STA 521 HW 2

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a)
TRUE . It is unbiased, since $E(\hat{\beta}) - \beta = 0$
b)
c)
TRUE. JUSTIFY
d)
TRUE . It is a projection matrix, and thus idempotent. $H = H^T$ so it is symmetric, and it is PSD since every $\lambda \geq 0$.

FALSE. tr(I-H) = tr(I) - tr(H) = n - p since the trace of an idempotent matrix

f)

equals it's rank, and H has rank = p.

1

FALSE. We saw an example in class of outliers that had low leverage scores.

g)

answer

h)

 $\textbf{TRUE.} \ Per \ wikipedia: \ https://en.wikipedia.org/wiki/Ggplot2\#:\sim: text = Created\%20 by\%20 Hadley\%20 Wickley Wikipedia.org/wiki/Ggplot2\#:\sim: text = Created\%20 by\%20 Hadley\%20 Wickley Wikipedia.org/wiki/Ggplot2\#:\sim: text = Created\%20 by\%20 Hadley\%20 Wickley Wikipedia.org/wikipe$

2

a)

```
projection = [x_1, 0, 0]^T
```

b)

```
projection = [\boldsymbol{x}_1, \boldsymbol{x}_2, \boldsymbol{0}]^T - just did this visually.
```

c)

```
    \text{part A:} \begin{bmatrix} 1 & 0 & 0 \\ 0 & 0 & 0 \\ 0 & 0 & 0 \end{bmatrix} \\
    \text{part B:} \begin{bmatrix} 1 & 0 & 0 \\ 0 & 1 & 0 \\ 0 & 0 & 0 \end{bmatrix}
```

d)

```
# proj matrices
part_A = matrix(c(1,0,0,0,0,0,0,0,0),3,3)
part_B = matrix(c(1,0,0,0,1,0,0,0,0),3,3,byrow=TRUE)

# making random unif vectors
set.seed(345)
v1 = matrix(runif(3,0,1))
v2 = matrix(runif(3,0,1))

proj1 = part_A %*% v1
proj2 = part_B %*% v2
```

```
proj1
           [,1]
[1,] 0.2162537
[2,] 0.0000000
[3,] 0.0000000
  proj2
           [,1]
[1,] 0.6557397
[2,] 0.4358664
[3,] 0.0000000
e)
                                       \frac{a^Tx}{a^Ta}a
  and the matrix is:
f)
  proj = function(x,a) {
    proj_matrix = (a %*% t(a)) / as.double((t(a) %*% a))
    x_onto_a = proj_matrix %*% x
    x_onto_a
  }
  proj(
    x=matrix(c(3,2,-1)),
    a=matrix(c(1,0,1))
  )
     [,1]
[1,]
        1
[2,]
        0
[3,]
        1
```

g)

$$P = A(A^T A)^{-1} A^T$$

Where the columns of A are a_1, a_2 .

h)

As hinted by the problem, we can use Gram-Schmidt to find an orthonormal basis that spans the same subspace given by a_1, a_2 . With A = QR, and Q being the orthonormal basis, we can construct a projection matrix:

$$P = Q(Q^T Q)^{-1} Q^T$$

i)

```
A = matrix(c(1,0,1,1,-1,0),3,2)
qr_decomp = qr(A)
Q = qr.Q(qr_decomp)

x = matrix(c(3,2,-1))

projection_matrix = Q %*% solve(t(Q) %*% Q) %*% t(Q)
projection_matrix %*% x
```

[,1]

[1,] 1.000000e+00

[2,] 1.665335e-16

[3,] 1.000000e+00

j)

I believe this is the same as my answer in h) (and looking ahead to part k) we can construct a projection matrix onto the k-dimensional subspace spanned by $a_1, ..., a_k$ by computing the QR decomposition, taking the Q, and constructing:

$$P = Q(Q^T Q)^{-1} Q^T$$

1)

The answer below looks the same as part i).

A %*% solve(t(A)%*% A) %*% t(A) %*% x

[,1]

[1,] 1.000000e+00

[2,] -5.551115e-17

[3,] 1.000000e+00

m)

$$P = Q(Q^TQ)^{-1}Q^T = Q(I)Q^T = QQ^T = I$$

3

a)

Joint Distributions:

$$p(X = x, Z = 1; \theta) = \pi_1 N(x; \mu_1, 1) p(X = x, Z = 2; \theta) = \pi_2 N(x; \mu_2, 1)$$

Marginal likelihood:

$$p(X=x;\theta) = \sum_{i=1}^2 p_{\theta}(X=x|Z=j) p_{\theta}(Z=j)$$

log-likelihood:

$$l(X=x;\theta) = \sum_{i=1}^{N} log p_{\theta}(X=x_i)$$

w/ n i.i.d. samples:

$$l(X=x;\theta) = -\frac{n}{2}ln(2\pi) - \frac{1}{2}\sum_{i=1}^{n}(x_i - \mu)^2$$

b)

c)

d)

4

1.

We know the distribution of $\hat{\beta}$ must be normal because we assumed the errors are normally distributed. (Why?) Then it is sufficient to find the mean and variance of $\hat{\beta}$

$$\begin{split} E(\hat{\beta}) &= E[(X^TX)^{-1}X^Ty] \\ &= (X^TX)^{-1}X^TE[x\beta^* + \epsilon] \to \text{since X is constant} \\ &= (X^TX)^{-1}X^TXE(\beta^*) \\ &= \beta^* \end{split}$$

(Note for self: below mirrors form of $Var(X) = E(X - E(X))^2$

$$\begin{split} Var(B^*) &= E[(\hat{\beta} - E(\hat{\beta}))(\hat{\beta} - E(\hat{\beta}))^T] \\ &= E[(\hat{\beta} - (X^TX)^{-1}X^T\epsilon)(\hat{\beta} - (X^TX)^{-1}X^T\epsilon)^T] \\ &\text{next line follows since} E(\hat{\beta}) = \beta \text{and} X^TX \text{is symmetric} \\ &= E[(X^TX)^{-1}X^T\epsilon\epsilon^TX(X^TX)^{-1}] \\ &= (X^TX)^{-1}X^TE(\epsilon\epsilon^T)X(X^TX)^{-1} \\ &= \sigma^2(X^TX)^{-1} \end{split}$$

Noting that $E(\epsilon \epsilon^T)$ is the covariance matrix of ϵ with mean = 0. Thus we have the mean and the variance of $\hat{\beta}$ and it is distributed $\sim N(\beta^*, \sigma^2(X^TX)^{-1})$ Now, since the data X is constant, $E(X\hat{\beta}) = E(X)E(\hat{\beta}) = X\beta^*$

2.

$$\begin{split} E||e||_2^2 &= ETr(ee^T) \\ &= ETr((I_n - X(X^TX)^{-1}X^T)\epsilon\epsilon^T(I_n - X(X^TX)^{-1}X^T)) \\ &= \sigma^2 Tr(I_n - X(X^TX)^{-1}X^T) \\ \text{below using the cyclic property of trace} \\ &= \sigma^2 Tr(I_n) - Tr(X^TX(X^TX)^{-1}) \\ &= \sigma^2 Tr(I_n) - Tr(I_p) \\ &= \sigma^2 (n-p) \end{split}$$

3. optional

 5

 1.

 2.

 3.

 4.

 5.

 6.

4. optional