

# On the fraught inference of historical human generation times

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## Abstract

WANG *et al.* (2023) recently proposed an approach to infer the history of human generation intervals from changes in mutation profiles over time. As the relative proportions of different mutation types depend on the ages of parents, stratifying variants by the time they arose allows for the inference of average paternal and maternal generation times at past times. Applying this approach to published allele age estimates, WANG *et al.* (2023) inferred long-lasting sex differences in average generation times and surprisingly found that ancestral generation times of West African populations remained substantially higher than those of Eurasian populations extending tens of thousands of generations into the past. Here we show that the results and interpretations in WANG *et al.* (2023) are primarily driven by noise and biases in input data and a lack of validation using independent approaches for estimating allele ages. With the recent development of methods to reconstruct genome-wide gene genealogies, coalescence times, and allele ages, we caution that downstream analyses may be strongly influenced by uncharacterized biases in their output.

Recent years have seen the rapid development of methods for reconstructing genealogical structures of large cohorts (SPEIDEL *et al.*, 2019; WOHNS *et al.*, 2022; HUBISZ *et al.*, 2020), which are comprised of a series of gene genealogies (or trees) along the genome. Reconstructed genealogies (or informative summaries of them) have the potential to transform population genetic inference, as biological and evolutionary processes impact the shape and correlation of gene trees and the distribution of variation that arises in the lineages they represent. In particular, the age of a variant can be estimated by mapping its mutation to the portion of the gene tree in which it is inferred to have occurred (ALBERS and McVEAN, 2020).

The past few years have also seen efforts to sequence large sets of pedigrees, providing increased resolution of important parameters of genome biology, including direct measurements of mutation rates and profiles. Through high-coverage sequencing of multiple generations of families, *de novo* mutations can be determined as maternally or paternally inherited, and with a large enough sample size both the number of mutations and proportions of mutation types (e.g., A→C, A→G, etc.) can be correlated with parental age and sex (JÓNSSON *et al.*, 2017; HALLDORSSON *et al.*, 2019). WANG *et al.* (2023) combined these two sets of inferences, the estimated ages of mutations and the parental age- and sex-dependence of the mutation spectrum, to infer the history of average maternal and paternal generation intervals for human populations of diverse ancestries. In order to avoid overfitting, this approach requires making a number of assumptions about the constancy of the mutational process over time ??, its similarity across populations, and of negligible effect of selection on shaping the profile of surviving mutation types.

This approach has recently been criticized (GAO *et al.*, 2022). Notably, GAO *et al.* (2022) show that observed changes in the mutation spectrum over time cannot be explained by changes in maternal and paternal generation intervals alone, as specific mutational signatures would require unique and divergent generation time histories to simultaneously explain them. GAO *et al.* (2022) also point out that pedigree-based estimates of the *de novo* mutation spectrum do not agree with the mutation spectrum among young

variants in existing population-level datasets, potentially biasing such approaches. They argue that factors other than changes in generation intervals, including genetic modifiers and environmental exposure, must explain observed variation in mutation profiles.

In this note, we examine both the results and conclusions of WANG *et al.* (2023). We first consider the reported inferred generation time histories and find that they are inconsistent with current understanding of human population history, in particular population structure within Africa. In exploring the source of this inconsistency, we show that allele age estimates are not just noisy, but age-stratified mutation spectra reconstructed using independent methods do not agree, with mutation profiles diverging in opposing directions. Thus, the results from WANG *et al.* (2023) do not reproduce. We further discuss the disagreement between the mutation rate profile found in pedigree studies and from young variants [APR: which cannot be accounted for in a satisfactory manner](#). In conclusion, we suggest that downstream analyses using estimated allele ages and mutation profiles should more carefully validate their results and such results should be interpreted with a heavy dose of skepticism.

## Long-lasting differences in population-specific generation intervals

Applying the proposed inference approach to multiple populations of different continental ancestries, WANG *et al.* (2023) estimated that the ancestors of European, East Asian, and South Asian populations included in the 1000 GENOMES PROJECT CONSORTIUM *et al.* (2015) dataset (1KGP) have a history of significantly reduced average generation times compared to West African populations. These differences extend to over 10,000 generations, the time period highlighted in this study. In discussing this result the authors state, “[T]he difference among populations beyond 2000 generations ago reflects population structure in humans before their dispersal out of Africa, a structure that is not fully captured by the 1000 Genomes AFR sample. This implies that the simple labels of ‘African’ and ‘non-African’ for these populations conceal differences in generation times that existed on our ancestral continent.” Indeed, a number of recent genetic studies suggest that human population structure within Africa extending hundreds of thousands of years into the past has in part shaped present-day genetic variation (HAMMER *et al.*, 2011; HSIEH *et al.*, 2016; HEY *et al.*, 2018; RAGSDALE and GRAVEL, 2019; DURVASULA and SANKARARAMAN, 2020; LORENTE-GALDOS *et al.*, 2019).

However, in extending their analysis deeper into the past, WANG *et al.* (2023) find that ancestral generation intervals do not converge until many 10s of thousands of generations ago. Assuming an average generation time of 25–30 years, this corresponds to well over one million years. This observation would require some portion of the ancestries of Eurasian and West African populations to have remained isolated for many hundreds of thousands of years, for those structured ancestral populations to have had large differences in average generation times over the course of this history, and for those groups to have contributed substantively to different contemporary human populations. While such a scenario of very long-lasting isolation among ancestral populations is not impossible, it is not supported by genetic (RAGSDALE *et al.*, 2022; ?) or archaeological (SCERRI *et al.*, 2018; ?) evidence, which rather suggest at least periodic connectivity of ancestral human populations within Africa.

Genetic studies have estimated the Eurasian–West African divergence (i.e., the time of recent shared ancestry) at only  $\approx 75$ ka (thousand years ago) (e.g., PAGANI *et al.*, 2015; ?). While population genetic studies vary considerably in estimated population split times, even those that infer deeper human divergences place the Eurasian–West African divergence at 100–150ka (e.g., SCHLEBUSCH *et al.*, 2017). If such estimated divergence times represent the majority of ancestry of the two groups (while allowing for a smaller portion to be due to long-lasting structure), then the shared portion of ancestry should be subject to the same generation intervals prior to the divergence time. Any differences in the mutation spectrum from those epochs would be driven by differences in generation times affecting the minority of ancestry that remained isolated.

As a simple test of such a scenario, we consider a demographic model of archaic admixture within Africa (e.g., DURVASULA and SANKARARAMAN, 2020), allowing for some proportion (up to 10%) of admixture from a diverged lineage into West African populations. At 10,000 generations ago, paternal and maternal

generation intervals in the ancestors of Eurasians were inferred to both be  $\approx 20$  years, while the ancestral African generation intervals were at least 28 and 23 (Figure S4 in WANG *et al.* (2023)). Using the same mutation model (JÓNSSON *et al.*, 2017), we can determine the generation intervals in the “ghost” lineage that is needed to result in a mutation spectrum matching that of the inferred generation times.

We assume that ancestry proportions of, for example, 10% and 90% from the diverged and Eurasian-shared lineages result in surviving variation from those epochs having similar proportions of contributions (in effect, ignoring differences in total mutation rate and demographic effects that may distort the magnitude of the contributed mutation spectra). With 10% admixture into West Africans from a diverged lineage, the mean paternal age of conception would need to be 92 and the mean maternal age 48. These are unreasonably long generation times for *Homo* species. With 20% admixture from this diverged lineage (which is larger than has been proposed or inferred in previous genetic studies), mean ages would still need to be 58 and 34.

Therefore, even assuming a model of long-lasting population structure with strict isolation within Africa, we find the reconstructed generation time intervals over the past 10,000 years from WANG *et al.* (2023) to be incompatible with plausible life histories of early humans. Given this, it is natural to ask what may be causing such mis-inferences. Below we show that multiple sources of uncertainty, namely noise and bias in allele age inference and inconsistencies in trio-base estimates of mutation profiles, confound inferences of generation times from time series of mutation spectra.

## Inconsistencies in inferred mutation spectra over time

Central to the inference of generation intervals from time-stratified mutation spectra is the dating of variant ages. WANG *et al.* (2023) used published allele ages from GEVA (ALBERS and MCVEAN, 2020). GEVA estimates allele ages by considering the number of mutations that have accumulated on the ancestral haplotype carrying the focal variant, as well as the effect of recombination in reducing the size of that ancestral haplotype. Singletons are excluded from analysis and are not assigned an age. Partitioning variants by their estimated ages shows that the mutation spectrum (i.e., the distribution of six mutation types) has changed over time, assuming that the observed spectrum of segregating variation is not biased with respect to the spectrum of *de novo* mutations occurring during that time (Figure 1A and Figure 1C in WANG *et al.* (2023)). **APR: Would require a selection (or genotyping or methodological error) argument.**

In fitting the generation time history to data from the past 10,000 generations, we find that the inferred generation times provide a poor fit to the data (Figure 1X). The relative proportions of the predicted mutation spectrum trend in opposite directions for some mutation classes, confirming the concern from GAO *et al.* (2022) that a single generation time history cannot simultaneously explain each of the six mutation class frequency changes. For alleles older than 10,000 generations, the mutation spectrum fluctuates by large amounts (Figure S4). While “age estimates of mutations in the very distant past have decreased accuracy” (ALBERS and MCVEAN, 2020; WANG *et al.*, 2023), the large differences in proportions suggests a bias in GEVA’s reported dates that correlates with mutation type.

Given the poor fit of the model to the GEVA data and the known uncertainty in age estimation for older variants (ALBERS and MCVEAN, 2020), we attempted to reproduce the inferred generation interval histories using allele age estimates from independent sources, **Relate** (SPEIDEL *et al.*, 2019) and **tsdate** (WOHNS *et al.*, 2022), both state-of-the-art genealogical reconstruction methods. Allele age estimates between the three methods are only moderately correlated (GEVA and **Relate**:  $r^2 \approx 0.28$ , GEVA and **tsdate**:  $r^2 \approx 0.34$ , **tsdate** and **Relate**:  $r^2 \approx 0.64$ , and see Figure S20 from WOHNS *et al.* (2022)). Despite the low to moderate correlation, we would hope that differences are unbiased with respect to mutation type. However, allele ages provided by each method result in unlike mutation spectrum histories (Figures 1A-C and S1-S3), with mutation spectrum changes often trending in opposite directions over the same time periods. In turn, these divergent spectrum histories provide qualitatively different inferred generation time histories (Figure 1X). Notably, none of them provide a reasonable fit to the data (Figure 1Y).

Published allele ages from both ALBERS and MCVEAN (2020) (GEVA) and SPEIDEL *et al.* (2019) (**Relate**)

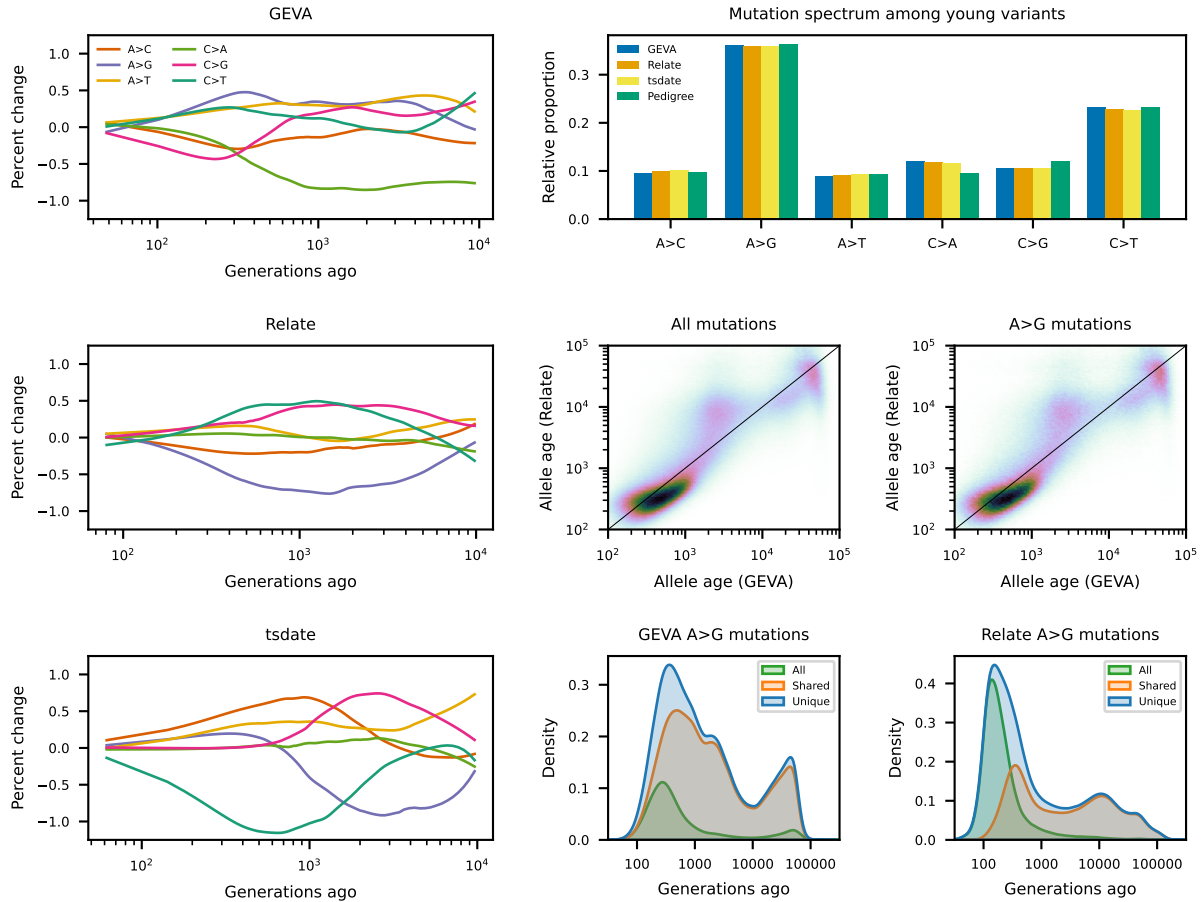


Figure 1: **caption** APR: I think I want to show a few other things APR: Need panel labels APR: This figure should just be left three panels, and model predictions and fit between model and data, move right side stuff to Fig 2

used the same input dataset, phase 3 of 1KGP. However, they differ in the total number of variants that were assigned ages (roughly 30 million and 48 million, respectively). Of those variants, 25 million variants were dated by each,

Q: Did Gao or Wang

- **GEVA** and **Relate** work with the same input data, but they keep and discard different proportions of mutations depending on class.
- **TODO: Quantify**
  - total numbers of mutations with estimated ages
  - proportions of each mutation type given an age
  - among non-shared mutations, which mutation types are over- and under-represented by each method?

Conclusions from this section:

- Mutation spectrum histories stratified by estimated allele ages are unreliable, as methods disagree even for fairly young mutations, and it’s not clear whether *any* of the methods get it right (relevant to GAO *et al.* (2022)).
- It is not obvious where the discrepancies are coming from (need to look into BRANDT *et al.* (2022))

## Mutation spectra differ between *de novo* mutations and young alleles

The large disagreements in mutation spectrum histories between multiple variant age-estimation methods are concerning for down-stream inferences that rely on them. Independent of biases in allele age estimate, there are further problems in comparing age-stratified mutation spectra to those estimated from pedigree studies (JÓNSSON *et al.*, 2017; HALLDORSSON *et al.*, 2019). As WANG *et al.* (2023) acknowledge, the spectrum of *de novo* mutations identified in Icelandic trios (JÓNSSON *et al.*, 2017) differs from the spectrum of young segregating variation (e.g., variants estimated to be less than 100 generations old, Table S1). GAO *et al.* (2022) argue that these differences are unlikely to be driven by biological processes.

For some mutation classes, the relative proportion of *de novo* mutations in the trio-based study differs from the young-variant spectrum by up to 0.02, which would imply a large over- or under-count of different mutation types. GEVA, tsdate and Relate, while their estimated mutation spectra differ at older times, very closely agree for mutations inferred to be less than 100 generations old (Table S1). In discussing this discrepancy, WANG *et al.* (2023) state, “We found that the mutation spectrum from the large pedigree study consistently differed from the variant spectrum inferred from the 1000 Genomes Project data, possibly because we removed singletons from the polymorphism dataset to reduce errors.” Rather, GEVA does not provide estimates of allele ages for singletons, so this suggested source of discrepancy cannot be checked with their published allele ages. Both tsdate and Relate report allele ages for singletons, and their inclusion does not strongly affect the mutation spectrum in the most recent time period (Table S1), though it does impact the mutation profiles in older time periods (Figures S5, S6). Of note, reported ages from GEVA and Relate both used the low-coverage phase 3 1KGP data while tsdate used the more recent independently sequenced high-coverage 1KGP data (BYRSKA-BISHOP *et al.*, 2022), so the accuracy of mutation profiles among young variants is unlikely to be driven by differences in coverage.

To account for the differences between the *de novo* spectrum from pedigree studies (JÓNSSON *et al.*, 2017) and the spectrum among young variants, WANG *et al.* (2023) subtracted this difference from each historical mutational spectrum. This choice “has the effect of assuming that parental ages in the pedigreed mutation dataset reflect generation times in the most recent historical bin” (WANG *et al.*, 2023). [APR: Is it reasonable to assume that generation intervals in present-day icelanders should be the same as the average intervals over the past 3-5 thousand years?](#) However, more concerning is that we do not know the source of the discrepancy between the Iceland and 1KGP spectra. Without knowing this, it is unclear that simply subtracting the differences between them properly account for it. Instead, mutation class-specific genotyping differences likely distorts the underlying mutation model, potentially driving deviations in predicted mutation profiles in unexpected ways [APR: this is purely speculative.](#)

What could be driving the large disagreement between the spectrum of *de novo* mutations from Icelandic pedigrees and that of young variants in the 1KGP dataset? First, there could be true differences in the mutation spectrum between the Iceland population and 1KGP populations, although this is unlikely as populations of different ancestries in 1KGP are more closely aligned than any one is to the Iceland spectrum, including the EUR populations. If it were the case, then the Icelandic pedigree data is an inappropriate calibration for the mutation spectrum in other populations. Second, there could be differences in selective pressures between mutations of different classes, although selection would need to be very strong for some classes compared to others in order to see the observed difference among young variants. [APR: we could maybe use moments to test this?](#) Third, the signal is driven by genotyping error or bioinformatics choices, though the agreement between high- and low-coverage 1KGP datasets suggests that genotyping error does not have a strong effect in the 1KGP data. Instead, filtering and bioinformatics choices in the pedigree approach are the likely culprit (BERGERON *et al.*, 2022). Until the source of these differences are understood

and properly accounted for, population-genetic inferences should avoid calibration using mutation rates and profiles from pedigree studies.

## Conclusions

1. Allele age estimates are noisy, and probably shouldn't be used for such detailed inferences. You'll end up fitting the noise and bias of each method.
2. DNM estimates from trios have their own sets of problems. Do we know where the discrepancy between trio-estimated DNM spectrum and observations from pop-gen data come from? Probably needs to be sorted out.
3. Finally, WANG *et al.* (2023) gives us an excellent example of the need for validation in population genetics studies, especially when inferences are built upon previous inferences that are known to be noisy and that need additional validations in their own right.

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