STAT 4130: Homework4

Wang Boqian

2023-07-19

```
knitr::opts_chunk$set(echo = TRUE)
# if you are using libraries, it's good practice to load them here
library(ggplot2)

## Warning: 'ggplot2'R4.3.1
library(car)

## warning: 'car'R4.3.1

## carData

## Warning: 'carData'R4.3.1
library(tibble)
library(locfit)

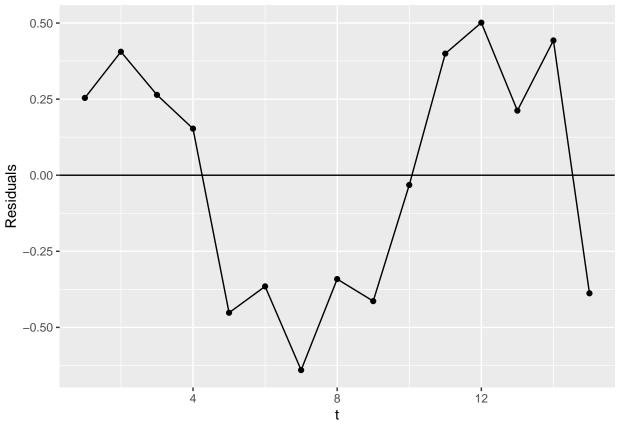
## Warning: 'locfit'R4.3.1

## locfit 1.5-9.8 2023-06-11
```

Question 1

```
# please do your coding inside a code chunk
# unless otherwise stated, feel free to do all computations in R
# commented code is always appreciated

#print("Hello")
# 1a
# Simple Linear Regression
df <- read.csv("hw4.csv")
lm1 = lm(y ~ x,data = df)
# Residual Plot
ggplot(df, aes(x=1:15, y=lm1$res)) + geom_point() + geom_line() + xlab("t") + ylab("Residuals")+ geom_h</pre>
```



```
# 1b
durbinWatsonTest(lm1, alt="positive")
## lag Autocorrelation D-W Statistic p-value
               0.5410964
                              0.8182972
## Alternative hypothesis: rho > 0
# Cochrane-Orcutt Iteration
x = df$x
y = df y
n = length(y)
res = lm1$res
 \text{rho.hat} = \text{sum}(\text{res}[1:(n-1)]*\text{res}[2:n]) / \text{sum}(\text{res*res}) 
ystar = y[2:n] - rho.hat*y[1:(n-1)]
xstar = x[2:n] - rho.hat*x[1:(n-1)]
lm2 = lm(ystar ~ xstar)
summary(lm2)
##
## Call:
## lm(formula = ystar ~ xstar)
##
## Residuals:
        Min
                   1Q Median
## -0.65214 -0.16183 0.04351 0.22211 0.40471
```

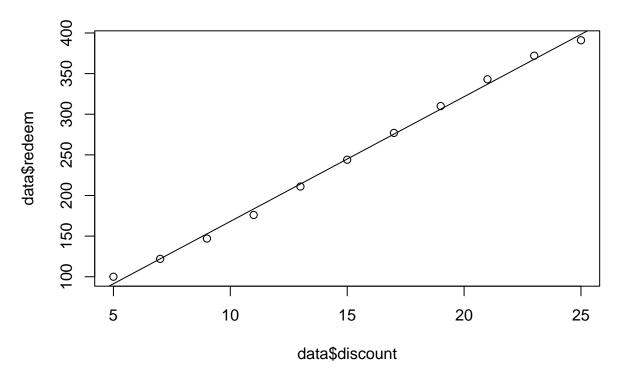
```
##
## Coefficients:
##
              Estimate Std. Error t value Pr(>|t|)
## (Intercept) 12.08544
                          0.55421 21.806 5.07e-11 ***
## xstar
              -0.11050
                          0.01403 -7.874 4.42e-06 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.3133 on 12 degrees of freedom
## Multiple R-squared: 0.8379, Adjusted R-squared: 0.8243
## F-statistic: 62.01 on 1 and 12 DF, p-value: 4.419e-06
# 1d
# DW Test
durbinWatsonTest(lm2, alt="positive")
##
   lag Autocorrelation D-W Statistic p-value
##
             0.2983033
                             0.90324
                                       0.004
##
   Alternative hypothesis: rho > 0
```

- 1a. The time-plot shows quite smooth, which indicates a sign of positive autocorrelation.
- **1b.** Since the p-value = 0.003 < 0.05, we reject H0 and accept H1: There is a positive autocorrelation.
- **1c.** According to the summary table, the standard error for intercept is 0.55421, and the standard error for slope is 0.01403.
- 1d, It has not been successful. Although the p-value is increased from 0.003 to 0.011, it's still far less than 0.05, where we still reject H0 and accept H1: There is a positive autocorrelation. We still need more iterations to reduce the autocorrelation.

Question 2

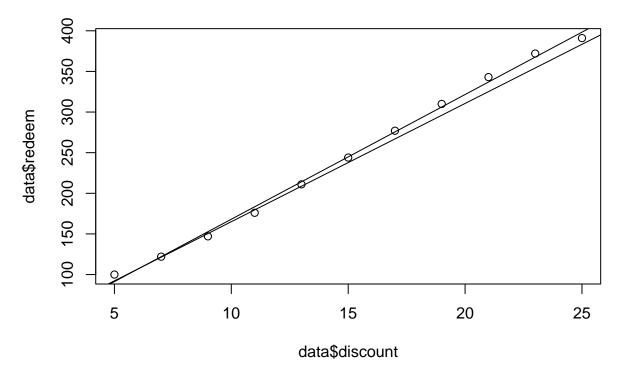
```
# Your code
# 2a
data <- tibble(</pre>
discount = c(5,7,9,11,13,15,17,19,21,23,25),
redeem = c(100, 122, 147, 176, 211, 244, 277, 310, 343, 372, 391)
lm3 <- glm(redeem ~ discount, data = data, family = gaussian())</pre>
summary(1m3)
##
## Call:
## glm(formula = redeem ~ discount, family = gaussian(), data = data)
##
## Coefficients:
##
             Estimate Std. Error t value Pr(>|t|)
## (Intercept) 14.5682
                          4.4786
                                  3.253 0.00995 **
## discount
               15.3500
                          0.2751 55.794 9.6e-13 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
```

```
## (Dispersion parameter for gaussian family taken to be 33.30404)
##
## Null deviance: 103973.64 on 10 degrees of freedom
## Residual deviance: 299.74 on 9 degrees of freedom
## AIC: 73.572
##
## Number of Fisher Scoring iterations: 2
## 2c
plot(data$discount,data$redeem)
abline(lm3)
```



```
# 2d
lm4 <- glm(redeem ~ discount + I(discount^2), data = data, family = gaussian())</pre>
summary(lm4)
##
## Call:
## glm(formula = redeem ~ discount + I(discount^2), family = gaussian(),
##
       data = data)
##
## Coefficients:
                 Estimate Std. Error t value Pr(>|t|)
##
## (Intercept)
                 19.47348
                            10.59431
                                        1.838
                                                 0.103
## discount
                 14.55455
                              1.56832
                                        9.280 1.48e-05 ***
## I(discount^2) 0.02652
                              0.05139
                                        0.516
                                                 0.620
```

```
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
  (Dispersion parameter for gaussian family taken to be 36.26061)
##
##
##
      Null deviance: 103973.64 on 10 degrees of freedom
                        290.08
## Residual deviance:
                               on 8 degrees of freedom
## AIC: 75.212
##
## Number of Fisher Scoring iterations: 2
plot(data$discount,data$redeem)
abline(lm3)
abline(lm4)
```



```
# 2f
confint(lm4)

## Waiting for profiling to be done...

## 2.5 % 97.5 %

## (Intercept) -1.29097708 40.2379468

## discount 11.48068981 17.6284011

## I(discount^2) -0.07421558 0.1272459
```

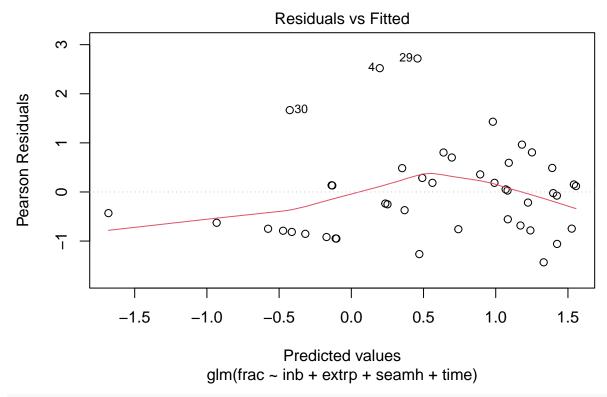
2b According to the summary table of lm3, we can calculate the R-squared, which equals to 0.997, which is close to 1. So we can't claim that the logistic regression model from part a is not adequate.

2d According to the summary table of lm4, we found that the p-value of the quadratic term = 0.620×0.05 . So this quadratic term is not required in the model as there is no strong evidence of non-linearity.

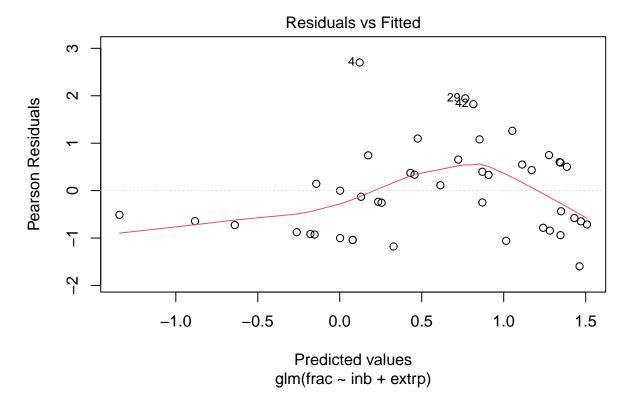
2e Yes, as the plot suggests, lm3 obviously underestimates most of the data, while data points in lm4 distributes more evenly.

Question 3

```
# Your Code
# 3a
data(mine)
data2 = mine
lm5 <- glm(frac ~ inb + extrp + seamh + time, data = data2, family = poisson(link = "log"))</pre>
# 3c
confint(lm5)
## Waiting for profiling to be done...
##
                      2.5 %
                                   97.5 %
## (Intercept) -5.69525976 -1.6604283925
               -0.00316198 0.0001348221
                0.03923837 0.0875324865
## extrp
## seamh
               -0.01287548 0.0070791852
## time
               -0.06418168 -0.0002029767
# 3d
lm6 <- glm(frac ~ inb + extrp, data = data2, family = poisson(link = "log"))</pre>
anova(lm5,lm6, test = "LRT")
## Analysis of Deviance Table
##
## Model 1: frac ~ inb + extrp + seamh + time
## Model 2: frac ~ inb + extrp
     Resid. Df Resid. Dev Df Deviance Pr(>Chi)
            39
## 1
                   37.856
## 2
            41
                    42.094 -2 -4.2377
                                         0.1202
# 3e
plot(lm5, which = 1)
```



plot(lm6, which = 1)



3b According to the summary table of lm5, we can calculate the R-squared, which equals to 0.4951. This indicates that the model can interpret less than 50% of the data points, which is far from our satisfaction.

3d lm5 is reduced to lm6 by removing the variable "seamh" and "time" because their p-values are larger than 0.05. I found that the p-value of the Analysis of Deviance is 0.1202 > 0.05, which indicates that the reduced model(lm6) is preferred.

3e Not really. Simply deleting features does not make the figure satisfactory from a residual analysis viewpoint. Variable transformations and interactions should be taken into account to modify the model.