# FINC 584 PS3

18 October, 2022

### Problem 1.a

	$beta\_hat$	beta_se
$\beta_0$	4.995182	0.0044644
$\beta_1$	0.070851	0.0003386

# Problem 1.b

```
ak<- ak_full %>% slice(1:5000)

## Compute beta_hats for conventional and robust
x <- matrix(data = c(rep(1,nrow(ak)), ak$edu), ncol = 2)
y <- matrix(data = ak$logwage)
beta_conventional <- solve(t(x) %*% x)%*%(t(x)%*%y)
resid <- y - x %*% beta_conventional

## Compute SEs for conventional method
sigma2_hat <- mean(resid^2)
beta_conventional_se <- sqrt(diag(sigma2_hat * (solve(t(x) %*% x))))[2]

## Compute SEs for robust method
sandwich_bread <- solve(t(x) %*% x)</pre>
```

	beta_hat_1_se
Conventional	0.0029033
Robust	0.0029218
Bootstrap	0.0029199

```
# ggplot(data = data.frame(beta_hat_bootstrap), aes(x = beta_hat_bootstrap)) +
# geom_histogram(fill = 'steelblue', color = 'black') +
# labs(title = 'Distribution of beta_hat_1') +
# theme_bw() +
# theme(plot.title = element_text(hjust = 0.5))
```

The results are pretty much what I expected. All three standard errors are roughly similar. The robust standard errors are larger than the conventional OLS estimators which is expected. The bootstrap method is similar to the other two methods in terms of efficiency.

# Problem 1C

```
confidence_interval <- function(b, b_se){
  return( abs((b - beta_conventional_full[2]))/b_se < qnorm(.975))
}

conventional_method <- function(x,y){

  beta_conventional <- solve(t(x) %*% x)%*%(t(x)%*%y)
  resid <- y - x %*% beta_conventional

## Compute SEs for conventional method
  beta_conventional_se <- sqrt(diag(mean(resid^2) * (solve(t(x) %*% x))))[2]</pre>
```

```
return(confidence_interval(beta_conventional[2], beta_conventional_se))
}
robust_method <- function(x,y){</pre>
  beta_conventional <- solve(t(x) %*\% x)%*\%(t(x)%*\%y)
  resid <- y - x %*% beta_conventional</pre>
  ## Compute SEs for robust method
  sandwich_bread <- solve(t(x) %*% x)</pre>
  sandwich_meat <- t(x) %*% diag(as.vector(resid)^2) %*% x</pre>
  beta_robust_se <- sqrt(diag(sandwich_bread %*% sandwich_meat %*% sandwich_bread))[2]
  return(confidence_interval(beta_conventional[2], beta_robust_se))
}
bootstrap_method <- function(data,n,m){</pre>
  bootstrap_estimates <- rep(NA, m)
  for(i in 1:length(bootstrap_estimates)){
    bootstrap_estimates[i] <- lm(logwage ~ edu,</pre>
                                 data = data[sample(1:n, n, replace = T),])$coefficients['edu']
  }
 # return((mean(bootstrap_estimates) - beta_conventional_full[2])/sd(bootstrap_estimates))
return(confidence_interval(mean(bootstrap_estimates), sd(bootstrap_estimates)))
}
problem1c <- function(n, m, data){</pre>
  mc_sample <- data[sample(1:nrow(data), n, replace = T),]</pre>
  x <- matrix(data = c(rep(1,n), mc_sample$edu), ncol = 2)</pre>
  y <- matrix(data = mc_sample$logwage)</pre>
  return(data.frame(n = n,
                     conventional = conventional_method(x,y),
                     robust = robust_method(x,y),
                     bootstrap = bootstrap_method(mc_sample,n,m)
}
iter <- 5000
results <- map_dfr(.x = c(rep(50, iter), rep(100, iter), rep(500, iter)), .f = problem1c, data = ak_ful
results %>% group_by(n) %>%
```

```
summarize_all( ~sum(.)*100/iter)
```

```
## # A tibble: 3 x 4
##
        n conventional robust bootstrap
##
          <dbl> <dbl>
    <dbl>
                 90.1 92.1
## 1
       50
                                  92.6
                        94.2
## 2
      100
                 91.3
                                  93.9
## 3
      500
                 92.8
                        95.6
                                  95
```

### Problem 1d

$$h(\hat{\beta}) = e^{\hat{\beta}_0 + 9\hat{\beta}_1} e^{\frac{\sigma^2}{2}} - e^{\hat{\beta}_0 + 9\hat{\beta}_1} e^{\frac{\sigma^2}{2}}$$
$$h(\hat{\beta}) = e^{\frac{\sigma^2}{2}} \left( e^{\hat{\beta}_0 + 9\hat{\beta}_1} - e^{\hat{\beta}_0 + 8\hat{\beta}_1} \right)$$

Thus,

$$\nabla h(\hat{\beta}) = \left(0, e^{\frac{\sigma^2}{2}} \left(9e^{\hat{\beta}_0 + 9\hat{\beta}_1} - 8e^{\hat{\beta}_0 + 8\hat{\beta}_1}\right)\right)$$

By the delta method, we have that

$$\sqrt(n)(h(\hat{\beta}) - h(\beta)) \to^D N\left(0, \nabla h(\hat{\beta}) * \Sigma * \nabla (h(\hat{\beta})^T\right)$$

```
e_sigma <- exp(sigma2_hat_full/2)
b0 <- beta_conventional_full[1]
b1 <- beta_conventional_full[2]

del_h <- matrix(c(0, e_sigma*(9*exp(b0 + 9*b1) - 8*exp(b0 + 8*b1))), ncol = 2)
S <- sigma2_hat_full * (solve(t(x_full) %*% x_full))

beta_hat_1d <- e_sigma*(exp(b0 + 9*b1) - exp(b0 + 8*b1))
beta_hat_1d_se <- sqrt(del_h %*% S %*% t(del_h))

ci_1d <- c(beta_hat_1d - qnorm(.975)*beta_hat_1d_se, beta_hat_1d + qnorm(.975)*beta_hat_1d_se)</pre>
```

The resulting confidence interval is 23.073871, 23.7771942.