_{t\in T}$  be two mean zero Gaussian processes. Assume  $\mathbb{E}(X_t-X_s)^2 \leq \mathbb{E}(Y_t-Y_s)^2$ . Then  $\mathbb{E}\sup_{t\in T}X_t \leq \mathbb{E}\sup_{t\in T}Y_t$ .

Proof Let 
$$f(x) = \frac{1}{\beta} \log \sum_{i=1}^{n} e^{\beta x_i}$$
.  $f(x) \to \max x_i$  as  $\beta \to \infty$ .  $\frac{\partial^2 f}{\partial x_i \partial x_i} \leq 0$ .

•  $A_{m \times n}$  independent  $\mathcal{N}(0,1)$  entries. Then  $\mathbb{E}||A|| \leq \sqrt{m} + \sqrt{n}$ .

Proof 
$$\max_{u \in \mathbb{S}^{n-1}, v \in \mathbb{S}^{m-1}} \langle Au, v \rangle := \max_{(u,v) \in T} X_{uv}$$
 where  $T = \mathbb{S}^{n-1} \times \mathbb{S}^{m-1}$  and  $X_{uv} \sim \mathcal{N}(0,1)$ .  $\mathbb{E}(X_{uv} - X_{wz})^2 = \mathbb{E}(\langle Au, v \rangle - \langle Aw, z \rangle)^2 = \mathbb{E}(\sum_{i,j} A_{ij}(u_j v_i - w_j z_i))^2 = \sum_{i,j} (u_j v_i - w_j z_i)^2 = \|uv^T - wz^T\|_F \le \|u - w\|_2^2 + \|v - z\|_2^2$ . Define  $Y_{uv} = \langle g, u \rangle + \langle h, v \rangle, g \sim \mathcal{N}(0, I_n), h \sim \mathcal{N}(0, I_m), g \perp h$ .  $\mathbb{E}(Y_{uv} - Y_{wz})^2 = \|u - w\|_2^2 + \|v - z\|_2^2$ . Then  $\mathbb{E}\|A\| = \mathbb{E}\sup_{(u,v) \in T} X_{uv} \le \mathbb{E}_{(u,v) \in T} Y_{uv} = \mathbb{E}\sup_{u \in \mathbb{S}^{n-1}} \langle g, u \rangle + \mathbb{E}\sup_{v \in \mathbb{S}^{m-1}} \langle g, v \rangle = \mathbb{E}\|g\|_2 + \mathbb{E}\|h\|_2 \le \sqrt{n} + \sqrt{m}$ .

•  $\mathbb{P}(\|A\| \ge \sqrt{m} + \sqrt{n} + t) \le 2e^{-ct^2}$ .

Proof 
$$A \sim \mathcal{N}(0, I_{nm}), f(A) = ||A|| \le ||A||_2 \Rightarrow ||f||_{\text{Lip}} \le 1 \Rightarrow ||f(A) - \mathbb{E}f(A)||_{\psi_2} \le C.$$

- Sudakov's minoration inequality:  $\{X_t\}_{t\in T}$  mean zero Gaussian process.  $\forall \epsilon > 0, \mathbb{E}\sup_{t\in T} X_t \geq C\epsilon\sqrt{\log \mathcal{N}(T,d,\epsilon)}$ . Proof Assume  $\mathcal{N}(T,d,\epsilon) = N < \infty$ . Let  $\mathscr{N}$  be a maximal  $\epsilon$ -separated subset of T.  $|\mathscr{N}| \geq N$ . It suffices to show  $\mathbb{E}\sup_{t\in\mathscr{N}} X_t \geq C\epsilon\sqrt{\log N}$ .  $Y_t := \frac{\epsilon}{\sqrt{2}}g_t, g_t \overset{\text{i.i.d.}}{\sim} \mathcal{N}(0,1)$ .  $\mathbb{E}(X_t - X_s)^2 = d(t,s)^2 \geq \epsilon^2 = \mathbb{E}(Y_t - Y_s)^2 \Rightarrow \mathbb{E}\sup_{t\in\mathscr{N}} X_t \geq \mathbb{E}\sup_{t\in\mathscr{N}} Y_t = C\epsilon\sqrt{\log N}$ .
- $X_t = \langle g, t \rangle, g \sim \mathcal{N}(0, I_n), d(s, t) = ||t s||_2, \mathbb{E} \sup_{t \in T} \langle g, t \rangle \ge C\epsilon \sqrt{\log \mathcal{N}(T, \epsilon)}.$
- P a polytope in  $\mathbb{R}^n$  with N vertices, diameter is bounded by 1. Then for  $\epsilon > 0, \mathcal{N}(P, \epsilon) \leq N^{c/\epsilon^2}$ .

Proof 
$$x_1, x_2, \dots, x_N$$
 vertices of  $P$ ,  $\mathbb{E} \sup_{t \in P} \langle g, t \rangle = \mathbb{E} \sup_{i < N} \langle g, x_i \rangle \leq C \sqrt{\log N}$ .

- Gaussian width:  $\omega(T) := \mathbb{E} \sup_{x \in T} \langle g, x \rangle, g \sim \mathcal{N}(0, I_n).$
- Properties of Gaussian width: (a)  $\omega(T) < \infty \Leftrightarrow T$  is bounded; (b) For every orthogonal matrix U and vector y,  $\omega(UT+y) = \omega(T)$ ; (c)  $\omega(\operatorname{conv}(T)) = \omega(T)$ ; (d)  $\omega(T+S) = \omega(T) + \omega(S), \omega(aT) = |a|\omega(T)$ ; (e)  $\omega(T) = \frac{1}{2}\omega(T-T) = \frac{1}{2}\mathbb{E}\sup_{x,y\in T}\langle g, x-y\rangle$ ; (f)  $\frac{1}{\sqrt{2\pi}}\operatorname{diam}(T) \leq \omega(T) \leq \frac{\sqrt{n}}{2}\operatorname{diam}(T)$ .

Proof (e):  $\omega(T) = \frac{1}{2}(\omega(T) + \omega(T)) = \frac{1}{2}(\omega(T) + \omega(-T)) \stackrel{\text{(d)}}{=} \frac{1}{2}\omega(T - T).$ 

(f): Lower bound. Fix  $x, y \in T, x - y, y - x \in T - T, \omega(T) \ge \frac{1}{2} \mathbb{E} \max(\langle x - y, g \rangle, \langle y - x, g \rangle) = \frac{1}{2} \mathbb{E} |\langle x - y, g \rangle| = \sqrt{\frac{1}{2\pi}} ||x - y||_2.$ 

Upper bound.  $\omega(T) = \frac{1}{2} \mathbb{E} \sup_{x,y \in T} \langle g, x - y \rangle \leq \frac{1}{2} \mathbb{E} \sup_{x,y \in T} \|g\|_2 \|x - y\|_2 = \frac{1}{2} \mathbb{E} \|g\|_2 \operatorname{diam}(T) \text{ and } \mathbb{E} \|g\|_2 \leq (\mathbb{E} \|g\|^2)^{\frac{1}{2}} = \sqrt{n}.$ 

- Spherical width:  $\omega_{\mathbb{S}}(T) = \mathbb{E} \sup_{x \in T} \langle \theta, x \rangle, \theta \sim \text{Unif}(\mathbb{S}^{n-1}).$
- $(\sqrt{n} C)\omega_{\mathbb{S}}(T) \le \omega(T) \le (\sqrt{n} + C)\omega_{\mathbb{S}}(T)$ .

 $Proof \ g = \|g\|_2 \cdot \frac{g}{\|g\|_2} := r \cdot \theta, r \perp \theta. \ \omega(T) = \mathbb{E} \sup_{x \in T} \langle r\theta, x \rangle = \mathbb{E} r \mathbb{E} \sup_{x \in T} \langle \theta, x \rangle = \mathbb{E} \|g\|_2 \omega_{\mathbb{S}}(T). \ \text{Ex } 3.1.4 \Rightarrow |\mathbb{E} \|g\|_2 - \sqrt{n}| \leq C.\square$ 

- Squared version of the Gaussian width:  $h(T)^2 = \mathbb{E} \sup_{t \in T} \langle g, t \rangle^2, g \sim \mathcal{N}(0, I_n)$ .
- Stable dimension: bounded  $T \subset \mathbb{R}^n$ ,  $d(T) := \frac{h(T-T)^2}{\operatorname{diam}^2(T)} \asymp \frac{\omega^2(T)}{\operatorname{diam}^2(T)}$ .
- $d(T) \leq \dim(T)$ .

$$Proof \ \ \text{Let} \ \dim(T) = k \ \text{and} \ T \subset \mathbb{R}^k. \ \ h(T-T)^2 = \mathbb{E} \sup_{x,y \in T} \langle g, x-y \rangle^2. \ \ x-y = \operatorname{diam}(T) \cdot z \ \text{for some} \ z \in B_2^k. \ \therefore \ h(T-T)^2 \leq \operatorname{diam}^2(T) \mathbb{E} \sup_{z \in B_2^k} \langle g, z \rangle^2 = \operatorname{diam}^2(T) \mathbb{E} \|g\|_2^2 = \operatorname{diam}^2(T) \cdot k.$$

- Stable rank:  $A_{m \times n}, r(A) := \frac{\|A\|_F^2}{\|A\|^2} = d(AB_2^n) \le \text{rank}(A) = \dim(AB_2^n)$ .
- Gaussian complexity:  $\gamma(T) := \mathbb{E} \sup_{x \in T} |\langle g, x \rangle|, g \sim \mathcal{N}(0, I_n).$
- $T \subset \mathbb{R}^n$ ,  $\mathcal{P}$  projection onto  $E \sim \text{Unif}(G_{n,m})$ .  $\forall m \leq n$ , with probability at least  $1 2e^{-m}$ ,  $\text{diam}(\mathcal{P}T) \leq C[\omega_{\mathbb{S}}(T) + \sqrt{\frac{m}{n}} \text{diam}(T)]$ .

#### **CHAINING**

Proof Step 1: Approximation. WLOG diam $(T) \leq 1$ .  $Q_{n \times n}$ : choosing the first m rows of  $U_{n \times n} \sim \text{Unif}(O(n))$ . Then  $\|\mathcal{P}x\|_2 \stackrel{\text{d}}{=} \|Qx\|_2, \forall x \in \mathbb{R}^n$ .  $Q^Tz \sim \text{Unif}(\mathbb{S}^{n-1}), \forall z \in \mathbb{S}^{m-1}$ . diam $(\mathcal{P}T) \stackrel{\text{d}}{=} \text{diam}(QT) = \sup_{x \in T-T} \|Qx\|_2 = \sup_{x \in T-T} \max_{z \in \mathbb{S}^{m-1}} \langle Qx, z \rangle = \sup_{x \in T-T} \max_{z \in \mathbb{S}^{m-1}} \langle x, Q^Tz \rangle$ . Choose an  $\frac{1}{2}$ -net  $\mathcal{N}$  of  $\mathbb{S}^{m-1}, |\mathcal{N}| \leq 5^m \Rightarrow \text{diam}(QT) \leq 2 \max_{z \in \mathcal{N}} \sup_{x \in T-T} \langle x, Q^Tz \rangle$ .

Step 2: Concentration. For  $z \in \mathcal{N}$ ,  $\mathbb{E}\sup_{x \in T-T} \langle Q^T z, x \rangle = \omega_{\mathbb{S}}(T-T) = 2\omega_{\mathbb{S}}(T)$ . The function  $f: Q^T z \to \sup_{x \in T-T} \langle Q^T z, x \rangle$  is Lipschitz on  $\mathbb{S}^{n-1} \Rightarrow \langle Q^T z, x \rangle$  is sub-gaussian  $\Rightarrow \mathbb{P}(\sup_{x \in T-T} \langle Q^T z, x \rangle \geq 2\omega_{\mathbb{S}}(T) + t) \leq 2e^{-cnt^2}$ .

Step 3: Union bound.  $\mathbb{P}(\max_{z \in \mathcal{N}} \sup_{x \in T - T} \langle Q^T z, x \rangle \ge 2\omega_{\mathbb{S}}(T) + t) \le 5^m 2e^{-cnt^2} \le 2e^{-m} \ (t = C\sqrt{\frac{m}{n}} \text{ and } C \text{ large enough}).$ 

- Phase transition: Equivalently write it as  $\operatorname{diam}(\mathcal{P}T) \leq C \max(\omega_{\mathbb{S}}(T), \sqrt{\frac{m}{n}} \operatorname{diam}(T))$ . Set  $\omega_{\mathbb{S}}(T) = \sqrt{\frac{m}{n}} \operatorname{diam}(T) \Rightarrow m = \frac{(\sqrt{n}\omega_{\mathbb{S}}(T))^2}{\operatorname{diam}^2(T)} \approx \frac{\omega^2(T)}{\operatorname{diam}^2(T)} \approx d(T)$ . That is, if m > d(T),  $\operatorname{diam}(\mathcal{P}T) \leq C\sqrt{\frac{m}{n}} \operatorname{diam}(T)$ ; if m < d(T),  $\operatorname{diam}(\mathcal{P}T) \leq C\omega_{\mathbb{S}}(T)$ .
- Random matrix  $G_{m \times n}$  with independent  $\mathcal{N}(0,1)$  entries.  $\forall m \leq n$ , with probability at least  $1-2e^{-m}$ , diam $(GT) \leq C[\omega(T) + \sqrt{m} \operatorname{diam}(T)]$ .

### 8 Chaining

- Sub-gaussian increments:  $\{X_t\}_{t\in T}, (T,d)$ . Exist  $K\geq 0$ , s.t. $\|X_t-X_s\|_{\psi_2}\leq Kd(t,s)$  for  $t,s\in T$ .
- $\{X_t\}_{t\in T}, \mathbb{E}X_t = 0, (T,d)$ , sub-gaussian increments. Then  $\mathbb{E}\sup_{t\in T}X_t \leq CK\sum_{k\in\mathbb{Z}}2^{-k}\sqrt{\log\mathcal{N}(T,d,2^{-k})}$ .

Proof Step 1: Chaining setup. WLOG assume K = 1 and T is finite.  $\epsilon_k = 2^{-k}, k \in \mathbb{Z}$ . Choose an  $\epsilon_k$ -net  $T_k$  of T so that  $|T_k| = \mathcal{N}(T, d, \epsilon_k)$ . T is finite  $\Rightarrow \exists$  small enough  $\kappa \in \mathbb{Z}$  and large enough  $K \in \mathbb{Z}$  s.t.  $T_{\kappa} = \{t_0\}, T_k = T$ . For a point  $t \in T$ ,  $\pi_k(t)$ : a closest point in  $T_k$ . Then  $d(t, \pi_k(t)) \leq \epsilon_k$ .  $\mathbb{E}X_{t_0} = 0 \Rightarrow \mathbb{E}\sup_{t \in T} X_t = \mathbb{E}\sup_{t \in T} (X_t - X_{t_0})$ . Since  $X_t - X_{t_0} = \sum_{k=\kappa+1}^K (X_{\pi_k(t)} - X_{\pi_{k-1}(t)})$ ,  $\mathbb{E}\sup_{t \in T} (X_t - X_{t_0}) \leq \sum_{k=\kappa+1}^K \mathbb{E}\sup_{t \in T} (X_{\pi_k(t)} - X_{\pi_{k-1}(t)})$ .

Step 2: Control the increments.  $||X_{\pi_k(t)} - X_{\pi_{k-1}(t)}||_{\psi_2} \le d(\pi_k(t), \pi_{k-1}(t)) \le \epsilon_k + \epsilon_{k-1} \le 2\epsilon_{k-1}$ . Ex 2.5.10  $\Rightarrow \mathbb{E}\sup_{t \in T} (X_{\pi_k(t)} - X_{\pi_{k-1}(t)}) \le C \cdot 2\epsilon_{k-1} \cdot \sqrt{\log(|T_k| \cdot |T_{k-1}|)} \le C \cdot 2\epsilon_{k-1} \sqrt{2\log|T_k|}$ .

Step 3: Summing up the increments.  $\mathbb{E}\sup_{t\in T}(X_t-X_{t_0}) \leq C\sum_{k=\kappa+1}^K \epsilon_{k-1}\sqrt{\log|T_k|}$ .

• Dudley's integral inequality:  $\mathbb{E} \sup_{t \in T} X_t \leq CK \int_0^\infty \sqrt{\log \mathcal{N}(T, d, \epsilon)} d\epsilon$ .

Proof 
$$2^{-k} = 2 \int_{2-k-1}^{2^{-k}} d\epsilon$$
.  $\sum_{k \in \mathbb{Z}} 2^{-k} \sqrt{\log \mathcal{N}(T, d, 2^{-k})} \le 2 \sum_{k \in \mathbb{Z}} \int_{2-k-1}^{2^{-k}} \sqrt{\log \mathcal{N}(T, k, \epsilon)} d\epsilon$ .

- $\mathbb{E} \sup_{t,s \in T} |X_t X_s| \le CK \int_0^\infty \sqrt{\log \mathcal{N}(T,d,\epsilon)} d\epsilon$ .
- $T \subset \mathbb{R}^n, \omega(T) \leq C \int_0^{+\infty} \sqrt{\log \mathcal{N}(T, \epsilon)} d\epsilon$ .
- $T \subset \mathbb{R}^n, s(T) := \sup_{\epsilon > 0} \epsilon \sqrt{\log \mathcal{N}(T, \epsilon)}$ . Then  $cs(T) \le \omega(T) \le Cs(T) \log n$ .
- Empirical process:  $f \in \mathcal{F}, f : \Omega \to \mathbb{R}, (\Omega, \Sigma, \mu)$ . X is a random point in  $\Omega$ .  $X \sim \mu$ .  $X_1, \dots, X_n$  i.i.d. copies of X.  $X_f := \frac{1}{n} \sum_{i=1}^n f(X_i) \mathbb{E}f(X)$  empirical process indexed by  $\mathcal{F}$ .
- $X_i \in [0,1], \mathscr{F} := \{ f : [0,1] \to \mathbb{R}, \|f\|_{\text{Lip}} \le L \}.$  Then  $\mathbb{E} \sup_{f \in \mathscr{F}} |\frac{1}{n} \sum_{i=1}^n f(X_i) \mathbb{E} f(X)| \le \frac{CL}{\sqrt{n}}.$  Proof WLOG  $\mathscr{F} = \{ f : [0,1] \to [0,1], \|f\|_{\text{Lip}} \le 1 \}.$

Step 1: Check sub-gaussian increments. Let  $Z_i := (f-g)(X_i) - \mathbb{E}(f-g)(X)$ . Then  $\forall f, g \in \mathscr{F}, \|X_f - X_g\|_{\psi_2} = \frac{1}{n} \|\sum_{i=1}^n Z_i\|_{\psi_2} \lesssim \frac{1}{n} (\sum_{i=1}^n \|Z_i\|_{\psi_2}^2)^{\frac{1}{2}}$ .  $\|Z_i\|_{\psi_2} \lesssim \|(f-g)(X_i)\|_{\psi_2} \lesssim \|f-g\|_{\infty}$ .

Step 2: Apply Dudley's inequality:  $\operatorname{diam}(\mathscr{F}) \leq 1$ .  $\mathbb{E}\sup_{f \in \mathscr{F}} |X_f| \leq \frac{1}{\sqrt{n}} \int_0^1 \sqrt{\log \mathcal{N}(\mathscr{F}, \|\cdot\|_{\infty}, \epsilon)} d\epsilon \leq \frac{1}{\sqrt{n}} \int_0^1 \sqrt{\frac{C}{\epsilon} \log \frac{C}{\epsilon}} d\epsilon \lesssim \frac{1}{\sqrt{n}}$ .