

Assignment 9 - MA212M

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1 Answer 1

Given:

- $X_1, X_2, \dots, X_n \sim U(\theta_1, \theta_2)$
- $-\infty < \theta_1 < \theta_2 < \infty$

To find: Moment Estimator of (θ_1, θ_2)

The algorithm to find Moment Estimator is as follows:

1. Calculate:

$$\begin{aligned}\mu'_1 &= g_1(\theta_1, \dots, \theta_k) \\ \mu'_2 &= g_2(\theta_1, \dots, \theta_k) \\ &\vdots \\ \mu'_k &= g_k(\theta_1, \dots, \theta_k)\end{aligned}$$

where $\theta = (\theta_1, \dots, \theta_k)$ are the parameters of the distribution from which X is drawn.

2. After calculating $\mu'_j \forall i \in \{1, \dots, k\}$, calculate the inverse functions, i.e.:

$$\begin{aligned}\theta_1 &= h_1(\mu'_1, \dots, \mu'_k) \\ \theta_2 &= h_2(\mu'_1, \dots, \mu'_k) \\ &\vdots \\ \theta_k &= h_k(\mu'_1, \dots, \mu'_k)\end{aligned}$$

3. Now just approximate μ'_j by the Monte Carlo Estimate as following:

$$\alpha_j = \frac{1}{n} \sum_{i=1}^n X_i^j$$

Therefore the Methods of Moments Estimator for θ_i becomes:

$$\hat{\theta}_i = h_i(\alpha_1, \dots, \alpha_k)$$

Now to find the Moment Estimator for (θ_1, θ_2) :

$$\begin{aligned}\mathbb{E}[X^1] &= \mu'_1 = \frac{1}{2}(\theta_1 + \theta_2) \\ \mathbb{E}[X^2] &= \mu'_2 = \int_{-\infty}^{\infty} x^2 f_X(x) dx \\ &= \int_{\theta_1}^{\theta_2} \frac{1}{\theta_2 - \theta_1} x^2 dx \\ &= \frac{1}{\theta_2 - \theta_1} \left[\frac{x^3}{3} \right]_{\theta_1}^{\theta_2} \\ &= \frac{1}{3}(\theta_1^2 + \theta_2^2 + \theta_1\theta_2)\end{aligned}$$

Now, to find the inverse functions h_1 and h_2 :

Manipulating the equations of μ'_1 and μ'_2 to get:

$$\begin{aligned} 4\mu_1^2 - 3\mu_2 &= \theta_1\theta_2 \\ &= \theta_1 \cdot (2\mu_1 - \theta_1) \\ \implies \theta_1^2 - 2\mu_1\theta_1 + 4\mu_1^2 - 3\mu_2 &= 0 \\ \implies \theta_1 &= \frac{2\mu_1 \pm \sqrt{4\mu_1^2 - 4 \cdot (4\mu_1^2 - 3\mu_2)}}{2} \end{aligned}$$

Let $m = \sqrt{4\mu_1^2 - 4 \cdot (4\mu_1^2 - 3\mu_2)}$.

Therefore,

$$\begin{aligned} \theta_1 &= \mu_1 \pm \frac{m}{2} \\ \theta_2 &= \mu_1 \mp \frac{m}{2} \end{aligned}$$

Now, to calculate α_1 and α_2 :

$$\begin{aligned} \alpha_1 &= \frac{1}{n} \sum_{i=1}^n x_i \\ \alpha_2 &= \frac{1}{n} \sum_{i=1}^n x_i^2 \end{aligned}$$

Therefore, the Method of Moments Estimator for (θ_1, θ_2) is (Note that $\theta_2 > \theta_1$):

$$\begin{aligned} \hat{\theta}_1 &= \frac{2\alpha_1 - \sqrt{4\alpha_1^2 - 4(4\alpha_1^2 - 3\alpha_2)}}{2} \\ \hat{\theta}_2 &= \frac{2\alpha_1 + \sqrt{4\alpha_1^2 - 4(4\alpha_1^2 - 3\alpha_2)}}{2} \end{aligned}$$

2 Answer 2

Given:

- $X_1, \dots, X_{10} \sim P_\theta$
- $f_P(x) = \frac{1}{2} \left(\frac{1}{\theta} e^{-\frac{x}{\theta}} + \frac{1}{10} e^{-\frac{x}{10}} \right)$
- $0 < x < \infty$
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$$\sum_{i=1}^{10} X_i = 150.$$

To Find: Method of Moments Estimator for θ .

First, calculating μ'_1 :

$$\begin{aligned}
 \mathbb{E}[X] = \mu'_1 &= \int_{-\infty}^{\infty} x f_P(x) dx \\
 &= \int_0^{\infty} \frac{1}{2} \left(\frac{1}{\theta} e^{-\frac{x}{\theta}} + \frac{1}{10} e^{-\frac{x}{10}} \right) .x .dx \\
 &= \frac{1}{2\theta} \int_0^{\infty} x e^{-\frac{x}{\theta}} dx + \frac{1}{20} \int_0^{\infty} x e^{-\frac{x}{10}} dx \\
 &= \frac{\theta}{2} \Gamma(2) + 5 \Gamma(2) \\
 &= \Gamma(2) \left(\frac{\theta}{2} + 5 \right)
 \end{aligned}$$

Now, finding the inverse function h_1 :

$$\theta = 2 \left(\frac{\mu'_1}{\Gamma(2)} - 5 \right)$$

Approximating μ'_1 by α_1 :

$$\alpha_1 = \frac{1}{10} \sum_{i=1}^{10} x_i = 15$$

Hence, the M.M.E. of θ will be:

$$\begin{aligned}
 \hat{\theta} &= 2 \left(\frac{\alpha_1}{\Gamma(2)} - 5 \right) \\
 &= 2 \left(\frac{15}{2} - 5 \right) \\
 &= 5
 \end{aligned}$$

3 Answer 3

Given:

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$$X_1, X_2, \dots, X_n \sim P_\theta.$$

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$$f_P(x; \theta) = \frac{1}{2} e^{-|x-\theta|} \text{ for } x \in \mathbb{R}.$$

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$$\theta \in \mathbb{R}.$$

To Find: Maximum Likelihood Estimator of θ .

Finding the Likelihood Function $L(\theta)$:

$$L(\theta) = \frac{1}{2^n} \exp \left[- \sum_{i=1}^n |x_i - \theta| \right]$$

To maximize $L(\theta)$ means to minimize $\sum |x_i - \theta|$.

Let's assume that x_1, \dots, x_n are arranged in ascending order. Thus we can formally write the target to minimize as:

$$L_1(\theta) = - \sum_{i=1}^n |x_i - \theta|$$

Now, two cases arise:

- **CASE I:** $\theta < x_1$

If that's the case, then:

$$L_1(\theta) = - \sum_{i=1}^n (x_i - \theta)$$

- **CASE II:** $\theta > x_n$

If that's the case then:

$$L_1(\theta) = \sum_{i=1}^n (x_i - \theta)$$

Remember that we want to minimize the value of $L_1(\theta)$. Therefore, we see now that in CASE I that the function $L_1(\theta)$ is decreasing whereas in CASE II, $L_1(\theta)$ is increasing.

Naturally, there will be a point of minima for θ between $\theta < x_1$ and $\theta > x_n$.

Thus our task now is to find i such that $x_i < \theta < \theta + d < x_{i+1}$ for some $d > 0$ and $i \in \{1, 2, \dots, n-1\}$. Hence, we can write $L_1(\theta + d)$ as:

$$\begin{aligned}
 L_1(\theta + d) &= - \sum_{k=1}^n |x_k - \theta - d| \\
 &= - \sum_{k=1}^i (x_k - \theta - d) + \sum_{k=i+1}^n (x_k - \theta - d) \\
 &= i\theta + id - \sum_{k=1}^i x_k - (n-i)\theta - (n-i)d + \sum_{k=1}^n x_k \\
 &= (2i-n)d - \sum_{k=1}^i (x_k - \theta) + \sum_{k=i+1}^n (x_k - \theta) \\
 &= (2i-n)d + L_1(\theta) \\
 \implies L_1(\theta + d) - L_1(\theta) &= (2i-n)d
 \end{aligned}$$

Therefore,

$$L_1(\theta + d) - L_1(\theta) = \begin{cases} < 0 & \text{if } i < \frac{n}{2} \\ = 0 & \text{if } i = \frac{n}{2} \\ > 0 & \text{if } i > \frac{n}{2} \end{cases}$$

Hence, $L_1(\theta)$ is decreasing for $i < \frac{n}{2}$, constant if $i = \frac{n}{2}$ and increasing for $i > \frac{n}{2}$.

Therefore, if n is even, then the M.L.E. of θ would be any $\theta \in [x_{\frac{n}{2}}, x_{\frac{n}{2}+1}]$.

Whereas, if n is odd, then the M.L.E of θ would be $x_{\frac{n+1}{2}}$.

4 Answer 4

Given:

- Samples drawn from a distribution P_θ

$$\{3, 3, 3, 3, 3, 7, 7, 7\}.$$

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$$P(3) = \theta \text{ and } P(7) = 1 - \theta.$$

To Find: MME and MLE of θ

For Methods of Moments Estimator for θ :

$$\begin{aligned}
 \mathbb{E}[X] &= \mu'_1 = 3\theta + 7(1 - \theta) \\
 &= 7 - 4\theta \\
 \implies \theta &= \frac{7 - \mu'_1}{4}
 \end{aligned}$$

Also,

$$\alpha_1 = \frac{1}{n} \sum_{i=1}^n x_i = \frac{1}{8} \times 36 = 4.5$$

Hence, the M.M.E. for θ would be:

$$\begin{aligned}\hat{\theta} &= \frac{7 - \alpha_1}{4} \\ &= \frac{7 - 4.5}{4} \\ &= \frac{25}{40} = \frac{5}{8} = 0.625\end{aligned}$$

Now, to find the Maximum Likelihood Estimator for θ :

First, the Likelihood Function is:

$$L(\theta) = \theta^5 (1 - \theta)^3$$

We need to maximize this Likelihood Function $L(\theta)$, therefore:

$$\begin{aligned}\frac{d}{d\theta} L(\theta) &= 5\theta^4 (1 - \theta)^3 - 3\theta^5 (1 - \theta)^2 = 0 \\ &= \theta^4 (1 - \theta)^2 [5 - 8\theta] = 0 \\ \implies \theta &= 0, 1, \frac{5}{8}\end{aligned}$$

To find the maximum, we check second order derivative, at the end of which, we see that if $\theta = \frac{5}{8}$, then $\frac{d^2}{d\theta^2} L(\theta = \frac{5}{8}) > 0$.

Hence, M.L.E. of θ is:

$$\hat{\theta} = \frac{5}{8} = 0.625.$$

5 Answer 5

Given:

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$$X_1, X_2, \dots, X_n \sim P_\theta.$$

• The PMF for P_θ is:

$$f(x; \theta) = \begin{cases} \frac{1-\theta}{2} & \text{if } x = 1 \\ \frac{1}{2} & \text{if } x = 2 \\ \frac{\theta}{2} & \text{if } x = 3 \\ 0 & \text{otherwise} \end{cases}.$$

To Find: Maximum Likelihood Estimator of $\theta \in (0, 1)$.

The Likelihood Function is:

$$L(\theta) = \prod_{i=1}^n \left[\left(\frac{1-\theta}{2} \right)^{\frac{(x_i-2)(x_i-3)}{2}} \left(\frac{1}{2} \right)^{-(x_i-1)(x_i-3)} \left(\frac{\theta}{2} \right)^{\frac{(x_i-1)(x_i-2)}{2}} \right]$$

Yikes.....

6 Answer 6

Given:

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$$X_1, X_2, \dots, X_n \sim P_N.$$

- PMF of P_N is:

$$P(X = k) = \begin{cases} \frac{1}{N} & \text{if } k = 1, 2, \dots, N \\ 0 & \text{otherwise} \end{cases}.$$

- N is a positive integer.

To Find: Maximum Likelihood Estimator of N .

Firstly, the Likelihood Function $L(N)$ is:

$$L(N) = \frac{1}{N^n} \quad \text{if } x_1, x_2, \dots, x_n \text{ are positive integers } \leq N$$

We need to maximize the Likelihood Function $L(N)$.

Maximizing $\frac{1}{N^n}$ is equivalent to minimizing N^n if all x_1, x_2, \dots, x_n are positive integers **and** are $\leq N$.

Now, if x_1, x_2, \dots, x_n are positive integers $\leq N$, or, $N \geq x_1, x_2, \dots, x_n$, therefore it's easy to see that N is minimized at the closest value of x_i to N , which is $\max(x_1, x_2, \dots, x_n)$.

Hence, the Maximum Likelihood Estimator \hat{N} is:

$$\hat{N} = \max(x_1, x_2, \dots, x_n)$$

7 Answer 7

Given:

- First M fishes are caught from the pond, tagged and returned to the pond. Next n fishes are caught at random out of which x fishes are found to be tagged.

To Find: Maximum Likelihood Estimator for N , the total number of fishes in the pond.

Since our task was to maximize the number of fishes that we tag, hence we first find the probability of tagging x fishes.

Firstly, the probability of tagging x fishes, where X is the Random Variable denoting the number of fishes tagged so far:

$$P(X = x) = \frac{{}^M C_x {}^{N-M} C_{n-x}}{{}^N C_n}$$

The Likelihood Function for N thus becomes:

$$L(N) = \frac{{}^M C_x {}^{N-M} C_{n-x}}{{}^N C_n} \quad \text{where } N \geq M, \text{ i.e. } N \in \{M, M+1, \dots\}$$

Now, to maximize $L(N)$, calculating $\frac{L(N+1)}{L(N)}$:

$$\begin{aligned} \frac{L(N+1)}{L(N)} &= \frac{{}^M C_x {}^{N+1-M} C_{n-x}}{N+1 C_n} \bigg/ \frac{{}^M C_x {}^{N-M} C_{n-x}}{N C_n} \\ &= \frac{{}^M C_x {}^{N-M+1} C_{n-x} N C_n}{{}^M C_x {}^{N-M} C_{n-x} N+1 C_n} \\ &= \frac{(N-M+1)(N+1-n)}{(N-M+1-n+x)(N+1)} \end{aligned}$$

Now, if the function $\frac{L(N+1)}{L(N)}$ is increasing, then:

$$\frac{L(N+1)}{L(N)} > 1$$

Therefore,

$$\begin{aligned} N^2 - NM + N + N - M + 1 - nN + nM - n &> N^2 - NM + N - nN + xN + N - M + 1 - n + x \\ nM &> xN + x \\ N &< \frac{nM - x}{x} \\ N &< \frac{nM}{x} - 1 \end{aligned}$$

Similarity, if function $\frac{L(N+1)}{L(N)}$ is decreasing, then it would be < 1 , which would transcribe to :

$$N > \frac{nM}{x} - 1$$

Hence, we see that N on left side of $\frac{nM}{x} - 1$, the ratio $\frac{L(N+1)}{L(N)}$ is increasing while decreasing on the right.

Therefore the maximum likelihood should occur at $N = \frac{nM}{x} - 1$

Now, few cases arises on $\frac{nM}{x}$

- **CASE I:** $\frac{nM}{x}$ is an Integer.

If that's the case, then M.L.E. is not unique, as the Likelihood remains maximum for for both $\frac{nM}{x} - 1$ and $\frac{nM}{x}$ as the likelihood Function's value between them is same, but since N has to be an integer therefore, only $\frac{nM}{x} - 1$ and $\frac{nM}{x}$ are taken.

- **CASE II:** $\frac{nM}{x}$ is not an Integer.

Note that $\left\lceil \frac{nM}{x} \right\rceil > \left\lceil \frac{nM}{x} \right\rceil - 1 > \left\lfloor \frac{nM}{x} \right\rfloor - 2$, therefore:

$$L\left(\left\lceil \frac{nM}{x} \right\rceil\right) < L\left(\left\lceil \frac{nM}{x} \right\rceil - 1\right) > L\left(\left\lfloor \frac{nM}{x} \right\rfloor - 2\right)$$

Hence, M.L.E. for N in the case when $\frac{nM}{x}$ is not an integer is $\left\lfloor \frac{nM}{x} \right\rfloor - 1$

8 Answer 8

Given:

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$$X_1, X_2, \dots, X_n \sim P_\theta$$

- PMF of P_θ , the lifetime of the integrated circuit is:

$$f(x, \theta) = \begin{cases} 2\lambda x e^{-\lambda x^2} & \text{if } x > 0 \\ 0 & \text{otherwise} \end{cases}$$

- X be the R.V. denoting the number of integrated circuits that fail before τ , where $\tau > 0$ is a known time.

To Find: Maximum Likelihood Estimator of the variance of X .

To find the probability of number of ICs which fail before τ , we should first find the probability of a I.C. failing before τ units if time:

$$\begin{aligned} P(X_i < \tau) &= \int_0^\tau f(x, \theta) dx \\ &= 2\lambda \int_0^\tau x e^{-\lambda x^2} dx \\ &= 1 - e^{-\lambda \tau^2} \end{aligned}$$

Now, the number of ICs that fail before τ , X is:

$$X \sim \text{Bin}\left(n, 1 - e^{-\lambda \tau^2}\right)$$

Therefore, the variance of X will become:

$$\begin{aligned} \text{Var}(X) &= \mathbb{E}\left[(X - \mathbb{E}[X])^2\right] \\ &= n e^{-\lambda \tau^2} (1 - e^{-\lambda \tau^2}) \end{aligned} \tag{1}$$

Remember that we want to find M.L.E. for (1), thus if we find the M.L.E. for λ , then we can use that to find M.L.E. for (1). Hence, we first try to find M.L.E for λ . The Likelihood function for λ then becomes:

$$L(\lambda) = 2^n \lambda^n \left(\prod_{i=1}^n x_i \right) e^{-\lambda \sum_{i=1}^n x_i^2}$$

To find critical points:

$$\begin{aligned} \frac{d}{d\lambda} L(\lambda) &= n 2^n \lambda^{n-1} \left(\prod_{i=1}^n x_i \right) e^{-\lambda \sum_{i=1}^n x_i^2} - 2^n \lambda^n \left(\prod_{i=1}^n x_i \right) \left(\sum_{i=1}^n x_i^2 \right) e^{-\lambda \sum_{i=1}^n x_i^2} = 0 \\ \implies \lambda &= \frac{n}{\sum_{i=1}^n x_i^2} \end{aligned}$$

Hence, the Maximum Likelihood Estimator for λ will be:

$$\hat{\lambda} = \frac{n}{\sum_{i=1}^n x_i^2}$$

Therefore, we can now use this M.L.E. of λ to find M.L.E. of another function of λ , which in this context is $\text{Var}(X) = ne^{-\lambda\tau^2} (1 - \lambda e^{-\lambda\tau^2})$, which then becomes:

$$\text{Var}(\hat{X}) = ne^{-\hat{\lambda}\tau^2} (1 - \hat{\lambda}e^{-\hat{\lambda}\tau^2})$$

which is the Maximum Likelihood Estimator for the Variance in the number of I.C. failing before time τ .