

# 统计方法与机器学习 理论作业3 参考答案

## 1

对任意  $j \in \{1, 2, \dots, p\}$ , 设

$$\mathbf{T}_j = \begin{pmatrix} & & 1 & & & \\ & & & \ddots & & \\ & & & & 1 & \\ 1 & & & & & \\ & & & & & 1 \\ & & & & & \ddots \\ & & & & & & 1 \end{pmatrix} \quad (1)$$

(即将第  $j$  列与前  $j - 1$  列交换), 则

$$\mathbf{T}_j^{-1} = \begin{pmatrix} & & & & 1 & \\ 1 & & & & & \\ & \ddots & & & & \\ & & 1 & & & \\ & & & 1 & & \\ & & & & 1 & \\ & & & & & \ddots \\ & & & & & & 1 \end{pmatrix} = \mathbf{T}_j^T \quad (2)$$

若令  $\mathbf{X}_t = (\mathbf{x}_j, \mathbf{X}_0)$ , 其中  $\mathbf{X}_0 = (\mathbf{x}_1, \dots, \mathbf{x}_{j-1}, \mathbf{x}_{j+1}, \dots, \mathbf{x}_p)$

则根据题意, 有  $\mathbf{X}_t = \mathbf{X}_s \mathbf{T}_j$

因此  $\mathbf{X}_s = \mathbf{X}_t \mathbf{T}_j^T$

于是

$$\begin{aligned} \mathbf{C} &= (\mathbf{X}_s^T \mathbf{X}_s)^{-1} \\ &= \left( (\mathbf{X}_t \mathbf{T}_j^T)^T (\mathbf{X}_t \mathbf{T}_j^T) \right)^{-1} \\ &= (\mathbf{T}_j \mathbf{X}_t^T \mathbf{X}_t \mathbf{T}_j^T)^{-1} \\ &= \mathbf{T}_j (\mathbf{X}_t^T \mathbf{X}_t)^{-1} \mathbf{T}_j^T \end{aligned} \quad (3)$$

而由于

$$\begin{aligned} \mathbf{X}_t^T \mathbf{X}_t &= \begin{pmatrix} \mathbf{x}_j^T \\ \mathbf{X}_0^T \end{pmatrix} (\mathbf{x}_j \quad \mathbf{X}_0) \\ &= \begin{pmatrix} \mathbf{x}_j^T \mathbf{x}_j & \mathbf{x}_j^T \mathbf{X}_0 \\ \mathbf{X}_0^T \mathbf{x}_j & \mathbf{X}_0^T \mathbf{X}_0 \end{pmatrix} \end{aligned} \quad (4)$$

故由分块矩阵求逆公式可知

$$(\mathbf{X}_t^T \mathbf{X}_t)^{-1} = \begin{pmatrix} \mathbf{X}_{11}^* & \mathbf{X}_{12}^* \\ \mathbf{X}_{21}^* & \mathbf{X}_{22}^* \end{pmatrix} \quad (5)$$

其中

$$\begin{aligned} \mathbf{X}_{11}^* &= \left( \mathbf{x}_j^T \mathbf{x}_j - \mathbf{x}_j^T \mathbf{X}_0 (\mathbf{X}_0^T \mathbf{X}_0)^{-1} \mathbf{X}_0^T \mathbf{x}_j \right)^{-1} \\ \mathbf{X}_{12}^* &= - \left( \mathbf{x}_j^T \mathbf{x}_j - \mathbf{x}_j^T \mathbf{X}_0 (\mathbf{X}_0^T \mathbf{X}_0)^{-1} \mathbf{X}_0^T \mathbf{x}_j \right)^{-1} \mathbf{x}_j^T \mathbf{X}_0 (\mathbf{X}_0^T \mathbf{X}_0)^{-1} \\ \mathbf{X}_{21}^* &= - (\mathbf{X}_0^T \mathbf{X}_0)^{-1} \mathbf{X}_0^T \mathbf{x}_j \left( \mathbf{x}_j^T \mathbf{x}_j - \mathbf{x}_j^T \mathbf{X}_0 (\mathbf{X}_0^T \mathbf{X}_0)^{-1} \mathbf{X}_0^T \mathbf{x}_j \right)^{-1} \\ \mathbf{X}_{22}^* &= (\mathbf{X}_0^T \mathbf{X}_0)^{-1} + (\mathbf{X}_0^T \mathbf{X}_0)^{-1} \mathbf{X}_0^T \mathbf{x}_j \left( \mathbf{x}_j^T \mathbf{x}_j - \mathbf{x}_j^T \mathbf{X}_0 (\mathbf{X}_0^T \mathbf{X}_0)^{-1} \mathbf{X}_0^T \mathbf{x}_j \right)^{-1} \mathbf{x}_j^T \mathbf{X}_0 (\mathbf{X}_0^T \mathbf{X}_0)^{-1} \end{aligned} \quad (6)$$

而由于  $\mathbf{X}_s$  为标准化后的特征，因此  $\mathbf{x}_j^T \mathbf{x}_j = 1$

于是由  $\mathbf{T}_j$  的定义可知

$$\begin{aligned} c_{jj} &= \mathbf{X}_{11}^* \\ &= \left( \mathbf{x}_j^T \mathbf{x}_j - \mathbf{x}_j^T \mathbf{X}_0 (\mathbf{X}_0^T \mathbf{X}_0)^{-1} \mathbf{X}_0^T \mathbf{x}_j \right)^{-1} \\ &= \left( 1 - \mathbf{x}_j^T \mathbf{X}_0 (\mathbf{X}_0^T \mathbf{X}_0)^{-1} \mathbf{X}_0^T \mathbf{x}_j \right)^{-1} \end{aligned} \quad (7)$$

另一方面，由于

$$\bar{x}_j = \frac{1}{n} \sum_{k=1}^n x_{jk} = 0 \quad (8)$$

因此

$$\begin{aligned} SS_R^j &= \sum_{k=1}^n (\hat{x}_{jk} - \bar{x}_j)^2 \\ &= \sum_{k=1}^n \hat{x}_{jk}^2 \\ &= \hat{\mathbf{x}}_j^T \hat{\mathbf{x}}_j \\ &= (\mathbf{X}_0 \hat{\boldsymbol{\beta}})^T \mathbf{X}_0 \hat{\boldsymbol{\beta}} \\ &= \mathbf{x}_j^T \mathbf{X}_0 (\mathbf{X}_0^T \mathbf{X}_0)^{-1} \mathbf{X}_0^T \mathbf{X}_0 (\mathbf{X}_0^T \mathbf{X}_0)^{-1} \mathbf{X}_0^T \mathbf{x}_j \\ &= \mathbf{x}_j^T \mathbf{X}_0 (\mathbf{X}_0^T \mathbf{X}_0)^{-1} \mathbf{X}_0^T \mathbf{x}_j \end{aligned} \quad (9)$$

而由标准化条件可知

$$SS_T^j = \sum_{k=1}^n (x_{jk} - \bar{x}_j)^2 = \sum_{k=1}^n x_{jk}^2 = 1 \quad (10)$$

故该回归模型的复决定系数

$$R^2 = \frac{SS_R^j}{SS_T^j} = \mathbf{x}_j^T \mathbf{X}_0 (\mathbf{X}_0^T \mathbf{X}_0)^{-1} \mathbf{X}_0^T \mathbf{x}_j \quad (11)$$

由此可得

$$c_{jj} = \left(1 - \mathbf{x}_j^T \mathbf{X}_0 (\mathbf{X}_0^T \mathbf{X}_0)^{-1} \mathbf{X}_0^T \mathbf{x}_j\right)^{-1} = \frac{1}{1 - R^2} \quad (12)$$

## 2

由线性回归最小二乘估计的性质可知

$$\begin{aligned} E[\hat{\boldsymbol{\beta}}] &= \boldsymbol{\beta} \\ \mathbf{Cov}(\hat{\boldsymbol{\beta}}) &= \sigma^2 (\mathbf{X}^T \mathbf{X})^{-1} \end{aligned} \quad (13)$$

因此

$$\begin{aligned} MSE(\hat{\boldsymbol{\beta}}) &= E\left[(\hat{\boldsymbol{\beta}} - \boldsymbol{\beta})^T (\hat{\boldsymbol{\beta}} - \boldsymbol{\beta})\right] \\ &= E\left[(\hat{\boldsymbol{\beta}} - E[\hat{\boldsymbol{\beta}}] + E[\hat{\boldsymbol{\beta}}] - \boldsymbol{\beta})^T (\hat{\boldsymbol{\beta}} - E[\hat{\boldsymbol{\beta}}] + E[\hat{\boldsymbol{\beta}}] - \boldsymbol{\beta})\right] \\ &= E\left[(\hat{\boldsymbol{\beta}} - E[\hat{\boldsymbol{\beta}}])^T (\hat{\boldsymbol{\beta}} - E[\hat{\boldsymbol{\beta}}])\right] \\ &= \text{tr}[\mathbf{Cov}(\hat{\boldsymbol{\beta}})] \\ &= \sigma^2 \text{tr}[(\mathbf{X}^T \mathbf{X})^{-1}] \end{aligned} \quad (14)$$

进一步的, 若  $\mathbf{X} \in \mathbb{M}_{n \times (p+1)}(\mathbb{R})$ , 设  $\mathbf{X}^T \mathbf{X}$  的特征值为  $\lambda_1, \lambda_2, \dots, \lambda_{p+1}$

则由  $\mathbf{X}^T \mathbf{X}$  为对称矩阵可知  $\lambda_1, \lambda_2, \dots, \lambda_{p+1} \in \mathbb{R}$

因此  $(\mathbf{X}^T \mathbf{X})^{-1}$  的特征值为  $\frac{1}{\lambda_1}, \frac{1}{\lambda_2}, \dots, \frac{1}{\lambda_{p+1}}$

此时

$$\begin{aligned} MSE(\hat{\boldsymbol{\beta}}) &= \sigma^2 \text{tr}[(\mathbf{X}^T \mathbf{X})^{-1}] \\ &= \sigma^2 \sum_{i=1}^n \frac{1}{\lambda_i} \end{aligned} \quad (15)$$