

Durrett: Probability: Theory and Examples

概率极限理论基础. 林正炎

Petrov (1995). Limit theorems of Probability Theory - Sequences of Independent R.V.s

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期末: WLLN, SLLN, CLT, Donsker, Martingale

WUN PF:  $P\left(\left|\frac{S_n - nEX}{\sqrt{n}}\right| > \varepsilon\right) \leq \frac{1}{\varepsilon^2} E\left(\frac{(S_n - nEX)^2}{n}\right) = \frac{\text{Var}(X)}{\varepsilon^2} \rightarrow 0 \quad (n \rightarrow \infty)$ , 若  $EX^2 < \infty$

補尾:  $S'_n = \sum X_i I(|X_i| \leq N)$ ,  $S''_n = \sum X_i I(|X_i| > N)$

$$\begin{aligned} P\left(\left|\frac{S_n - nEX}{\sqrt{n}}\right| > \varepsilon\right) &\leq P\left(\left|\frac{(S'_n - ES'_n)}{\sqrt{n}}\right| > \varepsilon\right) + P\left(\left|\frac{(S''_n - ES''_n)}{\sqrt{n}}\right| > \varepsilon\right) \\ &\leq \frac{\text{Var}(S'_n)}{N^2\varepsilon^2} + \frac{1}{\varepsilon} E[X'' - EX''] \\ &\leq \frac{N^2}{N^2\varepsilon^2} + \frac{2E[XI(|X| > N)]}{\varepsilon} \\ \lim_{n \rightarrow \infty} P\left(\left|\frac{S_n - nEX}{\sqrt{n}}\right| > 2\varepsilon\right) &\leq \frac{2E[XI(|X| > N)]}{\varepsilon} \rightarrow 0 \quad (N \rightarrow \infty) \Rightarrow \lim_{n \rightarrow \infty} P\left(\left|\frac{S_n - nEX}{\sqrt{n}}\right| > 2\varepsilon\right) = 0 \end{aligned}$$

PP2:  $Y_n = X_n I(|X_n| \leq n)$ ,  $T_n = \sum_{k=1}^n Y_k$

$$\begin{aligned} \sum P(X_n \neq Y_n) &= \sum P(|X_n| > n) = \sum E[I(|X| > n)] = E \sum I(|X| > n) \leq E|X| < \infty \\ \text{由 B-C 定理, } P(X_n \neq Y_n, i.o.) &= 0 \Rightarrow \frac{\sum X_k - \sum Y_k}{n} \rightarrow 0 \text{ a.s.} \\ \frac{S_n}{n} \xrightarrow{P} EX, \text{ 只须 } \frac{T_n}{n} \xrightarrow{P} EX. \\ \because P\left(\left|\frac{T_n - ET_n}{\sqrt{n}}\right| > \varepsilon\right) &\leq \frac{1}{\varepsilon^2} \sum_{k=1}^n \text{Var}(Y_k) \leq \frac{1}{\varepsilon^2} \sum_{k=1}^n EX^2 I(|X| \leq k) \stackrel{(*)}{\rightarrow} 0 \quad (n \rightarrow \infty) \\ \therefore \frac{T_n - ET_n}{\sqrt{n}} &\xrightarrow{P} 0 \quad \text{又: } \frac{ET_n}{n} = \frac{1}{n} \sum EX I(|X| \leq k) \rightarrow EX \quad (\text{数列收敛}) \\ \therefore \text{由 Slutsky, } \frac{T_n}{n} &\xrightarrow{P} EX \end{aligned}$$

(\*)欲证  $\frac{\text{Var}(Y_n)}{n^2} \rightarrow 0$ :  $\because \lim_{n \rightarrow \infty} \frac{EX^2 I(|X| \leq n)}{n^2 - (n-1)^2} = \lim_{n \rightarrow \infty} \frac{1}{2} E \frac{X^2}{n} I(|X| \leq n), \text{ (对称性)} \\ \therefore LHS = \frac{E \sum X^2 I(|X| \leq k)}{n^2} \leq \lim_{n \rightarrow \infty} \frac{E \sum X^2 I(|X| \leq n)}{n} + E \frac{X^2 I(N \leq |X| \leq n)}{n} \\ = \frac{EX^2 (n - [x]) I(|X| \leq n)}{n^2}$

$$\leq E \frac{X^2 (n+1 - |X|) I(|X| \leq n)}{n^2} \leq 2 E \frac{X^2}{n} I(|X| \leq n) \rightarrow 0 \quad (n \rightarrow \infty)$$

(\*\*)  $LHS = \frac{1}{n^2} \sum_{k=1}^n EX^2 I((k-1 < |X| \leq k))$

$$\begin{aligned} &\leq \frac{1}{n^2} \sum_{k=1}^n E|X| I((k-1 < |X| \leq k)) = \frac{1}{n^2} \sum_{k=1}^n E|X| I((k-1 < |X| \leq k)) + \frac{1}{n^2} E|X| I(|X| \leq k) \\ &\stackrel{?}{=} \frac{1}{n^2} \sum_{k=1}^n E|X| I((k-1 < |X| \leq k)) \leq \frac{1}{n^2} \sum_{k=1}^n E|X| I((k-1 < |X| \leq k)) + \frac{1}{n^2} E|X| I(|X| \leq k) \\ &\leq \frac{1}{n^2} \sum_{k=1}^n E|X|^2 I((k-1 < |X| \leq k)) + \frac{1}{n^2} E|X| I(|X| \leq k) \\ &\leq \frac{1}{n^2} E|X|^2 I(|X| \leq n) \downarrow 0 \\ &= \frac{E|X|^2}{n} I(|X| \leq n) \rightarrow 0 \end{aligned}$$

## · 大数定律

$X, X_1, X_2, \dots$  i.i.d. r.v.s,  $S_n \triangleq \sum_{i=1}^n X_i$

Thm WLLN:  $E|X| < \infty$  时,  $\frac{S_n}{n} \xrightarrow{P} EX$  (Rmk: 非必要, 即若  $\frac{S_n}{n} \xrightarrow{a.s.} a \neq E|X|$  时)

SLLN:  $E|X| < \infty$  时,  $\frac{S_n}{n} \xrightarrow{a.s.} EX$

Thm “ $\exists$  常数序列  $\{b_n\}$ , s.t.  $\frac{S_n - b_n}{\sqrt{n}} \xrightarrow{D} 0$ ”  $\Leftrightarrow n P(|X| > n) \rightarrow 0$ , 此时  $b_n = EX I(|X| > n)$

C.e.g.  $X_1, X_2, \dots$   $P(X_i \leq x) = \int_{-\infty}^x \frac{dt}{\pi(1+t^2)}$ , 有  $E|X|^r < \infty, r < 1$   
 $P(|X| > x) \sim \frac{1}{\pi x}$  不满足上述两个条件

$\frac{S_n}{n} \xrightarrow{d} X$

PF:  $X_{nk} \triangleq X_k I(|X_k| \leq n)$ ,  $T_n \triangleq \sum_{k=1}^n X_{nk}$ ,  $Y_n \triangleq EX_{nn}$

$$\begin{aligned} (\Leftarrow) P\left(\left|\frac{S_n}{n} - EX I(|X| \leq n)\right| > \varepsilon\right) &= P(S_n = T_n, \left|\frac{S_n}{n} - EX I(|X| \leq n)\right| > \varepsilon) \\ &\quad + P(S_n \neq T_n, \left|\frac{S_n}{n} - EX I(|X| \leq n)\right| > \varepsilon) \\ &\leq P\left(\frac{T_n - ET_n}{\sqrt{n}} > \varepsilon\right) + P(S_n \neq T_n) \quad (+\text{o}(n)) \end{aligned}$$

$$\therefore P(S_n \neq T_n) \leq P(\exists k, X_{nk} \neq X_k) \leq \sum_{k=1}^n P(X_{nk} \neq X_k) = n P(|X| > n) \rightarrow 0$$

$$P\left(\frac{T_n - ET_n}{\sqrt{n}} > \varepsilon\right) \leq \frac{1}{n^2 \varepsilon^2} \sum_{k=1}^n \text{Var}(Y_{nk})$$

$$\leq \frac{1}{n^2 \varepsilon^2} E|X| \xrightarrow{D} 0. \quad (\text{(*)}) = \int_0^\infty P\left(\frac{X^2}{n} I(|X| \leq n) > t\right) dt^2$$

$$\therefore \frac{S_n}{n} - EX I(|X| \leq n) \xrightarrow{P} 0$$

$$= \frac{1}{n} \int_0^n 2t P(|X| > t) dt$$

$$= \frac{1}{n} (\int_0^n + \int_{\frac{n}{2}}^n) 2t P(|X| > t) dt \rightarrow 0.$$

(\*) 对称化, 见群文件

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Eg.  $\{X_n\}$  独立且  $b_n \rightarrow 0$ . 若 (1)  $\frac{1}{b_n^2} P(|X_i| > b_n) \rightarrow 0$  (2)  $\frac{1}{b_n^2} \sum_{i=1}^n E X_i^2 I(|X_i| \leq b_n) \rightarrow 0$

$$\text{则 } a_n \frac{S_n}{b_n} I(|X_i| \leq b_n), \frac{1}{b_n}(S_n - a_n) \xrightarrow{P} 0.$$

$$\begin{aligned} \text{E.g.: } X_{1,n} &\stackrel{d}{=} X_1 I(|X_1| \leq b_n), T_n = \sum_{i=1}^n X_{1,i} \\ \text{则 } P\left(\frac{|S_n - a_n|}{b_n} > \varepsilon\right) &\leq P\left(\frac{|T_n - a_n|}{b_n} > \varepsilon\right) + P(S_n \neq T_n) \\ &\leq \frac{1}{\varepsilon^2 b_n^2} \sum_{i=1}^n E X_i^2 I(|X_i| \leq b_n) + \frac{1}{b_n} P(|X_i| > b_n) \rightarrow 0 \end{aligned}$$

$$\begin{aligned} \text{E.g. } X_1, X_2, \dots &\stackrel{i.i.d.}{\sim} U(1, 2, \dots, n), T_k \triangleq \inf\{m : |X_1, \dots, X_m| = k\}, T_n \triangleq T_n \\ \text{解: } P(X_{nk} = m) &= P(X_{nk} > m-1) - P(X_{nk} > m) = \left(\frac{k-1}{n}\right)^{m-1} \left(1 - \frac{k-1}{n}\right), k = 1, \dots, n. \\ \therefore T_n &= \sum_{k=1}^n X_{nk} \\ \therefore E T_n &= \sum \left(1 - \frac{k-1}{n}\right)^{-1} = n \sum_{m=1}^n \frac{1}{m} \sim n \log n \quad \begin{matrix} \text{服从几何分布且} \\ X_{nk}, k = 1, \dots, n \text{独立} \end{matrix} \\ \therefore \text{Var } T_n &= \sum \left(1 - \frac{k-1}{n}\right)^{-2} = n^2 \sum_{m=1}^n \frac{1}{m^2} \sim Cn^2 \\ \text{则 } \frac{T_n - E T_n}{E T_n} &\xrightarrow{P} 0 \quad \frac{T_n}{E T_n} \xrightarrow{P} 1, \quad \frac{\text{Var } T_n}{E T_n} \xrightarrow{P} 1 \\ \therefore P\left(\frac{T_n - E T_n}{E T_n} > \varepsilon\right) &\leq \frac{\text{Var } T_n}{\varepsilon^2 (E T_n)^2} \sim \frac{Cn^2}{\varepsilon^2 n^4} \xrightarrow{P} 0 \end{aligned}$$

$$\text{E.g. 设 } f \in C[0,1], f_n(x) = \sum_{m=0}^n \binom{n}{m} x^m (1-x)^{n-m} f\left(\frac{m}{n}\right), \text{ 则 } \lim_{n \rightarrow \infty} \sup_{x \in [0,1]} |f_n(x) - f(x)| = 0.$$

$$\text{PF: } X_1(x) \sim i.i.d. P(X_1(x)=1) = x, P(X_1(x)=0) = 1-x$$

$$\text{则 } f_n(x) = E f\left(\frac{S_n}{n}\right)$$

$$\text{Fix } x, \frac{S_n}{n} \xrightarrow{P} x \quad (\text{WLLN})$$

$$\Rightarrow f\left(\frac{S_n}{n}\right) \xrightarrow{P} f(x)$$

$$\Rightarrow E f\left(\frac{S_n}{n}\right) \xrightarrow{P} f(x) \quad \text{即 } f_n(x) \xrightarrow{P} f(x)$$

$$\text{由 } f \text{ 连续, 在 } [0,1] \text{ 上: } \forall \varepsilon > 0, \exists \delta > 0, \text{ s.t. } |f(x) - f(y)| < \varepsilon, \forall |x-y| < \delta$$

$$\begin{aligned} \sup_{x \in [0,1]} |f_n(x) - f(x)| &= \sup_{x \in [0,1]} |E f\left(\frac{S_n}{n}\right) - f(x)| \\ &\leq \sup_{x \in [0,1]} E |f\left(\frac{S_n}{n}\right) - f(x)| I\left(\left|\frac{S_n}{n} - x\right| < \delta\right) + \sup_{x \in [0,1]} E |f\left(\frac{S_n}{n}\right) - f(x)| I\left(\left|\frac{S_n}{n} - x\right| \geq \delta\right) \\ &\leq \varepsilon + \max f \cdot \sup_{x \in [0,1]} P\left(\left|\frac{S_n}{n} - x\right| > \delta\right) \\ &\leq \varepsilon + \max f \cdot \sup_{x \in [0,1]} \frac{\text{Var}(\frac{S_n}{n})}{\delta^2} \xrightarrow{n \rightarrow \infty} \varepsilon \end{aligned}$$

由  $\varepsilon$  任意性,  $\sup_{x \in [0,1]} |f_n(x) - f(x)| \rightarrow 0$

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Thm 三角不等式:  $X_1, \dots, X_{K_n}$  (独立)

$$X_{2,1}, \dots, X_{2,K_2}$$

$$\text{若 } S_n = \sum_{k=1}^{K_n} X_{nk}, b_n > 0 \rightarrow 0. \text{ 若 (1) } \sum_{k=1}^{K_n} P(|X_{nk}| > b_n) \rightarrow 0 \quad (2) \frac{1}{b_n^2} \sum_{k=1}^{K_n} E X_{nk}^2 I(|X_{nk}| \leq b_n) \rightarrow 0$$

$$\text{则 } \left| \frac{S_n - \sum_{k=1}^{K_n} E X_{nk} I(|X_{nk}| \leq b_n)}{b_n} \right| \xrightarrow{P} 0.$$

WLLN 基于 Chebyshev; SLIN 基于 B-C Lemma

$$\bullet X_n \xrightarrow{a.s.} X \Leftrightarrow \forall \varepsilon > 0, P(|X_n - X| > \varepsilon, i.o.) = 0$$

$$\bullet B-C: \sum P(A_n) < \infty \Rightarrow P(A_n, i.o.) = 0$$

$$\int \sum P(A_n) = 0, A_n \text{ 独立} \Rightarrow P(A_n, i.o.) = 1.$$

$$\text{E.g. (B-C 在 random walk 中应用) } S_n = X_1 + \dots + X_n, X_i \stackrel{i.i.d.}{\sim} U(a_i, \dots, d_i), a_i = \pm 1 \}$$

$$\text{则 } P(S_n = 0, i.o.) = \begin{cases} 1, & d_1 = 1, \\ 0, & d_1 = 3 \end{cases}$$

$$\text{若 } n=1 \text{ 时, } P(S_{2n}=0) = \left(\frac{1}{2}\right)^n \sim \frac{(1/2)^n}{\sqrt{n}} \underset{n \rightarrow \infty}{\sim} 0.$$

$$\therefore n \geq 3 \text{ 时, } P(S_{2n}=0) \sim C \frac{1}{n^{\frac{d_1}{2}}} \quad \because P(S_{2n}=0) < \infty \Rightarrow P(S_n=0, i.o.) = 0.$$

Thm (SLIN)  $X_1, X_2, \dots$  i.i.d.,  $E|X| < \infty$ , 则  $\frac{S_n}{n} \xrightarrow{a.s.} EX$

$$\bullet \text{若 } EX = 0, EX^2 < \infty, \text{ 证明如下:}$$

$$\sum_{n=1}^{\infty} P\left(\left|\frac{S_n}{n} - EX\right| > \varepsilon\right) \leq \frac{\text{Var} X}{n^2 \varepsilon^2} < \infty. \quad \therefore \frac{S_n}{n} \xrightarrow{a.s.} EX, \text{ 由 B-C. (1)}$$

$$\text{下证 } \frac{\sup_{1 \leq k \leq n} |S_k - S_{n+1}|}{(n-1)^2} \xrightarrow{a.s.} 0. \quad (\text{若 } |X| \leq C \text{ a.s. 易证})$$

$$\begin{aligned} \text{记 } D_n &= \sup_{1 \leq k \leq n} |S_k - S_{n+1}|, \text{ 则 } \sum_{n=1}^{\infty} P\left(\frac{D_n}{n^2} > \varepsilon\right) \leq \sum_{n=1}^{\infty} \frac{ED_n^2}{n^4 \varepsilon^2} \\ &\leq \sum_{n=1}^{\infty} \frac{1}{n^4 \varepsilon^2} \cdot n \cdot E(S_{(n)_1} - S_n)^2 \\ &\leq \sum_{n=1}^{\infty} \frac{1}{n^2} < \infty \end{aligned}$$

$$\bullet \text{若 } EX = 0, EX^4 < \infty, \text{ 则 } ES_n^4 \leq n EX^4 + n^3 (EX)^2 \leq n^2$$

$$\therefore \sum P\left(\left|\frac{S_n}{n}\right| > \varepsilon\right) \leq \frac{n^2}{n^4 \varepsilon^4} < \infty \quad \xrightarrow{a.s.} \frac{S_n}{n} \xrightarrow{a.s.} EX.$$

$$\therefore \sum P\left(\left|\frac{S_n}{n}\right| > \varepsilon\right) < \infty \Rightarrow \forall \varepsilon > 0, \exists N > 0, \forall n > N, P\left(\left|\frac{S_n}{n}\right| > \varepsilon\right) < \infty$$

$$\therefore \frac{S_n}{n} \xrightarrow{a.s.} 0, \text{ 用其它条件.}$$

$$\text{Rmk: } \frac{S_n}{n} \xrightarrow{a.s.} a \Leftrightarrow E|X| < \infty \text{ 的 (2) 方向: } \frac{X_n}{n} - \frac{S_{n-1}}{n} \xrightarrow{a.s.} 0 \Leftrightarrow \sum P\left(\left|\frac{X_n}{n}\right| > \varepsilon\right) < \infty, \forall \varepsilon > 0.$$

$$\Leftrightarrow \sum P(|X| > 1) < \infty$$

$$\Leftrightarrow \sum P(|X| > n) < \infty \Leftrightarrow E|X| < \infty$$

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Eg. 设  $A_1, A_2$  两两独立,  $\sum P(A_n) = \infty$ , 则当  $n \rightarrow \infty$ ,  $\frac{\sum_{k=1}^n P(A_k)}{\sum P(A_k)} \rightarrow 1$ , a.s.

证  $\forall n \geq 1$ ,  $S_n = \sum_{k=1}^n I_{A_k}$ , 由两两独立,  $P(|S_n - ES_n| > \varepsilon) \leq \varepsilon^2 (ES_n)^2$

$\frac{P(|S_n - ES_n| > \varepsilon)}{ES_n} = \frac{P(|S_n - ES_n| > \varepsilon)}{ES_n^2} \cdot ES_n \leq \frac{1}{\varepsilon^2}$

$\lim_{n \rightarrow \infty} \frac{P(|S_n - ES_n| > \varepsilon)}{ES_n} = \lim_{n \rightarrow \infty} \frac{\sum_{k=1}^n P(I_{A_k} - \frac{ES_{k-1}}{ES_n})}{ES_n} < \infty$

由B-C,  $\frac{S_n}{ES_n} \xrightarrow{a.s.} 1$

$\therefore S_n \leq S_m \leq S_{kn}, \forall m \in (k_n, k_n)$

$\frac{S_n}{ES_n} < \frac{S_m}{ES_m} = \frac{S_m}{ES_{k_n}} \cdot \frac{ES_{k_n}}{ES_m} \xrightarrow{a.s.} \frac{ES_{k_n}}{ES_m} \leq \frac{n^2 + 1}{(m+1)^2} \rightarrow 1$

$\lim_{n \rightarrow \infty} \frac{S_n}{ES_n} \xrightarrow{a.s.} 1$  同理  $\lim_{n \rightarrow \infty} \frac{S_n}{ES_m} \geq 1$

PF of SLN: 设  $X > 0$ ,  $Y_n \triangleq X_n I(X_n \leq n)$ ,  $T_n = \sum_{k=1}^n Y_k$

$\therefore \sum P(Y_n > n) = \sum P(|X| > n) < \infty$

$\therefore P(X_n > Y_n, i.o.) = 0 \Rightarrow \frac{S_n - T_n}{n} \xrightarrow{a.s.} 0$

为让  $\frac{T_n}{n} \xrightarrow{a.s.} EX$ , 先证  $\frac{T_n - ET_n}{n} \xrightarrow{a.s.} 0$ , 再由  $\frac{ET_n}{n} \xrightarrow{a.s.} EX$  和  $\frac{T_n}{n} \xrightarrow{a.s.} EX$ .

(1) Fix  $\alpha > 0$ , 全  $n_k = [\alpha^k]$

$\Pr\{T_{n_k} - ET_{n_k} > \varepsilon n_k\} \leq \sum_{k=1}^{\infty} \frac{1}{\varepsilon^2 n_k^2} \sum_{m=1}^{n_k} EY_m^2$

$\leq C_\alpha \sum_{m=1}^{\infty} \frac{1}{m^2} EY_m^2$

$= C_\alpha \varepsilon \sum_{m=1}^{\infty} \frac{1}{m^2} EX^2 I(|X| \leq m)$

$< \infty$  [CLAIZM]. 由B-C 可得  $\frac{T_{n_k} - ET_{n_k}}{n_k} \xrightarrow{a.s.} 0$

当  $n < n < n_{k+1}$  时,  $\frac{T_n}{n} \leq \frac{T_{n_{k+1}}}{n} \rightarrow EX \cdot \alpha \Rightarrow \lim \frac{T_n}{n} \leq EX$  同理  $\lim \frac{T_n}{n} \geq \frac{1}{\alpha} EX$

全  $\alpha > 1$  时  $\frac{T_n}{n} \xrightarrow{a.s.} EX \Rightarrow \frac{S_n}{n} \xrightarrow{a.s.} EX$ .

一般地,  $\frac{S_n}{n} = \frac{\sum X_k - \sum Y_k}{n} \rightarrow EX - EX = EX$ . a.s. #

(CLAI)  $\Pr[X^2 I(|X| \leq m)] \frac{1}{m^2} < \infty$

(mid1)  $= E \sum_{m=1}^{\infty} X^2 I(|X| \leq m) \frac{1}{m^2}$

$\leq EX^2 \cdot \frac{1}{\alpha^2} < \infty$

(mid2)  $= \sum_n \sum_{m \leq n} EX^2 I(|n - \lfloor X \rfloor| \leq n) \cdot \frac{1}{m^2}$

$\leq \sum_n \sum_{m \leq n} P(n - \lfloor X \rfloor \leq n) n^2 \cdot \frac{1}{m^2}$

$= \sum_n \sum_{m \leq n} \frac{1}{m^2} n^2 P(n - \lfloor X \rfloor \leq n) \approx EX < \infty$

Thm.  $X_1, X_2, \dots$  i.i.d.  $\# EX^t = \infty$ ,  $EX < \infty$ , 则  $\frac{S_n}{n} \xrightarrow{a.s.} EX$

Rmk.  $E|X| = \infty \Rightarrow \lim \frac{S_n}{n} = \infty$ , a.s.

pf.  $\forall M > 0$ ,  $X_i^M \triangleq \min(X_i, M) \xrightarrow{a.s.} E(X_i^M)$

$\therefore \lim \frac{S_n}{n} \geq E(X_i^M)$  a.s.  $\sum M > \infty$  且  $\lim \frac{S_n}{n} \geq \infty$  #

(Gnny's).  $X_i$  i.i.d.  $F$ ,  $F_n(x) = \frac{1}{n} \sum I(X_i \leq x)$ ,  $\Pr[\sup |F_n(x) - F(x)| > 0] \rightarrow 0$  a.s.

(Glivenko-Cantelli).  
 pf:  $F_n(x) \xrightarrow{a.s.} F(x)$  By SLN.

① 证明连续。  
 $\forall \varepsilon > 0$ ,  $\exists -\delta = x_0 < \dots < x_{m+1} = \infty$ , s.t.  $|F(x_i) - F(y_{i-1})| \leq \varepsilon$ ,  $\forall i$

$\exists n > n_0$ , s.t.  $\forall n > n_0$ ,  $|F_n(x_i) - F(x_i)| \leq \varepsilon$ ,  $\forall i$

$\Rightarrow \forall x \in [x_{i-1}, x_i], |F_n(x) - F(x)| \leq |F_n(x_i) - F(x_{i-1})| \leq 2\varepsilon$ .

同理  $F_n(x) - F(x) \geq 2\varepsilon$

② 一般证  $F$ ,  
 找  $F$  的所有跳跃点  $\geq \varepsilon$  的点  $x_0^*, \dots, x_u^*$ , 则  $F(x) - \sum \Delta F(x_i^*) := G(x)$  可用 ①  
 而  $x_i^*$  可由 SLN 逼近.

Thm. (Kolmogorov ineq.)  $X_k$  独立,  $EX_k = 0$ ,  $EX_k^2 < \infty$ , 则  $\forall \varepsilon > 0$ ,  $\Pr(\max_{k \leq n} |S_k| \geq \varepsilon) \leq \frac{1}{\varepsilon^2} \sum_{k=1}^n EX_k^2$

进一步, 若  $|X_k| \leq C < \infty$ , 则  $1 - \frac{(\varepsilon + c)^2}{EX_k^2} \leq \Pr(\max_{k \leq n} |S_k| \geq \varepsilon)$ .

pf:  $T \triangleq \inf\{k: |S_k| \geq \varepsilon\}$ , 则  $\Pr[\max_{k \leq n} |S_k| \geq \varepsilon] = \Pr[T \leq n] = \Pr[T = k]$

$ES_n^2 \geq ES_n^2 I(T \leq n) = \sum_{k=1}^n ES_k^2 I(T=k) = \sum_{k=1}^n ES_k^2 I(T=k) + ES_n I(T=k) \geq \sum_{k=1}^n ES_k^2 I(T=k) + \Pr(T=n)$

另一方面,  $ES_n^2 I(T \leq n) = \sum_{k=1}^n ES_k^2 I(T=k) + \frac{E(S_n - S_{k-1})^2 I(T=k)}{S_{(k-1)+1}^2}$

$\leq (\varepsilon + c)^2 \Pr(T \leq n) + ES_n^2 \Pr(T \leq n)$

$ES_n^2 I(T > n) \leq \varepsilon^2 \Pr(T > n)$

$\therefore ES_n^2 \leq (\varepsilon + c)^2 + ES_n^2 \Pr(T \leq n) + \varepsilon^2 (1 - \Pr(T \leq n))$

$\Rightarrow \Pr(T \leq n) \geq \frac{ES_n^2 - \varepsilon^2}{(\varepsilon + c)^2 + ES_n^2 - \varepsilon^2} = 1 - \frac{(\varepsilon + c)^2}{(\varepsilon + c)^2 + ES_n^2} \geq 1 - \frac{(\varepsilon + c)^2}{ES_n^2}$  #

Thm (Kolmogorov 0-1 Law).  $X_k$  独立,  $y_n = \sigma\{X_m, m \leq n\}$ ,  $y \triangleq \bigcap_{n=1}^{\infty} y_n$ , 则  $\forall A \in \mathcal{Y}$ ,  $P(A) = 0$  或 1

pf:  $\Pr[y \in \sigma\{X_1, \dots, X_n\}] \leq \Pr[y \in \sigma\{X_1, \dots, X_n\} \cap A] = \Pr[A \cap A] = \Pr[A]$  #

EE of 二级数定理  
 $\sum P(\max_{1 \leq k \leq n} |X_{n+1} + \dots + X_{n+k}| \geq \varepsilon) \geq 1 - \frac{(\varepsilon+c)^2}{E(S_{n+m}-S_n)^2} \rightarrow 1 \quad (m \rightarrow \infty)$

1° 对称化方法 ( $X-X'$  为  $X$  的对称化,  $\frac{1}{2}P(|X-m| \geq x) \leq P(|X-X'| \geq x) \leq 2P(X-1 \geq x)$ , 由)  
(从而,  $E|X| < \infty \Leftrightarrow E|X-X'| < \infty$ )

取独立复制  $\{X'_n\}$ ,  $X'_n \stackrel{\text{def}}{=} X_n - X_n$ , 则

$$|\tilde{X}_n| \leq 2C, \quad E\tilde{X}_n = 0, \quad \text{Var}(\tilde{X}_n) = 2\text{Var}(X_n)$$

CLAIM: 1°  $E\tilde{X}_n = 0$  成立

$$\sum \text{Var}(X_n) \text{ a.s. 收敛} \Rightarrow \sum \text{Var}(X_n - X'_n) \text{ a.s. 收敛} \Rightarrow \sum \text{Var}(X_n) < \infty \Rightarrow \sum \text{Var}(X'_n) < \infty$$

EE of 三级数:  $i \cup Y_n = X_n I(|X_n| \leq c)$   
 $\left(\begin{array}{l} \text{若 } \sum P(X_n \neq Y_n) < \infty \Rightarrow P(X_n \neq Y_n, i.o.) = 0 \\ \text{由 } \sum \text{Var}(Y_n) < \infty, \sum EY_n \text{ 收敛} \text{ 知 (二级数) } \sum Y_n \text{ a.s. 收敛.} \end{array}\right)$

(2) CLAIM: 1° 成立. 从而  $\sum Y_n \text{ a.s. 收敛} \Rightarrow \sum Y_n \text{ a.s. 收敛} \Rightarrow \sum EY_n, \sum \text{Var}Y_n < \infty$ .

下证 CLAIM.

$$\therefore \sum P(|X_n| > C) < \infty \stackrel{\text{Kolmogorov}}{\Leftrightarrow} P(|X_n| > C, i.o.) = 0, \forall C \Rightarrow X_n \xrightarrow{a.s.} 0 \Leftarrow \sum X_n < \infty.$$

EE of Kronecker:  $i \cup b_n = \sum_{k=1}^n \frac{X_k}{a_k} \rightarrow b$ ,  $b_0 = 0$

$$\begin{aligned} \text{RJ } X_{n+1} &= (b_{n+1} - b_n) a_{n+1} \\ \therefore \frac{1}{a_n} \sum_{k=1}^n X_k &= \frac{1}{a_n} \sum_{k=1}^n a_k (b_k - b_{k-1}) = \frac{1}{a_n} [b_1(a_1 - a_2) + \dots + b_{n+1}(a_{n+1} - a_n) + b_n a_n] \\ &= b_n - \frac{1}{a_n} \sum_{k=1}^n a_k (a_{k+1} - a_k) \\ &\rightarrow b - b = 0 \end{aligned}$$

EE of SLN:  $(\Rightarrow) \frac{S_n}{n} - \frac{S_{n-1}}{n} \rightarrow 0, \text{ a.s.} \Rightarrow \sum P(|X_n| > n) < \infty \Rightarrow E|X| < \infty$

由 SLN 定义,  $\frac{S_n}{n} \rightarrow EX$ , a.s. 则  $EX = a$ .

(X) ( $\Leftarrow$ ) 若  $EX = 0$ , 由 Kronecker, 只须  $\sum \frac{X_n}{n}$  收敛

$$\begin{cases} \text{由三级数, } \Leftrightarrow \sum P(|X_n| > n) < \infty \\ \text{由 } \sum E \frac{X_n}{n} I(|X_n| \leq n) \text{ 收敛} \\ \text{由 } \sum \text{Var} \left( \frac{X_n}{n} I(|X_n| \leq n) \right) < \infty \end{cases}$$

$$1^\circ \sum P(|X_n| > n) \leq E|X| < \infty$$

$$3^\circ \sum \text{Var} \left( \frac{X_n}{n} I(|X_n| \leq n) \right) \leq \sum \frac{X_n^2}{n^2} I(|X_n| \leq n) \lesssim E|X|^2 \frac{1}{n} \rightarrow 0 \quad \text{或} \quad \sum \text{Var} \left( \frac{X_n}{n} I(|X_n| \leq n) \right) \leq \sum E \frac{X_n^2}{n^2} I(|X_n| \leq n) \leq E|X|^2 \frac{1}{n} \rightarrow 0$$

但<sup>2</sup> 不对.

1.  $\frac{S_n}{n} \xrightarrow{a.s.} 0 \Leftrightarrow H \gg, P\left(\left|\frac{S_n}{n}\right| > \varepsilon, i.o.\right) = 0 \Leftrightarrow \sum P\left(\left|\frac{S_n}{n}\right| > \varepsilon\right) < \infty$

若先让  $\frac{S_{nk}}{nk} \xrightarrow{a.s.} 0$ ,  $\frac{S_n}{n} \xrightarrow{a.s.} \frac{S_{nk}}{nk}$

2.  $\frac{S_n}{n} \xrightarrow{a.s.} 0 \Leftrightarrow \sum_{k=1}^n \frac{X_k}{n}$  a.s. 收敛.

Def  $\sum_{n=1}^{\infty} X_n$  a.s. 收敛:  $P\left(\sum_{n=1}^{\infty} X_n(w) \text{ 收敛}\right) = 1$

则  $\sum_{k=1}^n X_k \xrightarrow{a.s.} S \Leftrightarrow \sum_{n=1}^{\infty} X_n \text{ a.s. 收敛.}$

Theorem.  $X_n$  独立,  $EY_n = 0$ ,  $\sum EY_k^2 < \infty$ , 则  $\sum X_n$  a.s. 收敛.

pf: 由 Kolmogorov inequality.  $P\left(\max_{1 \leq m \leq n} |S_m - S_{m-1}| \geq \varepsilon\right) \leq \frac{1}{\varepsilon^2} \sum_{m=1}^n EY_m^2$

$\sum_{m=1}^{\infty} EY_m^2 < \infty$ ,  $P\left(\max_{1 \leq m \leq n} |S_m - S_{m-1}| \geq \varepsilon\right) \leq \frac{1}{\varepsilon^2} \sum_{m=1}^{\infty} EY_m^2$

$\therefore \lim_{n \rightarrow \infty} P\left(\max_{1 \leq m \leq n} |S_m - S_{m-1}| \geq \varepsilon\right) = 0$

$\therefore S_n$  a.s. 收敛. (由  $S_n$  a.s. Cauchy 3) #

Cor  $\left(\begin{array}{l} \text{三级数, II} \\ \text{或 } \sum \text{Var}(X_n) < \infty \end{array}\right) \Rightarrow \sum (X_n - EX_n) \text{ a.s. 收敛} \Rightarrow \sum \text{Var}(X_n) < \infty \Rightarrow \sum X_n \text{ a.s. 收敛.}$

Theorem.  $X_n$  独立,  $\exists C > 0$ , s.t.  $|X_n| \leq C$ , a.s.

则 1° 若  $\sum X_n$  a.s. 收敛, 则  $\sum EX_n, \sum \text{Var}(X_n)$  收敛.

2° 若  $EX_n = 0$ ,  $\sum \text{Var}(X_n) = 0$ , 则  $\sum X_n$  a.s. 收敛.

Theorem (Kolmogorov 三级数定理)  $X_n$  独立,  $\sum X_n$  a.s. 收敛  $\Leftrightarrow \exists C > 0$ , s.t.

$$1^\circ \sum P(|X_n| > C) < \infty$$

$$2^\circ \sum EX_n I(|X_n| \leq C) \text{ 收敛}$$

$$3^\circ \sum \text{Var}(X_n I(|X_n| \leq C)) < \infty$$

$\Rightarrow: \forall C > 0, 1^\circ 2^\circ 3^\circ$  收敛.

lem. (Kronecker) 实数列  $\{X_n\}$ , 正数列  $\{a_n\}$  且  $a_n \rightarrow 0$ , 则当级数  $\sum \frac{X_n}{a_n}$  收敛, 有  $\frac{1}{a_n} \sum X_n \rightarrow 0$

Theorem (SLLN)  $X_n$  i.i.d., 则  $\exists a$ , s.t.  $\frac{S_n}{n} \xrightarrow{a.s.} a \Leftrightarrow EX = a$

$\Leftrightarrow: \exists Y_n = X_n I(|X_n| \leq n)$ , 由 1° 且  $P(X_n \neq Y_n, i.o.) = 0$ , 下证  $\frac{1}{n} \sum Y_n \xrightarrow{a.s.} EX$ .

由于 3° 仍成立,  $\sum \text{Var}(\frac{Y_n}{n}) < \infty \Rightarrow \sum (\frac{Y_n - EX}{n}) \text{ a.s. 收敛}$ , 由 Kronecker,  $\frac{1}{n} \sum (Y_n - EX) \rightarrow 0$  a.s.

$$\therefore \frac{1}{n} \sum Y_n = \frac{1}{n} \sum_{k=1}^n EX_k I(|X_k| \leq k) \rightarrow EX \quad \therefore \frac{1}{n} \sum Y_n \xrightarrow{a.s.} EX. \quad \#$$

Prf of M-Z SLLN: 设  $r \neq 1$ . 不妨  $EX=0$ , 则  $\frac{S_n}{n^r} \xrightarrow{a.s.} 0$ ,  
 由 Kronecker, 只须  $\sum \frac{Y_n}{n^r}$  a.s. 收敛.  
 由三级数,  $X$  须  $\begin{cases} 1^\circ \sum P(|X_n| > n^{\frac{1}{r}}) < \infty \\ 2^\circ E V_{\frac{1}{n^r}} I(|X_n| \leq n^{\frac{1}{r}}) \text{ 收敛.} \\ 3^\circ \sum \text{Var}(\frac{X_n}{n^r} I(|X_n| \leq n^{\frac{1}{r}})) < \infty \end{cases}$   
 $1^\circ \sum P(|X_n| > n^{\frac{1}{r}}) = \sum P(|X|^r > n) < \infty$   
 $3^\circ \sum \text{Var}(\frac{Y_n}{n^r} I(|X_n| \leq n^{\frac{1}{r}})) \leq \sum E \frac{X^2}{n^r} I(|X| \leq n^{\frac{1}{r}})$   
 $\leq EX^2 \sum \frac{1}{n^{\frac{2}{r}}} = EX^r < \infty$   
 $2^\circ \sum |E \frac{Y_n}{n^r} I(|X_n| \leq n^{\frac{1}{r}})| \leq \sum E \frac{|X_n|}{n^r} I(|X_n| \leq n^{\frac{1}{r}}) < \infty. \text{ if } 0 < r < 1, \text{ 方法同上.}$   
 $= \sum |E \frac{X_n}{n^r} I(|X_n| > n^{\frac{1}{r}})| \leq \sum E \frac{|X|}{n^r} I(|X| > n^{\frac{1}{r}}) \leq EX|X|^{\frac{1}{r}} < \infty$

Prf of Thm (用  $g_n(x)$  控制的 SLLN)  
 由 Kronecker, 只须  $\sum \frac{Y_n}{a_n}$  a.s. 收敛.  $\Leftrightarrow \begin{cases} 1^\circ \sum P(|X_n| > a_n) < \infty \\ 2^\circ \sum E \frac{X_n}{a_n} I(|X_n| \leq a_n) \text{ 收敛} \\ 3^\circ \sum \text{Var}(\frac{X_n}{a_n} I(|X_n| \leq a_n)) < \infty \end{cases}$

$$\begin{aligned} 1^\circ \sum P(|X_n| > a_n) &\leq \sum E \frac{g_n(X_n)}{g_n(a_n)} I(|X_n| > a_n) \leq \sum E \frac{g_n(X_n)}{g_n(a_n)} < \infty. \\ 3^\circ \sum \text{Var}(\frac{X_n}{a_n} I(|X_n| \leq a_n)) &\leq \sum E \left( \frac{g_n(X_n)}{g_n(a_n)} \right)^2 I(|X_n| \leq a_n) \\ &\leq \sum E \left( \frac{g_n(X_n)}{g_n(a_n)} \right)^2 \stackrel{(i)}{\leq} \frac{1}{a_n^2} \leq \frac{g_n(x)}{g_n(a_n)}; \text{ 若有 } (ii), \frac{X_n^2}{a_n^2} \leq \frac{g_n(X_n)}{g_n(a_n)} \leq \frac{g_n(x)}{g_n(a_n)} \\ &\leq \sum E \frac{g_n(a_n)}{g_n(a_n)} < \infty \end{aligned}$$

$$2^\circ \sum |E \frac{Y_n}{a_n} I(|X_n| \leq a_n)| \leq \sum E \frac{|X_n|}{a_n} I(|X_n| \leq a_n) \stackrel{(i)}{\leq} \sum E \frac{g_n(|X_n|)}{g_n(a_n)} < \infty$$

$$\text{或 } |E \frac{Y_n}{a_n} I(|X_n| \leq a_n)| = |E \frac{Y_n}{a_n} I(|X_n| > a_n)| \leq \sum E \frac{|X_n|}{a_n} I(|X_n| > a_n) \stackrel{(ii)}{\leq} \sum E \frac{g_n(|X_n|)}{g_n(a_n)} < \infty \quad \#$$

Thm (Marcin Riewics-Zygmund SLLN)  $X_n, i.i.d.$  存在常数  $a, b$  使  $\frac{S_n - na}{nb} \xrightarrow{a.s.} 0 \Leftrightarrow E|X|^r < \infty$   
 应用:  $E|X| < \infty \Leftrightarrow \max_{1 \leq k \leq n} |X_k| \xrightarrow{a.s.} 0 \Leftrightarrow \max_{1 \leq k \leq n} \frac{|X_k|^2}{n^2} \xrightarrow{a.s.} 0$   
 $\Rightarrow \max_{1 \leq k \leq n} \frac{|X_k|^2}{n^2} \leq \frac{1}{n^2} \sum_{k=1}^n |X_k|^2 \xrightarrow{a.s.} 0 \quad (r=\frac{1}{2}) \quad \#.$   
 $\therefore X_n, i.i.d. EX=0, EX^2 < \infty \Rightarrow \frac{S_n}{n} \xrightarrow{a.s.} 0, \forall 0.5 < r < 1, \text{ 但 } \frac{S_n}{n^r} \xrightarrow{a.s.} 0.$   
 $\text{Thm. } \frac{S_n}{\sqrt{n(\log n)^{\frac{1}{2}+\varepsilon}}} \xrightarrow{a.s.} 0. \text{ Pf: 由 Kronecker, 只须 } \sum \frac{Y_n}{\sqrt{n(\log n)^{\frac{1}{2}+\varepsilon}}} \text{ a.s. 收敛.}$

Thm (Hartman-Winter 重对数律)  $X_n, i.i.d.$ ,  $EX=0, E|X|^2 = 0 < \infty$ , 则  $\lim_{n \rightarrow \infty} \frac{S_n}{\sqrt{2 \log n}} = 0, a.s.$   
 $\text{证: } \lim_{n \rightarrow \infty} \frac{S_n}{\sqrt{2 \log n}} = 0, a.s.; \lim_{n \rightarrow \infty} \frac{|S_n|}{\sqrt{2 \log n}} = 0, a.s.; \lim_{n \rightarrow \infty} \frac{|S_n|}{\sqrt{2 \log n}} = 0, a.s. (\text{因为 } \lim_{n \rightarrow \infty} \frac{|S_n|}{n} = 0, a.s.)$

Thm.  $X_n$  独立,  $g_n(x) > 0$  偶在  $x > 0$  时成立且满足

(i)  $\frac{g_n(x)}{x}$  非升 in  $x > 0$ .

或 (ii)  $\frac{x}{g_n(x)} \frac{g_n(x)}{x^2}$  非升 in  $x > 0$ .

若  $\exists g_n \uparrow \infty$ ,  $\sum \frac{E g_n(x_n)}{g_n(a_n)} < \infty$ , (若满足(ii), 还要有  $E X_n = 0$ ), 则  $\frac{S_n}{a_n} \xrightarrow{a.s.} 0$ .

Rmk. 常取  $g_n(x) = |X|^r, 0 < r \leq 2$

•  $X_n, i.i.d. \frac{\log n}{n} \sum_{k=1}^n \frac{|X_k|}{\log k} \xrightarrow{a.s.} 0, a.s. \Leftrightarrow E|X| < \infty, EX=0$

• 若  $E X=0, E|X| \log |X| < \infty$ , 则  $\left( \frac{\log n}{n} \right) \sum_{k=1}^n |X_k| \xrightarrow{a.s.} 0$ . (自行练习!) (截尾)  
 (约定,  $|X| < 1$  时,  $\log |X| = \log (|X| \vee 1) = 0$ )

重要不等式:  $\forall x > 0 \quad \frac{x}{1+x^2} e^{-\frac{x^2}{2}} \leq \int_x^\infty e^{-\frac{t^2}{2}} dt \leq \frac{1}{x} e^{-\frac{x^2}{2}}$  特别  $1 - F(x) \approx \frac{1}{\sqrt{2\pi}x} e^{-\frac{x^2}{2}}$   
 $\text{Pf: } \int_x^\infty e^{-\frac{t^2}{2}} dt \leq \int_x^\infty \frac{t}{x} e^{-\frac{t^2}{2}} dt = RHS.$   
 $\int_x^\infty e^{-\frac{t^2}{2}} dt \geq \int_x^\infty \frac{1+x^2}{1+x^2} e^{-\frac{t^2}{2}} dt = \frac{x^2}{1+x^2} \cdot \frac{e^{-\frac{x^2}{2}}}{t} \Big|_{t=x}^\infty = LHS \quad \#$

• ( Levy 不等式)  $X_n$  独立, 则  $\forall x > 0, P(\max_{1 \leq k \leq n} (S_k + \text{med}(S_n - S_k)) > x) \leq 2P(S_n > x)$   
 $P(\max_{1 \leq k \leq n} |S_k + \text{med}(S_n - S_k)| > x) \leq 2P(S_n > x).$

Pf: 记  $T = \inf \{k : S_k + \text{med}(S_n - S_k) > x\}$

LHS =  $P(T \leq n) = \sum_{k=1}^n P(T=k) \leq \sum_{k=1}^n P(T=k) \cdot 2P(S_n - S_k \geq \text{med}(S_n - S_k))$

$\stackrel{\text{Pf.}}{=} \sum_{k=1}^n P(T=k, S_n - S_k \geq \text{med}(S_n - S_k))$

$\leq \sum_{k=1}^n P(T=k, S_n > x) = RHS.$

$\therefore$  (中位数估计)  $|\text{med}(S_n - S_k)| \leq \sqrt{2 \text{Var}(S_n)} \quad \text{Chernoff} \quad \therefore P(|S_n - S_k| \leq \sqrt{2 \text{Var}(S_n)}) \geq 1 - \frac{1}{2} = \frac{1}{2}.$

特别,  $EX_i=0, E|X_i|^2 < \infty, \forall i \quad P(\max_{1 \leq i \leq n} S_i > x) \leq 2P(S_n > x - \sqrt{2 \sum_{i=1}^n EX_i^2})$

$\sum_i S_n$  依 P 收斂  $\Leftrightarrow \sum_i X_n$  a.s. 收斂.

若  $S_n \xrightarrow{P} S$ , 则  $S_n \xrightarrow{a.s.} S$ .

$$\begin{aligned} & \exists \varepsilon, \exists N, \text{s.t. } \sum_i P(|S_{n+k} - S_n| > 2^{-k}) < \infty \Rightarrow S_{n+k} \xrightarrow{a.s.} S \\ & \sum_k P(\max_{n \geq N} |S_n - S_{n+k} + \text{med}(S_{n+k} - S_n)| > 2^{-k}) \stackrel{\text{Levy}}{\leq} \sum_k 2 P(|S_{n+k} - S_n| > 2^{-k}) < \infty \\ (\star) & \quad \max_{n \geq N} |S_n - S_{n+k} + \text{med}(S_{n+k} - S_n)| \rightarrow 0, \text{ a.s.} \\ & \quad \forall \varepsilon > 0, \max_{n \geq N} P(|S_{n+k} - S_n| > \varepsilon) \rightarrow 0 \\ (\star\star) & \quad \max_{n \geq N} |\text{med}(S_{n+k} - S_n)| \rightarrow 0. \quad (\text{对上式取 } k \text{ 充分大, 使 } \max P(\cdot) < \frac{1}{2}, \text{ 则 } \max |\text{med}(\cdot)| \leq \varepsilon) \end{aligned}$$

由 (†) 和 (‡) 得证. #

PF of Hoeffding: 不妨  $EX_i = 0$ .

$$\begin{aligned} & \forall t > 0, P(S_n \geq nx) \leq e^{-tnx} E[e^{tX_n}] \\ & \quad \because e^{tx} \leq \frac{x}{b_i - a_i} e^{t(b_i)} + \frac{b_i - x}{b_i - a_i} e^{ta_i}, \forall x \in [a_i, b_i] \\ & \quad \therefore E[e^{tx}] \leq -\frac{a_i}{b_i - a_i} e^{tb_i} + \frac{b_i}{b_i - a_i} e^{ta_i} = e^{g(u)}, \\ & \quad \text{其中 } g(u) = \ln(1 - \theta + \theta e^u) - \theta u, \theta = -\frac{a_i}{b_i - a_i}, u = t(b_i - a_i) \\ & \quad \because g'(0) = g(0) = 0, g''(u) \leq \frac{1}{4} \\ & \quad \therefore g(u) \leq \frac{u^2}{8} = \frac{t^2(b_i - a_i)^2}{8} \\ & \quad \therefore P(S_n \geq nx) \leq e^{-tnx + \frac{t^2(b_i - a_i)^2}{8}} \quad (\text{minimizer } t = \frac{4n}{\sum b_i - a_i}) = \text{RHS}. \quad \# \end{aligned}$$

5.4.  $f(a, a)$  的 ch.f.:  $\frac{\sin at}{at}$

分布律  $P(x) = \frac{a-x}{a^2}$  在  $[a, a]$  的 ch.f.:  $\left(\frac{\sin \frac{at}{2}}{\frac{at}{2}}\right)^2$

(练习) 设  $X$  有界连续,  $Eg(X) = \lim_{\delta \rightarrow 0} Eg(X + \delta Z) = \lim_{\delta \rightarrow 0} \int g(x) P_{X+\delta Z}(dx)$

根据  $f_X(t) \Rightarrow P_{X+\delta Z}(x) \Rightarrow Eg(X) \Rightarrow X$  的 d.f.

$\therefore P(x) = \lim_{\delta \rightarrow 0} P_{X+\delta Z}(x) = \frac{1}{2\pi} \int f_X(t) e^{itx} dt$ , 下让  $p(x)$  为  $X$  的 d.f.

④ 若有界区间 I,  $P(X \in E) = \lim_{\delta \rightarrow 0} P(X + \delta Z \in E) = \lim_{\delta \rightarrow 0} \int_I P_{X+\delta Z}(x) dx = \int_I p(x) dx$ . #

Theorem.  $X_n$  独立, 则  $\sum X_n$  依 P 收斂  $\Leftrightarrow \sum X_n$  a.s. 收斂  $\Leftrightarrow \sum X_n$  依 d 收斂.

Theorem (Hoeffding)  $X_i$  独立,  $a_i \leq X_i \leq b_i$ , 则  $\forall \varepsilon > 0$ ,  $P(S_n - ES_n \geq \varepsilon n) \leq \exp[-\frac{2n\varepsilon^2}{\sum (b_i - a_i)}]$ .

指數不等  
Fuk-Nagaev 不等  
Talagrand 不等  
Concentration 不等  
(概率不等)

Theorem.  $X_j$  独立. (1) 若  $0 < r \leq 1$ ,  $E|X_j|^r < \infty$ , 则  $E|S_n|^r \leq \sum E|X_j|^r$

(2)  $1 \leq r \leq 2$ ,  $E|X_j|^r < \infty$ ,  $EX_j = 0$ , 则  $E|S_n|^r \leq 2^{2-r} \sum E|X_j|^r$

(3)  $r > 2$ ,  $E|X_j|^r < \infty$ ,  $EX_j = 0$ , 则  $\exists C(r) > 0$ , s.t.

$$E|S_n|^r \leq C(r) \left( \sum E|X_j|^r + \left( \sum EX_j^2 \right)^{\frac{r}{2}} \right)$$

• 特征函数  $f(t) = Ee^{itX} = E\cos(tx) + iE\sin(tx)$

prop. (i)  $|f(t)| \leq 1$

(ii)  $f(-t) = \bar{f}(t)$

(iii)  $f(t)$  一致连续.  $\sup_t |f(t+h) - f(t)| = \sup_t |Ee^{ithX} - Ee^{itX}|$

(iv)  $f(t)$  非负定. (判定  $f$  为特征的必要条件).  $\leq \sup_t |Ee^{ithX} - 1| \xrightarrow{P.T.} 0$  ( $h \rightarrow 0$ )

i.e.,  $\forall t_1, \dots, t_n, \lambda_1, \dots, \lambda_n \in \mathbb{C}$ ,  $\sum_{k,j} f(t_k + t_j) \lambda_k \bar{\lambda}_j \geq 0$

(v)  $f_{X+tb}(t) = e^{itb} f_X(at)$ .  $\sup_t |f_{X+tb}(t)| \leq E(\sum_k f(t_k) \lambda_k)(\sum_j \bar{\lambda}_j)$

(vi)  $f_1, f_2$  为特征, 则  $f_1 f_2$  为特征  $= f_{X+Y}$

$f$  为 ch.f., 则  $|f|^2$  为 ch.f.  $= f_{X-X}$ .

$f_i$  为 ch.f.,  $a_i \in \mathbb{C}$ ,  $\sum a_i = 1$ , 则  $\sum a_i f_i$  为 ch.f. 分布为  $\sum a_i F_i(x)$ .

Theorem (Parseval) 设  $X, Y$  分布  $F_x, F_y$ , 特征  $f_x, f_y$ , 则

$$\int f_x(t) dF_y(t) = \int f_y(t) dF_x(t).$$

若  $X, Y$  独立,  $Ee^{iXt} = Ee^{iYt} = \int f_x(t) dF_x(t)$

若  $Y \sim N(0, \sigma^2)$ ,

$$\begin{aligned} LHS &= \int f_x(t) \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{t^2}{2\sigma^2}} dt, \quad RHS = \int e^{-\frac{\sigma^2 X^2}{2}} dF_x(x) \\ &= \int e^{itX} dF_x(x) \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{t^2}{2\sigma^2}} dt \\ &= \frac{1}{\sqrt{2\pi}\sigma} \int \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{t^2}{2\sigma^2}} dF_x(x) \\ &= \frac{1}{\sqrt{2\pi}\sigma} P_{X+\sigma Z}(0), \quad Z \sim N(0, 1) \text{ 与 } X \text{ 独立} \end{aligned}$$

$$\begin{aligned} &\Rightarrow P_{X+\sigma Z}(0) = \frac{1}{\sqrt{2\pi}} \int f_x(t) e^{\frac{\sigma^2 t^2}{2}} dt \\ &\Rightarrow P_{X+\sigma Z}(x) = P_{X+\sigma Z-x}(0) = \frac{1}{\sqrt{2\pi}} \int f_x(t) e^{-\frac{\sigma^2 t^2}{2}} dt \end{aligned}$$

由 Parseval

$$\int f_X(t) dF_Y(t) = \int f_Y(t) dF_X(t)$$

$$\text{取 } Y \sim U(-u, u), \text{ 则 } \int_{-u}^u f_Y(t) dt = \int_{-u}^u f_X(t) dt$$

$$= \int_{|t| \leq \frac{u}{2}} dF_X(t) + \int_{|t| > \frac{u}{2}} dF_X(t)$$

$$= 1 - \frac{1}{2} P(|X| \geq \frac{u}{2})$$

$$\therefore P(|X| \geq \frac{u}{2}) \leq \frac{1}{u} \int_{-u}^u (1 - f_X(t)) dt.$$

用  $u$  代替  $\frac{u}{2}$  得证.

由 (特征矩等):

$$|f(t) - \sum_{j=0}^{k-1} t^j EX^j| \leq E \left| e^{itX} - \sum_{j=0}^{k-1} t^j X^j \right|$$

$$\stackrel{\text{lem}}{\leq} \frac{1}{(k+1)!} E |X|^{k+1} (t) \times 2(k+1) \stackrel{\text{DCT}}{\rightarrow} o(1+t^k)$$

用归纳,  $k=1$  时,  $\frac{f(t+h)-f(t)}{h} = \int \frac{e^{itX}(e^{ih}-1)}{h} dF(x) \xrightarrow{\text{PCT}} \int (iX) e^{itX} dF(x)$

若  $(k-1)$  时成立.  $f^{(k-1)}(t) = \int_{-\infty}^{+\infty} (ix)^{k-1} e^{itX} dF(x)$ , 且有  $E|X|^k < \infty$ .

$$\frac{f^{(k-1)}(t+h) - f^{(k-1)}(t)}{h} = \int \frac{(ix)^{k-1} e^{itX} (e^{ih}-1)}{h} dF(x) \xrightarrow{\text{PCT}} \int (ix)^k e^{itX} dF(x)$$

由  $k$  (偶数) 所导  $\Rightarrow k$  阶矩:

$$\text{① } k=2 \quad f''(0) = \lim_{h \rightarrow 0} \frac{f(h) + f(-h) - 2f(0)}{h^2}$$

$$= \lim_{h \rightarrow 0} \int \frac{e^{ihx} + e^{-ihx} - 2}{h^2} dF$$

$$= -2 \lim_{h \rightarrow 0} \int \frac{1 - \cos hx}{h^2} dF$$

$$\therefore \int x^2 dF \stackrel{\text{PCT}}{\leq} \lim_{h \rightarrow 0} \int \frac{2(1 - \cos hx)}{h^2} dF = -f''(0) < \infty \quad (\text{进一步}, f''(0) = -\int x^2 dF)$$

② 设  $2k-2$  时成立,  $f$  在端附近有  $2k$  阶导, 则  $E|X|^{2k-2} < \infty$

$$f^{(2k-2)}(0) = (-1)^{k-1} \int X^{2k-2} e^{itX} dF$$

令  $G(x) = \int_{-\infty}^x y^{2k-2} dF(y)$ , 若  $G(\infty) \neq 0$ , 则  $\frac{G(x)}{G(\infty)}$  为 d.f.

其 c.f.  $\psi(t) = \frac{1}{G(\infty)} \int e^{itY} Y^{2k-2} dF = \frac{(-1)^{k-1} f^{(2k-2)}(0)}{G(\infty)}$  在 0 附近有  $2k$  阶导

由①知  $\int x^2 dG(x) < \infty \Rightarrow \int x^{2k} dF < \infty$

从而得证:

Thm. 若  $\int f_X(t) dt < \infty$ , 则  $X$  的密度  $p_X(x) = \frac{1}{2\pi} \int_{-\pi}^{\pi} f_X(t) e^{-itx} dt$

Thm (反演公式)  $\forall x_1 < x_2$ ,  $P((x_1, x_2)) = \frac{1}{2\pi} \int_{-\pi}^{\pi} f_X(t) dt = \lim_{T \rightarrow \infty} \frac{1}{2\pi} \int_{-T}^T \frac{e^{-itx_2} - e^{-itx_1}}{it} f_X(t) dt$

Thm.  $P(|X| \geq u) \leq \frac{u}{2} \int_{-\pi}^{\pi} (1 - f_X(t)) dt$

- 若  $|f_X(t)| \equiv 1$ , 则  $X$  退化. ( $\because |f_X(t)|^2 \equiv 1$  为  $X - \bar{X}$  的 ch.f.  $\Rightarrow P(|X - \bar{X}| \geq u) = 0, \forall u \Rightarrow X - \bar{X} \stackrel{a.s.}{=} 0 \Rightarrow X \stackrel{a.s.}{=} \bar{X}$ )
- 特征与矩:
 

c.e.g. Ch.f. 的导数不存在:  $P(X = 5^k) = 2^{-k+1}, k=0, 1, \dots$

Thm.  $\forall n \in \mathbb{N}$ ,  $\left| e^{ix} - \sum_{m=0}^n \frac{(ix)^m}{m!} \right| \leq \min \left\{ \frac{|x|^{n+1}}{(n+1)!}, \frac{2|x|^n}{n!} \right\}$

特别,  $|e^{ix} - 1| \leq |x|$

Thm (特征矩等)  $\left| e^{ix} - 1 - ix \right| \leq \frac{|x|^2}{2}$

Thm. 若  $E|X|^k < \infty$ , 则  $f(t) = \sum_{j=0}^k \frac{i^j t^j}{j!} EX^j + o(|t|^k)$

且  $f(t)$   $k$  阶可导, 则  $f^{(k)}(t) = \int_{-\infty}^{+\infty} (ix)^k e^{itX} dF(x)$

$f^{(k)}(0) = i^k EX^k$

Thm.  $k$  为偶, 若  $f$  在端附近有  $k$  阶导, 则  $X$  有  $k$  阶矩.

c.e.g.  $f$  有  $2k+1$  阶导,  $E|X|^{2k+2} < \infty$  但  $E|X|^{2k+1} < \infty$  不成立.

$P(X = \pm 1) = \frac{1}{2Cj^2 \log j}$

- 若  $f(t) = \sum_{k=0}^{\infty} \frac{(it)^k EX^k}{k!}$ , 则若  $EX^k = EY^k$ , 则  $X \stackrel{d}{=} Y$  ( $X$  的矩唯一确定分布)

Thm. (矩确定分布) 设  $X$  的矩能唯一确定分布  $EX_n^k \rightarrow EX^k, \forall k$ , 则  $X_n \xrightarrow{d} X$ .

- $X_n \xrightarrow{d} X \Leftrightarrow \{f_{n,k}\} \subset \{f_{n+1,k}\}, f_{n,k} \xrightarrow{d} f_{n+1,k}$
- Helly 定理:  $\{f_{n,k}\}, \exists \{f_{n,k}\} \subset \{f_{n+1,k}\}$ , s.t.  $X_n \xrightarrow{\text{ vague }} G$

当  $X_n$  脱壳,  $G$  是 d.f.

Thm. (特征 Taylor) 若  $\forall n \geq 1, V_n = E|X|^n < \infty$ , 且  $\limsup \frac{(V_n)^{\frac{1}{n}}}{n} = r < \infty$ ,

则当  $|t| < \frac{1}{er}$  时,  $f(t) = \sum_{n=0}^{\infty} \frac{t^n}{n!} f^{(n)}(0)$ ,  $\forall \theta \in \mathbb{R}$ .

Pf.  $|f(t+\theta) - \sum_{n=0}^{k-1} \frac{t^n}{n!} f^{(n)}(\theta)| \leq |E[e^{itX}(e^{i\theta X} - \sum_{n=0}^{k-1} \frac{(i\theta X)^n}{n!})]| \leq \frac{|t|^k}{k!} E|X|^k$

$$\left( \because \forall \varepsilon > 0, \text{ 当 } k \text{ 充分大, } (2k)^k \leq K(r+\varepsilon)^k \right)$$

$$\leq \frac{(r+\varepsilon)^k}{(k!)^k} \leq \frac{(r+\varepsilon)^k}{(er)^k} \leq \frac{1}{e^k} \quad \text{取 } \varepsilon, \text{s.t. } er/(r+\varepsilon) \leq 1/2$$

由③.  $X \sim N(0, 1)$ , 若  $EX^n = EY^n, \forall n$ , 则由下页 Thm,  $X \stackrel{d}{=} Y$ .

而  $EX_n^n \rightarrow EX^n \Rightarrow X_n \xrightarrow{d} X \Rightarrow \frac{S_n}{\sqrt{n}} \xrightarrow{d} N(0, 1)$

## 证明连续性定理:

- (1)  $Ee^{itX_n} \rightarrow Ee^{itX}$
- (2)  $P(|X_n| > r) \leq \frac{r}{2} \int_{-\frac{r}{2}}^{\frac{r}{2}} (1 - f_n(t)) dt$   
 $\lim P(|X_n| > r) \leq \frac{r}{2} \int_{-\frac{r}{2}}^{\frac{r}{2}} (1 - f(t)) dt \rightarrow 0 \quad (r \rightarrow \infty)$   $\therefore f_n(t) \xrightarrow{t \rightarrow 0} f(t) = \lim_{n \rightarrow \infty} f_n(0) = 1$
- $\lim_{n \rightarrow \infty} P(|X_n| > r) = 0 \quad \therefore X_n \text{ 为紧集}$
- 由 Helly 定理  $\forall \mu_n, \exists f_n \in \mathcal{G}, \text{ s.t. } X_{n_k} \xrightarrow{d} X$   
 由 (1)  $f_{n_k}(t) \xrightarrow{t \rightarrow 0} Ee^{itX} \Rightarrow f_{n_k}(t) = Ee^{itX} \Rightarrow X \text{ 分布与子列无关, 完全由 } f_{n_k}(t) \text{ 决定}$   
 $X_n \xrightarrow{d} X \text{ 且 } X \text{ 的 ch.f. 为 } f_n$

PF of 1em: CLAIM 1:  $|b| \leq 1$  时,  $|e^{ib} - 1 - b| \leq |b|^2$

(1) LHS =  $\left| \sum_{k=2}^{\infty} \frac{b^k}{k!} \right| \leq |b|^2 \sum_{k=2}^{\infty} \frac{1}{2^k k!} = |b|^2$ .

CLAIM 2:  $|f_m - f_m^n| \leq \theta$  时,  $|f_m - f_m^n| \leq \theta^{m-n} |f_m - f_m^n| \leq \theta^{m-n} |f_m - f_m^n|$   
 $f_m \triangleq (1 + \frac{c_m}{n}), f_m^n = e^{cn}, |c| < \gamma$   
 则当  $n$  充分大时  $|f_m^n| < \frac{\gamma}{n} < 1$ .  $\Rightarrow |f_m^n|, |f_m^n| \leq e^{\frac{\gamma}{n}}$   
 $\therefore \text{由 CLAIM 2, } |(1 + \frac{c_m}{n})^n - e^{cn}| \leq (e^{\frac{\gamma}{n}})^{n-m} \left| 1 + \frac{c_m}{n} - e^{\frac{c_m}{n}} \right|$   
 $\stackrel{\text{CLAIM 1.}}{\leq} e^{\gamma} \cdot n \cdot \left| \frac{c_m}{n} \right|^2 \leq e^{\gamma} \cdot \frac{\gamma^2}{n} \rightarrow 0 \quad (n \rightarrow \infty)$

$\therefore \lim_{n \rightarrow \infty} (1 + \frac{c_m}{n})^n = \lim_{n \rightarrow \infty} e^{cn} = e^c$

PF of CLT: WLOG,  $\mu=0$ .  $f_X(t) = 1 - \frac{t^2}{2} + o(t^2)$   
 $Ee^{it(\frac{S_n}{\sqrt{n}})} = (1 - \frac{t^2}{2n} + o(\frac{1}{n})) \xrightarrow{n \rightarrow \infty} e^{-\frac{1}{2}t^2}$

PF of Lindeberg CLT. [粗暴分析法]:  $\frac{1}{k_n} Ee^{itX_{nk}} = \frac{1}{k_n} \left( 1 - \frac{t^2}{2} + \Delta_k \right) = \frac{1}{k_n} \left( e^{-\frac{t^2}{2}} + \Delta_k \right) \xrightarrow{k_n \rightarrow \infty} e^{-\frac{t^2}{2}}$ , 但  $\Delta_k$  bounded  
 设  $X_{nk}$  为 ch.f.  $f_{nk}(t)$   
 $\Rightarrow \sigma_{nk}^2 = E[X_{nk}^2]$ .  $G_{nk} \sim N(0, \sigma_{nk}^2)$  with ch.f.  $\varphi_{nk}(t) = e^{-\frac{t^2}{2\sigma_{nk}^2}}$   
 $|Ee^{it(X_{nk} + X_{nkl})} - e^{-\frac{t^2}{2}}| = \left| \prod_{k=1}^n f_{nk}(t) - \prod_{k=1}^n \varphi_{nk}(t) \right|$   
 $\leq \sum_{k=1}^n |f_{nk}(t) - \varphi_{nk}(t)|$   
 $\leq \sum_{k=1}^n \left| \left( f_{nk}(t) - 1 + \frac{1}{2}t^2 \sigma_{nk}^2 \right) + \left| \varphi_{nk}(t) - 1 + \frac{1}{2}t^2 \sigma_{nk}^2 \right| \right|$   
 $\leq C + \sum_{k=1}^n (E[X_{nk}^2] \wedge |X_{nk}|^3) + E[G_{nk}^2 \wedge |G_{nk}|^3]$   
 $\stackrel{\text{伊顿 Elon}}{\leq} \sum_{k=1}^n E[X_{nk}^2 \wedge |X_{nk}|^3] \leq \sum_{k=1}^n \left[ E[X_{nk}^2 \wedge I(|X_{nk}| \leq \varepsilon)] + E[X_{nk}^2 \wedge I(|X_{nk}| > \varepsilon)] \right] \leq \varepsilon \quad (n \rightarrow \infty)$   
 $\leq E[X_{nk}^2] \cdot \varepsilon.$  由大数律知  $\sum_{k=1}^n E[X_{nk}^2 \wedge |X_{nk}|^3] \rightarrow 0$ .

Date: \_\_\_\_\_

Thm. 若  $\{\mu_n\}$  有  $\limsup_{n \rightarrow \infty} \frac{(\mu_n)^2}{2n} = r < \infty$ , 则  $\exists \rightarrow d.f. F$ , s.t.  $\mu_n = \int x^n dF$ ,  $n \geq 1$

从而, 若  $\{EX^n\} = \{EY^n\}$  满足上式, 则  $X \stackrel{d}{=} Y$

(从) 证明: 首先, 由  $|EX|^{2n} \leq \sqrt{EX^{2n} EY^{2n}}$  和  $\limsup_{n \rightarrow \infty} \frac{(\mu_n)^2}{2n} = r < \infty$ , 其中  $\mu_n = EX^n = EY^n$   
 由上一题知, 当  $|t| < \frac{K}{\sqrt{r}}$  时,  $f_X(t) = \sum_{k=0}^{\infty} \frac{t^k}{k!} \mu_k = f_Y(t)$

遂推知当  $|t| < \frac{K}{\sqrt{r}}$  时,  $f_X(t) = f_Y(t) \Rightarrow X \stackrel{d}{=} Y$

Thm. (连续性定理) (1)  $X_n \xrightarrow{d} X_\infty$ , 则  $\forall t \in \mathbb{R}, f_n(t) \rightarrow f_\infty(t)$ . (且在有限区间上一致收敛)  
 (2) 若  $\exists f(t)$ , s.t.  $f_n(t) \rightarrow f(t), \forall t \in \mathbb{R}$ , 且  $f(t)$  在 0 点连续,

(待与 type & Law) 例  $\exists r.v. X$ , s.t.  $X_n \xrightarrow{d} X$  且  $X$  的 ch.f. 为  $f_\infty$

Thm. 常数列  $\{f_n, g_n\}, f_n \xrightarrow{(n \rightarrow \infty)} F$  退化, 则

(1) 若  $f_n(a_n + b_n) \xrightarrow{d} G(x)$  退化, 则  $\exists a, b$ , s.t.  $G(x) = f(a + bx)$ ,  $a_n \rightarrow a, b_n \rightarrow b$ .

(2) 若  $a_n \rightarrow a, b_n \rightarrow b$ , 则  $f_n(a_n + b_n) \xrightarrow{d} f(a + bx)$ .

•  $X_n \xrightarrow{d} X \Leftrightarrow f_{X_n}(t) \rightarrow f_X(t)$

•  $\frac{S_n}{n} \xrightarrow{P} 0 \Leftrightarrow \frac{S_n}{n} \xrightarrow{d} 0 \Leftrightarrow Ee^{it\frac{S_n}{n}} \rightarrow 1$  但  $f_{X_n}(t) \rightarrow f_X(t) \not\Rightarrow X_n \xrightarrow{P} X$

证明  $X_i, i.i.d. EX_i = 0$  由 CLT:

$$Ee^{it\frac{S_n}{n}} = (Ee^{it\frac{X_1}{n}})^n = (1 + it\frac{EX_1}{n} + o(\frac{1}{n}))^n \rightarrow e^{itEX} = 1. \Rightarrow \frac{S_n}{n} \xrightarrow{P} 0. \quad \#$$

由设  $C_n \rightarrow c \in \mathbb{C}$ , 则  $(1 + \frac{C_n}{n})^n \rightarrow e^c$ .

Thm (CLT)  $X_i, i.i.d. EX_i = \mu, \text{Var}(X_i) = \sigma^2$ , 则  $\frac{S_n - n\mu}{\sqrt{n}\sigma} \xrightarrow{d} N(0, 1)$

• 三角阵列:  $X_{11} \cdots X_{1k_1} \quad \dots \quad X_{n_1} \cdots X_{n_k} \text{ 独立}$

$X_{21} \cdots X_{2k_2} \quad \dots \quad E[X_{kk}] > 0$

(3)  $E[X_{kk}]^2 = 1$ .

Thm (Lindeberg CLT). 上述 setting, 且  $\forall \varepsilon > 0$ ,  $\sum_{k=1}^{\infty} E[X_{kk}^2] I(|X_{kk}| \geq \varepsilon) \rightarrow 0$ ,  
 则  $\frac{1}{\sqrt{n}} X_{nk} \xrightarrow{d} N(0, 1)$

Rmk. (Feller 条件):  $\limsup_{n \rightarrow \infty} \max_k X_{kk}^2 = 0$  ② 无穷小条件:  $\lim_{n \rightarrow \infty} \max_k P(|X_{kk}| > \varepsilon) = 0, \forall \varepsilon > 0$ .

Thm (Lindeberg-Feller) 在 Lindeberg CLT setting 下, 以下等价.

(1)  $\sum_{k=1}^{\infty} X_{kk} \xrightarrow{d} N(0, 1)$  且  $\limsup_{n \rightarrow \infty} \max_k X_{kk}^2 = 0$

(2)  $\forall \varepsilon > 0, \sum_{k=1}^{\infty} E[X_{kk}^2] I(|X_{kk}| \geq \varepsilon) \rightarrow 0$ .

Cor (Lyapunov) 若  $\sum_{k=1}^{\infty} E[X_{kk}^3] \rightarrow 0$ , 则  $\frac{1}{\sqrt{n}} X_{nk} \xrightarrow{d} N(0, 1)$ .

$\sum_{k=1}^{\infty} E[G_{nk}^2 \wedge |G_{nk}|^3] \leq \sum_{k=1}^{\infty} E[G_{nk}^3] = C \cdot \sum_{k=1}^{\infty} \sigma_{nk}^3 \leq C \cdot \max_k \sigma_{nk}^3 \rightarrow 0$

(3)  $\max_k \sigma_{nk}^2 \leq \max_k E[X_{nk}^2] I(|X_{nk}| \leq \varepsilon) + \max_k E[X_{nk}^2] I(|X_{nk}| > \varepsilon) \leq \varepsilon \quad (n \rightarrow \infty), \forall \varepsilon > 0$ .

of lemma:

step 1:  $F_1, F_2$  均值为 0, ch  $f_1, f_2$  可积, 则  $F_1(x) - F_2(x) = \frac{1}{2\pi} \int e^{-ix} (f_1(t) - f_2(t)) dt$

② 密度  $P_j(x) = \frac{1}{2\pi} \int e^{-ixt} f_j(t) dt, j=1, 2$  (由  $f_j$  可积知有密度)

$$F_1(x) - F_2(x) = (P_1(x) - P_2(x))$$

$$= \int_{-\infty}^x (P_1(y) - P_2(y)) dy$$

$$= \frac{1}{2\pi} \int_{-\infty}^x e^{-ity} (f_1(t) - f_2(t)) dt dy$$

$$= \frac{1}{2\pi} \int_{-\infty}^x dt (f_1(t) - f_2(t)) \frac{e^{-ity}}{it}$$

由 R-L 定理, 只须  $\frac{f_1(t) - f_2(t)}{it}$  可积.

(在 0 邻域,  $f_i(t) = 1 + o(t)$ ,  $\frac{f_i(t)-1}{t} \rightarrow 0 \Rightarrow \frac{(f_1(t)-1)-(f_2(t)-1)}{t} \rightarrow 0$  在  $[0, 1]$  上 bdd)

step 2:  $h_L(x) \triangleq \frac{1-w\pi x}{\pi L^2 x^2}, x \in \mathbb{R}$  为密度且  $f_L(t) = (1-\frac{t}{L})^+$ , 记分布函数为  $H_L$ .

$$\text{取 } F_L = F * H_L, G_L = G * H_L \quad (\text{均值仍为 0}), \mathbb{E}[|F_L(x) - G_L(x)|] \leq \frac{1}{2\pi} \int_{-\infty}^x |f_L(t) - g_L(t)| dt$$

step 3: 设  $G'(x) \leq \lambda < \infty$ , 则  $\sup_x |F(x) - G(x)| \leq 2 \sup_x |F_L(x) - G_L(x)| + \frac{24\lambda}{\pi L}$  由 step 2.3.8 lemma.

③ 记  $\Delta(x) = F(x) - G(x), \eta = \sup_x |\Delta(x)|$ , 由于  $\Delta(x) \rightarrow 0, x \rightarrow \infty$  且  $G$  连续,  $WLOG \Delta(0) = \eta$

$$\begin{aligned} \sup_x |F(x) - G(x)| &\geq F_L(0) - G_L(0) \quad \text{for some } t_0 \\ &= P(X \leq t_0) - P(Y \leq t_0) \\ &= \int [P(X \leq t_0 - x) - P(Y \leq t_0 - x)] h_L(x) dx. \end{aligned}$$

$$\Delta(x_0 + s) \geq F(x_0) - G(x_0 + s) \geq \eta - s\lambda$$

$$\text{设 } \delta = \frac{1}{2}\lambda, t_0 = x_0 + \delta, \text{ 则当 } |x| \leq s \text{ 时, } \Delta(t_0 - x) \geq \eta - \lambda(\delta - x) = \frac{1}{2} + \lambda x.$$

$$\begin{aligned} \sup_x |F(x) - G(x)| &\geq \int_{|x| \leq s} \left( \frac{1}{2} + \lambda x \right) h_L(x) dx - \eta \int_{|x| > s} h_L(x) dx \\ &= \frac{1}{2} \int_{|x| \leq s} h_L(x) dx - \eta \int_{|x| > s} h_L(x) dx \\ &= \frac{1}{2} - \frac{3}{2}\eta. \int_{|x| > s} \frac{2}{\pi L^2 x^2} dx \\ &= \frac{1}{2} - \frac{6\eta}{\pi L^2 s} = \frac{1}{2} - \frac{12\lambda}{\pi L}. \end{aligned}$$

证 of lemma 2:  $|W_k| = |E e^{it \frac{X_k}{B_n}}| \leq |E X_k|^{\frac{1}{2}} |t|^{\frac{3}{2}}$

$$\begin{aligned} &\leq \left| -\frac{\partial_k^2 t^2}{2B_n} + 6B_n^{\frac{3}{2}} \right|^{\frac{1}{2}} \leq \exp \left| -\frac{\partial_k^2 t^2}{2B_n} + 6B_n^{\frac{3}{2}} \right|^{\frac{1}{2}} \\ &\leq \frac{1}{m!} \left( W_m \cdot \prod_{m=k+1}^n B_m \right)^{\frac{1}{2}} \leq \exp \left| -\sum_{m=k+1}^n \frac{\partial_m^2 t^2}{2B_m} + \sum_{m=k+1}^n \frac{E[X_m]^{\frac{3}{2}} |t|^3}{B_m} \right|^{\frac{1}{2}} \end{aligned}$$

$$\begin{aligned} &\leq \exp \left| -\frac{\partial_k^2 t^2}{2B_n} + 2B_n^{\frac{3}{2}} + \frac{|t|^3}{B_n} \ln \frac{1}{n} \right|^{\frac{1}{2}} \quad \text{④} \quad \frac{\partial_k^2 t^2}{B_n} \leq \frac{(E[X_k])^{\frac{3}{2}}}{B_n} \leq \frac{1}{L^{\frac{3}{2}}} \leq \frac{1}{4} \\ &\leq \exp \left\{ -\frac{t^2}{4} + \frac{6B_n^{\frac{3}{2}}}{2B_n} + \frac{|t|^3}{8B_n^{\frac{3}{2}}} \right\} \end{aligned}$$

$$|W_k - \frac{\partial_k^2 t^2}{2B_n}| \leq \frac{6B_n^{\frac{3}{2}}}{8B_n^{\frac{3}{2}}} \Rightarrow |W_k - \bar{W}_k| \leq \frac{6B_n^{\frac{3}{2}}}{8B_n^{\frac{3}{2}}} + \frac{6B_n^{\frac{3}{2}}}{8B_n^{\frac{3}{2}}} \quad \text{⑤} \quad \bar{W}_k \in \Omega_k, (E[X_k])^{\frac{3}{2}} \leq \bar{W}_k \cdot E[X_k]^{\frac{3}{2}} \leq \bar{W}_k$$

Thm (CLT 充要) 设  $X_k$  满足无穷小条件 ( $\forall \varepsilon > 0, \lim_{n \rightarrow \infty} \max_k P(|X_{nk}| > \varepsilon) = 0$ ),  $b \in \mathbb{R}, C > 0$ ,

$$\text{则 } \sum_{k=1}^n X_{nk} \xrightarrow{d} N(b, C) \Leftrightarrow \begin{cases} 1. \forall \varepsilon > 0, \mathbb{P}(X_{nk} > \varepsilon) \rightarrow 0 \\ 2. \sum_k \mathbb{E}(X_{nk} I(|X_{nk}| \leq 1)) \rightarrow b \\ 3. \mathbb{E} \text{Var}(X_{nk} I(|X_{nk}| \leq 1)) \rightarrow C. \end{cases}$$

Thm (Berry-Esseen Ineq)  $X_i \stackrel{i.i.d.}{\sim} \mathbb{E}X_i = 0, \mathbb{E}X_i^2 = \sigma^2, \mathbb{E}|X_i|^3 = \rho < \infty, \forall i$

$$\sup_x |\mathbb{P}(\frac{S_n}{\sqrt{n}} \leq x) - \Phi(x)| \leq \frac{C\rho}{\sqrt{n}\sigma^3}$$

lem 若  $F, G$  均值为 0,  $\lambda = \sup_x G'(x), \forall t > 0$ ,

$$\sup_x |F(x) - G(x)| \leq \frac{1}{\pi} \int_{-L}^L \left| \frac{f_F(t) - f_G(t)}{t+1} \right| dt + \frac{24\lambda}{\pi L}$$

pf of B-E:  $\lambda = \sup_x \mathbb{P}(\frac{S_n}{\sqrt{n}} \leq x) = \frac{1}{\sqrt{n}} < \frac{\lambda}{\sqrt{n}}$  WLOG  $\sigma = 1$ .

$$\begin{aligned} \therefore \sup_x |\mathbb{P}(\frac{S_n}{\sqrt{n}} \leq x) - \Phi(x)| &\leq C \cdot \int_{-L}^L \left| \frac{(f_X(\frac{t}{\sqrt{n}}))^n - e^{-\frac{t^2}{2}}}{t+1} \right| dt + \frac{C}{L} \\ &\quad (\because \Delta(t) \geq (f_X(\frac{t}{\sqrt{n}}))^n - e^{-\frac{t^2}{2}}) \\ &\leq C \cdot \int_{-L}^L \frac{n}{t+1} \left| f_X(\frac{t}{\sqrt{n}}) - e^{-\frac{t^2}{2}} \right| \Delta(t) dt + \frac{C}{L} \\ &\quad \therefore |f_X(\frac{t}{\sqrt{n}}) - 1 + \frac{t^2}{2n}| \leq \frac{1+t^2}{6n^{\frac{3}{2}}} \leq |e^{-\frac{t^2}{2}} - 1 + \frac{t^2}{2n}| \leq \frac{t^4}{8n^2} \\ &\quad \therefore |f_X(\frac{t}{\sqrt{n}})| \leq | -\frac{t^2}{2n} + \frac{1+t^2}{6n^{\frac{3}{2}}} | \leq \exp \left| -\frac{t^2}{2n} + \frac{1+t^2}{6n^{\frac{3}{2}}} \right|, t^2 \leq 2n \text{ 时成立.} \\ &\quad \leq \exp \left| -\frac{t^2}{18n} \right|, \frac{1+t^2}{\sqrt{n}} \leq \frac{4}{3} \text{ 时成立.} \\ &\quad \therefore \sup_x |\mathbb{P}(\frac{S_n}{\sqrt{n}} \leq x) - \Phi(x)| \leq C \cdot \int_{-L}^L n \cdot \left( \exp \left| -\frac{t^2}{18n} \right| \right)^{n-1} \left( \frac{1+t^2}{6n^{\frac{3}{2}}} + \frac{t^4}{8n^2} \right) dt + \frac{C}{L} \\ &\quad \bar{W}_L = \frac{3\sqrt{n}}{2\pi} \leq \frac{C\rho}{\sqrt{n}} + \frac{C}{n} \leq \frac{C\rho}{\sqrt{n}}. \end{aligned}$$

Thm (Berry-Esseen, II).  $X_n$  独立,  $\mathbb{E}X_n = 0, \mathbb{E}X_n^2 = \sigma_n^2 < \infty, \beta_n = \sum_{k=1}^n \Omega_k^2, \forall i$

$$\sup_x |\mathbb{P}(\frac{S_n}{\sqrt{\beta_n}} \leq x) - \Phi(x)| \leq A \cdot \frac{\sqrt{n} \mathbb{E}|X_n|^3}{\beta_n^{\frac{3}{2}}}$$

$$\text{pf: } \frac{1}{\sqrt{\beta_n}} \int_{-\infty}^x (t - \frac{X_k}{\beta_n}) dt = \prod_{k=1}^n \omega_k, \omega_k \triangleq E e^{\frac{X_k}{\sqrt{\beta_n}}}$$

$$\psi(t) = e^{-\frac{t^2}{2}} \prod_{k=1}^n \beta_k, \beta_k \triangleq E e^{-\frac{X_k^2}{2\beta_n}}$$

$$\therefore |\int_{-\infty}^x (t - \frac{X_k}{\beta_n}) dt - \psi(t)| \leq \sum_{k=1}^n \left( \prod_{m=1}^k \frac{1}{\omega_m} \prod_{m=k+1}^n \beta_m \right) \cdot (|\omega_k - \beta_k|)$$

$$\text{lem 2 } L_n \triangleq \beta_n^{\frac{3}{2}} \sum_{k=1}^n \Omega_k^2, \text{ 若 } |t| \leq \frac{1}{4L_n}, |L_n| \leq \frac{1}{8}, \text{ 则 } \forall k, \text{ 有 } \prod_{m=1}^k \frac{1}{\omega_m} \prod_{m=k+1}^n \beta_m \leq e^{-\frac{t^2}{4}}$$

$$\text{由 lem 2, } |\int_{-\infty}^x (t - \frac{X_k}{\beta_n}) dt - \psi(t)| \leq e^{-\frac{t^2}{4}} L_n \cdot \left( \frac{t^2}{6} + \frac{t^4}{8} \right), |t| \leq \frac{1}{4L_n}, |L_n| \leq \frac{1}{8}, |\omega_k - \beta_k| \leq \frac{E|X_k|^3}{\beta_n^{\frac{3}{2}}} \left( \frac{|t|^3}{6} + \frac{|t|^5}{8} \right).$$

下设  $L_n \leq \frac{1}{8}$ , 否则令  $A = 8$  则得证.

$$\text{令 } L = \frac{1}{4L_n}, \text{ 由 B-E 的 lem, } \sup_x |\mathbb{P}(\frac{S_n}{\sqrt{\beta_n}} \leq x) - \Phi(x)| \leq \frac{1}{\pi} \int_{-L}^L e^{-\frac{t^2}{4}} L_n \left( \frac{|t|^3}{6} + \frac{|t|^5}{8} \right) dt + \frac{C}{L}$$

$$\leq C L_n.$$

伊顿 Elon #9

Date .

$$\begin{aligned} \text{Pf of lem } & \frac{\sum_{n=1}^{\infty} 2^n}{x^{2^n}} E[X]^n I(|X|>x) = x^{2^n} \sum_{n=1}^{\infty} E[X]^n I\left(\frac{|X|}{x} \in (2^n, 2^{n+1}] \right) \\ & \leq \sum_{n=1}^{\infty} E[X]^n I\left(\frac{|X|}{x} \in (2^n, 2^{n+1}] \right) 2^{n(P-2)} \\ & = \sum_{n=1}^{\infty} 2^{(P-2)n} [L(2^{n+1}x) - L(2^n x)] \end{aligned}$$

$$\begin{aligned} (\text{Hire}(1, 2^{P-1}), \exists x_0, \text{s.t. } \forall x > x_0, \frac{L(x)}{L(x)} \leq r) \\ \leq \sum_{n=1}^{\infty} 2^{(P-2)n} (r-1) r^n L(x) \end{aligned}$$

$$= \frac{(r-1)L(x)}{1-r^{P-2}}$$

令  $r \rightarrow 1$ , 则  $\forall p \in [0, 1]$ ,  $\lim_{x \rightarrow \infty} \frac{x^{2^n} E[X]^n I(|X|>x)}{L(x)} = 0$

step 2 若  $E[X^2] < \infty$ , 则  $E[(x-m)^2] < \infty$ ,  $L_m(x) / E(x-m)^2$  为缓变化函数

若  $E[X^2] = \infty$ ,  $L_m(x) = E[X^2] I(|X-m| \leq x) + m E[(m-2)X] I(|X-m| \leq x)$   
在  $L(x \pm m)$  之间。有界

$$\begin{aligned} (\text{写好一点: } \frac{L(x \pm m)}{L(x)} \rightarrow 1) \quad \frac{m E[(m-2)X] I(|X-m| \leq x)}{L(x)} \rightarrow 0 \\ \therefore \frac{L_m(x)}{L(x)} \rightarrow 1 \quad \therefore \frac{L_m(x)}{L_m(x)} \rightarrow 1. \quad \# \end{aligned}$$

Pf of Thm (弱退化3): (2)  $\Rightarrow$  (3) 由 Lem.

$$(3) \Rightarrow (2), \frac{L(2x)}{L(x)} - 1 = \frac{E[X^2] I(x < |X| \leq 2x)}{L(x)} \stackrel{(2) \Rightarrow P(|X|>x) \rightarrow 0}{\leq} \frac{(2x)^2 P(|X|>x)}{L(x)} \rightarrow 0.$$

(2)  $\Rightarrow$  (1): 取  $m_k = EX \stackrel{def}{=} 0$

$$q_n = 1/n \sum_{k=1}^n P(X_k > x) = n L(x) \geq x^2, \text{ 则 } q_n \nearrow \infty \quad \text{(*)}$$

$$\text{且 } X_{nk} = \frac{X_k}{a_n}, \text{ 且 } \sum_{k=1}^n X_{nk} \stackrel{d}{\rightarrow} N(0, 1).$$

$$\Leftrightarrow (a) \forall \varepsilon > 0, \sum_{k=1}^n P(|X_{nk}| \geq \varepsilon) \rightarrow 0$$

$$(b) \sum_{k=1}^n E[X_{nk} I(|X_{nk}| \leq 1)] \rightarrow 0$$

$$(c) \sum_{k=1}^n \text{Var}(X_{nk} I(|X_{nk}| \leq 1)) \rightarrow 0.$$

$$(a), \sum_{k=1}^n P(|X_{nk}| > \varepsilon) = n P(|X_k| > \varepsilon a_n) \sim \frac{a_n^2}{n} P(|X| > \varepsilon a_n) \xrightarrow{n \rightarrow \infty} 0$$

$$(b), |LHS| = \frac{n}{a_n} |EX \cdot I(|X| \leq a_n)|$$

$$= \frac{n}{a_n} |EX \cdot I(|X| > a_n)|$$

$$\leq \frac{n}{a_n} E|X| \cdot I(|X| > a_n) \sim \frac{a_n}{L(a_n)} E|X| I(|X| > a_n) \xrightarrow{n \rightarrow \infty} 0$$

$$(c), |LHS| = \frac{n}{a_n} E[X^2] I(|X| \leq a_n) - \frac{n}{a_n^2} (E[X I(|X| \leq a_n)])^2$$

$$= \frac{nL(a_n)}{a_n^2} - \frac{n}{a_n^2} (E[X I(|X| > a_n)])^2$$

$$\rightarrow 1 \quad \#$$

$$B-E: \sup_{x \in \mathbb{R}} |P\left(\frac{S_n}{\sqrt{n}} \leq x\right) - \bar{F}(x)| \leq C \frac{P}{\sqrt{n}},$$

Rmk. 1. B-E 中  $E[X^3] < \infty$  可改为  $E[X^{2.5}] < \infty$ , 收敛速度  $\frac{1}{n^{1/2}}$

$$\sup_{x \in \mathbb{R}} |P\left(\frac{S_n}{\sqrt{n}} \leq x\right) - \bar{F}(x)| \leq C \cdot \left( \frac{1}{\sqrt{n}} E[X^3] I(|X| \leq \sqrt{n}) + \frac{1}{\sqrt{n}} E[X^2] I(|X| > \sqrt{n}) \right)$$

可改进.

2.  $\sup_{x \in \mathbb{R}} \frac{1}{\sqrt{n}} \text{ in B-E 不可改进.}$

Cex.  $X_i \stackrel{iid}{\sim} \text{sym Bernoulli}$ :  $P(S_{2n}=0) = \binom{2n}{n} \left(\frac{1}{2}\right)^{2n} = \frac{1}{\sqrt{\pi n}} (1+o(1))$

$$P(S_{2n} \neq 0) = \frac{1}{2} (1 - \frac{1}{\sqrt{\pi n}} (1+o(1)))$$

$$\therefore P(S_{2n} \neq 0) - \bar{F}(0) = \frac{1}{2\sqrt{\pi n}} (1+o(1))$$

3. Edgeworth 展开:  $|P\left(\frac{S_n}{\sqrt{n}} \leq x\right) - \bar{F}(x)| - \sum \dots | \leq \frac{1}{n^2}$

Thm. 设  $X_1, \dots$  独立,  $EX_i = 0$ ,  $E[X_i]^3 < \infty$ ,  $V_{3r}(X_i) = o_r^2$ ,  $B_n = \sum_{i=1}^n \sigma_i^2$ ,  $L_n = B_n^{-\frac{1}{2}} \sum_{i=1}^n E[X_i]^3$

$$\text{则 } |P\left(\frac{S_n}{\sqrt{n}} \leq x\right) - \bar{F}(x)| \leq \frac{A L_n}{\sqrt{n}}$$

$$\left| \frac{P\left(\frac{S_n}{\sqrt{n}} > x\right)}{1 - \bar{F}(x)} - 1 \right| \leq \frac{C P\left(\frac{S_n}{\sqrt{n}} > x\right)}{1 - \bar{F}(x)} \quad \text{当 } 0 \leq x \leq (-\varepsilon) \sqrt{\log n} \text{ 时, } \leq \frac{C P\left(\frac{S_n}{\sqrt{n}} > x\right)}{e^{-\frac{1}{2}(1-\varepsilon)\sqrt{\log n}/(-\varepsilon)\sqrt{\log n}}} \leq \frac{\frac{1}{n^2} \sqrt{\log n}}{n^{\frac{1}{2}-\frac{1}{2}(1-\varepsilon)}} \rightarrow 0.$$

从而,  $\frac{P\left(\frac{S_n}{\sqrt{n}} > x\right)}{1 - \bar{F}(x)} \rightarrow 1$  在  $[0, 1 - \varepsilon \sqrt{\log n}]$  改成立.

\*考高“大偏差”:  $x_n \rightarrow \infty$ ,  $P\left(\frac{S_n}{\sqrt{n}} > x_n\right)$  的值.

Thm (Cramer 大偏差) 若  $0 \leq x = o(\sqrt{n})$ ,  $EX = 0$ ,  $EX^2 = 1$ , 则  $F_n \stackrel{d}{\rightarrow} \frac{S_n}{\sqrt{n}}$  d.f.

$$\frac{P(S_n > x)}{1 - \bar{F}(x)} = \exp \left( \frac{x^2}{n} \lambda \left( \frac{x}{\sqrt{n}} \right) \right) [1 + o\left( \frac{x}{\sqrt{n}} \right)], \text{ 其中 } \lambda(t) = \sum_{k=0}^{\infty} C_k t^k \text{ 为 Cramer 系数}$$

Def. (缓变化函数)  $L$  在  $[0, \infty)$  上正值可测. 若  $\forall c > 0$ ,  $\frac{L(cx)}{L(x)} \rightarrow 1$  ( $x \rightarrow \infty$ ), 则称  $L$  是( $x$ 处的)缓变化函数. e.g. 常值,  $\log x$ ,  $\log \log x$ ,  $\ln \ln x$ ,  $\gamma$

Thm (弱退化3) 设  $X_i \stackrel{iid}{\sim}$  非退化, 则  $\sum_{k=1}^n \frac{(X_k - m_k)}{a_n} \xrightarrow{d} N(0, 1)$  \* 称  $L$  属于正态吸引场.

(1)  $\exists a_n, m_n, \text{s.t. } \sum_{k=1}^n \frac{(X_k - m_k)}{a_n} \xrightarrow{d} N(0, 1)$

$\Leftrightarrow (2) L(x) = E[X^2] I(|X| \leq x)$  ( $x$ 处的)缓变化函数.

$$\Leftrightarrow (3) \lim_{x \rightarrow \infty} \frac{x^2 P(|X| > x)}{E[X^2] I(|X| \leq x)} = 0$$

Thm 设  $X$  非退化,  $L(x)$  缓变化函数, 则  $\forall m \in \mathbb{R}$ ,  $L_m(x) = E[(x-m)^2] I(|X-m| \leq x)$  也缓变化

且  $\forall p \in [0, 2]$ ,  $\lim_{x \rightarrow \infty} \frac{x^p E[X^p I(|X| > x)]}{L(x)} = 0$

E.g.  $X_i \stackrel{iid}{\sim} P(|X| > x) = \frac{1}{x^2}$ ,  $(X > 1)$ , 对称

且  $EX^2 = \infty$ ,  $EX^2 I(|X| \leq x) = 2 \log x \sim o(\ln x) \sim a_n^2 \Rightarrow a_n \sim \sqrt{\ln x}$

$$\therefore \frac{S_n}{\sqrt{\ln x}} \xrightarrow{d} N(0, 1). \text{ 也可直接 Ch.f. 证之.}$$

Def. of Thm (Ber-Poi收敛):  $Ee^{itS_n} = \prod_{k=1}^n Ee^{itX_{nk}} = \prod_{k=1}^n (1 + p_{nk}(e^{it} - 1))$

 $\left| \prod_{k=1}^n (1 + p_{nk}(e^{it} - 1)) - \prod_{k=1}^n e^{p_{nk}(e^{it} - 1)} \right| \leq \sum_{k=1}^n |1 + p_{nk}(e^{it} - 1) - e^{p_{nk}(e^{it} - 1)}|$ 
 $\Leftrightarrow |e^{p_{nk}(e^{it} - 1)}| = e^{p_{nk}(e^{it} - 1)} \leq 1, |1 + p_{nk}(e^{it} - 1)| \leq 1 \Rightarrow |1 + p_{nk}(e^{it} - 1)| \approx 1$ 

由  $|e^b - 1 - b| \leq |b|^2$  ( $|b| < 1$ ) 知:  $\leq \sum_{k=1}^n p_{nk}^2 |e^{it} - 1|^2$  (当  $p$  很大, 由条件)

 $\therefore \lim_{n \rightarrow \infty} e^{itS_n} = \lim_{n \rightarrow \infty} e^{2p_{nk}(e^{it} - 1)} = e^{\lambda(e^{it} - 1)} \xrightarrow{|b| < 1} \text{满足}$

Def. of Lem (全变差距离): 设  $B = \{j : \mu_j > \nu_j\}$ ,  $\forall A \subset B$ ,

$$\begin{aligned} |\mu(A) - \nu(A)| &= \mu(AB) - \nu(AB) + \sum_{i \in A \setminus B} (\mu_i - \nu_i) \\ &\leq \mu(AB) - \nu(AB) \\ &\leq \mu(AB) - \nu(AB) + \sum_{i \in A \setminus B} (\mu_i - \nu_i) = \mu(B) - \nu(B). \end{aligned}$$

类似,  $\nu(A) - \mu(A) \leq \mu(B) - \nu(B)$ .  $\Rightarrow |\mu(A) - \nu(A)| \leq \mu(B) - \nu(B), \forall A \subset B$ .

 $\therefore ||\mu - \nu|| = \mu(B) - \nu(B) = \frac{1}{2} (\nu(B^c) - \mu(B^c) + \mu(B) - \nu(B)) = \frac{1}{2} \sum |\mu_i - \nu_i|$

Def. of Lem 2: LHS  $= \frac{1}{2} \sum_x |\mu_1 * \mu_2(x) - \nu_1 * \nu_2(x)|$

 $= \frac{1}{2} \sum_y |\sum_x (\mu_1(y)\mu_2(x-y) - \nu_1(y)\nu_2(x-y))|$ 
 $\leq \frac{1}{2} \sum_y |\mu_1(y)\mu_2(x-y) - \nu_1(y)\nu_2(x-y)|$ 
 $= \frac{1}{2} \sum_y |\mu_1(y)\mu_2(x) - \nu_1(y)\nu_2(x)|$ 
 $\leq \frac{1}{2} \sum_y |\mu_1(y)(\mu_2(x) - \nu_2(x))| + |\mu_1(y) - \nu_1(y)|\nu_2(x)|$ 
 $= \frac{1}{2} \sum_y |\mu_2(x) - \nu_2(x)| + \frac{1}{2} \sum_y |\mu_1(y) - \nu_1(y)| \stackrel{\text{由 Lem 1}}{\leq} \text{RHS.}$

Def. of Lem 3:  $2||\mu - \nu|| \stackrel{\text{def}}{=} |\mu(0) - \nu(0)| + |\mu(1) - \nu(1)| + \sum_{i \geq 2} \nu(i)$

 $= |1-p - e^{-p}| + |p - pe^{-p}| + |1-e^{-p} - pe^{-p}|$ 
 $= e^{-p} - (1-p) + p - pe^{-p} + 1 - e^{-p}(1+p)$ 
 $= 2p(1-e^{-p}) \leq 2p^2.$

Ber-Poi 收敛 Thm:  $X_{nk} \sim \text{Unif}_k, S_n \sim \mu_n, \nu_n \sim \text{Poi}(p_{nk}), Poi(\sum_k p_{nk}), \text{Poi}(\lambda)$

设  $\mu_n = \mu_1 * \dots * \mu_{n_k}, \nu_n = \nu_1 * \dots * \nu_{n_k}$

 $\therefore \|\mu_n - \nu_n\| \stackrel{\text{def}}{=} \sum_{k=1}^{n_k} \|\mu_{nk} - \nu_{nk}\| \stackrel{\text{def}}{=} \sum_{k=1}^{n_k} p_{nk}^2 \leq \max_{1 \leq k \leq n_k} p_{nk} \cdot \sum_{k=1}^{n_k} p_{nk} \rightarrow 0$

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- Poisson 收敛:

若  $B(n, p)$ . 若  $p$  fixed,  $\frac{B(n, p) - np}{np(p)} \xrightarrow{d} N(0, 1)$ .

若  $n p \rightarrow \lambda \in (0, \infty)$ ,  $P(S_n = k) = \binom{n}{k} p_n^k (1-p_n)^{n-k} \sim \frac{1}{k!} n^k p_n^k (1-p_n)^{n-k} \xrightarrow{k!} \frac{\lambda^k}{k!} e^{-\lambda}$

Thm. 设  $X_{nk} \sim \text{Unif}_k$  独立.  $P(X_{nk}=1) = p_{nk} = 1 - P(X_{nk}=0)$ .

若 (1)  $\sum_{k=1}^{n_k} p_{nk} \rightarrow \lambda \in (0, \infty)$

(2)  $\max_{1 \leq k \leq n_k} p_{nk} \rightarrow 0$ ,

$\Rightarrow S_n \xrightarrow{d} \text{Poi}(\lambda)$

Def. 全变差距离: prob. measure  $\mu, \nu$  on measurable space  $(\Omega, \mathcal{F})$ ,

 $||\mu - \nu|| \stackrel{\text{(全变差距离)}}{=} \max_{A \in \mathcal{F}} |\mu(A) - \nu(A)|$ 

Lem 1:  $\mu, \nu$  空间  $\Omega$ ,  $||\mu - \nu|| = \frac{1}{2} \sum_i |\mu_i - \nu_i|$

Lem 2:  $||\mu_1 * \mu_2 - \nu_1 * \nu_2|| \leq ||\mu_1 - \nu_1|| + ||\mu_2 - \nu_2||$ .

Lem 3: 设  $\mu(\cdot) = p = 1 - \mu(\emptyset)$ ,  $\nu \sim \text{Poi}(p)$ , 则  $||\mu - \nu|| \leq p^2$

Thm. (Poisson 收敛定理-必要) 独立非负整数值  $\nu$ . 则  $\sum_{k=1}^{n_k} P(X_{nk} > 1) \rightarrow 0$  足以推得  $\sum_{k=1}^{n_k} P(X_{nk} = 1) \rightarrow \lambda$ .

证: (1) 不妨  $X'_n = X_{nk} I(X_{nk} \neq 0)$ , 则由 Ber-Poi 收敛 + 充分条件,  $S'_n \xrightarrow{d} \text{Poi}(\lambda)$ .

$\therefore P(S_n \neq S'_n) \leq \sum P(X_{nk} > 1) \rightarrow 0 \Rightarrow S_n \xrightarrow{d} \text{Poi}(\lambda)$ .

Ex  $X_i \text{ i.i.d. Cauchy}$ ,  $f(x) = \frac{\alpha}{\pi(x+\alpha^2)}, \alpha > 0$ ,  $\mathbb{P}[\frac{S_n}{n} \stackrel{d}{=} X]$

Def:  $\int_{-\infty}^{\infty} \frac{\alpha}{2} e^{-\alpha|x|} dt = \int_0^{\infty} \alpha e^{-\alpha x} e^{i\alpha t} dt = \frac{\alpha}{2} (\frac{1}{q-i} + \frac{1}{q+i}) = \frac{\alpha^2}{1+t^2}$

反演:  $\frac{\alpha}{2} e^{-\alpha|x|} = \frac{1}{2\pi} \int_{-\infty}^{\infty} \frac{\alpha^2}{1+t^2} e^{i\alpha t} dt$

$\therefore f_X(t) = \int_{-\infty}^{\infty} \frac{\alpha}{\pi(1+t^2)} e^{i\alpha t} dt = e^{-\alpha|t|}$

$\therefore f_{\bar{X}}(t) = Ee^{it\bar{X}} = (f_X(\frac{t}{n}))^n = e^{-\alpha|t|} = f_X(t)$ .

Rmk: 若  $f(t) = e^{-\alpha|t|}$ , 则  $f_{\bar{X}}(t) = (f(\frac{t}{n}))^n = f(t)$ ,  $\mathbb{P}[\frac{S_n}{n} \stackrel{d}{=} X]$ .

Ex  $X_i \text{ i.i.d. 对称}$ ,  $P(|X|>x) = x^{-\alpha}, x \geq 1, 0 < \alpha < 2$ .  $\mathbb{P}[\exists \text{ r.v. } Y, \text{ s.t. } \frac{S_n}{n^{\alpha}} \stackrel{d}{=} Y]$

Def:  $1-f_X(t) = \int_{|X|>t} (1-e^{itx}) \frac{\alpha}{2|x|^{1+\alpha}} dx$

$= \alpha \int_1^\infty (1-\cos(tx)) \frac{1}{x^{1+\alpha}} dx$

$\stackrel{y=tx}{=} \alpha \int_1^\infty \frac{1-\cos y}{y^{1+\alpha}} dy \underset{t \rightarrow 0}{\sim} C_\alpha t^\alpha \quad (\int_0^\infty \frac{1-\cos y}{y^{1+\alpha}} dy < \infty)$

$Ee^{it\frac{S_n}{n^{\alpha}}} = (f(\frac{t}{n^{\alpha}}))^n \sim [1 - C_\alpha t^\alpha]^n \sim e^{-C_\alpha t^\alpha}$

Rmk: 若  $f(t) = e^{-C_\alpha t^\alpha}$ , 则  $f_{\bar{X}}(t) = (f(\frac{t}{n^{\alpha}}))^n = f(t)$ ,  $\mathbb{P}[\frac{S_n}{n^{\alpha}} \stackrel{d}{=} X]$ .

由 Thm (弱大数律): ( $\Leftarrow$ ) 全  $X_i \sim F$ , 由  $F$  稳定,  $\exists \{a_n\}, \{b_n\}$ , s.t.  $(f(t))^n = e^{ib_n t} f(a_n + t)$

$\mathbb{P}[\frac{S_n}{n} \stackrel{d}{=} b_n + a_n X] \Rightarrow \frac{S_n - b_n}{a_n} \stackrel{d}{=} X \Rightarrow \frac{S_n - b_n}{a_n} \stackrel{d}{=} F$ .

( $\Rightarrow$ ) 若  $F$  退化, 则  $F$  稳定

若  $F$  非退化, 记  $Z_n = \frac{S_n - b_n}{a_n}$ ,  $\mathbb{P}[Z_n \stackrel{d}{=} F]$

$Z_{kn} = \frac{(X_1 + \dots + X_n + (X_{kn+1} + \dots + X_{kn+k}) + \dots + (X_{kn+n} + \dots + X_{kn+n+k}) - b_{kn})}{a_{kn}}$

$\stackrel{a_{kn}}{=} \frac{S_n + t S_{kn}^k - b_{kn}}{a_{kn}}$

$\therefore \frac{a_{kn} Z_{kn} + b_{kn} + b_{kn}}{a_{kn}} = \frac{S_n - b_n}{a_n} + \dots + \frac{S_{kn} - b_{kn}}{a_{kn}} \stackrel{d}{=} F^{*k}(x)$

( $\Leftarrow$ )  $Ee^{it(Z_{kn} + b_{kn})} = (Ee^{itZ_n})^k = (f_{Z_n})^k = (f_F)^k$

$\therefore Z_{kn} \stackrel{d}{=} F$

由 律与型,  $\exists a_k, b_k$ , s.t.  $F^{*k}(x) = F(a_k x + b_k)$   $\therefore F$  为稳定的  $\#$

Ex Lévy 过程 (独立平稳增量过程)

$\overset{\text{独立}}{\longleftarrow} \overset{\text{平稳}}{\longleftarrow} \overset{\text{增量}}{\longleftarrow} Y_0, X_1 = \frac{1}{i\pi} (Y_{\frac{1}{i\pi}} - Y_{\frac{-1}{i\pi}}), \text{ 其中 } \{Y_{\frac{k}{i\pi}} - Y_{\frac{-k}{i\pi}}\}_{k \in \mathbb{Z}} \text{ i.i.d. 从而 } F \text{ 为无宽可分分布}$

1. 平稳  
2. Poisson 分布  $f(t) = e^{\lambda(e^{it}-1)} = (e^{\lambda(e^{it}-1)})^n$

3. Gamma 分布

Date.

• 稳定分布:  $\frac{S_n - b_n}{a_n} \stackrel{d}{\rightarrow} ?$  若  $EX \geq 0$ , 则 CLT  $\Rightarrow$  正态分布

Def:  $f(x)$  的 ch. f.  $f(t)$  是稳定的, 若  $\forall k \in \mathbb{N}, \exists c_k > 0, \forall t, \text{ s.t. } (f(t))^k = e^{itkt} f(ct)$

Thm:  $\exists \{X_i\}_{i \in \mathbb{N}}, \{f_{b_n}\}, \text{ s.t. } \frac{S_n - b_n}{a_n} \stackrel{d}{\rightarrow} F \Leftrightarrow F$  是稳定的.

Thm: 稳定分布的 ch. f.  $f(t) = \exp[fitc - bt|t|^{\alpha} (1 + iK \operatorname{sgn}(t) w_{\alpha}(t))]$  称具有指数  $\alpha$  的稳定分布

其中  $-1 \leq K \leq 1$ , 指数  $\alpha \in [0, 2]$ ,  $w_{\alpha}(t) = \begin{cases} \tan(\frac{\pi\alpha}{2}), & \alpha \neq 1 \\ \frac{2}{\pi} \log|t|, & \alpha = 1 \end{cases}$

Rmk: 1.  $\alpha = 2$ , 则  $N(c, \sigma^2)$ .  
 2.  $\alpha = 1, c = k = 0$ , 则 Cauchy 分布.  
 3.  $p(x) = (2\pi x^2)^{-\frac{1}{2}} e^{-\frac{x^2}{2}}$ ,  $x > 0$  为稳定分布 ( $\alpha = \frac{1}{2}, k = 1, c = 0, b = 1$ ).  
 4.  $\int f(t) dt < \infty$  时有密度, 但  $1 \sim 3$  是现在可直接写出显式表示的.

Thm:  $X$  的分布  $G$  属于某个具有指数  $\alpha \in (0, 2)$  稳定分布的吸引子  $\Leftrightarrow$

(1)  $P(|X| > x) = x^{-\alpha} L(x)$ ,  $L$  为缓变化函数,  $0 < \alpha < 2$ .

(2)  $\lim_{x \rightarrow \infty} \frac{P(X > x)}{P(|X| > x)} = \theta \in [0, 1]$

进一步, (1) 成立时, 记  $S_n = X_1 + \dots + X_n$  (i.i.d.),  $a_n = \inf\{x : P(|X| > x) \leq \frac{1}{n}\}$ ,  $b_n = nE[X \mid |X| \leq a_n]$  则  $\exists$  指数为  $\alpha$  的稳定分布  $F$ , s.t.  $\frac{S_n - b_n}{a_n} \stackrel{d}{\rightarrow} F$ .  
 $F$  具有  $K = 2\theta - 1$  的 ch. f.

Def: (无穷可分分布)  $F$  的 ch. f.  $f(t)$  满足  $\forall n, \exists f_n(t), \text{ s.t. } f(t) = (f_n(t))^n$ .

Thm: 分布  $F$ , 则  $\exists \{Y_{nk}\}_{k \in \mathbb{N}}$ , s.t. (1)  $\forall n, Y_{nk}, k = 1 \dots n$ , i.i.d. (2)  $Y_{1n} + \dots + Y_{nn} \stackrel{d}{\rightarrow} F$ .  $\Leftrightarrow F$  为无穷可分的.

Ex: ( $\Leftarrow$ ) trivial ( $\Rightarrow$ ) 只证  $n=2$ , 即  $F$  可写成两个 i.i.d. r.v. 之和的分布.

$\therefore S_{2n} = (X_1 + \dots + X_{2n,n}) + (X_{2n+1,n+1} + \dots + X_{2n+2n,2n}) \stackrel{d}{=} Y_n + Y'$

且  $P(Y_n > x) = P(Y_n > x, Y' > x) \leq P(S_{2n} > 2x) \stackrel{d}{\rightarrow} S_{2n}$  胜算  $\Rightarrow Y_n$  胜算

$P^2(Y_n < -x) \leq P(S_{2n} < -2x)$

$\therefore$  by Helly,  $\exists \{f_{nk}\}$ , s.t.  $Y_{nk} \xrightarrow{d} Y$ ,  $Y$  为一个 r.v., 从而  $F = Y + Y'$

对  $\forall n \in \mathbb{N}$  为类似处理  $\Rightarrow F$  为无穷可分的.

Thm: 独立 r.v. 组  $\{X_{nk}\}$  满足无穷小条件, 则  $\sum_{k=1}^n X_{nk}$  的极限分布族与无穷可分分布族相重

Prop.  $f(t)$  i.d.,  $\forall t \in \mathbb{R}, f(t) \neq 0$ .

Thm (Lévy-Khintchine 表示)  $f(t)$  i.d.  $\Leftrightarrow f(t) = \exp\left\{ict - \frac{\sigma^2 t^2}{2} + \int (e^{itx} - 1 - \frac{itx}{\lambda x^2}) \nu(dx)\right\}$ .

其中测度  $\nu$  满足  $\nu(\{0\}) = 0$  且  $\int (1 \wedge x^2) \nu(dx) < \infty$ , 称  $\nu$  为 Lévy 测度.

(注意 poisson 进步  $\int (e^{itx} - 1) \nu(dx) = e^{a(e^{ita}-1)}$ , 这定理可看作分布可拆成 normal + - by poisson 分布)

Lemma 1: ①  $E f'(z) = \frac{1}{\sqrt{\pi}} \int_0^z e^{-\frac{t^2}{2}} f'(t) dt$

$$= \frac{1}{\sqrt{\pi}} \int_0^z f'(t) dt + \int_t^\infty w e^{-\frac{w^2}{2}} dw - \frac{1}{\sqrt{\pi}} \int_{-\infty}^0 f'(t) dt + \int_0^t w e^{-\frac{w^2}{2}} dw$$

$$= \frac{1}{\sqrt{\pi}} \int_0^z w e^{-\frac{w^2}{2}} dw \int_0^w f(t) dt - \frac{1}{\sqrt{\pi}} \int_{-\infty}^0 w e^{-\frac{w^2}{2}} \int_w^0 f'(t) dt$$

$$= E Z f'(z)$$

Lemma 2. i:  $(E^{\frac{w^2}{2}} f(w))' = E^{\frac{w^2}{2}} (f'(w) - w f(w)) \stackrel{\text{def}}{=} E^{\frac{w^2}{2}} (I(w \leq x) - \phi(x))$

 $\text{而 } \int_t^\infty (E^{\frac{w^2}{2}} f(w))' dw = -E^{\frac{w^2}{2}} f(t), \text{ 若 } f \text{ 有界.}$ 
 $\therefore E^{\frac{w^2}{2}} f(w) = - \int_w^\infty E^{\frac{w^2}{2}} (I(w \leq t) - \phi(t)) dt$ 
 $\text{由 } \int_{-\infty}^0 E^{\frac{w^2}{2}} (\phi(x) - I(t \leq x)) dt = \phi(x) \sqrt{\pi} - \int_{-\infty}^x E^{\frac{w^2}{2}} dt = 0 \text{ 得第二个等号.}$

Check 这个解有界:

若  $w < x$ ,  $f(w) = -E^{\frac{w^2}{2}} \int_w^\infty E^{\frac{t^2}{2}} (\phi(t) - 1) dt = E^{\frac{w^2}{2}} (1 - \phi(x)) \sqrt{\pi} \phi(w)$

若  $w > x$ ,  $f(w) = E^{\frac{w^2}{2}} \int_w^\infty E^{\frac{t^2}{2}} \phi(t) dt = E^{\frac{w^2}{2}} \phi(x) \sqrt{\pi} (1 - \phi(w))$

下证  $|f| \leq \sqrt{\frac{\pi}{2}}$

由于  $w > 0$  时,  $1 - \phi(w) \leq \min\left\{\frac{1}{2}, \frac{1}{\sqrt{\pi} w}\right\} e^{-\frac{w^2}{2}}$  [CLAIM].

从而  $w > 0$  时,  $f(w) \leq E^{\frac{w^2}{2}} \sqrt{\pi} (1 - \phi(w)) \leq \sqrt{\frac{\pi}{2}}$

$w \leq x$ ,  $f(w) \leq E^{\frac{w^2}{2}} (1 - \phi(w)) \sqrt{\pi} \leq \sqrt{\frac{\pi}{2}}$

$w < 0$  时类似可证

Lemma 1 ii: ② 由 Stein 方程,  $\forall x$ ,  $f'_x(z) - E f'_x(z) = I(z \leq x) - \phi(x)$  且  $E f'_x(z) = E Z f'_x(z)$

 $\therefore E I(z \leq x) = \phi(x) \Rightarrow z \sim N(0, 1)$

#### References:

△ Fundamentals of Stein's Method, Ross, N. Probability Surveys, 219-293, 2011.

Barbour, Holst and Janson, Poisson Approximation, (1992)

Chen, Goldstein and Shao, Normal Approximation by Stein's Method, (2011).

Date:

• Stein 方法 (证明极限分布, 常参考 ch.f., 矩方法, Stein 方法)

Lemma 1:  $Z \sim N(0, 1)$ , 则  $\forall AC$  函数  $f: \mathbb{R} \rightarrow \mathbb{R}$  且  $E f(Z) < \infty$ , 有  $E f'(Z) = E Z f'(Z)$

反之, 若  $f$  有界连续、逐段可微,  $E|f'(z)| < \infty$  的函数  $f$  都有  $E f'(Z) = E Z f'(Z)$ , 则  $Z \sim N(0, 1)$ .

Lemma 2:  $\forall x \in \mathbb{R}$ , Stein 方程:

$$f'(w) - w f(w) = I(w \leq x) - \phi(x)$$

有唯一有界解.

$$f(w) = E^{\frac{w^2}{2}} \int_w^\infty e^{-\frac{t^2}{2}} (\phi(x) - I(t \leq x)) dt$$

$$= -E^{\frac{w^2}{2}} \int_w^\infty e^{-\frac{t^2}{2}} (\phi(x) - I(t \leq x)) dt$$

[CLAIM]:  $w > 0$ ,  $1 - \phi(w) \leq \min\left\{\frac{1}{2}, \frac{1}{\sqrt{\pi} w}\right\} e^{-\frac{w^2}{2}}$

$$\begin{aligned} \text{if: } \int_w^\infty E^{\frac{t^2}{2}} dt &= \frac{1}{\sqrt{\pi}} \int_w^\infty e^{-\frac{t^2}{2}} e^{-\frac{1}{2}(t-w)^2} e^{-w(t-w)} dt \\ &\leq e^{-\frac{w^2}{2}} \cdot \frac{1}{\sqrt{\pi}} \int_w^\infty e^{-\frac{1}{2}(t-w)^2} dt = \frac{1}{2} e^{-\frac{w^2}{2}} \end{aligned}$$

另一部分:  $\int_w^\infty E^{\frac{w^2}{2}} dt \leq \int_w^\infty \frac{t}{w} E^{\frac{w^2}{2}} dt = \frac{1}{w} e^{-\frac{w^2}{2}}$

Lemma 3: Stein 方程  $f'(w) - w f(w) = h(w) - Nh$ , 其中  $Nh = E h(Z)$ ,  $Z \sim N(0, 1)$ .

有唯一有界解.

$$\begin{aligned} f_h(w) &= E^{\frac{w^2}{2}} \int_w^\infty e^{-\frac{t^2}{2}} (Nh - h(t)) dt \\ &= -E^{\frac{w^2}{2}} \int_w^\infty e^{-\frac{t^2}{2}} (Nh - h(t)) dt \end{aligned}$$

若  $h$  有界,  $\|f_h\| \leq \sqrt{\frac{\pi}{2}} \|h\| - Nh\|$

$\|f'_h\| \leq 2 \|h\| - Nh\|$

若  $h$  AC,  $\|f_h\| \leq 2 \|h'\|$ ,  $\|f'_h\| \leq \sqrt{\frac{\pi}{2}} \|h'\|$ ,  $\|f''_h\| \leq 2 \|h''\|$

Def. BE 距离  $d_E(y, v) = \sup_{h \in \mathcal{H}} |\int_h(y) d\mu(y) - \int_h(v) d\nu(y)|$

e.g. ①  $\mathcal{H} = \{I(\cdot \leq x): x \in \mathbb{R}\}$ , Kolmogorov BE 距离

②  $\mathcal{H} = \{I(\cdot \in A): A \in \mathcal{B}(\mathbb{R})\}$ , 全变差距离

③  $\mathcal{H} = \{h: |h(x) - h(y)| \leq c|x-y|\}$ , Wasserstein BE 距离

$$\begin{aligned} d_W(W, N(\mu, \Sigma)) &= \sup_{h \in \mathcal{H}} |E h(W) - E h(Z)| \\ &\stackrel{\text{Stein}}{=} \sup_{h \in \mathcal{H}} |E(f'_h(W) - W f_h(W))| \\ &\leq \sup_{\substack{|F| \leq 2 \\ \|F\| \leq 2}} |E(f'(W) - W f(W))| \end{aligned}$$

$\text{Prop. } d_K(W, Z) \leq d_W(W, Z)$

④  $d_K(W, Z) \leq \sqrt{2} d_W(W, Z)$ ,

若  $h$  的密度有界  $C$ .

Recall:

$$Z \sim N(0, 1) \Leftrightarrow E f'(z) = E z f(z)$$

$$\text{Stein 方程 } f'(x) - x f(x) = I(x \leq z) - E(z) \Rightarrow f'(x) - x f(x) = h(x) - Eh(z)$$

Date.

我们常考虑 Wasserstein 度量:  $d_W(X, Y) = \sup_{h: |h(x+y)| \leq |x-y|} |Eh(X) - Eh(Y)|$

$$\text{则有 } d_W(W, N(0, 1)) \leq \sup_{\substack{\|f'\|, \|f''\| \leq 2 \\ \|f\| \leq \sqrt{n}}} |Ef'(W) - Ef(W)| \leftarrow \text{这个}$$

考虑  $X_1, \dots, X_n$  独立,  $E X_i = 0$ ,  $E X_i^2 = 1$ , 记  $W \triangleq \frac{\sum X_i}{\sqrt{n}}$

$$\begin{aligned} \text{III} \quad E W f(W) &= \frac{1}{\sqrt{n}} \sum_i E X_i f(W) = \frac{1}{\sqrt{n}} \sum_i E X_i (f(W) - f(W_i)), \quad W_i = \frac{1}{\sqrt{n}} \sum_j X_j \\ &= \frac{1}{\sqrt{n}} \sum_i E X_i (f(W) - f(W_i) - (W - W_i) f'(W_i)) + \frac{1}{\sqrt{n}} \sum_i E X_i^2 f(W) \end{aligned}$$

$$\Rightarrow |Ef'(W) - Ef(W)| \leq \frac{1}{\sqrt{n}} |E \sum_i X_i (f(W) - f(W_i) - (W - W_i) f'(W_i))| + |E f'(W) (1 - \frac{1}{n} \sum_i E X_i^2)| \\ \triangleq |\text{I}| + |\text{II}|.$$

$$\text{其中 } |\text{I}| \leq \frac{1}{\sqrt{n}} |E \sum_i X_i \cdot \frac{(W-W_i)^2}{2} f''(W_i)|, \quad W_i \text{ 在 } W \text{ 和 } W_i \text{ 之间}$$

$$\leq \frac{1}{2n^{\frac{3}{2}}} E \sum_i |X_i|^3 \|f''\| = \frac{1}{n^{\frac{3}{2}}} \sum_i E |X_i|^3$$

$$|\text{II}| \leq \|f'\| \cdot |E|1 - \frac{1}{n} \sum_i X_i^2| \leq \sqrt{\frac{2}{\pi}} \cdot \sqrt{E(1 - \frac{1}{n} \sum_i X_i^2)^2} = \sqrt{\frac{2}{\pi}} \cdot \frac{1}{n} \sqrt{\text{Var}(\sum_i X_i^2)} \leq \sqrt{\frac{2}{\pi}} \cdot \frac{1}{n} \cdot \sqrt{\sum_i E X_i^4}$$

$$\Rightarrow \boxed{d_W(W, N(0, 1)) \leq \frac{1}{n^{\frac{3}{2}}} \sum_i E |X_i|^3 + \sqrt{\frac{2}{\pi}} \cdot \frac{1}{n} \sqrt{\sum_i E X_i^4}}$$

$$\text{Mtd 2: } E W f(W) \triangleq \sum_i E \xi_i (f(W) - f(W_i)), \quad W_i = \sum_j \xi_j$$

$$= \sum_i E (\xi_i \int_0^{\xi_i} f'(W_i + t) dt)$$

$$= \sum_i E \xi_i \int_{-\infty}^{+\infty} (I_{(0 \leq t \leq \xi_i)} - I_{(0 > t \geq \xi_i)}) f'(W_i + t) dt$$

$$= \sum_i \int_{-\infty}^{+\infty} (E f'(W_i + t)) K_i(t) dt, \quad \text{其中 } K_i(t) = E \xi_i (I_{(0 \leq t \leq \xi_i)} - I_{(0 > t \geq \xi_i)})$$

$$\because \sum_i \int_{-\infty}^{+\infty} K_i(t) dt = \sum_i E \xi_i^2 = 1$$

$$\therefore |E(f(W) - W f(W))| = \left| \sum_{i=1}^n \int_{-\infty}^{+\infty} E f'(W) - f'(W_i + t) \cdot K_i(t) dt \right|$$

$$\leq \frac{1}{\|f'\| \leq 2, \|f''\| \leq 2} 2 \sum_{i=1}^n \int_{-\infty}^{+\infty} (|E \xi_i| + |t|) K_i(t) dt$$

$$\stackrel{\text{代入 } K_i}{=} 2 \sum_{i=1}^n \left( E \xi_i \int_{-\infty}^{+\infty} |t| (I_{(0 \leq t \leq \xi_i)} - I_{(0 > t \geq \xi_i)}) dt + E(\xi_i |E \xi_i|^2) \right)$$

$$\leq 2 \sum_{i=1}^n \left( \frac{1}{2} E(\xi_i)^3 + E(\xi_i)^3 \right) = 3 \sum E(\xi_i)^3 = \frac{3}{n^{\frac{3}{2}}} \sum E |X_i|^3$$

$$\rightarrow \text{Mtd 3: } \because |f'(W) - f'(W_i + t)| \leq \min\{2, 2(|t| + |\xi_i|)\}$$

$$\leq 2(|t| + |\xi_i|)$$

$$\therefore |E f'(W) - W f(W)| \leq 2 \sum_{i=1}^n E \int_{-\infty}^{+\infty} (|t| + |\xi_i|) K_i(t) dt$$

$$= 2 \sum_i (E \xi_i^2 I(|\xi_i| > 1) - \frac{1}{2} E|\xi_i| I(|\xi_i| > 1) + \frac{1}{2} E|\xi_i|^3 I(|\xi_i| > 1))$$

$$\leq \sum_i (4 E \xi_i^2 I(|\xi_i| > 1) + 3 E |\xi_i|^3 I(|\xi_i| > 1) + E \xi_i^2 E(|\xi_i| > 1))$$

$$\left( \because E \xi_i^2 E(|\xi_i| > 1) \leq E \xi_i^2 I(|\xi_i| > 1) + E \xi_i^3 I(|\xi_i| > 1) \right) \text{ Eton}$$

$$d_W(W, N(0, 1)) \leq \frac{4}{n} \sum_{i=1}^n E X_i^2 I(|X_i| > \sqrt{n}) + \frac{3}{n^{\frac{3}{2}}} \sum_i E |X_i|^3 I(|X_i| > \sqrt{n}) \quad 15$$

Date: \_\_\_\_\_

Def.  $\sup_{n \in \mathbb{N}} |P\left(\frac{S_n}{\sqrt{n}} \leq x\right) - \bar{F}(x)| \leq \frac{6.5 E|X|^3}{\sqrt{n}}$ . 记  $\xi_i = \frac{X_i}{\sqrt{n}}$ ,  $W_k = \sum_{i=1}^k \xi_i$

$$\begin{aligned} \text{def } E W_n f(W_n) &= \sum E \xi_i f(W_{i-1}) \\ &= n E \xi_n f(W_{n-1} + \xi_n) \\ &= n E \xi_n (f(W_{n-1} + \xi_n) - f(W_{n-1})) \\ &= n E \xi_n \int_0^{\xi_n} f'(W_{n-1} + t) dt \\ &= E \int_{-\infty}^{+\infty} f'(W_{n-1} + t) K(t) dt, \text{ 其中 } K(t) = n E \xi_n (I_{(0 \leq t < \xi_n)} - I_{(\xi_n \leq t < 0)}) \end{aligned}$$

$\therefore K(t) \geq 0, K(t=0)=0, \int_{-\infty}^t K(t) dt = 1 \Rightarrow \int_{-\infty}^t K(t) dt = \frac{E|X|^3}{2\sqrt{n}} \quad (\star)$

$\therefore E(f(W_n) - W_n f(W_n)) = E \int_{-\infty}^{+\infty} (f(W_{n-1} + \xi_n) - f(W_{n-1})) K(t) dt$

Prop 1.  $E g_x(w) = \begin{cases} -\frac{1}{2}(b-a) - X, & w \leq a-X \\ w - \frac{1}{2}(b+a), & a-X \leq w \leq b+a \\ \frac{1}{2}(b-a) + X, & w \geq b+X \end{cases}$

Prop 2. 若  $g_x \in AC$  且  $g'_x(w) = I(a-X \leq w \leq b+X)$

将  $g_{E|X|^3/\sqrt{n}}$  代入  $E W_n f(W_n) = E \int_{-\infty}^{+\infty} f'(W_{n-1} + t) K(t) dt$  得

$$\begin{aligned} E W_n g_{E|X|^3/\sqrt{n}}(W_n) &= E \int_{(a-\frac{E|X|^3}{\sqrt{n}}) \leq W_{n-1} + t \leq b + \frac{E|X|^3}{\sqrt{n}}} K(t) dt \\ &\geq E I(a \leq W_{n-1} \leq b) \int I(|t| \leq \frac{E|X|^3}{\sqrt{n}}) K(t) dt \\ &\geq P(a \leq W_{n-1} \leq b) \cdot \frac{1}{2} \end{aligned}$$

且  $LHS \leq (\frac{1}{2}(b-a) + \frac{E|X|^3}{\sqrt{n}}) E|W_n| \leq \frac{1}{2}(b-a) + \frac{E|X|^3}{\sqrt{n}}$ .

对于  $f_g$ ,  $|E f_g(W_n) - W_n f_g(W_n)| = |E \int (W_{n-1} + \xi_n) f_g(W_{n-1} + \xi_n) - (W_{n-1} + t) f_g(W_{n-1} + t)|$

$$\begin{aligned} &\stackrel{(\star)}{\leq} \int E \left( W_{n-1} + \frac{E|X|^3}{\sqrt{n}} \right) (E |\xi_n| + |t|) K(t) dt \\ &\quad + \int P(|\xi_n| + |t| \leq W_{n-1} \leq |\xi_n| + |t|) K(t) dt \end{aligned}$$

Prop 3.  $|P\left(\frac{S_n}{\sqrt{n}} \leq x\right) - \bar{F}(x)| \stackrel{E|X|^3 \leq 1}{\leq} \frac{1}{2} \int E |\xi_n| + |t| K(t) dt$

$$\begin{aligned} &\stackrel{(\star)}{\leq} \frac{1}{2} \int (|\xi_n| + |t|) + \frac{2E|X|^3}{\sqrt{n}} K(t) dt \\ &\leq 3 \int (|\xi_n| + |t|) K(t) dt + \frac{2E|X|^3}{\sqrt{n}} \\ &\stackrel{(\star)}{\leq} \frac{9}{2\sqrt{n}} E|X|^3 + \frac{2E|X|^3}{\sqrt{n}} = \frac{6.5 E|X|^3}{\sqrt{n}} \quad (\star) \end{aligned}$$

Date: \_\_\_\_\_

(Berry-Esseen)  
(Kolmogorov-Smirnov)

Thm  $X_1, \dots, X_n$  iid.  $EX=0, EX^2=1, \forall i \sup_{x \in \mathbb{R}} |P(S_n \leq \sqrt{n}x) - \bar{F}(x)| \leq \frac{6.5 E|X|^3}{\sqrt{n}}$

lem.  $a < b, \forall i P(a \leq W_{n-1} \leq b) \leq b-a + \frac{2E|X|^3}{\sqrt{n}}$

lem 2. 设  $f_g$  为  $f(w) - w f'(w) = I(w \leq z) - \bar{F}(z)$  的解, 则  $|f_g(w+t) - f_g(w+|t|)| \leq |t| + |t| |w+|t|| + I(z-t \leq w-z) I(z \leq t) + I(z-s \leq w-s-t) I(s > t)$

$|w+|t|| f_g(w+|t|) - |w+|t|| f_g(w+|t|) | \leq (|w| + \frac{|t|}{2}) (|w| + |t|)$

Def.  $\alpha$ -交换对:  $(W, W') \stackrel{\alpha}{=} (W, W)$   
 $\alpha$ -Stein  $\alpha$ -交换对:  $E(W'W) = (1-\alpha)W, 0 < \alpha \leq 1$ . 且  $(W, W')$  可交换对.

Prop. 1.  $EW=0$

若  $\text{Var}(W)=0^2$ , 则  $E(W'-W)^2 = 2\alpha \sigma^2$

考虑  $F(W) = \int_0^W f(t) dt, f \in C^2$

$$0 = E(f(W) - f(W)) \stackrel{\text{Taylor}}{=} E[(W-W)f(W) + \frac{1}{2}(W-W)^2 f'(W) + \frac{1}{6}(W-W)^3 f''(W)]$$

又  $E(W-W)f(W) = E[(W-W)f(W)|W] = -\alpha EWF(W)$

$$\Rightarrow EWf(W) = \frac{1}{2\alpha} E(W-W)^2 f(W) + \frac{1}{6\alpha} E(W-W)^3 f''(W)$$

$$\Rightarrow |E(f(W)-Wf(W))| \leq \|f'\| E[|1 - \frac{1}{2\alpha} E(W-W)^2|W|] + \|f''\| \cdot \frac{E|W-W|^3}{6\alpha}$$

若  $\text{Var}(W)=1$ , 则  $E[1 - \frac{1}{2\alpha} E((W-W)^2|W)] = \frac{1}{2\alpha} E[(W-W)^2|W] - E(W-W)^2 / \leq \frac{1}{2\alpha} \sqrt{\text{Var}(E(W-W|W))}$

Thm. 若  $(W, W')$  为  $\alpha$ -Stein 交换对,  $\text{Var}(W)=1$ , 则  $d_W(W, Z) \leq \frac{\sqrt{\text{Var}(E(W-W|W))}}{\sqrt{2\alpha}} + \frac{E|W-W'|^3}{3\alpha}$

e.g. 设  $X_1, \dots, X_n$  独立,  $EX_i^4 < \infty, EX_i=0, \text{Var}X_i=1, W = \frac{1}{\sqrt{n}} \sum_{i=1}^n X_i$ . 估计  $d_W(W, Z)$ .

Sol. 设  $Z \sim \text{Unif}\{1, \dots, n\}$ ,  $(X_1, \dots, X_n)$  独立,  $E(X_i, \dots, X_n) = W' = W - \frac{X_1}{\sqrt{n}} + \frac{X_Z}{\sqrt{n}}$

则  $(W, W')$  可交换,  $E(W'-W) = E\left(\frac{X'_1 - X_1}{\sqrt{n}}\right) = \frac{1}{\sqrt{n}} \cdot E\left(\frac{1}{n} \sum (X'_i - X_i)\right) |W| = -\frac{1}{n} W$

$\therefore (W, W')$  为  $\frac{1}{n}$ -Stein 交换对

由  $\text{Thm. } d_W(W, Z) \leq \frac{\sqrt{\text{Var}(E(W-W|W))}}{\sqrt{2\alpha}} + \frac{E|W-W'|^3}{3\alpha}$

其中  $\frac{E|W-W'|^3}{\alpha} = \frac{1}{n^{\frac{3}{2}}} \sum_{i=1}^n E|X'_i - X_i|^3 \leq \frac{8}{n^{\frac{3}{2}}} \sum_{i=1}^n E|X_i|^3$

$$\begin{aligned} E[(W-W)^2 | X_1, \dots, X_n] &= E\left[\frac{(X'_1 - X_1)^2}{n} | X_1, \dots, X_n\right] = \frac{1}{n^2} E\left[\sum_{i=1}^n (X'_i - X_i)^2 | X_1, \dots, X_n\right] = \frac{1}{n^2} \sum_{i=1}^n (EX_i^2) \\ &\Rightarrow \sqrt{\text{Var}(E(W-W|W))} = n \sqrt{\frac{1}{n^4} \sum_{i=1}^n \text{Var}(X_i^2)} \leq \frac{1}{n} \sqrt{\sum_{i=1}^n EX_i^4} \\ &\therefore d_W(W, Z) \leq \frac{1}{\sqrt{2\alpha n}} \sqrt{\sum_{i=1}^n EX_i^4} + \frac{8}{3n^{\frac{3}{2}}} \sum_{i=1}^n E|X_i|^3 \end{aligned}$$

由离散情况:  $E f(X^S) = \sum f(k) P(X^S=k) = \frac{E X f(x)}{\mu}$

(e) 取  $f = X_{\{k\}}$

由构造的  $X^S$  为  $X$  的 size-bias 分布:  $h(x) = E(f(X)|X_i=x) = E(f(\sum_j X_j^{(i)} + X_i^S | X_i^S=x)$   
 $\therefore E(X_i f(X)) = E(X_i h(X_i)) = \mu_i E(h(X_i^S)) = \mu_i E(f(\sum_j X_j^{(i)} + X_i^S)) = \mu E(f(X^S) X_{\{i\}})$   
 $\therefore E(X f(X)) = \mu E f(X^S)$

Ex (Erdos图)  $G(n, p)$ ,  $X$ : 图中孤立点数,  $X = \sum_i X_i$ ,  $X_i = \chi_{V_i \text{是孤立点}}$ .

由构造方法,  $I \sim \text{Unif}\{1, \dots, n\}$  与其它 r.v. 独立.

取  $i=1$ , 构造  $X_j^{(i)}$ , s.t.  $(X_j^{(i)}, j \neq 1) \stackrel{d}{=} (X_j, j \neq 1 | X_1 = 1)$

$\Rightarrow X_j^{(i)} = \begin{cases} X_j & \text{若掉与 } V_j \text{ 相连的边后 } V_j \text{ 孤立} \\ X_j & \text{否则} \end{cases} = X_j \chi_{V_j \text{ 与 } V_i \text{ 相连}}$

$\Rightarrow X^S = \text{去掉与 } V_1 \text{ 相关的边后的孤立点个数.}$

### • Size-bias coupling

Def.  $X \geq 0$ ,  $E X = \mu < \infty$ , 称  $X^S$  为关于  $X$  的 size-bias 分布, 若  $\forall f$ ,  $E X f(X) < \infty$ , 有

$$E X f(X) = \mu E f(X^S)$$

Thm. 若  $X$  离散非负整值, 则  $P(X^S=k) = \frac{k P(X=k)}{\mu} \Leftrightarrow X^S$  有关于  $X$  的 size-bias 分布.

Thm (证明) 若  $X \sim F$ , 则  $dF^S(x) = \frac{x}{\mu} dF$ .

若  $X \geq 0$ ,  $E X = \mu$ ,  $\text{Var} X = \sigma^2$ , 则  $\forall W = \frac{X-\mu}{\sigma}$

$$E W f(W) = E \frac{X-\mu}{\sigma} f\left(\frac{X-\mu}{\sigma}\right) = \frac{1}{\sigma} E X f\left(\frac{X-\mu}{\sigma}\right) - \frac{\mu}{\sigma} E f\left(\frac{X-\mu}{\sigma}\right)$$

$$\stackrel{X \text{ size-bias}}{\leq} \frac{\mu}{\sigma} E\left[f\left(\frac{X-\mu}{\sigma}\right) - f\left(\frac{X-\mu}{\sigma}\right)\right] \stackrel{\text{Taylor}}{\leq} \frac{\mu}{\sigma} E\left[\frac{X^S-X}{\sigma} f'\left(\frac{X^S-X}{\sigma}\right) + \frac{(X^S-X)^2}{2\sigma^2} f''\left(\frac{X^S-X}{\sigma}\right)\right]$$

$$\Rightarrow |E[f(W) - W f(W)]| \leq |E f'(W)(1 - \frac{\mu}{\sigma^2} E(X^S - X))| + \frac{\mu}{\sigma^2} |E f''(\frac{X^S-X}{\sigma})(X^S - X)^2|$$

$$\Rightarrow d_W(W, Z) \leq \sqrt{\frac{\mu}{\sigma^2} E |f'(W)(X^S - X)|^2} + \frac{\mu}{\sigma^2} E (X^S - X)^2 \quad (\because E(X^S - X) = \frac{EX^2 - EX}{\mu} = \frac{\sigma^2}{\mu})$$

$\Rightarrow$  Thm  $X \geq 0$ ,  $E X = \mu$ ,  $\text{Var} X = \sigma^2$ ,  $X^S$  为  $X$  的 size-bias 分布.  $W = \frac{X-\mu}{\sigma}$ ,  $|Z|$

$$d_W(W, Z) \leq \sqrt{\frac{\mu}{\sigma^2} \frac{\mu}{\sigma^2} \sqrt{\text{Var}(E(X^S|X))} + \frac{\mu}{\sigma^2} E (X^S - X)^2}$$

• 如何构造  $X^S$ ? 考虑  $X = \frac{\mu}{\sigma} X_i$ ,  $X_i \geq 0$ ,  $E X_i = \mu$ .

(1) 设  $X_i^S$  为  $X_i$  的 size-bias 分布, 独立于  $X_{i-1}, X_{i+1}$ . 定义  $(X_j^{(i)})_{j \neq i}$ , s.t.  $(X_j^{(i)}, j \neq i | X_i^S=x) \stackrel{d}{=} (X_j, j \neq i | X_i=x)$

(2) 取  $I$ , s.t.  $P(I=i) = \frac{\mu_i}{\mu}$  与其它 r.v. 独立.

(3)  $X^S \triangleq \sum_{j \neq I} X_j^{(I)} + X_I^S$

Ex (特殊构造) (1)  $X_i$  独立,  $X_i^S$  与  $X_{i-1}, X_{i+1}$  独立, 则  $X^S = X - X_I + X_I^S$

(2)  $X_i$  为 0-1 r.v.,  $X_i^S = 1$ ,  $(X_j^{(i)}, j \neq i) \stackrel{d}{=} (X_j, j \neq i | X_i=1)$ , 则  $X^S = \sum_{j \neq I} X_j^{(I)} + 1$ .

### • Zero-bias coupling.

Def.  $E W = 0$ ,  $\text{Var} W = \sigma^2$ , 称  $W^S$  为关于  $W$  的 zero-bias coupling, 若  $\forall f$ ,  $E|W f(W)| < \infty$ , 有  $E W f(W) = \sigma^2 E f(W^S)$

Thm. 若  $\sigma^2 = 1$ , 则  $|d_W(W, Z)| \leq 2E|W-W^S| \quad \circledcirc |E(f(W)-W f(W))| = |E(f(W)-f(W^S))| \leq |f'(W)| E|W-W^S|$

Thm ( $W^S$  为 1 性)  $W^S \sim P^S(W) = \sigma^{-2} E(W I(W \leq \omega)) = -\sigma^{-2} E(W I(W > \omega))$

• 如何构造  $W^S$ ? 考虑  $W = \frac{\mu}{\sigma} X_i$ ,  $X_i$  独立,  $E X_i = 0$ ,  $\text{Var} X_i = \sigma_i^2$ ,  $\sum \sigma_i^2 = 1$

(1)  $\forall i$ ,  $X_i^S$  为  $X_i$  的 zero-bias coupling, 独立于  $X_{i-1}, X_{i+1}$

(2) 选取  $I$ , s.t.  $P(I=i) = \sigma_i^2$  且与其它 r.v. 独立

(3)  $W^S = W - X_I + X_I^S = \sum_i X_i^S + X_I^S$

由 of Lem (1).  $Ef(z+1) = E(zf(z))$

$$\begin{aligned} LHS &= \sum_{k=0}^{\infty} f(k+1) e^{-\lambda} \frac{\lambda^{k+1}}{k!} \\ &= \sum_{k=0}^{\infty} f(k+1) e^{-\lambda} \frac{\lambda^{k+1}}{(k+1)!} \cdot (k+1) = E(zf(z)). \end{aligned}$$

(2)  $\forall k, f(k) = I(k \geq j)$ , 由  $Ef(W+1) = E(Wf(W))$  得

$$\begin{aligned} \lambda P(W=j-1) &= P(Z=j) \\ \Rightarrow P(W=j) &= \frac{\lambda^j}{j!} P(W=0) \Rightarrow P(W=0) = e^{-\lambda} \Rightarrow W \sim P(\lambda). \end{aligned}$$

由 of Thm:  $d_{TV}(W, Z) \leq \sup_{f \in F} |E(\lambda f(W+1) - \lambda f(W))|$

$$\begin{aligned} \text{且 } Ef(W+1) &= \sum_{i=1}^n p_i Ef(W_i+1) \\ &\because Ew f(w) \\ &= \sum_i E X_i f(X_i + (w - X_i)) \\ &= \sum_i p_i E f(W_i+1), \quad W_i \triangleq w - X_i = \sum_{j \neq i} X_j \\ &\therefore |Ef(W+1) - Ef(W)| \leq \sum_{i=1}^n p_i |Ef(W_i+1) - f(W_i+1)| \\ &\leq \sum_{i=1}^n p_i \cdot EX_i \cdot \|\Delta f\| \leq (1/\lambda) \sum p_i^2 \end{aligned}$$

由 of Thm(不独立性):  $\because Ef(W) = \sum EX_i f(W_i+1), \quad W_i = W - X_i$

$$\begin{aligned} &= \sum EX_i (f(W_i+1) - f(V_i+1)), \quad V_i = \sum_{j \neq i} X_j \\ &+ \sum EX_i f(V_i+1) \\ &= \sum EX_i (f(W_i+1) - f(V_i+1)) + \sum p_i E f(V_i+1) \end{aligned}$$

$$\therefore \lambda Ef(W+1) = \sum p_i \lambda E f(W_i+1)$$

$$\therefore |Ef(W+1) - Ef(W)| \leq \left( \sum_i \frac{1}{\lambda} EX_i (f(W_i+1) - f(V_i+1)) \right) + \sum p_i |Ef(W_i+1) - Ef(W_i)|$$

$$\leq \|\Delta f\| \left( \sum EX_i (W_i - V_i) + \sum p_i E(W - V_i) \right)$$

由  $n$  次抛硬币, 正面  $p$  为  $p$ ,  $X$ : 不少于  $k$  个连续正面出现的次数

设  $Y_i$  为第  $i$  次出现正面

$$\begin{cases} Y_1 \sim Ber(p) \\ Y_i = (1 - Y_{i-1}) \prod_{j=1}^{i-1} Y_{i-j}, \quad i=2, 3, \dots, n-k+1 \end{cases}$$

记  $\lambda = EX = p^k (n-k) (1-p+1)$ , 则  $d_{TV}(X, P(\lambda)) \leq \sqrt{\frac{2k}{n-k+1}} + 2\lambda p^k$ .

由 Thm(不独立性),  $d_{TV}(X, P(\lambda)) \leq (1/\lambda) \left( \sum_{j \in N} p_j + \sum_{j \in N \setminus \{i\}} p_j \right)$

设  $N = \{j | j \leq k\}$ , 而  $j \in N$ ,  $i \in N$  时  $p_j = 0$ ,  $p_i = p^k$ ,  $p_j = (1-p)p^{k-1}$

$$d_{TV}(X, P(\lambda)) \leq 0 \left( \frac{\log n}{n} \right).$$

Poisson 逼近 Stein-Chen 法:  $\lambda f(k+1) - kf(k) = h(k) - Ef(z)$ , 考虑  $TV : h(k) = I(k \in A) - P_\lambda(A)$

Lem  $\lambda > 0$ ,  $\forall f(k) = \lambda f(k+1) - kf(k)$ .

- (1)  $Z \sim P(\lambda)$ , 则  $f$  有界, 有  $Ef(Z) = 0$ .
- (2) 若  $\exists$  非负整数  $W$ , s.t.  $E \not\propto f(W) = 0$ ,  $f$  有界, 则  $W \sim P(\lambda)$ .

Thm Stein 方程  $\begin{cases} \lambda f(k+1) - kf(k) = I(k \in A) - P_\lambda(A) \\ f(0) = 0 \end{cases}$

有唯一解  $f_A(k) = \lambda^{-k} e^{\lambda} (\lambda - k)!$  ( $P_\lambda(A \cup U_k) - P_\lambda(A) P(U_k)$ ).  $U_k = \{0, 1, \dots, k-1\}$

由  $\|f_A\| \leq \min\{\frac{1}{\lambda}, \frac{1}{k}\}$ ,  $\|\Delta f_A\| \leq \frac{1-e^{-\lambda}}{\lambda} \leq \min\{\frac{1}{\lambda}, \frac{1}{k}\}$ , 其中  $\Delta f_A(k) = f_A(k+1) - f_A(k)$ ,  $\|f\| = \|f_A\|$

$\because Z \sim P(\lambda)$ , 关心  $|P(W \in A) - P(Z \in A)| = |E(\lambda f_A(W+1) - \lambda f_A(W))|$ ,

$$\begin{aligned} d_{TV}(W, Z) &= \sup_{f \in F} |P(W \in A) - P(Z \in A)| \leq \sup_{f \in F} |E(\lambda f(W+1) - \lambda f(W))|, \\ \text{其中 } F &= \{f : Z \geq 0 \rightarrow \mathbb{R} / \|f\| \leq \min\{\frac{1}{\lambda}, \frac{1}{k}\}, \|f'\| \leq \frac{1-e^{-\lambda}}{\lambda}\}. \end{aligned}$$

2 cases: (1)  $W = \sum I_i$ ,  $I_i$  独立 (2) size-bias coupling:  $Ew f(w) = \lambda E f(w)$

Thm  $X_i \sim Ber(p_i)$ ,  $W = \sum_{i=1}^n X_i$ ,  $\lambda = EW = \sum_{i=1}^n p_i$ . 若  $Z \sim P(\lambda)$ , 则  $d_{TV}(W, Z) \leq (1/\lambda) \sum p_i^2$

Cor (独立性)  $\lambda = np \equiv C$  与  $n$  无关, 则  $d_{TV}(W_n, Z) \leq \frac{\lambda^2}{n} \rightarrow 0$ .  $W_n \xrightarrow{d} P(\lambda)$

②  $p = n^{-\frac{1}{4}}$ ,  $np \geq n^{\frac{1}{4}}$ , 但  $d_{TV}(W_n, Z) \leq n^{-\frac{1}{2}} \rightarrow 0$ , 其中  $Z_n = X_1 + \dots + X_{EW_n}$ ,  $X_i \stackrel{iid}{\sim} P(\lambda)$

又由  $\frac{Z_n - E Z_n}{\sqrt{Var(Z_n)}} \xrightarrow{d} N(0, 1)$  及  $\frac{W_n - E W_n}{\sqrt{Var(W_n)}} \xrightarrow{d} N(0, 1)$

$$(1) P\left(\frac{W_n - E W_n}{\sqrt{Var(W_n)}} \leq x\right) = P\left(\frac{Z_n - E Z_n}{\sqrt{Var(Z_n)}} \leq x\right) + \underbrace{[P\left(\frac{W_n - E W_n}{\sqrt{Var(W_n)}} \leq x\right) - P\left(\frac{Z_n - E Z_n}{\sqrt{Var(Z_n)}} \leq x\right)]}_{\text{绝对值} \leq d_{TV}(W_n, Z)} \rightarrow 0$$

Thm  $X_i \sim Ber(p_i)$ ,  $W = \sum_{i=1}^n X_i$ ,  $\lambda = EW = \sum_{i=1}^n p_i$ . 设  $N_i \subset \{1, \dots, n\}$ , s.t.  $i \in N_i$  且  $i$  与  $\{j | j \in N_i\}$  独立.

记  $p_j = EX_j$ , 若  $Z \sim P(\lambda)$ , 则

$$d_{TV}(W, Z) \leq (1/\lambda) \left( \sum_{i \in N} \sum_{j \in N_i} p_i p_j + \sum_{i \in N} \sum_{j \in N \setminus N_i} p_i p_j \right)$$

Eq. 2 续上例  $R_n, R_n$ : 最长连续正面长度  $P(R_n < k) = P(X < k \text{ 个连续正}) = 0$

$$\begin{aligned} &|P(X \leq k) - P(P(\lambda) = 0)| = O\left(\frac{\log n}{n}\right) \\ &\Rightarrow |P(R_n - \frac{\log n (1-p)}{\log p} < k) - e^{-pk}| \leq C \cdot \frac{\log n}{n}. \\ &\Rightarrow \frac{R_n}{\frac{\log n (1-p)}{\log p}} \xrightarrow{P} 1 \Rightarrow \frac{R_n}{\log n} \xrightarrow{P} \frac{1}{\log p} \quad \text{进一步还可证 a.s. 收.} \end{aligned}$$

• size-bias coupling:  $E W f(W) = \lambda E f(W^s)$ .

Thm.  $d_{TV}(W, P(\lambda)) \leq \min\{1, \lambda\} E|W+1-W^s|$

$$\leq \sup_{f \in \mathcal{F}} |E W f(W) - \lambda E f(W^s)| \quad //$$

$$\leq \sup_{f \in \mathcal{F}} \lambda \|f'\| \cdot E|W^s - (W+1)| \leq \min(1, \lambda) \cdot E|W+1-W^s|.$$

$W^s$  的构造:

0. 设  $X_i$  互不独立,  $W = \sum_{i=1}^n X_i$ :

1. 定义  $I$  与其它独立,  $P(I=i) = \frac{p_i}{\lambda}$

2.  $W^s = \sum_{j \notin I} X_j + 1$ . Check:  $\lambda E f(W^s) = \lambda \sum_{i=1}^n E f(W^s)_{\{I=i\}} = \sum_{i=1}^n p_i E f(\sum_{j \neq i} X_j + 1) = \sum_{i=1}^n E f(W)$

$\Rightarrow$  Cor.  $d_{TV}(W, P(\lambda)) \leq \min\{1, \lambda\} E|X_I| = \min\{1, \lambda\} \sum p_i^2$  与前文一致!

定理 of Portmanteau: (1)  $\Rightarrow$ (2), (3) $\Rightarrow$ (4) trivial.  
(2) $\Rightarrow$ (3):  $\forall F, \int^{\varepsilon} f(x) dP(x, F) < \varepsilon$   
 $\exists$ 一致连续  $f$ , s.t.  $I_F(x) \leq f_{\varepsilon}(x) \leq I_F(x)$  ( $f = (1 - \frac{P(x, F)}{\varepsilon})^+$ )  
 $\therefore \lim P(X_n \in F) \leq \lim E f_{\varepsilon}(X_n) = E f_{\varepsilon}(X) \leq P(X \in F)$  全  $\varepsilon > 0$  得证.  
(3) $\nRightarrow$ (4):  $\exists (F, E) \ni P_X(F) = P_X(E)$   
 $P(X \in E) \leq \lim P(X_n \in E) \leq \lim P(X_n \in \bar{E}) = P(X \in \bar{E}) \Rightarrow \lim P(X_n \in E) \neq P(X \in E)$   
 $\Downarrow \lim P(X_n \in E) \neq \lim P(X_n \in \bar{E})$   
(5) $\Rightarrow$ (1): WLOG,  $f \in [0, 1]$ .  
 $E f(X_n) = \int_0^1 P(f(X_n) > t) dt, E f(X) = \int_0^1 P(f(X) > t) dt.$   
若  $E_t = \{f(x) > t\}$ , 记  $T = \{t: P(f=t) > 0\}$ , 则  $T$  可数, 且  
 $\forall t \in T^c, E_t$  为  $X$ -连续集  $\Rightarrow P(f(X_n) > t) \rightarrow P(f(X) > t)$ .

定理 of 连续映射: 由 Portmanteau,  $\forall G$  为  $\mathcal{G}$  中的  $\sigma$ -代数,  $P(f(X) \in G) \leq \lim P(f_n(X_n) \in G)$   
注意  $f^{-1}(G) \cap D \subset \bigcup_n (f_n^{-1}(G))^o$   
 $P(f(X) \in G) \leq P\left(X \in \bigcup_n (f_n^{-1}(G))^o\right)$   
 $= \sup_m P\left(X \in \bigcup_n (f_n^{-1}(G))^o\right)$   
 $\leq \sup_m \lim P(X_n \in f_n^{-1}(G))$   
 $\leq \sup_m \lim P(X_n \in f_n^{-1}(G)) = \lim P(f_n(X_n) \in G)$ .

定理的  $\Leftarrow$  与  $\Rightarrow$  关系:

① 由 Portmanteau,  $\forall X$  为  $\mathcal{G}$  中的  $\sigma$ -代数,  $P(\{X_n \in E\} \Delta \{X \in E\})$   
 $= P(X_n \in E, X \in E^c) + P(X_n \in E^c, X \in E)$   
 $\leq 2P(P(X_n, X) \geq \varepsilon) + P(P(X_n, X) < \varepsilon, X_n \in E, X \in E^c)$   
 $+ P(P(X_n, X) < \varepsilon, X_n \in E^c, X \in E)$   
 $\leq 2P(P(X_n, X) \geq \varepsilon) + P(X \in E^c, P(X, E) < \varepsilon) + P(X \in E, P(X, E^c) < \varepsilon)$   
 $\xrightarrow{\substack{\downarrow n \rightarrow \infty \\ 0}} 0 \xrightarrow{\varepsilon \rightarrow 0} P(X \in \partial E) = 0$   
 $\rightarrow P(X \in \partial E) = 0$

② ( $\Leftarrow$ ):  $P(P(X_n, a) \geq \varepsilon) = P(X_n \in A)$ , 其中  $A = \{x \in S: P(x, a) \geq \varepsilon\}$ , 由  $\exists A = \{x \in S: P(x, a) = 1\}$  且  
 $\therefore X_n \xrightarrow{a}$ .

度量空间的弱收敛.

How to extend  $X_n \xrightarrow{d} X$  to  $(X'_1, \dots, X'_n) \xrightarrow{d} (X^1, \dots, X^k)$ ? (由  $Ee^{i\sum_k X'_k} \rightarrow Ee^{i\sum_k X^k}$  )  
 $(X'_1, X'_2, \dots) \xrightarrow{d} X$   
 $(X'_n, t \geq 0) \xrightarrow{d} (X_t, t \geq 0)$

•  $(S, P)$ : 度量空间,  $\mathcal{B} = \mathcal{B}(S)$  为  $S$  中所有开集生成的 Borel  $\sigma$ -域.  $X: (\Omega, \mathcal{F}) \rightarrow (S, \mathcal{B})$ .

Def.  $(\Omega, \mathcal{F}, P)$ : prob. space,  $X$ : mapping  $\Omega \rightarrow S$ . 称  $X$  可测, 且称为  $(S, \mathcal{B})$  上随机元.

若  $\forall B \in \mathcal{B}(S)$ , 有  $X^{-1}(B) \triangleq \{w \in \Omega: X(w) \in B\} \in \mathcal{F}$ .

Def. 随机元  $X$  的 prob. dist:  $P_X(B) = P(X^{-1}(B)) = P(X \in B)$

Def. (概率收敛):  $C = C(S) \triangleq \{S$  上有界连续函数全体}.

若  $\forall f \in C(S)$ ,  $E f(X_n) \rightarrow f(X)$ ,

则称  $X_n \xrightarrow{d} X$ ,  $X_n$  依分布收敛于  $X$ .

$P_{X_n} \Rightarrow P_X: \int_S f dP_{X_n}(x) \rightarrow \int_S f dP_X$ ,  $\forall f \in C(S)$

(收敛性-性)  $\nu_n \Rightarrow \nu: \int_S f d\nu_n \rightarrow \int_S f d\nu$ ,  $\forall f \in C(S)$ .  
prop. 若  $X_n \xrightarrow{d} Y$ ,  $X_n \xrightarrow{a.s.} Y$ , 则  $X \xrightarrow{d} Y$ . ( $\forall f \in C(S)$ ,  $E f(X_n) \rightarrow E f(X)$ ,  $E f(X_n) \rightarrow E f(Y) \Rightarrow E f(X) = E f(Y)$ )

Thm (Portmanteau)  $\{X_n\}$  为  $S$  上随机元, 则

(1)  $X_n \xrightarrow{d} X$  ( $\forall f$  为  $\mathcal{G}$  中的  $\sigma$ -连续,  $E f(X_n) \rightarrow E f(X)$ )

$\Leftrightarrow$  (2)  $\forall f$  有界一致连续,  $E f(X_n) \rightarrow E f(X)$

$\Leftrightarrow$  (3)  $\forall F$  闭,  $P(X \in F) \geq \lim P(X_n \in F)$

$\Leftrightarrow$  (4)  $\forall G$  开,  $P(X \in G) \leq \lim P(X_n \in G)$

$\Leftrightarrow$  (5)  $\forall X$  为  $\mathcal{G}$  中的  $\sigma$ -代数,  $P(X \in E) = P(X_n \in E)$ ,  $X$  为  $\mathcal{G}$  中的  $\sigma$ -代数,  $E: P_X(\partial E) = 0, \partial E = \bar{E} - E^c$

Thm (连续映射) 设度量空间  $S, T$ ; 随机元  $X_n \xrightarrow{d} X$ ,  $f: S \rightarrow T$  可测映射

( $f$  为  $\mathcal{G}$  中的  $\sigma$ -连续, 若  $\exists D \subset S$ ,  $P(X \in D) = 1$ , 且  $s_n \rightarrow s \in D$  且  $f(s_n) \rightarrow f(s)$   
 $f(s_n) \xrightarrow{a.s.} f(s)$ ) 则  $f(X_n) \xrightarrow{d} f(X)$ .

Cor.  $X_n \xrightarrow{d} X$  且  $f: S \rightarrow T$  在  $X$  上 a.s. 为  $\mathcal{G}$  中的  $\sigma$ -连续, 则  $f(X_n) \xrightarrow{d} f(X)$

Def.  $X_n \xrightarrow{P} X$ : 若  $\forall \varepsilon > 0$ ,  $P(P(X_n, X) \geq \varepsilon) \xrightarrow{n \rightarrow \infty} 0$ , 则称  $X_n$  依概率收敛于  $X$ .

Thm. ①  $X_n \xrightarrow{P} X \Rightarrow X_n \xrightarrow{d} X$

②  $X_n \xrightarrow{P} a \Leftrightarrow X_n \xrightarrow{d} a$ .

Def:  $\forall F \in \mathcal{F} \subset S$ , 记  $F_\varepsilon = \{x : p(x, F) \leq \varepsilon\}$

$$P(X_n \in F) \leq P(P(Y_{n_k}, X_n) \geq \varepsilon) + P(Y_{n_k} \in F_\varepsilon)$$

$$\liminf_n P(X_n \in F) \leq \liminf_n P(P(Y_{n_k}, X_n) \geq \varepsilon) + P(Y_{n_k} \in F_\varepsilon) \quad (\text{Portmanteau})$$

$$\text{全 } k \rightarrow \infty, \liminf_n P(X_n \in F) \leq P(X \in F_\varepsilon) \quad (\text{Portmanteau}).$$

全  $\varepsilon > 0$ , 则  $F_\varepsilon \downarrow F$ ,  $\liminf_n P(X_n \in F) \leq P(X \in F)$  由 Portmanteau,  $X_n \xrightarrow{d} X$ . #

Prf of Lem1: 设  $X_n \sim P_n, X \sim P$ .  $\forall A_i \in \mathcal{F}$ ,

$$P_n(\bigcup_i A_i) = \sum_i P_n(A_i) - \sum_i P_n(A_i \cap A_j^c) + \sum_{i < j} P_n(A_i \cap A_j) - \dots$$

$$\rightarrow P(\bigcup_i A_i)$$

$$\forall G \in \mathcal{F}, G = \bigcup_i A_i, A_i \in \mathcal{F}. \quad \forall \varepsilon > 0, \exists r, \text{s.t. } P(\bigcup_i A_i) > P(G) - \varepsilon.$$

$$\therefore P(G) - \varepsilon < \liminf_n P_n(\bigcup_i A_i) \leq \liminf_n P_n(G)$$

由  $\varepsilon$  任意  $P(G) \leq \liminf_n P_n(G)$ , 由 Portmanteau,  $X_n \xrightarrow{d} X$ . #

Prf of Lem2: 需证  $\forall G \in \mathcal{F}$ , 可写为  $\mathcal{F}$  中可数并, 从而用 Lem1 即证.

$\forall x \in G, \exists \varepsilon > 0, \text{s.t. } B(x, \varepsilon) \subset G$ .

$\exists A_x \in \mathcal{F}, \text{s.t. } x \in A_x^\circ \subset A_x \subset B(x, \varepsilon) \subset G$ .

又  $G = \bigcup_{x \in G} A_x^\circ$ . 由  $S$  可分,  $G$  有可数覆盖, 由  $P(G) = \bigcup_{i=1}^{\infty} A_{x_i}^\circ = \bigcup_{i=1}^{\infty} A_{x_i}$ . #

Prf of  $\lim_{n \rightarrow \infty}$  中  $X_n \xrightarrow{d} X \Leftrightarrow \forall x \in C_F, F_n(x) \rightarrow F(x)$ :

( $\Rightarrow$ )  $\forall x \in C_F$ , 有  $F(x_n) \rightarrow F(x), \forall n \rightarrow \infty$  (高维, 注意!).

记  $A_x = (-\infty, x] \triangleq \{y \in \mathbb{R}^k, y \leq x\}$

$P(X \in A_x) \leq P(X \in A_{x+\varepsilon} \setminus A_{x-\varepsilon})$

$\leq F(x+\varepsilon) - F(x-\varepsilon) \rightarrow 0 \quad (\varepsilon \rightarrow 0) \Rightarrow A_x$  为  $X$ -连续集.

$\therefore F_n(x) = P(X_n \in A_x) \xrightarrow{d} P(X \in A_x) = F(x)$ .

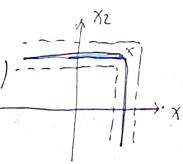
( $\Leftarrow$ ) 记  $\mathcal{F} = \{f(a, b) : 2k \text{ 个包含 } (a, b) \text{ 表面的 } k \text{ 维超平面的 } P_X \text{-测度为 } 0\}$ .

Check Lem2:

1° 有限支封闭  $\exists r > 0, \forall x \in \mathbb{R}^k, \exists (a, b) \subset (a, b) \subset B(x, r)$  且

2° 包含  $(a, b)$  表面的  $k$  维超平面的  $P_X$ -测度为 0.

(因为测度>0的超平面至多1个, 把1+2+3+...+n+...个超平面去掉, 剩下的测度为0)



Thm:  $\exists \frac{d}{d} Y_{n_k} \xrightarrow{k \rightarrow \infty} X$ . 若  $\forall \varepsilon > 0, \liminf_n \lim_k P(P(Y_{n_k}, X_n) \geq \varepsilon) = 0$ , 则  $X_n \xrightarrow{d} X$ .

Cor (Slusky):  $Y_n \xrightarrow{d} X$  且  $P(X_n, Y_n) \xrightarrow{d} 0$ , 则  $X_n \xrightarrow{d} X$ .

Thm (Skorokhod 表示定理): 若  $S$  可分,  $X_n \xrightarrow{d} X$ , 则  $\exists (\tilde{X}, \tilde{F}, \tilde{P})$  上  $S$ -值随机元  $(Y, Y_n)$  st.  $Y \xrightarrow{d} X, Y_n \xrightarrow{d} Y$  且  $\tilde{P}(Y_n \in A) \rightarrow \tilde{P}(Y \in A)$ .

Lem 1: 设可测集类  $\mathcal{F} \subset B$  满足: 1°  $\mathcal{F}$  有限支封闭 2°  $\forall x \in S, \exists A \in \mathcal{F}$ , st.  $x \in A \subseteq B(x, \varepsilon)$  且  $\forall A \in \mathcal{F}, P(X_n \in A) \rightarrow P(X \in A)$ , 则  $X_n \xrightarrow{d} X$ .

下面考虑  $\mathbb{R}^k$  情形.

Thm:  $X_n \xrightarrow{d} X \Leftrightarrow$  对  $\mathbb{R}^k$  的  $k$  連續点  $x \in \mathbb{R}^k$ , 有  $f_n(x) \rightarrow f(x)$

该乘积空间  $(\prod_{k=1}^{\infty} S_k, P)$ :  $P(x, y) = \sum_{k=1}^{\infty} \frac{1}{2^k} (P_k(x^k, y^k) \wedge 1)$  由  $(S_k, P_k)$  生成.

Prop:  $P(x, y) \rightarrow 0 \Leftrightarrow P_k(x^k, y^k) \rightarrow 0, \forall k$ . 与 正交量 等价

Thm: 若  $S_k$  可分, 则  $B(S) = \prod_{k=1}^{\infty} B(S_k) \cong \sigma(B_{(x, \prod_{k=1}^{\infty} S_k)})$

Prf: 令  $\pi_k : S \rightarrow S_k$ , 则  $\pi_k$  连续  $\Rightarrow$   $\exists B_k \in \mathcal{B}(S_k), \pi_k^{-1}(B_k) \in \mathcal{B}(S) \Rightarrow B \times \prod_{k=1}^{\infty} S_k \in \mathcal{B}(S)$

$$\therefore \prod_{k=1}^{\infty} B(S_k) \subset \mathcal{B}(S)$$

设  $A \in \mathcal{B}(S, P), \forall x \in G, \exists r, \text{s.t. } B(x, r) \subset G$

若  $A_x = B(x^1, \frac{r}{2^1}) \times \dots \times B(x^n, \frac{r}{2^n}) \times \prod_{k=n+1}^{\infty} S_k$ , 其中  $\frac{1}{2^1} \leq \frac{r}{2^n}$

则  $\forall y \in A_x, P(y, x) \leq \sum_{k=1}^n \frac{1}{2^k} \cdot \frac{r}{2} + \frac{r}{2} < r \Rightarrow A_x \subset B(x, r) \subset G$

由  $S_k$  可分  $\Rightarrow S$  可分, 且  $G \subset \bigcup_{x \in G} A_x \Rightarrow G \subseteq \bigcup_{x \in G} A_x$  且  $G$  是  $S_k$ -值随机元  $\therefore B(S) \subset \prod_{k=1}^{\infty} B(S_k)$ . #

$\mathcal{B}(B(S)) = \sigma(f_k)$ , 而  $f_k$  为  $\pi_k$ -类,  $f_k^{-1} \subset \prod_{k=1}^{\infty} B(S_k)$ .

Thm:  $\forall S_k$  可分,  $X, X_n = \prod_{k=1}^{\infty} S_k$ -值随机元 ( $\Leftrightarrow \forall k, X_k, X_{n_k}$  是  $S_k$ -值随机元)

则  $X_n \xrightarrow{d} X \Leftrightarrow (X'_1, X'_2, \dots, X'_n) \xrightarrow{d} (X^1, X^2, \dots, X^n), \forall m \in \mathbb{N}$ .

Prf: (1) 令  $\pi_1, \dots, \pi_m : x_1 \rightarrow (x^1, \dots, x^m)$ , 则 它连续, 由 连续映射定理 得证

(2)  $\forall a_k \in S_k, (x'_1, \dots, x'_m, a_{m+1}, \dots) \xrightarrow{d} (x^1, \dots, x^m, a_{m+1}, \dots)$

$$x'_{n+m} \xrightarrow{d} Y_m$$

则  $P(X, Y_m) \leq \frac{1}{2^m}, P(X_n, Y_{n+m}) \leq \frac{1}{2^m}, Y_m \xrightarrow{P} X$

$\forall \varepsilon > 0, \liminf_n \lim_k P(P(Y_{n_k}, X_n) \geq \varepsilon) = 0$ , 由 最上面的 Thm,  $X_n \xrightarrow{d} X$ . #

PF of Donsker. 由上项定理，只须证如下两点。

① 有限维分布收敛:  $\forall 0 \leq t_1 < \dots < t_k \leq 1$ ,

$$\begin{aligned} & \frac{1}{\sqrt{n}} (S_{[nt_1]}, \dots, S_{[nt_k]}) \\ &= \frac{1}{\sqrt{n}} (S_{[nt_1]}, S_{[nt_1]} - S_{[nt_2]}, \dots, S_{[nt_k]} - S_{[nt_{k-1}]}) \begin{pmatrix} 1 & \dots & 1 \\ 0 & \dots & 1 \end{pmatrix} \\ &\stackrel{\text{② } S_{[nt_1]} \xrightarrow{t \rightarrow 0} N(0, 1), S_{[nt_i]} \xrightarrow{t \rightarrow 0} B_t, \text{ 且有}}{\rightarrow} (B_{t_1}, B_{t_2} - B_{t_1}, \dots, B_{t_k} - B_{t_{k-1}}) \begin{pmatrix} 1 & \dots & 1 \\ 0 & \dots & 1 \end{pmatrix} = (B_{t_1}, \dots, B_{t_k}). \end{aligned}$$

lem.  $X_1, \dots, X_n$  独立,  $X_n \xrightarrow{d} X_\infty$ , 则  $(X_1^k, \dots, X_n^k) \xrightarrow{d} (X_\infty^k, \dots, X_\infty^k)$  (证明可用 ch.f.)

② 相关性:

(Levy):  $\forall \varepsilon > 0$ ,  $P(\max_{1 \leq k \leq n} |S_k - \text{med}(S_n - S_k)| \geq \varepsilon) \leq 2P(|S_n| \geq \varepsilon)$

$$\therefore P(|S_n - S_k| \geq \sqrt{2\text{Var}(S_n - S_k)}) \leq \frac{1}{2}$$

$$\therefore |\text{med}(S_n - S_k)| \leq \sqrt{2\text{Var}(S_n - S_k)} \leq \sqrt{2n}\sigma$$

$$\therefore \forall \lambda > 2, P(\max_{1 \leq k \leq n} |S_k| \geq \lambda \sqrt{n}\sigma) \leq 2P(|S_n| \geq (\lambda - \sqrt{2})\sqrt{n}\sigma)$$

$$\rightarrow 2P(|N(0, 1)| \geq \lambda - \sqrt{2}) \leq C \frac{1}{\sqrt{2\pi}} \frac{(\lambda - \sqrt{2})^2}{2}$$

$$\begin{aligned} |X_n(t) - X_n(s)| &\leq \frac{1}{\sqrt{n}} \max_{1 \leq k \leq n} |S_{[nt]} - S_{[ns]}|, |S_{[nt]} - S_{[ns+1]}|, |S_{[ns+1]} - S_{[ns]}|, |S_{[ns]} - S_{[ns]}| \\ \therefore P(\sup_{t \leq s \leq t+1} |X_n(t) - X_n(s)| \geq \varepsilon) &\leq P(2 \sup_{0 \leq k \leq n([t]+1)} |S_k - S_{[k+1]}| \geq \sqrt{n}\sigma \varepsilon) \leq P(\sup_{0 \leq k \leq n} |S_k| \geq \sqrt{n}\sigma \varepsilon) \\ &\leq \frac{1}{\sqrt{n}} \lim_{\delta \rightarrow 0} \sup_{0 \leq k \leq n} P(\sup_{s \leq t \leq s+1} |X_n(t) - X_n(s)| \geq \varepsilon) \\ &\leq \frac{1}{\sqrt{n}} \lim_{\delta \rightarrow 0} P(2 \sup_{0 \leq k \leq n([t]+1)} |S_k| \geq \sqrt{n}\sigma \varepsilon) \\ &\leq \frac{1}{\sqrt{n}} C \cdot e^{-\frac{(\lambda - \sqrt{2})^2}{2}} \rightarrow 0 \quad (\delta \rightarrow 0) \end{aligned}$$

#

PF of  $\frac{\max S_i}{\sqrt{n}} \xrightarrow{d} |N|$ :

记  $h: C[0, 1] \rightarrow \mathbb{R}$ , 则  $h$  连续

$$\begin{aligned} & x \mapsto \sup_{t \in [0, 1]} x(t) \\ & \frac{\max S_i}{\sqrt{n}} = h(X_n) \xrightarrow{d} h(B) = \sup_{t \in [0, 1]} B_t \end{aligned}$$

下证  $\sup B_t \xrightarrow{d} |N|$  且  $P\bar{0}$ .

设  $\xi_i \sim \text{sym Bern}$ ,  $\tau = \inf\{t \geq 1 : S_i \geq k\}$ , 求  $\max_{1 \leq i \leq n} S_i$  分布

则  $P(\max_{1 \leq i \leq n} S_i \geq k) = P(\tau \leq n) = P(S_n \geq k) + \sum_{i=1}^n P(S_n < k, \tau = i)$

$$\text{reflection} \quad \bar{P}(S_n \geq k) + \sum_{i=1}^n P(S_n < k, \tau = i) = P(S_n \geq k) + P(S_n > k)$$

伊顿 Eton:  $P(\max_{1 \leq i \leq n} S_i > \chi \sqrt{n}\sigma) = 2P(S_n > \chi \sqrt{n}\sigma) + P(S_n = \chi \sqrt{n}\sigma) \rightarrow 2P(N(0, 1) > \chi) + P(N(0, 1) = \chi)$

(有的教科书证明:  $P(S_n < k, \tau = i) = P(S_i = k, S_n - S_i < 0, \tau = i) = P(S_i = k, \tau = i)P(S_n - S_i < 0) = P(S_i = k)P(S_n - S_i < 0) = P(S_i = k, \tau = i)$ )

Thm  $X, X_1, X_2, \dots: C[0, 1]$ -值随机元, 则  $X_n \xrightarrow{d} X \Leftrightarrow \int X_n \frac{dx}{dt} X$   
其中  $w(x, h)$  为  $X$  的连续核,

$$w(x, h) \triangleq \sup_{t \in [0, 1]} |x(t) - h(t)|, \forall t \in C[0, 1], h > 0.$$

Rmk. 1. "lim" 可换成 "sup"

2.  $w(X_n, h)$  为 r.v. ( $\Rightarrow w(x, h)$  关于  $x$  连续)

3. (\*)  $\Leftrightarrow \forall \varepsilon > 0, \forall \eta > 0, \exists \delta(\varepsilon, \eta), \exists n_0(\varepsilon, \eta)$ , s.t.  $P(w(X_n, \delta) \geq \varepsilon) \leq \eta, \forall n \geq n_0$ .

$$\Leftrightarrow \forall \varepsilon > 0, \forall \eta > 0, \exists \delta(\varepsilon, \eta), \exists n_0(\varepsilon, \eta), \text{s.t. } \forall 0 \leq t \leq 1, \forall n \geq n_0, P\left(\sup_{t \leq s \leq t+1} |X_n(s) - X_n(t)| \geq \varepsilon\right) \leq \eta$$

$$\Leftrightarrow \forall \varepsilon > 0, \lim_{\delta \rightarrow 0} \sup_t P\left(\sup_{t \leq s \leq t+1} |X_n(s) - X_n(t)| \geq \varepsilon\right) = 0.$$

BM

Def. (标准 Brown 运动).  $[0, 1]$  上随机过程  $\{B_t\}_{t \in [0, 1]}$ , 若

①  $B_0 = 0$ ,  $B_t$  独立平稳增量

②  $B_t \sim N(0, t)$

③  $B_t$  连续样本路径

称  $\{B_t\}$  为标准 BM.

Def. (Gauss 过程)  $\forall$  有限维分布都是 Gaussian, 分布由均值  $\text{Cov}$  决定.

Rmk. 1. BM 是均值 0,  $\text{Cov}(X_s, X_t) = s \wedge t$  的 Gauss 过程

2.  $\{B_t\}_{t \geq 0}, \{CB_{C^2 t}\}_{t \geq 0}$  都为 BM

· 设  $\xi_1, \xi_2, \text{ind}, E\xi_i = 0, E\xi_i^2 = \sigma^2$ ,  $X_n(t) \triangleq \frac{1}{\sqrt{n}} \left( \sum_{k=1}^{[nt]} \xi_k + (n-t)[n+1] \right) \xi_{[nt+1]} = \frac{1}{\sqrt{n}} (S_{[nt]} + (n-t)\xi_{[nt+1]})$

Thm (Donsker)  $X_n \xrightarrow{d} B$ , 其中  $B$  是  $[0, 1]$  上标准 BM.

Thm  $\frac{\max S_i}{\sqrt{n}} \xrightarrow{d} |N|$

Def: 全  $\mathcal{D} = \text{Supp } f: [0, t]$  可分成有限个满足  $W_X([t_{i-1}, t_i])$  的小区间了, 只须证  $t=1$ .  
 首先,  $t>0$ . 因为在  $t=0$  处右连续.  
 若  $t<1$ , 则  $[0, t]$  可分成有限个满足的小区间, 由  $X(t)$  右连续, 则  
 $\exists \delta, [0, t+\delta]$  也可分成  $\cdots$  的小区间, 矛盾!  
 $\therefore t=1$ .

并

Def:  $D[0, 1]$ : r.c.l.l 函数全体, 即  $\forall x \in D[0, 1], \forall t \in [0, 1], t_n \downarrow t, x(t_n) \rightarrow x(t)$   
 $, x \in C[0, 1] \Leftrightarrow w(x, h) = \sup_{|s-t| \leq h} |x(s) - x(t)| \rightarrow 0, h \rightarrow 0$ .  
 $\lim_{t_n \uparrow t} x(t_n) = x(t)$  存在有限  
lem:  $\forall x \in D, \forall \varepsilon > 0, \exists 0=t_0 < \dots < t_r = 1$ , s.t.  $W_X([t_{i-1}, t_i]) = \sup_{s \in [t_{i-1}, t_i]} |x(s) - x(t_i)| < \varepsilon, \forall i=1, \dots, r$   
 $\Rightarrow$  (1) 至多可数个  $t$ , s.t.  $\Delta t = |x(t) - x(t_i)| > \varepsilon$   
 (2) 至多可数个  $t$ , s.t.  $\Delta t > 0 \Rightarrow$  至多可数个间断点  
 (3)  $\sup_{t \in [0, 1]} |x(t)| < \infty$   
 (4)  $x$  是 Borel 可测的(阶梯逼近):

• 另一刻画: 定义  $W'_X(\delta) \triangleq \inf_{\substack{\text{分段} \\ \text{分点} t_0 < \dots < t_r = 1}} \max_{s \in [t_{i-1}, t_i]} |x(s) - x(t_i)|$ , 其中  $t_{i-1}$  满足  $0=t_0 < \dots < t_r = 1$ , 则  
 $\Rightarrow x \in D[0, 1] \Leftrightarrow W'_X(\delta) \rightarrow 0 (\delta \rightarrow 0)$ .  
 $t_i - t_{i-1} > \delta$ .

Def: 称  $\lambda: [0, 1] \rightarrow [0, 1]$  为时间变换, 若它单增、双射,  $\lambda(0)=0, \lambda(1)=1$ , 连续.

Thm: (Skorokhod J<sub>1</sub>拓扑)  $\exists D[0, 1]$  上的完备可分度量  $d$ , s.t.  $d(x_n, x) \rightarrow 0$   
 $\Leftrightarrow \exists [0, 1]$  上时间变换  $\lambda_n$ , s.t.  $\sup_{0 \leq s \leq 1} |\lambda_n(s) - s| + \sup_{0 \leq s \leq 1} |x_n(\lambda_n(s)) - x(s)| \rightarrow 0$ .

Rmk:  $J(x_n, x) = \sup_{0 \leq s \leq 1} |\lambda_n(s) - s| + \sup_{0 \leq s \leq 1} |x_n(\lambda_n(s)) - x(s)|$  定义的度量不完备可分.  
 (从而定义距离:  $\Lambda = \|\lambda\| = \sup_{0 \leq s \leq 1} |\log \frac{x_n(\lambda(s))}{s} - \log \frac{x(\lambda(s))}{s}| < \infty$ )  
 $d(x, y) = \inf \{ \varepsilon: \exists \lambda \in \Lambda, \text{s.t. } \|\lambda\| \leq \varepsilon, \sup_{0 \leq s \leq 1} |x_n(\lambda(s)) - y(\lambda(s))| \leq \varepsilon \}.$

Rmk:  $d(x, y) \leq p(x, y)$ , 反之不对.  
 且  $x \in C[0, 1]$  时,  $d(x_n, x) \rightarrow 0 \Leftrightarrow p(x_n, x) \rightarrow 0$ .

$\therefore (\Rightarrow) \sup_{t \in [0, 1]} |x_n(t) - x(t)| = \sup_{t \in [0, 1]} |x_n(\lambda_n(t)) - x(\lambda_n(t))| \leq \sup_{t \in [0, 1]} |x_n(\lambda_n(t)) - x(t)| (d(x_n, x))$   
 $+ \sup_{t \in [0, 1]} |x(t) - x(\lambda_n(t))| (\rightarrow 0)$

Proposition: 设  $X_n, X$  是  $C[0, 1]$ -值随机元, 则在  $(D[0, 1], d)$  上  $X_n \xrightarrow{d} X$ ,  
 $\Leftrightarrow$  在  $(D[0, 1], P)$  上  $X_n \xrightarrow{P} X$ , 其中  $d$  为  $D[0, 1]$  上 Skorokhod J<sub>1</sub>拓扑  
 $P$  为  $D[0, 1]$  上一致拓扑.

Proposition: 投影  $\pi_t$  不连连续 (在 Skorokhod J<sub>1</sub>拓扑下), 但  $\pi_0, \pi_1$  在  $D[0, 1]$  上连续.  
 $\forall 0 < t < 1, \pi_t$  连续  $\Leftrightarrow x$  在  $t$  连续.

但  $\pi_{t_1}, \dots, \pi_{t_n}, \pi_{t_1}, \dots, \pi_{t_n}$  可测.

Thm: 设  $X, X_1, X_2, \dots$  为  $(D[0, 1], d)$  随机元, 则  $X_n \xrightarrow{d} X \Leftrightarrow$   
 ① 在稠密  $T = \{t \geq 0: d(X(t), X_n(t)) = 0, \forall s \in [0, t]\}$  上有  $X_n \rightharpoonup X$ , 且  
 ②  $\lim_{s \rightarrow 0} \lim_{n \rightarrow \infty} E(w(X_n, s) \wedge 1) = 0$ .  
 $\quad \text{②: } \forall \varepsilon > 0, \forall \eta > 0, \exists \delta(\varepsilon, \eta) \in (0, 1), N_0(\varepsilon, \eta) \in \mathbb{N}, \text{s.t. } \forall n \geq N_0, P(w(X_n, \delta) \geq \varepsilon) \leq \eta$   
 $\quad (\Leftarrow) \lim_{s \rightarrow 0} \lim_{n \rightarrow \infty} E(w(X_n, s) \wedge 1) = 0, \text{ 由 } w(X_n, \delta) \geq w'(X_n, \delta) \quad \text{若 (*) 成立且 } X_n \xrightarrow{P} X,$   
 $\quad \text{则 } P(X \in C) = 1$

Date:

Def of Thm:  $X_t^n = \frac{S_{n+1}}{\sqrt{n}} + (n-t) \frac{S_{n+1}}{\sqrt{n}} +$   
由 Donsker,  $X_n \xrightarrow{d} B$  on  $(C[0,1], \rho)$   
 $\Rightarrow X_n \xrightarrow{d} B$  on  $(D[0,1], d)$   
又  $d(X^n, \frac{S_{n+1}}{\sqrt{n}}) \leq p(X^n, \frac{S_{n+1}}{\sqrt{n}}) \leq \frac{\max|S_{n+1}|}{\sqrt{n}} \xrightarrow{a.s.} 0$   
由 Slutsky,  $\frac{S_{n+1}}{\sqrt{n}} \xrightarrow{d} B$ .

Prf of Prop 1.  $\{S_{t \wedge T} \leq t\} = \{S_{t \wedge T} \leq t\} \in \mathcal{F}_t$   
 $\{S_{t \wedge T} = t\} = \{S_{t \wedge T} \leq t\} \cup \{S_{t \wedge T} > t\} \in \mathcal{F}_t$   
 $\{S_{t \wedge T} > t\} = \{T < t\} \cup \{S_{t \wedge T} > t, S_{t \wedge T} \leq t\} \in \mathcal{F}_t$

Def of Prop (i)  $\forall A \in \mathcal{F}_S, A \cap \{T \leq t\} = A \cap \{S \leq t\} \cap \{T \leq t\} \in \mathcal{F}_t$   
(iii)  $\mathcal{F}_S \cap \mathcal{F}_{S \wedge T} = \mathcal{F}_{S \wedge T}$ . (by (i))  
另一方面,  $\forall A \in \mathcal{F}_S \cap \mathcal{F}_T, A \cap \{S \leq T\} = (A \cap \{S \leq t\}) \cup (A \cap \{S > t\}) \in \mathcal{F}_t$   
(iv) ①  $\{S < T\} \cap \{S \leq t\} = \bigcup_{r < t} \{S < r < T\} \in \mathcal{F}_t$ .  $\Rightarrow \{S < T\} \in \mathcal{F}_S \Rightarrow \{S < T\} \in \mathcal{F}_T$   
且  $\{S < T\} \cap \{T \leq t\} = \bigcup_{r < t} \{S < r < T \leq t\} \in \mathcal{F}_t \Rightarrow \{S < T\} \in \mathcal{F}_{S \wedge T}$   
②  $\{S \leq T\} = \{T < S\} \in \mathcal{F}_{S \wedge T}$  ③  $\{S = T\} = \{S \leq T\} \setminus \{S < T\} \in \mathcal{F}_{S \wedge T}$   
"且" ① 只须证 "C":  
 $V \in \mathcal{F}_{S \wedge T}, C = A \vee B, A \in \mathcal{F}_S, B \in \mathcal{F}_T, A \cap B = \emptyset$   
 $\therefore C \cap \{S \leq T\} = (A \cap \{S \leq T\}) \cup (B \cap \{S \leq T\})$   
而  $A \cap \{S \leq T\} \cap \{T \leq t\} = A \cap \{V \in \mathcal{F}_S \cap \{T \leq t\}\} \in \mathcal{F}_t \Rightarrow A \cap \{S \leq T\} \in \mathcal{F}_t$   
 $A \cap \{S \leq T\} \cap \{T > t\} = A \cap \{V \in \mathcal{F}_S \cap \{T > t\}\} \in \mathcal{F}_t \Rightarrow A \cap \{S \leq T\} \in \mathcal{F}_t$   
 $\Rightarrow C \cap \{S \leq T\} \in \mathcal{F}_t \cap \{S \leq T\}$ .  
(4) 用严格等价的证明中, 同理.

即设  $S, T$  为停时,  $X \in \mathcal{F}_T$  且  $X \in L'$ , 则  $E(X|\mathcal{F}_S) = E(X|\mathcal{F}_{S \wedge T})$ , a.s.

若:  $\{S < T\}$ , 显然.

on  $\{S > T\}$ ,  $E(X|\mathcal{F}_S) = E(X|\mathcal{F}_{S \wedge T}) = X$  a.s. 而  $E(X|\mathcal{F}_{S \wedge T}) = E(X|\mathcal{F}_T) = X$  a.s. #  
由伊藤定理  $E(E(Y|\mathcal{F}_T)|\mathcal{F}_S) = E(E(Y|\mathcal{F}_T)|\mathcal{F}_{S \wedge T}) = E(Y|\mathcal{F}_{S \wedge T}) \Rightarrow E(E(Y|\mathcal{F}_S)|\mathcal{F}_T) = E(E(Y|\mathcal{F}_S)|\mathcal{F}_T)$ .  
即  $E(E(Y|\mathcal{F}_T)|\mathcal{F}_S) \neq E(E(Y|\mathcal{F}_S)|\mathcal{F}_T)$ , 但停时却可以!

Date:

Thm.  $\exists_1, \exists_2$  i.i.d.  $E\exists_i = 0, E\exists_i^2 = 0$ ,  $B: [0,1]$  上标准 BM, 则在  $(DC_0, \mathcal{J}, d)$  上有  
 $\frac{S_{n+1}}{\sqrt{n}} \xrightarrow{d} B$ ,  $d$  是 Skorohod J, 拓扑.

Proposition  $X_i$  i.i.d. 若有稳定分布  $\frac{S_{n+1}}{\sqrt{n}} \xrightarrow{d} Y$ , 则  $\frac{S_{n+1}-\mu_{n+1}}{\sqrt{n}} \xrightarrow{d} Y(t)$  (稳定过程)  
 $Y(t)$  为独立平稳增量过程,  $Y(1) \equiv Y$ .

[reference] Chris Heyde & Allan Potter, Martingale Limit Theorem

事实:  $\{X_n, n \geq 0\}$  ( $\Omega, \mathcal{F}, P, \mathcal{F}_n, t \in T\}$ )  $T = \{0, \infty\}$   $\mathcal{F}_S \subset \mathcal{F}_t$  ( $S \leq t$ ):  $\sigma$ -域流

Def. (停时) 称  $T: \Omega \rightarrow \mathbb{R}$  为  $\{\mathcal{F}_t\}$  的停时, 若  $\forall t \in T$ ,  $\{T \leq t\} \in \mathcal{F}_t$ .

Remark: 若  $S, T$  为停时, 则  $SVT, SAT, S+T$  均为停时.

2. 若  $T_n$  为停时, 则  $\bigvee T_n$  为停时, 但  $\bigwedge T_n$  不一定.

Def. ( $\mathcal{F}_T$ ) 设  $T$  为停时, 则  $T$  前事件  $\sigma$ -域  $\mathcal{F}_T = \{A \in \mathcal{F}_S : A \cap \{T \leq t\} \in \mathcal{F}_t, \forall t \in T\}$ .  
其中  $\mathcal{F}_0 = \bigvee \mathcal{F}_t = (\sigma \mathcal{F}_t, t \in T)$

Prop. 设  $S, T$  为停时, 则

(i)  $T \in \mathcal{F}_T$

(ii)  $S \leq T \Rightarrow \mathcal{F}_S \subset \mathcal{F}_T$

(iii)  $\mathcal{F}_S \cap \mathcal{F}_T = \mathcal{F}_{S \wedge T}$

(iv)  $\mathcal{F}_S \vee \mathcal{F}_T = \mathcal{F}_{S \wedge T} \triangleq \{A \vee B : A \in \mathcal{F}_S, B \in \mathcal{F}_T, A \cap B = \emptyset\}$ .

(v)  $\{S \leq T\}, \{S < T\}, \{S = T\} \in \mathcal{F}_{S \wedge T}$ , 且

$\mathcal{F}_T \cap \{S \leq T\} = \mathcal{F}_T \cap \{S \leq T\}$

$\mathcal{F}_{S \wedge T} \cap \{S < T\} = \mathcal{F}_S \cap \{S < T\}$

$\mathcal{F}_{S \wedge T} \cap \{S = T\} = \mathcal{F}_S \cap \{S < T\}$

$\mathcal{F}_{S \wedge T} \cap \{S = T\} = \mathcal{F}_{S \wedge T} \cap \{S = T\} = \mathcal{F}_S \cap \{S = T\} = \mathcal{F}_T \cap \{S = T\}$ .

Thm. (条件期望) 设  $\sigma$ -域  $\mathcal{G}, \mathcal{H} \subset \mathcal{F}$ ,  $X, Y$  为可积 r.v.,  $A \in \mathcal{G} \cap \mathcal{H}$ .

若  $g \cap A = \mathcal{H} \cap A$  且  $X = Y$ , a.s. on  $A$ , 则  $E(X|g) = E(Y|\mathcal{H})$ , a.s. on  $A$ .

Ex. 记  $X_n = E(X|\mathcal{F}_n)$ , 则  $\forall$  有限停时  $T$ ,  $X_T = E(X|\mathcal{F}_T)$

#:  $\forall n \in T, \{A = \{T = n\}\}$  则在  $\{T = n\}$  上, a.s. 有  $X_T = X_n := E(X|\mathcal{F}_n) = E(X|\mathcal{F}_T)$   
在  $\{T \neq n\}$  上 a.s. 有  $X_T = E(X|\mathcal{F}_T)$ .

Date.  
设  $(X_t, t \in T)$  关于流  $(F_t, t \in T)$  适应, 定义  $D_A = \inf\{t \geq 0, X_t \in A\}$  进入时. (称为停时).

$\exists i, X_i$  独立,  $EY_i = 0$ , 则  $(S_n = \sum_{i=1}^n X_i, F_{X_n})$  为鞅.

$\exists i, X_i$  独立,  $EY_i = 1$ , 则  $(\sum_{i=1}^n X_i, F_{X_n})$  为鞅.

$X_1, X_2, \dots$  独立, 且  $\exists \lambda > 0$ , s.t.  $Ee^{\lambda X_i} < \infty$ , 则  $(\frac{e^{\lambda X_n}}{Ee^{\lambda X_n}}, F_{X_n})$  为鞅.

$\exists Y_0, Y_1, \dots$  i.i.d.,  $f_0, f_1$  为密度函数, 则  $Y_n \sim f_0$  时,  $(\frac{f_1(Y_1) \cdots f_n(Y_n)}{f_0(Y_1) \cdots f_n(Y_n)}, F_{Y_n})$  为鞅.

④ (罐子模型)  $n=0$  时一红一绿, 每次抽一个, 放回两个同色的球,

$X_n$ :  $n$  时刻红球占比, 则  $X_n$  为鞅.

(证明)  $Y_n = (n+2)X_n$ :  $n$  时刻红球数

⑤  $\forall (F_n)$  适应过程  $(Y_n)$  ( $\forall n, Y_n \in F_n$ ).  $(\sum_{i=1}^n (Y_i - E(Y_i | F_{i-1})), F_n)$  为鞅.

$(\sum_{i=1}^n (Y_i - E(Y_i | F_{i-1})) a_i (Y_1, \dots, Y_{i-1}), F_n)$  也为鞅

Pr of prop: 即证  $\forall \varepsilon > 0$ ,  $\exists M$ , s.t.  $E(E(Y|g) 1_{\{|E(Y|g)| > M\}}) < \varepsilon$ ,  $\forall g$

$\therefore \forall \varepsilon > 0$ ,  $\exists S > 0$ , s.t.  $\forall p(A) \leq S$ ,  $E|Y| 1_A < \varepsilon$ .

且  $P(|E(Y|g)| > M) \leq \frac{E|Y|}{M}$  aim LHS.

又取  $M = \frac{E|Y|}{\varepsilon}$  时,  $E|Y| 1_{\{|E(Y|g)| > M\}} < \varepsilon$

#

Pr of Doob分解:  $A_0 = 0$ ,  $A_n = \sum_{i=1}^n E(X_i - X_{i-1} | F_{i-1}) \in F_n$  可预报.

$M_n = X_n - A_n = X_n + \sum_{i=1}^n (X_i - E(X_i | F_{i-1}))$  为鞅.

$X$  为下鞅  $\Leftrightarrow A_{n+1} - A_n = E(X_{n+1} | F_n) - X_n \geq 0$ , a.s.

#

Eng时  $X_n = \sum_{k=1}^n 1_{B_k}$ ,  $B_k \in \mathcal{F}_k$  作 Doob分解:

$A_n = \sum_{k=1}^n P(B_k | F_{k-1})$ ,  $M_n = \sum_{k=1}^n (1_{B_k} - P(B_k | F_{k-1}))$ .

Rmk.  $\{ \sum_{k=1}^n 1_{B_k} = \infty \} = \{ B_n, \text{i.o.} \} \stackrel{\text{a.s.}}{=} \{ \sum_{k=1}^n P(B_k | F_{k-1}) = \infty \}.$

Date.  
Def. 设  $(X_t, t \in T)$  关于流  $(F_t, t \in T)$  适应, 定义  $D_A = \inf\{t \geq 0, X_t \in A\}$  进入时. (称为停时).  
 $T_A = \inf\{t \geq 0, X_t \in A\}$  着陆时

Rmk.  $T$  离散时,  $D_A, T_A$  为停时

连续时, 若  $A$  开,  $X$  连续,  $D_A = T_A$  不一定为停时

$A$  闭,  $X$  连续,  $D_A$  为停时.

Def. (鞅) 实值随机过程  $(X_n)_{n \geq 0}$ , 0-域流  $(F_n)_{n \geq 0}$ . 若  $\forall n \geq 0$ ,

①  $E|X_n| < \infty$  ②  $X_n \in F_n$  ③  $E(X_{n+1} | F_n) = X_n$ , a.s.

则称  $(X_n, F_n, n \geq 0)$  为鞅. 若③中“=”变成“ $\geq$ ”/“ $\leq$ ”, 则称为 下/上鞅

• 自然0-域流  $F_n = \sigma(X_1, \dots, X_n)$ , 则  $(X_n, F_n)$  为鞅  $\Rightarrow (X_n, F_n)$  为鞅.  $\oplus E(X_{n+1} | F_n) = E(X_{n+1} | F_n) | F_n$

•  $\forall Y \in F_n$ ,  $E(X_n | Y) = E(X_n | F_n)$

•  $(X_n, F_n)$  下鞅  $\Leftrightarrow (-X_n, F_n)$  上鞅.

• ③  $\Leftrightarrow \forall m > n$ ,  $E(X_m | F_n) = X_n$

•  $(X_n, F_n)$  下鞅, 则  $EY_n \uparrow$  且 下鞅  $(X_n, F_n)$  为鞅  $\Leftrightarrow EY_n = \text{const.}$

•  $Y \in L^1$ ,  $(F_n)$  0-域流, 则  $(X_n = E(Y | F_n), F_n)$  为一改可积鞅.

prop.  $f \in E(Y | g)$ ,  $g$  为子0-域流 一改可积, 若  $Y \in L^1$ .

Thm. 若  $(X_n, F_n)$  鞅,  $\phi: \mathbb{R} \rightarrow \mathbb{R}$ ,  $E|\phi(X_n)| < \infty$ , 则  $(\phi(X_n), F_n)$  为下鞅.

若  $(X_n, F_n)$  鞅,  $\phi: \mathbb{R} \rightarrow \mathbb{R}$  递增凸,  $\lim_n |\phi(X_n)| < \infty$ , 则  $(\phi(X_n), F_n)$  为下鞅.  $E(\phi(X_{n+1}) | F_n) \geq E(\phi(X_n) | F_n) | F_n$

Def.  $(X_n)$  关于  $(F_n)$  适应,  $H = (H_n)$  可预报 (即  $H_n \in F_{n-1}, n \geq 1$ ).

$H$  关于  $X$  的随机积分  $H \cdot X: (H \cdot X)_n = H_n X_n$

(也称作乘法变換)

$(H \cdot X)_n = \sum_{i=1}^n H_i (X_i - X_{i-1}) + H_0 X_0, n \geq 1$

Thm. 设  $H \cdot X$  可积. 若  $X$  为鞅, 则  $H \cdot X$  为鞅;  $(H \cdot X)_n \in F_n$

若  $X$  为下鞅,  $H$  非负, 则  $H \cdot X$  为下鞅.  $E((H \cdot X)_n - (H \cdot X)_{n-1} | F_n) = E(H_{n+1}(X_{n+1} - X_n) | F_n) = 0$

Rmk.  $N$  为停时,  $H_n = 1_{\{N \geq n\}} \in F_n$  为可预报过程  $\Rightarrow (X_n)$  为  $(F_n)$  鞅时,  $(X_{N+n})$  为  $(F_n)$  鞅.

2.  $\exists N = \inf\{n \geq 1, |X_n| > M\}$ . 则若  $|X_n - X_{n-1}| \leq M$ , 则  $|X_{N+n} - X_N| \leq M + \Delta$  为鞅, 但性质较好.

Thm. (Doob分解) 设  $(X_n)$  是  $(F_n)$  适应的, 则 存  $(A, g)$  分解:  $X_n = M_n + A_n$ , s.t.

$(M_n, F_n)$  为鞅,  $(A_n)$  可预报且  $A_0 = 0$ .

特别,  $X$  为下鞅  $\Leftrightarrow A$  a.s. 非降.

图 伊顿 Elton

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PF of upcrossing: 全  $Y_n = a + (Y_{n-1})^+$ , 则  $Y_n$  为下鞅,  $U_n$  为  $Y_n$  upcrossing time.

$$N_0 = 0, N_{2k+1} = \inf\{n \geq N_{2k}: Y_n = a\}, N_{2k} = \inf\{n \geq N_{2k}: Y_n \geq b\}$$

则  $N_n$  为停时 且  $Y_{N_{2k+1}} = a, Y_{N_{2k}} \geq b$ .

全  $H_n = \begin{cases} 1, & \exists k, s.t. N_{2k+1} < n \leq N_{2k} \\ 0, & \text{else} \end{cases}$ , 则  $\{H_n\} = \{Y_n \geq a\} \in \mathcal{F}_{n-1}$  为独立

$$\Rightarrow \{H_n\} \geq (b-a) U_n$$

$$\text{记 } V_n = 1 - H_n, \text{ 则 } Y_n = (H-Y)_n + (V-Y)_n$$

$\therefore (V-Y)_n$  为下鞅

$$\therefore E(V-Y)_n \geq E(V-Y)_0 = 0$$

$$\therefore (b-a) E(U_n) \leq E(H-Y)_n \leq E(Y_n - Y_0) = E(Y_n - a)^+ - E(Y_0 - a)^+$$

PF of 鞅收敛:  $a < b$ ,  $U$  表示上穿  $[a, b]$  次数, 则  $U_n \uparrow U$

$$E U_n \leq \frac{E(X_n - a)^+ - E(X_0 - a)^+}{b-a}$$

$$\leq \frac{E(Y_n + |a|)}{b-a} \leq \infty \Rightarrow U \leq \infty, a.s.$$

$$\therefore P(\lim_{n \rightarrow \infty} X_n < \lim_{n \rightarrow \infty} X_n)$$

$$= P(\lim_{n \rightarrow \infty} X_n < a < b < \lim_{n \rightarrow \infty} X_n)$$

$\therefore U_{a,b} < \infty, a.s. \therefore P(\lim_{n \rightarrow \infty} X_n < a < b < \lim_{n \rightarrow \infty} X_n) = 0. \therefore X_n a.s. \text{ 有极限} X$

$$\text{且 } E|X| = E \lim_{n \rightarrow \infty} |X_n| \leq \lim_{n \rightarrow \infty} E|X_n| \leq \sup_{n \in \mathbb{N}} E|X_n| < \infty$$

PF of Thm: (1)  $\Rightarrow$  (2):  $\sup_{n \in \mathbb{N}} E|X_n| < \infty \Rightarrow X_n \xrightarrow{a.s.} X + \text{一致可积} \Rightarrow Y_n \xrightarrow{L^1} X$

$$(2) \Rightarrow (3) \cdot (4)$$

(3)  $\Rightarrow$  (1): 自证.

若  $(Y_n)$  为鞅, (4)  $\Rightarrow$  (1) ( $\vee$ )

(3)  $\Rightarrow$  (5): 设  $X_n \xrightarrow{L^1} X$ , 下证  $X_n = E(X|\mathcal{F}_n)$ , 即  $\forall A \in \mathcal{F}_n, E(X_A|I_A) = E(X|I_A)$ .

$$\text{由 } X_n \text{ 鞅, } E(X_n|I_A) = E(X_m|I_A), \forall m > n = E(X|I_A).$$

$$\therefore E(X_n|I_A) = E(X|I_A) \xrightarrow{a.s.} E(X|I_A) \quad \#$$

PF of Thm:  $X_n = E(X|\mathcal{F}_n)$  一致可积  $\Rightarrow \exists \eta, s.t. E(X|\mathcal{F}_n) \xrightarrow{a.s.} \eta$  且  $\eta \in L^1, \eta \in \mathcal{F}_\infty$ .

下证  $\eta = E(X|\mathcal{F}_\infty)$ , a.s. 即  $\int_A \eta = \int_A X, \forall A \in \mathcal{F}_\infty$ .

由条件期望,  $\int_A \eta = \int_A E(X|\mathcal{F}_\infty) dP, \forall m > n$ , 而右边  $\rightarrow \int_A \eta dP$  由π-入定理得证. #

### 鞅收敛定理.

Thm (下鞅上穿不模式)  $-\infty < a < b < \infty, (X_n)$  下鞅 (sub). 若  $U_n$  为  $X_0, \dots, X_n$  从下往上穿过  $[a, b]$  的次数,

则  $(b-a) E(U_n) \leq E(Y_n - a)^+ - E(Y_0 - a)^+$ .

Thm (鞅收敛)  $(X_n)$ : 下鞅,  $\sup_n E X_n^+ < \infty$ , 则  $X_n$  a.s. 收敛  $\rightarrow X$ , 且  $E|X| < \infty$ .  
(在下鞅条件下,  $\sup_n E X_n^+ < \infty \Leftrightarrow \sup_n E|X_n| < \infty$ )

Cor  $(Y_n)$  非负上鞅, 则  $Y_n \rightarrow X$  a.s. 且  $EY \leq EX$ .

e.g. 1.  $\sup_n E X_n^+ < \infty \Rightarrow X_n \xrightarrow{L^1} X$ :

SRW,  $S_n = 1, T_0 < \infty$  a.s. 则  $S_{T_0+n}$  是鞅  $\xrightarrow{a.s.} 0$ . 但  $E S_{T_0+n} = \frac{n}{T_0+n} \rightarrow 0$ .

$X_n \xrightarrow{a.s.} X$  加  $X_n$  一致可积 即有  $L^1$ .

又  $X_n$  一致可积  $\Rightarrow \sup_n E|X_n| < \infty \Rightarrow X_n \xrightarrow{a.s.} X$  及  $L^1$

e.g. 2.  $\xi_n$  独立,  $P(\xi_n = 1) = 1 - \frac{1}{n^2}, P(\xi_n = -n^2 + 1) = \frac{1}{n^2}$ , 则  $S_n$  为鞅

$$\therefore \sum P(\xi_n = -n^2 + 1) < \infty \therefore P(\sum \xi_n = -n^2 + 1, i.o.) = 0$$

$$\therefore a.s. \xi_n = 1 \therefore X_n \rightarrow \infty, a.s.$$

e.g. 3. 疑在鞅  $X_n$ ,  $X_n \rightarrow \infty$  但有有限收敛极强. (鞅收敛定理条件有余必要)

$(X_n)$  为 MC,  $P(X_i = n+1 | X_0 = n) = \frac{2n+1}{2n+2}, P(X_i = -(n+1) | X_0 = n) = \frac{1}{2n+2}, P(X_i = -k | X_0 = n) = \frac{1}{2k+2}$

设  $X_0 = 0$ , 则  $X_n$  为鞅,

$P(Y_n \text{ 不收敛}) = P(X_0 = 0, X_1 = 1, X_2 = 2, \dots) = \prod_{n=0}^{\infty} \frac{2n+1}{2n+2} = 0 \quad \therefore X_n \text{ a.s. 有有限极限}$

但  $P(X_n = -k) = P(X_0 = 0, \dots, X_{k-1} = k-1, X_k = -k) \geq \frac{1}{2k(2k-2)}$

$$\therefore E|X_n| \geq \sum_{k=1}^n \frac{k}{2k(2k-2)} \rightarrow \infty$$

一致可积鞅.

Thm  $(X_n)$ : 下鞅, 则 (1)  $X_n$  一致可积

$\Leftrightarrow$  (2)  $X_n$  a.s. 及  $L^1$  收敛

$\Leftrightarrow$  (3)  $X_n$   $L^1$  收敛.

若  $(Y_n)$  鞅, 则  $\Leftrightarrow$  (4)  $\exists X \in L^1, s.t. X_n = E(X|\mathcal{F}_n)$ .

Thm.  $X \in L^1, (\mathcal{F}_n)$  递增  $\sigma$ -域流, 则  $E(X|\mathcal{F}_n) \xrightarrow{a.s. L^1} E(X|\mathcal{F}_\infty), \mathcal{F}_\infty = \sigma(\cup \mathcal{F}_n)$

Cor  $A \in \mathcal{F}_\infty, P(A|\mathcal{F}_n) \rightarrow 1_A, a.s.$

Cor 设  $T$  为独立 r.v.  $\{X_1, X_2, \dots\}$  属于  $\sigma$ -域  $\{T = \sigma(X_1, X_2, \dots)\}$  则  $\mathcal{F}_n = \sigma(X_1, \dots, X_n, T)$

(Kolmogorov)  $P(A|\mathcal{F}_n) = P(A), \forall A \in T \Rightarrow P(A) = 1_A, a.s. \Rightarrow P(A) = 0 \text{ 或 } 1$  伊顿 Eton 7

Date:

证明 Thm:  $E(Y|F_n) \xrightarrow{a.s. \& L^1} E(Y|F_\infty)$

$$\text{s.t. } E|E(Y|F_n) - E(Y|F_\infty)| \leq E|Y_n - Y| \xrightarrow{PCT} 0$$

$$\therefore E(Y_n|F_n) \xrightarrow{L^1} E(Y|F_\infty).$$

对 a.s. 收敛,  $\forall n \geq m$ ,  $E(\inf_{k \geq m} Y_k|F_n) \leq E(Y_n|F_n) \leq E(\sup_{k \geq m} Y_k|F_n)$   
 固定  $m$ , 对  $n \rightarrow \infty$ ,  $E(\inf_{k \geq m} Y_k|F_\infty) \leq \liminf E(Y_n|F_n) \leq \limsup E(Y_n|F_n) \leq E(\sup_{k \geq m} Y_k|F_\infty)$   
 全  $m \rightarrow \infty$  时,  $E(Y|F_\infty) \leq \liminf E(Y_n|F_\infty) \leq \limsup E(Y_n|F_\infty) \leq E(Y|F_\infty)$  a.s. #

证明  $L^p$  收敛:  $\because \sup E|X_n|^p < \infty \quad \because (X_n)$ -一致可积  $\Rightarrow X_n \xrightarrow{a.s. \& L^1} X_\infty$

下证  $X_n \xrightarrow{L^p} X_\infty$ , 只须  $X_n$  在  $L^p$ -一致可积.

①  $X_n$  轴, 则  $(X_n)$ -一致可积且  $X_n \geq E(X_n|F_n)$  a.s.

$$|X_n|^p = |E(X_n|F_n)|^p \leq E(|X_n|^p|F_n). \quad \text{只须 } X_\infty \in L^p$$

由  $E|X_n|^p \leq \lim E|X_n|^p \leq \sup E|X_n|^p < \infty$ . 得证.  $X_\infty \in L^p \Rightarrow X_n$ -一致可积  $\Rightarrow X_n \xrightarrow{L^p} X_\infty$ .

②  $X_n$  非负下轴, 则  $0 \leq X_n \leq E(X_n|F_n)$

$$|X_n|^p \leq |E(X_n|F_n)|^p \leq E(|X_n|^p|F_n). \quad \text{由①中类似可得 } |X_n|^p \text{-一致可积.} \quad \#$$

证明 Thm (上轴):  $\because \{X_n\}, \{Y_n\}$  为下轴, 记  $U_{a,b,N}$  为其上穿  $[a,b]$  的次数,

$$\forall j \exists U_{a,b,N} \leq \frac{H(y_j-a)}{b-a} < \infty$$

$$\Rightarrow EU_{a,b,N} \leq \frac{E(H(x-a))}{b-a} < \infty \Rightarrow U_{a,b,N} < \infty \text{ a.s.} \Rightarrow X_n \text{ a.s. 上轴.} \quad \#$$

若  $X_n$  为轴, 则  $X_n = E(X_n|F_n), \forall n \in \mathbb{N}$

若  $X_n$  下轴, 则  $X_n = M_n + A_n$ , 其中  $E(X_n|F_{n+1}) = M_{n+1} + E(A_n|F_{n+1})$   
 $\left\{ \begin{array}{l} \exists d_n = A_n - A_{n-1} = E(X_n|F_{n-1}) - X_{n-1} \geq 0, \text{ 则 } A_n = \sum_{k=n}^{\infty} d_k. \\ \text{下证 } \inf EY_n > -\infty \text{ 时 } (X_n) \text{-一致可积.} \end{array} \right.$

即  $d_n = E(Y_n - Y_{n-1}|F_n) \geq 0, \forall n \in \mathbb{N}$ , 且  $E \sum_{n=0}^{\infty} d_n = \lim_{K \rightarrow \infty} \sum_{n=K}^{\infty} E d_n = E X_0 - \lim_{K \rightarrow \infty} E \sum_{n=K}^{\infty} d_n < \infty$   
 $\therefore \sum_{n=0}^{\infty} d_n < \infty$ , a.s.

定义  $A_n = \sum_{k=n}^{\infty} d_k, M_n = Y_n - A_n$ , 则  $\{M_n\}$  为轴且  $X_n = M_n + A_n$ .

$\because \{M_n\}$ -一致可积,  $\exists C < \infty$  且  $E \sum_{n=0}^{\infty} d_n < \infty$   
 $\therefore \{A_n\}$ -一致可积,  $\therefore \{X_n\}$ -一致可积. #

Thm.  $Y_n \xrightarrow{a.s.} Y, |Y_n| \leq Z, Y, Z \in L^1$ , 则  $E(Y_n|F_n) \xrightarrow{a.s. \& L^1} E(Y|F_\infty)$

•  $L^p (p > 1)$  收敛.

Thm.  $X_n$  轴/非负下轴, 且  $\sup E|X_n|^p < \infty, p > 1$ , 则  $X_n \xrightarrow{a.s. \& L^p} X_\infty$

逆轴:

$$Y_1, Y_2, \dots, F_m, E(Y_m|F_{m-1}) = Y_{m-1}$$

Thm.  $(X_n, F_n)_{n \in \mathbb{N}}$ : 下轴, 则  $\exists X_\infty$ , s.t.  $X_n \xrightarrow{a.s.} X_\infty (n \rightarrow \infty)$

且若  $\inf EX_n > -\infty$ , 则  $X_n$ -一致可积,  $X_n \xrightarrow{L^1} X_\infty \in L^1$ . 此时,  $\forall n, E(X_n|F_n) \geq X_\infty$ .

Rmk (Cor)  $(X_n)_{n \in \mathbb{N}}$  轴, 则  $X_n \xrightarrow{a.s. \& L^1} X_\infty$  且  $E(X_n|F_n) = X_\infty$ . 其中  $F_\infty = \bigcap_{n \in \mathbb{N}} F_n$ .

(Cor)  $\exists g_n \downarrow F_\infty, \forall x \in L^1$ , 有  $E(X|F_n) \xrightarrow{a.s. \& L^1} E(X|F_\infty) (n \rightarrow \infty)$

(②)  $(X_n = E(X|F_n))_{n \in \mathbb{N}}$  独立, 则  $\exists X_\infty$ , s.t.  $X_n \xrightarrow{a.s.} X_\infty (n \rightarrow \infty)$ .

$$\forall g_\infty = \bigcap_{n \in \mathbb{N}} g_n = \bigcap_{n \in \mathbb{N}} F_n = F_\infty$$

$$X_\infty = E(X|g_\infty) = E(E(X|g_n)|g_\infty) = E(X|g_\infty) = E(X|F_\infty) \text{ a.s.}$$

(Cor)  $\Rightarrow Y_n \rightarrow Y, |Y_n| \leq Z, \text{ 则 } E(Y_n|F_n) \xrightarrow{a.s. \& L^1} E(Y|F_\infty) \quad (n \rightarrow \infty)$

应用 (SLUV)

Thm.  $X_1, X_2, \dots$  i.i.d.,  $E|X| < \infty$ , 则  $\frac{S_n}{n} \xrightarrow{a.s.} EX$ .

pf:  $S_n \triangleq \{S_1, S_2, \dots\} \downarrow$  且  $E(X_i|F_n) = E(X_k|F_n), \forall k \leq n$ .

$$\therefore E(X_i|F_n) = E\left(\frac{S_n}{n}\right) = \frac{S_n}{n}. \quad (\text{②} \Leftrightarrow \int_A X_i = \int_A X_k, \forall A \in \sigma(S_1, \dots) \Leftrightarrow E(X_i|F_n) = E(X_k|F_n) \Leftrightarrow (X_1, S_2, \dots) \triangleq (X_k, S_n, \dots))$$

$\therefore S_n \xrightarrow{a.s.} E(X|F_\infty) \in T$

$\therefore \exists C, \text{s.t. } E(X|F_\infty) = C \text{ a.s.} \quad \text{且 } C = E(E(X|F_\infty)) = EX_1$

$\therefore \frac{S_n}{n} \xrightarrow{a.s.} EX$ .

$$\sup_{n \in \mathbb{N}} (M_n - M_{n-1}) \quad \#.$$

Ex. 设  $\{M_n\}_{n \in \mathbb{N}}$  为轴且  $\exists C < \infty$ , s.t.  $|\Delta M| \leq C$ .

记  $C = \{\lim M_n \text{ 为有限}\}$ ,  $D = \{\lim M_n = \infty \text{ 且 } \lim M_n = -\infty\}$ ,  $\Omega \setminus D \cap C = \emptyset$ .

pf: 只须  $\lim M_n < \infty$  时  $C$  a.s. 成立.

(对  $\forall n$ ,  $\sup E M_n < \infty$  或  $\sup E M_n < \infty$  或  $\sup E M_n < \infty$ , 就有 a.s. 收敛)

考虑  $T_m = \inf \{n: M_n \geq m\}$ .

则  $M_{n \wedge T_m} \leq m + C \Rightarrow M_{n \wedge T_m}$  a.s. 收敛  $\Rightarrow \{\sup M_n < m\}$  时,  $M_n$  a.s. 收敛  $\Rightarrow \{\sup M_n < m\}$  时,  $M_n$  a.s. 收敛

$\Rightarrow \lim M_n < \infty$  时,  $M_n$  a.s. 收敛  $\Rightarrow D^c$  上  $C$  a.s. 成立

图 伊顿 Elon

Def of de Finetti:  $\forall$  有界可测  $\varphi: \mathbb{R}^k \rightarrow \mathbb{R}$ ,  $A_n(\varphi) \triangleq \frac{(n+k)!}{n!} \sum_{\{i_1, \dots, i_k \in [1, n]\}} \varphi(X_{i_1}, \dots, X_{i_k})$ , 则  $A_n(\varphi) \in \mathcal{E}_n$ .

$$\therefore A_n(\varphi) = E(A_n(\varphi) | \mathcal{E}_n) = \frac{(n+k)!}{n!} \sum \varphi(X_{i_1}, \dots, X_{i_k}) | \mathcal{E}_n$$

$$= E(\varphi(X_1, \dots, X_k) | \mathcal{E}_n) \rightarrow E(\varphi(X_1, \dots, X_k) | \mathcal{E})$$

下证  $X_1, X_2, \dots$  在  $\mathcal{E}$  下条件 i.i.d., 即

$$E(f_1(X_1), \dots, f_k(X_k) | \mathcal{E}) = E(f_1(X_1) | \mathcal{E}) \cdots E(f_k(X_k) | \mathcal{E}) \quad \dots \textcircled{1}$$

$$\text{且 } E(f_i(X_i) | \mathcal{E}) = E(f_i(X_i) | \mathcal{E}) \quad \dots \textcircled{2}$$

令  $\varphi(X_1, \dots, X_k) = f_1(X_1, \dots, X_k) g(X_k)$  有界, 其中  $f, g$  有理

$$\begin{aligned} A_n(\varphi) A_n(g) &= \frac{(n+k)!}{n!} \sum f(X_{i_1}, \dots, X_{i_k}) \cdot \frac{1}{n} \sum g(X_m) \\ &= \frac{(n+k)!}{n!} \frac{1}{n} \sum \varphi(X_{i_1}, \dots, X_{i_k}) + \frac{(n+k)!}{n!} \frac{1}{n} \sum f(X_{i_1}, \dots, X_{i_k}) \sum_{j=1}^k g(X_j) \\ &= \frac{n+k!}{n!} A_n(\varphi) + \frac{1}{n} \sum_{j=1}^k A_n(g_j), \text{ 其中 } g_j(X_{i_1}, \dots, X_{i-k}) = f_{i_1} \cdots f_{i_{k-1}} g(X_j) \end{aligned}$$

令  $n \rightarrow \infty$ ,  $E(\varphi(X_1, \dots, X_k) | \mathcal{E}) = E(f_1(X_1, \dots, X_k) | \mathcal{E}) \cdot E(g(X_m) | \mathcal{E})$

$\Rightarrow \textcircled{1}$

对  $\textcircled{2}$ , 若证  $\exists I, T \subset \mathcal{E}$ , 则  $E(\varphi(X_1, \dots, X_k) | \mathcal{E})$

$$2. E(\varphi(X_1, \dots, X_k) | \mathcal{E}) \in T = E(E(\varphi(X_1, \dots, X_k) | \mathcal{E}) | T)$$

$$\text{下证 } 2. \text{ 即 } \lim A_n(\varphi) \in T.$$

$$\forall m, n > m \text{ 时, } \exists i_1, i_2, \dots, i_m \in [1, n], \exists i_{m+1}, i_{m+2}, \dots, i_n \in [m+1, n]$$

$$\text{且 } |I_{m+1} - I_{m+2}| = \frac{n!}{(n-m)!} - \frac{(n-m)!}{(n-m-1)!}$$

$$\frac{(n+k)!}{n!} \left| \sum_{i_{m+1}}^{i_n} \varphi(X_{i_1}, \dots, X_{i_m}) \right| \leq \frac{(n+k)!}{n!} \left( \frac{n!}{(n-m)!} - \frac{(n-m)!}{(n-m-1)!} \right) \sup |\varphi| \rightarrow 0$$

$$\therefore \lim A_n(\varphi) = \lim \frac{1}{n!} \sum_{i_1, i_2, \dots, i_n} \varphi(X_{i_1}, \dots, X_{i_n}) \in \sigma(X_{m+1}, X_{m+2}, \dots)$$

$$\therefore \lim A_n(\varphi) \in \sigma(X_{m+1}, \dots) \subset T$$

R证缺:

即  $\forall A \in \mathcal{F}_S$ ,  $M \triangleq S1_A + T1_{Ac}$ , 则  $\{M \leq t\} = A \cap \{S \leq t\} + A^c \cap \{T \leq t\} \in \mathcal{F}_T$

$\therefore M$  为停时,  $M \leq T \Leftrightarrow EX_M = EX_T$

$$\Rightarrow EX_S 1_A = EX_T 1_A, \forall A \in \mathcal{F}_S \Rightarrow E(X_T | \mathcal{F}_S) \underset{\text{a.s.}}{=} X_{SAT}$$

$$(2) \Rightarrow (3) \{S \leq T\} \text{ 上, } E(X_T | \mathcal{F}_S) = E(X_T | \mathcal{F}_{SAT}) \stackrel{(2)}{=} X_{SAT}, \text{ a.s.}$$

$$\{S > T\} \text{ 上, } E(X_T | \mathcal{F}_S) = E(X_{SAT} | \mathcal{F}_S) = X_{SAT}, \text{ a.s.}$$

(3)  $\Rightarrow$  (2)  $\Rightarrow$  (1). trivial.

#

Ex. 2  $\forall \{f_n\}$  流,  $A_n \in \mathcal{F}_n$ ,  $\{A_n, i.o.\} = \{\sum P(A_n | \mathcal{F}_n) = \infty\} \text{ a.s.}$

即:  $\{A_n, i.o.\} = \{\sum I_{A_n} = \infty\}$

$M_n = \sum_{k=1}^n (1_{A_k} - P(A_k | \mathcal{F}_{n-1})) \geq \text{mar. 且 } (\Delta M) \leq 1$

由上 -  $\epsilon$ -记号, 在  $D$  上,  $\{\sum I_{A_n} = \infty\}$  且  $\{\sum P(A_n | \mathcal{F}_n) = \infty\}$  成立

在  $C$  上, 若  $\sum I_{A_n} = \infty$ , 则  $\{\sum P(A_n | \mathcal{F}_n) = \infty\}$  也成立

由上 -  $\epsilon$ -  $P(C \cup D) = 1$  知  $\{A_n, i.o.\} = \{\sum P(A_n | \mathcal{F}_n) = \infty\} \text{ a.s.}$

### · 可交换序列

Def.  $(X_1, \dots, X_N)$  称为 exchangeable, 若  $(X_{\pi(1)}, \dots, X_{\pi(N)}) \triangleq (X_1, \dots, X_N)$ ,  $\forall \pi: N \rightarrow N$  为换元

无穷  $(X_1, X_2, \dots)$  可交换, 若  $\forall N \in \mathbb{N}, (X_1, \dots, X_N)$  可交换

$\mathcal{E}_n = \{A: \exists R^\infty \text{ 可测 } B, \text{ s.t. } A = f(X_1, X_2, \dots) \in B\} \text{ 且 } \forall f_1, \dots, f_n \text{ 置换 } \pi, \text{ 有 } A = f(X_{\pi(1)}, \dots, X_{\pi(n)}, \dots) \in B\}$

Thm (de Finetti) 设  $X_1, X_2, \dots$  可交换, 则在  $\mathcal{E}$  的条件下,  $X_1, X_2, \dots$  条件独立同分布.

Rank:  $X_1, X_2, \dots$  关于  $\mathcal{E}$  条件 i.i.d., 还关于尾事件  $T$  条件 i.i.d.

Thm (Hewitt-Savage 0-1 律),  $X_1, X_2, \dots$  i.i.d.,  $A \in \mathcal{E}$ . 则  $P(A) = 0$  或 1.

即:  $\forall \varphi$  可测有界,  $E(\varphi(X_1, \dots, X_n) | \mathcal{E}) = E(\varphi(X_1, \dots, X_n) | T) \text{ a.s.}$

(i.i.d. 且 r.v. 为 const)  $= E\varphi(X_1, \dots, X_n)$

$\therefore X_1, \dots, X_n \perp \mathcal{E} \Rightarrow \mathcal{E} \perp \sigma(X_1, X_2, \dots) \Rightarrow \mathcal{E} \text{ 与 } \mathcal{E} \text{ 独立} \Rightarrow P(A) = P(A \cap \mathcal{E})$

$\Rightarrow P(A) = 0$  或 1.

Thm (鞅停时)  $(X_t, \mathcal{F}_t)$  为鞅则 (1)  $\forall S \leq T, EX_S \neq EX_T, \text{ a.s.}$

(下节) (2)  $\forall S \leq T, E(X_T | \mathcal{F}_S) = X_S, \text{ a.s.}$

(3)  $\forall S, T, E(X_T | \mathcal{F}_S) = X_{SAT}, \text{ a.s.}$

\*  $EX_T = EX_0$ ,  $\forall$  停时  $T$ , 则  $X_T$  为鞅

CBG (停时定理不成立) SRW,  $S_n = \sum_{i=1}^n X_i$ ,  $S \equiv 1$ ,  $T = \inf\{n: S_n = 1\}$ , 则  $\lim S_n = \infty$ ,  $\lim_{n \rightarrow \infty} S_n = \infty$  (a.s.)

但  $ES_T \neq ES_S$ .

Thm  $(X_t, \mathcal{F}_t)$  为鞅, 停时  $S, T$ ,  $T$  有界, 则  $E(X_T | \mathcal{F}_S) = X_{SAT}, \text{ a.s.}$  即  $\forall S \leq T, EX_S = EX_T, \text{ a.s.}$

PF:  $(X_n) \text{ 为 mar.} \Rightarrow EX_n = EX_0$ , 设  $S \leq T \leq M$  且  $X_T = EX_0, EX_S = EX_0$ , 则  $EX_T = EX_S$  # 29

$(X_n) \text{ 为 sub, } K_n = \inf_{S \leq t \leq n} E[X_t], \text{ 则 } (K_n) \text{ 为 sub, } (K_X)_n = \inf_{S \leq t \leq n} X_t \Rightarrow EX_{K_X} = EX_{K_0}$

Def of Lem:  $E[X_{Tn} | \mathcal{F}_s] \leq E[X_T | \mathbb{1}_{\{X_{Tn} > k\}}, T \leq n] + E[X_n | \mathbb{1}_{\{X_n > k\}}, T > n]$

 $\leq E[X_T | \mathbb{1}_{\{X_T > k\}}] + E[X_n | \mathbb{1}_{\{X_n > k\}}]$ 

由  $(X_n)$ -一致可积,  $\mathbb{P}\{X_T < \infty\} \Rightarrow \sup_n E[X_{Tn} | \mathbb{1}_{\{X_{Tn} > k\}}] \leq E[X_T | \mathbb{1}_{\{X_T > k\}}]$

下证  $E[X_n] < \infty$

 $\sup_n E[X_n] \leq \sup_n E[X_n^+] \leq \sup_n E[X_n] < \infty$ 

由单侧收敛,  $X_{Tn} \xrightarrow{\text{a.s.}} X_T$

Def of 互斥  $\Rightarrow E(X_T | \mathcal{F}_s) = (\exists) X_{SAT}$ , a.s. : P领证  $(X_{Tn})_n$ -一致可积 (由上一定义)

① trivial

②  $\lim E[X_n | \mathbb{1}_{\{T > n\}}] = 0 \Rightarrow E[X_n | \mathbb{1}_{\{T > n\}}] \rightarrow 0$ .

 $E[X_{Tn} | \mathbb{1}_{\{X_{Tn} > M\}}] = E[X_T | \mathbb{1}_{\{X_T > M, T \leq n\}}] + E[X_n | \mathbb{1}_{\{X_n > M, T > n\}}]$ 
 $\leq E[X_T | \mathbb{1}_{\{X_T > M\}}] + E[X_n | \mathbb{1}_{\{T > n\}}] \rightarrow 0$ 

(由  $\lim_{M \rightarrow \infty} \lim_{n \rightarrow \infty} E[X_{Tn} | \mathbb{1}_{\{X_{Tn} > M\}}] = 0$ )  $\Rightarrow (X_{Tn})$ -一致可积.

③  $E[X_{Tn} | \mathbb{1}_{\{X_{Tn} > M\}}] \leq E[X_T | \mathbb{1}_{\{X_T > M\}}] + E[X_n | \mathbb{1}_{\{X_n > M, T > n\}}]$

 $\stackrel{\text{由互斥}}{=} E[X_T] + E[Y_n | \mathbb{1}_{\{Y_n > M\}}} \Rightarrow (X_{Tn})$ -一致可积

④  $\Rightarrow$  ②  $\Rightarrow$  ③:  $|X_T| \leq |X_0| + \sum_{n=1}^T |X_n - X_{n-1}|$

$\sum Y_n = X_n - X_0$ , 则  $E[\sum_{n=1}^T |Y_n|] = E[\sum_{n=1}^T \mathbb{1}_{\{T \geq n\}}]$

 $= \sum_{n=1}^T E[Y_n | \mathbb{1}_{\{T \geq n\}}]$ 
 $= \sum_{n=1}^T E(E(Y_n | \mathbb{1}_{\{T \geq n\}}) | \mathcal{F}_{n-1})$ 
 $= \sum_{n=1}^T E(\mathbb{1}_{\{T \geq n\}} E(Y_n | \mathcal{F}_{n-1}))$ 
 $= E\left(\sum_{n=1}^T E(Y_n | \mathcal{F}_{n-1})\right) < \infty$ 

$\therefore E[X_T] < \infty$ .  $|X_n | \mathbb{1}_{\{T \geq n\}} \leq |X_0| + \sum_{n=1}^T |X_n - X_{n-1}| \in L'$   $\therefore |X_n | \mathbb{1}_{\{T \geq n\}}$ -一致可积, 由②得

Thm:  $X$ : 非负上鞅,  $S \leq T$ ,  $E(X_T | \mathcal{F}_S) \leq X_S$ , a.s. (若  $T = \infty$ ,  $X_T = X_\infty = \lim X_n$ )

pf:  $E(X_{Tn} | \mathcal{F}_S) \leq X_{SAT}$  a.s. (停时)

而  $E(X_T | \mathcal{F}_S) \leq \lim_{n \rightarrow \infty} E(X_{Tn} | \mathcal{F}_S) \leq \lim_{n \rightarrow \infty} X_{SAT} = X_S$ , a.s. #

Thm.  $(X_{Tn})$ -一致可积鞅, 则  $E(X_T | \mathcal{F}_S) = X_{SAT}$  a.s. Rmk. 只要  $\exists n$   $(X_{Tn})$ -一致可积.

lem ( $X_n$ )-一致可积下鞅, 则  $\forall T$ ,  $X_{Tn}$ -一致可积

Rmk.  $(X_n)$ -一致可积鞅,  $X_n \xrightarrow{\text{L'as.}} X_\infty$ ,  $X_\infty = E(X_\infty | \mathcal{F}_n)$ ,  $X_{SAT} = E(X_\infty | \mathcal{F}_{SAT})$ .

pf: 由停时,  $E(X_{Tn} | \mathcal{F}_S) = (\exists) X_{SAT}$ , a.s.

$\therefore X_{SAT} \xrightarrow{\text{L'as.}} X_{SAT}$  (由一致可积)

$E(X_{Tn} | \mathcal{F}_S) \xrightarrow{\text{L'as.}} E(X_T | \mathcal{F}_S)$  (由  $X_{Tn} \xrightarrow{\text{as.}} X_T$  且一致可积)

$\therefore E(X_T | \mathcal{F}_S) = (\exists) X_{SAT}$ , a.s. (取3引理 L'as. 4次).

Thm.  $(X_n)$ -鞅,  $S, T$  停时, 若下述之一成立,

v ①  $X_n$ -一致可积,  $\forall T$

②  $E[X_T] < \infty$ , 且  $\lim E[X_n | \mathbb{1}_{\{T > n\}}] = 0$

③  $E[X_T < \infty]$ , 且  $X_n | \mathbb{1}_{\{T > n\}}$ -一致可积

v ④  $\forall T < \infty$  且  $E[X_{n+1} - X_n | \mathcal{F}_n] \leq B$ , a.s.

⑤  $T$  有限, 且  $E[\sum_{n=1}^T E(|X_n - X_{n-1}| | \mathcal{F}_{n-1})] < \infty$

则  $E(X_T | \mathcal{F}_S) = X_{SAT}$ , a.s.

\*  $X_i \stackrel{iid}{\sim} EX_i = 0$ ,  $EX_i^2 = 1$ ,  $ET < \infty$ , 则  $ES_T^2 = ET$  ( $S_{n \rightarrow \infty}$  为 martingale)

△ 总结: 停时定理  $E(X_T | \mathcal{F}_S) = (\exists) X_{SAT}$  成立条件

①  $T$  有界

或 ②  $(X_n)$ -一致可积.

Eg.  $Y \in L'$ ,  $S, T$  停时, 则  $E(E(Y | \mathcal{F}_S) | \mathcal{F}_T) = E(E(Y | \mathcal{F}_T) | \mathcal{F}_T) = E(Y | \mathcal{F}_{SAT})$

Pf:  $X_n = E(Y | \mathcal{F}_n)$ , 则  $LHS = E(X_T | \mathcal{F}_S) = X_{SAT}$  #.

Eg.  $X \geq 0$  super,  $b_n \geq 0$ , 有  $P(X_n > a, \inf_{k \leq n} X_k = 0) = 0$ .

Pf:  $\sigma \triangleq \inf\{n: X_n = 0\}, b_n \geq 0, \tau \triangleq \inf\{n > \sigma: X_n \geq a\}$ , 则  $\tau = \infty$ , a.s.

$0 \geq EX_{Tn} - EX_{\sigma n} = E(X_T - X_\sigma; \tau \leq n) + E(X_n - X_\sigma; \tau > n)$

 $\geq a \cdot P(\tau \leq n) + 0$ 

$\therefore P(\tau \leq n) = 0 \Rightarrow \tau = \infty$ , a.s. #

PF of Thm: 设  $T = \inf\{k \geq 0, X_k \geq \lambda\}$ , 则  
 $LHS = P(T \leq n) = E[I(T \leq n)] \leq E[X_n^+] I(T \leq n)$   
 $\leq E[X_n] I(T \leq n) \leq E[X_n^+]$

$$\text{设 } \sigma = \inf\{k \geq 0, X_k \geq \lambda\}, \text{ 则 } \\ \lambda P(\max_{0 \leq k \leq n} (-X_k) \geq \lambda) = P(\sigma \leq n) = \lambda E[I(\sigma \leq n)] \leq E(-X_{\sigma \wedge n}) I(\sigma \leq n) \\ = -E X_{\sigma \wedge n} + E X_{\sigma \wedge n} I(\sigma > n) \\ \leq -E X_0 + E X_n I(\sigma > n) \leq E X_n^+ - E X_0$$

PF of  $L^p$ -maximal: 设  $X^* = \max(X_k)$

$$\begin{aligned} \textcircled{1} \quad E(X^*)^p &= P \int_0^\infty x^{p-1} P(X^* \geq x) dx \\ &= P \int_0^\infty x^{p-2} E(X_n^+; X^* \geq x) dx \\ &= P E \left[ X_n^+ \int_0^\infty x^{p-2} 1(X^* \geq x) dx \right] \\ &= \frac{P}{P} E X_n^+ (X^*)^{p-1} \\ &\stackrel{\text{Holder}}{\leq} \frac{P}{P} (E(X^*)^p)^{\frac{1}{p}} (E(X^*)^p)^{\frac{p-1}{p}} \end{aligned}$$

$$\begin{aligned} \textcircled{2} \quad E X^* &= \int_0^\infty P(X^* \geq x) dx \\ &\leq \int_0^\infty \frac{1}{x} E(X_n^+; X^* \geq x) dx + 1 \\ &\leq 1 + E X_n^+ \log^+ X^* \\ (\# \quad a \log^+ b \leq a \log^+ a + \frac{1}{e} b, \forall a, b > 0) \quad &\leq 1 + E Y_n^+ \log^+ X_n^+ + \frac{1}{e} E X^*. \end{aligned}$$

Eg. (Wald 等式)  $X_n$  i.i.d.  $S_n = \sum_{k=1}^n X_k$ .  $T$  停时,  $ET < \infty$ , 则  $EST = EX_1 \cdot ET$ .  
若  $EX_1 = 0$ ,  $EX_1^2 = 1$ , 则  $EST^2 = ET$ .

Pf:  $\{S_n - nEX_1\}$  为 mar. 由  $ET < \infty$ ,  $E(X_n - EX_1)/P_{n+1} \leq E/X_1 - EX_1 / < \infty$   
由弱大数律,  $E(S_T - TEX_1) = 0$

$$\{S_n^2 - n\} \text{ 为 mar. } E(S_{Tn}^2) = E(T \wedge n) \\ \downarrow \\ ET$$

$$\begin{aligned} \therefore E(S_{T(n+1)}^2) &= E(S_T^2 \cdot 1_{T \leq n}) + E(S_{n+1}^2 \cdot 1_{T > n+1}) \\ &= ES_T^2 \cdot 1_{T \leq n} + ES_n^2 \cdot 1_{T > n+1} + E X_{n+1}^2 \cdot 1_{T > n+1} \\ &= ES_{Tn}^2 + P(T \geq n+1) \end{aligned}$$

$$\therefore E(S_{Tn}^2 - S_{Tn}^2)^2 = ES_{Tn}^2 - ES_{Tn}^2 = \sum_{k=n+1}^m P(T \geq k) \rightarrow 0 (n \rightarrow \infty)$$

(由于  $ET = \sum_{k=1}^\infty P(T \geq k) < \infty$ .)

由 Fatou,  $E(S_T^2 - ES_{Tn}^2)^2 = E(S_T - S_{Tn})^2$   
 $\leq \lim_{n \rightarrow \infty} E(S_{Tn}^2 - S_{Tn}^2)^2 \leq \sum_{k=n+1}^\infty P(T \geq k) \rightarrow 0 (n \rightarrow \infty)$

Thm.  $(X_k)$ : sub,  $\mathbb{R}_+$ ,  $\forall \lambda > 0, n \geq 0$ , 有

$$\lambda P(\max_{0 \leq k \leq n} X_k \geq \lambda) \leq E(\max_{0 \leq k \leq n} X_k \geq \lambda) \leq E X_n^+$$

$$\lambda P(\max_{0 \leq k \leq n} X_k \geq \lambda) \leq 2E X_n^+ - E X_0 \leq 3 \max_{0 \leq k \leq n} E|X_k|$$

Rmk. 若  $(X_k)$  mar:  $|X_k|$  sub  $\Rightarrow \lambda P(\max |X_k| \geq \lambda) \leq E|X_n|$

Thm ( $L^p$  最大值不等式)  $(X_n)$ : sub,  $\textcircled{1} P > 1$ , 有  $E(\max_{0 \leq k \leq n} X_k)^p \leq (\max_{0 \leq k \leq n} X_k^p)^{\frac{p}{p-1}} E(X_n^+)^p$   
 $\textcircled{2} p=1, \log^+ x = \log x \vee 0$ , 有  $E(\max_{0 \leq k \leq n} X_k^+) \leq \frac{e}{e-1} (1 + E X_n^+ \log^+ X_n^+)$