

# Chapter 1

## 1.1 Basics of Stochastic Processes

We will refer  $X_t$  to be real or  $\mathbb{R}^d$ -valued continuous-time stochastic processes defined on a probability space  $(\Omega, \mathcal{F}, P)$ . For every fixed  $\omega \in \Omega$ ,  $t \mapsto X_t(\omega)$  is called a trajectory or sample path of the process.

For a real-valued stochastic process, let  $-\leq t_1 < \dots < t_n$  be fixed. Then we know

$$P_{t_1, \dots, t_n} = P \circ (X_{t_1}, \dots, X_{t_n})^{-1}$$

is a probability distribution on  $\mathbb{R}^n$ , which is called the finite-dimensional marginal distribution of the process.

### Theorem 1.1

(Kolmogorov's extension theorem) Consider a family of probability measures

$$\{P_{t_1, \dots, t_n}, t_1 < \dots < t_n, n \geq 1, t_i \geq 0\}$$

such that

- $P_{t_1, \dots, t_n}$  is a probability on  $\mathbb{R}^n$ .
- For  $\{t_{k_1} < \dots < t_{k_m}\} \subset \{t_1 < t_2 < \dots < t_n\}$ ,  $P_{t_{k_1}, \dots, t_{k_m}}$  is required to be a marginal of  $P_{t_1, \dots, t_n}$ , then there exists a real-valued stochastic process  $X_t$  owning finite-dimensional marginal distributions of  $\{P_{t_1, \dots, t_n}\}$ .



### Definition 1.1

A real-valued process  $X_t$  is a second-order process iff  $EX_t^2 < \infty, t \geq 0$ , define

$$m_X(t) = EX_t, \Gamma_X(s, t) = \text{cov}(X_s, X_t)$$



### Definition 1.2

A real-valued process  $X_t$  is said to be Gaussian if its finite-dimensional marginal distributions are multidimensional Gaussian laws.



### Proposition 1.1

A Gaussian process is determined by  $m_X$  and  $\Gamma_X$ , conversely, for any  $m : \mathbb{R}_+ \rightarrow \mathbb{R}$  and a symmetric  $\Gamma : \mathbb{R}_+ \times \mathbb{R}_+ \rightarrow \mathbb{R}$  which is nonnegative definite, there always exists a Gaussian process with mean  $m$  and covariance function  $\Gamma$  by Kolmogorov's extension theorem.



### Definition 1.3

We call two processes  $X, Y$  are equivalent if for all  $t \geq 0$ ,  $X_t = Y_t$  a.s. And we call them indistinguishable if  $X_t(\omega) = Y_t(\omega)$  for all  $t \geq 0$  and for all  $\omega$  in some set with probability 1.



### Proposition 1.2

Two equivalent processes with right-continuous trajectories are indistinguishable.



**Proof** Let  $X_q = Y_q$  on  $\Omega_q$  for  $q \in \mathbb{Q}$  and let  $\Omega' = \bigcap_{q \in \mathbb{Q}} \Omega_q$  and we know  $\Omega'$  has the probability 1. And it is easy to check that  $X_t = Y_t$  on  $\Omega'$  for all  $t$ .

### Theorem 1.2


(Kolmogorov's continuity theorem) Suppose that  $X = X_t, t \in [0, T]$  satisfies

$$E(|X_t - X_s|^\beta) \leq K|t - s|^{1+\alpha}$$


for all  $s, t \in [0, T]$  and for some constants  $\beta, \alpha, K > 0$ . Then there exists a version  $\tilde{X}$  of  $X$  such that, if  $\gamma < \alpha/\beta$ ,

then

$$|\tilde{X}_t| - \tilde{X}_s \leq G_\gamma |t - s|^\gamma$$


for all  $s, t \in [0, T]$ , where  $G_\gamma$  is a random variable. The trajectories of  $\tilde{X}$  are Hölder continuous of  $\gamma$  for any  $\gamma < \alpha/\beta$ . 

#### Definition 1.4

$\mathcal{F}_t$  is an increasing family of sub- $\sigma$ -field of  $\mathcal{F}$ . A process  $X_t$  is  $\mathcal{F}_t$ -adapted if  $X_t$  is  $\mathcal{F}_t$ -measurable for all  $t \geq 0$ . 


#### Definition 1.5

An adapted process  $X_t, t \geq 0$  is a Markov process w.r.t. a filtration  $\mathcal{F}_t$  if for any  $s \geq 0, t > 0$  and any measurable and bounded function  $f : \mathbb{R} \rightarrow \mathbb{R}$ ,

$$E(f(X_{s+t})|\mathcal{F}_s) = E(f(X_{s+t})|X_s) \text{ a.s.}$$


#### Proposition 1.3

A  $\mathcal{F}_t$ -Markov process  $X_t$  is also a  $\mathcal{F}_t^X$ -Markov process where

$$\mathcal{F}_t^X = \sigma\{X_u, 0 \leq u \leq t\}$$


**Proof** Notice

$$E(f(X_{s+t})|\mathcal{F}_s^X) = E(E(f(X_{s+t})|\mathcal{F}_s)|\mathcal{F}_s^X) = E(E(f(X_{s+t})|X_s)|\mathcal{F}_s^X) = E(f(X_{s+t})|X_s)$$

since  $\sigma(X_s) \subset \mathcal{F}_s^X \subset \mathcal{F}_t$ .


#### Definition 1.6

Assume a filtration  $\mathcal{F}_t$  on  $(\Omega, \mathcal{F}, P)$  satisfies that for any  $P(A) = 0, A \in \mathcal{F}, A \in \mathcal{F}_0$  and it is right-continuous, i.e.

$$\mathcal{F}_t = \bigcap_{n \geq 1} \mathcal{F}_{t+n^{-1}}$$

Then consider a r.v.  $T : \Omega \rightarrow [0, \infty]$  is a stopping time w.r.t. to the filtration if

$$\{T \leq t \in \mathcal{F}_t\}$$

for any  $t \geq 0$ . 

#### Proposition 1.4

a.  $T$  is a stopping time iff  $\{T < t\} \in \mathcal{F}_t$  for all  $t \geq 0$ .

b.  $S \vee T$  and  $S \wedge T$  are stopping times.

c. Given a stopping time  $T$ ,

$$\mathcal{F}_T = \{A, A \cap \{T \leq t\} \in \mathcal{F}_t, \forall t \geq 0\}$$

is a  $\sigma$ -algebra.

d. If  $S \leq T$ , then  $\mathcal{F}_S \subset \mathcal{F}_T$ .

e. Let  $X_t, t \geq 0$  be a continuous and adapted process. The hitting time of a set  $A \subset \mathbb{R}$  is defined by

$$T_A = \inf\{t \geq 0, X_t \in A\}$$

and whether  $A$  is open or closed,  $T_A$  is a stopping time.

f. Let  $X_t$  be an adapted stochastic process with right-continuous paths and let  $T < \infty$  be a stopping time. Then the random variable

$$X_T(\omega) = X_{T(\omega)}(\omega)$$

is  $\mathcal{F}_T$ -measurable. 

**Definition 1.7**

An adapted process  $M = M_t, t \geq 0$  is called a martingale w.r.t. a filtration  $\mathcal{F}_t, t \geq 0$  if

- a. for all  $t \geq 0$ ,  $E(|M_t|) < \infty$
- b. for each  $s \leq t$ ,  $E(M_t | \mathcal{F}_s) = M_s$

**Proposition 1.5**

- a. For any integrable random variable  $X$ ,  $E(X | \mathcal{F}_t)$  is a martingale.
- b. If  $M_t$  is a submartingale then  $t \rightarrow E(M_t)$  is nondecreasing.
- c. If  $M_t$  is a martingale and  $\varphi$  is a convex function such that  $E|\varphi(M_t)| < \infty$  for all  $t \geq 0$  then  $\varphi(M_t)$  is a submartingale.



**Proof** Only (c) is needed to be proved. Consider  $ax + b \leq \varphi(x)$  and we know

$$E(\varphi(M_t) | \mathcal{F}_s) \geq aE(M_t | \mathcal{F}_s) + b$$

for any such  $a, b$  and hence

$$E(\varphi(M_t) | \mathcal{F}_s) \geq \varphi(M_s)$$

**Definition 1.8**

An adapted process  $M_t, t \geq 0$  is called a local martingale if there exists a sequence of stopping times  $\tau_n \uparrow \infty$  such that, for any  $n \geq 1$   $M_{t \wedge \tau_n}$  is a martingale.

**Theorem 1.3**

Let  $M_t, t \geq 0$  be a continuous local martingale such that  $M_0 = 0$ . Let  $\pi = \{0 = t_0 < t_1 < \dots < t_n = t\}$  be a partition of  $[0, t]$ . Then we have

$$\sum_{j=0}^{n-1} (M_{t_{j+1}} - M_{t_j})^2 \rightarrow \langle M \rangle_t, |\pi| \rightarrow 0$$

in probability, where  $\langle M \rangle_t, t \geq 0$  is called the quadratic variation of the local martingale. Moreover, if  $M_t, t \geq 0$  is a martingale then the convergence holds in  $L^1(\Omega)$ .

**Theorem 1.4**

The quadratic variation is the unique continuous and increasing process satisfying  $\langle M \rangle_0 = 0$  and

$$M_t^2 - \langle M \rangle_t$$

is a local martingale.



## 1.2 Brownian Motion

**Definition 1.9**

A real-valued stochastic process  $B = (B_t)_{t \geq 0}$  defined on a probability space  $(\Omega, \mathcal{F}; P)$  is called a Brownian motion if it satisfies the following conditions:

- a. Almost surely  $B_0 = 0$ .
- b. For all  $0 \leq t_1 < \dots < t_n$  the increments  $B_{t_n} - B_{t_{n-1}}, \dots, B_{t_2} - B_{t_1}$  are independent random variables.
- c. If  $0 \leq s < t$ , the increment  $B_t - B_s$  is a Gaussian random variable with mean zero and variance  $t - s$ .
- d. With probability one, the map  $t \rightarrow B_t$  is continuous.

A  $d$ -dimensional Brownian motion is defined as an  $\mathbb{R}^d$ -valued stochastic process  $B = (B_t)_{t \geq 0}$ ,  $B_t = (B_t^1, \dots, B_t^d)$ , where  $B^1, \dots, B^d$  are  $d$  independent Brownian motions.



**Proposition 1.6**

Properties (a),(b),(c) are equivalent to that  $B$  is a Gaussian process, i.e. for any finite set of indices  $t_1, \dots, t_n$ ,  $(B_{t_1}, \dots, B_{t_n})$  is a multivariate Gaussian random variable, equivalently, any linear combination of  $B_{t_i}$  is normal distributed r.v., with mean zero and covariance function

$$\Gamma(s, t) = \min(s, t)$$

**Proof**

Suppose (a),(b),(c) holds, then we know  $(B_{t_1}, \dots, B_{t_n})$  is normal for any finite indices and then

$$m(t) = E(B_t) = 0$$

$$\Gamma(s, t) = E(B_s B_t) = E(B_{\min(s, t)}^2) = \min(s, t)$$

Conversely, we know  $E(B_0^2) = 0$  and hence  $B_0 = 0$  a.s., then we know  $E(B_s^2) = s$  and for any  $0 < s < t$ ,

$$E(B_s(B_t - B_s)) = 0$$

and it is easy to check (c), and (b) is deduced by computing the covariance of the increments, notice that two r.v.s are independent iff  $\phi_{(X_1, X_2, \dots, X_n)} = \phi_{X_1} \phi_{X_2} \dots \phi_{X_n}$  which implies that normal r.v.s are independent iff they have zero covariances.

**Theorem 1.5**

(Kolmogorov's continuity theorem) Suppose that  $X = (X_t)_{t \in [0, T]}$  satisfies

$$E(|X_t - X_s|^\beta) \leq K|t - s|^{1+\alpha}$$

for all  $s, t \in [0, T]$  and some constant  $\beta, \alpha, K > 0$ . Then there exists a version  $\tilde{X}$  of  $X$  such that if

$$\gamma < \alpha/\beta$$

then

$$|\tilde{X}_t - \tilde{X}_s| \leq G_\gamma |t - s|^\gamma$$

for all  $s, t \in [0, T]$  where  $G_\gamma$  is a random variable. The trajectories of  $\tilde{X}$  are Holder continuous of order  $\gamma$  for any  $\gamma < \alpha/\beta$ .

**Proposition 1.7**

There exists a version of  $B$  with Holder-continuous trajectories of order  $\gamma$  for any  $\gamma < (k-1)/2k$  on any interval  $[0, T]$ .

**Proof**

Since we know  $B_t - B_s$  has the normal distribution  $\mathcal{N}(0, t-s)$  and then we know

$$E((B_t - B_s)^{2k}) = \frac{1}{\sqrt{2\pi(t-s)}} \int_{-\infty}^{\infty} x^{2k} \exp\left(-\frac{x^2}{2(t-s)}\right) dx = (2k-1)!!(t-s)^k = \frac{(2k)!}{2^k k!} (t-s)^k$$

and by the theorem 1.1, the proposition holds.

**Proposition 1.8**

Brownian motion are basic properties:

- For any  $a > 0$ , the process  $(a^{-1/2} B_{at})_{t \geq 0}$  is a Brownian motion.
- For any  $h > 0$ , the process  $(B_{t+h} - B_h)_{t \geq 0}$  is a Brownian motion.
- The process  $(-B_t)_{t \geq 0}$  is a Brownian motion.
- Almost surely  $\lim_{t \rightarrow \infty} B_t/t = 0$  and the process  $X_t = tB_{1/t}$  for  $t > 0$ ,  $X_t = 0$  for  $t = 0$  is a Brownian motion.

**Proof**

a. Consider  $0 \leq t_1 < t_2 < \dots < t_n$  and we may calculate the covariance matrix for

$$a^{-1/2}B_{at_n} - a^{-1/2}B_{at_{n-1}}, \dots, a^{-1/2}B_{at_2} - a^{-1/2}B_{at_1}$$

by

$$\begin{aligned} & E[(a^{-1/2}B_{at_j} - a^{-1/2}B_{at_{j-1}})(a^{-1/2}B_{at_k} - a^{-1/2}B_{at_{k-1}})] \\ &= a^{-1}(at_j \wedge at_k) - a^{-1}(at_j \wedge at_{k-1}) - a^{-1}(at_{j-1} \wedge at_k) + a^{-1}(at_{j-1} \wedge at_{k-1}) \\ &= \begin{cases} t_j - t_{j-1} - t_{j-1} + t_{j-1} = t_j - t_{j-1} & \text{if } j = k \\ t_j - t_j - t_{j-1} + t_{j-1} = 0 & \text{if } j < k \\ 0 & \text{if } j > k \end{cases} \end{aligned}$$

and hence  $(a^{-1/2}B_{at})_{t \geq 0}$  satisfies the property (b) in definition 1.1, a,d are obvious and c is easy to be checked.

b. Obvious.

c. Obvious.

d. Notice  $B$  is Holder continuous. Now we only need to check that

$$E(tB_{1/t}sB_{1/s}) = ts(1/t \wedge 1/s) = (t \wedge s)$$

and the rest is easy to be checked.

### Theorem 1.6

(The law of the iterated logarithm)

$$\limsup_{t \rightarrow s^+} \frac{|B_t - B_s|}{\sqrt{2|t - s| \ln \ln |t - s|}} = 1, \quad a.s.$$



### Proposition 1.9

Fix a time interval  $[0, t]$  and consider the following subdivision  $\pi$  of this interval:

$$0 = t_0 < t_1 < \dots < t_n = t$$

The norm of the subdivision  $\pi$  is defined as  $|\pi| = \max_{0 \leq j \leq n-1} (t_{j+1} - t_j)$ . Then

$$\lim_{|\pi| \rightarrow 0} \sum_{j=0}^{n-1} (B_{t_{j+1}} - B_{t_j})^2 = t$$

in  $L^2(\Omega)$ .



### Proof

Consider let  $\xi_j = (B_{t_{j+1}} - B_{t_j})^2 - (t_{j+1} - t_j)$  and we know  $\xi_j$  are independent with mean 0 and hence

$$\begin{aligned} E\left(\sum_{j=0}^{n-1} (B_{t_{j+1}} - B_{t_j})^2 - t\right)^2 &= \sum_{j=0}^{n-1} E\xi_j^2 = \sum_{j=0}^{n-1} (3(t_{j+1} - t_j)^2 - 2(t_{j+1} - t_j)^2 + (t_{j+1} - t_j)^2) \\ &= 2 \sum_{j=0}^{n-1} (t_{j+1} - t_j)^2 \leq 2t|\pi| \rightarrow 0 \end{aligned}$$

### Proposition 1.10

The total variation of Brownian motion on an interval  $[0, t]$  defined by

$$V = \sup_{\pi} \sum_{i=1}^{n-1} [B_{t_{i+1}} - B_{t_i}]$$

where  $\pi$  is any partition of  $[0, t]$ , is infinite with probability 1.



### Proof

Here we know

$$\sum_{j=0}^{n-1} (B_{t_{j+1}} - B_{t_j})^2 \leq V \sup_j |B_{t_{j+1}} - B_{t_j}|$$

and hence if  $V < \infty$ , then

$$\lim_{|\pi| \rightarrow 0} \sum_{j=0}^{n-1} (B_{t_{j+1}} - B_{t_j})^2 = 0$$

which means  $P(V < \infty) = 0$ .

#### Definition 1.10

(Wiener integral) Let  $\mathcal{E}_0$  be the set of step functions in  $\mathbb{R}_+$ , i.e.

$$\phi(t) = \sum_{j=0}^{n-1} a_j \chi_{t_j, t_{j+1}}(t)$$

where  $n \geq 1$  is an integer,  $a_i \in \mathbb{R}$  and  $0 = t_0 < \dots < t_n$ . And we may define Wiener integral of a step function by

$$\int_0^\infty \phi dB_t = \sum_{i=0}^{n-1} a_i (B_{t_{i+1}} - B_{t_i})$$



#### Proposition 1.11

The Wiener integral is a linear isometry from  $\mathcal{E}_0 \subset L^2(\mathbb{R}^+)$  to  $L^2(\Omega)$ .



**Proof** Notice

$$E[(\int_0^\infty \phi dB_t)^2] = \sum_{i=0}^{\infty} a_i^2 (t_{i+1} - t_i) = \|\phi\|_2^2$$

#### Definition 1.11

We have already know Wiener integral is a linear isometry from a dense subspace from  $L^2(\mathbb{R}_+)$  to  $L^2(\Omega)$ , and hence we may call the extension of the linear isometry to be the Wiener integral and for any  $\phi \in L^2(\mathbb{R}_+)$ , denote

$$\int_0^\infty \phi dB_t$$

to be its image of the isometry.



#### Definition 1.12

Let  $D$  be a Borel subset of  $\mathbb{R}^m$ , a white noise on  $D$  is a centered Gaussian family of random variables

$$\{W_A, A \subset \mathcal{B}(\mathbb{R}^m), A \subset D, m(A) < \infty\}$$

such that

$$E(W_A W_B) = m(A \cap B)$$



#### Proposition 1.12

$\chi_A \rightarrow W_A$  is a linear isometry from  $L^2(D) \rightarrow L^2(\Omega)$ .



#### Definition 1.13

Similarly, we may define the integral r.s.t.  $W$  of  $\phi \in {}^2(D)$  denoted by

$$\phi \mapsto \int_D \phi W(dx)$$

by extending the linear isometry.



**Definition 1.14**

Consider a Brownian motion  $B$  defined on a probability space  $(\Omega, \mathcal{F}, P)$ . For any time  $t \geq 0$ , define  $\mathcal{F}_t$  the  $\sigma$ -algebra by  $B_s, 0 \leq s \leq t$  and the null events in  $\mathcal{F}$ , we call  $\mathcal{F}_t$  the natural filtration of Brownian motion on the probability space  $(\Omega, \mathcal{F}, P)$ .

**Lemma 1.1**

Suppose  $X$  and  $Y$

**Theorem 1.7**

For any measurable and bounded (or nonnegative) function  $f : \mathbb{R} \rightarrow \mathbb{R}, s \geq 0$  and  $t \geq 0$ , we have

$$E(f(B_{s+t})|\mathcal{F}_s) = (P_t f)(B_s)$$

where

$$(P_t f)(x) = \int_{\mathbb{R}} f(y) p_t(x - y) dy$$

where

$$p_t = \frac{1}{\sqrt{2\pi t}} e^{-\frac{x^2}{2t}}$$



Check Durrett Theorem 7.2.1.

**Proposition 1.13**

The family of operators  $P_t$  satisfies the semigroup property  $P_t \circ P_s = P_{t+s}$  and  $P_0 = Id$ .

**Proof**

$$\begin{aligned} P_t \circ P_s(f)(x) &= \int_{\mathbb{R}} P_s f(y) p_t(x - y) dy \\ &= \int_{\mathbb{R}} \int_{\mathbb{R}} f(z) p_s(y - z) p_t(x - y) dz dy \\ &= \int_{\mathbb{R}} f(z) \int_{\mathbb{R}} \frac{1}{2\pi\sqrt{st}} e^{-\left(\frac{(y-z)^2}{2s} + \frac{(x-y)^2}{2t}\right)} dy dz \\ &= \int_{\mathbb{R}} f(z) \int_{\mathbb{R}} \frac{1}{2\pi\sqrt{st}} e^{-\left(\frac{(\sqrt{s+t}y - (2tz+2sx)/\sqrt{s+t})^2 - (tz+sx)^2/(s+t) + tz^2 + sx^2}{2st}\right)} dy dz \end{aligned}$$

and the rest is easy to be checked.

**Theorem 1.8**

The processes  $B_t, (B_t^2 - t)$  and  $e^{aB_t - a^2 t/2}, a \in \mathbb{R}$  are  $\mathcal{F}_t$  martingales.



**Proof**  $B_t$  is obviously a  $\mathcal{F}_t$  martingale. Notice

$$E[(B_t^2 - t)|\mathcal{F}_s] = E[(B_t - B_s)^2 + B_s^2 + 2B_s(B_t - B_s) - t|\mathcal{F}_s] = t - s + B_s^2 - t = B_s^2 - s$$

and

$$E(e^{aB_t - a^2 t/2}|\mathcal{F}_s) = e^{-a^2 t/2} E(e^{a(B_t - B_s)} e^{aB_s}|\mathcal{F}_s) = e^{aB_s} E(e^{a(B_t - B_s)}|\mathcal{F}_s) = e^{aB_s} E(e^{a(B_t - B_s) - a^2(t-s)/2}) = e^{aB_s - a^2 s/2}$$

**Definition 1.15**

The Brownian hitting time is defined by

$$\tau_a = \inf\{t \geq 0, B_t = a\}$$



**Proposition 1.14**

Fix  $a > 0$ . Then, for all  $\alpha > 0$

$$E(e^{-\alpha\tau_a}) = e^{-\sqrt{2\alpha}a}$$



## 1.3 Stochastic Integrals

**Definition 1.16**

We say that a stochastic process  $u_t$  is progressively measurable if, for any  $t \geq 0$  the restriction of  $u$  to  $\Omega \times [0, t]$  is  $\mathcal{F}_t \times B([0, t])$ -measurable.

**Definition 1.17**

Let  $P$  be the  $\sigma$ -field of sets  $A \subset \Omega \times \mathbb{R}^+$  such that  $\chi_A$  is progressively measurable. We denote by  $L^2(P)$  the Hilbert space  $L^2(\Omega \times \mathbb{R}^+, P, P \times m)$  equipped the norm

$$\|u\|^2 = E\left(\int_0^\infty u_s^2 ds\right) = \int_0^\infty Eu_s^2 ds$$

by Fubini's theorem.

**Definition 1.18**

A process  $u = u_t$  is called a simple process if it is of the form

$$u_t = \sum_{j=0}^{n-1} \phi_j \chi_{(t_j, t_{j+1}]}(t)$$

where  $0 \leq t_0 < t_1 < \dots < t_n$  and the  $\phi_j$  are  $\mathcal{F}_{t_j}$ -measurable random variables such that  $E\phi_j^2 < \infty$  and denote the space of simple processes as  $\mathcal{E}$ .

We define the stochastic integral of a process  $u \in \mathcal{E}$  of  $u_t$  as

$$I(u) = \int_0^\infty u_t dB_t = \sum_{j=0}^{n-1} \phi_j (B_{t_{j+1}} - B_{t_j})$$



Here it is easy to see that  $\phi_j \chi_{(t_j, t_{j+1}]}(t)$  is progressively measurable.

**Proposition 1.15**

Here are some properties of stochastic integrals.

a. For any  $a, b \in \mathbb{R}$  and simple process  $u, v \in \mathcal{E}$ , we know

$$\int_0^\infty (au_t + bv_t) dB_t = a \int_0^\infty u_t dB_t + b \int_0^\infty v_t dB_t$$

b. For any  $u \in \mathcal{E}$ , we have

$$E\left(\int_0^\infty u_t dB_t\right) = 0$$

c. For any  $u \in \mathcal{E}$ , we know

$$E\left(\left(\int_0^\infty u_t dB_t\right)^2\right) = E\left(\int_0^\infty u_t^2 dt\right)$$



**Proof** (a) is trivial. (b) can be shown by the independence of  $\phi_j$  and  $B_{t_{j+1}} - B_{t_j}$ . (c) Can be shown by

$$E\left(\left(\int_0^\infty u_t dB_t\right)^2\right) = E\left(\left(\sum_{j=0}^{n-1} \phi_j (B_{t_{j+1}} - B_{t_j})\right)^2\right) = \sum_{j=0}^{n-1} E\phi_j^2 (t_{j+1} - t_j)$$



**Proposition 1.16**

The space  $\mathcal{E}$  of simple processes is dense in  $L^2(P)$ . 

**Proof** For  $u \in L^2(P)$ , we define

$$u_t^{(n)} = n \int_{(t-1/n) \wedge 0}^t u_s ds$$

we know  $u_t^{(n)}$  is continuous as  $\mathbb{R} \rightarrow L^2(\Omega)$  and hence we also know

$$\lim_{n \rightarrow \infty} E \left( \int_0^\infty |u_t - u_t^{(n)}|^2 dt \right) = 0$$

since  $\lim_{n \rightarrow \infty} u_t^{(n)} = u_t$  a.s. by Lebesgue differential, then we may know the limit above holds by the DCT since

$$\int_0^\infty |u_t^{(n)}(\omega)|^2 dt = \int_0^\infty n^2 \left| \int_{t-1/n}^t u_s(\omega) ds \right|^2 dt \leq \|u\|_2^2$$

where  $f(s, t) = u_s \chi_{(t-1/n, t]}$  by the Minkowski's inequality of integrals.

For  $u \in L^2(P)$  continuous in  $L^2(\Omega)$ , we define


$$u_t^{(n, N)} = \sum_{j=0}^{(n, N)} u_{t_j} \chi_{(t_j, t_{j+1}]}(t)$$

where  $t_j = jN/n$ . The continuity in  $L^2(\Omega)$  implies that

$$E \left( \int_0^J |u_t - u_t^{(n, N)}|^2 dt \right) \leq E \left( \int_N^\infty u_t^2 dt \right) + N \sup_{|t-s| \leq N/n, t, s \leq N} E(|u_t - u_s|^2)$$

and we let  $N \rightarrow \infty$  and  $n \rightarrow \infty$  we may get the conclusion.

**Proposition 1.17**

The stochastic integral can be extended to a linear isometry. 


**Proposition 1.18**

Here are some properties, for any  $u, v \in L^2(P)$ , we know

$$E(I(u)) = 0, \quad E(I(u)I(v)) = E\left(\int_0^\infty u_s v_s ds\right)$$

and for any  $T$  we set

$$\int_0^T u_s dB_s = \int_0^\infty u_s \chi_{[0, T]}(s) dB_s$$


which is the indefinite integral of  $u$  with respect to  $B$ , where requiring  $u \in L_T^2(P)$ . 

**Definition 1.19**

Define  $L_\infty^2(P)$  the set of progressively processes such that

$$E \left( \int_0^t u_s^2 ds \right) < \infty$$

for each  $t > 0$ , for any process  $u \in L_\infty^2(P)$ , we can define the indefinite integral process

$$\int_0^t u_s dB_s$$


**Proposition 1.19**

Here are some properties of indefinite integral process.

a. For any  $a \leq b \leq c$ , we have

$$\int_a^b u_s dB_s + \int_b^c u_s dB_s = \int_a^c u_s dB_s$$

b. If  $a < b$  and  $F$  is a bounded and  $\mathcal{F}_a$ -measurable random variable then

$$\int_a^b F u_s dB_s = F \int_a^b u_s dB_s$$

c. Let  $u \in L^2 \infty(P)$ , then the indefinite stochastic integral

$$M_t = \int_0^t u_s dB_s, t \geq 0$$

is a square integrable martingale w.r.t. the filtration  $\mathcal{F}_t$  and admits a continuous version.

**Proof** (a) is trivial. For (b), we only need to consider  $u_t^{(n)}$  simple and converging to  $u_t$  in  $L^2(P)$ , and we have

$$\int_a^b F u_s^{(n)} dB_s = F \sum \phi_{t_j}(B_{t_{j+1}} - B_{t_j}) = F \int_a^b u_s^{(n)} dB_s$$

and we are done since  $F$  is bounded.

(c) For a simple process

$$u_t = \sum_{j=0}^{n-1} \phi_j \chi_{(t_j, t_{j+1}]}(t)$$

then we know

$$\begin{aligned} E\left(\int_0^t u_v dB_v \middle| \mathcal{F}_s\right) &= \sum_{j=0}^{n-1} E(\phi_j(B_{t_{j+1} \wedge t} - B_{t_j \wedge t}) \middle| \mathcal{F}_s) \\ &= \sum_{j=0}^{n-1} E(E(\phi_j(B_{t_{j+1} \wedge t} - B_{t_j \wedge t}) \middle| \mathcal{F}_{t_j \wedge s}) \middle| \mathcal{F}_s) \\ &= \int_0^s u_v dB_v \end{aligned}$$

and hence we know  $M_t$  is an  $\mathcal{F}_t$ -martingale if  $u$  is simple. For  $T > 0$ , let  $u^{(n)}$  converges to  $u$  in  $L_T^2(P)$ , then we know  $t \in [0, T]$ , we have

$$\int_0^t u_s^{(n)} dB_s \rightarrow \int_0^t u_s dB_s$$

in  $L^2(\Omega)$  and we know the convergence of the conditional expectations by  $E(Z(X - E(X|\mathcal{F}))) = 0$  for any  $Z \in L^2(\mathcal{F})$  and hence  $\int_0^t u_s dB_s$  is a martingale.

To show that the indefinite integral has a continuous version. Let  $u \in L^2$  and fix  $T > 0$ . Consider a sequence of simple processes  $u^{(n)}$  which converges to  $u$  in  $L_T^2(P)$ , by the continuity of the paths of Brownian motion, we know

$$M_t^{(n)} = \int_0^t u_s^{(n)} dB_s$$

has continuous trajectories. Then since  $M^{(n)}$  is a martingale, so we know

$$\begin{aligned} P\left(\sup_{0 \leq t \leq T} |M_t^{(n)} - M_t^{(m)}| > \lambda\right) &\leq \frac{1}{\lambda^2} E(|M_t^{(n)} - M_t^{(m)}|^2) \\ &= \frac{1}{\lambda^2} E\left(\int_0^T |u_t^{(n)} - u_t^{(m)}|^2\right) \rightarrow 0 \end{aligned}$$

as  $n, m \rightarrow \infty$  and we may choose  $n_k$  such that

$$P\left(\sup_{0 \leq t \leq T} |M_t^{(n_{k+1})} - M_t^{(n_k)}| > 2^{-k}\right) \leq 2^{-k}$$

and we may know that for any  $\omega \in \Omega$ , there is  $k_1(\omega)$  such that

$$\sup_{0 \leq t \leq T} |M_t^{(n_{k+1})}(\omega) - M_t^{(n_k)}(\omega)| \leq 2^{-k}$$

has probability of 1 and then we know  $M_t^{(n_k)}(\omega)$  is uniformly convergent to a continuous function  $J_t(\omega)$  a.s. and then we know  $J_t(\omega) = \int_0^t u_s dB_s$  a.s. Then we may choose different  $t$  and construct a continuous version of  $\int_0^t u_s dB_s$  inductively.

**Proposition 1.20**

For any  $T, \lambda > 0$  and  $u \in L^2_\infty(P)$ , we know

$$P\left(\sup_{t \in [0, T]} |M_t| > \lambda\right) \leq \frac{1}{\lambda^2} E\left(\int_0^T u_t^2 dt\right)$$

and

$$E\left(\sup_{t \in [0, T]} |M_t|^2\right) \leq 4E\left(\int_0^T u_t^2 dt\right)$$

by Doob's inequality and  $L^p$  maximum inequalities.

**Proposition 1.21**

Let  $u \in L^2_\infty(P)$ . Consider a subdivision of the interval  $[0, t]$

$$\pi = \{0 = t_0 < t_1 < \dots < t_n = t\}$$

then as  $|\pi| \rightarrow 0$ , we have

$$S_\pi^2(u) = \sum_{j=0}^{n-1} \left( \int_{t_j}^{t_{j+1}} u_s dB_s \right)^2 \rightarrow \int_0^t u_s^2 ds$$

in  $L^1(\Omega)$ .



## 1.4 Derivative and Divergence Operators

**Definition 1.20**

For this chapter, we will consider the probability space  $(\Omega, \mathcal{F}, P)$  such that  $\Omega = \mathbb{R}^n$ ,  $\mathcal{F} = \mathcal{B}(\mathbb{R}^n)$  to be the Borel  $\sigma$ -field of  $\mathbb{R}^n$  and  $P$  to be the standard Gaussian probability with density

$$p(x) = (2\pi)^{-n/2} e^{-|x|^2/2}$$

and we define the derivative operator for differentiable function  $F : \mathbb{R}^n \rightarrow \mathbb{R}$  as

$$\nabla F = \left( \frac{\partial F}{\partial x_1}, \dots, \frac{\partial F}{\partial x_n} \right)$$

and the divergence operator for differentiable vector-valued functions  $u : \mathbb{R}^n \rightarrow \mathbb{R}^n$  as

$$\delta(u) = \sum_{i=1}^n \left( u_i x_i - \frac{\partial u_i}{\partial x_i} \right) = \langle u, x \rangle - \text{div} u$$

**Proposition 1.22**

The operator  $\delta$  is the adjoint of  $\nabla$  that is

$$E(\langle u, \nabla F \rangle) = E(F \delta(u))$$

for  $F : \mathbb{R}^n \rightarrow \mathbb{R}$  and  $u : \mathbb{R}^n \rightarrow \mathbb{R}^n$  continuously differentiable functions satisfying integral by parts.



**Proof** Since  $\partial p / \partial x_i = -x_i p$  and we have

$$\begin{aligned} \int \langle \nabla F, u \rangle p dx &= \sum_{i=1}^n \int \frac{\partial F}{\partial x_i} u_i p dx \\ &= \sum_{i=1}^n \left( - \int F \frac{\partial u_i}{\partial x_i} p dx + \int F u_i x_i p dx \right) \\ &= \int F \delta(u) p dx \end{aligned}$$

**Definition 1.21**

Consider the Hilbert space  $h \in H = L^2(\mathbb{R}^+)$  and the Wiener integral

$$B(h) = \int_0^\infty h(t)dB_t$$

and  $S$  the set of smooth and cylindrical random variables of the form

$$F = f(B(h_1), \dots, B(h_n))$$

where  $f \in C_p^\infty(\mathbb{R}^n)$  and  $h_i \in H$ .

**Definition 1.22**

For  $F \in S$ , we define the Malliavin derivative  $DF$  to be the  $H$ -valued random variable defined by

$$D_t F = \sum_{i=1}^n \frac{\partial f}{\partial x_i}(B(h_1), \dots, B(h_n)) h_i(t)$$

which is well-defined and a linear but unbounded operator from  $S$  into  $L^2(\Omega; H)$ .

Let  $S_H$  to be the element  $u = (u_t)$  with the form

$$u_t = \sum_{j=1}^n F_j h_j(t)$$

with  $F_j \in S$  and  $h_j \in H$ . And the divergence of an element  $u$  in  $S_H$  will be given by

$$\delta(u) = \sum_{j=1}^n F_j B(h_j) - \sum_{j=1}^n \langle DF_j, h_j \rangle_H$$

**Proposition 1.23**

For  $F \in S$  and  $u \in S_H$ , then

$$E(F\delta(u)) = E(\langle DF, u \rangle_H)$$



**Proof** Consider  $h_i, h_j$  orthonormal and then we may know that  $B(h_i)$  are i.i.d., then for  $F = f(B(h_1), \dots, B(h_n))$  and

$$u = \sum_{j=1}^n g_j(B(h_1), \dots, B(h_n)) h_j$$

and

$$\begin{aligned} E(\langle DF, u \rangle_H) &= \int \sum_{j=1}^n E(D_t F g_j(B(h_1), \dots, B(h_n)) h_j(t)) \\ &= \sum_{i,j=1}^n \int E\left(\frac{\partial f}{\partial x_i}(B(h_1), \dots, B(h_n)) g_j(B(h_1), \dots, B(h_n)) h_i(t) h_j(t)\right) \\ &= \sum_{i=1}^n E\left(\frac{\partial f}{\partial x_i}(B(h_1), \dots, B(h_n)) g_j(B(h_1), \dots, B(h_n))\right) \\ &= \sum_{j=1}^n E\left(F g_j(B(h_1), \dots, B(h_n)) B(h_j) - F \frac{\partial g_j}{\partial x_j}(B(h_1), \dots, B(h_n))\right) \\ &= E(F\delta(u)) \end{aligned}$$

since

$$\langle DF_j, h_j \rangle = \frac{\partial g_j}{\partial x_j}(B(h_1), \dots, B(h_n))$$

**Proposition 1.24**

Suppose that  $u, v \in S_H$ ,  $F \in S$  and  $h \in H$ . Then for a complete orthonormal system in  $H$  denoted as  $e_i$ , we have

$$E(\delta(u)\delta(v)) = E(\langle u, v \rangle_H) + E\left(\sum_{i,j=1}^{\infty} D_{e_i}\langle u, e_j \rangle_H D_{e_j}\langle v, e_i \rangle_H\right)$$

$$D_h(\delta(u)) = \delta(D_h(u)) + \langle h, u \rangle_H$$

$$\delta(Fu) = F\delta(u) - \langle DF, u \rangle_H$$



**Proof** We know

$$\begin{aligned} D_h(\delta(u)) &= D_h\left(\sum_{j=1}^n F_j B(h_j) - \sum_{j=1}^n \langle DF_j, h_j \rangle_H\right) \\ &= \sum_{j=1}^n D_h F_j B(h_j) + \sum_{j=1}^n F_j \langle h, h_j \rangle_H - \sum_{j=1}^n D_h \left\langle \sum_{i=1}^n \frac{\partial F_j}{\partial x_i} (B(h_1), \dots, B(h_n)) h_i, h_j \right\rangle_H \\ &= \sum_{j=1}^n F_j \langle h, h_j \rangle_H + \sum_{j=1}^n \left( D_h F_j B(h_j) - \langle D_h(DF_j), h_j \rangle_H \right) \\ &= \langle u, h \rangle_H + \delta(D_h u) \end{aligned}$$

by Fubini's Theorem.

Then

$$\begin{aligned} E(\delta(u)\delta(v)) &= E\left(\langle v, D(\delta(u)) \rangle_H\right) \\ &= \sum_{n \geq 1} \left( \langle v, e_i \rangle_H D_{e_i}(\delta(u)) \right) \\ &= E(\langle u, v \rangle_H) + \sum_{i=1}^{\infty} E(\langle v, e_i \rangle_H \delta(D_{e_i} u)) \\ &= E(\langle u, v \rangle_H) + \sum_{i=1}^{\infty} E(\langle D\langle v, e_i \rangle_H, D_{e_i} u \rangle_H) \end{aligned}$$