Linear Regression

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Geometric Interpretation

Column Space of the Design Matrix

- ▶ $\mathbf{1}_n : n \times 1$ vector of ones; $\mathbf{x} = (X_1, \dots, X_n)^T : n \times 1$ vector of X values.
- ▶ Design matrix: $\mathbf{X} = (\mathbf{1}_n, \mathbf{x})$.
- $ightharpoonup \operatorname{col}(X)$: linear subspace of \mathbb{R}^n generated by the columns of **X**

$$col(X) = \{ \mathbf{v} \in \mathbb{R}^n : \exists c_0, c_1 \in R, s.t., \mathbf{v} = c_0 \mathbf{1}_n + c_1 \mathbf{x} \}.$$

Projection to col(X)

- $\mathbf{H} = \mathbf{X}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'$ projects a vector in \mathbb{R}^n to $\operatorname{col}(\mathbf{X})$: For any $\mathbf{w} \in \mathbb{R}^n$
 - ► **Hw** ∈ col(X), i.e., there exists $c_0, c_1 \in \mathbf{R}$ such that **Hw** = $c_0 \mathbf{1}_n + c_1 \mathbf{x}$.
 - ▶ $\mathbf{w} \mathbf{H}\mathbf{w} \perp \operatorname{col}(X)$, i.e., for any $\mathbf{v} \in \operatorname{col}(X)$, the inner product $< \mathbf{w} \mathbf{H}\mathbf{w}, \mathbf{v} > = (\mathbf{w} \mathbf{H}\mathbf{w})^T \mathbf{v} = 0$.

Fitted Values and Residuals

$$\mathbf{\widehat{Y}} = \mathbf{HY} = \hat{\beta}_0 \mathbf{1}_n + \hat{\beta}_1 \mathbf{x} \in \operatorname{col}(X)$$

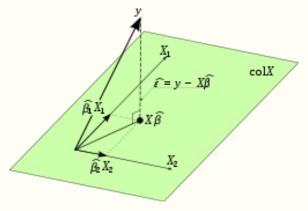
- ightharpoonup e = Y HY \perp col(X)
- ▶ Since $\mathbf{1}_n, \mathbf{x}, \widehat{\mathbf{Y}} \in \operatorname{col}(\mathbf{X})$, so

$$\langle \mathbf{e}, \mathbf{1}_n \rangle = \sum_{i=1}^n e_i = 0$$

 $\langle \mathbf{e}, \mathbf{x} \rangle = \sum_{i=1}^n X_i e_i = 0$
 $\langle \mathbf{e}, \widehat{\mathbf{Y}} \rangle = \sum_{i=1}^n \hat{Y}_i e_i = 0$

Geometric Interpretation

Figure: Orthogonal projection of response vector \mathbf{Y} onto the linear subspace of \mathbb{R}^n generated by the columns of the design matrix \mathbf{X}



Sums of Squares: Matrix Form

Error Sum of Squares

$$SSE = \sum_{i=1}^{n} e_i^2$$

can be expressed in matrix form:

$$SSE = e'e = Y'(I_n - H)'(I_n - H)Y = Y'(I_n - H)Y.$$

- ▶ I_n **H** is a projection matrix.
- $df(SSE) = rank(I_n H) = n 2.$

Total Sum of Squares

SSTO =
$$\sum_{i=1}^{n} (Y_i - \overline{Y})^2 = \sum_{i=1}^{n} Y_i^2 - n(\overline{Y})^2$$

can be expressed in matrix form:

$$SSTO = \mathbf{Y}'\mathbf{Y} - \frac{1}{n}\mathbf{Y}'\mathbf{J}_n\mathbf{Y} = \mathbf{Y}'\left(\mathbf{I}_n - \frac{1}{n}\mathbf{J}_n\right)\mathbf{Y}.$$

▶ $I_n - \frac{1}{n}J_n$ is a projection matrix.

$$\mathbf{J}_n = \mathbf{1}_n \mathbf{1}_n' = \begin{bmatrix} 1 & 1 & \cdots & 1 \\ 1 & 1 & \cdots & 1 \\ \vdots & \vdots & \cdots & \vdots \\ 1 & 1 & \cdots & 1 \end{bmatrix}$$

• $df(SSTO) = rank(\mathbf{I}_n - \frac{1}{n}\mathbf{J}_n) = n - 1.$

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Regression Sum of Squares

$$SSR = \sum_{i=1}^{n} (\widehat{Y}_i - \overline{Y})^2$$

can be expressed in matrix form:

$$SSR = (\widehat{\mathbf{Y}} - \overline{\mathbf{Y}})' (\widehat{\mathbf{Y}} - \overline{\mathbf{Y}}), \qquad \overline{\mathbf{Y}} = \frac{1}{n} \mathbf{J}_n \mathbf{Y}$$
$$= \mathbf{Y}' \left(\mathbf{H} - \frac{1}{n} \mathbf{J}_n \right)' \left(\mathbf{H} - \frac{1}{n} \mathbf{J}_n \right) \mathbf{Y}$$
$$= \mathbf{Y}' (\mathbf{H} - \frac{1}{n} \mathbf{J}_n) \mathbf{Y}$$

- ▶ $\mathbf{H} \frac{1}{n} \mathbf{J}_n$ is a projection matrix.
- $df(SSR) = rank(\mathbf{H} \frac{1}{n}\mathbf{J}_n) = 1.$

Expectation of SSE

$$E(SSE) = E(\mathbf{Y}'(\mathbf{I}_n - \mathbf{H})\mathbf{Y}) = E(Tr((\mathbf{I}_n - \mathbf{H})\mathbf{Y}\mathbf{Y}'))$$

$$= Tr((\mathbf{I}_n - \mathbf{H})E(\mathbf{Y}\mathbf{Y}'))$$

$$= Tr((\mathbf{I}_n - \mathbf{H})(\sigma^2\mathbf{I}_n + \mathbf{X}\boldsymbol{\beta}\boldsymbol{\beta}'\mathbf{X}'))$$

$$= \sigma^2Tr(\mathbf{I}_n - \mathbf{H}) + Tr((\mathbf{I}_n - \mathbf{H})\mathbf{X}\boldsymbol{\beta}\boldsymbol{\beta}'\mathbf{X}')$$

$$= (n-2)\sigma^2$$

The last equality is because $Tr(\mathbf{I}_n - \mathbf{H}) = n - 2$ and $(\mathbf{I}_n - \mathbf{H})\mathbf{X} = \mathbf{0}$.

Confidence Interval for σ^2

Under the Normal error model:

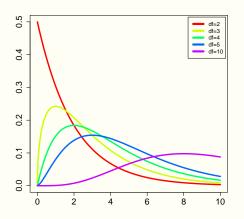
- ► SSE ~ $\sigma^2 \chi^2_{(n-2)}$
- Pivotal quantity:

$$\frac{SSE}{\sigma^2} \sim \chi^2_{(n-2)}$$

▶ $(1 - \alpha)100\%$ -confidence interval for σ^2 :

$$\left[\frac{SSE}{\chi^2(1-\alpha/2;n-2)}, \frac{SSE}{\chi^2(\alpha/2;n-2)}\right]$$

Probability Density Curves of χ^2 Distributions



Properties of Projection Matrices

Optional material.

- Eigen-decomposition: QΛQ^T, where Q is an orthogonal matrix of eigenvectors and Λ is a diagonal matrix of eigenvalues.
- Eigenvalues are either 1 or 0.
- The number of nonzero eigenvalues equals trace of the matrix equals the rank.
- For example. In simple linear regression:

$$rank(\mathbf{H}) = 2$$
, $rank(\mathbf{I}_n - \mathbf{H}) = n - 2$

Sampling Distribution of SSE under Normal Error Model

Optional material (Cont'd).

▶ $I_n - H$ is a projection matrix with rank $n - 2 \Longrightarrow$

$$\mathbf{I}_n - \mathbf{H} = \mathbf{Q}^T \wedge \mathbf{Q},$$

where $\Lambda = diag\{1, \dots, 1, 0, 0\}$ and **Q** is an orthogonal matrix.

 $ightharpoonup (I_n - H)X = 0 \Longrightarrow$

$$\mathbf{e} = (\mathbf{I}_n - \mathbf{H})\mathbf{Y} = (\mathbf{I}_n - \mathbf{H})(\mathbf{X}\boldsymbol{\beta} + \boldsymbol{\epsilon}) = (\mathbf{I}_n - \mathbf{H})\boldsymbol{\epsilon}$$

Optional material (Cont'd).

▶ Let $\mathbf{z} = \mathbf{Q} \boldsymbol{\epsilon}$, then

$$SSE = \sum_{i=1}^{n-2} z_i^2$$

Moreover

$$\mathbf{E}(\mathbf{z}) = \mathbf{Q}\mathbf{E}\{\boldsymbol{\epsilon}\} = \mathbf{0}, \quad \boldsymbol{\sigma^2}\{\mathbf{z}\} = \mathbf{Q}\boldsymbol{\sigma^2}\{\boldsymbol{\epsilon}\}\mathbf{Q}^T = \boldsymbol{\sigma^2}\mathbf{Q}\mathbf{Q}^T = \boldsymbol{\sigma^2}\mathbf{I}_n$$

So under Normal error model, z_i s are i.i.d. $N(0, \sigma^2)$.

► Thus $SSE \sim \sigma^2 \chi^2_{(n-2)}$.