

## Homework 1

**Due: Wed, Feb. 28, 10PM CT.** Homework should be submitted as a single PDF file via Canvas. Please read the additional instructions in the syllabus. Late homework will **not** be accepted.

[Durrett] refers to the course textbook **Richard Durrett: Probability: Theory and Examples, 5th edition, 2019**

**Exercise 1** ( $k$ th hitting time). Let  $(X_n)_{n \geq 0}$  be a stochastic process defined on a measurable space  $(\mathcal{S}, \mathcal{G})$  adapted to a filtration  $(\mathcal{F}_n)_{n \geq 0}$ . Let  $A \subseteq \mathcal{S}$  be a measurable subset of the state space. For each  $k \geq 1$ , let  $T_A^{(k)}$  denote the  $k$ th time that the process  $X_n$  visits some state in  $A$ . That is,

$$T_A^{(m)} = \begin{cases} \inf\{n > T_A^{(m-1)} : X_n \in A\} & \text{if } T_A^{(m-1)} < \infty \\ \infty & \text{otherwise.} \end{cases}$$

Show that  $T_A^{(k)}$  is a stopping time for all  $k \geq 1$ .

**Exercise 2** (Casino always win). Let  $X = (X_n)_{n \geq 0}$  be a supermartingale w.r.t. a filtration  $\mathcal{F}_n$  and let  $H = (H_n)_{n \geq 1}$  be any predictable sequence w.r.t.  $(\mathcal{F}_n)_{n \geq 1}$ . Suppose that  $H_n$  is bounded and nonnegative for  $n \geq 1$ . Show that  $\int_0^n H dX$  is a supermartingale w.r.t.  $\mathcal{F}_n$ . (*Hint: Mimic the proof of Theorem 5.2.18.*) Also show the similar results for submartingales and martingales. (For the martingale case, it holds without assuming  $H_n \geq 0$ .)

**Exercise 3.** Find an instance of martingale converging in probability but not almost surely.

**Exercise 4.** Use your favorite programming language (e.g., python, R, matlab, C++) and reproduce plots similar to the ones in Figure 5.3.1.

**Exercise 5** (A variational Jensen's inequality). Let  $X$  be a mean zero RV taking values from an interval  $[-A, B]$ . Fix a convex function  $\varphi : \mathbb{R} \rightarrow \mathbb{R}$ . We will show that

$$\mathbb{E}[\varphi(X)] \leq \varphi(-A) \frac{B}{A+B} + \varphi(B) \frac{A}{A+B}. \quad (1)$$

In words, over all possible distributions of  $X$  over  $[-A, B]$ , the most extreme distribution that maximizes  $\mathbb{E}[\varphi(X)]$  is the one that puts point mass on  $-A$  and  $B$  as in the right-hand side.

(i) Let  $Y$  be a RV taking values from  $[0, 1]$  and mean  $p \in [0, 1]$ . Suppose that for any convex function  $\psi : \mathbb{R} \rightarrow \mathbb{R}$ , we have

$$\mathbb{E}[\psi(Y)] \leq (1-p)\psi(0) + p\psi(1). \quad (2)$$

Then deduce (1) from this. (*Hint: Rescale  $X$  and make appropriate change to  $\varphi$ .*)

(ii) Here we will deduce (2). Let  $Y$  be as before. Let  $U \sim \text{Uniform}(0, 1)$  independent from  $Y$ . Argue that

$$\mathbf{1}(U \leq Y) \mid Y \sim \text{Bernoulli}(Y) \quad \text{and} \quad \mathbf{1}(U \leq Y) \sim \text{Bernoulli}(p).$$

(You may use Ex. 5.1.14 for the first part.) Then use Jensen's inequality to deduce

$$(1-p)\varphi(0) + p\varphi(1) = \mathbb{E}[\varphi(\mathbf{1}(U \leq Y))] \geq \mathbb{E}[\varphi(Y)].$$

(iii) (Hoeffding's lemma) Let  $\varphi(x) = e^{\theta x}$  for a fixed  $\theta > 0$  and assume  $A = B > 0$ . Deduce that

$$\mathbb{E}[\exp(\theta X)] \leq \frac{\mathbb{E}[\exp(-\theta A)] + \mathbb{E}[\exp(\theta A)]}{2} \leq \exp(\theta^2 A^2 / 2).$$

**Exercise 6** (Number of triangles in  $G(n, p)$ ). Let  $T = T(n, p)$  denote the total number of triangles in  $G(n, p)$ .

- (i) For each three distinct nodes  $i, j, k$  in  $G$ , let  $Y_{ijk} := \mathbf{1}(ij, jk, ki \in E)$ , which is the indicator variable for the event that there is a triangle with node set  $\{i, j, k\}$ . Show that

$$Y_{ijk} \sim \text{Bernoulli}(p^3).$$

- (ii) Show that we can write

$$T = \sum_{1 \leq i < j < k \leq n} \mathbf{1}(ij, jk, ki \in E). \quad (3)$$

Deduce that the expected number of triangles is

$$\mathbb{E}[T] = \binom{n}{3} p^3.$$

- (iii) Show that

$$\text{Var}(T(n, p)) = \binom{n}{3} (p^3 - p^6) + 12 \binom{n}{4} (p^5 - p^6) \sim \frac{n^4}{2} (p^5 - p^6).$$

(Hint: First compute  $\mathbb{E}[T^2]$  and use the fact that  $\text{Var}(T) = \mathbb{E}[T^2] - \mathbb{E}[T]^2$ . For computing  $\mathbb{E}[T^2]$ , use (3) and consider possible cases according to the number of overlapping edges.) Thus  $\text{Std}(T(n, p)) = \Theta(n^2)$ . If CLT holds for  $T(n, p)$ , then  $T(n, p)$  should fluctuate around its mean by  $\Theta(n^2)$ . Can we conclude this by CLT?

- (iv) Show that for each  $t \geq 0$ ,

$$\mathbb{P} \left( \left| T(n, p) - \binom{n}{3} p^3 \right| \geq t \right) \leq 2 \exp \left( - \frac{t^2}{n(n-1)(n-2)^2} \right).$$

Deduce that the above probability is  $o(1)$  if  $t \gg n^2$ . Specifically, for any  $\varepsilon > 0$ ,

$$\mathbb{P} \left( \left| T(n, p) - \binom{n}{3} p^3 \right| \geq n^{2+\varepsilon} \right) \leq 2 \exp(-n^{2\varepsilon}).$$

Thus, McDiarmid's inequality almost confirms the upper tail of fluctuation of  $T(n, p)$  predicted by CLT. (Hint: Let  $X_1, \dots, X_{\binom{n}{2}}$  denote the indicator of there being an edge for the  $k$ th pair of distinct nodes. Let  $f(X_1, \dots, X_{\binom{n}{2}})$  denote the number of triangles using the edges indicated by  $X_k$ s. Consider the "edge exposure filtration"  $(\mathcal{F}_n)_{0 \leq n \leq \binom{n}{2}}$ , where we reveal the connectedness of every pair of distinct nodes  $(i, j)$  sequentially. Argue that there at most  $n-2$  triangles that contains a given edge. Then use Theorem 5.4.3.)

**Exercise 7** (Durrett). 4.3.1, 4.3.2, 4.3.3, 4.3.4