Announcements

- Make sure to sign in on the google form (linked here)
- Midterm October 11

Slope independent of outcome mean

- 1. Find the distribution of \bar{Y} . Recall that $Y_i = \beta_0 + \beta_1 X_i + \epsilon_i$ with $\epsilon_i \sim \mathcal{N}(0, \sigma^2)$.
- 2. Show that $E(\bar{Y}\hat{\beta}_1) = E(\bar{Y})E(\hat{\beta}_1)$. Recall that

$$\hat{\beta}_1 = \frac{\sum_{i=1}^n (X_i - \bar{X})(Y_i - \bar{Y})}{\sum_{i=1}^n (X_i - \bar{X})^2}$$

Hint: Use the fact that $\bar{Y} = \beta_0 + \beta_1 \bar{X} + \bar{\epsilon}$ and think about what is fixed and what is random.

- 3. Find the covariance of \bar{Y} and $\hat{\beta}_1$.
- 4. Apply 7.5.7 from the Stat 110 textbook to show that \bar{Y} and $\hat{\beta}_1$ are independent.

Redundant summary information

Here's a bunch of useful information (also available here, but be careful of what they call p): Definitions:

- Sum of squares model (SSM): $\sum_{i=1}^{n} (\hat{Y}_i \bar{Y})^2$ Sum of squares error (SSE): $\sum_{i=1}^{n} (Y_i \hat{Y}_i)^2$ Sum of squares total (SST): $\sum_{i=1}^{n} (Y_i \bar{Y})^2$
- Degrees of freedom for model with p predictors and an intercept (df_M) : p
- Degrees of freedom for error with p predictors and an intercept (df_E): n-p-1
- Residual standard error: $\sqrt{SSE/df_E}$
- R^2 : 1 SSE/SST
- Adjusted R^{2} : $1 (1 R^{2}) \frac{n-1}{\text{dfg}}$

Facts:

- SSE + SSM = SST
- $\hat{\sigma}^2 = SSE/df_E$
- Under the null (all coefficients are 0),

$$\frac{\mathrm{SSM}/\mathrm{df}_M}{\mathrm{SSE}/\mathrm{df}_E} \sim F_{\mathrm{df}_M,\mathrm{df}_E}$$

From the partial output above, calculate the following:

- 1. How many non-NA data points were included.
- 2. The t-statistics for the intercept and mad_gdppc coefficient.
- 3. The p-values of the two t-tests for the intercept and mad gdppc coefficient being 0.

```
Call:
lm(formula = spi_ospi ~ mad_gdppc, data = countries)
Residuals:
     Min
                    Median
               1Q
                                 3Q
                                         Мах
-61.1775
          -7.3751
                    2.7922
                             9.2159
                                     15.7948
Coefficients:
              Estimate Std. Error t value Pr(>|t|)
(Intercept) 5.9866e+01 1.2971e+00
            4.6565e-04 4.6213e-05
mad_gdppc
Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' '1
Residual standard error: 11.745 on 151 degrees of freedom
  (41 observations deleted due to missingness)
Multiple R-squared: 0.40205,
                                Adjusted R-squared:
F-statistic:
                         and 151 DF, p-value:
```

Figure 1: Lm output with missing information

```
# TODO: intercept
# TODO: mad_gdppc coefficient
```

4. A 95% confidence interval for the mad_gdppc coefficient.

```
# TODO: CI
```

- 5. The adjusted R^2 .
- 6. The sum of squares error, the sum of squares total, and the sum of squares model.
- 7. The f-statistic and p-value for the test that all coefficients are equal to 0.

```
# TODO: f statistic
```

8. Note that the hypothesis tested in $7 (H_0: \beta_1 = 0 \text{ vs } H_a: \beta_1 \neq 0)$ was the same as one of the hypotheses tested in 2. If our framework is consistent, these should give the same answer. Recall from week 2's section that if $T_n \sim t_n$, $T_n^2 \sim F_{1,n}$. Show (numerically) that your calculated t statistic squared is your f statistic, and explain how this shows that the two tests are the same. (Note that this only works because we have a single predictor.)

Regression on real data

This section will deal with a data set of country-level statistics from this source with an explanation of the data encoding found here.

```
countries <- read.csv("data/countries.csv")</pre>
```

1. Fit a linear model to predict the percent of individuals using the internet in a country (wdi_internet) from the log of its GDP per capita (mad_gdppc), and formally test whether this association is significant. Provide a visual to support your conclusion.

```
library(ggplot2)
```

Warning: package 'ggplot2' was built under R version 4.1.3

```
# TODO: Make linear model
# TODO: Make a plot
```

2. Check the assumptions of the model.

```
# TODO: Visualize assumptions
```

- Linearity:
- Constant variance:
- Normality:
- Independence:
- 3. Uganda has a GDP per capita listed but no statistic for internet access. Provide a point estimate and 90% prediction interval.

```
# TODO: Provide prediction
```

How bad are correlated residuals?

Let $Y_i = \beta_0 + \beta_1 X_i + \epsilon_i$ with marginal $\epsilon_i \sim \mathcal{N}(0,1)$ and $\operatorname{Corr}(\epsilon_i,\epsilon_{i+1}) = \rho$ for $i \in \{1,...,n-1\}$ and $\operatorname{Corr}(\epsilon_i,\epsilon_{i-1}) = \rho$ for $i \in \{2,...,n\}$ and $\operatorname{Corr}(\epsilon_i,\epsilon_j) = 0$ otherwise. Write a function to use simulation to find the probability of rejecting the null $H_0: \beta_1 = 0$, the expected value $E(\hat{\beta}_1)$, and the standard deviation $\operatorname{SD}(\hat{\beta}_1)$ in the following situations:

```
library(MASS)
```

Warning: package 'MASS' was built under R version 4.1.3

```
nsims = 1000
n = 10
b0 = 1

run_sim = function(nsims, n, p, b0, b1, sorted=FALSE) {
    # Covariance matrix
    Sigma = matrix(0, nrow = n, ncol = n)
    diag(Sigma) <- 1
    for (i in 2:n) {
        Sigma[i, i-1] <- p
        Sigma[i-1, i] <- p</pre>
```

```
pval = vector(length = nsims)
  coef = vector(length = nsims)
  for (i in 1:nsims) {
     # Generate x
     if (sorted) {
       x \leftarrow sort(rgamma(n, 3, 2/5))
     } else {
       x \leftarrow rgamma(n, 3, 2/5)
     # Generate y with multivariate normal
     y \leftarrow b0 + b1 * x + mvrnorm(n = 1, rep(0, n), Sigma)
     # TODO: Get p-value and coefficient
  # TODO: Return probability of rejecting the null, mean, sd
  return()
   1. n = 10, X_i \sim \text{Gamma}(3, 2/5), \rho = 0, \beta_0 = 1, \beta_1 = 1.
run_sim(nsims, n, 0, b0, 1, sorted=FALSE)
## NULL
   2. n = 10, X_i \sim \text{Gamma}(3, 2/5), \rho = 0.5, \beta_0 = 1, \beta_1 = 1.
run_sim(nsims, n, 0.5, b0, 1, sorted=FALSE)
## NULL
   3. n = 10, X_i \sim \text{Gamma}(3, 2/5) \text{ sorted}, \rho = 0, \beta_0 = 1, \beta_1 = 1.
run_sim(nsims, n, 0, b0, 1, sorted=TRUE)
## NULL
  4. n = 10, X_i \sim \text{Gamma}(3, 2/5) \text{ sorted}, \rho = 0.5, \beta_0 = 1, \beta_1 = 1.
run_sim(nsims, n, 0.5, b0, 1, sorted=TRUE)
## NULL
   5. n = 10, X_i \sim \text{Gamma}(3, 2/5), \rho = 0, \beta_0 = 1, \beta_1 = 0.
```

```
run_sim(nsims, n, 0, b0, 0, sorted=FALSE)
```

NULL

6. $n = 10, X_i \sim \text{Gamma}(3, 2/5), \rho = 0.5, \beta_0 = 1, \beta_1 = 0.$

```
run_sim(nsims, n, 0.5, b0, 0, sorted=FALSE)
```

NULL

7. $n = 10, X_i \sim \text{Gamma}(3, 2/5) \text{ sorted}, \rho = 0, \beta_0 = 1, \beta_1 = 0.$

```
run_sim(nsims, n, 0, b0, 0, sorted=TRUE)
```

NULL

8. $n = 10, X_i \sim \text{Gamma}(3, 2/5) \text{ sorted}, \ \rho = 0.5, \ \beta_0 = 1, \ \beta_1 = 0.$

```
run_sim(nsims, n, 0.5, b0, 0, sorted=TRUE)
```

NULL

9. What conclusions can you draw?