# 602\_hw2

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#### 3.1

a)

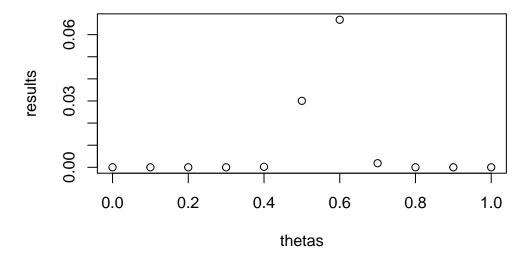
$$\begin{split} &P(Y_1=y_1,\dots,Y_{100}=y_{100}|\theta)=\text{by independence}\\ &=\prod_{i=1}^n P(Y_i|\theta)\\ &=\prod_{i=1}^n \theta^{y_i}(1-\theta)^{1-y_i}\\ &=\theta^{\Sigma_{i=1}^n y_i}(1-\theta)^{100-\Sigma_{i=1}^n y_i};y=0,1 \end{split}$$

Finding the distribution of  $P(\sum_{i=1}^n Y_i = y|\theta)$ 

$$\begin{split} &M_{\sum Y_i=y|\theta}(t)\\ &=\text{by independence}\\ &=\prod_{i=i}^n M_{Y_i|\theta}(t)=\\ &=\prod_{i=i}^n (1-p+pe^t)\\ &=(1-p+pe^t)^n\\ &=\binom{n}{x}\theta^x(1-\theta)^{n-x}\\ &=\binom{100}{57}\theta^{57}(1-\theta)^{43};\theta\in[0,1] \text{assuming a uniform prior} \end{split}$$

b)

```
thetas = seq(0.0,1.0,by=0.1)
results = dbinom(57,100,thetas)
plot(thetas,results)
```

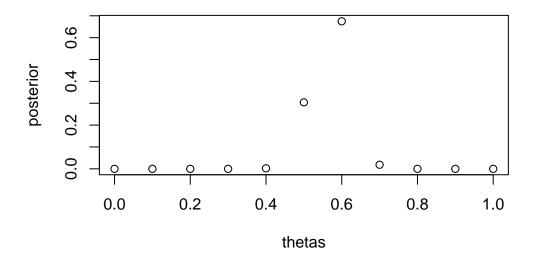


c)

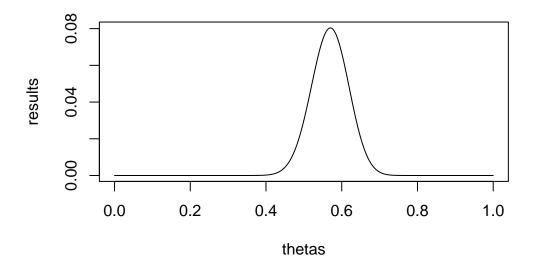
$$p(\theta|\Sigma_{i=1}^{n}y_{i}=57) = \frac{P(\Sigma_{i=1}^{n}y_{i}=57|\theta)P(\theta)}{P(\Sigma_{i=1}^{n}y_{i}=57)} \text{each } P(\Theta=\theta) = \frac{1}{11}$$

The posterior distribution and marginal distribution of Y are just scaling constants since the denominator does not depend on theta and we have equal belief for each of  $P(\theta)$ .

```
marginal_y = sum((1/11) * dbinom(57,100,thetas))
posterior = (results * (1/11))/marginal_y
plot(thetas,posterior)
```

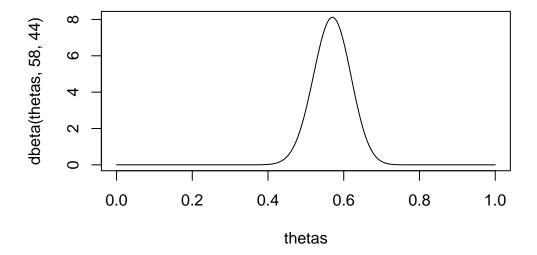


d)
thetas = seq(0,1,by=0.001) #U(0,1)
results = dbinom(57,100,thetas)
plot(thetas,results,type="l")



e)

plot(thetas,dbeta(thetas,58,44),type='1')



### 3.2

```
theta0 = seq(0.1,0.9,by=0.1)
n0 = c(1,2,8,16,32)

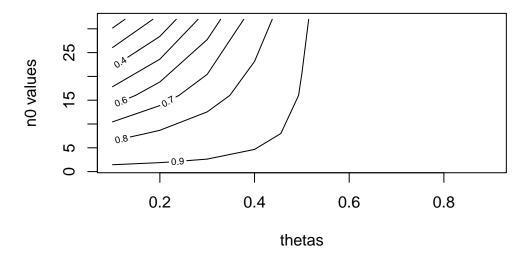
data=c()

for (i in theta0) {
    for (j in n0) {
        a = i * j
        b = (1-i)*j

        p = pbeta(.5,a+57,b+43,lower.tail=FALSE) #posterior (theta > .5 | sum = 57)
        data = append(data,p)

}

probability_data = matrix(data,nrow=9,ncol=5,byrow=TRUE)
contour(theta0,n0,probability_data, xlab="thetas",ylab='n0 values')
```



3.4

$$p(\theta|y) = \frac{p(y|\theta)p(\theta)}{p(y)}$$

a)

calculations for posterior with prior beta(2,8) and beta(8,2) are here, with plots for part a) and part b) following this chunk.

```
beta_mean = function(a,b){
  print("mean:")
  a / (a+b)
}

beta_mode = function(a,b){
  print('mode:')
  (a-1) / (a+b-2)
}

beta_sd = function(a,b){
  print('standard deviation:')
  var = (a*b) / ((a+b)^2 * (a+b+1))
  sd = sqrt(var)
  return(sd)
}
```

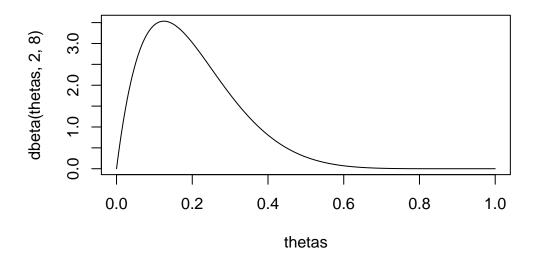
```
CI_28 = c(qbeta(.025,17,36),qbeta(.975,17,36))
  CI_82 = c(qbeta(.025,23,30),qbeta(.975,23,30))
  #data for the posterior w/2,8 prior and posterior a = 17, posterior b = 36
  print("using alpha = 17 and beta = 36 with beta(2,8) prior")
[1] "using alpha = 17 and beta = 36 with beta(2,8) prior"
  beta_mean(17,36)
[1] "mean:"
[1] 0.3207547
  beta_mode(17,36)
[1] "mode:"
[1] 0.3137255
  beta_sd(17,36)
[1] "standard deviation:"
[1] 0.0635189
  print(c("95% CI",CI_28))
[1] "95% CI"
                        "0.203297787819103" "0.451023982216632"
  #with 8,2 prior
  print(" ")
[1] " "
```

```
[1] "-----"
  print("using alpha = 23, beta = 30 with beta(8,2) prior")
[1] "using alpha = 23, beta = 30 with beta(8,2) prior"
  beta_mean(23,30)
[1] "mean:"
[1] 0.4339623
  beta_mode(23,30)
[1] "mode:"
[1] 0.4313725
  beta_sd(23,30)
[1] "standard deviation:"
[1] 0.06744532
  print(c("95% CI",CI_82))
[1] "95% CI"
                 "0.304695624711747" "0.567952795996458"
plots for part a)
```

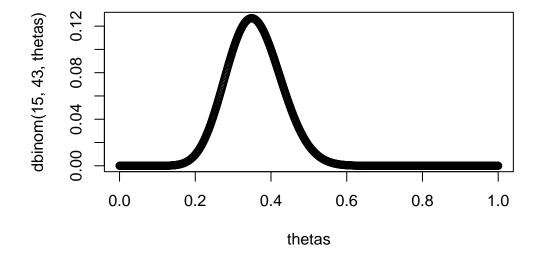
```
#plotting prior p(\theta)
thetas = seq(0,1,by=0.001) #U(0,1)

plot(thetas, dbeta(thetas, 2,8), type='l',main="p(theta)")
```

## p(theta)



#plotting p(y=15|\theta)
#plot a binomial here
plot(thetas, dbinom(15,43,thetas))



```
#posterior which is beta(2 + success, 8 + failure) = beta()
a=2+15
b=8+28
plot(thetas, dbeta(thetas,a,b),type='l',main="posterior model")
abline(v=beta_mean(a,b), col='red') #mean

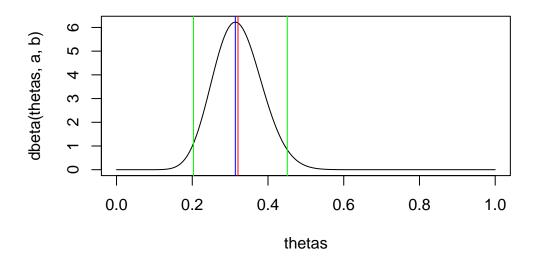
[1] "mean:"

abline(v=beta_mode(a,b), col='blue') #mode

[1] "mode:"

# CI
abline(v=qbeta(.975,a,b),col='green') #lower bound
abline(v=qbeta(.025,a,b),col='green') #upper bound
```

### posterior model

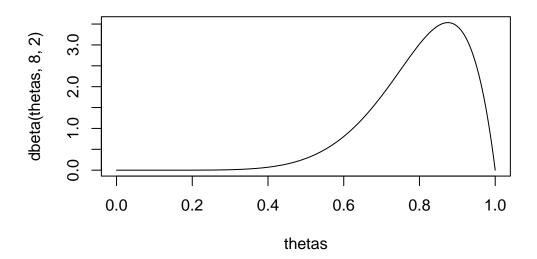


```
plots for part b)

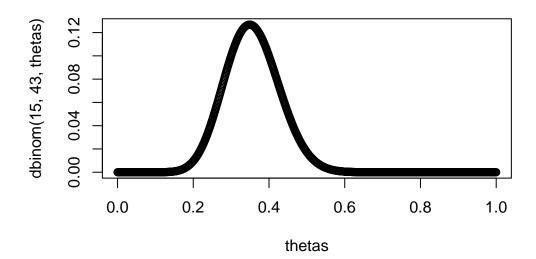
#plotting prior p(\theta)
thetas = seq(0,1,by=0.001) #U(0,1)

plot(thetas, dbeta(thetas, 8,2), type='l',main="p(theta)")
```

## p(theta)



```
#plotting p(y=15|\theta)
#plot a binomial here
plot(thetas, dbinom(15,43,thetas))
```



```
#posterior which is beta(2 + success, 8 + failure) = beta()
a=8+15
b=2+28
plot(thetas, dbeta(thetas,a,b),type='l',main="posterior model")
abline(v=beta_mean(a,b), col='red') #mean
```

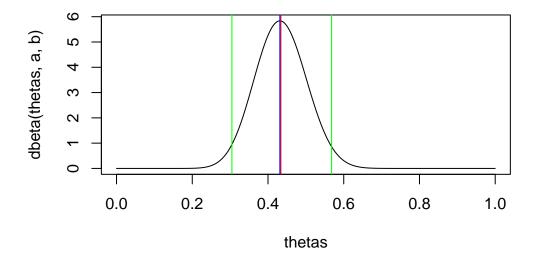
```
[1] "mean:"

abline(v=beta_mode(a,b), col='blue') #mode

[1] "mode:"

# CI
abline(v=qbeta(.975,a,b),col='green') #lower bound
abline(v=qbeta(.025,a,b),col='green') #upper bound
```

#### posterior model

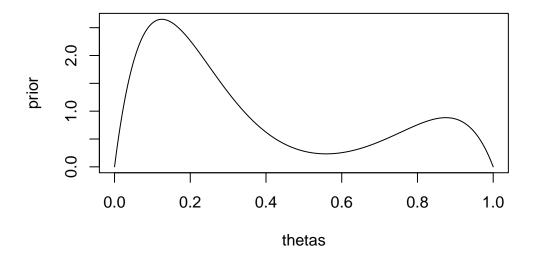


c)

This may represent that you have about 25% confidence that there are going to be 8 cases of recidivism and 2 cases of not, while the beta(2,8) represents you're 75% confident that there will be 2 cases of recidivism and 8 cases of failure respectively. This is if you've only seen 10 prior cases.

Or maybe there were two previous studies with 2 successes and 8 failures or 2 failures and 8 successes respectively.

```
prior = 0.75 * dbeta(thetas,2,8) + 0.25 * dbeta(thetas,8,2)
plot(thetas,prior,type="1")
```

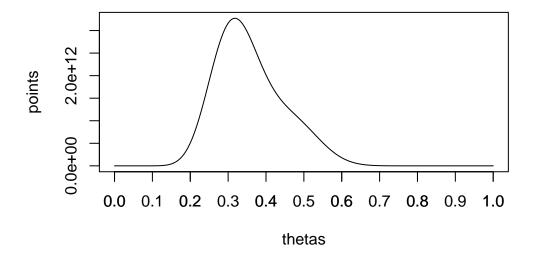


d) i)

$$\begin{split} &p(\theta)*p(y|\theta)\\ &=\frac{1}{4}\frac{\Gamma(10)}{\Gamma(2)\Gamma(8)}{43 \choose 15}[3\theta^{16}(1-\theta^{35})+\theta^{22}(1-\theta)^{25}] \end{split}$$

- ii) This is a mixture of beta(17,36) and beta(23,26)
- iii) Plot of  $p(\theta|y)$  is below. It looks like the mode is about 0.32 (approximately). Since this is more heavily weighted towards the prior of beta(2,8), it makes sense that this mode is closer to the mode of the previous example we saw with the same prior, though pulled slightly to the right by the beta(8.2) prior.

```
coefficient = .25 * 18 * choose(43,15)
thetas = seq(0,1,by=0.001)
points = coefficient* (0.75 * dbeta(thetas,17,36) + 0.25 * dbeta(thetas,23,26))
plot(thetas,points,type="l")
axis(1, at = seq(0.0,1,by=0.1))
```



e)