FORECASTING ECON 492

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August 15, 2018

Models vs Reality

Many macro models have the following **autoregressive** form:

$$x_t = \rho_0 + \rho_1 x_{t-1}$$

x and ρ_0 can be vectors and ρ_1 can be a matrix

Example: neo-classical growth model

► capital accumulation equation:

$$k_{t+1} = Ak_t^{\alpha} \quad \Rightarrow \quad \underbrace{\ln k_{t+1}}_{x_t} = \underbrace{\ln A}_{\rho_0} + \underbrace{\alpha \ln k_t}_{\rho_1 x_{t-1}}$$

Models vs Reality

Many macro models have the following autoregressive form:

$$x_t = \rho_0 + \rho_1 x_{t-1}$$

x and ρ_0 can be vectors and ρ_1 can be a matrix

Example: two-state labor market model

▶ employment/unemployment dynamics:

$$\left[\begin{array}{c} e_t \\ u_t \end{array}\right] = \left[\begin{array}{cc} 1-s & f \\ 1-f & s \end{array}\right] \left[\begin{array}{c} e_{t-1} \\ u_{t-1} \end{array}\right]$$

Models vs Reality

Many macro *models* have the following autoregressive form:

$$x_t = \rho_0 + \rho_1 x_{t-1}$$

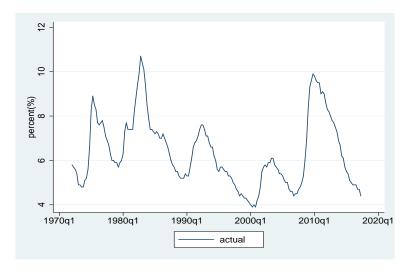
Reality is more accurately described as a vector autoregression (VAR):

$$x_t = \rho_0 + \rho_1 x_{t-1} + \rho_2 x_{t-2} + \rho_3 x_{t-3} + \dots + \epsilon_t$$

- ▶ system may be higher than just first-order autoregressive
- system is hit with shocks
- we don't know the values of ρ_0, ρ_1, \ldots
- \triangleright if x_t is **stationary**, we can estimate the VAR using OLS!

DATA TRANSFORMATIONS

- ► Seasonally adjusted?
- ► Transform quantities in logs?
- ► Stationary?



Civilian Unemployment Rate, 1972q1-2017q2. Source: fred.stlouisfed.org.

STATA commands:

insheet using "data20170923.csv", clear

▶ loads the data from data20170923.csv of the Excel file

```
generate qtr = tq(1967q1) + time - 1
format qtr %tq
```

▶ uses the variable "time" (=1,2,...) and creates new variable "qtr" (=1967q1,1967q2,...)

tsset qtr

▶ tells STATA that the data is time series data, and "qtr" defines the time dimension

Unit root tests

Formal tests are developed to test for non-stationarity Dickey-Fuller (DF) and augmented DF are examples of such tests

- ▶ DF runs: $y_t = \alpha + \rho y_{t-1} + \delta t + \epsilon_t$
 - H_0 is $\rho = 1$
- Augmented DF controls for possible serial correlation: $\Delta y_t = \alpha + \beta y_{t-1} + \delta t + \psi_1 \Delta y_{t-1} + \psi_2 \Delta y_{t-2} + \dots + \epsilon_t$
 - H_0 is $\beta = 0 \Leftrightarrow \rho = 1$

var unemp if qtr > = tq(1972q1), lags(1 2)

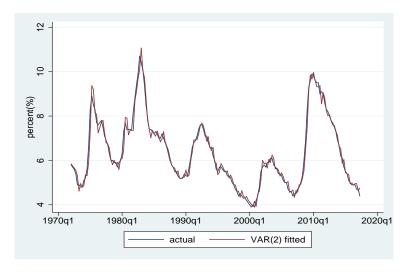
▶ runs an AR(2) on "unemp" using data starting in 1972q2

```
. var unemp if qtr>=tq(1972q1), lags(1/2)
```

```
Vector autoregression
Sample: 1972q1 - 2017q2
Log likelihood = -9.110109
                                               ATC
                                               HQIC
Det(Sigma_ml) = .0647148
                                               SBIC
                                                     P>chi2
Equation
                 Parms
                            RMSE
                                    R-sq
                                             chi2
                    3
                          .256514
                                   0.9731
unemp
                                            6582.07 0.0000
```

| | unemp | Coef. | Std. Err. | z | P> z | [95% Conf | . Interval] |
|-------|-------|----------|-----------|--------|--------|-----------|-------------|
| unemp | | | | | | | |
| | unemp | | | | | | |
| | L1. | 1.636853 | .0549505 | 29.79 | 0.000 | 1.529152 | 1.744554 |
| | L2. | 676608 | .0551195 | -12.28 | 0.000 | 7846401 | 5685758 |
| | _cons | .2497777 | .0806643 | 3.10 | 0.002 | .0916786 | .4078768 |

So: $u_t = 0.250 + 1.637u_{t-1} - 0.677u_{t-2}$, $R^2 = 0.973$



In forecasting we care about "fit" (i.e. \mathbb{R}^2), not precision (std err)

We have unemployment data that ends in 2017q2

▶ forecast for 2017q3?

$$\hat{u}_{2017q3} = \hat{\rho_0} + \hat{\rho_1}u_{2017q2} + \hat{\rho_2}u_{2017q1}$$

We have unemployment data that ends in 2017q2

▶ forecast for 2017q3?

$$\hat{u}_{2017q3} = \hat{\rho_0} + \hat{\rho_1}u_{2017q2} + \hat{\rho_2}u_{2017q1}$$

▶ forecast for 2017q4?

We have unemployment data that ends in 2017q2

▶ **forecast** for 2017q3?

$$\hat{u}_{2017q3} = \hat{\rho_0} + \hat{\rho_1} u_{2017q2} + \hat{\rho_2} u_{2017q1}$$

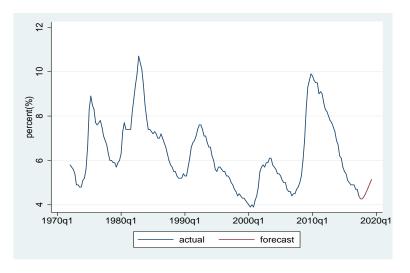
▶ forecast for 2017q4?

$$\hat{u}_{2017q4} = \hat{\rho_0} + \hat{\rho_1}\hat{u}_{2017q3} + \hat{\rho_2}u_{2017q2}$$

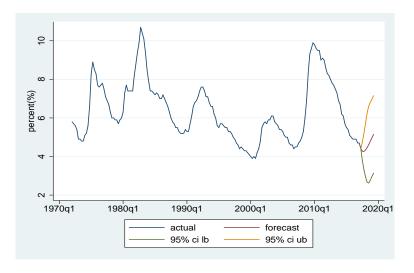
▶ forecast for 2017q5?

$$\hat{u}_{2017q5} = \hat{\rho_0} + \hat{\rho_1}\hat{u}_{2017q4} + \hat{\rho_2}\hat{u}_{2017q3}$$

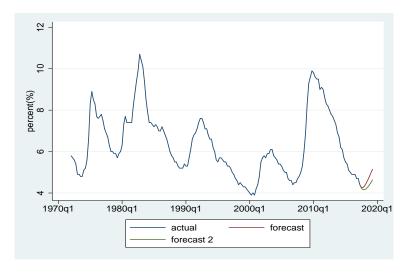
and so on ...



Forecast is for a return to sample average

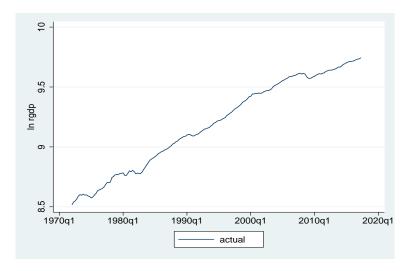


Forecasts get less accurate the farther you go



New "forecast 2" based on 1989q1-2017q2 data

SCALAR EXAMPLE: (NATURAL LOG OF) RGDP



Real GDP (2009\$s), 1972q1-2017q2. Source: fred.stlouisfed.org.

Autoregressive time series models do **not** behave well for variables that are growing (i.e. non-stationary)

In addition, they're not very informative for forecasting:

$$y_t = (1 + g_t)y_{t-1}$$
$$\ln y_t = \underbrace{\ln(1 + g_t)}_{\approx g_t} + \ln y_{t-1}$$

What if I run regression: $\ln y_t = \rho_0 + \rho_1 \ln y_{t-1} + \epsilon_t$?

var lrgdp if qtr>=tq(1972q1), lags(1)

▶ I should get $\rho_1 \approx 1$

```
var lrgdp if qtr>=tq(1972q1), lags(1)
Vector autoregression
Sample: 1972q1 - 2017q2
                                                     No. of obs
                                                                              1.82
Log likelihood = 623.8009
                                                     ATC
FPF
                  .0000631
                                                     HQIC
                                                                      = -6.818704
Det(Sigma_ml) = .0000617
                                                     SBIC
                                                                      = -6.797768
Equation
                                                            P>chi 2
                    Parms
                               RMSE
                                         R-sa
                                                   chi2
1rgdp
                                                 411020.2
                               .0079
                                        0.9996
                                                            0.0000
                     coef.
                                                            [95% Conf. Interval]
       1radp
                             Std. Err.
                                                  P>|z|
                                             z
1rgdp
       1rgdp
         L1.
                  .9963956
                             .0015542
                                         641.11
                                                  0.000
                                                            . 9933495
                                                                         .9994418
       _cons
                  .0399435
                             .0142906
                                           2.80
                                                  0.005
                                                            .0119343
                                                                         .0679526
```

If y_t is growing:

$$\Delta y_t \equiv \ln y_t - \ln y_{t-1} \approx g_t$$

is stationary!

I can run regression:

$$\Delta y_t = \rho_0 + \rho_1 \Delta y_{t-1} + \ldots + \epsilon_t$$

- ▶ ρ_1 : if Δy_{t-1} was lower than usual last quarter, is Δy_t likely to be higher/lower than usual today?
 - ▶ in other words, are output growth fluctuations *persistent*?

var growth if qtr>=tq(1972q1), lags(1 2)

L1.

_cons

.3286373

.1175306

```
. var growth if gtr>=tg(1972g1), lags(1/2)
Vector autoregression
Sample: 1972q1 - 2017q2
                                                    No. of obs
Loa likelihood = -202,1218
                                                    AIC
                                                    HOIC
Det(Sigma ml) = .5396911
                                                    SBIC
                                                                    = 2,306899
Equation
                   Parms
                               RMSE
                                       R-sq
                                                  chi 2
                                                           P>chi2
growth
                            .740767
                                       0.1508
                                                32, 30694
                                                           0.0000
                                                           [95% Conf. Interval]
      arowth
                    coef.
                            Std. Err.
                                                 P> |z|
arowth
      arowth
```

4.47 0.000

1.60

0.110

.1844638

-.0267244

.2256456

.4728109

.2617856

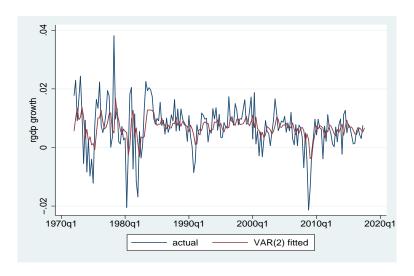
. 5323717

▶ RGDP growth hard to explain: $R^2 = 0.151$ ⇒ be cautious forecasting!

.0735593

.0736008

.3790087 .0782479 4.84 0.000



Can use growth rates to get RGDP in (log) levels, plot that instead $\,$

$$x_t = \rho_0 + \rho_1 x_{t-1} + \epsilon_t$$

$$\begin{bmatrix} u_t \\ \Delta y_t \end{bmatrix} = \begin{bmatrix} \rho_0^u \\ \rho_0^y \end{bmatrix} + \begin{bmatrix} \rho_1^{uu} & \rho_1^{uy} \\ \rho_1^{yu} & \rho_1^{yy} \end{bmatrix} \begin{bmatrix} u_{t-1} \\ \Delta y_{t-1} \end{bmatrix} + \begin{bmatrix} \epsilon_t^1 \\ \epsilon_t^2 \end{bmatrix}$$

- ▶ u_t depends not just on u_{t-1} but also Δy_{t-1}
- ▶ forecast for how quickly unemployment returns to "usual" depends on whether output growth is above/below "usual"
- \triangleright same goes for Δy_t
- obvously, you can consider more lags

A more interesting example (with only one lag):

$$\begin{bmatrix} r_t \\ u_t \\ \Delta y_t \\ \Delta p_t \end{bmatrix} = \rho_0 + \rho_1 \begin{bmatrix} r_{t-1} \\ u_{t-1} \\ \Delta y_{t-1} \\ \Delta p_{t-1} \end{bmatrix} + \epsilon_t$$

- ightharpoonup suppose r_t is the Central Bank's policy rate (overnight rate, federal funds rate)
- ightharpoonup now u_t depends on interest rate policy
- VAR allows us to estimate the Central Bank's "policy rule"
 - ▶ how r_t responds to past values of Δy_t , Δp_t , etc. and today's unexpected shocks

var ffr unemp growth infl if qtr>=tq(1983q1), lags(1)

| | Coef. | Std. Err. | z | P> z | [95% Conf. | Interval] |
|---------------|------------------------------|--|-------|--|------------|---|
| | | | | | | |
| | | | | | | |
| L1. | .956098 | .0133266 | 71.74 | 0.000 | .9299783 | .9822178 |
| unemp L1. | 015216 | .0231139 | -0.66 | 0.510 | 0605184 | .0300864 |
| growth L1. | .3741846 | .0622191 | 6.01 | 0.000 | . 2522374 | .4961318 |
| inf1 | | | | | | |
| L1. | 005795 | .0818168 | -0.07 | 0.944 | 1661529 | .154563 |
| _cons | 0402746 | .1673516 | -0.24 | 0.810 | 3682777 | . 2877285 |
| | growth L1. infl L1. | ffr 1956098 unemp 1015216 growth 13741846 infl 1005795 | ffr | ffr956098 .0133266 71.74 unemp t1015216 .0231139 -0.66 growth t13741846 .0622191 6.01 infl t1005795 .0818168 -0.07 | ffr | ffr L1956098 .0133266 71.74 0.000 .9299783 unemp L1015216 .0231139 -0.66 0.5100605184 growth L13741846 .0622191 6.01 0.000 .2522374 infl L1005795 .0818168 -0.07 0.9441661529 |

▶ Signs of coefficients make sense

Data ends in 2017q2, so forecasting proceeds in analogous manner to scalar case

- **▶ forecast** for 2017q3?
 - use coeff estimates and **actual** data on $r, u, \Delta y, \Delta p$ from 2017q2 and 2017q1
- **▶ forecast** for 2017q4?
 - use coeff estimates and **forecasted** data on r, u, Δy , Δp from 2017q3 and 2017q2
- ▶ and so on ...