

Neyman-Pearson Lemma

Animesh Renanse

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1 Most Powerful Test

Def: Consider the **collection C of all level α tests** for $H_0 : \theta \in \Theta_0$ against $H_1 : \theta \in \Theta_1$. A test belonging to C with power function $\beta(\cdot)$ is called **uniformly most powerful (UMP) level α test** if :

$$\beta(\theta) \geq \beta^*(\theta) \text{ for all } \theta \in \Theta_1.$$

where $\beta^*(\cdot)$ is the power function of any other test in C . Note the $\theta \in \Theta_1$ in above equation

If the alternative hypothesis is simple (that means that Θ_1 is singleton), the test is called **most powerful (MP) level α test**.

Remark: Note what we are doing here is that, *across all $\theta \in \Theta_1$, we are finding that test amongst all tests in C with level α which returns maximum value of it's own test function $\beta(\cdot)$* . In this way, we maximize the $\beta(\cdot)$ for all parameters from Θ_1 , which is the parameter space of Alternative Hypothesis. Therefore, in reality, we are maximizing the value of $1 - P(\text{Type-II Error})$, in-turn, **minimizing $P(\text{Type-II Error})$** .

At the same time, we are **keeping a bound on probability of Type-I Error** because level being α , i.e. $P(\text{Type-I Error}) \leq \alpha$

Remark: When $H_1 : \theta = \theta_1$ for some fixed θ_1 , that is, H_1 is simple, it boils down to checking whether $\beta(\theta_1) \geq \beta^*(\theta_1)$ amongst all tests in C . That is the reason why word *uniformly* is removed from the definition when H_1 is simple.

2 Neyman-Pearson Lemma

Theorem: Let $\theta_0 \neq \theta_1$ be two fixed numbers in Θ . The **MP level α test** for $H_0 : \theta = \theta_0$ **against** $H_1 : \theta = \theta_1$ is given by,

$$\varphi(\mathbf{x}) = \begin{cases} 1 & \text{if } L(\theta_1) > kL(\theta_0) \\ \gamma & \text{if } L(\theta_1) = kL(\theta_0) , \\ 0 & \text{if } L(\theta_1) < kL(\theta_0) \end{cases}$$

where $k \geq 0$ and $\gamma \in [0, 1]$ such that $\beta(\theta_0) = \mathbb{E}_{\theta_0}(\varphi(\mathbf{X})) = \alpha$. **Here $L(\cdot)$ is the Likelihood Function.**

Note: This theorem says that if Θ_0 and Θ_1 are singleton sets, then the above $\varphi(\mathbf{x})$ is the most powerful test *automatically*. Since being *most powerful* means that the value of $\beta(\cdot)$ is greater than or equal to the value of all other tests, therefore it has been set to the maximum possible value of $\beta(\cdot)$ under the constraints of the definition, which is α .

Remark: In the theorem, both Null and Alternative Hypotheses are simple.

Remark: $L(\theta_1) > kL(\theta_0)$ can be expressed as $\frac{L(\theta_1)}{L(\theta_0)} > k$ if $L(\theta_0) > 0$. Hence the MP test rejects the null hypothesis for large values of $\frac{L(\theta_1)}{L(\theta_0)}$.

3 Examples

Question: Let $X_1, X_2, \dots, X_n \stackrel{i.i.d.}{\sim} N(\mu, \sigma^2)$, where σ is known. let $\mu_0 < \mu_1$ be two real numbers. We are interested to test $H_0 : \mu = \mu_0$ against $H_1 : \mu = \mu_1$.

First of all,

$$\frac{L(\mu_1)}{L(\mu_0)} = \exp \left[\frac{1}{2\sigma^2} \{2n\bar{x}(\mu_1 - \mu_0) + n(\mu_0^2 - \mu_1^2)\} \right].$$

Now, if $\frac{L(\mu_1)}{L(\mu_0)} > k \iff \bar{x} > k_1$ for some constant k_1 , as $\mu_1 > \mu_0$.

Hence the MP level α test is given by:

$$\varphi(\mathbf{x}) = \begin{cases} 1 & \text{if } \bar{x} > k_1 \\ \gamma & \text{if } \bar{x} = k_1, \\ 0 & \text{if } \bar{x} < k_1 \end{cases}$$

where k_1 and γ are such that,

$$\mathbb{E}_{\mu_0}(\varphi(\mathbf{x})) = \alpha.$$

i.e.,

$$\begin{aligned} \mathbb{E}_{\mu_0}(\varphi(\mathbf{X})) &= P_{\mu_0}[\bar{X} > k_1] = \alpha \\ \implies k_1 &= \mu_0 + \frac{\sigma}{\sqrt{n}} z_\alpha \end{aligned}$$

Now note that $P_{\mu_0}(\bar{X} = k_1) = 0$ as \bar{X} is a (continuous) Normal Distribution. It implies that $\gamma = 0$, **which makes it a non-randomized test.**

Hence the MP level α test is given by:

$$\varphi(\mathbf{x}) = \begin{cases} 1 & \text{if } \bar{x} > \mu_0 + \frac{\sigma}{\sqrt{n}} z_\alpha \\ 0 & \text{otherwise} \end{cases}$$

Remark: Note the way of solving the problem:

- We try to simplify $\frac{L(\mu_1)}{L(\mu_0)} > k$ so that we can *condition on a **statistic whose distribution under H_0 is known or can be found.***
- If this statistic is a continuous random variable, **then we will have a non-randomized test.** Otherwise, we may need to consider $\gamma \in (0, 1)$, making the test a randomized one.

Question: Let $X_1, X_2, \dots, X_n \stackrel{i.i.d.}{\sim} \text{Bernoulli}(\theta)$. Let $0 < \theta_1 < \theta_0 < 1$ be two real numbers. We are interested to test $H_0 : \theta = \theta_0$ against $H_1 : \theta = \theta_1$.

First of all,

$$\frac{L(\theta_1)}{L(\theta_0)} = \left(\frac{\theta_1}{\theta_0} \times \frac{1-\theta_0}{1-\theta_1} \right)^t > k \iff t < k_1$$

where $t \sim T = \sum_{i=1}^n X_i \sim \text{Bin}(n, \theta_0)$ and for some constant k_1 as $\theta_0 > \theta_1$. Hence the MP level α test is:

$$\varphi(\mathbf{x}) = \begin{cases} 1 & \text{if } t < k_1 \\ \gamma & \text{if } t = k_1 \\ 0 & \text{if } t > k_1 \end{cases},$$

where $\mathbb{E}_{\theta_0}(\varphi(\mathbf{x})) = \alpha = P_{\theta_0}(T < k_1) + \gamma P_{\theta_0}(T = k_1)$

Now, take $\hat{K} \in \{1, 2, \dots, n\}$ such that

$$P_{\theta_0}(T < \hat{K}) \leq \alpha < P_{\theta_0}(T \leq \hat{K}).$$

so that if we take $k_1 = \hat{K}$, we get::

$$\gamma = \frac{\alpha - P_{\theta_0}(T < \hat{K})}{P_{\theta_0}(T = \hat{K})}.$$

Therefore, the MP level α test is given by:

$$\varphi(\mathbf{x}) = \begin{cases} 1 & \text{if } t < \hat{K} \\ \frac{\alpha - P_{\theta_0}(T < \hat{K})}{P_{\theta_0}(T = \hat{K})} & \text{if } t = \hat{K} \\ 0 & \text{if } t > \hat{K} \end{cases}.$$

Question: Let $X_1, X_2, \dots, X_n \stackrel{i.i.d.}{\sim} N(\mu, \sigma^2)$, where σ is known. Let μ_0 be a real number. We are interested to test $H_0 : \mu = \mu_0$ against $H_1 : \mu \neq \mu_0$.

We know that the MP level α test for testing $H_0 : \mu = \mu_0$ against $H_1 : \mu = \mu_1 (> \mu_0)$ is given by:

$$\varphi_1(\mathbf{x}) = \begin{cases} 1 & \text{if } \frac{\sqrt{n}}{\sigma}(\bar{x} - \mu_0) > z_\alpha \\ 0 & \text{otherwise} \end{cases}.$$

In the similar way, the MP level α test for testing $H_0 : \mu = \mu_0$ against $H_1 : \mu = \mu_1 (< \mu_0)$ is given by:

$$\varphi_2(\mathbf{x}) = \begin{cases} 1 & \text{if } \frac{\sqrt{n}}{\sigma}(\bar{x} - \mu_0) < -z_\alpha \\ 0 & \text{otherwise} \end{cases}.$$

Now, notice that:

- For $\mu_1 > \mu_0$, the $\varphi_1(\cdot)$ has the maximum power among level α tests.
- For $\mu_1 < \mu_0$, the $\varphi_2(\cdot)$ has the maximum power among level α tests.
- As these two tests are different, therefore UMP level α test for $H_0 : \mu = \mu_0$ against $H_1 : \mu \neq \mu_0$ does not exist for all $\alpha \in (0, 1)$

Remark: There are cases when UMP test does not exist.

Remark: However, the problem of hypothesis testing $H_0 : \mu = \mu_0$ against $H_1 : \mu \neq \mu_0$ is practically quite meaningful. Hence we need some alternatives. One such alternative called **Likelihood Ratio Test** which depends on the concept of Maximum Likelihood Estimator (MLE) is discussed next.