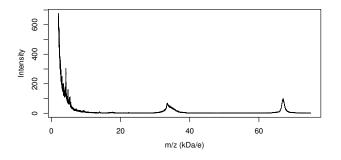
## Bayesian Adpative Regression Kernels

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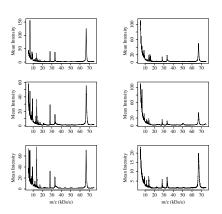
April 11, 2017

### MALDI-TOF Mass Spectroscopy



Location of peaks/proteins in spectra in the presence of noise Non-Gaussian Noise

### Multiple Spectra



Learning "features" that are common versus those that separate groups of spectra but do not know a prior the number of ions



# **Problem Setting**

Regression problem

$$E[Y \mid x] = f(x), \quad x \in \mathcal{X}$$

with unknown function  $f(\mathbf{x}): \mathcal{X} \to \mathbb{R}$ 

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$$f(\mathbf{x}_i) = \sum_{j=0}^{J} \psi(\mathbf{x}_i, \boldsymbol{\omega}_j) \beta_j$$

#### in terms of an (over-complete) dictionary where

- $\{\beta_j\}$ : unknown coefficients
- ▶ *J*: number of terms in expansion (finite or infinite)
- lacksquare  $\psi({m x}, {m \omega}_j)$  Dictionary elements from a "generator function" g
  - cubic splines

$$\psi(x_i,\omega_j)=(x_i-\omega_j)_+^3$$

$$\psi(\mathbf{x}_i, \omega_j) = g(\mathbf{\Lambda}_j(\mathbf{x} - \mathbf{\chi}_j)) = \exp\left\{-\frac{1}{2}(\mathbf{x} - \mathbf{\chi}_j)^T \mathbf{\Lambda}_j(\mathbf{x} - \mathbf{\chi}_j)\right\}$$

- translation and scaling wavelet families
- ▶ Need not be symmetric!



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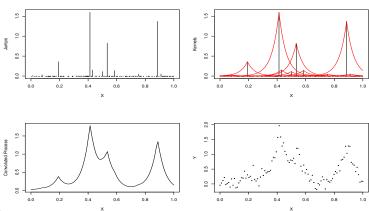
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#### Kernel Convolution



Easy to generate \*non-stationarity processes

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Introduce Lévy Measure  $\nu(\beta, \omega)$  (measure)

$$\Rightarrow J \sim P(\nu_{+}), \qquad \nu_{+} \equiv \nu(\mathbb{R} \times \Omega) = \iint \nu(\beta, \omega) d\beta d\omega$$
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- ▶ Possibly infinite number of  $\beta \in [-\epsilon, \epsilon]$
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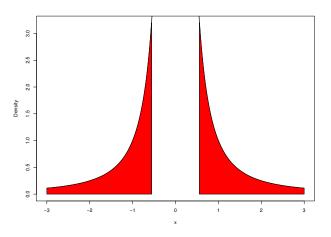
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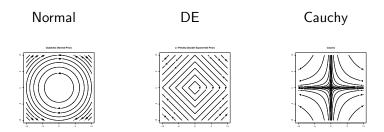
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# Truncated Cauchy Process

Restriction  $|\beta| > \epsilon$ 



# Contours of Log Prior (in $\mathbb{R}^2$ ) – Penalties



#### Penalized Likelihood:

$$-\frac{1}{2\sigma^2}\sum_i (Y_i - f(\mathbf{x}_i))^2 - (\alpha + 1)\sum_j \log(|\beta_j|) - \nu_{\epsilon}^+ \dots$$

- ▶ Birth: generate coefficients  $\beta_j$  near  $\epsilon$  in absolute value and generate kernel parameters  $\omega_j$  given increase in  $J \to J+1$
- ▶ Death:  $\beta_j = 0$  drop dictionary element when J decrements by 1.
- ▶ Update: Random-Walk update, but also leads to deaths with coefficients that wander out of bounds and cross ϵ boundary!
- Merge-Split: allow "neighboring points" to merge into one dictionary element (an alternative death) or a split into new dictionary elements

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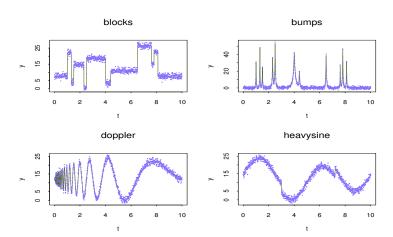
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### Computation - Reversible Jump MCMC

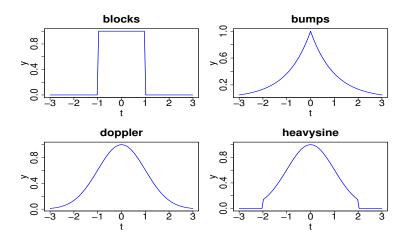
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- ▶ Death:  $\beta_j = 0$  drop dictionary element when J decrements by 1.
- Update: Random-Walk update, but also leads to deaths with coefficients that wander out of bounds and cross ε boundary!
- Merge-Split: allow "neighboring points" to merge into one dictionary element (an alternative death) or a split into new dictionary elements

Advantage over fixed dimensional over-complete methods (frames)

# Wavelet Test Functions (SNR = 7)



#### **Kernel Functions**



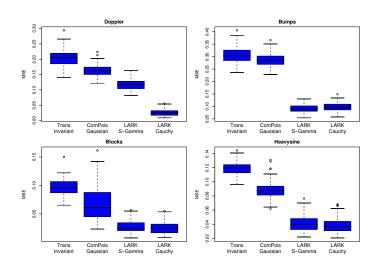
# Comparisons of OCD Methods

- Translational Invariant Wavelets Laplace Priors (Johnstone & Silverman 2005)
- Continuous Wavelet Dictionary Compound Poisson with Gaussian Priors (Chu, Clyde, Liang 2007)
- LARK Symmetric Gamma
- LARK Cauchy

Range of Over-complete Dictionaries and Priors

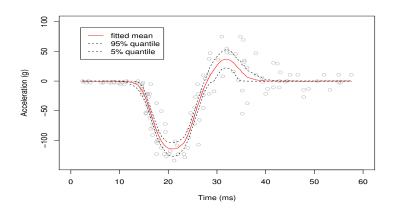
# Comparison of Mean Square Error w/ OCDs

#### 100 realizations of each function



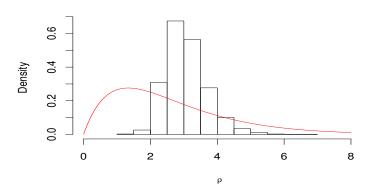
### Motorcycle Crash Data: A Real Example

On average, only  $E[J \mid Y] \approx 4$  jumps are needed for fit:



#### Form of Kernel

$$k(t_i; \tau_j, \lambda_j) = e^{-\lambda_j |t_i - \tau_j|^{\rho}}$$



# Higher Dimensional ${\mathcal X}$

#### RJ-MCMC is too slow in higher dimensional space to allow

- $ightharpoonup \chi$  to be completely arbitrary; restrict support to observed  $\{x_i\}$
- ightharpoonup use diagonal  $\Lambda$

Kernels take form:

$$\psi(\mathbf{x}, \omega_j) = \prod_{d} \exp\{-\frac{1}{2}\lambda_d(x_d - \chi_d)^2\}$$
$$f(\mathbf{x}) = \sum_{i} \psi(\mathbf{x}, \omega_j)\beta_j$$

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# Approximate Lévy Prior II

Continuous Approximation Student  $t(\alpha, 0, \epsilon)$  approximation:

$$\nu_{\epsilon}(d\beta, d\omega) = c_{\alpha}(\beta^2 + \alpha \epsilon^2)^{-(\alpha+1)/2} d\beta \ \gamma(d\omega)$$

Based on the following hierarchical prior

$$eta_j \mid \phi_j \stackrel{ind}{\sim} \operatorname{N}(0, \varphi_j^{-1})$$
 $\phi_j \stackrel{ind}{\sim} \operatorname{G}\left(\frac{\alpha}{2}, \frac{\alpha \epsilon^2}{2}\right)$ 
 $J \sim \operatorname{P}(\nu_{\epsilon}^+)$ 

where 
$$\nu_{\epsilon}^{+} = \nu_{\epsilon}(\mathbb{R}, \Omega) = \frac{\alpha^{1-\alpha/2} \Gamma(\alpha) \Gamma(\alpha/2)}{\epsilon^{\alpha} \pi^{1/2} \Gamma(\frac{\alpha+1}{2})} \sin(\frac{\pi \alpha}{2}) \gamma(\Omega)$$

Key: need to have variance of coefficients decrease as J increases

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- lacktriangle Cauchy process corresponds to lpha=1
- ▶ Tipping's "Relevance Vector Machine" corresponds to  $\alpha = 0$  (improper posterior!)
- ► Provides an extension of **Generalized Ridge Priors to** infinite dimensional
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- ▶ Poisson number of points  $J_{\epsilon} \sim P(\nu_{\epsilon}^{+}(\alpha, \gamma))$  with  $\nu_{\epsilon}^{+}(\alpha, \gamma) = \frac{\gamma \alpha^{1-\alpha/2}}{2^{1-\alpha}\epsilon^{\alpha}} \frac{\Gamma(\alpha/2)}{\Gamma(1-\alpha/2)}$
- ▶ Given J,  $[n_1:n_n] \sim MN(J,1/(n+1))$  points supported at each kernel located at  $x_i$

The regression mean function can be rewritten as

$$f(\mathbf{x}) = \sum_{i=0}^{n} \tilde{\beta}_{i} \psi(\mathbf{x}, \boldsymbol{\omega}_{i}), \quad \tilde{\beta}_{i} = \sum_{\{j \mid \boldsymbol{\chi}_{j} = \boldsymbol{x}_{i}\}} \beta_{j}.$$

In particular, if  $\alpha=1$ , not only the Cauchy process is infinitely divisible, the approximated Cauchy prior distributions on the regression coefficients are also infinitely divisible:

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- ▶  $n_i$  may be 0 which drops dictionary elements from representation in finite representation for fixed  $\epsilon$
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- Many input variables may be irrelevant
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## Regression Out of Sample Prediction

#### Average Relative MSE to best procedure

| BARK |  |   | C1/N/I   | BART   |
|------|--|---|--|--|
| D    | S + E  | S + D   | 20101  | DAILI  |
| 1.22 | 2.26   | 1.93  | 5.36   | 1.97   |
| 1.07 | 1.09   | 1.04  | 4.36   | 3.64   |
| 1.46 | 2.30   | 1.44  | 2.70   | 1.00   |
| 1.09 | 1.23   | 1.20  | 1.56   | 1.01   |
| 1.81 | 1.01   | 2.19  | 4.04   | 1.68   |
| 1.01 | 1.01   | 1.02  | 1.16   | 1.10   |
|      | 1.22<br>1.07<br>1.46<br>1.09<br>1.81<br>1.01 | D         S + E           1.22         2.26           1.07         1.09           1.46         2.30           1.09         1.23           1.81         1.01           1.01         1.01 | D         S + E         S + D           1.22         2.26         1.93           1.07         1.09         1.04           1.46         2.30         1.44           1.09         1.23         1.20           1.81         1.01         2.19           1.01         1.02 | D         S + E         S + D         SVM           1.22         2.26         1.93         5.36           1.07         1.09         1.04         4.36           1.46         2.30         1.44         2.70           1.09         1.23         1.20         1.56           1.81         1.01         2.19         4.04           1.01         1.02         1.16 |

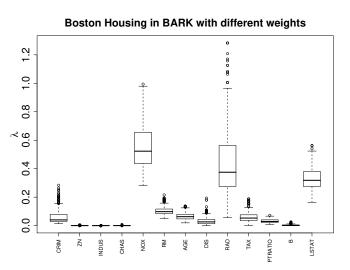
D: dimension specific scale  $\lambda_d$ 

E: equal scales  $\lambda_d = \lambda \, \forall \, d$ 

S: selection  $\lambda_d=0$  with probability  $\rho$ 

### Feature Selection in Boston Housing Data

Posterior Distribution of  $\lambda_d$ 



## Classification Examples

| Name             | d  | data type  | <pre>n (train/test)</pre> |  |
|------------------|----|------------|---------------------------|--|
| Circle           | 2  | simulation | 200/1000                  |  |
| Circle (3 null)  | 5  | simulation | 200/1000                  |  |
| Circle (18 null) | 20 | simulation | 200/1000                  |  |
| Swiss Bank Notes | 6  | real data  | 200 (5 <i>cv</i> )        |  |
| Breast Cancer    | 30 | real data  | 569 (5 <i>cv</i> )        |  |
| Ionosphere       | 33 | real data  | 351 (5 <i>cv</i> )        |  |

- ► Add latent Gaussian Z<sub>i</sub> for probit regression (as in Albert & Chib)
- Same model as before conditional on Z
- lacktriangle Advantage: Draw eta in a block from full conditional
- Can extend to Logistic

## Predictive Error Rate for Classification

| Data Sets  | BARK  |       |        | SVM     | BART   |
|------------|-------|-------|--------|---------|--------|
| Data Sets  | D     | S + E | S + D  | 2 4 141 | DAILI  |
| Circle 2   | 4.91% | 1.88% | 1.93%  | 5.03%   | 3.97%  |
| Circle 5   | 4.70% | 1.47% | 1.65%  | 10.99%  | 6.51%  |
| Circle 20  | 4.84% | 2.09% | 3.69%  | 44.10%  | 15.10% |
| Bank       | 1.25% | 0.55% | 0.88%  | 1.12%   | 0.50%  |
| BC         | 4.02% | 2.49% | 6.09%  | 2.70%   | 3.36%  |
| Ionosphere | 8.59% | 5.78% | 10.87% | 5.17%   | 7.34%  |

D: dimension specific scale  $\lambda_d$ 

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- ► NP Bayes of many flavors often does better than frequentist methods (BARK, BART, Treed GP, more)
- ► Hyper-parameter specification theory & computational approximation
- need faster code for BARK that is easier for users (BART & TGP are great!) (library(bark) or github
- Can these models be added to JAGS, STAN, etc instead of stand-alone R packages
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- Adaptive bandwidth for kernel regression
- Allow flexible generating functions
- ▶ Provide sparser representations compared to SVM & RVM, with coherent Bayesian interpretation
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