

Generic Representation of Self-Similarity via Structure Sensitive Sampling of Noisy Imagery

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Abstract

An adaptive sampling scheme is presented for discrete representation of complex patterns in noisy imagery. In this paper, patterns to be observed are assumed to be generated as fractal attractors associated with a fixed set of unknown contraction mappings. To maintain geometric complexity, the brightness distribution of self-similar patterns are counted on 2D array of Gaussian probability density functions. By solving a diffusion equation on the Gaussian array, capturing probability of unknown fractal attractor is generated as a multi-scale image. The totality of local maxima of the capturing probability, then, yields a pattern sensitive sampling of fractal attractors. For eliminating background noise in this sampling process, two filters are introduced: input filter based on local structure analysis on the Gaussian array, and, output filter based on probabilistic complexity analysis at feature points. The sampled image through these filters are structure sensitive so that extracted feature patterns support invariant subset with respect to mapping sets associated with observed patterns. As the main result, a generic model is established for unknown self-similar patterns in background noise. The detectability of the generic model has been verified through simulation studies.

1 Introductory Remarks

Complex patterns are captured as computable entities through coding on discrete image plane. In many practical applications, discrete representation should maintain complete information for exact restoration of complex imagery. However, it is not easy to generate such “visible” code of random imagery within conventional statistical – computational frameworks. For instance, sampling on “very fine” lattice often yields “fragile” discrete representation that is susceptible to non-essential pattern deformation. To recognize

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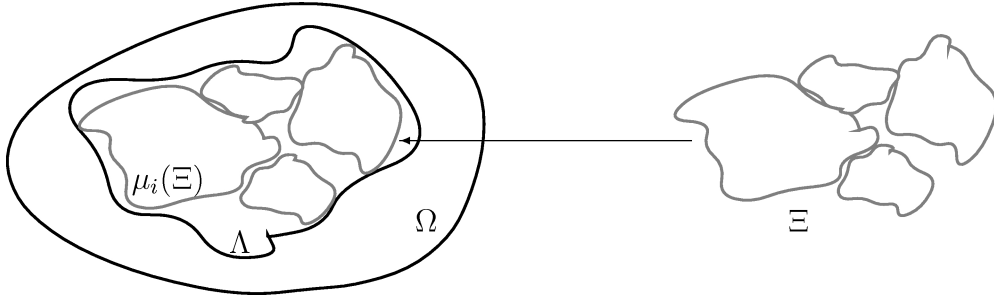


Fig. 1. Collage

For any pattern $\Lambda \subset \Omega$, there exists a set of contraction mapping $\nu = \{\mu_i\}$, $\mu_i : \Omega \rightarrow \Omega$ that yields an invariant subset $\Xi \subset \Omega$ approximating the pattern Λ within arbitrary small imaging error.

intrinsic features out of representation noise, thus, such susceptible representation must be handled through sophisticated matching processes.

The representation difficulty arises from discrete image modeling. Logically, image model must be independent on pattern structures to be detected because pattern grammar should be applied to *a priori* fixed lattice. Geometrically, however, sampling process of discrete image should be adapted to specific feature distribution. Thus, without “generic” representation, pattern sampling easily falls into a serious self-contradiction: *to adjust lattice to not-yet-identified patterns to be represented*.

A potential way to bypass the self-contradiction is to introduce the self-similarity as *a priori* pattern structure. Noticing logical – geometric coordination in self-similarity imaging processes, in this paper, we assume that patterns to be observed are generated as fractal attractors associated with unknown set of contraction mappings. Without serious loss of generality, in the following discussions, we suppose that the number of contraction mappings can be guessed. The assumption of self-similarity is not so restrictive because we can approximate any patterns in terms of the following “Fractal Collage[1]”: *For arbitrary pattern Λ in a fixed image plane Ω , there exists a set of contraction mappings $\nu = \{\mu_i\}$, that yields an invariant subset $\Xi \subset \Omega$ for approximating the pattern within arbitrary small imaging error (Fig. 1)*. This implies that any observed patterns can be coded in terms of finite symbols. The finite code completely specifies imaging process for generating fractal attractor of infinite geometric complexity.

In contrast with conventional statistical – computational representation, fractal model conveys complete information to specifies invoked contraction mappings. In fact, we have enough data for determining mapping parameter as the distribution of attractor points. Hence, we have logical bases for pattern coding as the following “Structural Observability[2]”: *The attractor*

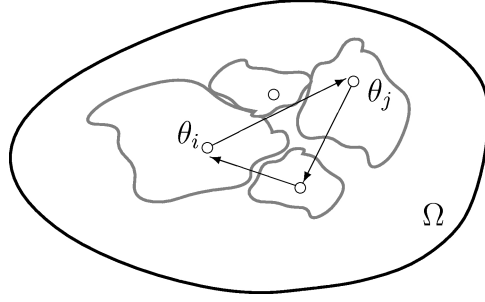


Fig. 2. Finite Composite

The attractor Ξ is covered by the totality of fixed points $\theta_i, \theta_j, \dots$, associated with all finite composite of the mappings $\dots \mu_j \mu_i$.

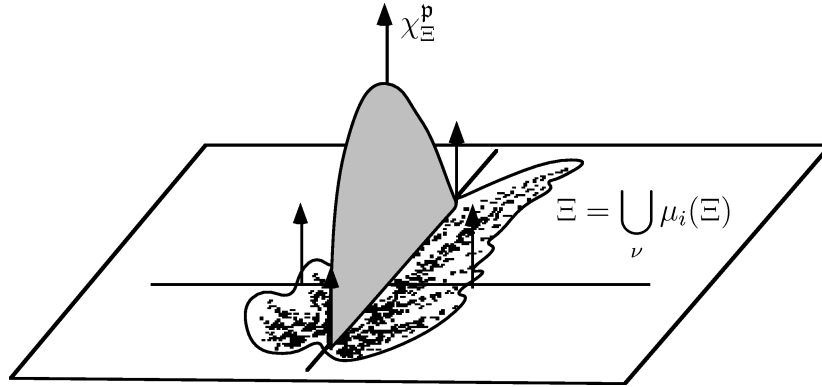


Fig. 3. Invariant Measure

For arbitrary self-similar pattern Ξ generated by random application of fixed contraction mappings μ_i , there exists a measure χ_Ξ^p that is invariant with respect to the transform by the mappings.

Ξ is covered by the totality of fixed points $\theta_i, \theta_j, \dots$, associated with all finite composite of the mappings (Fig. 2). Thus, pattern coding results in identifying origin – destination pairs in complex attractors. Since each attractor point deterministically “jumps” into the attractor by a contraction mapping, we have exact origin – destination associations in observed imagery.

In addition, the self-similarity induces definite association between geometric order, i.e., spatial distribution of attractor points, and probability for pattern capturing, i.e., gray level distribution. This implies that we can analyze pattern structure via the estimation of the “Invariant Measure[1]”: *For arbitrary attractor generated by random application of fixed contraction mappings, there exists a measure χ_Ξ^p that is invariant with respect to transform by the mappings (Fig. 3).* The existence of invariant measure implies the association

between the distribution pattern and the density function of fractal attractors. The self-similarity of the density function introduces the self-similarity in the distribution of statistical parameter.

We can exploit the collage theorem as a general framework for fractal pattern coding. Fractal code is described in terms of finite contraction mappings that restore observed patterns of infinite complexity. By invoking the structural observability, we can design the mappings through origin – destination association on discrete points. To discriminate the discrete image from background noise, the invariance of observed “brightness distribution” with respect to the mappings to be designed should be analyzed. In this paper, hence, we consider the integration of these three aspects of the self-similarity to develop a unified sampling scheme for unknown complex patterns.

2 Self-Similarity on Continuous Image

Let $\Omega \subset R^2$ be a continuous image plane and suppose that patterns are generated within the Borel field $\mathcal{F}[\Omega]$ of the totality of subsets of Ω . The disparity between patterns $A, B \in \mathcal{F}[\Omega]$ is indexed in terms of the Hausdorff distance $\eta[A, B]$ defined by

$$\eta[A, B] = \max \{ \overleftarrow{\eta}[A, B], \overleftarrow{\eta}[B, A] \}, \quad (1a)$$

$$\overleftarrow{\eta}[A, B] = \max_{\omega \in A} \left\{ \min_{\lambda \in B} |\omega - \lambda| \right\}. \quad (1b)$$

Consider a fixed set of unknown contraction mappings $\nu = \{\mu_i, i = 1, 2, \dots, m\}$ with length $\|\nu\| = m$ where $\mu_i : \Omega \xrightarrow{s_{\mu_i}} \Omega$ is a mapping from Ω into itself with contractivity factor s_{μ_i} , $0 < s_{\mu_i} < 1$, i.e.,

$$(2) \quad |\mu_i(\omega_1) - \mu_i(\omega_2)| \leq s_{\mu_i} |\omega_1 - \omega_2|,$$

for any $\omega_1, \omega_2 \in \Omega$. By the contractivity, we can *program* pattern generation processes as the collage of mapping images $\mu_i(\Omega)$. For self-similar patterns, particularly, we have the following exact pattern generation scheme on program set ν :

Proposition 2.1 (Fractal Attractor) *Let Ξ be the attractor generated by the “program” ν to satisfy*

$$(3) \quad \Xi = \bigcup_{\mu_i \in \nu} \mu_i(\Xi).$$

The attractor Ξ can be successively approximated by the dynamical system on $\mathcal{F}[\Omega]$:

$$(4a) \quad \xi_{t+1} \in \bigcup_{\mu_i \in \nu} \mu_i(\xi_t),$$

$$\Xi_t = \{\xi_\tau \in \Omega, \tau \leq t\}. \quad (4b)$$

The sequence Ξ_t converges to the attractor Ξ in the following sense

$$(5) \quad \lim_{t \rightarrow \infty} \eta[\Xi_t, \Xi] = 0.$$

If the initial value of the process (4) is confined within target attractor

$$\Xi_0 \subset \Xi,$$

the imaging process is monotone, i.e.,

$$(6) \quad \Xi_0 \subset \Xi_1 \subset \Xi_2 \subset \cdots \subset \Xi_t \subset \cdots \subset \Xi.$$

3 2D Gaussian Sampling

By assigning the following basic measure to the image plane

$$(7) \quad dP(\omega) = \frac{d\omega}{\int_{\Omega} d\omega},$$

we can introduce a probability space $(\Omega, \mathcal{F}[\Omega], P)$ as the basis of image analysis. For instance, the “brightness” of the patterns Ξ_t and Ξ are represented by distributions on the probability space as follows:

Proposition 3.1 (Convergent Distribution) [3] *Let Ξ_t , $t = 1, 2, \dots$, be a sequence of point sets with initial value $\Xi_0 \subset \Xi$. Assume that $\Xi_t \subset \Omega$ for any $t \geq 1$. Define*

$$(8) \quad \chi_{\Xi_t} = \frac{1}{\|\Xi_t\|} \sum_{\xi_t \in \Xi_t} \delta_{\xi_t}.$$

Then $\Xi_t \subset \Xi$ and there exists a distribution χ_{Ξ} satisfying

$$\chi_{\Xi_t} \rightarrow \chi_{\Xi} \text{ as } t \rightarrow \infty, \quad (9a)$$

in the following sense:

$$\lim_{t \rightarrow \infty} \chi_{\Xi_t}(f) = \chi_{\Xi}(f), \quad (9b)$$

for arbitrary locally summable “test function” f .

By using a system of test functions, we can extend sampling mechanism to distributions. For this purpose, consider the following one parameter family of test functions:

$$(10) \quad \mathfrak{T} = \{f_{\xi}(\omega), \xi \in \Omega\}.$$

Generally, sample values of the distribution Ξ on \mathfrak{T} is represented as the following linear functional:

$$(11) \quad \chi_{\Xi}(f_{\xi}) = f_{\xi} * \chi_{\Xi}(\omega) = \int_{\Omega} f(\xi - \omega) \chi_{\Xi} dP(\omega).$$

The representation error can be reduced arbitrary by introducing the following δ -convergent sequence:

$$(12) \quad f_{\xi}^{\epsilon} \rightarrow \delta_{\xi} \text{ as } \epsilon \rightarrow 0,$$

in the space of test functions. For instance, let the following simple averaging functions be introduced as test functions:

$$(13) \quad f_{\xi}^{\epsilon} = \begin{cases} \frac{1}{2\pi\epsilon^2}; & \text{for } |\omega - \xi| \leq \epsilon, \\ 0; & \text{otherwise.} \end{cases}$$

By “counting” the distribution on f_{ξ}^{ϵ} , we have “average value” of distribution χ_{Ξ} in small region around ξ as the the intensity of image Ξ such that

$$(14) \quad f_{\xi}^{\epsilon} \rightarrow \delta_{\xi} \text{ then } \chi_{\Xi}(f_{\xi}^{\epsilon}) \rightarrow \chi_{\Xi}.$$

By this convergence, we have discrete representation for the image Ξ with brightness distribution χ_{Ξ} as follows:

Definition 3.2 (Generalized Sampling) Let D be discrete subset of Ω . The set of the values of measure χ_{Ξ} on one parameter family of test functions $\mathfrak{F}_D = \{f_d, d \in D\}$

$$(15) \quad \chi_{\Xi}^D = \left\{ \chi_{\Xi}(f_d) \mid d \in D \right\}$$

is called generalized sampling of imagery χ_{Ξ} .

Consider the representation of the self-similarity on the generalized sampling χ_{Ξ}^D . Let the discrete subset D be *a priori* given as a “uniform” lattice with resolution $\epsilon/\sqrt{2}$

$$(16) \quad \mathfrak{L} = \begin{bmatrix} \cdots & \cdots & \cdots & \cdots & \cdots \\ \cdots & (i - \frac{\epsilon}{\sqrt{2}}, j + \frac{\epsilon}{\sqrt{2}}) & (i, j + \frac{\epsilon}{\sqrt{2}}) & (i + \frac{\epsilon}{\sqrt{2}}, j + \frac{\epsilon}{\sqrt{2}}) & \cdots \\ \cdots & (i - \frac{\epsilon}{\sqrt{2}}, j) & (i, j) & (i + \frac{\epsilon}{\sqrt{2}}, j) & \cdots \\ \cdots & (i - \frac{\epsilon}{\sqrt{2}}, j - \frac{\epsilon}{\sqrt{2}}) & (i, j - \frac{\epsilon}{\sqrt{2}}) & (i + \frac{\epsilon}{\sqrt{2}}, j - \frac{\epsilon}{\sqrt{2}}) & \cdots \\ \cdots & \cdots & \cdots & \cdots & \cdots \end{bmatrix}$$

and consider partitioning of the image plane Ω by unit disks $\{d_{\ell}, \ell \in \mathfrak{L}\}$:

$$(17) \quad d_{\ell} = \left\{ \omega \in \Omega \mid |\omega - \ell| \leq \epsilon \right\}.$$

By evaluating the "brightness" of the image χ_Ξ at the "point" $\ell \in \mathfrak{L}$ in terms of the mean value

$$(18) \quad \chi_\Xi * \delta_\ell \sim \chi_\Xi(f_\ell^\epsilon),$$

we have local representation of the self-similarity in d_ℓ as follows:

$$\chi_\Xi(f_\ell^\epsilon)P(D_\ell^+) \geq \|\nu\|\chi_\Xi(f_\ell^\epsilon)P(d_\ell), \quad (19a)$$

where D_ℓ^+ denotes discrete support of the brightness distribution given by

$$D_\ell^+ = \left\{ d_k, k \in \mathfrak{L} \mid \eta[k, \ell] \leq \epsilon \text{ and } \chi_\Xi(f_k^\epsilon) > 0 \right\}. \quad (19b)$$

Hence, we have the following "input filter" to discrete image plane \mathfrak{L} for observing unknown self-similar patterns with complexity factor $\|\nu\|$:

$$(20) \quad P(D_\ell^+) \geq \|\nu\|P(d_\ell).$$

Due to the loss of information by simple averaging, however, it is not easy to evaluate the expansion $P(D_\ell^+)$ relative to the measuring unit $P(d_\ell)$. As another version of sampling mechanism, consider the family of Gaussian probability density functions $\{g_\sigma^\ell, \sigma > 0\}$, where

$$(21) \quad g_\sigma^\ell(\omega) = \frac{1}{2\pi\sigma} \exp \left[-\frac{|\omega - \ell|^2}{2\sigma} \right].$$

Noticing g_σ^ℓ , $\sigma > 0$ yields a δ -convergent sequence:

$$(22) \quad g_\sigma^\ell \rightarrow \delta_\ell \text{ as } \sigma \rightarrow 0,$$

we have the following stochastic sampling scheme on deterministic image plane \mathfrak{L} :

$$(23) \quad \mathfrak{G} = \{g_\sigma^\ell, \ell \in \mathfrak{L}\}.$$

By testing the value of distributions on \mathfrak{G} , we have the following stochastically sampled image:

$$(24) \quad \chi_\Xi^\mathfrak{G} = \{\chi_\Xi(g_\sigma^\ell)_\omega, \ell \in \mathfrak{L}\},$$

where $\chi_\Xi(g_\sigma^\ell)_\omega = g_\sigma^\ell * \chi_\Xi(\omega)$. In the sampling scheme $(\mathfrak{G}, \mathfrak{L})$, complete information χ_Ξ of infinite resolution is associated with discrete image plane \mathfrak{L} . Following the zero-cross method[5], for instance, the boundary point ω' associated with a point image δ_ξ should be detected by

$$\frac{1}{2}\Delta g_\sigma(\omega' - \xi) = \frac{1}{2\pi\sigma^2} \exp \left[-\frac{|\omega' - \xi|^2}{2\sigma} \right] \left(\frac{|\omega' - \xi|^2}{2\sigma} - 1 \right)$$

$$(25) \quad = 0.$$

The combination of the local complexity (20) with pixel boundary evaluation (25) yields the following “counting rule” for input filtering on digital image plane \mathfrak{L} :

$$(26) \quad P(N_\ell^+) \geq \|\nu\|,$$

where N_ℓ^+ means the number of pixels located in D_ℓ^+ (Fig. ??).

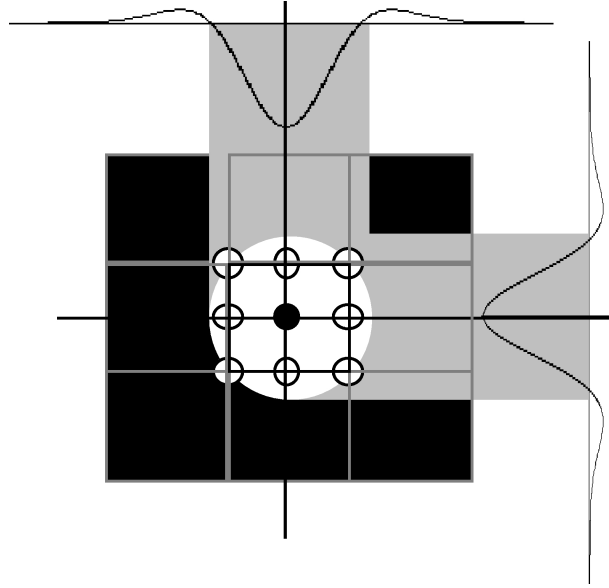


Fig. 4. 2D Gaussian Array for Visualizing Point Images

The pixel of interest (\bullet) is connected with eight neighborhood pixels (\circ) in the lattice \mathfrak{L} . To each pixel, a Gaussian distribution is assigned to clarify local area in which connectedness of pixels is tested. The variance parameter of the Gaussian distribution specifies the zero-cross boundary of the pixel (\bullet).

4 Self-Similarity on Measures

The “painting” process (4) combined with the generalized “brightness” control (8) induce the self-similarity on the distribution χ_Ξ . Noticing static constraint (3), we have

Proposition 4.1 (Invariant Measure) *For the programs μ_i to be selected with probability p_{μ_i} , the measure χ_Ξ on $\mathcal{F}[\Omega]$ is invariant with respect to the*

following Markov operation:

$$(27) \quad \chi_{\Xi}^{\mathfrak{p}}(\cdot) = \sum_{\mu_i \in \nu} p_{\mu_i} \chi_{\Xi}^{\mathfrak{p}}[\mu_i^{-1}(\cdot)], \quad (\cdot) \in \mathcal{F}[\Omega],$$

where p_{μ_i} and $\mu_i \in \nu$ are nonnegative constants such that $\sum_{\mu_i \in \nu} p_{\mu_i} = 1$.

Let the measure χ_{Ξ} be regularized by the following one-parameter family of test functions

$$(28) \quad \mathfrak{g} = \{g_{\tau}, \tau \geq 0\} = \left\{ \frac{e^{-\frac{|\omega|^2}{2\tau}}}{2\pi\tau}, \tau \geq 0 \right\},$$

and consider the adaptation of “scale parameter” τ to the self-similar pattern Ξ to be detected.

The imaging process (4) expands initial points Ξ_0 through nondeterministic scattering within a fixed domain Ξ . This implies that the process should be modeled by 2D dynamical system with the following antagonistic imaging mechanisms

- diffusion of point image δ_{ξ} within image plane Ω , and,
- successive reduction of imaging domain via not-yet-identified contraction mappings $\mu_i \in \nu$.

Let the model be described in terms of the following system

$$(29) \quad \exists \mu_i \in \nu : \omega_{t+1} = \mu_i(\omega_t),$$

where random shift of a point image is considered to be observed as a sample path on a “tectonic plate” successively reduced by randomly selected mapping $\mu_i \in \nu$. By identifying “observation error” with 2D Brownian motion, we have the following stochastic evaluation for capturing the point image within an ordinary domain $\Gamma \in \mathcal{F}[\Omega]$:

$$\begin{aligned} P^w(t, \omega, \Gamma) &= \int_{\Gamma} g_t(\gamma - \omega) d\gamma \\ &= \int_{\Omega} c_{\Gamma}(\gamma) \frac{e^{-\frac{|\omega - \gamma|^2}{2t}}}{2\pi t} d\gamma = g_t * c_{\Gamma}(\omega), \end{aligned} \quad (30a)$$

where t is the time elapse for capturing the point ξ and c_{γ} denotes the characteristic function of “regular” set Γ :

$$c_{\Gamma}(\omega) = \begin{cases} 1; & \text{for } \omega \in \Gamma, \\ 0; & \text{otherwise.} \end{cases} \quad (30b)$$

In (30), the point image is assumed to be emitted from $\omega \in \Omega$. The evaluation (30) can be extended to self-similar patterns of geometric singularity as follows:

$$P^w(t, \omega, \Xi) = g_t * \chi_\Xi(\omega), \quad (31a)$$

and to its mapping image, as well:

$$P^w(t, \omega, \mu_i(\Xi)) = g_t * \chi_{\mu_i(\Xi)}. \quad (31b)$$

Suppose that the time elapse is counted in terms of the activation of the painting process (4). For such situation, we have

$$(32) \quad P^{[1]}(t, \omega, \Xi) = \sum_{\mu_i \in \nu} p_{\mu_i} P^w(t, \omega, \mu_i(\Xi)),$$

where p_{μ_i} denotes the probability for selecting a mapping $\mu_i \in \nu$. Noticing the Chapman's Equation

$$(33) \quad g_t(\gamma - \omega) = g_t((\cdot) - \omega) * g_0(\gamma - (\cdot)),$$

$$(34) \quad \forall \xi, \gamma \in \Xi : g_0(\gamma - \xi) = \delta(\gamma - \xi),$$

we have the following stochastic evaluation for capturing the point image at $\gamma \in \Omega$ under the selection of μ_i :

$$(35) \quad p_{\mu_i} \sim p_\xi(\gamma | \mu_i) = \chi_\Xi(\xi) \cdot \delta(\gamma - \mu_i(\xi)),$$

where $p_\xi(\gamma | \mu_i)$ denotes the transition density function from ξ to γ conditioned by the selection $\mu_i \in \nu$ with *a priori* probability p_{μ_i} . Hence, the transition function associated with the imaging process (4) is computed by

$$(36) \quad \begin{aligned} P_\omega^t(\Xi | \mu_i) &= \int_{\Omega} \chi_\Xi(\mu_i^{-1}(\gamma)) \cdot g_t(\gamma - \omega) d\gamma \\ &= \chi_\Xi(\mu_i^{-1}(\cdot)) * g_t(\omega) = \chi_{\mu_i(\Xi)} * g_t(\omega) \\ &= P^w(t, \omega, \mu_i(\Xi)), \end{aligned}$$

for one step reduction by fixed $\mu_i \in \nu$. Assume that the mapping μ_i is selected uniformly in a fixed set ν with size $\|\nu\|$. By setting $p_{\mu_i} = \frac{1}{\|\nu\|}$, it follows that

$$(37) \quad P_\omega^{[1]}(\Xi | \mu_i) = \|\nu\|^{-1} P^w(t, \omega, \mu_i(\Xi)).$$

This implies that, the transition function conditioned by a fixed “programming framework” ν is obtained as follows:

$$(38) \quad P_\omega^{[1]}(\Xi|\nu) = \frac{1}{\|\nu\|} \sum_{\mu_i \in \nu} P^w(t, \omega, \mu_i(\Xi)) = \frac{1}{\|\nu\|} P^w(t, \omega, \Xi).$$

Iterating such capturing process, we have

$$(39) \quad P_\omega^{[n]}(\Xi|\nu) = \|\nu\|^{-1} P^{[n-1]}(n, \omega, \Xi) = \cdots = \|\nu\|^{-n} P^w(n, \omega, \Xi),$$

or, equivalently,

$$P_\omega^t(\Xi|\nu) = \exp[-\rho t] g_t * \chi_\Xi(\omega), \quad (40a)$$

where ρ is the complexity parameter defined by

$$\rho = \log \|\nu\|. \quad (40b)$$

Noticing that the transition function $P_\omega^t(\Xi|\nu)$ is generated by the following evolution equation

$$\frac{\partial P_\omega^t(\Xi|\nu)}{\partial t} = \frac{1}{2} \Delta P_\omega^t(\Xi|\nu) - \rho P_\omega^t(\Xi|\nu), \quad (41a)$$

with the initial distribution

$$P_\omega^0(\Xi|\nu) = \chi_\Xi, \quad (41b)$$

we have the following

Proposition 4.2 (Multi-Scale Image, Fig. 5) *Let ϵ be a small positive constant and consider the weighted average of Gaussian probability density function*

$$(42) \quad G_t^\rho = \rho \int_0^{t-\epsilon} e^{-\rho(t-\epsilon-\tau)} g_\tau d\tau.$$

Then the multi-scale image of Gaussian distribution G_t^ρ satisfies the following equation:

$$(43) \quad \mathcal{M}_\epsilon : \frac{\partial G_t^\rho}{\partial t} = \frac{1}{2} \Delta G_t^\rho + \rho[g_\epsilon - G_t^\rho].$$

The dynamical system \mathcal{M}_ϵ generates stochastic evaluation for a point image observed through Gaussian array. The system \mathcal{M}_ϵ converges to the following

system \mathcal{M} to generate the multi-scale image associated with “exact” point image δ :

$$(44) \quad \mathcal{M} : \frac{\partial G_t^\rho}{\partial t} = \frac{1}{2} \Delta G_t^\rho + \rho[\delta - G_t^\rho],$$

as $\epsilon \rightarrow 0$.

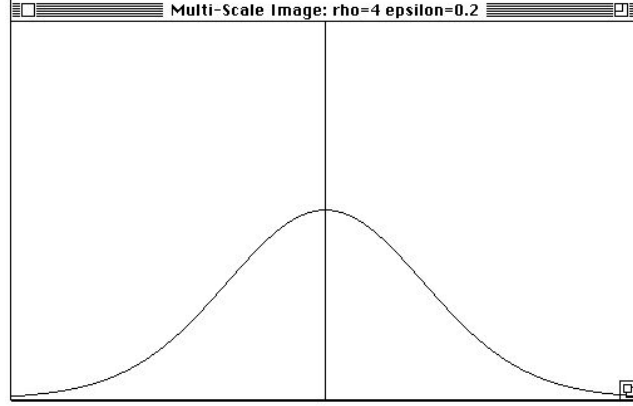


Fig. 5. Multi-Scale Image

A version of the weighted average (42) with $\rho = 4$ and $\epsilon = 0.2$ is indicated. G_t^ρ visualizes the stochastic evaluation of a point image emitted from the origin as a smooth field.

The steady state of the dynamical system (43) yield the final estimate for capturing the point images Ξ as a computable entity structured by ν . By superimposing the evaluation on the initial distribution, thus, we have the following

Proposition 4.3 (Capturing Probability) *Let χ_Ξ be a given brightness distribution to be collaged by $\nu = \{\mu_i\}$. Assume that the distribution is observed through the Gaussian array \mathfrak{G} . Then the probability for regenerating Ξ within the framework of maximum entropy capturing is visualized as a smooth field $\varphi(\omega|\nu)$ satisfying*

$$(45) \quad \frac{1}{2} \Delta \varphi(\omega|\nu) + \rho[\chi_\Xi(g_\epsilon^\ell)_\omega - \varphi(\omega|\nu)] = 0, \quad \ell \in \mathfrak{L}.$$

Generally, the probability distribution can be generated via the following equation:

$$(46) \quad \frac{1}{2} \Delta \varphi(f|\nu) + \rho[\chi_\Xi(f) - \varphi(f|\nu)] = 0,$$

where f denotes a test function.

5 Pattern Boundary on Invariant Measures

By invoking zero-cross criterion (25), we have the estimate of the capturing probability at pattern boundary as follows:

$$\varphi(\omega'|\nu) = \gamma_\tau, \quad (47a)$$

$$\gamma_\tau \sim g_\tau(\omega'), \quad |\omega'| = \sqrt{2\tau}. \quad (47b)$$

It should be noted that the level γ_τ can be specified without specification of mappings. For instance, the level γ_τ of imaging process with $\|\nu\| = 3$ can be computed by

$$\gamma_\tau \sim \frac{e^{-1}}{2\pi\tau} \simeq 0.04829.$$

This implies that the expansion $\hat{\Xi}_\nu$ and the boundary $\partial\hat{\Xi}_\nu$ of the distribution χ_Ξ can be specified by estimating the scale parameter τ . By paraphrasing the generator (46) as

$$\varphi(g_\epsilon^\ell|\nu) = \chi_\Xi(g_\epsilon^\ell) + \frac{1}{2}\Delta\varphi(g_\epsilon^\ell|\nu) \cdot \tau(\rho), \quad \ell \in \mathfrak{L}, \quad (48a)$$

$$\tau(\rho) = \frac{1}{\rho} = \frac{1}{\log \|\nu\|}, \quad (48b)$$

we have the following association:

$$(49) \quad \varphi((\cdot)|\nu) \sim g_{\tau(\rho)} \iff g_\epsilon^\ell + \int_0^{\tau(\rho)} \frac{1}{2}\Delta g_t dt,$$

on Gaussian array $(\mathfrak{G}, \mathfrak{L})$. Hence, we have the following estimates, $\hat{\Xi}_\nu$ and $\partial\hat{\Xi}_\nu$ on continuous image plane, respectively:

$$\hat{\Xi}_\nu = \left\{ \omega \in \Omega \mid \varphi(\omega|\nu) \geq \gamma(\rho) \right\}, \quad (50a)$$

$$\partial\hat{\Xi}_\nu = \left\{ \omega \in \Omega \mid \varphi(\omega|\nu) = \gamma(\rho) \right\}. \quad (50b)$$

Consider self-similar patterns generated in noisy background. For such patterns, we can generate the capturing probability and a version of conditional probability for evaluating possible variation of brightness, as follows:

$$p(\omega|\nu) = \frac{\varphi(\omega|\nu)}{C_\Omega^\varphi}, \quad (51a)$$

$$C_{\Omega}^{\varphi} = \int_{\Omega} \varphi(\omega|\nu) d\omega. \quad (51b)$$

By using the conditional probability, we can index the complexity of brightness variation in terms of the following Shannon's entropy

$$(52) \quad - \int_{\Omega} p(\omega|\nu) \log p(\omega|\nu) d\omega = -\mathcal{E} \left\{ \log p(\omega|\nu) \mid \nu \right\} = \hat{H}_{\nu}.$$

The existence of self-similarity structure should be verified through the comparison with the entropy evaluation under "null condition":

$$(53) \quad \mathcal{E} \left\{ \log p(\omega|\emptyset) \mid \emptyset \right\} = -\hat{H}_{\emptyset},$$

where $p(\omega|\emptyset) = \text{const.}$ on Ω . Hence, we have

Proposition 5.1 *Assume the background noise χ_{Ω} is uniformly distributed in the image plane Ω and suppose that observed measure χ_{Λ} is represented by*

$$(54) \quad \chi_{\Lambda} = \chi_{\Xi} + \chi_{\Omega}.$$

Then the boundary level is given by

$$\gamma = C_{\Omega}^{\varphi} \bar{p}_{\nu}, \quad (55a)$$

where

$$\log \bar{p}_{\nu} = 1 - \frac{1}{2} (1 - e^{\hat{H}_{\nu} - \hat{H}_{\emptyset}}) - \hat{H}_{\emptyset}. \quad (55b)$$

Proof. Noticing the maximum entropy estimate of variance is given by

$$(56) \quad \sigma_{(\cdot)} = \frac{1}{2\pi e} \exp[\hat{H}_{(\cdot)}],$$

we have

$$(57) \quad \bar{p}_{\nu} = \sup_{|\xi_{\Omega}|^2 > \sigma_{\Omega}} \frac{1}{2\pi\sigma_{\emptyset}} \exp \left[-\frac{|\xi_{\Omega}|^2}{2\sigma_{\emptyset}} \right] = \exp \left[1 - \frac{\sigma_{\Omega}}{2\sigma_{\emptyset}} - \hat{H}_{\emptyset} \right].$$

Since $\sigma_{\emptyset} = \sigma_{\nu} + \sigma_{\Omega}$, and

$$\begin{aligned} \hat{H}_{\nu} &= -\mathcal{E} \left\{ \log \left(\frac{1}{2\pi\sigma_{\nu}} \exp \left[-\frac{|\omega|^2}{2\sigma_{\nu}} \right] \right) \right\} \\ &= \log 2\pi e \sigma_{\emptyset}, \end{aligned} \quad (58a)$$

in the exterior of the support of χ_Ξ for mutually independent generalized random fields χ_Ξ and χ_Ω , it follows that

$$(59) \quad \frac{\sigma_\Omega}{\sigma_\emptyset} = 1 - \frac{\sigma_\nu}{\sigma_\emptyset} = 1 - \exp[\hat{H}_\nu - \hat{H}_\emptyset],$$

as was to be proved. \square

6 Self-Similarity on Stochastic Features

The capturing probability $\varphi(\omega|\nu)$ is the smoothing of gray level distribution of self-similar patterns. By modeling the imaging via unknown contraction mappings in terms of 2D Brownian motion on dynamically regenerated domain, a unified framework is introduced for information compression: the maximum entropy. Due to the infinite differentiability of generated field, on the other hand, the capturing probability maintains complete information of self-similarity processes. The association (49), particularly, implies that the generator of the capturing probability is adapted to the complexity of the patterns to be observed.

Consider a discrete image defined by

$$(60) \quad \tilde{\Theta} = \left\{ \tilde{\theta} \in \Omega \mid \nabla \varphi(\tilde{\theta}|\nu) = 0, \det [\nabla \nabla^T \varphi] (\tilde{\theta}|\nu) > 0, \Delta \varphi(\tilde{\theta}|\nu) < 0 \right\}.$$

Through the adaptation of the field $\varphi(\omega|\nu)$ to unknown generator, the image $\tilde{\Theta}$ yields a version of structurally sensitive sampling. The mapping structure is said to be uniformly observable if, for arbitrary $\xi \in \Xi$, there exists a finite composite $\langle \mu_i \rangle_t$ generating the fixed point ξ_t

$$\xi_t = \langle \mu_i \rangle_t (\xi_t), \quad \mu_i \in \nu, \quad (61a)$$

satisfying the following condition[4]:

$$|\xi - \xi_t| < \epsilon. \quad (61b)$$

The observability condition can be tested on discrete image $\tilde{\Theta}$ as follows:

Proposition 6.1 (Invariant Features) *Assume that there exists an subset $\Theta \subset \tilde{\Theta}$ invariant with respect to ν , i.e.,*

$$(62) \quad \Theta = \left\{ \theta \in \tilde{\Theta} \mid \exists \mu_i \in \nu : \mu_i^{-1}(\theta) \in \Theta \right\}.$$

Suppose that for arbitrary $\mu_i \in \nu$ there exist $\theta^o, \theta^d \in \Theta$ such that

$$(63) \quad \theta^d = \mu_i(\theta^o).$$

Then the imaging process (4) is uniformly observable.

Proof. By definition, invariant features Θ is made up of two types of feature points: periodic points Θ^p and destination points Θ^d . Periodic points are fixed points of some finite composite on ν , i.e.,

$$\Theta^p = \left\{ \theta_t^p \in \tilde{\Theta} \mid \exists \langle \mu_i \rangle_t : \theta_t^p = \langle \mu_i \rangle_t (\theta_t^p) \right\}, \quad (64a)$$

where

$$\langle \mu_i \rangle_t = \mu_{i_t} \mu_{i_{t-1}} \mu_{i_{t-2}} \cdots \mu_{i_2} \mu_{i_1}, \quad \mu_{i_\tau} \in \nu. \quad (64b)$$

On the other hand, each destination point is a mapping image of some finite composite, i.e.,

$$(65) \quad \Theta^d = \left\{ \theta_t^d \in \tilde{\Theta} \mid \exists \langle \mu_i \rangle_t : \theta_t^d = \langle \mu_i \rangle_t (\theta^p) \right\},$$

where $\theta^p \in \Theta^p$. Since $\Theta^p \subset \Xi$, and $\Theta^d \subset \Xi$, as well, it follows that

$$(66) \quad \Theta \subset \Xi.$$

By applying imaging process (4) to initial set Θ , we have a point $\xi_t \in \Xi$, for any $\xi \in \Xi$, such that,

$$(67) \quad \xi_t = \langle \mu_i \rangle_t (\theta), \quad \theta \in \Theta,$$

with $|\xi - \xi_t| < \frac{\epsilon}{2}$, and

$$\left(\max_{\mu_i \in \nu} s_{\mu_i} \right)^t \cdot \left[\max_{\gamma \in \Xi} |\gamma - \xi_t| \right] < \frac{\epsilon}{2}.$$

Since $\xi_t^f \in \langle \mu_i \rangle_t (\Xi)$, where $\xi_t^f = \langle \mu_i \rangle_t (\xi_t^f)$, it follows that

$$(68) \quad |\xi_t - \xi_t^f| < \frac{\epsilon}{2}.$$

Hence

$$(69) \quad |\xi - \xi_t^f| \leq |\xi - \xi_t| + |\xi_t - \xi_t^f| < \epsilon,$$

as was to be proved. \square

Obviously, $\Theta \subset \Xi$ if $\tilde{\Theta} \subset \Xi$. This implies that we can restrict the domain for extracting invariant features by the following pointwise “output filter”:

$$(70) \quad \hat{\Theta} = \left\{ \hat{\theta} \in \tilde{\Theta} \mid p(\hat{\theta}|\nu) \geq \bar{p}_\nu \right\}.$$

Thus, we have generic representation of the self-similarity on noisy discrete imagery as follows:

$$(71) \quad \Theta = \left\{ \theta \in \hat{\Theta} \mid \exists \mu_i \in \nu : \mu_i^{-1}(\theta) \in \Theta \right\}.$$

In this representation, the constraint for imaging process is grammatically specified on discrete pattern $\hat{\Theta}$. The discrete pattern $\hat{\Theta}$ is extracted within sampled image $\hat{\Theta}$ through pointwise filtering. The discrete information $\hat{\Theta}$, conversely, is generated through adaptive sampling based on stochastic evaluation $\varphi(\omega|\nu)$ for unknown mappings ν .

7 Experiments

Pattern detection on proposed sampling scheme was verified via simulation studies. In these simulations, fractal attractors were generated by Monte-Carlo simulation on continuous image model. To each attractors, uniformly distributed random dots were added as background noise. Result of simulation studies are illustrated in Figs. 6 – 9.

Figure 6 illustrate an observation of a fractal pattern χ_{Ξ} in background noise χ_{Ω} satisfying $\|\chi_{\Omega}\| = 2\|\chi_{\Xi}\|$. Pattern detection results in this situation are shown in Fig. 7. In these figures, the distribution of attractor points are “counted” on test functions f_{ξ}^{ϵ} on 2D lattice \mathcal{L} with $\epsilon = 1$ (Observables view). By selecting locally connected lattice point satisfying (26) with $\|\nu\| = 3$, we have the initial value for generating capturing probability $\chi_{\Xi}(g_{\epsilon}^{\ell})$. Extracted stochastic features are illustrated in “Features View” where $\hat{\Theta}$ is estimated via in-out discriminator (70) and indicated by (■) in background noise (□). As shown in Fig. 7, the generator of observed self-similar pattern is observable so that the generator yields invariant subset $\Theta \subset \hat{\Theta}$ (Coding View) and regenerates fractal attractor (Restoration View). Thus, we can detect the generator of observed pattern via structure sensitive sampling $\hat{\Theta}$ on discrete image $(\mathfrak{G}, \mathfrak{L})$.

Figures 8 and 9 illustrate another results where background noise satisfies $\|\chi_{\Omega}\| = 4\|\chi_{\Xi}\|$. As shown in this figure, the generator of observed self-similar pattern was detected successfully on sampled distribution $(\mathfrak{G}, \mathfrak{L})$.

The results of simulation studies are summarized as follows:

- Proposed input- and the output-filters jointly generate discrete subset of unknown fractal attractors.
- Sampled patterns is well structured to support origin – destination associations with respect to not-yet-identified mapping set.
- Structural consistency of sampled pattern with mapping descriptions can be evaluated by invariance test.

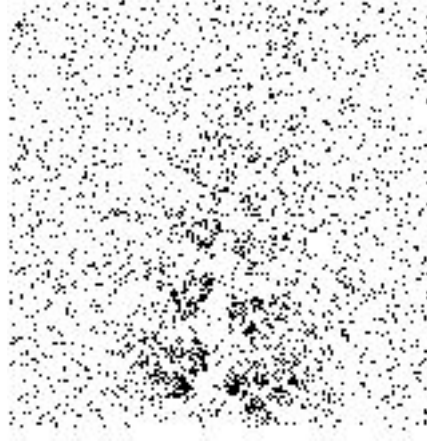


Fig. 6. Noisy Observation of Leaves

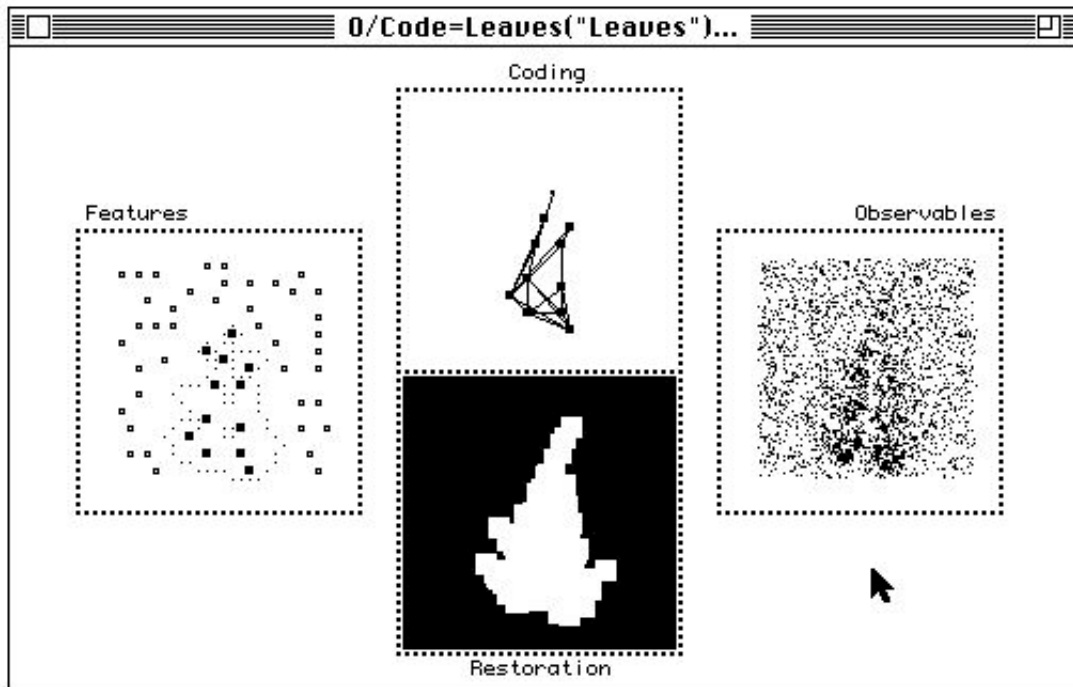


Fig. 7. Uniformly Observable Fractal Model

8 Concluding Remarks

A method was presented for structure sensitive sampling of unknown self-similarity in noisy imagery. By counting locally connected distribution on 2D



Fig. 8. Noisy Observation of Leaves

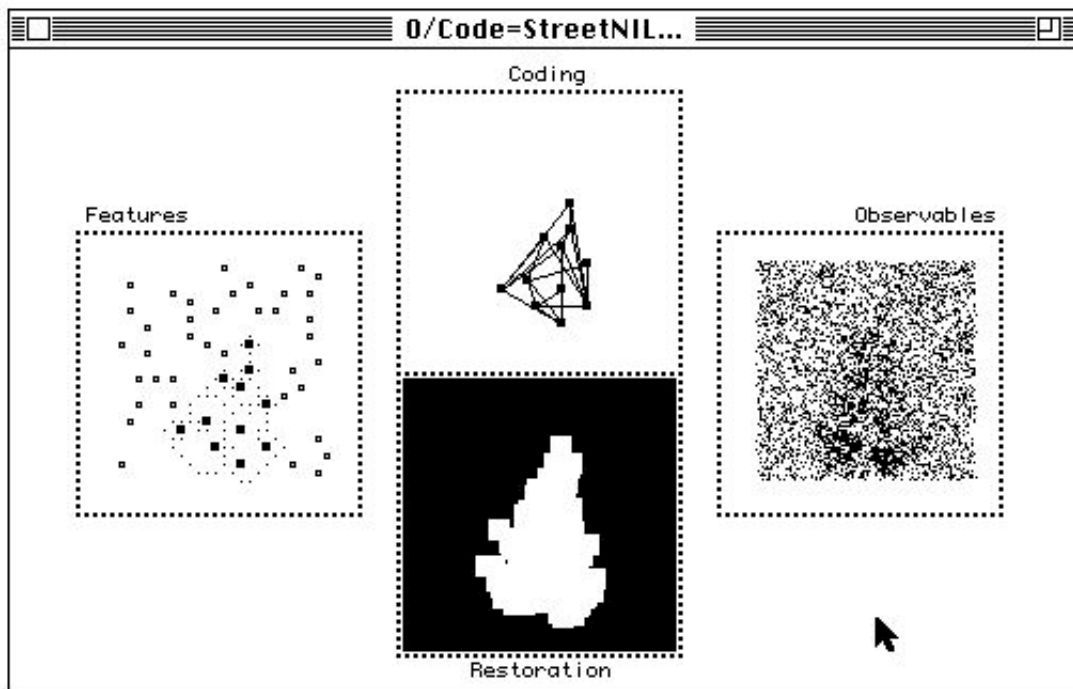


Fig. 9. Uniformly Observable Fractal Model

Gaussian array, the capturing probability for self-similar region is evaluated to generate discrete feature patterns. The capturing probability is sensitive to self-similar structure so that generated discrete pattern specifies the totality

of most probable attractor points. As the main result, a generic model is established for unknown self-similar patterns in background noise. Through simulation studies, extracted discrete patterns have been verified to maintain sufficient information to regenerate observed attractors.

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