

Homework04 - MATH 734

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Notation

Here I use $X \wedge Y$ for $\min(X, Y)$ and $X \vee Y$ for $\max(X, Y)$. r.v. for random variable.

Before Reading:

To make the proof more readable, I will miss or gap some natural or not important facts or notations during my writing. If you feel it hard to see, you can refer the appendix after the proof, where I will try to explain some simple conclusions (will be marked) more clearly. In case that you misunderstand the mark, I will add the mark just after those formulas between \$ and before those between \$\$.

And I have to claim that the appendix is of course a part of my assignment, so the reference of it is required. Enjoy your grading!

Ex.1

Let X_n be a martingale with $EX_n^2 < \infty$ for all n . Show that

- a. If $m \leq n$ and $Y \in \mathcal{F}_m$, then $E[(X_n - X_m)Y] = 0$.
- b. If $l \leq m \leq n$, then $E[(X_n - X_m)(X_m - X_l)] = 0$.
- c. If $m \leq n$, then

$$E[(X_n - X_m)^2 | \mathcal{F}_m] = E(X_n^2 | \mathcal{F}_m) - X_m^2$$

Sol.

- a. We know

$$E[(X_n - X_m)Y] = E[E[(X_n - X_m)Y | \mathcal{F}_m]] = E[Y E[(X_n - X_m) | \mathcal{F}_m]] = 0$$

- b. Let $Y = X_m - X_l \in \mathcal{F}_m$ in (a).

- c. We know

$$\begin{aligned} E[(X_n - X_m)^2 | \mathcal{F}_m] &= E(X_n^2 - 2X_n X_m + X_m^2 | \mathcal{F}_m) \\ &= E(X_n^2 | \mathcal{F}_m) - 2X_m E(X_n | \mathcal{F}_m) + X_m^2 \\ &= E(X_n^2 | \mathcal{F}_m) - X_m^2 \end{aligned}$$

□

Ex.2

Consider supercritical branching process $(Z_n)_{n \geq 0}$ with mean offspring number $\mu = E\xi_n^i > 1$ and suppose $\text{var}(\xi_n^i) = \sigma^2 < \infty$. We know that $\zeta = P(\tau < \infty)$ is nonzero, where τ denotes the extinction time. And we also know that $X_n = Z_n/\mu^n$ converges a.s. to some limiting r.v. X . It is reasonable that on the survival event $\tau = \infty$, X should be positive so the population grows asymptotically exponentially as $X\mu^n$. The goal of this exercise is to justify this: except for a set of probability zero,

$$\{X > 0\} = \{\tau = \infty\}$$

a. Show that

$$E(X_n^2 | \mathcal{F}_{n-1}) = X_{n-1}^2 + \mu^{-2n} E[(Z_n - \mu Z_{n-1})^2 | \mathcal{F}_{n-1}]$$

b. Show that

$$E[(Z_n - \mu Z_{n-1})^2 | \mathcal{F}_{n-1}] = \sigma^2 Z_{n-1}$$

c. Deduce that for all $n \geq 1$,

$$EX_n^2 = EX_{n-1}^2 + \sigma^2 / \mu^{n+1}$$

and by induction, show that

$$EX_n^2 = 1 + \sigma^2 \sum_{k=2}^{n-1} \mu^{-k}$$

d. Show that $X_n \rightarrow X$ in L^2 and $EX_n \rightarrow EX$ for some RV X .

e. Deduce that $EX = 1$, so $\theta = P(X = 0) < 1$. Show that θ satisfies the fixed point equation

$$\theta = \sum_{k=0}^{\infty} p_k k^\theta = \phi(\theta)$$

where $\phi(s) = E(s^{Z_1})$ is the generating function of the offspring distribution. Deduce that $\theta = \zeta = P(\tau < \infty)$ and conclude

$$\{X > 0\} = \{\tau = \infty\}$$

Sol.

a. We know

$$E(X_n^2 | \mathcal{F}_{n-1}) - X_{n-1}^2 = E[(X_n - X_{n-1})^2 | \mathcal{F}_{n-1}] = \mu^{-2n} E[(Z_n - \mu Z_{n-1})^2 | \mathcal{F}_{n-1}]$$

by EX.1.(c) since X_n is a martingale and hence we get

$$E(X_n^2 | \mathcal{F}_{n-1}) = X_{n-1}^2 + \mu^{-2n} E[(Z_n - \mu Z_{n-1})^2 | \mathcal{F}_{n-1}]$$

b. We know

$$\begin{aligned} E[(Z_n - \mu Z_{n-1})^2 \chi_{Z_{n-1}=k} | \mathcal{F}_{n-1}] &= \chi_{Z_{n-1}=k} E[(\sum_{i=1}^k \xi_n^i - k\mu)^2 | \mathcal{F}_n] \\ &= \chi_{Z_{n-1}=k} \sum_{i=1}^k \text{var}(\xi_n^i) \\ &= \chi_{Z_{n-1}=k} k\sigma^2 \\ &= \chi_{Z_{n-1}=k} \sigma^2 Z_{n-1} \end{aligned}$$

and notice

$$(Z_n - \mu Z_{n-1})^2 = (Z_n - \mu Z_{n-1})^2 \sum_{k \geq 0} \chi_{Z_{n-1}=k}$$

and we are done by the MCT.

c. We know

$$\begin{aligned} EX_n^2 &= EX_{n-1}^2 + E[\mu^{-2n} E[(Z_n - \mu Z_{n-1})^2 \chi_{Z_{n-1}=k} | \mathcal{F}_{n-1}]] \\ &= EX_{n-1}^2 + \sigma^2 \mu^{-2n} EZ_{n-1} = EX_{n-1}^2 + \sigma^2 \mu^{-n-1} \end{aligned}$$

and since $EX_0 = 1$ and we may know that

$$EX_n^2 = 1 + \sigma^2 \sum_{i=2}^{n+1} \mu^{-i}$$

by the induction.

d. Obviously we have

$$E[X_n^2] \leq 1 + \frac{\sigma^2}{\mu^2(1-\mu)}$$

and we may use the L^p martingale convergence theorem to X_n and hence $X_n \rightarrow X$ a.s. and in L^2 for some r.v. X . Then we know

$$(E|X_n - X|)^2 \leq E|X_n - X|^2 \rightarrow 0$$

by the Jensen's inequality and hence

$$E|X_n - X| \rightarrow 0, n \rightarrow \infty$$

which means $EX_n \rightarrow EX, n \rightarrow \infty$.

e. Notice $EX_n = 1$ and hence we know $EX = 1$. Obviously if $\theta = 1$ then we know $\mu = 0$ which is a contradiction and hence $\theta < 1$. Then notice

$$\theta = P(\lim_{n \rightarrow \infty} X_n = 0) = \sum_{k=1}^{\infty} P(\lim_{n \rightarrow \infty} X_n = 0 | Z_1 = k) p_k = p_k P(\lim_{n \rightarrow \infty} X_n = 0)^k = \sum_{k=0}^{\infty} p_k k^\theta = \phi(\theta)$$

and hence $\theta = \tau$ since the fixed point is unique on $[0, 1)$, and the required conclusion holds.

Ex.3

In theorem 5.6.11, further assume that $X \in L^p$ for some $p \geq 1$. Conclude that in Theorem 5.6.11., $X_n = E[X | \mathcal{F}_n] \rightarrow E[X | \mathcal{F}_\infty]$ in L^p .

Sol.

We know

$$|E(X | \mathcal{F}_n)|^p \leq E(|X|^p | \mathcal{F}_n)$$

by the Jensen's inequality and hence

$$E(|X_n|^p) \leq E(|X|^p)$$

so we know

$$E(|\bar{X}_n|^p) \leq \left(\frac{p}{p-1}\right)^p E|X_n|^p \leq \left(\frac{p}{p-1}\right)^p E(|X|^p)$$

and hence

$$E(\sup |X_n|)^p \leq \left(\frac{p}{p-1}\right)^p E(|X|^p) < \infty$$

by the MCT and hence

$$\lim_{n \rightarrow \infty} E|X_n - X_\infty|^p = E \lim_{n \rightarrow \infty} |X_n - X_\infty|^p = 0$$

by the DCT since $|X_n - X_\infty|^p \leq 2^p(\sup |X_n|)^p \in L^1$. \square

Ex.4

Let $f : [0, 1) \rightarrow \mathbb{R}$ be a Borel measurable function. In this exercise, we will show that for each $k \geq 1$, there exists a step function g_k with stepsize 2^{-k} such that $\|f - g_k\|_1 \rightarrow 0$ as $k \rightarrow \infty$, a well-known fact in real analysis. We will use a filtration given by dyadic partition and Levy's upward convergence theorem.

a. Fix an integer $L \geq 1$ and denote the intervals $I_{L,i} = [\frac{i-1}{L}, \frac{i}{L})$ for $i = 1, \dots, L$ the partition $[0, 1)$. Let U be an independent Uniform($[0,1)$) r.v. and let \mathcal{F}_L denote the σ -algebra generated by the events $(U \in I_{L,i})$ for $i = 1, \dots, L$. Define

$$f_L = E[f(U)|\mathcal{F}_L]$$

Show that f_L is the block average of f over the interval partition $[0, 1) = I_{L,1} \sqcup \dots \sqcup I_{L,L}$ that is for each $\omega \in I_i$ for $i = 1, \dots, L$

$$f_L(\omega) = \frac{1}{|I_i|} \int_{I_i} f(x) dx$$

b. Now take $L = 2^k$ for $k = 1, 2, \dots$. Show that $(\mathcal{F}_{2^k})_{k \geq 2}$ defines a filtration and that f_{2^k} is a martingale w.r.t. this filtration. Conclude that

$$\|f_{2^k} - f\|_1 \rightarrow 0, k \rightarrow \infty$$

Sol.

a. It suffices to check that

$$\int_{U \in I_{L,i}} f(U) dP = \int_{I_{L,i}} f(x) dx$$

and we know

$$\int_{U \in I_{L,i}} f(U) dP = \int f(U) \chi_{I_{L,i}}(U) dP = \int f(x) \chi_{I_{L,i}}(x) dx = \int_{I_{L,i}} f(x) dx$$

and hence the conclusion holds.

b. Obviously \mathcal{F}_{2^k} is a filtration and

$$E(f_{2^k} | \mathcal{F}_{2^{k-1}}) = \sum_{i=1}^{2^{k-1}} \frac{1}{|I_{2^{k-1},i}|} \chi_{U \in I_{2^{k-1},i}} \int_{I_{2^{k-1},i}} f_{2^k} dx = \sum_{i=1}^{2^{k-1}} \frac{1}{|I_{2^{k-1},i}|} \chi_{U \in I_{2^{k-1},i}} \int_{I_{2^{k-1},i}} f(x) dx = f_{2^{k-1}}$$

and hence f_{2^k} is a martingale w.r.t \mathcal{F}_{2^k} and hence $\|f_{2^k} - E(f(U)|\mathcal{F}_\infty)\|_1 \rightarrow 0$. Since $\mathcal{F}_\infty = \{U \in \mathcal{B}\}$ and hence

$$\|f_{2^k} - f(U)\|_1 \rightarrow 0, k \rightarrow \infty$$

Now we let $g_{2^k} = \int_{i=1}^{2^k} \chi_{I_{2^k,i}} \frac{1}{|I_{2^k,i}|} \int_{I_{2^k,i}} f(x) dx$ and it is easy to check that $g_{2^k}(U) = f_{2^k}$ and we know

$$\int |g_{2^k} - f| dx = \int |g_{2^k}(U) - f(U)| dP \rightarrow 0, k \rightarrow \infty$$

and we are done. \square

Ex.5

A symmetric integrable function $W : [0, 1]^2 \rightarrow [0, 1]$ is called a graphon, a continuum generalization of graphs which also arise as the limit object for sequences of dense graphs. A 'block graphon' is a special graphon that takes constant values over rectangles that partition $[0, 1]^2$. Use the approach in Ex.5.6.15 to show that, for each $k \geq 1$, there exists a block graphon W_k with square blocks of side lengths 2^{-k} such that

$$\|W - W_k\|_1 \rightarrow 0, k \rightarrow \infty$$

Sol.

We know

$$I = \int W(dx dy) = \int \int W(x, y) dx dy = \int \int W(x, y) dy dx$$

by the Fubini's theorem and hence assume

$$g(x) = \int W(u, x) du = \int W(x, u) du$$

and we know $g(x) \in L^1([0, 1])$, and hence $W(x, \cdot) \in L^1([0, 1])$ a.s. for any $\epsilon > 0$ and $y \in [0, 1]$, we may find $g_{2^k}^y$ such that

$$\|g_{2^k}^y - W(\cdot, y)\|_1 < \epsilon$$

and it is easy to check that

$$\int g_{2^k}^y(x) dy = \frac{1}{I_{2^k,i}} \int_{I_{2^k,i} \times [0,1]} W$$

for $y \in I_{2^k,i}$ and hence $g_{2^k}^y(x)$ is in L^1 respect to y and for any $\epsilon > 0$ we may find $\|\phi_{2^m,x}(y) - g_{2^k}^y(x)\|_1 < \epsilon, m \geq k$, then assume $W_k = \phi_{2^m,x}|_{I_{2^m,i} \times I_{2^m,j}}$ if $x \in I_{2^m,j}$, then we know

$$\begin{aligned} \int |W - W_k|(dx dy) &\leq \int \int |W - g_{2^k}^x| dy dx + \int \int |g_{2^k}^x(y) - W_k(x, y)| dx dy \\ &< \epsilon + \epsilon = 2\epsilon \end{aligned}$$

and notice ϵ is arbitrary and we are done. \square

Durrett Ex.4.6.4.

Sol.

For any $\omega \in \{\limsup X_n < \infty\}$, we know

$$M(\omega) < \infty$$

and then there exists $N(\omega)$ such that

$$X_n(\omega) < 2M(\omega)$$

for all $n \geq N(\omega)$. Then

$$P(D|X_1, \dots, X_n)(\omega) \geq \delta(x)$$

for all $n \geq N(\omega)$. Since $D \in \sigma(X_1, X_2, \dots)$ by Theorem 4.6.9., we know LHS converges to $\chi_D(\omega)$ for all $\omega \in \{\limsup X_n < \infty\}$ a.s., then we know

$$\chi_D \geq \delta(x) > 0$$

for all $\omega \in \{\limsup X_n < \infty\} - E$ where $P(E) = 0$. Then $\chi_D(\omega) = 1$ and hence $\omega \in D$, the rest is easy to be checked. \square

Durrett 4.6.5.

Sol.

Notice

$$E(\chi_D|X_1, X_2, \dots, X_n) = \sum_{i=0}^{[x]} \chi_{X_n=i} E(\chi_D|X_1, X_2, \dots, X_n) \geq \sum_{i=0}^n p_0^i > 0$$

so we can let $X_n = Z_n$ and we are done by Ex.4.6.4. \square

Durrett 4.6.6.

Sol.

Notice X_n is a martingale such that

$$E[X_{n+1}|\mathcal{F}_n] = X_n(\alpha + \beta X_n) + (1 - X_n)\beta X_n = X_n$$

Since $|X_n| \leq 1$ for all n , it is uniformly integrable and hence $X_n \rightarrow X$ a.s. and in L^1 for some $X \in L^1$. Then let

$$B_n = \{X_{n+1} = \alpha + \beta X_n\}$$

and $B = \limsup B_n$. For $\omega \in B$,

$$X_{n+1}(\omega) - X_n(\omega) = \alpha(1 - X_n(\omega))$$

and hence X_n converges a.s. Then we know

$$\alpha|1 - X_n| < \epsilon, i.o.$$

and hence

$$X = 1$$

a.s. on B . For $\omega \in B^c$, we know

$$X_{n+1} = \beta X_n$$

which means $X = 0$ on B^c a.s. and hence $X \in \{0, 1\}$.

Since X_n is a martingale, we know $EX_0 = EX_n$, and $EX_n \rightarrow EX$ by the DCT and hence

$$\theta = EX_0 = EX = P(X = 1)$$

□

Durrett 4.6.7.

Sol.

Notice that

$$\begin{aligned} E|E(Y_n|\mathcal{F}_n) - E(Y|\mathcal{F}_\infty)| &\leq E|E(Y_n|\mathcal{F}_n) - E(Y|\mathcal{F}_n)| + E|E(Y|\mathcal{F}_n) - E(Y|\mathcal{F}_\infty)| \\ &\leq E|Y_n - Y| + E|E(Y|\mathcal{F}_n) - E(Y|\mathcal{F}_\infty)| \rightarrow 0, n \rightarrow \infty \end{aligned}$$

by the Jensen's inequality.

□

Durrett 4.8.5.

Sol.

a. By Ex.4.8.3., we know

$$(S_{V_0 \wedge n} - (V_0 \wedge n)(p - q))^2 - (V_0 \wedge n)(1 - (p - q)^2)$$

is a uniformly integrable martingale, and hence

$$(1 - (p - q)^2)EV_0 = E(S_{V_0} - V_0(p - q))^2$$

and notice $p < 1/2$, $V_0 < \infty$ a.s. and $S_{V_0} = 0$, we know

$$(1 - (p - q)^2)EV_0 = (p - q)^2EV_0^2$$

and hence

$$EV_0^2 = \frac{1 - (p - q)^2}{(q - p)^3}x$$

b. We know $EV_0^2 = 0$ where $x = 0$ and it is easy to check that EV_0^2 is linear respect to x by Theorem 4.8.9.

□

Durrett 4.8.6.

Sol.

a. Assume $\phi(\theta) = Ee^{\theta \xi_i}$ and $X_n = e^{\theta S_n} / \phi(\theta)^n$. Then X_n will become a martingale and hence

$$e^{\theta x} = EX_0 = EX_{V_0 \wedge n} = E \frac{e^{\theta S_{V_0 \wedge n}}}{\phi(\theta)^{V_0 \wedge n}}$$

and notice $\theta \leq 0$ and $S_{V_0 \wedge n} \geq 0$, $e^{\theta S_{V_0 \wedge n}} \leq 1$ and $\phi(\theta) \geq 1$. Then we know

$$e^{\theta x} = E \frac{e^{\theta S_{V_0 \wedge n}}}{\phi(\theta)^{V_0 \wedge n}} \rightarrow E \frac{e^{\theta S_{V_0}}}{\phi(\theta)^{V_0}} = E\phi(\theta)^{-V_0}$$

by the DCT.

b. Since $\phi(\theta) = 1/s$, we get $pe^{2\theta} - e^\theta/s + q = 0$ and we have

$$e^\theta = \frac{1 - \sqrt{1 - 4pqs^2}}{2ps}$$

and then

$$Es^{V_0} = \left(\frac{1 - \sqrt{1 - 4pqs^2}}{2ps} \right)^x$$

□

Durrett 4.8.7.

Sol.

Notice that $E(S_{n+1}^4 | \mathcal{F}_n) = S_n^4 + 6S_n^2 + 1$ and $E(S_{n+1}^2 | \mathcal{F}_n) = S_n^2 + 1$. Since $E(Y_{n+1} | \mathcal{F}_n) = Y_n$, we know

$$(2b - 6)n + (b + c - 5) = 0$$

and hence $b = 3, c = 2$. Then by $EY_0 = EY_{T \wedge n}$ and we know

$$0 = ES_{T \wedge n}^4 - 6E(T \wedge n)S_{T \wedge n}^2 + 3E(T \wedge n)^2 + ET \wedge n$$

and hence

$$0 = a^4 - 6a^2 + 3ET^2 + 2ET$$

by the DCT and MCT and then

$$ET^2 = \frac{5a^4 - 2a^2}{3}$$

since $ET = a^2$.

□

Durrett 4.8.11.

Sol.

Assume

$$\phi(\theta) = Ee^{\theta \xi} = e^{((c-\mu)\theta + \sigma^2 \theta^2 / 2)}$$

and notice $\theta_0 = -2(c - \mu)/\sigma^2$ satisfies $\phi(\theta_0) = 1$. For this θ_0 , $X_n = e^{\theta_0 S_n}$ will be a martingale and let $T = \inf \{S_n \leq 0\}$. By Ex.4.8.8. we know $X_{T \wedge n}$ is a uniformly integrable martingale and hence

$$Ee^{\theta_0 S_0} = Ee^{\theta_0 S_T} \geq E(e^{\theta_0 S_T}; T < \infty) \geq P(T < \infty)$$

and let $\theta_0 = -2(c - \mu)/\sigma^2$ will be fine.

□